

# Returns to Education in the Russian Federation: Application of a machine learning instrumental variable technique for policy development in priority regions

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## Data and Code

Thanks are due to the Higher School of  
Economics, Moscow for making the Russian  
Longitudinal Monitoring Study (RLMS)  
Household data readily available for  
researchers around the world. The code  
used for this paper is made freely available  
for all researchers at [https://bitbucket.  
org/zagamog/edreru/src/master/](https://bitbucket.org/zagamog/edreru/src/master/)

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## KEYWORDS

keyword 1, keyword 2, keyword 3, keyword 4, keyword 5,  
keyword 6, keyword 7

## 1 | MOTIVATION FOR THIS PAPER

### 1.1 | Stylized fact 1: Returns trend and looking at cohorts

### 1.2 | Stylized fact 2: Priority regions and policy options for education

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**First level:**

$$\text{LogWage}_{ij} = b_{0j} + b_{1j} \cdot \text{Educ} + b_{2j} \cdot \text{Exp} + b_{3j} \cdot \text{Exp}^2 + b_{4j} \cdot \text{Gender} + \epsilon_{ij} \quad (1)$$

**Second Level:**

$$b_{0j} = \gamma_{00} + \gamma_{0n} \cdot Z + u_{00}; \quad b_{1j} = \gamma_{10} + \gamma_{1n} \cdot Z + u_{10}; \quad b_{ij} = \gamma_{i0} \quad \text{for } i \neq 0 \quad (2)$$

where an individual  $i$  is nested withing a region  $j$ ,  $\text{LogWage}$  is a logarithm of monthly wage,  $\text{Educ}$  stands for highest attained level of education,  $\text{Exp}$  and  $\text{Exp}^2$  reflect the years of working experience and its quadratic term respectively,  $\text{Gender}$  is a dummy variable for gender,  $Z$  is an  $n \times i$  matrix of regional characteristics,  $\epsilon$  and  $u_{00}$ ,  $u_{10}$  are the first- and second-level errors respectively.

We refer to `card_070_1999` and `belloni052.2011`.

## 2 | LITERATURE REVIEW: USE OF INSTRUMENTAL VARIABLES IN RETURNS TO EDUCATION STUDIES

General Literature

Russia Specific

The usage of instrumental variables as a way of correcting for the endogeneity of educational attainments in the estimation of returns to schooling in Russia is relatively rare. Nevertheless, there are several papers that leveraged IV technique and made a number of important conclusions.

One research ascertained that during the transition returns to education in Russia were not improving and remained among the most deficient in the world (`cheidvasser_educated_2007`). The researchers instrumented years of schooling, employing a policy experiment in Russia from the 1950s to 1960s, and corrected for a selectivity bias by adding an equation for the labor market participation. It was highlighted that the excess of well-educated workers seemed to be the main underpinning factor of wage differentials in Russia after Soviet Union dissolution. Additionally, the study showed that heterogeneity in rates of returns to education in Russia also hails from gender differences similar to the global patterns: women receive greater returns to higher education than men.

Utilizing the same instruments as `cheidvasser_educated_2007`, `akhmedjonov_higher_2011` evaluated the magnitude of the education premiums in Russia between 2000 and 2002. Both OLS and 2SLS estimates (8% and 19.1% respectively) were indicative of significant returns and the formation of a more flexible wage structure in the Russian Federation.

`arabsheibani_returns_2012` introduced the age of sexual intercourse as another instrumental variable to tackle endogeneity in schooling. In line with some past research, the scholars show that OLS undervalues returns to education compared to the IV estimates. The study adopts an instrumental variable quantile approach over the wage distribution in addition to the conditional mean estimation.

**kyui\_expansion\_2016**, using the amount available slots as an instrument, showed that returns to schooling in Russia declined for those who took advantage of higher education expansion in a post-communist Russia (1990-2005) in comparison to youths who obtained university degree in preceding periods. In an earlier study, **kyui\_returns\_2010** used the accessibility of tertiary education and the education levels of other household members as instruments in a wage equation (2010). The scholar demonstrated that a growth in returns to education in the Russian Federation does not indicate its closeness to the developed countries (**kyui\_returns\_2010**).

