

Hidden Markov Models

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Biometrics Security and Privacy, Idiap

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- Discrete Markov Models
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Introduction

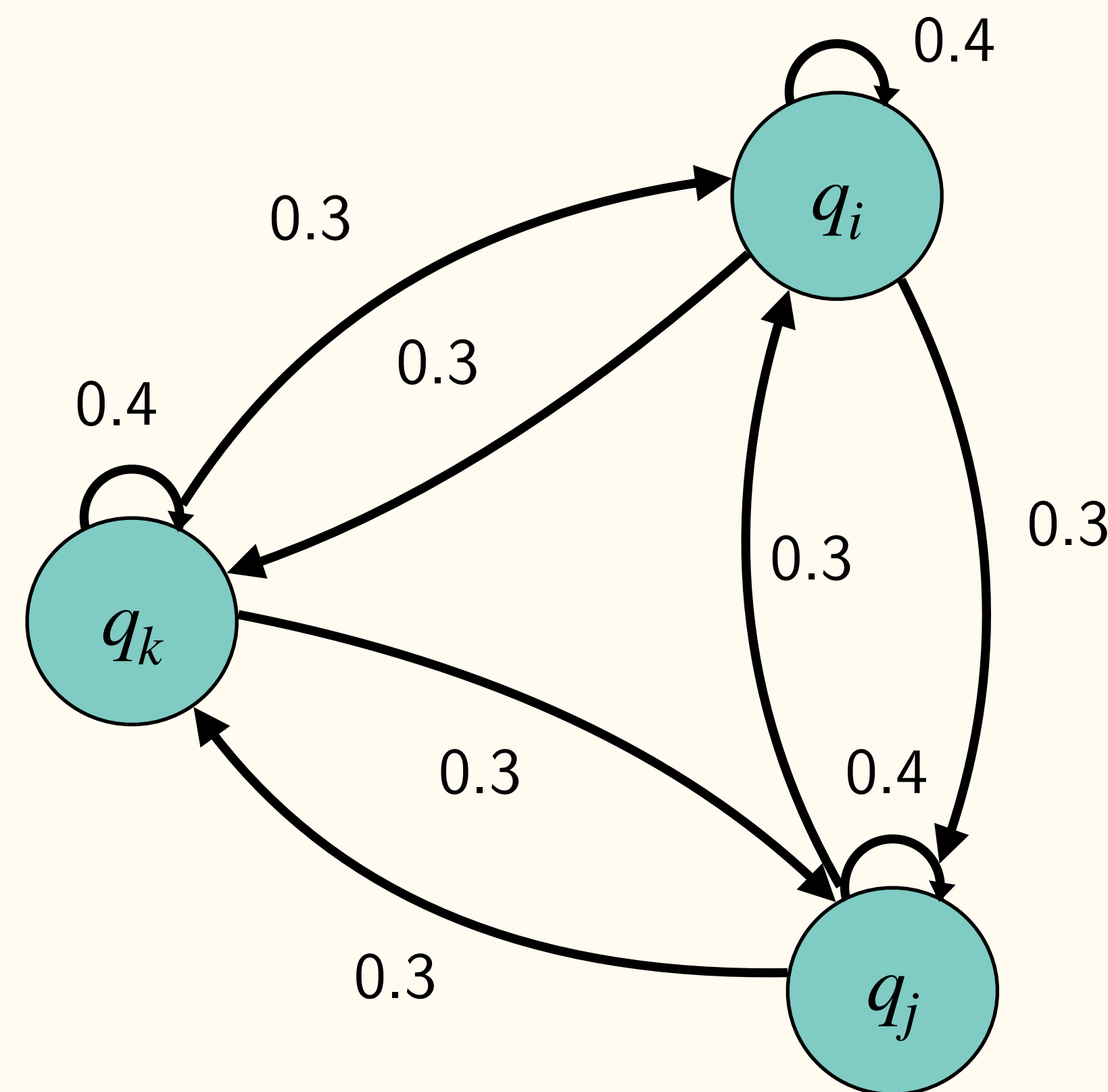
- Sequence processing:
 - Input: sequence X
 - Goal: estimate a sequence of outputs M
 - $P(M|X)$
- Tool: Hidden Markov Models (HMMs)
 - Introduced and studied in 1960-70s
 - Lawrence R. Rabiner. *A tutorial on Hidden Markov Models and selected applications in speech recognition.*



L. R. Rabiner

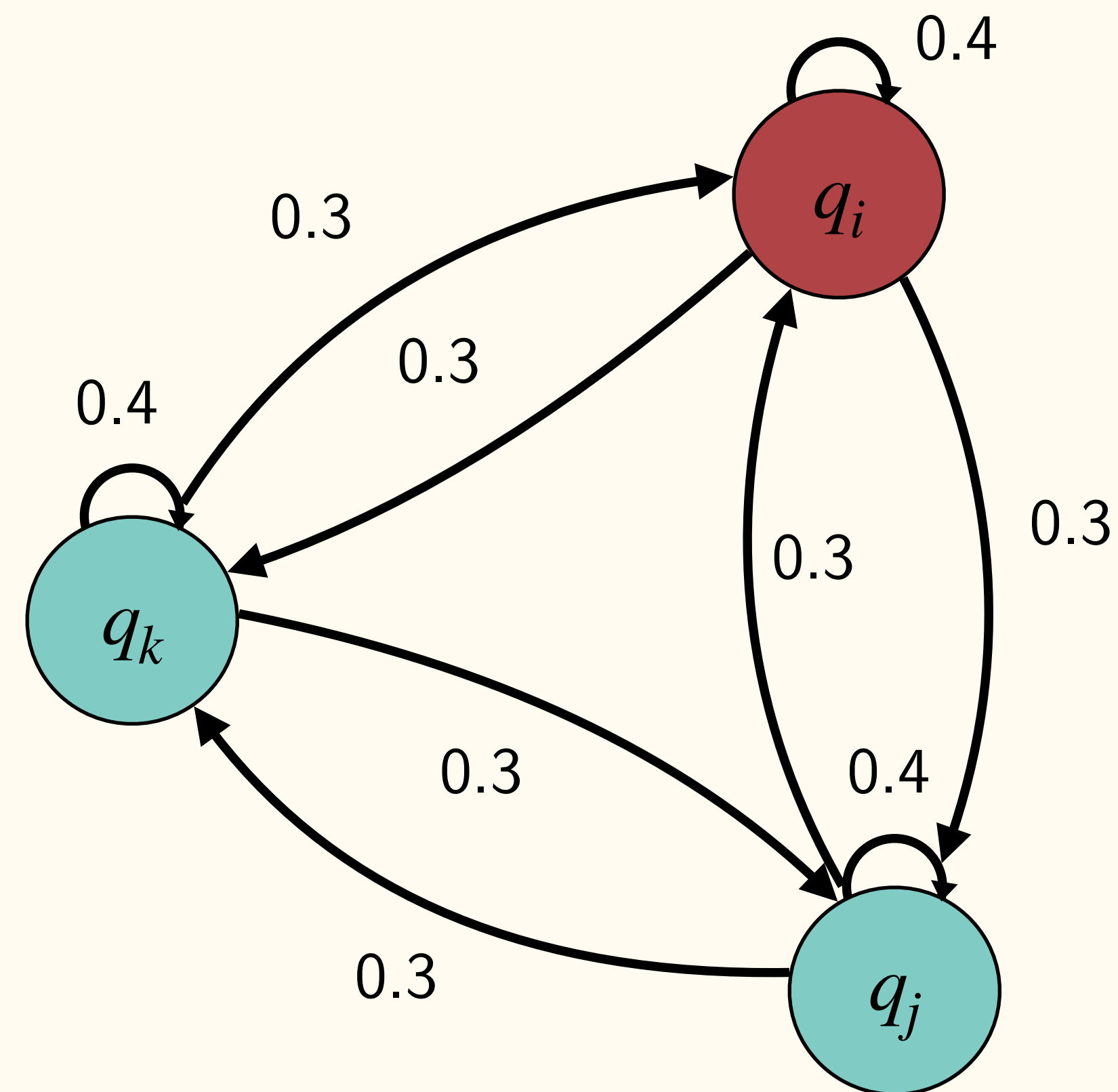
Discrete Markov Models (DMMs)

- Model M_k
- Composed of states $Q = \{q_1, \dots, q_k, \dots, q_K\}$
 - q_j^t denotes state q_j at time t
- Transition probabilities:
 - $A = \{a_{ij}\} = \frac{C(i \rightarrow j)}{\sum_k C(i \rightarrow k)}$
- First-order Markov Models
- Time independent
- $X = \{\}$



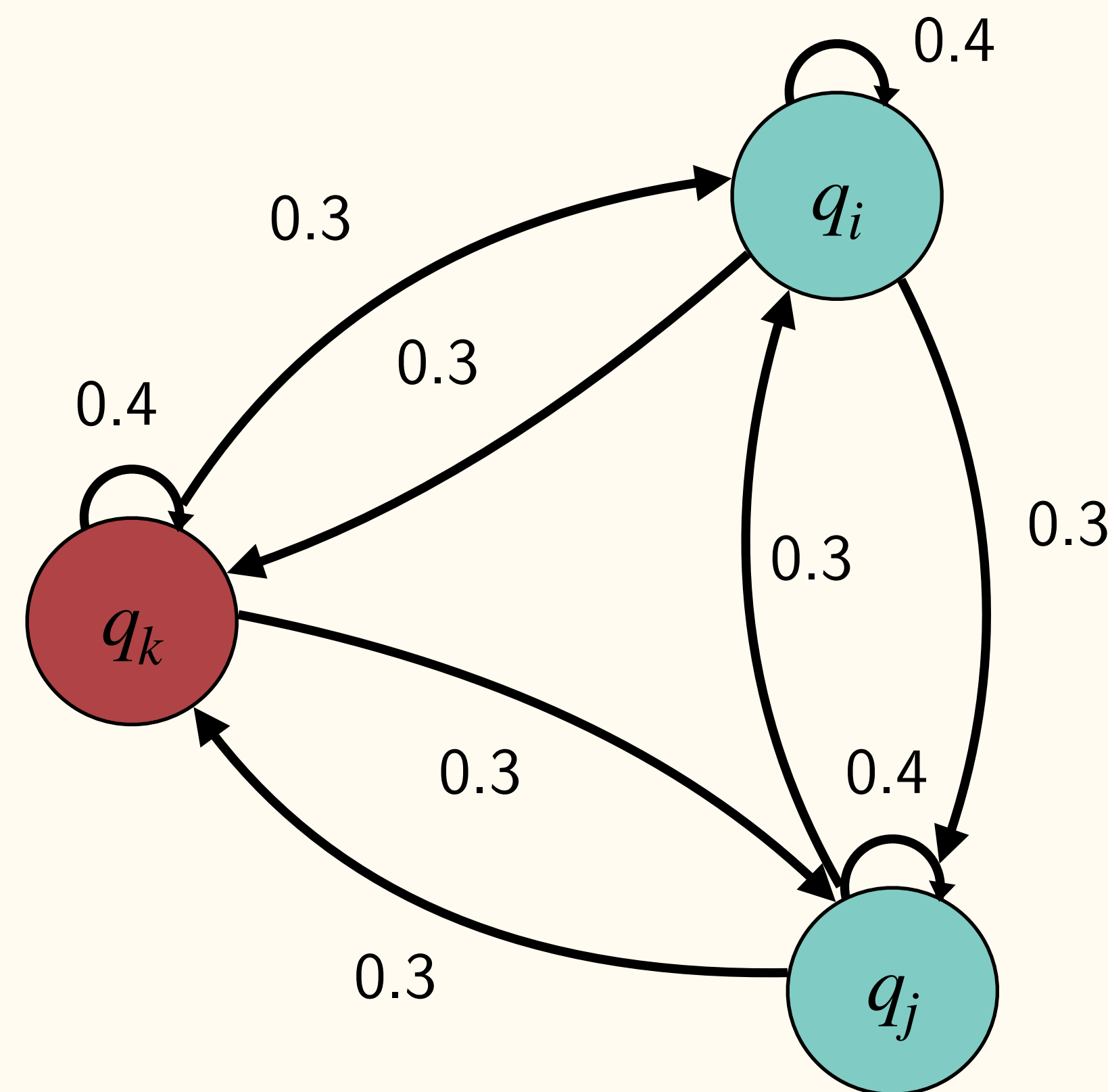
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- Time independent
- $X = \{q_i\}$



