

# Chap 4

31. 设  $\{X_n, n = \dots, -1, 0, 1, \dots\}$  为平稳序列, 协方差函数为  $R(\tau)$ ,

(1) 求  $X_{n+1}$  的形如  $\hat{X}_{n+1}^{(1)} = aX_n$  的最小误差方差预报, 这里  $a$  是待定常数,

(2) 求  $X_{n+1}$  的形如  $\hat{X}_{n+1}^{(2)} = aX_n + bX_{n-1}$  的最小均方误差预报, 这里  $a$  和  $b$  是待定常数.

(1) 设  $\hat{X}^* = aX_n$  为  $X_{n+1}$  的最佳预报. 则由投影定理可知.  $E[(X_{n+1} - \hat{X}^*)X_n] = 0$

$$\begin{aligned} E[(X_{n+1} - \hat{X}^*)X_n] &= E(X_{n+1}X_n - aX_n^2) \\ &= E(X_{n+1} - \mu)(X_n - \mu) - aE(X_n - \mu)(X_n - \mu) \quad (\mu = 0) \\ &= R(1) - aR(0) = 0 \end{aligned}$$

$$\Rightarrow a = \frac{R(1)}{R(0)}$$

(2) 设  $\hat{X}^* = aX_n + bX_{n-1}$  为  $X_{n+1}$  的最佳预报. 则由投影定理可知.  $E[(X_{n+1} - \hat{X}^*)X_n] = 0$ ,  $E[(X_{n+1} - \hat{X}^*)X_{n-1}] = 0$

$$\begin{aligned} E[(X_{n+1} - \hat{X}^*)X_n] &= E(X_{n+1}X_n - aX_n^2 - bX_nX_{n-1}) \\ &= R(1) - aR(0) + bR(1) = 0 \end{aligned}$$

$$\Rightarrow R(0)a + R(1)b = R(1)$$

$$\begin{aligned} E[(X_{n+1} - \hat{X}^*)X_{n-1}] &= E(X_{n+1}X_{n-1} - aX_nX_{n-1} - bX_{n-1}^2) \\ &= R(2) - aR(1) - bR(0) \end{aligned}$$

$$\Rightarrow R(1)a + R(0)b = R(2)$$

$$\Rightarrow a = \frac{R(1)(R(0) - R(2))}{R^2(0) - R^2(1)} \quad b = \frac{R(2)R(0) - R^2(1)}{R^2(0) - R^2(1)}$$

34. 证明如下两个滑动平均序列

$$\begin{aligned} X_n &= \varepsilon_n + \alpha\varepsilon_{n-1}, \\ Y_n &= \varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1} \end{aligned}$$

有相同的自相关函数.

由题可知 自相关函数  $r(h) = E[X(t)X(t+h)] \quad t \in \mathbb{T}$

对于  $X_n = \varepsilon_n + \alpha\varepsilon_{n-1}$

$$h=0 \text{ 时 } R_X(0) = E[(\varepsilon_n + \alpha\varepsilon_{n-1})(\varepsilon_n + \alpha\varepsilon_{n-1})] = E[\varepsilon_n^2] + \alpha^2 E[\varepsilon_{n-1}^2] = (1 + \alpha^2)\sigma^2$$

$$h=1 \text{ 时 } R_X(1) = E[(\varepsilon_n + \alpha\varepsilon_{n-1})(\varepsilon_{n+1} + \alpha\varepsilon_n)] = \alpha E[\varepsilon_n^2] = \alpha\sigma^2$$

$$h \geq 2 \text{ 时 } R_X(h) = E[(\varepsilon_n + \alpha\varepsilon_{n-1})(\varepsilon_{n+h} + \alpha\varepsilon_{n+h-1})] = 0$$

$$\Rightarrow \rho_X(h) = \begin{cases} 1 & h=0 \\ \frac{\alpha}{1+\alpha^2} & h=1 \\ 0 & h \geq 2 \end{cases}$$

对于  $Y_n = \varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1}$

$$h=0 \text{ 时 } R_Y(0) = E[(\varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1})(\varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1})] = E[\varepsilon_n^2] + \frac{1}{\alpha^2} E[\varepsilon_{n-1}^2] = (1 + \frac{1}{\alpha^2})\sigma^2$$

$$h=1 \text{ 时 } R_Y(1) = E[(\varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1})(\varepsilon_{n+1} + \frac{1}{\alpha}\varepsilon_n)] = \frac{1}{\alpha} E[\varepsilon_n^2] = \frac{1}{\alpha}\sigma^2$$

$$h \geq 2 \text{ 时 } R_Y(h) = E[(\varepsilon_n + \frac{1}{\alpha}\varepsilon_{n-1})(\varepsilon_{n+h} + \frac{1}{\alpha}\varepsilon_{n+h-1})] = 0$$

$$\Rightarrow \rho_Y(h) = \begin{cases} 1 & h=0 \\ \frac{\alpha}{1+\alpha^2} & h=1 \\ 0 & h \geq 2 \end{cases}$$

$\therefore X$  与  $Y$  具有相同的自相关函数

35. 设  $\{X_n, n = 0, \pm 1, \dots\}$  为 AR(p) 模型:

$$X_n = \alpha_1 X_{n-1} + \dots + \alpha_p X_{n-p} + \varepsilon_n, \quad n = \dots, -1, 0, 1, \dots$$

试导出 Yule-Walker 方程:

$$R(h) = \alpha_1 R(h-1) + \dots + \alpha_p R(h-p), \quad h > 0.$$

提示: AR(p) 模型两边同乘以  $X_{n-k}$ , 然后取期望.

