Assignment #1

Elliot Smith 8/26/2018

Problem 1

Part i

f(x) is the pdf of any of the x_i , and since they are *iid*, they all follow one pdf.

The joint pdf is: $f_{X_1,...,X_n}(x_1,...,x_n) = f(x_1) \times ... \times f(x_n)$ because they are *iid*.

Since you can multiply the right-hand side in any order, this implies symmetry in regards to the left-hand side.

As a result, $X_1, ..., X_n$ are exchangeable.

Part ii

$$p(y_1, ..., y_n) = \int p(y_1, ..., y_n | \theta) p(\theta) d(\theta)$$

$$= \int \left(\prod_{i=1}^n p(y_i | \theta) \right) p(\theta) d(\theta))$$

$$= \int \left(\prod_{i=1}^n p(y_{\pi_i} | \theta) \right) p(\theta) d(\theta))$$

$$= \int p(y_{\pi_1, ..., y_{\pi_n}} | \theta) p(\theta) d(\theta)$$

$$= p(y_{\pi_1, ..., y_{\pi_n}} | \theta)$$

Where:

- The first line is the definition of marginal probability
- The second line is because the Y_i 's are conditionally *iid*
- The third line is because the product does not depend on order
- The fourth line we are converting back to the form used in the first line
- Finally, the last line is the definition of marginal probability

Problem 2

Part a

Prior Density: $p(\theta) = 1$

Sampling Distribution: $p(y|\theta) = \binom{n}{y} \theta^y (1-\theta)^{n-y}$

Posterior Density: $p(\theta|y) \propto p(y|\theta)p(\theta) = \binom{n}{y}\theta^y(1-\theta)^{n-y}$ Posterior Distribution: $\theta|y \sim Beta(y+1,(n-y)+1)$

$$\binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma((y+1)+(n-y+1))} = \binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma(n+2)}$$

$$= \frac{n!}{y!(n-y)!} \frac{y!(n-y)!}{(n+1)!}$$

$$= \frac{n!}{(n+1)!}$$

$$= \frac{1}{n+1}$$

Part b

Prior Distribution: $\theta \sim Beta(\alpha, \beta)$

Sampling Distribution: $y|\theta \sim Binomial(n,\theta)$

Posterior Density:

$$p(\theta|y) = \binom{n}{y} \theta^{y} (1-\theta)^{n-y} \times \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$
$$= \binom{n}{y} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{y+\alpha-1} (1-\theta)^{n-y+\beta-1}$$
$$\propto Beta(y+\alpha, (n-y)+\beta)$$

Posterior Mean:

$$\frac{y+\alpha}{\alpha+\beta+n} = \frac{y}{n} + \lambda \left(\frac{\alpha}{\alpha+\beta} - \frac{y}{n}\right)$$

$$\frac{y+\alpha}{\alpha+\beta+n} - \frac{y}{n} = \lambda \left(\frac{\alpha}{\alpha+\beta} - \frac{y}{n}\right)$$

$$\frac{ny+n\alpha-\alpha y - \beta y - ny}{(\alpha+\beta+n)n} = \lambda \left(\frac{n\alpha-\alpha y - \beta y}{(\alpha+\beta)n}\right)$$

$$\lambda = \frac{\alpha+\beta}{\alpha+\beta+n}$$

Since λ will always be between 0 and 1 the Posterior Mean will act as a weighted average between our Prior Mean, $\frac{y}{n}$, and the data.

Part c

Posterior Distribution: $\theta|y \sim Beta(y+\alpha,(n-y)+\beta)$ Prior Variance $(\alpha=1,\beta=1)$: $\frac{1}{12}$ Posterior Variance $(\alpha=1,\beta=1)$:

$$\frac{(y+1)(n-y+1)}{(n+2)^2(n+3)} = \frac{ny-y^2+y+n-y+1}{(n^2+4n+4)(n+3)}$$
$$= \frac{ny-y^2+y+n-y+1}{n^3+3n^2+4n^2+12n+4n+12}$$
$$= \frac{-y^2+ny+n+1}{n^3+7n^2+16n+12}$$

Now, we can deduce that smaller values of n will maximize this quantity; since $n \geq y$, we will set n = y = 1.