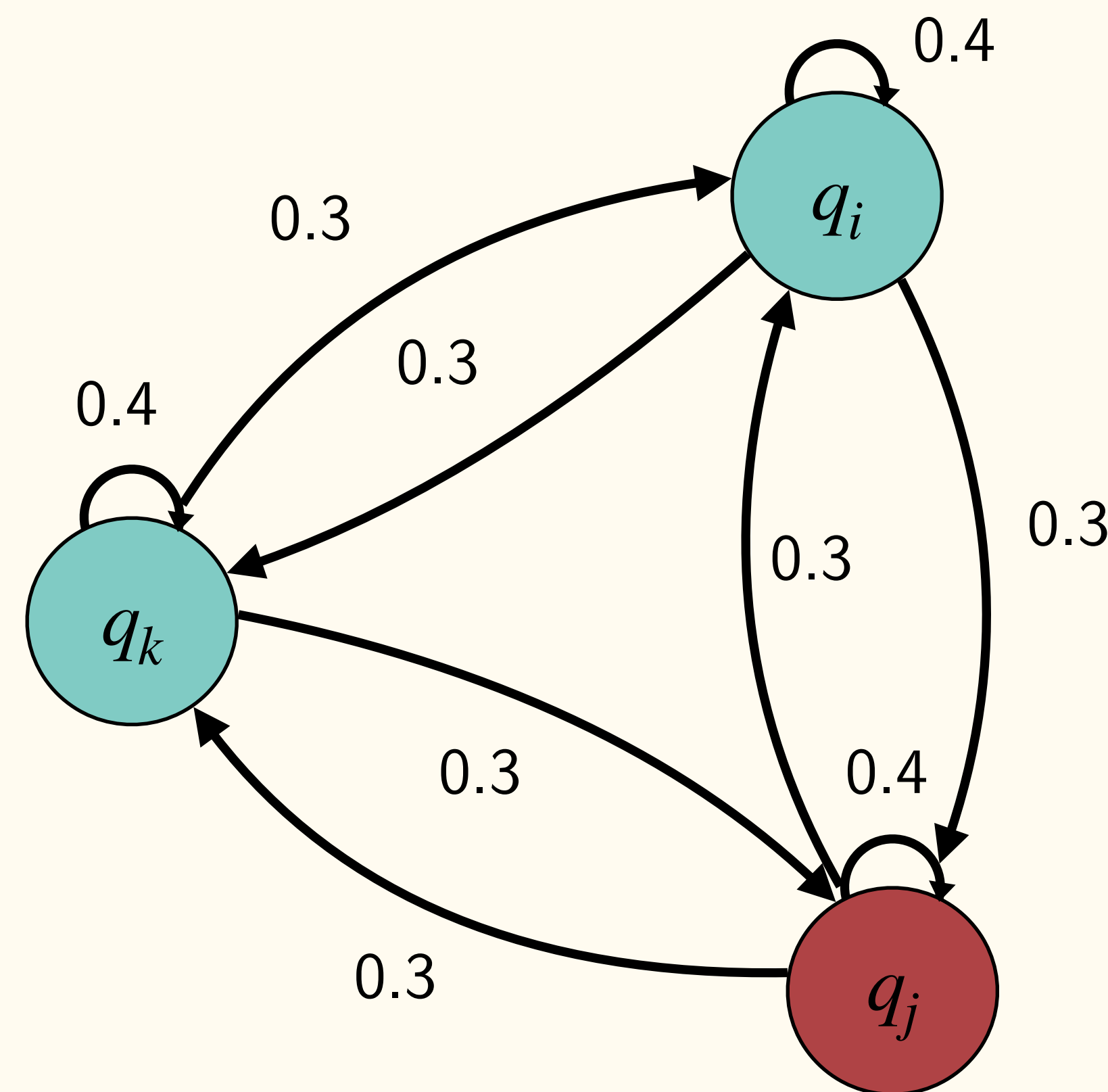
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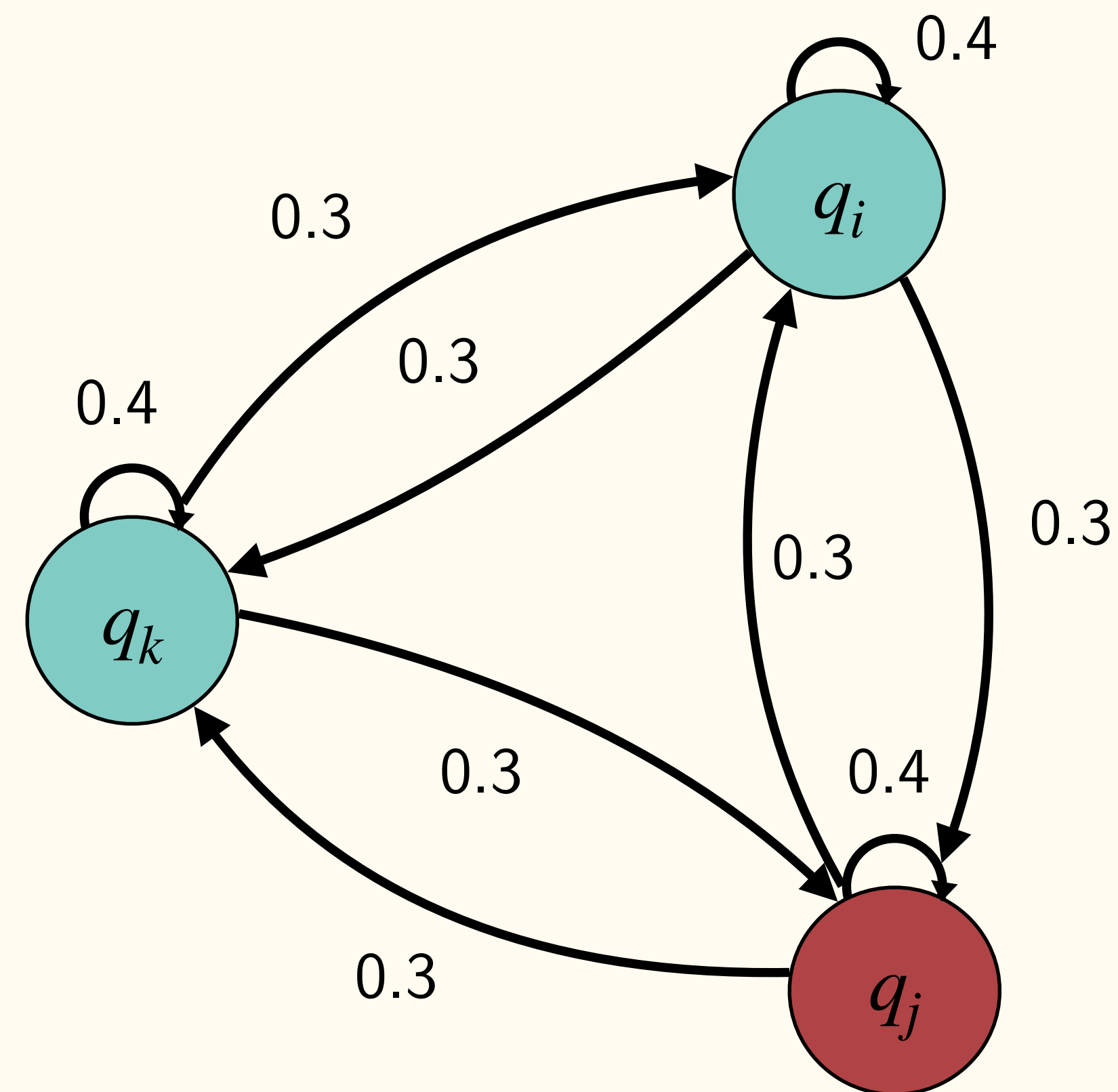
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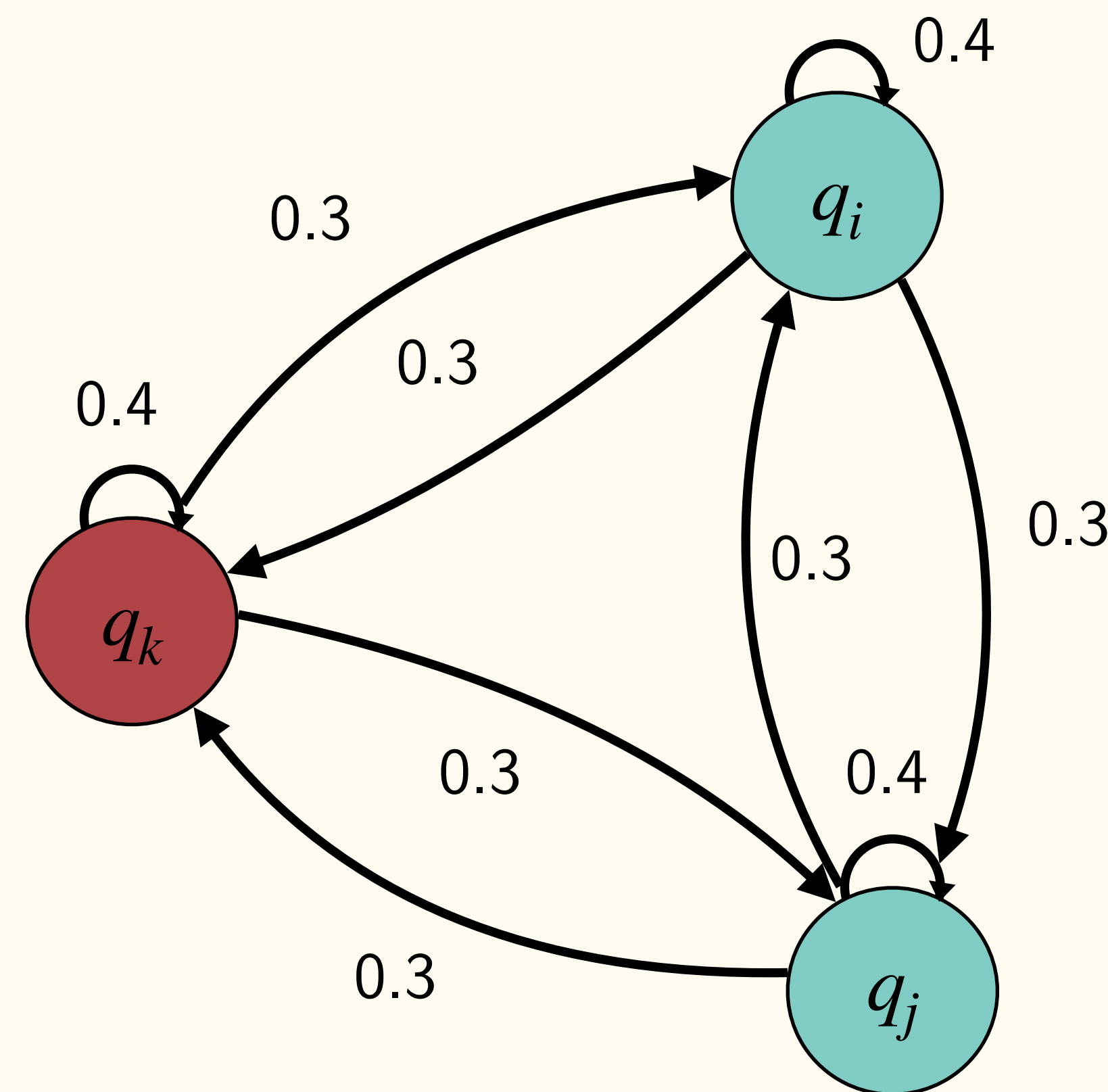
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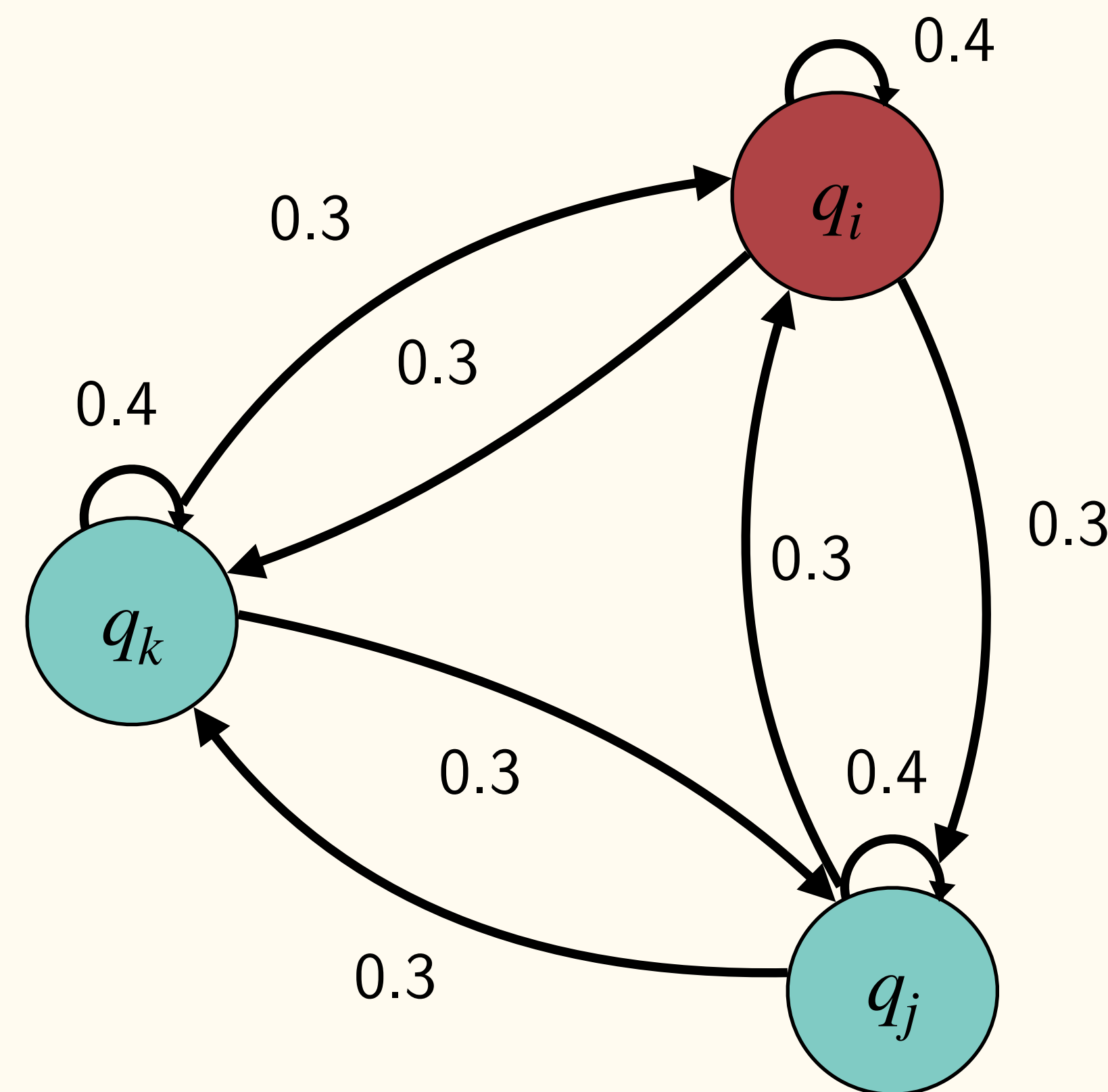
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DMMs - Sequence Probability

$$X = \{q_i, q_k, q_j, q_j, q_k, q_i\}$$

$$P(X|M)$$

$$= P(q_i, q_k, q_j, q_j, q_k, q_i | M)$$

$$= P(q_i | q_i, q_k, q_j, q_j, q_k) \cdot P(q_i, q_k, q_j, q_j, q_k)$$

$$= P(q_i | q_i, q_k, q_j, q_j, q_k) \cdot P(q_k | q_i, q_k, q_j, q_j) \cdot P(q_i, q_k, q_j, q_j)$$

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DMMs - Sequence Probability

$$X = \{q_i, q_k, q_j, q_j, q_k, q_i\}$$

$$P(X|M)$$

$$= P(q_i, q_k, q_j, q_j, q_k, q_i | M)$$

$$= P(q_i | q_i, q_k, q_j, q_j, q_k) \cdot P(q_i, q_k, q_j, q_j, q_k)$$

$$= P(q_i | q_i, q_k, q_j, q_j, q_k) \cdot P(q_k | q_i, q_k, q_j, q_j) \cdot P(q_i, q_k, q_j, q_j)$$

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DMMs - Sequence Probability

$$P(q_i | \cancel{q_i}, \cancel{q_k}, \cancel{q_j}, q_k) \cdot P(q_k | \cancel{q_i}, \cancel{q_k}, \cancel{q_j}, q_j) \cdot P(q_j | \cancel{q_i}, \cancel{q_k}, q_j) \cdot P(q_j | \cancel{q_i}, q_k) \cdot P(q_k | q_i) \cdot P(q_i)$$

↓ First-Order Markov Property $X = \{q_i, q_k, q_j, q_j, q_k, q_i\}$

$$\Rightarrow P(q_i | q_i) \cdot P(q_k | q_j) \cdot P(q_j | q_j) \cdot P(q_j | q_k) \cdot P(q_k | q_i) \cdot P(q_i) \quad \leftarrow \text{Only need transition probabilities}$$

$$= a_{ii} \cdot a_{kj} \cdot a_{jj} \cdot a_{jk} \cdot a_{ki} \cdot \pi_{q_i}$$

$$= 0.4 \cdot 0.3 \cdot 0.4 \cdot 0.3 \cdot 0.3 \cdot 1$$

$$= 0.00432$$

$$A = \{a_{ij}\} = \begin{bmatrix} 0.4 & 0.3 & 0.3 \\ 0.3 & 0.4 & 0.3 \\ 0.4 & 0.3 & 0.4 \end{bmatrix}$$

