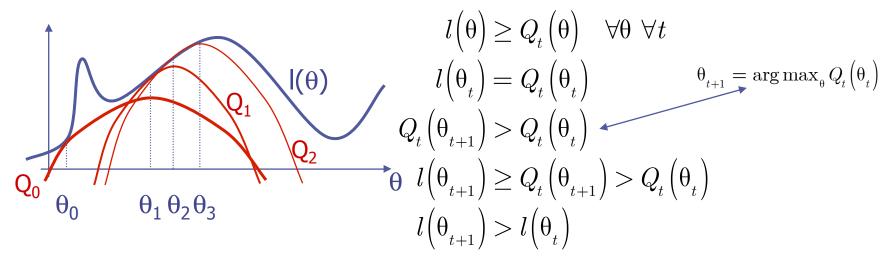
Machine Learning

Topic 13

- Expectation Maximization as Bound Maximization
- •EM for Maximum A Posteriori

EM as Bound Maximization

- Let's now show that EM indeed maximizes likelihood
- •Bound Maximization: optimize a lower bound on $I(\theta)$
- •Since log-likelihood $I(\theta)$ not concave, can't max it directly
- •Consider an auxiliary function $Q(\theta)$ which is concave
- •Q(θ) kisses I(θ) at a point and is less than it elsewhere



- Monotonically increases log-likelihood
- •But how to find a bound and guarantee we max it?

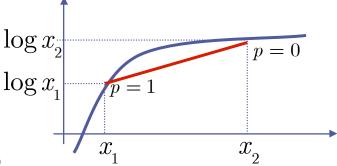
Jensen's Inequality



- •An important general bound from Jensen (1906)
- •For convex f: $f(E\{x\}) \le E\{f(x)\}$ •For concave f: $f(E\{x\}) \ge E\{f(x)\}$
- Expectation in discrete case is sum weight by probability
- •For convex f: $f\left(\sum_{i=1}^{M}p_{i}x_{i}\right) \leq \sum_{i=1}^{M}p_{i}f\left(x_{i}\right) \ when \ \sum_{i=1}^{M}p_{i}=1, \ p_{i}\geq 0$ •For concave f: $f\left(\sum_{i=1}^{M}p_{i}x_{i}\right) \geq \sum_{i=1}^{M}p_{i}f\left(x_{i}\right) \ when \ \sum_{i=1}^{M}p_{i}=1, \ p_{i}\geq 0$
- •Example: f(x) = log(x) = concave and M=2

$$\log \left(px_{_{\! 1}} + \left(1-p\right)x_{_{\! 2}}\right) \geq p\log x_{_{\! 1}} + \left(1-p\right)\log x_{_{\! 2}} \ \log x_{_{\! 2}}$$

Bound log(sum) with sum(log)



•How to apply this to mixture models?

Expectation-Maximization

EM as Bound Maximization

•Now have the following bound and maximize it:

$$\begin{split} l\left(\theta\right) &\geq Q\left(\theta \mid \theta_{t}\right) - \sum\nolimits_{n=1}^{N} \sum\nolimits_{z} p\left(z \mid x_{n}, \theta_{t}\right) \log p\left(z \mid x_{n}, \theta_{t}\right) \\ \theta^{t+1} &= \arg \max_{\theta} Q\left(\theta \mid \theta_{t}\right) = \arg \max_{\theta} \sum\nolimits_{n=1}^{N} \sum\nolimits_{z} p\left(z \mid x_{n}, \theta_{t}\right) \log p\left(x_{n}, z \mid \theta\right) \\ &= \arg \max_{\theta} \sum\nolimits_{n=1}^{N} \sum\nolimits_{z} \tau_{n, z} \log p\left(x_{n}, z \mid \theta\right) \end{split}$$

- $\bullet Q(\theta | \theta_t)$ is called Auxiliary Function... take derivatives of it
- •This is easy for e-families... just weighted max likelihood!
- For example, Gaussian mixture:

EM as Expected Log-Likelihood

IncompleteLog-Likelihood

$$l\!\left(\theta\right) = \log p\!\left(observed \mid \theta\right) = \sum\nolimits_{n=1}^{N} \log \sum\nolimits_{z} p\!\left(x_{n}, z \mid \theta\right)$$

