

# Mathematics IV

(Probability and Statistics)

## MATH 403

### Lecture 6

**Dr. Phoebe Edward Nashed**

# Lecture 6 - Outline

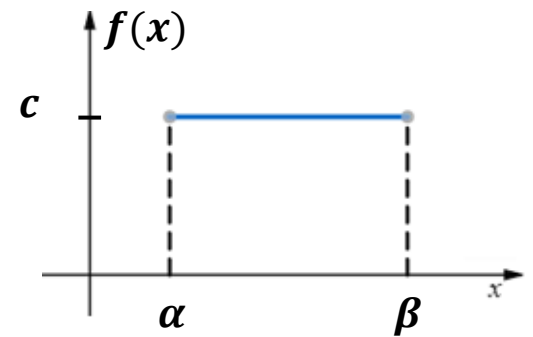
## ❖ Special Continuous Distributions

- Uniform Distribution
- Exponential Distribution
- Normal (Gaussian) Distribution

# I. Uniform Distribution

❖ Suppose that a continuous random variable  $X$  can assume values only in a bounded interval, say the open interval  $(\alpha, \beta)$ , and suppose that the probability density function (PDF) is **constant**, say

$$f(x) = c \quad \text{for} \quad \alpha \leq x \leq \beta$$



❖ This implies that  $c = 1/(\beta - \alpha)$ , since

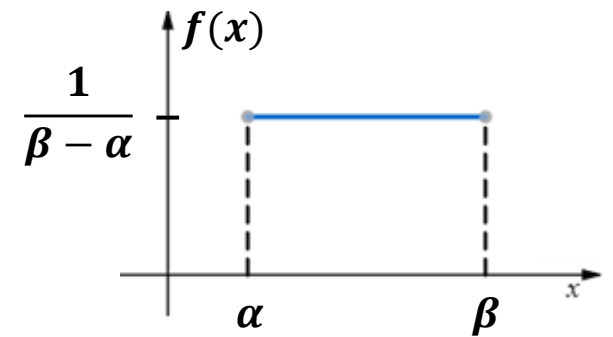
$$\int_{-\infty}^{\infty} f(x) \, dx = \int_{\alpha}^{\beta} c \, dx = c (\beta - \alpha) = 1$$

❖ If we define  $f(x) = 0$  outside the interval, then we obtain  $f(x) \geq 0$  for all  $x$ .

# I. Uniform Distribution – Def. - PDF

- ❖ Definition 1: A continuous random variable  $X$  is said to have *uniform distribution* with parameters  $\alpha$  and  $\beta$  ( $\alpha < \beta$ ) if and only if its [PDF](#) is given by

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & \text{for } \alpha \leq x \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

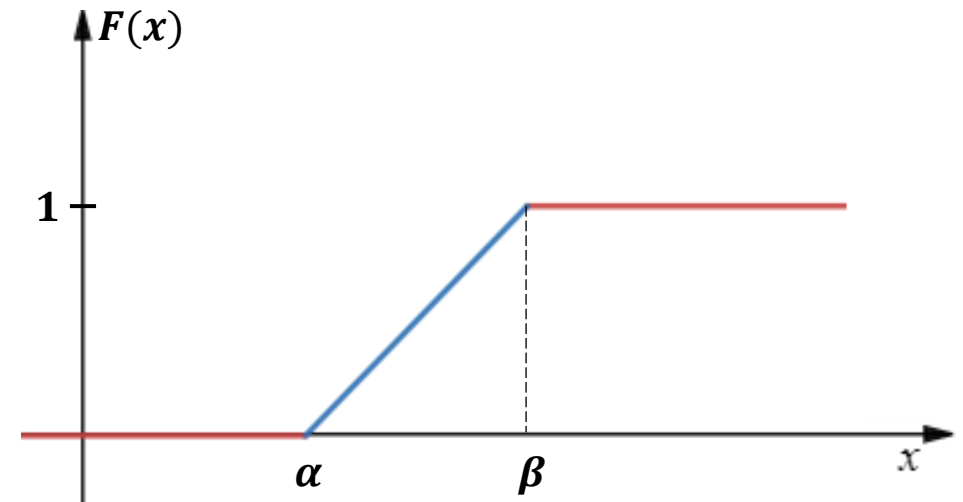


- ❖ This distribution provides a probability model for selecting a point "at random" from an interval  $(\alpha, \beta)$ .