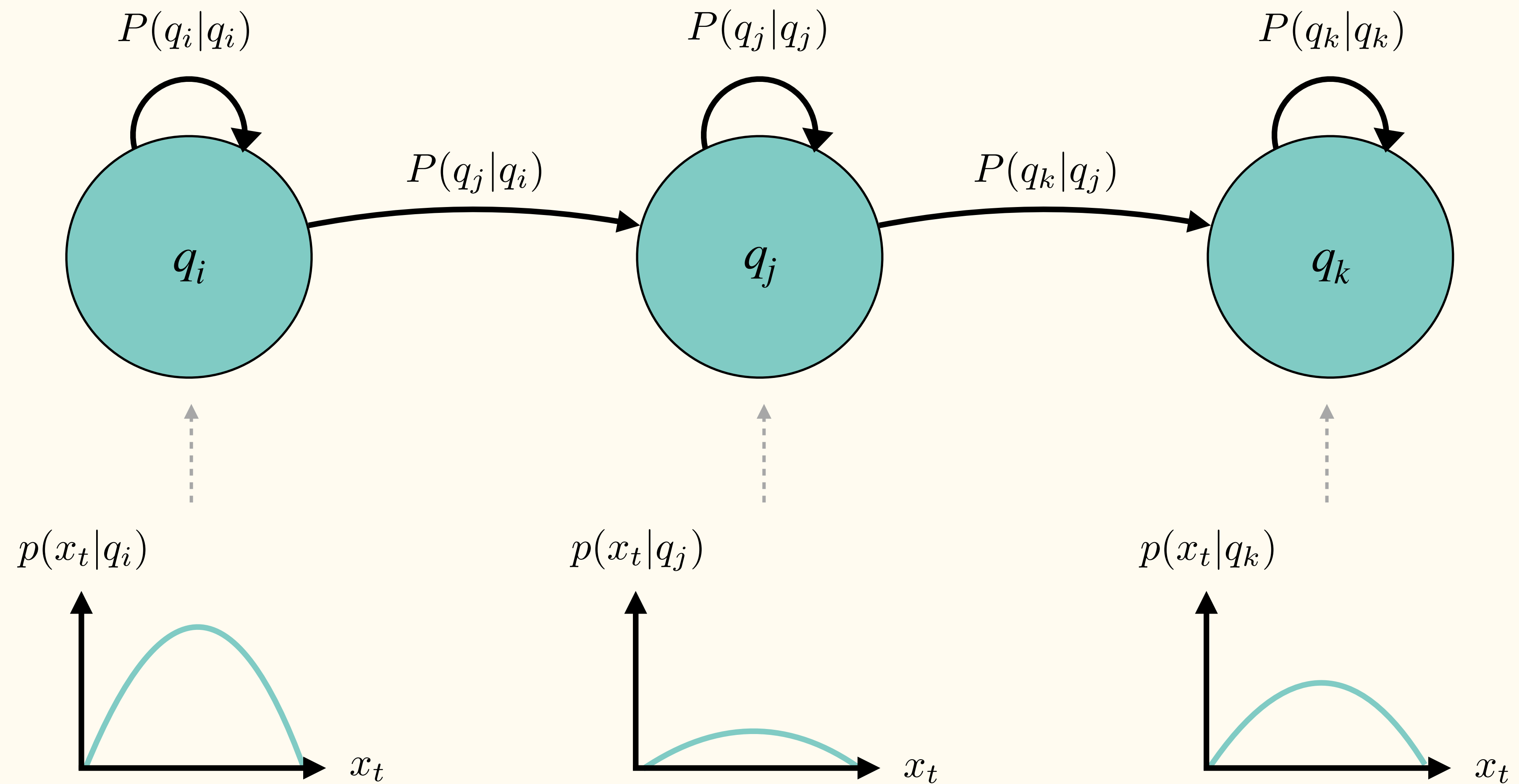
DMMs - Consecutive Sequence Probability

Given the model M is in a known state, what is the probability it stays in the **same state** for exactly d days ?

- $X = \{q_i^1, q_j^2, q_j^3, \dots, q_i^d, q_j^{d+1} \neq q_i\}$
- Discrete probability density function of duration d in state i :
 - $P(X|M, q^1 = q_i) = (a_{ii})^{d-1} \cdot (1 - a_{ii}) = p_i(d)$
- Expected number of observations (duration) in a state:

- $$\bar{d} = \sum_{d=1}^{\infty} d p_i(d) = \sum_{d=1}^{\infty} d \cdot (a_{ii})^{d-1} \cdot (1 - a_{ii}) = \frac{1}{1 - a_{ii}}$$

Hidden Markov Models (HMMs)



HMMs

- Sequence of observations: $X = \{x_1, \dots, x_t, \dots, x_T\}$
- Sequence of states: $Q = \{q_1, \dots, q_k, \dots, q_K\}$, q_j^t is state q_j at time t
- Transition probabilities: $A = \{a_{ij}\} : a_{ij} = P(q_j^{t+1} | q_i^t), \quad 1 \leq i, j \leq K$
- Emission probability distribution: $B = \{b_j(k)\} : b_j(k) = p(v_k^t | q_j^t), \quad 1 \leq j \leq K$
- Initial state distribution: $\pi = \{\pi_i\} : \pi_i = P(q_i^1), \quad 1 \leq i \leq K$

$$\Theta = \{\pi, A, B\}$$

- Observations now also described by emission probabilities, characterized by different stochastic distributions for each state q_i , $i \in [1, \dots, K]$.
 - Discrete, Gaussians, GMMs, ANNs (MLPs, or RNNs).

HMMs - Steps

1. Choose an initial state $q_1 = q_i$ according to initial state distribution π .
2. Set $t = 1$.
3. Choose $X_t = v_k$ according to **emission probability distribution** in state q_i i.e. $b_i(k)$.
4. Transit to a new state q_j^{t+1} according to state **transition probabilities** i.e. a_{ij} .
5. Set $t = t + 1$
 - If $t < T$:
 - **Return to step 3)**
 - Else:
 - **Terminate.**

HMM-based Pattern Classification

Bayes Theorem

$$P(M|X, \Theta) = \frac{p(X|M, \Theta) P(M|\Theta)}{p(X|\Theta)}$$

- M : Sequential (sentence) model
- Θ : Model Parameters
- $P(X, M|\Theta)$: HMM (acoustic model)
- $P(X|\Theta)$: Assumed constant
- $P(M|\Theta)$: Prior knowledge (language model). $P(M|\Theta) \Rightarrow P(M|\Theta^*)$

Three HMM Problems

1. Definition and estimation of transition a_{ij} and emission $b_i(x)$ probabilities:
 - Computing likelihood $P(X|M, \Theta)$ for a given M_k and fixed Θ

2. **Training** a HMM:

- Estimating Θ such that: $\operatorname{argmax}_{\Theta} \prod_{j=1}^J P(X_j|M_j, \Theta)$

3. Classification (decoding) of an observed sequence X :

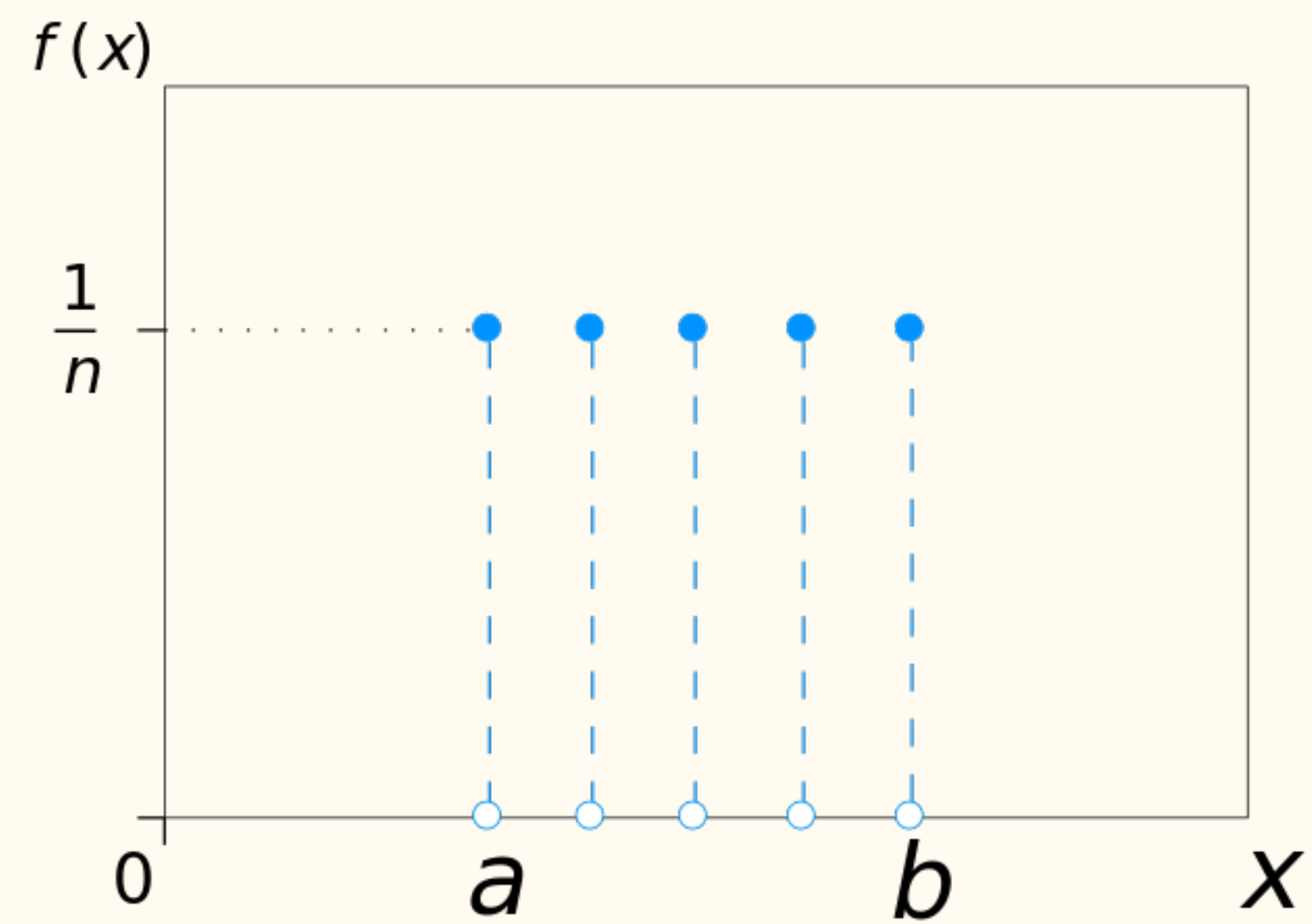
- $X \in M_j$ if $M_j = \operatorname{argmax}_{M_j} P(X|M_k, \Theta)P(M_k)$

Training Problem

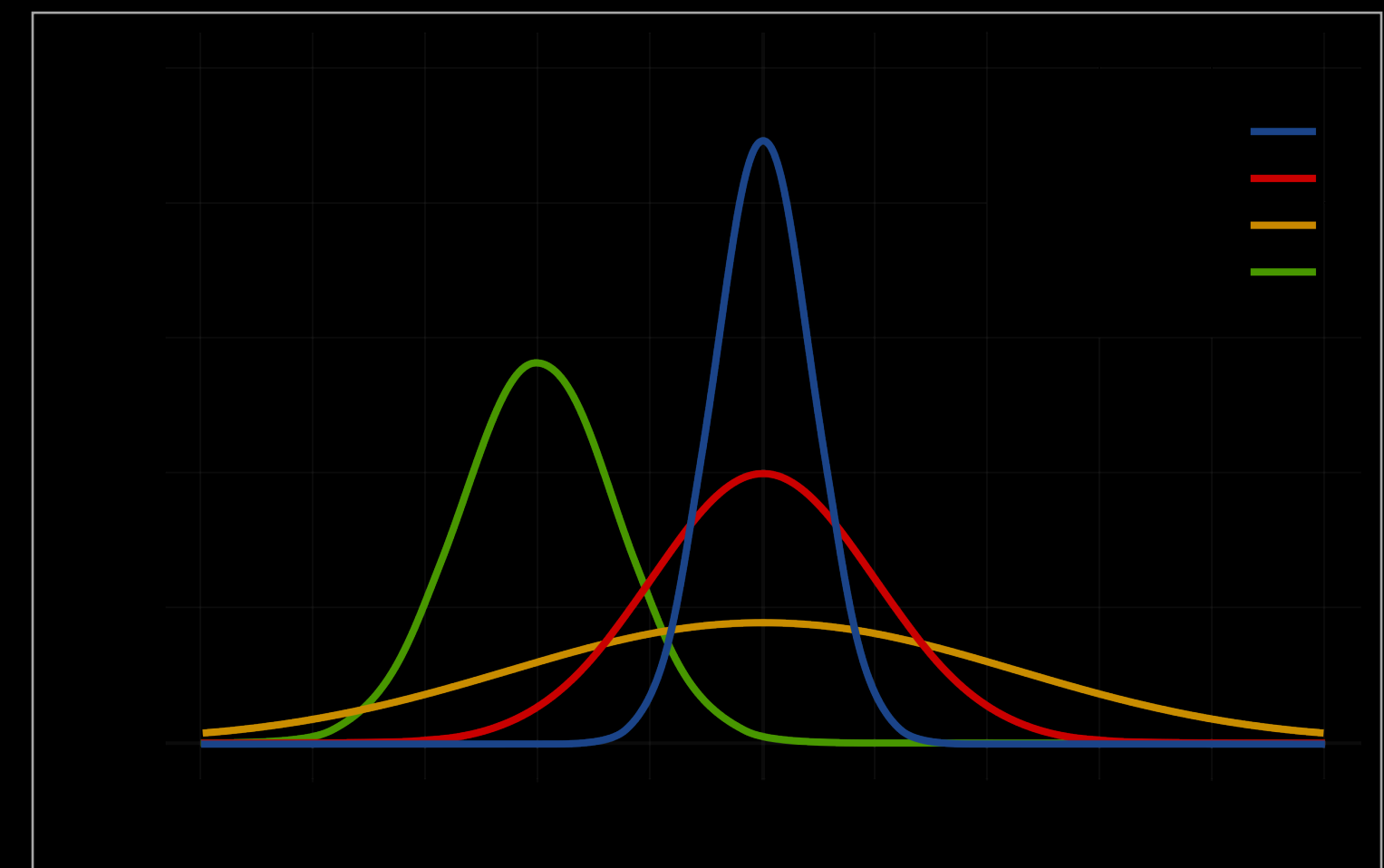
HMM Training Problem

- We want accurate parameters Θ from the observations sequence.
- States in HMM are hidden \rightarrow no closed-form equation for estimating parameters.
- We need to estimate the parameters $\Theta = \{\pi, A, B\}$ with a **maximum likelihood framework** on $p(X | \Theta)$.
 - Transition matrix A
 - Emission's underlying PDF B (discrete or continuous)

Discrete



Continuous



HMM Training Problem

- We estimate these parameters such that $\operatorname{argmax}_{\Theta} \prod_{j=1}^J P(X_j | M_j, \Theta)$
- We use the **Forward-Backward** algorithm
 - Iterative procedure of re-estimations
 - Efficient:
 - $\mathcal{O}(TK^T) \rightarrow \mathcal{O}(TK^2)$
 - Greatly reduces computation of the likelihood of a sequence given parameters.
 - Stores intermediate values that lead to a given state at a given time.
- Can also use *embedded* **Viterbi** approximation.

I. Forward-Backward Training

Forward-Backward Training

- Algorithms and variables:
 - Forward algorithm and variable $\alpha_t(i)$
 - Backward algorithm variable $\beta_t(i)$
 - Sequence of events $\xi_t(i, j)$
 - Gamma variable $\gamma_t(i)$

Forward Recurrence

We define the following variable:

- $\alpha_t(i) = p(x_1, \dots, x_t, q^t = q_i | \Theta)$

i.e. the probability of having observed the partial sequence $\{x_1, \dots, x_t\}$ and being at state i at time t , given the parameters Θ .

- Requires π, A, B
- Complexity: $\mathcal{O}(TK^2)$

1. Initialization:

- $\alpha_1(i) = \pi_i b_i(x_1), \quad 1 \leq i \leq K$

Join probability of state q_i and initial observation x_i .