**belskaya\_college\_2014** evaluated a large-scale college expansion in Russia after the breakdown of the Soviet Union (2014). Using the number of campuses in the municipality of residence at age 17 as an instrument, the research contended that as the number of university campuses grew, individuals with low returns to schooling grew as well. But for a marginal person, who switched into a treatment group as a result of new campuses opening, the total gains from attending a college are considerable and positive. Furthermore, the scholars found that students with higher returns are attracted more intensively by new campuses opened in constrained municipalities (small non-capital cities or those lacking higher education institutions before college expansion) in comparison to the unconstrained ones.

### 3 | LASSO AND POST-LASSO INSTRUMENTAL VARIABLES: OVERVIEW OF ESTIMATION METHOD

This study uses post-Lasso IV estimation, which was elaborated by **belloni\_high\_2011** and **belloni\_sparse\_2012** a decade ago. The method performs an optimal IV selection, using Lasso technique.

The Lasso is a regression analysis method, popularized by Tibshirani **tibshirani\_regression\_1996**, that does both variable selection and shrinkage (unlike Ridge, which only shrinks). The Lasso solves a regression problem with L1 penalization of finding (**diebold\_econometric\_2019**):

$$\hat{\beta}_{LASSO} = \operatorname{argmin}_{\beta} \left( \sum_{i=1}^N \left( y_i - \sum_{i=1}^K \beta_i x_{it} \right)^2 + \lambda \sum_{i=1}^K |\beta_i| \right) \quad (3)$$

or equivalently:

$$\hat{\beta}_{LASSO} = \operatorname{argmin}_{\beta} \sum_{i=1}^N \left( y_i - \sum_{i=1}^K \beta_i x_{it} \right)^2 \quad (4)$$

s.t.

$$\sum_{i=1}^K |\beta_i| \leq c \quad (5)$$

The equations portray that the Lasso loss function is composed of the least squares estimator (or can be extended to many other estimator) and a penalty term in a form of the absolute coefficient value. The main distinctive property of this technique is the ability conducting feature selection procedure by nullifying regression coefficients of low importance, which is of particular relevance in situations with large feature sets. In case of zero lambda, the method gets back to OLS, while huge lambdas reduce coefficients to zero and may lead to under-fitting.

## 4 | EMPIRICAL STUDIES USING POST-LASSO IV METHODOLOGY

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## 5 | FINDINGS 1: TRADITIONAL INSTRUMENTAL VARIABLE TECHNIQUE

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## 6 | FINDINGS 2: POST LASSO IV RESULTS

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## 7 | CONCLUSIONS

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# Appendix

**TABLE A1** Results of Estimating Human Capital Depreciation for the Female sample, RLMS

	1994	1998	2003	2006	2012	2018
	(1)	(2)	(3)	(4)	(5)	(6)
Constant	9.725*** (0.381)	3.786*** (0.322)	5.464*** (0.301)	6.946*** (0.247)	8.133*** (0.186)	8.767*** (0.242)
Educ, years ( <i>S</i> )	0.122*** (0.025)	0.153*** (0.022)	0.158*** (0.020)	0.118*** (0.016)	0.087*** (0.012)	0.066*** (0.015)
Educ X Exper ( <i>T</i> <i>S</i> )	−0.002* (0.001)	−0.002*** (0.001)	−0.002** (0.001)	−0.0002 (0.001)	−0.0001 (0.0005)	0.0004 (0.001)
Exper ( <i>T</i> )	0.074*** (0.019)	0.080*** (0.016)	0.055*** (0.015)	0.013 (0.013)	0.020** (0.010)	0.020* (0.011)
Exper squared ( <i>T</i> <sup>2</sup> )	−0.001*** (0.0002)	−0.001*** (0.0002)	−0.001*** (0.0002)	−0.0003** (0.0001)	−0.0005*** (0.0001)	−0.001*** (0.0001)
Observations	1,645	1,667	2,093	2,630	4,057	3,312
R <sup>2</sup>	0.051	0.089	0.110	0.139	0.104	0.092
Adjusted R <sup>2</sup>	0.049	0.087	0.108	0.138	0.103	0.091
Residual Std. Error	0.853	0.728	0.731	0.664	0.641	0.597
F Statistic	22.179***	40.520***	64.342***	106.385***	117.366***	83.993***