# I. Uniform Distribution - CDF

- ❖ The cumulative distribution function (CDF) of a uniformly distributed random variable  $X$  has the form:

$$F(x) = \begin{cases} 0 & \text{for } x \leq \alpha \\ \frac{x - \alpha}{\beta - \alpha} & \text{for } \alpha < x < \beta \\ 1 & \text{for } x \geq \beta \end{cases}$$



# I. Uniform Dist. – Mean and Var.

❖ The mean and the variance of *uniform distribution* are given by:

$$E(X) = \mu = \frac{\beta + \alpha}{2} \quad \text{and} \quad \text{Var}(X) = \sigma^2 = \frac{(\beta - \alpha)^2}{12}$$

# I. Uniform Dist. – Mean

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{\alpha}^{\beta} x \left( \frac{1}{\beta - \alpha} \right) dx = \left( \frac{1}{\beta - \alpha} \right) \frac{x^2}{2} \Big|_{\alpha}^{\beta} = \left( \frac{1}{\beta - \alpha} \right) \frac{\beta^2 - \alpha^2}{2} = \frac{\beta + \alpha}{2}$$

# I. Uniform Dist. – Variance

$$\diamond E(X) = \frac{\beta + \alpha}{2}$$

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$\begin{aligned} & b^3 - a^3 \\ &= (b - a)(b^2 + ab + a^2) \end{aligned}$$

$$= \int_{\alpha}^{\beta} x^2 \left( \frac{1}{\beta - \alpha} \right) dx = \left( \frac{1}{\beta - \alpha} \right) \frac{x^3}{3} \Big|_{\alpha}^{\beta} = \left( \frac{1}{\beta - \alpha} \right) \frac{\beta^3 - \alpha^3}{3} = \frac{\beta^2 + \beta\alpha + \alpha^2}{3}$$

$$\diamond \therefore \text{Var}(X) = E(X^2) - [E(X)]^2$$

$$\text{Var}(X) = \frac{\beta^2 + \beta\alpha + \alpha^2}{3} - \left( \frac{\beta + \alpha}{2} \right)^2 = \frac{\beta^2 - 2\beta\alpha + \alpha^2}{12} = \frac{(\beta - \alpha)^2}{12}$$



# Example 1

- ❖ Suppose a train arrives at a subway station regularly every **10 min**. If a passenger arrives at the station without knowing the timetable, then the waiting time to catch the train is uniformly distributed with density

$$f(x) = \begin{cases} \frac{1}{10} & \text{for } 0 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

- i. Calculate the “average” waiting time.
- ii. Calculate the probability of waiting for the train for less than 3 min.

# Example 1 – Sol.

❖ Since  $X$  is uniformly distributed with PDF:

$$f(x) = \begin{cases} \frac{1}{10} & \text{for } 0 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

❖ Therefore,  $\alpha = 0, \beta = 10$ .

❖ The “average” waiting time is

$$E(X) = \frac{\beta + \alpha}{2} = (10 + 0)/2 = 5 \text{ min.}$$

# Example 1 – Sol.

❖ Since  $X$  is uniformly distributed with PDF:

$$f(x) = \begin{cases} \frac{1}{10} & \text{for } 0 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

❖ Therefore,  $\alpha = 0, \beta = 10$ .

❖ The probability of waiting for the train for **less than 3 min** is obviously **0.3** (or **30 %**) which can be calculated as follows:

$$P(X \leq 3) = F(3) = \frac{3 - 0}{10 - 0} = 0.3$$

$$F(x) = \frac{x - \alpha}{\beta - \alpha}$$

## Example 2

- ❖ Suppose the time it takes a data collection operator to fill out an electronic form for a database is uniformly between 1.5 and 2.2 minutes.
- (a) What is the mean and the variance of time, it takes an operator to fill out the form?
- (b) What is the probability that it takes less than two minutes to fill out the form?