2. Recursion:

- $\alpha_{t+1}(j) = \left[\sum_{i=1}^K \alpha_t(i) a_{ij} \right] b_j(x_{t+1})$

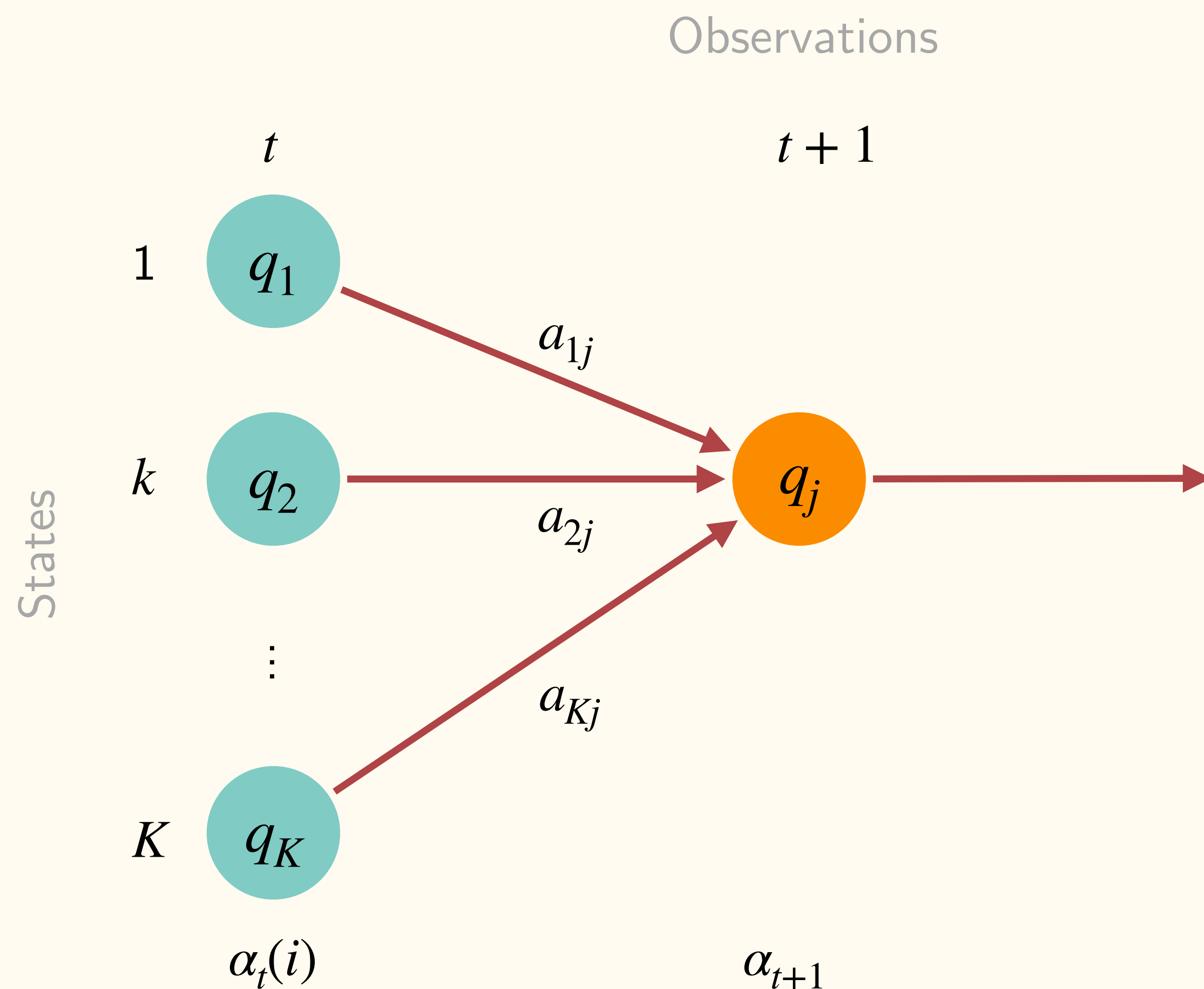
All possible ways to reach j · probability to generate x_{t+1}

3. Termination:

- $P(X | \Theta) = \sum_{i=1}^K \alpha_T(i)$

Sum over all possible states one could've ended up in.

Forward Algorithm - Recursion



Variable:

$$\underbrace{\alpha_t(i)} = p(x_1, \dots, x_t, q^t = q_i | \Theta)$$

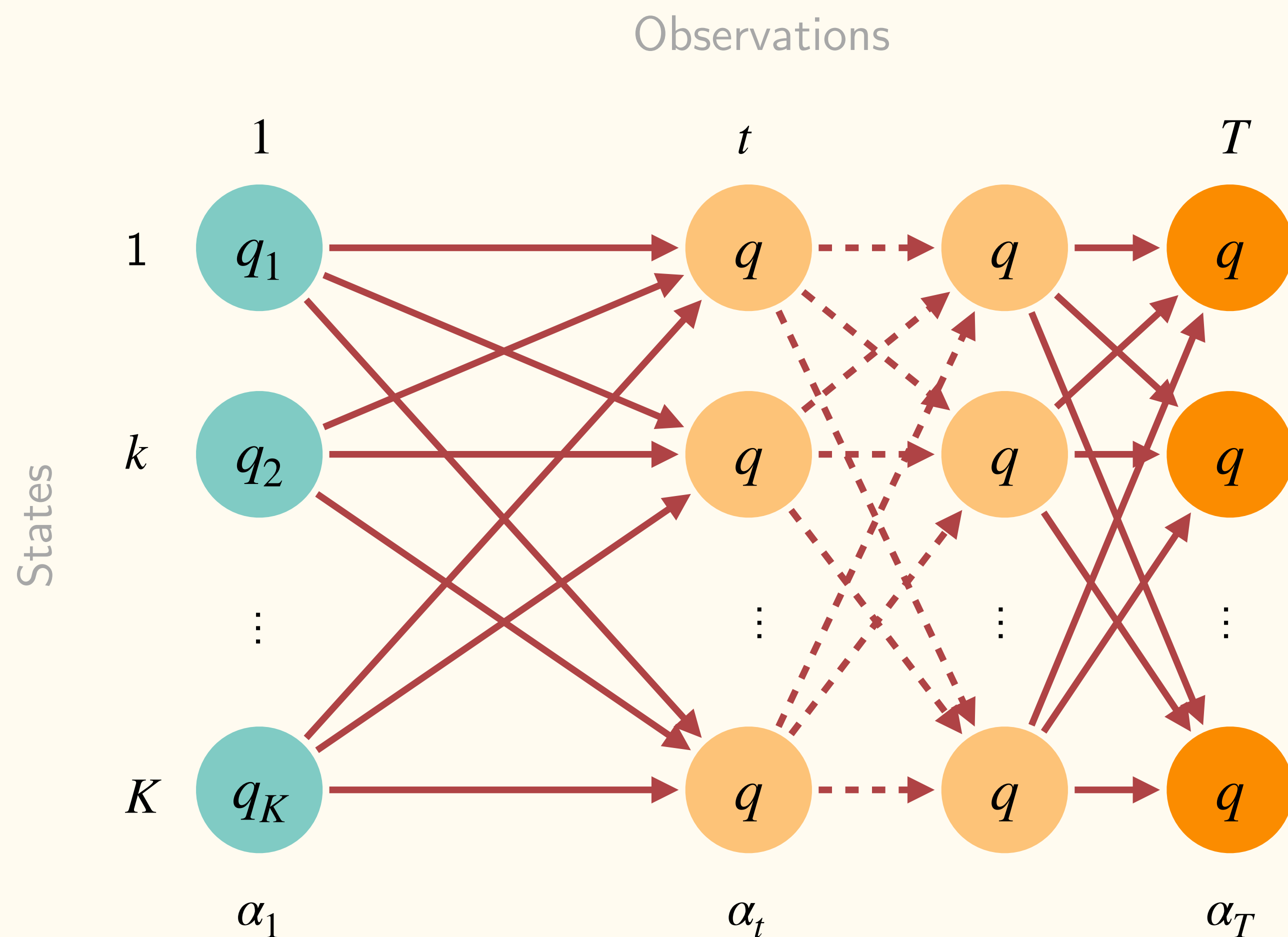
Probability of joint event that X is observed and the state at time t is q_i .

Recursion step: Generate observation

$$\alpha_{t+1}(j) = \left[\sum_{i=1}^K \underbrace{\alpha_t(i) a_{ij}} \right] \overbrace{b_j(x_{t+1})}$$

Probability of joint event that X is observed and q_j is reached at time $t+1$ via q_i at time t .

Forward Algorithm - Termination



Variable:

$$\underbrace{\alpha_t(i)} = p(x_1, \dots, x_t, q^t = q_i | \Theta)$$

Probability of joint event that X is observed and the state at time t is q_i .

Termination step:

$$P(X | \Theta) = \sum_{i=1}^K \alpha_T(i)$$

Backward Algorithm

We define the following variable:

- $\beta_t(i) = p(x_{t+1}, \dots, x_T | q^t = q_i, \Theta)$

i.e. the probability of having observed the partial sequence $\{x_{t+1}, \dots, x_T\}$, given the state i at time t and the parameters Θ .

- Requires π, A, B
- Complexity: $\mathcal{O}(TK^2)$

1. Initialization:

- $\beta_T(i) = 1$

Arbitrarily defined to be 1 for all i .

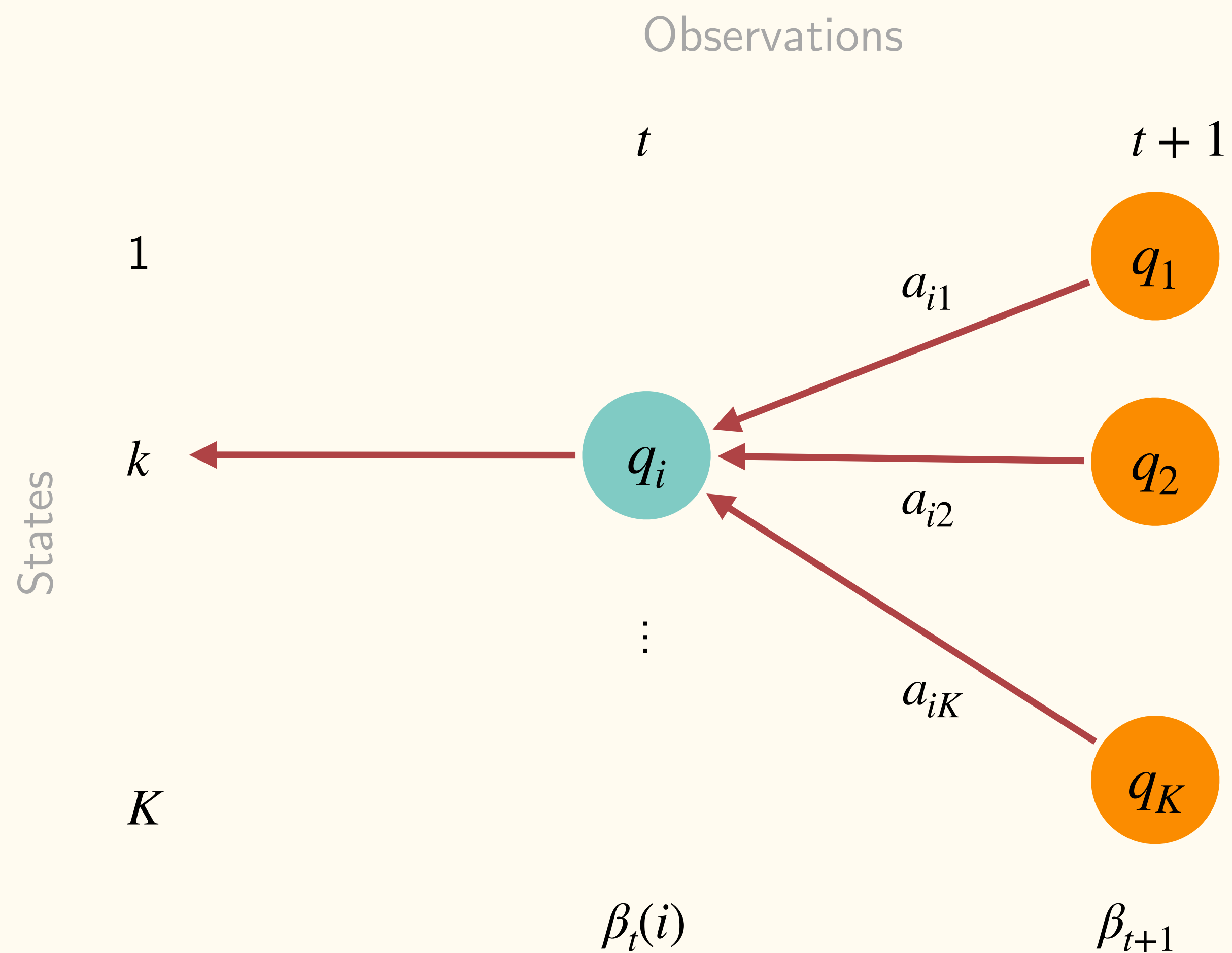
2. Recursion:

- $\beta_t(j) = \left[\sum_{i=1}^K \beta_{t+1}(i) a_{ij} \right] b_j(x_{t+1})$

3. Termination:

- $\beta_0 = P(X | \Theta) = \sum_{i=1}^K \pi_i b_i(x_1) \beta_1(i)$

Backward Algorithm - Recursion



Variable:

$$\underbrace{\beta_t(i)} = p(x_{t+1}, \dots, x_T | q^t = q_i, \Theta)$$

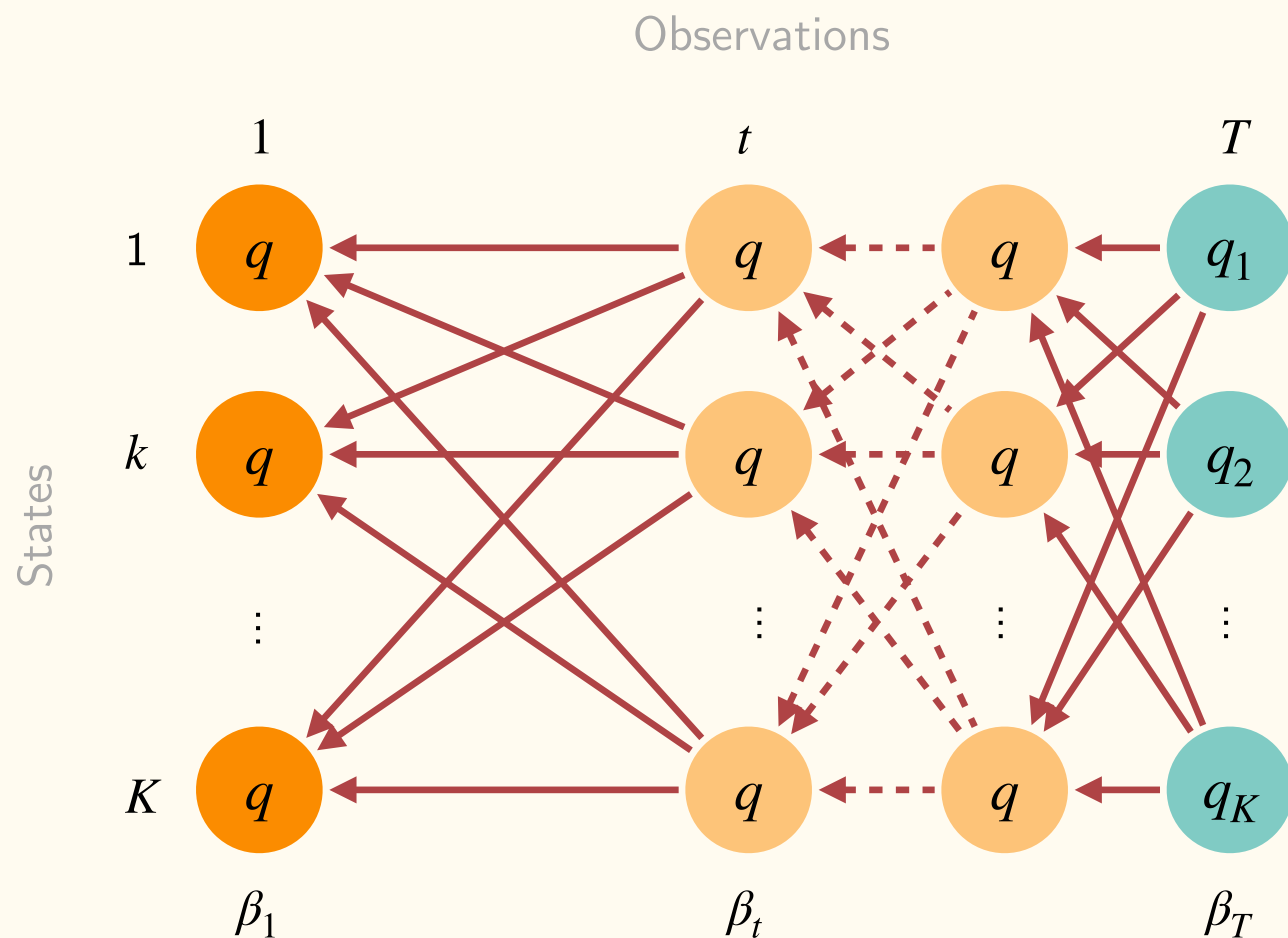
Probability that X is observed given the state q_i at time t and model parameters Θ .