$$\text{由 } X_n = \alpha_1 X_{n-1} + \dots + \alpha_p X_{n-p} + \varepsilon_n$$

$$\text{可知 } X_n X_{n-k} = \alpha_1 X_{n-1} X_{n-k} + \dots + \alpha_p X_{n-p} X_{n-k} + \varepsilon_n X_{n-k}$$

$$\text{两边取期望, } E X_n X_{n-k} = \alpha_1 E X_{n-1} X_{n-k} + \dots + \alpha_p E X_{n-p} X_{n-k} + E \varepsilon_n X_{n-k}.$$

$$\text{记 } R(\tau) \text{ 是 } X_n \text{ 的协方差函数. 由 } X \text{ 是平稳过程. 则 } R(\tau) = E X_n X_{n+\tau}.$$

$$X_{n-k} = \sum_{i=0}^{\infty} \beta_i \varepsilon_{n-k-i} \Rightarrow E \varepsilon_n X_{n-k} = 0$$

$$\therefore R(k) = \alpha_1 R(k-1) + \dots + \alpha_p R(k-p) \quad \text{即 } R(h) = \alpha_1 R(h-1) + \dots + \alpha_p R(h-p)$$

37. 考虑如下 AR(2) 模型:

$$(1) X_n = 0.5X_{n-1} + 0.3X_{n-2} + \varepsilon_n,$$

试用 Yule-Walker 方程导出协方差函数, 证明它们的谱密度函数  $S(\omega)$  是周期函数, 并作出  $S(\omega)$  在  $(-\pi, \pi)$  上的图形.

由 Yule-Walker 方程可得

$$R(h) = 0.5R(h-1) + 0.3R(h-2)$$

$$\therefore R(0) = E[X_n^2] = E[(0.5X_{n-1} + 0.3X_{n-2} + \varepsilon_n)^2] = 0.34R(0) + 0.3R(1) + \sigma^2$$

$$R(1) = E[X_n X_{n-1}] = E[0.5X_{n-1}^2 + 0.3X_{n-2}X_{n-1} + \varepsilon_n X_{n-1}] = 0.5R(0) + 0.3R(1)$$

$$\Rightarrow R(0) = \frac{175}{78} \sigma^2 \quad R(1) = \frac{125}{78} \sigma^2$$

$$S_1(\omega) = \sum_{\tau=-\infty}^{+\infty} e^{-j\omega\tau} R(\tau) = 0.5 \sum_{\tau=-\infty}^{+\infty} e^{-j\omega(\tau+1)} R(\tau) + 0.3 \sum_{\tau=-\infty}^{+\infty} e^{-j\omega(\tau+2)} R(\tau)$$

$$S_1(2\pi + \omega) = \sum_{\tau=-\infty}^{+\infty} e^{-j(2\pi + \omega)\tau} R(\tau) = \sum_{\tau=-\infty}^{+\infty} e^{-j2\pi} e^{-j\omega\tau} R(\tau) = \sum_{\tau=-\infty}^{+\infty} e^{-j\omega\tau} R(\tau) = S_1(\omega)$$

其中  $e^{-j2\pi} = \cos 2\pi - j\sin 2\pi = 1$  故  $S_1(\omega)$  具有周期性.

2. 设  $\{W(t), t \geq 0\}$  为 Brown 运动, 验证下列过程仍为  $[0, \infty)$  上的 Brown 运动.

(1)  $T(t) = tW(1/t)$ ,

(2)  $W(t) = W(a^2 t)/a, a > 0$ .

$$(1) T(t) = tW(\frac{1}{t})$$

$$\textcircled{1} \lim_{t \rightarrow 0} T(t) = 0$$

$$\textcircled{2} T(t+s) - T(s) = (t+s)W(\frac{1}{t+s}) - sW(\frac{1}{s}) \\ = tW(\frac{1}{t+s}) + s[W(\frac{1}{t+s}) - W(\frac{1}{s})]$$

$$tW(\frac{1}{t+s}) \sim N(0, \frac{t^2}{t+s})$$

$$s[W(\frac{1}{t+s}) - W(\frac{1}{s})] \sim N(0, \frac{ts}{t+s})$$

$$\therefore T(t+s) - T(s) \sim N(0, t)$$

③  $T(t)$  具有独立增量.

$$(2) Y(t) = \frac{1}{a} W(a^2 t)$$

$$\textcircled{1} Y(0) = 0.$$

$$\textcircled{2} Y(t+s) - Y(s) = \frac{1}{a} W(a^2(t+s)) - \frac{1}{a} W(a^2 s) \\ = \frac{1}{a} (W(a^2(t+s)) - W(a^2 s)) \\ \sim \frac{1}{a} N(0, a^2 t) \sim N(0, t)$$

③  $Y(t)$  具有独立增量.

则  $Y(t) = \frac{1}{a} W(a^2 t)$  为 Brown 运动

5. 设  $\{Z(t), t \geq 0\}$  为 Brown 桥, 验证  $W(t) = (t+1)Z(\frac{t}{t+1})$  是 Brown 运动.

$$EZ(t) = 0.$$

$$EW(t) = (t+1)E[Z(\frac{t}{t+1})] = 0$$

$$\therefore E[W(t)W(s)] = (t+1)(s+1)E[Z(\frac{t}{t+1})Z(\frac{s}{s+1})] \\ = (t+1)(s+1)(1 - \frac{t}{t+1})\frac{s}{s+1} \\ = s$$

$\therefore W(t)$  是 Brown 运动