Note:

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01

**TABLE A2**   Results of Estimating Human Capital Depreciation for the Male sample, RLMS

	1994	1998	2003	2006	2012	2018
	(1)	(2)	(3)	(4)	(5)	(6)
Constant	10.357*** (0.433)	5.029*** (0.360)	7.334*** (0.282)	8.067*** (0.243)	8.771*** (0.157)	9.094*** (0.185)
Educ, years ( <i>S</i> )	0.136*** (0.028)	0.123*** (0.024)	0.080*** (0.019)	0.077*** (0.016)	0.077*** (0.010)	0.077*** (0.012)
Educ X Exper ( <i>TS</i> )	−0.002* (0.001)	−0.001 (0.001)	0.0004 (0.001)	−0.0003 (0.001)	−0.0004 (0.0005)	−0.001 (0.001)
Exper ( <i>T</i> )	0.054** (0.023)	0.032* (0.017)	0.002 (0.014)	0.007 (0.013)	0.035*** (0.009)	0.037*** (0.010)
Exper squared ( <i>T</i> <sup>2</sup> )	−0.001*** (0.0003)	−0.0004** (0.0002)	−0.0003* (0.0002)	−0.0003* (0.0001)	−0.001*** (0.0001)	−0.001*** (0.0001)
Observations	1,392	1,433	1,763	2,170	3,360	2,800
R <sup>2</sup>	0.057	0.070	0.078	0.074	0.153	0.110
Adjusted R <sup>2</sup>	0.054	0.067	0.076	0.072	0.152	0.108
Residual Std. Error	0.951	0.803	0.754	0.688	0.598	0.570
F Statistic	20.989***	26.879***	37.362***	43.281***	151.868***	86.125***

Note:
 \*p<0.1; \*\*p<0.05; \*\*\*p<0.01



**TABLE A3** Results of Multilevel Modeling with Coverage by Vocational Education, Rosstat 2018

	Null model	Mincerian	Random Slope	Cross-Level Interaction
	(1)	(2)	(3)	(4)
Constant	10.178*** (0.034)	10.032*** (0.034)	10.056*** (0.036)	10.065*** (0.036)
Vocational		0.283*** (0.009)	0.279*** (0.021)	0.267*** (0.021)
Higher		0.638*** (0.009)	0.641*** (0.025)	0.622*** (0.025)
Coverage VE X Vocational				0.050** (0.025)
Coverage VE X Higher				0.083*** (0.030)
Experience		-0.026*** (0.002)	-0.027*** (0.002)	-0.027*** (0.002)
Experience squared		-0.065*** (0.002)	-0.065*** (0.002)	-0.065*** (0.002)
Females		-0.403*** (0.005)	-0.404*** (0.005)	-0.404*** (0.005)
Coverage VE			-0.101*** (0.039)	-0.142*** (0.043)
Variance of Intecept	0.09	0.08	0.09	0.09
Variance of Vocational			0.02	0.02
Variance of Higher			0.04	0.04
Residual Deviance	0.45	0.35	0.34	0.34
Observations	49,187	49,187	49,187	49,187
Log Likelihood	-59,755.060	-53,289.500	-53,094.620	-53,096.640
Akaike Inf. Crit.	119,516.100	106,595.000	106,217.200	106,225.300
Bayesian Inf. Crit.	119,542.500	106,665.400	106,340.500	106,366.100

Note:

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01