Recursion step:

$$\beta_t(j) = \left[\sum_{i=1}^K \beta_{t+1}(i) a_{ij} \right] b_j(x_{t+1})$$

q_j can be reached at time $t+1$ from the K possible states.

Backward Algorithm - Termination



Variable:

$$\underbrace{\beta_t(i)}_{\text{Probability that } X \text{ is observed given the state } q_i \text{ at time } t \text{ and model parameters } \Theta.}$$

Probability that X is observed given the state q_i at time t and model parameters Θ .

Termination step:

$$\beta_0 = P(X | \Theta) = \sum_{i=1}^K \pi_i b_i(x_1) \beta_1(i)$$

Transition Probabilities Re-Estimation

- Forward and backward algorithms used to isolate states within HMM
- These variables let us estimate:
 - Transition probabilities between states
 - Emission probability distribution
- Start with re-estimation of A :
 - $\overline{a_{ij}} = \frac{\text{Expected number of times from state } q_i \text{ to } q_j}{\text{Expected number of transitions from } q_i}$
 - Need ξ

Sequence of Events

Forward Backward

We define the following variable:

- $\xi_t(i, j) = P(q^t = q_i, q^{t+1} = q_j | X, \Theta)$

i.e. the probability of being in state i at time t and in state j at time $t + 1$, given the observations and parameters Θ .

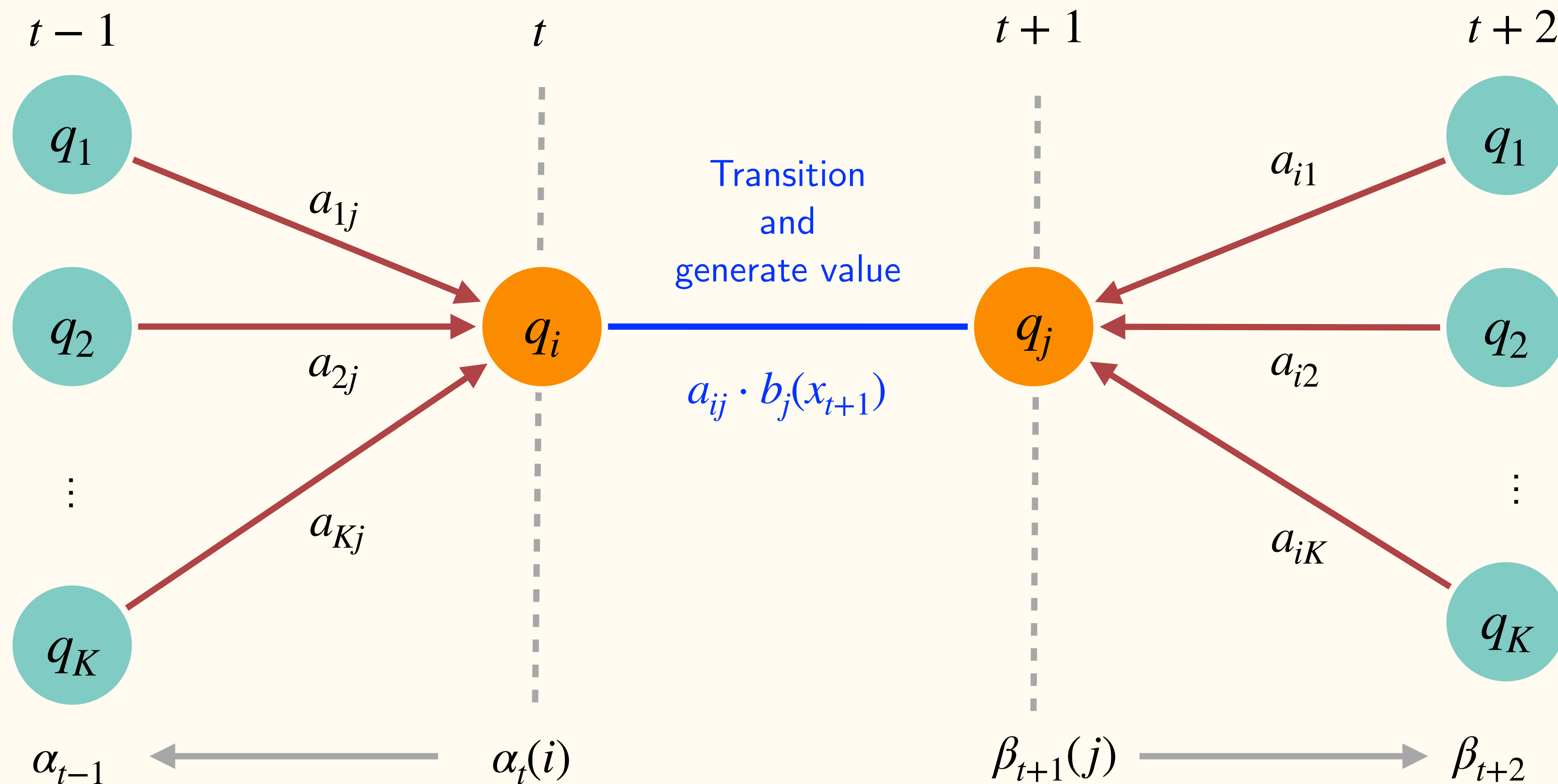
Can be expressed in terms of both forward and backward variables as:

$$\xi_t(i, j) = \frac{P(q_i^t, q_j^{t+1} | X, \Theta)}{P(X | \Theta)} \quad \} \text{ Normalization factor}$$

$$\begin{aligned} &= \frac{\alpha_t(i) a_{ij} b_j(x_{t+1}) \beta_{t+1}(j)}{\sum_{i=1}^K \alpha_t(i) \beta_t(i)} \\ &= \frac{\alpha_t(i) a_{ij} b_j(x_{t+1}) \beta_{t+1}(j)}{\sum_{i=1}^K \sum_{j=1}^K \alpha_t(i) a_{ij} b_j(x_{t+1}) \beta_{t+1}(j)} \end{aligned}$$