Posterior Variance (n = y = 1):
$$\frac{-y^2 + ny + n + 1}{n^3 + 7n^2 + 16n + 12} = \frac{2}{36} = \frac{1}{18}$$

Thus, the Posterior Variance, which we just maximized, is always less than the Prior Variance of $\frac{1}{12}$.

Part d

- n = y = 1
- $\alpha = 1$
- $\beta = 10$

Prior Distribution: $\theta \sim Beta(\alpha, \beta)$

Posterior Distribution: $\theta | y \sim Beta(y + \alpha, (n - y) + \beta)$

Prior Variance: $\frac{(1)(10)}{(11)^2(12)} = \frac{10}{1452} = 0.0069$ Posterior Variance: $\frac{(2)(10)}{(12)^2(13)} = \frac{20}{1872} = 0.0107$

Problem 3

Part a

Prior Distribution: $\theta \sim Beta(\alpha, \beta)$

Prior Mean: 0.6 Prior Variance: 0.09

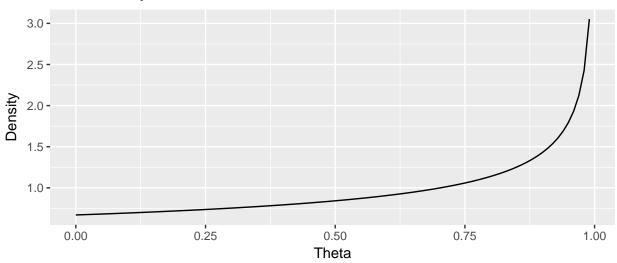
Sampling Distribution: $y|\theta \sim Binomial(n,\theta)$

$$\mu = \frac{\alpha}{\alpha + \beta}$$
$$(\alpha + \beta)\mu = \alpha$$
$$\alpha\mu + \beta\mu = \alpha$$
$$\beta\mu = \alpha - \alpha\mu$$
$$\beta = \alpha(\frac{1}{\mu} - 1)$$

$$\sigma^{2} = \frac{\alpha\beta}{(\alpha+\beta)^{2}(\alpha+\beta+1)}$$
$$\alpha = \left(\frac{1-\mu}{\sigma^{2}} - \frac{1}{\mu}\right)\mu^{2}$$

$$\alpha = 1$$
$$\beta = 0.67$$

Beta Density



Part b

n=1000

y = 650

Prior Distribution: $\theta \sim Beta(\alpha, \beta)$

Sampling Distribution: $y|\theta \sim Binomial(n, \theta)$

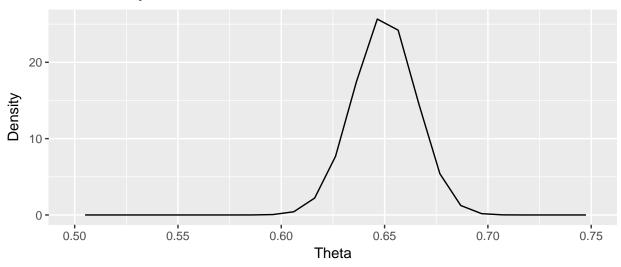
Posterior Density:

$$p(\theta|y) = \binom{n}{y} \theta^{y} (1-\theta)^{n-y} \times \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$
$$= \binom{n}{y} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{y+\alpha-1} (1-\theta)^{n-y+\beta-1}$$
$$\propto Beta(y+\alpha, (n-y)+\beta)$$

 $\alpha_{post} = y + \alpha$

 $\beta_{post} = (n - y) + \beta$ Posterior Mean: $\mu_{post} = \frac{y + \alpha}{(y + \alpha) + ((n - y) + \beta)} = 0.6499$ Posterior Variance: $\sigma_{post}^2 = \frac{(y + \alpha)((n - y) + \beta)}{((y + \alpha) + ((n - y) + \beta))^2((y + \alpha) + ((n - y) + \beta) + 1)} = 0.0151$

