

Variational Inference: An Introduction

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November 18, 2019

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References

- Maximum Likelihood from Incomplete Data via the EM Algorithm, Dempster et al., JRSSB, 1977. [paper]
- Main reference: **Pattern recognition and machine learning, C. Bishop, Springer, 2006. [Chapter 9, 10]**
- Elements of statistical learning, R. Tibshirani et al, Springer, 2009. [Chapter 8]
- Modern multivariate statistical techniques, A. Izenman, Springer, 2013. [Chapter 12]
- Numerical Analysis for Statisticians, K. Lange, Springer, 2010. [Chapter 12]

Model set-up

Suppose the following:

- $\mathbf{X} \in \mathbb{R}^{n \times p}$: observed variables.
- $\mathbf{Z} \in \mathbb{R}^{n \times k}$: latent variables.
- $\boldsymbol{\beta}$: parameters of interest.
- $\log p(\mathbf{X}|\boldsymbol{\beta})$: observed log-likelihood.

Lower bound of log-likelihood

Denote $Q(\mathbf{Z})$ as any positive measurable pmf/pdf of \mathbf{Z} , we have

$$\begin{aligned}\log p(\mathbf{X}|\boldsymbol{\beta}) &= \log \sum_{\mathbf{z}} p(\mathbf{X}, \mathbf{z}|\boldsymbol{\beta}) \\ &= \log \left[\sum_{\mathbf{z}} \frac{p(\mathbf{X}, \mathbf{z}|\boldsymbol{\beta})}{Q(\mathbf{z})} Q(\mathbf{z}) \right] \\ &= \log \mathbb{E}_{\mathbf{Z}} \left[\frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})}{Q(\mathbf{Z})} \right] \\ &\geq \mathbb{E}_{\mathbf{Z}} [\log \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})}{Q(\mathbf{Z})}]\end{aligned}$$

By Jensen's Inequality.

Note that the lower bound is attained iff $\frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})}{Q(\mathbf{Z})}$ is a constant w.r.t. \mathbf{Z} (it can depend on \mathbf{X}).

Derivation of $Q(\mathbf{Z})$

Assume $k = \frac{p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\beta})}{Q(\mathbf{Z})}$, then

$$\begin{aligned} k &= \int_{\Omega} k Q(\mathbf{Z}) d\mathbf{z} \\ &= \int_{\Omega} p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\beta}) d\mathbf{z} \\ &= p(\mathbf{X} | \boldsymbol{\beta}) \end{aligned}$$

Therefore, the lower bound of $\log p(\mathbf{X} | \boldsymbol{\beta})$ is attained if we set $Q(\mathbf{Z})$ to be the posterior distribution of \mathbf{Z} , i.e.

$$Q(\mathbf{Z}) = p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\beta})$$

MM Algorithm

- E-step: At t^{th} iteration, compute the lower bound of $\log p(\mathbf{X}|\boldsymbol{\beta})$, which is

$$\begin{aligned} & \mathbb{E}_{\mathbf{Z}|\mathbf{X}, \boldsymbol{\beta}^{(t)}} \left[\log \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})}{p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\beta}^{(t)})} \right] \\ &= \mathbb{E}_{\mathbf{Z}|\mathbf{X}, \boldsymbol{\beta}^{(t)}} [\log p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})] + H(p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\beta}^{(t)})) \end{aligned}$$

Where $H(\cdot)$ denotes the entropy of a distribution.

- M-step: Note that $H(\cdot)$ does not involve $\boldsymbol{\beta}$, thus maximizing $\log p(\mathbf{X}|\boldsymbol{\beta})$ is equivalent to maximize the first conditional expectation:

$$\boldsymbol{\beta}^{(t+1)} = \arg \max_{\boldsymbol{\beta}} \mathbb{E}_{\mathbf{Z}|\mathbf{X}, \boldsymbol{\beta}^{(t)}} [\log p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\beta})]$$