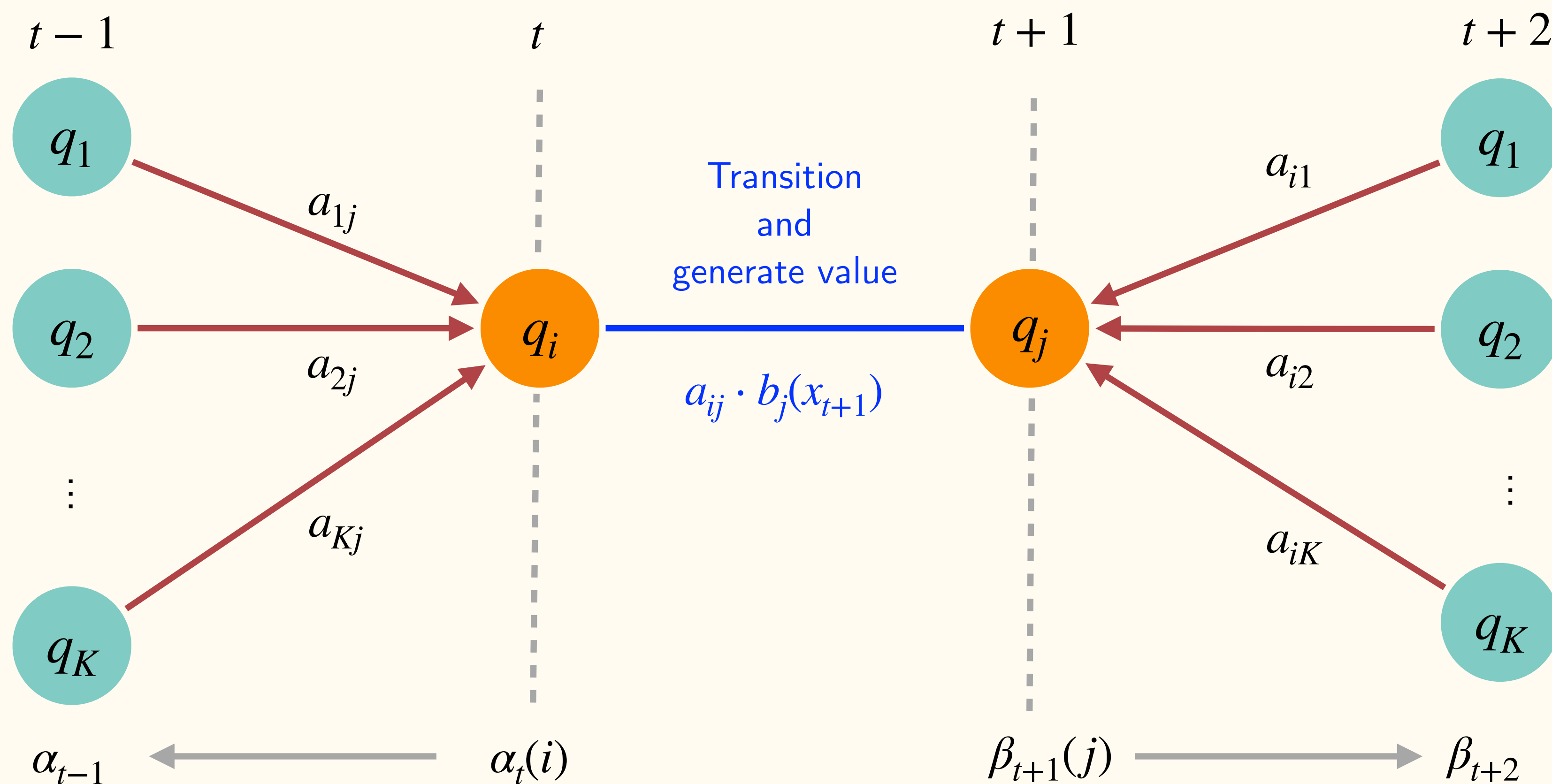
Sequence of Events

$$\xi_t(i, j) = P(q^t = q_i, q^{t+1} = q_j | X, \Theta)$$



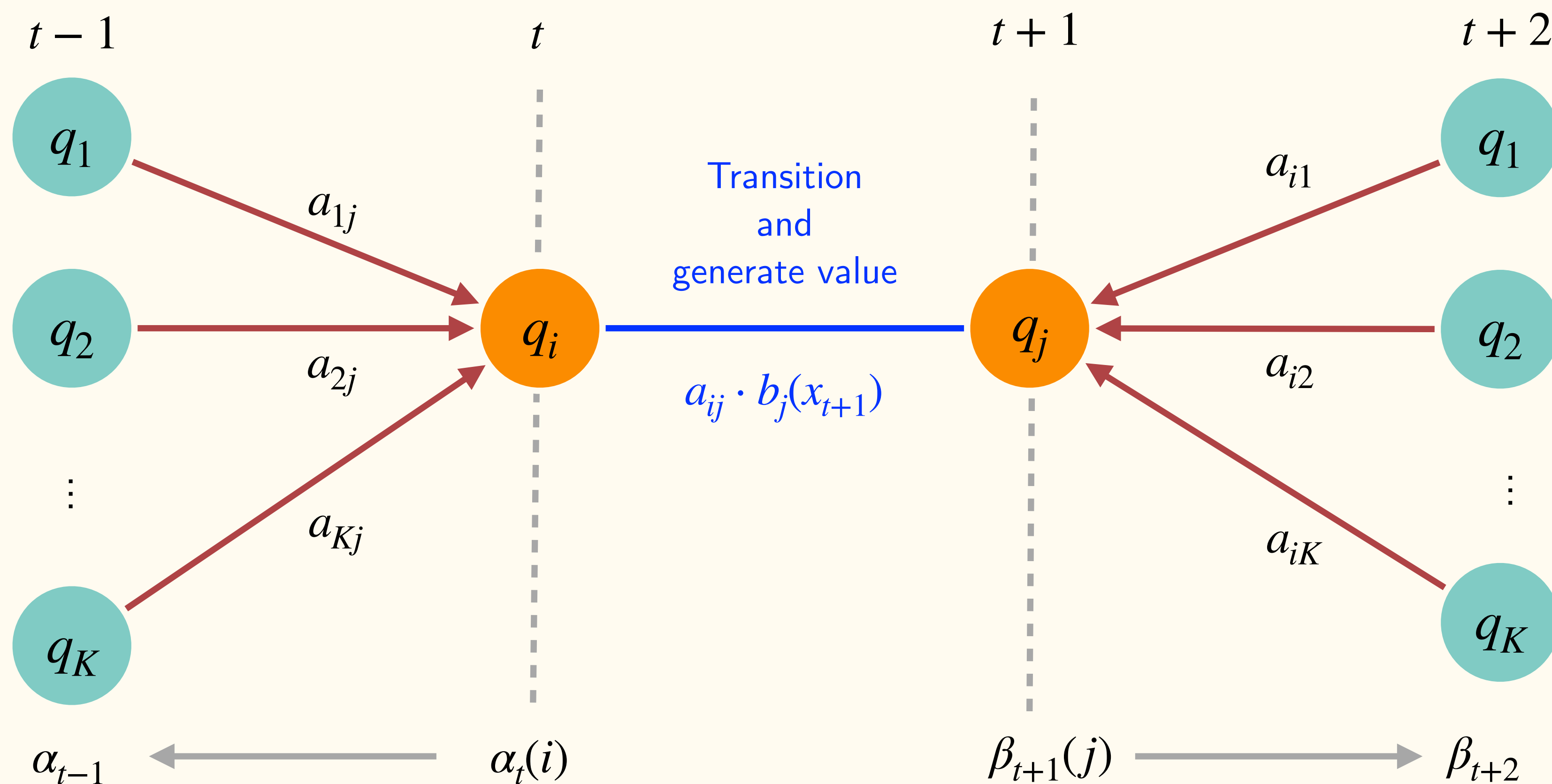
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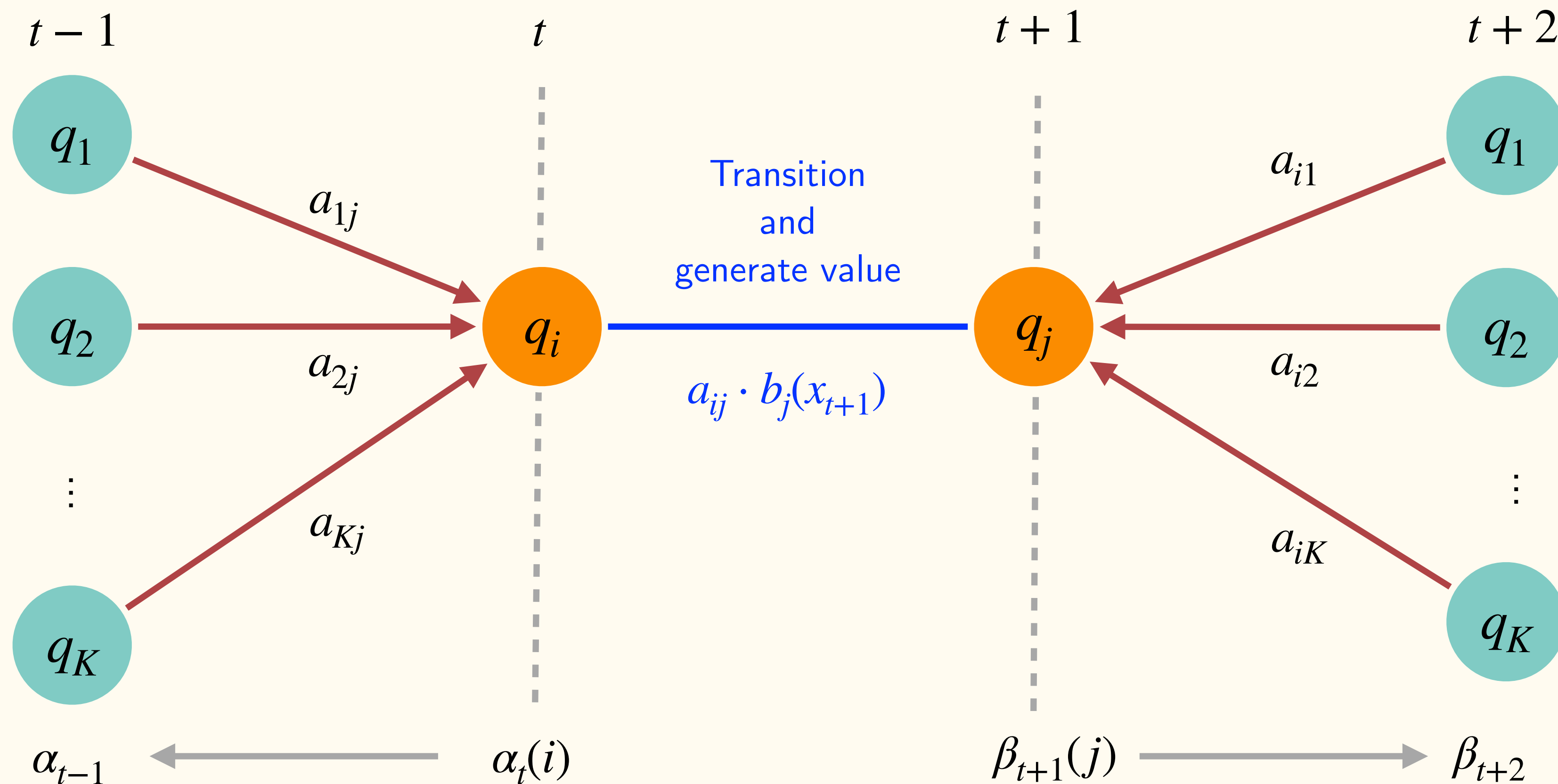
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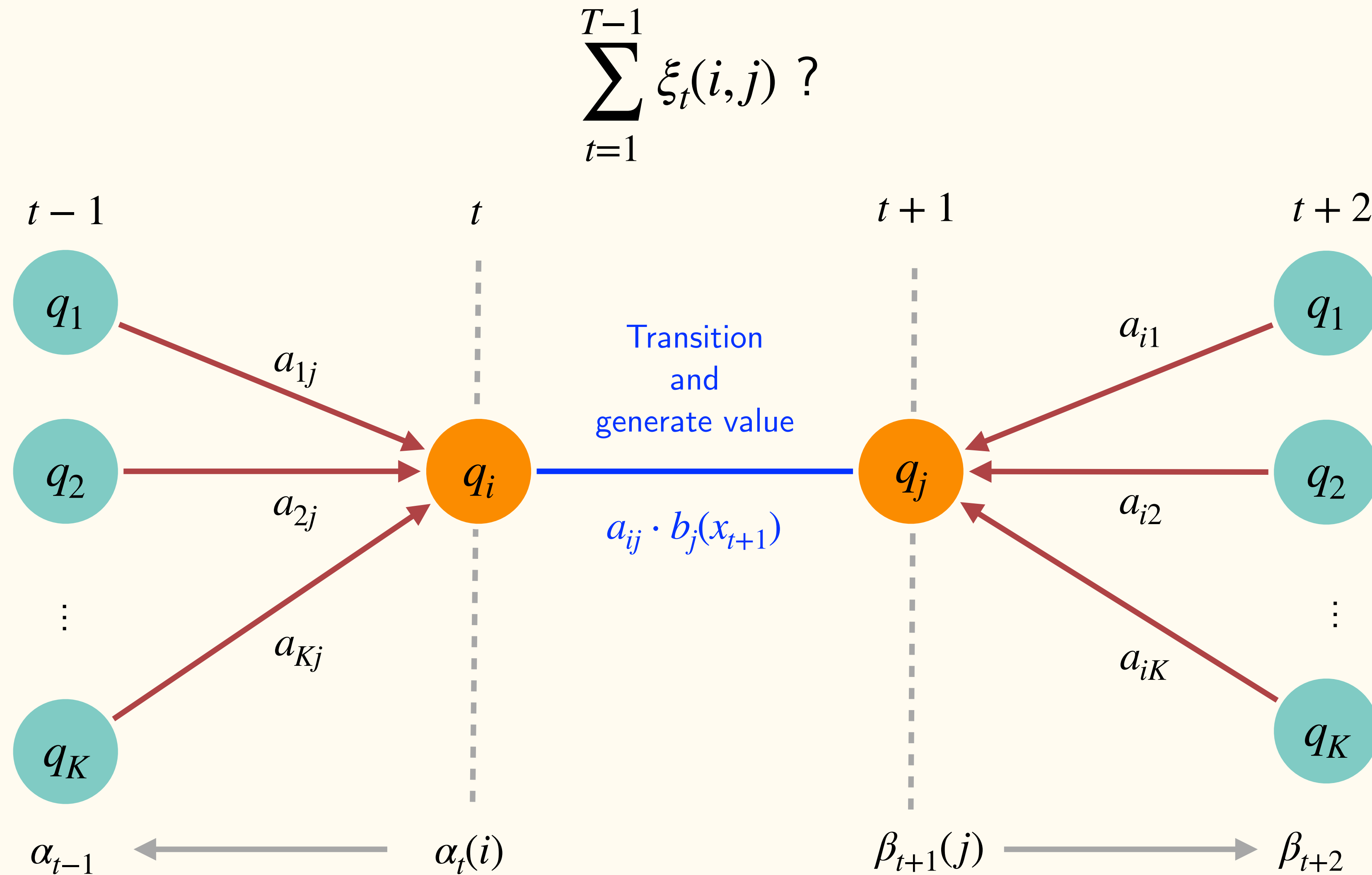


Sequence of Events

$$\xi_t(i, j) = \frac{\alpha_t(i) a_{ij} b_j(x_{t+1}) \beta_{t+1}(j)}{\sum_{i=1}^K \sum_{j=1}^K \alpha_t(i) a_{ij} b_j(x_{t+1}) \beta_{t+1}(j)}$$



Sequence of Events



Transition Matrix Re-Estimation

- $$\bar{a}_{ij} = \frac{\text{Expected number of times from state } q_i \text{ to } q_j}{\text{Expected number of transitions from } q_i} = \frac{\sum_{t=1}^{T-1} \xi_t(i, j)}{\sum_{t=1}^{T-1} \sum_{k=1}^K \xi_t(i, k)}$$

- Compute for all pairs (i, j) :
$$\bar{A} = \begin{bmatrix} \bar{a}_{11} & \bar{a}_{12} & \bar{a}_{13} \\ \bar{a}_{21} & \bar{a}_{22} & \bar{a}_{23} \\ \bar{a}_{31} & \bar{a}_{32} & \bar{a}_{33} \end{bmatrix}$$

Emission Probability Distribution Re-Estimation (Discrete)

- At each state q , we have an observation x which is a discrete value in the 'observation vocabulary' V .

- ▶
$$\overline{b_j(v_k)} = \frac{\text{Expected number of times in state } q_j \text{ and observing } v_k}{\text{Expected number of times in state } q_j}$$
- ▶ Need γ

Gamma Variable

[Forward](#)[Backward](#)

We define the following variable:

- $\gamma_t(i) = P(q^t = q_i | X, \Theta)$

i.e. the probability of being in state i at time t , given the observations and parameters Θ .

Can be expressed in terms of both forward and backward variables as:

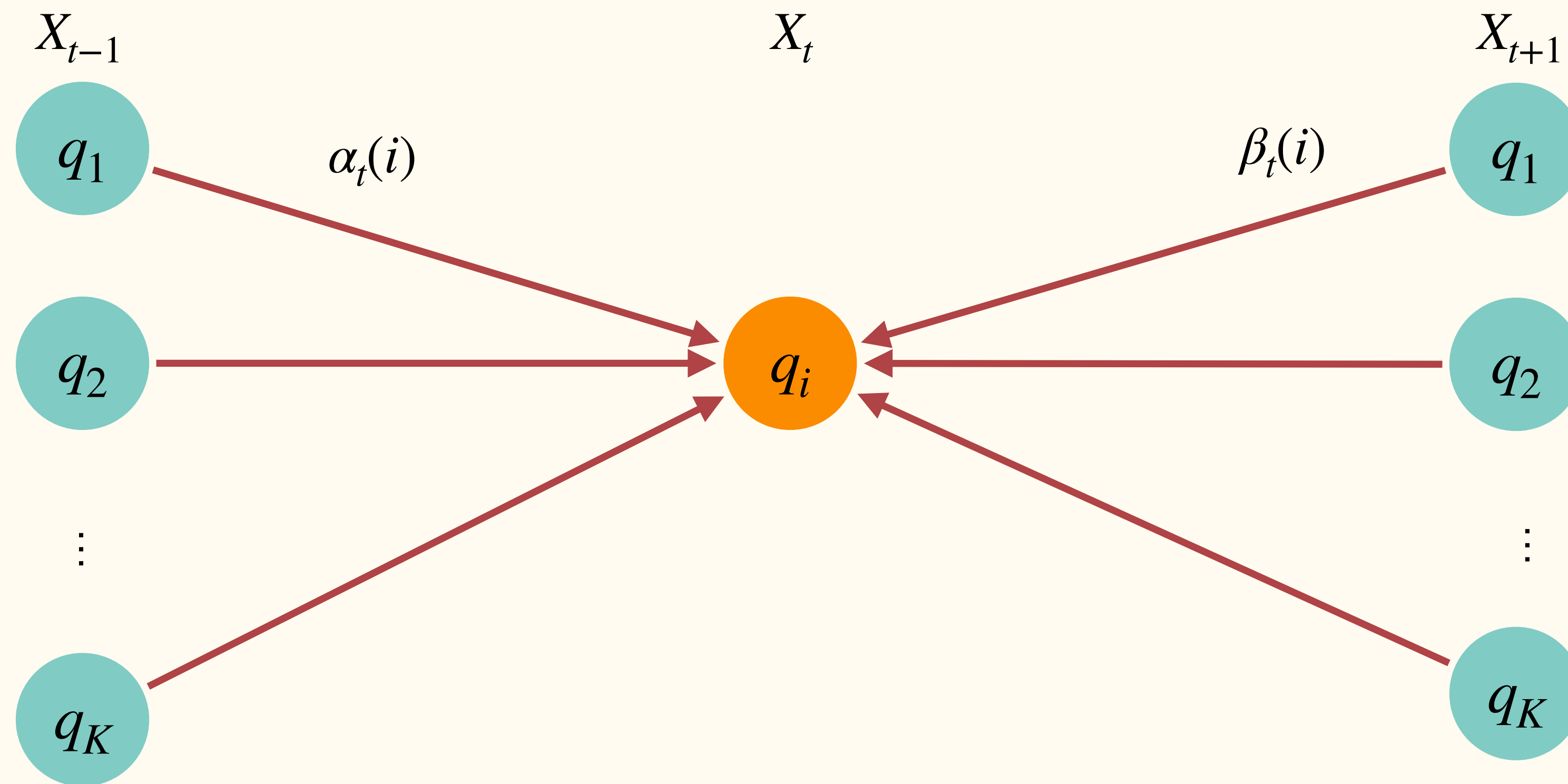
$$\gamma_t(i) = \frac{P(q_i^t, X | \Theta)}{P(X | \Theta)} = \frac{\alpha_t(i) \beta_t(i)}{P(X | \Theta)}$$

We can relate $\gamma_t(i)$ to $\xi(i, j)$:

$$\gamma_t(i) = \sum_{j=1}^K \xi(i, j)$$

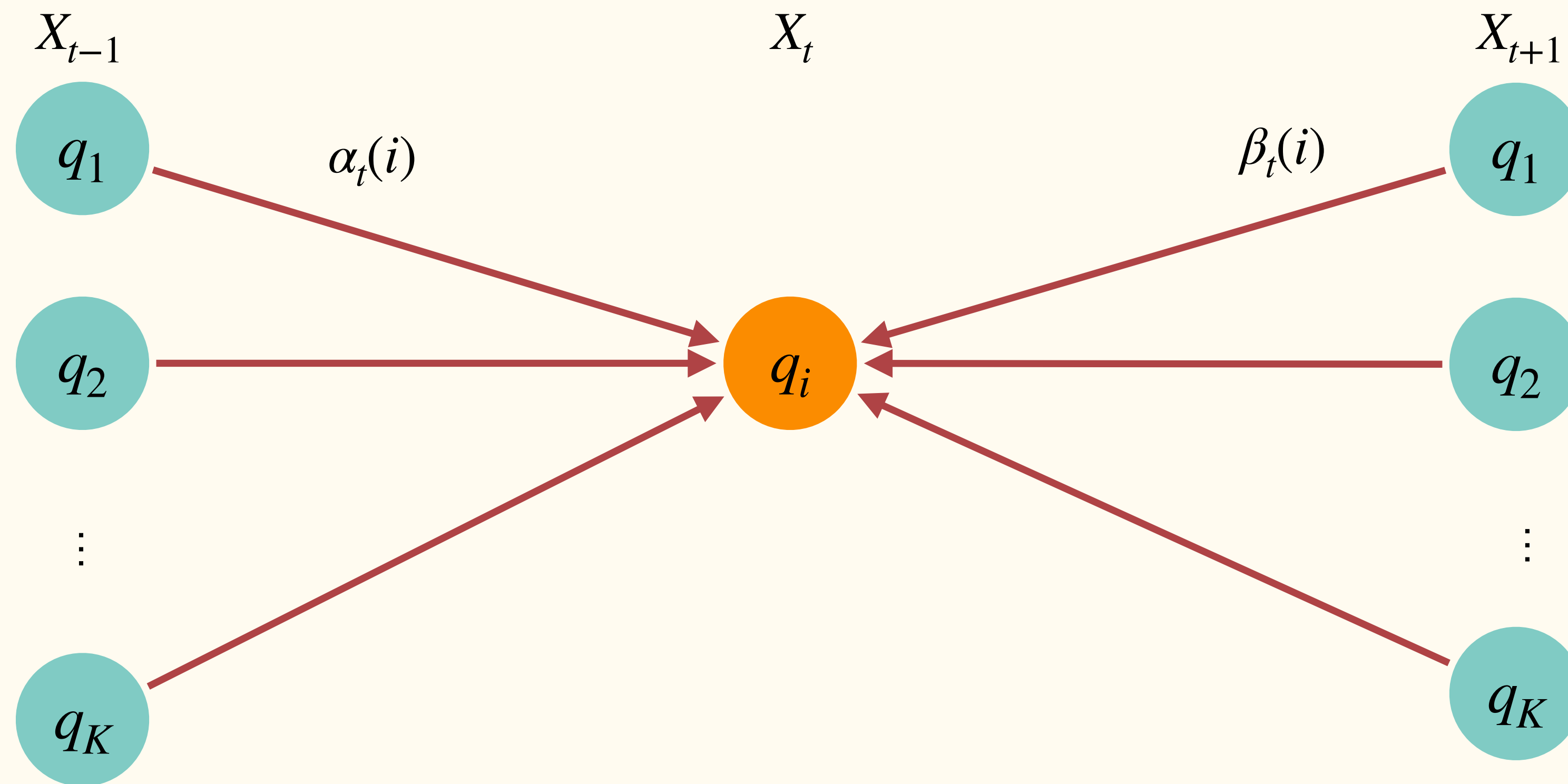
Gamma Variable

$$\gamma_t(i) = P(q^t = q_i | X, \Theta)$$



Gamma Variable

$$\gamma_t(i) = \frac{\alpha_t(i) \beta_t(i)}{P(X | \Theta)}$$



Emission Probability Distribution Re-Estimation (Discrete)

$$\overline{b_j(v_k)} = \frac{\text{Expected number of times in state } q_j \text{ and observing } v_k}{\text{Expected number of times in state } q_j} = \frac{\sum_{t=1 \& x_t=v_k}^T \gamma_t(i)}{\sum_{t=1}^T \gamma_t(i)}$$

Gamma Variable

We can express $\gamma_t(i)$ in 2 ways.