Beta Density



Part c

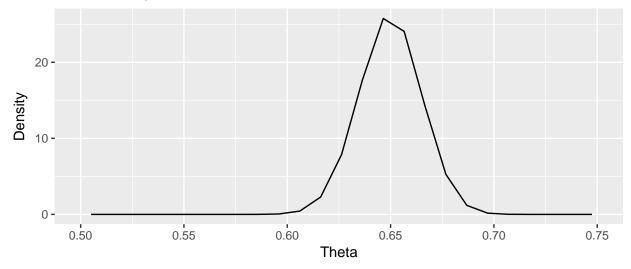
First Sensitivity Check - Uniform Prior

Prior Mean: $\frac{1}{2}$ Prior Variance: $\frac{1}{12}$ Prior Distribution: $\theta \sim Uniform(0,1)$ Sampling Distribution: $y|\theta \sim Binomial(n,\theta)$

Posterior Density:

$$p(\theta|y) = \binom{n}{y} \theta^y (1-\theta)^{n-y} \times 1$$
$$= \binom{n}{y} \theta^y (1-\theta)^{n-y} \times 1$$
$$\propto Beta(y+1, (n-y)+1)$$

Beta Density



Second Sensitivity Check - Beta Prior

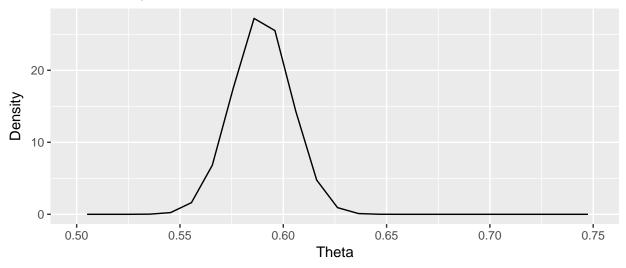
Prior Mean: 0.3 Prior Variance: 0.001

Prior Distribution: $\theta \sim Beta(62.7, 146.3)$ Sampling Distribution: $y|\theta \sim Binomial(n, \theta)$

Posterior Density:

$$p(\theta|y) = \binom{n}{y} \theta^{y} (1-\theta)^{n-y} \times \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$
$$= \binom{n}{y} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{y+\alpha-1} (1-\theta)^{n-y+\beta-1}$$
$$\propto Beta(y+\alpha, (n-y)+\beta)$$

Beta Density



Third Sensitivity Check - Beta Prior

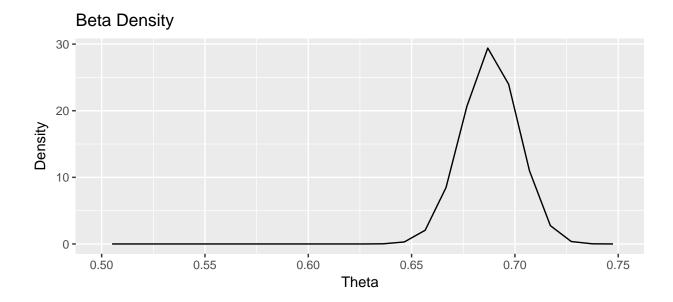
Prior Mean: 0.9

Prior Variance: 0.0005

Prior Distribution: $\theta \sim Beta(161.1, 17.9)$ Sampling Distribution: $y|\theta \sim Binomial(n, \theta)$

Posterior Density:

$$p(\theta|y) = \binom{n}{y} \theta^{y} (1-\theta)^{n-y} \times \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$
$$= \binom{n}{y} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{y+\alpha-1} (1-\theta)^{n-y+\beta-1}$$
$$\propto Beta(y+\alpha, (n-y)+\beta)$$



Conclusion

We can see that a prior distribution such as Uniform(0,1) has a negligible affect on the posterior distribution, whereas, our $Beta(\alpha,\beta)$ distributions with prior means and variances that are substantially different enough from the observed data will have a noticeable affect on the posterior distribution. This shows us how our prior distribution may have a large affect on our posterior distribution, depending on the family, and parameters of said family, that we select.