EM as variational inference

We can rewrite

$$\log p(\mathbf{X}) = \mathcal{L}(q) + KL(q\|p)$$

where we have defined

$$\mathcal{L}(q) = \int_{\Omega} q(\mathbf{Z}) \log \left\{ \frac{p(\mathbf{X}, \mathbf{Z})}{q(\mathbf{Z})} \right\} d\mathbf{Z}$$

$$KL(q\|p) = - \int_{\Omega} q(\mathbf{Z}) \log \left\{ \frac{p(\mathbf{Z}|\mathbf{X})}{q(\mathbf{Z})} \right\} d\mathbf{Z}$$

Note that we absorb β into \mathbf{Z} . Also, $\mathcal{L}(q)$ is a lower bound of $\log p(\mathbf{X})$ and such bound is attained iff

$$KL(q\|p) = 0 \text{ iff } q(\mathbf{Z}) = p(\mathbf{Z}|\mathbf{X}) \text{ a.e.}$$

Intractable posterior distribution

What can we do if $p(\mathbf{Z}|\mathbf{X})$ is intractable ?

- Sampling methods via Markov Chain Monte Carlo (MCMC)
- Consider a restricted family of distributions $q(\mathbf{Z})$ and then seek the member of this family for which the KL divergence is minimized.

Example 1 (parametric models as restricted families)

We assume $q(\mathbf{Z})$ can be specified by a set of parameters ω . Then the lower bound $\mathcal{L}(q)$ becomes a function of ω .

Factorized distributions

- **Mean field theory:** we can partition \mathbf{Z} into M disjoint groups, that is,

$$q(\mathbf{Z}) = \prod_{i=1}^M q_i(\mathbf{Z}_i)$$

- **Rewrite $\mathcal{L}(q)$ as a function of $q_i(\mathbf{Z}_i)$:**

$$\begin{aligned}\mathcal{L}(q) &= \int_{\Omega} q(\mathbf{z}) \log \left\{ \frac{p(\mathbf{X}, \mathbf{z})}{q(\mathbf{z})} \right\} d\mathbf{z} \\ &= \int_{\Omega} \prod_{i=1}^M q_i[\log \mathbb{P}(\mathbf{X}, \mathbf{z}) - \sum_{i=1}^M \log q_i] d\mathbf{z}\end{aligned}$$

Where $q_i = q_i(\mathbf{Z}_i)$ in short.

Factorized distributions cont'd

$$\begin{aligned}
 \mathcal{L}(q) &= \int_{\Omega} \prod_{i=1}^M q_i [\log \mathbb{P}(\mathbf{X}, \mathbf{Z}) - \sum_{i=1}^M \log q_i] d\mathbf{Z} \\
 &= \int q_j \left(\int \prod_{i \neq j} q_j \log p(\mathbf{X}, \mathbf{Z}) d\mathbf{z}_{\{-j\}} \right) d\mathbf{z}_j \\
 &\quad - \int \prod_{i=1}^M q_i \left(\sum_{i=1}^M q_i \right) d\mathbf{z}
 \end{aligned}$$

Key step: the first can be rewritten as a probability density plus a constant and the second term can be considered as the entropy of $p(\mathbf{z}_j)$ plus a constant (see next slide for detail).

Factorized distributions: Minimizing KL divergence

- First term can be written as (left as an exercise):

$$\int q_j \mathbb{E}_q(\log p(\mathbf{X}, \mathbf{Z}) | \mathbf{Z}_{\{-j\}}) d\mathbf{z}_j = \int q_j \log \tilde{p}(\mathbf{X}, \mathbf{z}_j) d\mathbf{z}_j + C_1$$

Where $\tilde{p}(\mathbf{X}, \mathbf{z}_j)$ is a new pdf w.r.t. \mathbf{z}_j and C_1 is a constant.

- Second term can be written as (left as an exercise):

$$\int q_j \log q_j d\mathbf{z}_j + C_2$$

Where C_2 is also a constant.