## Example 2 – Sol.

- ❖ Since  $X$  is uniformly between 1.5 and 2.2 minutes;  $\alpha = 1.5, \beta = 2.2$
- ❖ Therefore,

$$E(X) = \mu = \frac{\beta + \alpha}{2} = \frac{2.2 + 1.5}{2} = 1.85$$

$$Var(X) = \sigma^2 = \frac{(\beta - \alpha)^2}{12} = \frac{(2.2 - 1.5)^2}{12} = 0.041$$

- ❖ The probability that it takes less than two minutes to fill out the form:

$$P(X \leq 2) = F(2) = \frac{2 - 1.5}{2.2 - 1.5} = 0.714$$

$$F(x) = \frac{x - \alpha}{\beta - \alpha}$$

## Example 3

- ❖ Suppose  $X$  has a continuous uniform distribution over the interval  $[-1, 1]$ . Find the value for  $k$  such that

$$P(-k \leq X \leq 2k) = 0.6$$

## Example 3 – Sol.

- ❖ Suppose  $X$  has a continuous uniform distribution over the interval  $[-1, 1]$ . Find the value for  $k$  such that

$$P(-k \leq X \leq 2k) = 0.6$$

- ❖ Since  $X$  is uniformly distributed,  $\alpha = -1, \beta = 1$ ,
- ❖ Therefore, its PDF is:

$$f(X) = \frac{1}{\beta - \alpha} = \frac{1}{2} \quad -1 \leq x \leq 1$$

$$P(-k \leq X \leq 2k) = \int_{-k}^{2k} 0.5 \, dx = 0.5(2k + k) = 1.5k = 0.6$$

$$\therefore k = 0.6/1.5 = 0.4$$

## Example 3 – Another Sol.

- ❖ Suppose  $X$  has a continuous uniform distribution over the interval  $[-1, 1]$ . Find the value for  $k$  such that

$$P(-k \leq X \leq 2k) = 0.6$$

- ❖ Since  $X$  is uniformly distributed,  $\alpha = -1, \beta = 1$ ,
- ❖ Therefore, its PDF is:

$$f(X) = \frac{1}{\beta - \alpha} = \frac{1}{2} \quad -1 \leq x \leq 1$$

$$P(-k \leq X \leq 2k) = F(2k) - F(-k) = \left(\frac{2k + 1}{2}\right) - \left(\frac{-k + 1}{2}\right) = 1.5k$$
$$\therefore k = 0.6/1.5 = 0.4$$



# Poisson Distribution - Review

- ❖ Poisson distribution is used to approximate Binomial probabilities when the number of trials  $n$  is large ( $n \rightarrow \infty$ ), the probability of success  $p$  is small ( $p \rightarrow 0$ ) and  $np$  remains constant.
- ❖ Such an approximation is acceptable, say, for  $n \geq 30$  and  $p \leq 0.05$ , and it becomes more accurate for larger  $n$ .

# Poisson Distribution – Def. - Review

- ❖ A random variable  $X$  is said to have *Poisson distribution* with parameter  $\lambda > 0$  if and only if its probability mass function (PMF) is given by

$$f(x) = \frac{e^{-\lambda} \lambda^x}{x!} \quad \text{for } x = 0, 1, 2, \dots$$

# Poisson Distribution – Mean and Variance - Review

❖ Hence, the mean and the variance of Poisson distribution are given by

$$E(X) = \mu = \lambda$$

$$Var(X) = \sigma^2 = \lambda$$

As  $q \rightarrow 1$

# Illustrative Example

- ❖ Imagine that you are a traffic engineer and you need to model the random variable that represents the number of cars that pass a certain point in an hour.
- ❖ Assume the following:
  - The number of cars passing in any hour does not affect another hour (**independent** events/occurrences)
  - The probability of success (car passing) is **constant** among different hours, it is  $p$ .
- ❖ You can try to model it by a Binomial distribution, because every hour, either success (car passes) or failure (car does not pass) occurs.
- ❖ So,  $E(X) = \lambda = np = 60 \cdot \frac{\lambda}{60}$

# Illustrative Example

- ❖ You can try to model it by a Binomial distribution, because every hour, either success (car passes) or failure (car does not pass) occurs.
- ❖ So,  $E(X) = \lambda = np = 60 \cdot \frac{\lambda}{60}$
- ❖  $P(X = k) = \binom{60}{k} \left(\frac{\lambda}{60}\right)^k \left(1 - \frac{\lambda}{60}\right)^{60-k}$
- ❖ But, what if two or more cars pass within the same minute?! We have to increase granularity... using seconds...
- ❖  $P(X = k) = \binom{3600}{k} \left(\frac{\lambda}{3600}\right)^k \left(1 - \frac{\lambda}{3600}\right)^{3600-k}$
- ❖ But, what if two or more cars pass within the same second?! We have to increase granularity...
- ❖ As  $n \rightarrow \infty$  and  $p \rightarrow 0$ , the Binomial distribution will become **Poisson Distribution** (as proved last lecture).