Code Appendix

```
## Load in necessary packages
library(ggplot2)
##### Problem 3 #####
## Set mean and variance
mu <- 0.6
var <- 0.09
##### Part a
## Estimate the beta parameters
est_beta_params <- function(mu, var) {</pre>
    alpha \leftarrow (((1 - mu) / var) - (1 / mu)) * mu^2
    beta <- alpha * ((1 / mu) - 1)
    return(round(c(alpha, beta), 2))
}
beta_params <- est_beta_params(mu, var)</pre>
## Establish the x-axis values
theta \leftarrow seq(0, 1, length.out = 100)
## Set the beta density
beta_density <- dbeta(x = theta,</pre>
                       shape1 = beta_params[1],
                       shape2 = beta_params[2])
beta_density <- beta_density[-length(beta_density)]</pre>
theta <- theta[-length(theta)]</pre>
## Build the density plot
ggplot() +
    geom_line(aes(x = theta, y = beta_density)) +
    labs(x = "Theta", y = "Density", title = "Beta Density")
```

```
##### Part b
## Set the sample size
n <- 1000
## Set the number of yes votes
y < -n * 0.65
## Set post alpha and beta
post_alpha <- y + beta_params[1]</pre>
post_beta <- (n - y) + beta_params[2]</pre>
## Posterior mean and variance
post_mean <- round(post_alpha / (post_alpha + post_beta), 4)</pre>
post_var <- round(sqrt((post_alpha * post_beta) /</pre>
                            ((post_alpha + post_beta)^2 * (post_alpha + post_beta + 1))), 4)
## Calculate the posterior density
post_density <- dbeta(x = theta,</pre>
                       shape1 = post_alpha,
                       shape2 = post_beta)
## Build the density plot
ggplot() +
    geom_line(aes(x = theta, y = post_density)) +
    labs(x = "Theta", y = "Density", title = "Beta Density") +
    xlim(c(0.5, 0.75))
##### Part c
## First sensitivity check
## Set mean and variance
mu <- 0.5
var < (1/12)
```

```
## Set post alpha and beta
post_alpha <- y + 1</pre>
post_beta \leftarrow (n - y) + 1
## Establish the x-axis values
theta \leftarrow seq(0, 1, length.out = 100)
## Set the beta density
beta_density <- dbeta(x = theta,</pre>
                        shape1 = post_alpha,
                        shape2 = post_beta)
beta_density <- beta_density[-length(beta_density)]</pre>
theta <- theta[-length(theta)]</pre>
## Build the density plot
ggplot() +
    geom_line(aes(x = theta, y = beta_density)) +
    labs(x = "Theta", y = "Density", title = "Beta Density") +
    xlim(c(0.5, 0.75))
## Second sensitivity check
## Set mean and variance
mu <- 0.3
var < -0.001
## Estimate the beta parameters
beta_params <- est_beta_params(mu, var)</pre>
## Set post alpha and beta
post_alpha <- y + beta_params[1]</pre>
post_beta <- (n - y) + beta_params[2]</pre>
## Establish the x-axis values
theta \leftarrow seq(0, 1, length.out = 100)
## Set the beta density
beta_density <- dbeta(x = theta,</pre>
                        shape1 = post_alpha,
```

```
shape2 = post_beta)
beta_density <- beta_density[-length(beta_density)]</pre>
theta <- theta[-length(theta)]</pre>
## Build the density plot
ggplot() +
    geom_line(aes(x = theta, y = beta_density)) +
    labs(x = "Theta", y = "Density", title = "Beta Density") +
    xlim(c(0.5, 0.75))
## Third sensitivity check
## Set mean and variance
mu <- 0.9
var < -0.0005
## Estimate the beta parameters
beta_params <- est_beta_params(mu, var)</pre>
## Set post alpha and beta
post_alpha <- y + beta_params[1]</pre>
post_beta <- (n - y) + beta_params[2]</pre>
## Establish the x-axis values
theta \leftarrow seq(0, 1, length.out = 100)
## Set the beta density
beta_density <- dbeta(x = theta,
                       shape1 = post_alpha,
                       shape2 = post_beta)
beta_density <- beta_density[-length(beta_density)]</pre>
theta <- theta[-length(theta)]</pre>
## Build the density plot
ggplot() +
    geom_line(aes(x = theta, y = beta_density)) +
    labs(x = "Theta", y = "Density", title = "Beta Density") +
    xlim(c(0.5, 0.75))
```