The expected number of times q_i is visited:

$$\sum_{t=1}^{T-1} \gamma_t(i)$$

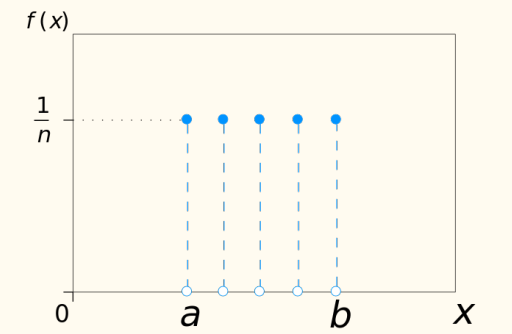
Useful for re-estimating
transition probabilities.

The expected number of times transitions are made from q_i :

$$\sum_{t=1}^T \gamma_t(i)$$

Useful for re-estimating
emission probability distribution.

Parameters Re-Estimation (Discrete)



We define the following formulas, as estimators for the:

- Initial state:** $\bar{\pi}_i = \gamma_1(i)$ \leftarrow ----- Expected frequency in state q_i at time $t = 1$
- Transition probabilities:** $\bar{a}_{ij} = \frac{\sum_{t=1}^{T-1} \xi_t(i, j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$

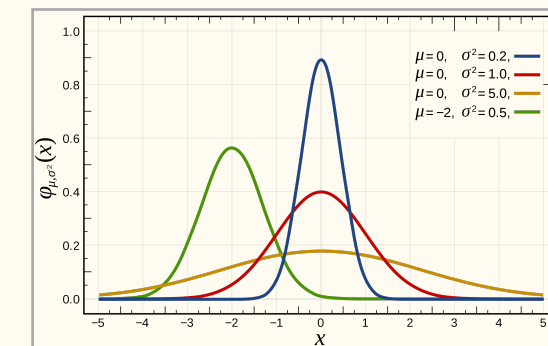
\leftarrow Expected number of transitions from state q_i to q_j

\leftarrow ----- Expected number of transitions from state q_i
- Emission PDF:** $\bar{b}_j(v_k) = \frac{\sum_{t=1 \& x_t=v_k}^T \gamma_t(i)}{\sum_{t=1}^T \gamma_t(i)}$

\leftarrow ----- Expected number of times in state q_j and observing v_k

\leftarrow ----- Expected number of times in state q_j

Parameters Re-Estimation (Continuous)

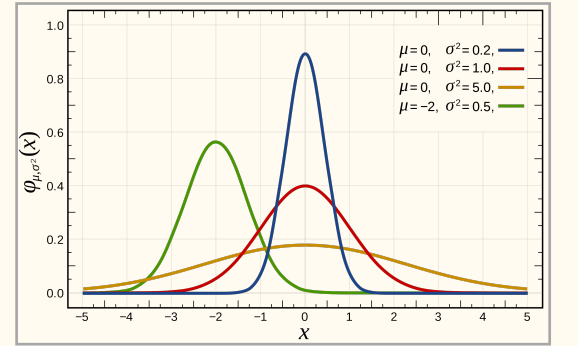


- The re-estimate for the **transition probabilities** \bar{a}_{ij} are the same.
- We are not interested in the emission probabilities, but the *parameters* that describe its distribution, e.g.: $\mathcal{N}(\mu, \sigma^2)$ or $\mathcal{N}(\mu, \Sigma)$
 - **Emission PDF**: $\bar{b}_j = \{\bar{\mu}, \bar{\sigma}^2\}$ or $\bar{b}_j = \{\bar{\mu}, \bar{\Sigma}\}$ for a Gaussian distribution.

- $$\bar{\mu}_{jk} = \frac{\sum_{t=1}^T \gamma_t(j, k) \cdot X_t}{\sum_{t=1}^T \gamma_t(j, k)}$$

- $$\bar{\sigma}_{jk}^2 = \frac{\sum_{t=1}^T \gamma_t(j, k) \cdot (X_t - \bar{\mu}_{jk})^2}{\sum_{t=1}^T \gamma_t(j, k)} \quad \text{or} \quad \bar{\Sigma}_{jk} = \frac{\sum_{t=1}^T \gamma_t(j, k) \cdot (X_t - \bar{\mu}_{jk}) \cdot (X_t - \bar{\mu}_{jk})^T}{\sum_{t=1}^T \gamma_t(j, k)}$$

Parameters Re-Estimation (Continuous)



Probability of being in state j at time t with k -th mixture component accounting for X_t :

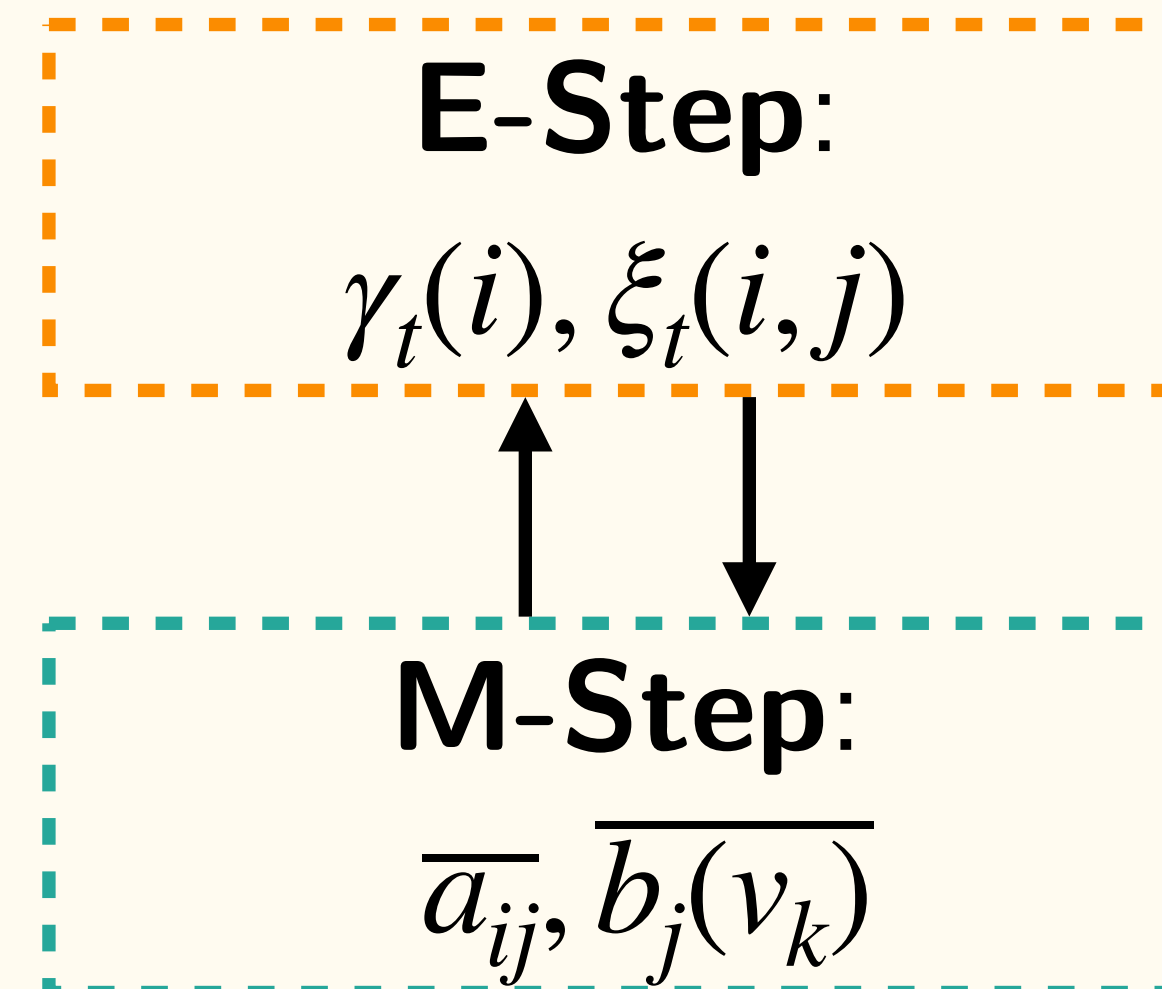
- $\bar{b}_j = \{\bar{c}, \bar{\mu}, \bar{\Sigma}\}$

- $$\bar{c}_{jk} = \frac{\sum_{t=1}^T \gamma_t(j, k)}{\sum_{t=1}^T \sum_{k=1}^M \gamma_t(j, k)}$$

- $$\gamma_t(j, k) = \frac{\alpha_t(j) \beta_t(j)}{\sum_{j=1}^K \alpha_T(j) \beta_t(j)} \cdot \frac{c_{jk} \mathcal{N}(X_t, \mu_{jk}, \Sigma_{jk})}{\sum_{m=1}^M c_{jm} \mathcal{N}(X_t, \mu_{jm}, \Sigma_{jm})}$$

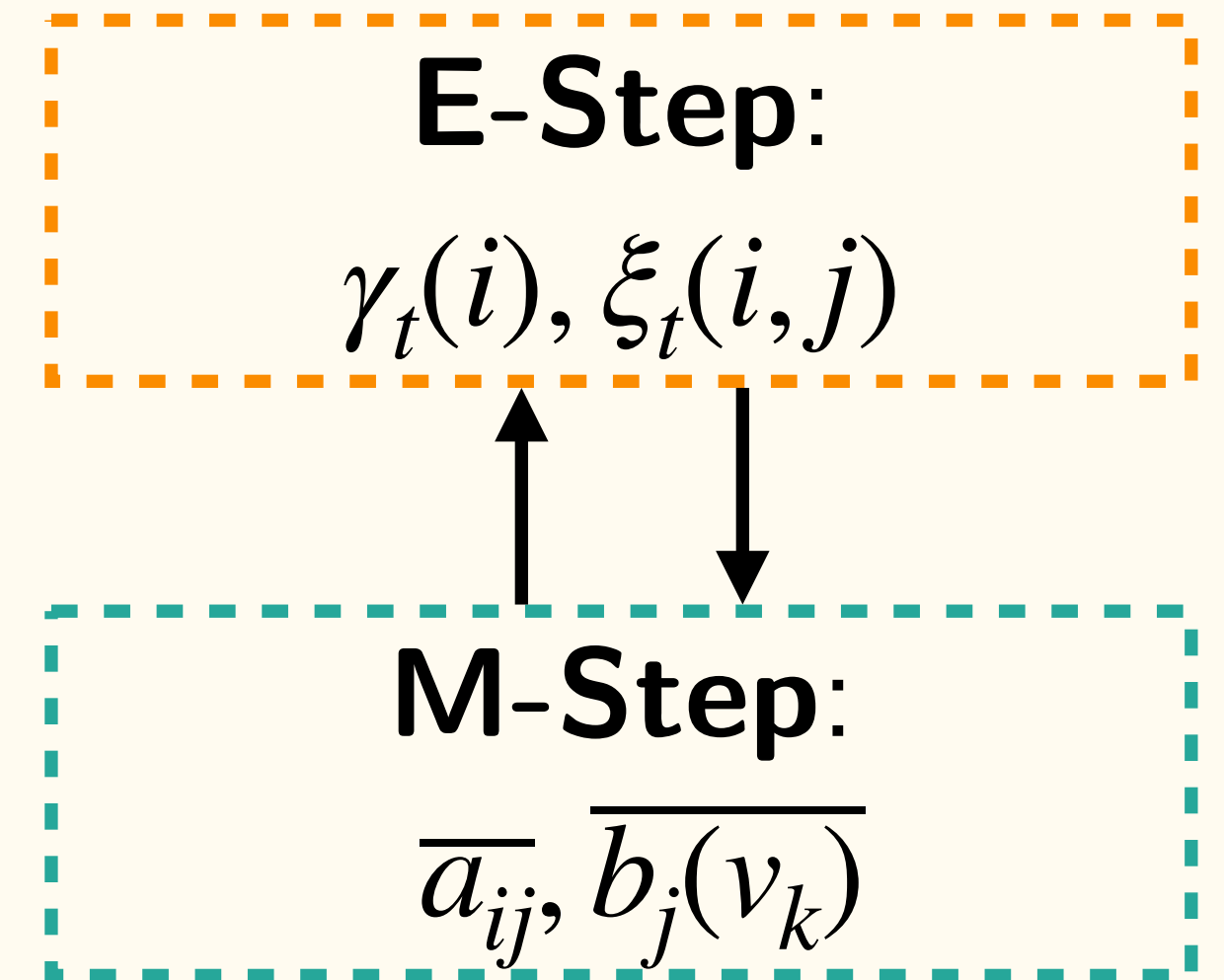
Baum-Welch EM Algorithm

- $\overline{a_{ij}}$ and $\overline{b_j(v_k)}$
 - Re-compute $\alpha_t, \beta_t, \gamma_t, \xi_t$
 - New values $\overline{a_{ij}}$ and $\overline{b_j(v_k)}$
 - ...