## II. Exponential Distribution

- ❖ Exponential distribution is often used to model time: **waiting time**, **interarrival time**, hardware **lifetime**, failure time, time between telephone calls, etc. As we shall see below, in a sequence of rare events, when the number of events is **Poisson**, the time between events is **Exponential**.

# Relation between Poisson Dist. and Exponential Dist.

- ❖ Conditions for a Poisson Process:
  - Events must occur at a **constant rate** (constant  $\lambda$ ).
  - Events must be **independent**.
  - Events **cannot** occur **simultaneously**.

# Relation between Poisson Dist. and Exponential Dist.

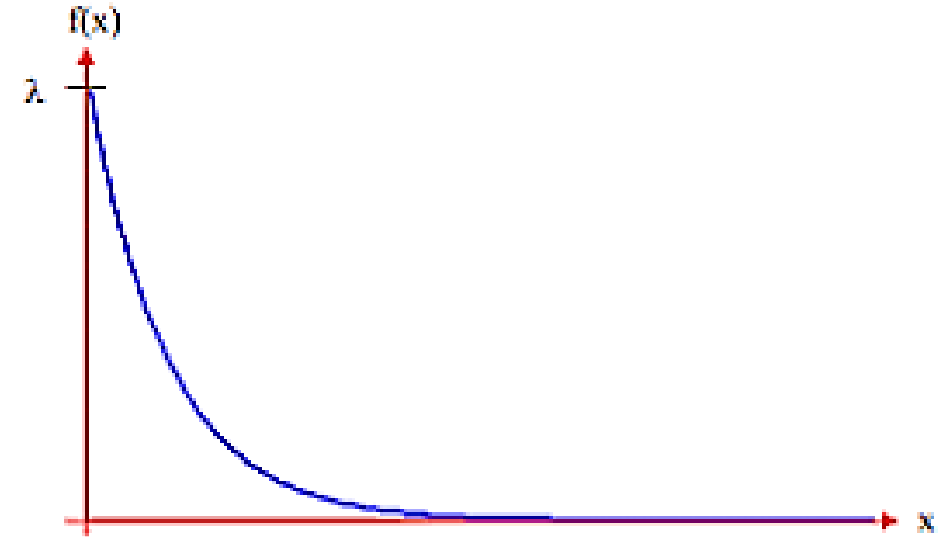
Exponential Distribution	Poisson Distribution
Continuous distribution	Discrete distribution
It represents <b>time</b> between a number of events in a Poisson Process (inter-arrival time)	It represents <b>number</b> of events (occurrences) in a time interval (rare events, $p \rightarrow 0$ ).
Parameter: $\lambda$ : rate/frequency of occurrences per unit time – unit: 1/time ( $sec^{-1}$ , $min^{-1}$ , ... etc.)	Parameter: $\lambda$ : average (expected) number of events (occurrences)
Example: Time between two cars' arrivals.	Example: Number of cars passing a certain square in an hour.



## II. Exponential Distribution – Def. - PDF

- ❖ A random variable  $X$  is said to have *Exponential distribution* with parameter  $\lambda > 0$  if and only if its PDF is given by

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$



## II. Exponential Distribution – Def. - CDF

❖ Its PDF is:

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

❖ Its CDF can be obtained by:

$$F(x) = \int_{-\infty}^x f(t) dt = \int_0^x \lambda e^{-\lambda t} dt = \lambda \left( \frac{e^{-\lambda t}}{-\lambda} \bigg|_0^x \right) = 1 - e^{-\lambda x}$$

$$\therefore F(x) = \begin{cases} 1 - e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{for } x \leq 0 \end{cases}$$

## II. Exponential Distribution – Mean

❖ Its PDF is:

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

❖ Its Mean is:

$$E(X) = \mu = \int_{-\infty}^{\infty} x f(x) dx = \int_0^{\infty} x (\lambda e^{-\lambda x}) dx = \frac{1}{\lambda}$$

Using  
integration by  
parts

## II. Exponential Distribution – Variance

❖ Its PDF is:

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

❖ Its Mean is:  $E(X) = \frac{1}{\lambda}$

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_0^{\infty} x^2 (\lambda e^{-\lambda x}) dx = \frac{2}{\lambda^2}$$