CompleteLog-Likelihood

$$l^{\scriptscriptstyle C}\!\left(\theta\right) = \log p\!\left(observed, hidden \mid \theta\right) = \sum\nolimits_{n=1}^{N} \log p\!\left(x_{_{\!n}}, z_{_{\!n}} \mid \theta\right)$$

- We don't know the hidden variables z
- •EM computes expected values of hidden z under current θ_t
- EM chooses Q to be the Expected Complete Log-Likelihood

$$\begin{split} E\left\{l^{C}\left(\theta\right)\right\} &= \sum\nolimits_{hidden} p\left(hidden\mid observed, \theta_{t}\right) l^{C}\left(\theta\right) \\ &= \sum\nolimits_{z_{1}} \cdots \sum\nolimits_{z_{N}} p\left(z_{1}, \ldots, z_{n} \mid x_{1}, \ldots, x_{n}, \theta_{t}\right) l^{C}\left(\theta\right) \\ &= \sum\nolimits_{z_{1}} \cdots \sum\nolimits_{z_{N}} \prod\nolimits_{n} p\left(z_{n} \mid x_{n}, \theta_{t}\right) l^{C}\left(\theta\right) \\ &= \sum\nolimits_{z_{1}} \cdots \sum\nolimits_{z_{N}} \prod\nolimits_{n} p\left(z_{n} \mid x_{n}, \theta_{t}\right) \sum\nolimits_{n} \log p\left(x_{n}, z_{n} \mid \theta\right) \\ &= \sum\nolimits_{z_{1}} \sum\nolimits_{z_{N}} p\left(z_{n} \mid x_{n}, \theta_{t}\right) \log p\left(x_{n}, z_{n} \mid \theta\right) \sum\nolimits_{z_{1}} \cdots \sum\nolimits_{z_{i \neq n}} \cdots \sum\nolimits_{z_{N}} \prod\nolimits_{i \neq n} p\left(z_{i} \mid x_{i}, \theta_{t}\right) \\ &= \sum\nolimits_{n} \sum\nolimits_{z} p\left(z_{n} \mid x_{n}, \theta_{t}\right) \log p\left(x_{n}, z_{n} \mid \theta\right) = Q\left(\theta \mid \theta_{t}\right) \end{split}$$

EM for Max A Posteriori

- •We can also do MAP instead of ML with EM (stabilizes sol'n) $\log posterior(\theta) = \sum_{n=1}^{N} \log \sum_{z} p(x_n, z \mid \theta) + \log p(\theta)$
- Prior doesn't have log-sum
- •The E-step remains the same: lower bound log-sum $\log posterior(\theta) = l(\theta) + \log p(\theta) \ge E\{l^{C}(\theta)\} + const + \log p(\theta)$
- •The M-step becomes slightly different for each model
- •For example, mixture of Gaussians with prior on covariance

$$\begin{split} &\log posterior \left(\boldsymbol{\theta} \right) = \sum\nolimits_{n = 1}^N {\log \sum\nolimits_k {\pi _k} N{\left({{\vec{x}_n} \mid {\vec{\mu }_k},{\Sigma _k}} \right)} + \log \prod\nolimits_k {p\left({{\Sigma _k} \mid S,\eta } \right)} \\ &\log posterior{\left(\boldsymbol{\theta} \right)} \ge \sum\nolimits_{n = 1}^N {\sum\nolimits_k {\tau _{n,k}} \log {\pi _k} N{\left({{\vec{x}_n} \mid {\vec{\mu }_k},{\Sigma _k}} \right)} + \sum\nolimits_k \log p{\left({{\Sigma _k} \mid S,\eta } \right)} + const \end{split}$$

•Updates on π and μ stay the same, only Σ is:

$$\Sigma_k \leftarrow \frac{1}{\sum_{n=1}^N \tau_{n,k} + \eta} \left(\sum_{n=1}^N \tau_{n,k} \left(\vec{x}_n - \vec{\mu}_k \right) \left(\vec{x}_n - \vec{\mu}_k \right)^T + \eta S \right)$$

•Typically, we use the identity matrix I for S and a small eta.