Iterative Training

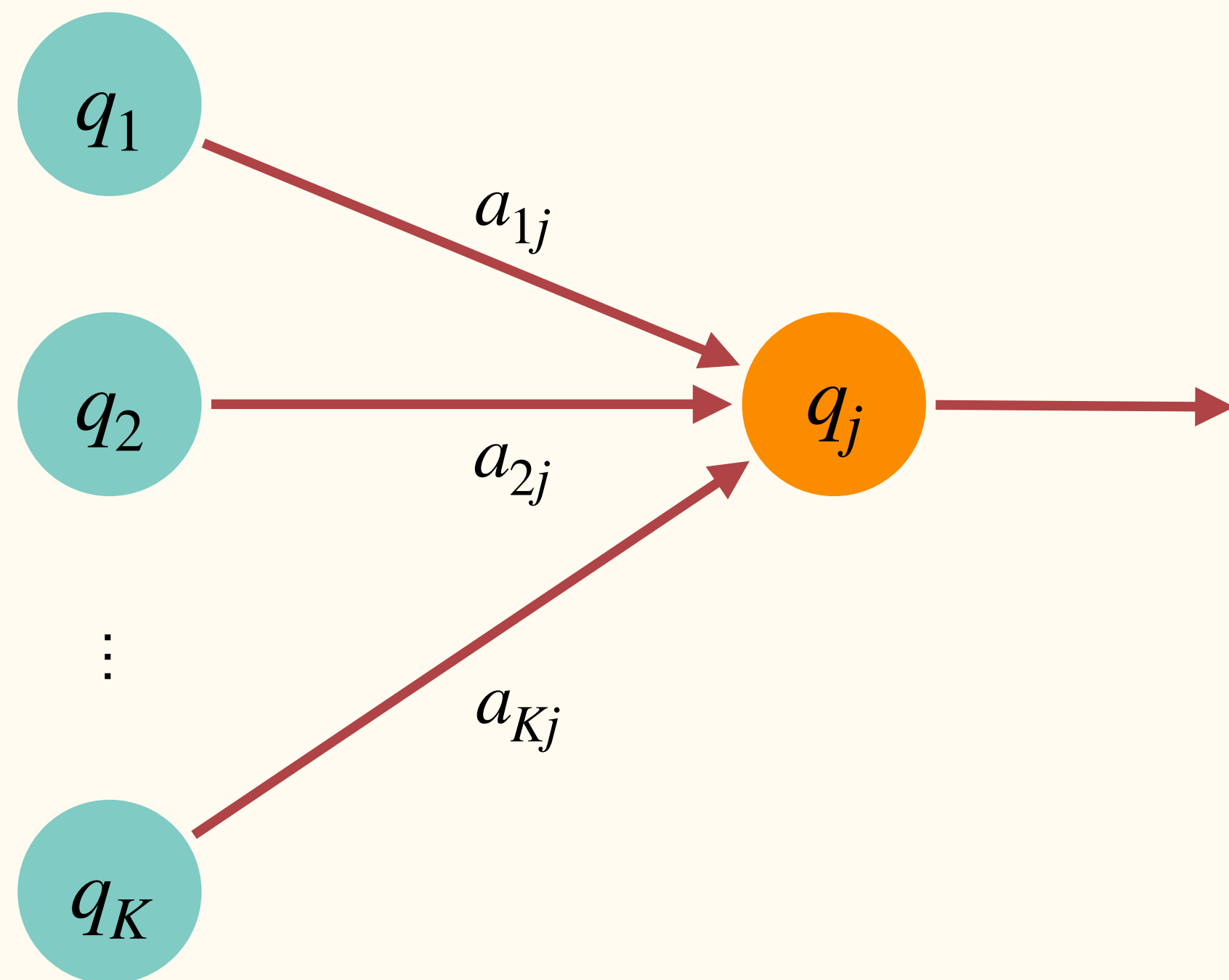
1. Estimate $p(X | \Theta)$ and $p(X | \bar{\Theta})$.
2. If $p(X | \bar{\Theta}) \geq p(X | \Theta)$:
 - Replace Θ with new estimate of parameters $\bar{\Theta}$.
 - Repeat the **E** and **M** steps of EM algorithm.
 - Go to step 1.
3. Else:
 - Terminate with $\bar{\Theta}$ as trained parameters (**convergence**).



II. Embedded Viterbi Training

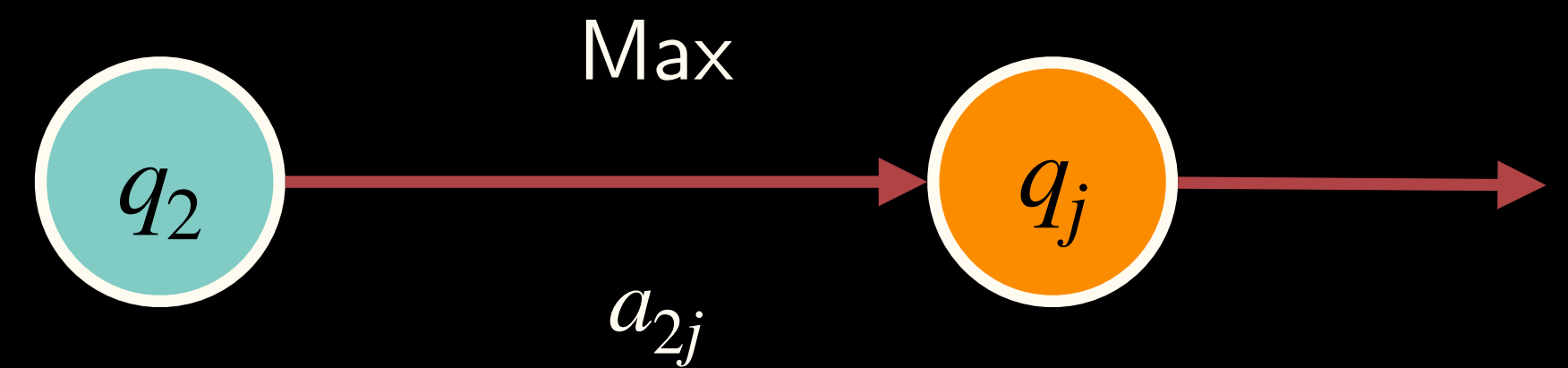
Forward-Backward

$$\alpha_t(i) = p(x_1, \dots, x_t, q^t = q_i | \Theta)$$



Viterbi

$$\delta_t(i) = \max p(q^1, \dots, q_i^t, x^1, \dots, x^t | \Theta)$$



Viterbi Algorithm

We define 2 variables:

1. $\delta_t(i)$: **highest likelihood** along a side path among all paths ending in state q_i at time t :

$$\triangleright \delta_t(i) = \max p(q^1, \dots, q_i^t, x^1, \dots, x^t \mid \Theta)$$

2. $\psi_t(i)$: variable to keep track of '**best path**' ending in state q_i at time t :

$$\triangleright \psi_t(i) = \operatorname{argmax} p(q^1, \dots, q_i^t, x^1, \dots, x^t \mid \Theta)$$

Viterbi Algorithm

1. Initialization:

- $\delta_1(i) = \pi_i b_i(x_1)$
- $\psi_1(i) = 0$

2. Recursion:

- $\delta_t(j) = \max_{1 \leq i \leq K} [\delta_{t-1}(i) a_{ij}] b_j(x_t)$
- $\psi_t(j) = \operatorname{argmax}_{1 \leq i \leq K} [\delta_{t-1}(i) a_{ij}]$

3. Termination:

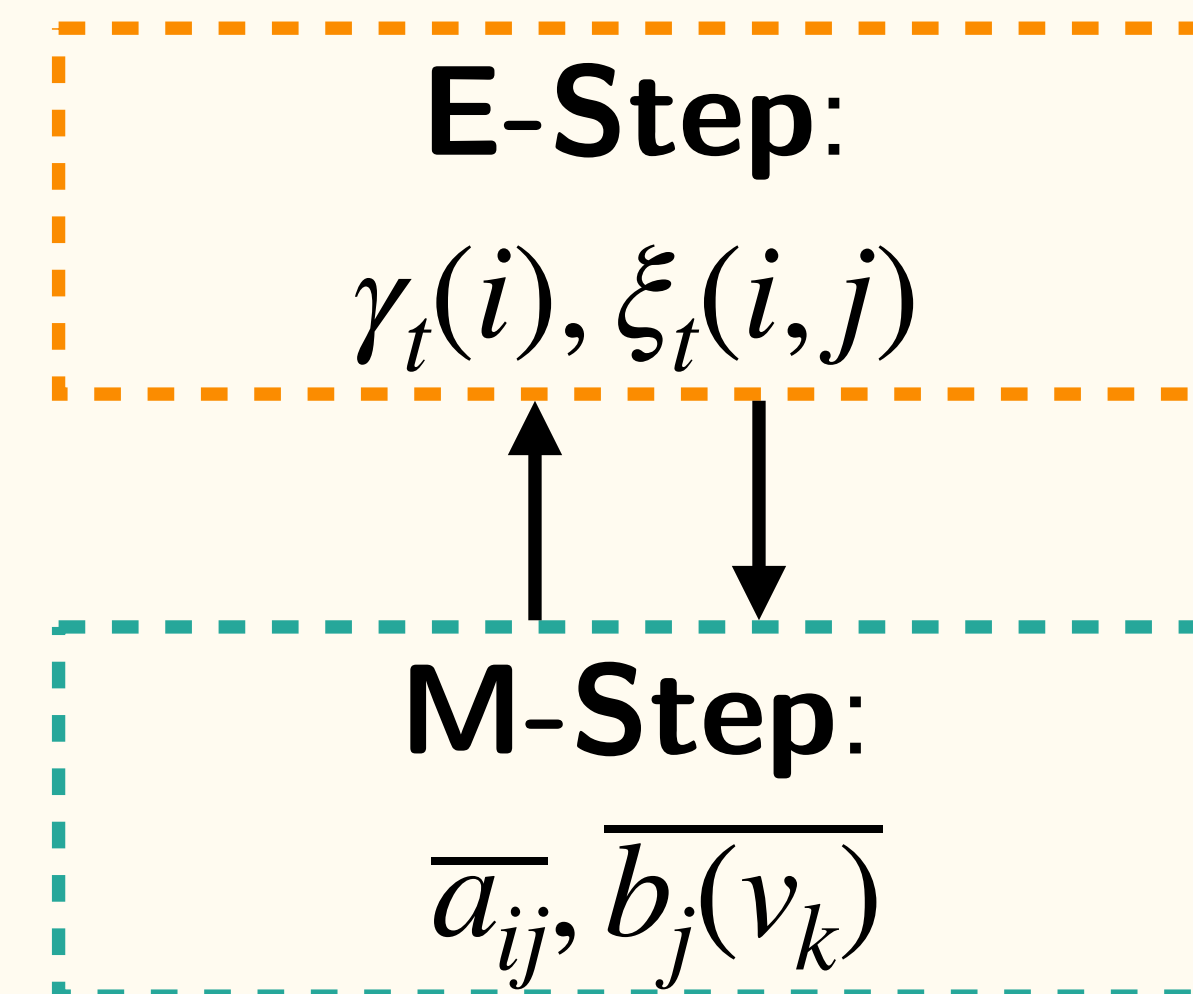
- $P^*(X | \Theta) = \max_{1 \leq i \leq K} \delta_T(i)$
- $q^{T*} = \operatorname{argmax}_{1 \leq i \leq K} [\delta_T(i)]$

4. Backtracking:

- $q^{t*} = \psi_{t+1}(q^{t+1*})$

Embedded Viterbi Approximation

1. Estimate $p(X|\Theta)$ and $p(X|\bar{\Theta})$.
2. If $p(X|\bar{\Theta}) \geq p(X|\Theta)$:
 - Replace Θ with new estimate of parameters $\bar{\Theta}$.
 - Repeat the **E** and **M** steps of EM algorithm.
 - Obtain optimal state sequence.
 - γ_t and ξ_t are either 0 or 1.
3. Else:
 - Terminate with $\bar{\Theta}$ as trained parameters (**convergence**).
- Faster than BW, as computational cost is less.



Solved !

Summary

Pros:

- Flexible topology.
- Rich mathematical framework.
- Wide range of applications.
- Powerful learning and decoding methods.
- Good abstraction for sequences, temporal aspects.

Cons:

- A priori selection of model topology and statistical distributions.
- First order Markov model for state transition.
- Lack of contextual information as correlation between successive acoustic vectors is ignored.
- Assumption of independence for computational efficiency.

Thank you !



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Parameters Re-Estimation

- Transition probabilities: $\overline{a}_{ij} = \frac{C(i \rightarrow j)}{\sum_k C(i \rightarrow k)}$
- Emission PDF:
 - $\overline{\mu}_j = \frac{\sum_{x \in Z_j} x}{|Z_j|}$
 - $\overline{\Sigma}_j = \frac{\sum_{x \in Z_j} (X_t - \overline{\mu}_j) \cdot (X_t - \overline{\mu}_j)^T}{|Z_j|}$
 - Z_j : Set of observed features assigned to q_j

Old Slides

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Likelihood Problem

—

Likelihood Estimation Problem

$$P(M|X, \Theta) = \frac{p(X|M, \Theta) P(M|\Theta)}{p(X|\Theta)}$$

- Computing $P(X|M, \Theta)$
- Fixed Θ
- Likelihood of a sequence of observations w.r.t. a HMM:
- Complexity: $\mathcal{O}(TK^T)$
 - Infeasible !

$$\begin{aligned} P(X|M) &= \sum_{Q \in M} P(X, Q|M) \quad \begin{array}{c} \vdots \\ \boxed{\text{Bayes Theorem}} \\ \downarrow \end{array} \\ &= \sum_{Q \in M} P(X|Q, M) P(Q|M) \\ &= \sum_{Q \in M} \prod_{t=1}^T p(x_t|q^t) \prod p_{q^{t-1}, q^t} \\ &= \sum_{Q \in M} \prod_{t=1}^T p(x_t|q^t) p_{q^{t-1}, q^t} \end{aligned}$$

Forward Recurrence - Log Space

1. Initialization:

- $\alpha_1(i) = \pi_i b_i(x_1), \quad 1 \leq i \leq K$

2. Recursion:

- $\alpha_{t+1}(j) = \left[\sum_{i=1}^K \alpha_t(i) a_{ij} \right] b_j(x_{t+1})$

3. Termination:

- $P(X|M) = \sum_{i=1}^K \alpha_T(i)$

1. Initialization:

- $\alpha_1^{(\log)}(i) = \log \pi_i + \log b_i(x_1)$

3. Recursion:

- $\alpha_{t+1}^{(\log)}(j) = [\text{logsum}_{i=1}^K (\alpha_t^{(\log)}(i) + \log a_{ij})] + \log b_j(x_{t+1})$

6. Termination:

- $\log P(X|M) = [\text{logsum}_{i=1}^K \alpha_T^{(\log)}(i)]$

Decoding Problem

Decoding Problem

- Estimating an optimal sequence of states given a sequence of observations and the parameters of a model.
 - Viterbi algorithm

Viterbi Algorithm - Log Space

1. Initialization:

- $\delta_1^{(\log)}(i) = \log \pi_i + \log b_i(x_1)$
- $\psi_1(i) = 0$

2. Recursion:

- $\delta_t^{(\log)}(i) = \max_{1 \leq j \leq K} [\delta_{t-1}^{(\log)}(j) + \log a_{ji}] + \log b_i(x_t)$
- $\psi_t(j) = \operatorname{argmax}_{1 \leq i \leq K} [\delta_{t-1}^{(\log)}(i) + \log a_{ij}]$

3. Termination:

- $\log P^*(X | \Theta) = \max_{1 \leq i \leq K} \delta_T^{(\log)}(i)$
- $q_T^* = \operatorname{argmax}_{1 \leq i \leq K} [\delta_T^{(\log)}(i)]$

4. Backtracking:

- $q^{t*} = \psi_{t+1}(q^{t+1*})$

Viterbi Algorithm

In summary, given a:

- Sequence of observations $X = \{x_1, \dots, x_n, \dots x_T\}$
- Parameters Θ

The Viterbi algorithm returns the:

- Optimal path $Q^* = \{q_1^*, \dots, q_T^*\}$
- Likelihood along the best path $P^*(X | \Theta)$