Using  
integration by  
parts twice

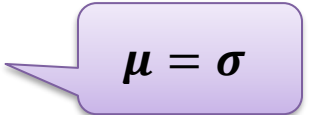
❖ Its Variance:

$$\therefore \text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

## II. Exponential Dist. – Mean and Var.

$$E(X) = \mu = \frac{1}{\lambda}$$

$$Var(X) = \sigma^2 = \frac{1}{\lambda^2}$$


$$\mu = \sigma$$

## II. Exponential Dist. - $\lambda$ Parameter

- ❖ The quantity  $\lambda$  is a parameter of *Exponential distribution*.
- ❖ It is clear from the fact that  $E(X) = 1/\lambda$ , that  $\lambda$  is the **reciprocal** of the mean of the exponentially distributed random Variable  $X$ .
- ❖ If  $X$  is time, measured in minutes, then  $\lambda$  is a **frequency (rate)**, measured in  $\text{min}^{-1}$ .
- ❖ **For example**, if arrivals occur every half a minute, on the average, then  $E(X) = 0.5$ , therefore  $\lambda = 2$ , meaning that they occur with a frequency (arrival rate) of 2 arrivals per minute.
- ❖ This  $\lambda$  has the same meaning as the parameter of **Poisson distribution**.
- ❖  $\lambda$  = Frequency parameter, the number of events per time unit.

## Example 4

❖ Jobs are sent to a printer at an average rate of 3 jobs per hour.

(a) What is the expected time between jobs?

(b) What is the probability that the next job is sent within 5 minutes?

## Example 4 – Sol.

❖ Job arrivals represent rare events, thus the time  $X$  between them is Exponential with the given parameter  $\lambda = 3 \text{ hr}^{-1}$  (jobs per hour).

a)  $E(X) = 1/\lambda = 1/3 \text{ hr}$  or 20 min between jobs;

b) Convert to the same measurement unit: 5 min =  $(1/12) \text{ hr}$ . Then,

$$P(X < 1/12) = F(1/12) = 1 - e^{-\lambda(1/12)} = 1 - e^{-1/4} = 0.2212$$



## Example 5

- ❖ The time to failure of fans in a PC can be modeled by an exponential distribution with mean **4000** hours. What proportion of the fans will last at least **10,000** hours?

## Example 5 – Sol.

- ❖ The time to failure of fans in a PC can be modeled by an exponential distribution with mean **4000** hours. What proportion of the fans will last at least **10,000** hours?

$$\mu = 4000 \quad \rightarrow \quad \lambda = 1/4000$$

$$\begin{aligned} P(X \geq 10,000) &= 1 - F(10,000) \\ &= 1 - [1 - e^{-(10000/4000)}] = e^{-(2.5)} = \mathbf{0.082} \end{aligned}$$

- ❖ Another Approach

$$P(X \geq 10,000) = \int_{10000}^{\infty} \frac{1}{4000} e^{-(x/4000)} dx = e^{-(2.5)} = \mathbf{0.082}$$

## II. Exp. Dist. - Memoryless Property

### ❖ Theorem

❖ If the random variable  $X$  has *Exponential distribution*, then

$$P(X > T + t | X > T) = P(X > t)$$

for all  $T > 0$  and  $t > 0$ .

❖ Proof:

$$\begin{aligned} P(X > T + t | X > T) &= \frac{P(X > T + t \cap X > T)}{P(X > T)} = \frac{P(X > T + t)}{P(X > T)} \\ &= \frac{1 - F(T + t)}{1 - F(T)} = \frac{e^{-\lambda(T+t)}}{e^{-\lambda(T)}} = e^{-\lambda(t)} = P(X > t) \end{aligned}$$

## II. Exp. Dist. - Memoryless Property - Meaning

- ❖ **Note that:** If  $X$  is the lifetime of a component has *Exponential distribution*, then the probability that the component will last more than  $T + t$  time's units given that it has already lasted more than  $T$  units is the same as that of a new component lasting more than  $t$  units.
- ❖ In other words, an old component which is still working is **just as reliable as** a new component.

## Example 6

- ❖ Suppose a certain solid-state component has a lifetime time  $X$  with mean **100** hours. Find the probability that the component will last at least **80** hours given that it is already worked more than **30** hours.

## Example 6 – Sol.

- ❖ Suppose a certain solid-state component has a lifetime time  $X$  with mean **100** hours. Find the probability that the component will last at least **80** hours given that it is already worked more than **30** hours.

$$\mu = 100 \quad \rightarrow \quad \lambda = 0.01$$

$$P(X > 80 | X > 30) = P(X > 50)$$

$$= 1 - F(50) = e^{-0.5} = 0.6065$$



# Thank You 😊