Minimizing KL divergence

Therefore, taking $\mathcal{L}(q)$ as a function of q_j , we have

$$\mathcal{L}(q) = -KL(q_j \parallel \tilde{p}(\mathbf{X}, \mathbf{z}_j)) + C$$

Where C is a constant. Thus, maximizing $\mathcal{L}(q)$ is equivalent to minimizing KL divergence which occurs when:

$$q_j(\mathbf{z}_j) = \tilde{p}(\mathbf{X}, \mathbf{z}_j))$$

In fact, we have more (but this is not an explicit solution):

$$q_j^*(\mathbf{z}_j) = \tilde{p}(\mathbf{X}, \mathbf{z}_j)) = \frac{\exp\{\mathbb{E}_q(\log p(\mathbf{X}, \mathbf{Z}) | \mathbf{Z}_{\{-j\}})\}}{\int \exp\{\mathbb{E}_q(\log p(\mathbf{X}, \mathbf{Z}) | \mathbf{Z}_{\{-j\}})\} d\mathbf{z}_j}$$

2 components factorized Gaussian

Example 2

- Suppose $\mathbf{Z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Lambda}^{-1})$ where

$$\boldsymbol{\mu} = (\mu_1^T, \mu_2^T)^T, \boldsymbol{\Lambda} = \begin{bmatrix} \Lambda_{11} & \Lambda_{12} \\ \Lambda_{21} & \Lambda_{22} \end{bmatrix}$$

- Goal: find $q_1^*(Z_1)$ where

$$\log q_1^*(Z_1) = \mathbb{E}_{q_2}(\log \mathbb{P}(\mathbf{X}, \mathbf{Z}) | Z_1) + C$$

- Solution: the log likelihood is

$$-\frac{1}{2} Z_2^T \Lambda_{11} Z_1 + \mu_1^T \Lambda_{11} Z_1 - (\mathbb{E}_{q_2}(Z_2) - \mu_2)^T \Lambda_{21} Z_1 + C$$

2 components factorized Gaussian

Example 3 (cont'd)

- Recall solution:

$$-\frac{1}{2}Z_2^T \Lambda_{11} Z_1 + \mu_1^T \Lambda_{11} Z_1 - (\mathbb{E}_{q_2}(Z_2) - \mu_2)^T \Lambda_{21} Z_1 + C$$

- This is a quadratic form $\Rightarrow q_1^*(Z_1)$ is Gaussian

$$\mathcal{N}(Z_1 | m_1, \Lambda_{11}^{-1})$$

where

$$m_1 = \mu_1 - \Lambda_{11}^{-1} \Lambda_{12} (\mathbb{E}_{q_2}(Z_2) - \mu_2)$$

2 components Gaussian

Example 4 (cont'd)

- Besides, knowing that $\mathbb{E}_{q_2}(Z_2) = m_2$ and $\mathbb{E}_{q_1}(Z_1) = m_1$:

$$m_1 = \mu_1, m_2 = \mu_2$$

- Solution:

$$q_1^*(Z_1) = \mathcal{N}(Z_1 | \mu_1, \Lambda_{11}^{-1})$$

$$q_2^*(Z_2) = \mathcal{N}(Z_2 | \mu_2, \Lambda_{22}^{-1})$$

Minimizing the reverse KL divergence

Previously, we are minimizing $KL(q_j \parallel \tilde{p}(\mathbf{X}, Z_j))$. Now suppose we want to minimize

$$KL(\tilde{p}(\mathbf{X}, Z_j) \parallel q_j)$$

Which is

$$KL(p \parallel q) = - \int p(\mathbf{z}) \left(\sum_{j=1}^M \log q_j(Z_j) \right) d\mathbf{z} + C$$

This gives us

$$q_1^*(Z_1) = p(Z_1) = \mathcal{N}(\mu_1, \Sigma_{11})$$

$$q_2^*(Z_2) = p(Z_2) = \mathcal{N}(\mu_2, \Sigma_{22})$$

Sidenotes

- Factorized VI leads to estimation that is too compact (i.e. variance is too concentrated).
- Minimizing reversed KL leads to too broad estimation (i.e. variance is too high).

Sidenotes

Example 5 (α -family of divergence)

$$D_\alpha(p\|q) = \frac{4}{1-\alpha^2} \left(1 - \int p(x)^{\frac{1+\alpha}{2}} q(x)^{\frac{1-\alpha}{2}} dx\right)$$

- $\alpha = 0$: Hellinger distance: $D_H(p\|q) = \int (p^{\frac{1}{2}} - q^{\frac{1}{2}})^2 dx$
- $\alpha \rightarrow 1$: $KL(p\|q)$: zero-forcing
- $\alpha \rightarrow -1$: $KL(q\|p)$: zero-avoiding