

---

---

---

---

---



# Lec 13

## Bayesian Est. in LM

$$\mathbb{E} Y = X\beta, \text{Cov}(e) = \sigma^2 I$$

70-80s: Frequentist vs. Bayesian

$f(y|\theta)$  density;  $g(\theta)$  prior density

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} \ln L(y|\theta), \quad L(y|\theta) = \prod_{i=1}^n f(y_i|\theta)$$

Posterior  $p(\theta|y) = f(y, \theta)/f(y) \propto f(y|\theta)g(\theta)$

(Ex) Let  $y_1, \dots, y_n \sim N(\mu, \sigma^2)$ ,  $\theta = (\mu, \sigma^2)$ .

$$L(y|\theta) = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n e^{-\frac{1}{2\sigma^2} \sum_i (y_i - \mu)^2} \\ = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^{\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \left\{ \sum_i (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2 \right\}}$$

If  $\sigma$  is known,

$$L(y|\mu) \propto e^{-\frac{n}{2\sigma^2}(\bar{y} - \mu)^2}$$

$$\text{Assume } f(\mu) \propto e^{-\frac{1}{2\sigma^2}(\mu - \mu_0)^2}$$

Posterior density is

$$p(\mu|y) \propto e^{-\frac{n}{2\sigma^2}(\bar{y} - \mu)^2} e^{-\frac{1}{2\sigma^2}(\mu - \mu_0)^2} \\ = \exp \left\{ -\frac{n}{2\sigma^2}(\bar{y} - \mu)^2 - \frac{1}{2\sigma^2}(\mu - \mu_0)^2 \right\}$$

Let's see if  $\exists \mu_n$  &  $\sigma_n^2$  s.t. RHS is  $\propto$

$$\exp \left\{ -\frac{1}{2\sigma_n^2}(\mu - \mu_n)^2 \right\}$$

Equate coefficients:

$$\mu^2: -\frac{1}{2\sigma_n^2} = -\frac{1}{2\sigma^2} - \frac{n}{2\sigma^2}$$

$$\sigma_n^2 = \left( \frac{1}{\sigma^2} + \frac{n}{\sigma^2} \right)^{-1}$$

$$\mu: \frac{\mu_n}{\sigma_n^2} = \frac{\mu_0}{\sigma^2} + \frac{\bar{x}_n}{\sigma^2}$$

$$\Rightarrow \mu_n = \frac{\sigma^2}{\sigma^2 + n\sigma_0^2} \mu_0 + \frac{n\sigma_0^2}{\sigma^2 + n\sigma_0^2} \bar{x}$$

posterior density:  $p(\mu|y) \propto N(\mu_n, \sigma_n^2)$

(Ex)  $Y \sim \text{Bin}(n, \theta)$ ,  $P(Y=y) = \binom{n}{y} \theta^y (1-\theta)^{n-y}$

Let  $g(\theta)$  be a prior density for  $\theta$ .

Posterior for  $\theta \propto L(y|\theta)g(\theta)$

$$\text{If } g(\theta) = \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

$$\text{Posterior is } \propto \frac{\theta^{y+\alpha-1} (1-\theta)^{n-y+\beta-1}}{B(\alpha, \beta)}$$

Relative mean & variance of the Beta-Binomial r.v.

$$\mathbb{E} Y = X\beta, \text{Var}(e) = \sigma^2 I, \quad \theta^T = (\beta^T, \sigma^2)$$

$$g(\theta) = g(\beta, \sigma^2) = g_1(\beta) g_2(\sigma^2)$$

Preliminaries:

$$(a) \int_0^\infty e^{-\frac{k}{x}} x^{-v-1} dx = \frac{1}{k^v \Gamma(v)}$$

$$(b) \int_0^\infty e^{-\frac{a}{x}} x^{-b-1} dx = \frac{\Gamma(b/a)}{2a^{b/2}}$$

$$(c) \|Y - X\beta\|_2^2 = (n-p)S^2 + \|X(\beta - \bar{\beta})\|_2^2$$

$$Y = X\beta + \varepsilon, \quad \varepsilon \sim N(0, \sigma^2 I)$$

$$f(\theta|y) \propto f(y|\theta)g(\theta)$$

$$\text{what is } g(\theta): \underbrace{g_1(\beta)g_2(\sigma^2)}_{\text{constant } \propto \frac{1}{\theta}}$$

$$f(\theta|y) \propto \frac{1}{\theta} \left( \frac{1}{\sqrt{2\pi}\sigma} \right)^n e^{-\frac{1}{2\sigma^2} \|Y - X\beta\|_2^2}$$

If interest is in  $\beta$  alone,

$$f(\beta|y) = \int_0^\infty \frac{1}{\theta^n} \frac{1}{\theta^{\frac{n}{2}}} e^{-\frac{1}{2\theta^2} \|Y - X\beta\|_2^2} d\theta$$

$$\stackrel{(b)}{=} \frac{1}{2} \left( \frac{\|Y - X\beta\|_2^2}{2} \right)^{-\frac{n}{2}} \Gamma\left(\frac{n}{2}\right)$$

$$\Rightarrow f(\beta|y) \propto \|Y - X\beta\|_2^{-n}$$

$$\stackrel{(c)}{=} ((n-p)S^2 + \|X(\beta - \bar{\beta})\|_2^2)^{-\frac{n}{2}}$$

$$\propto \left( 1 + \frac{(X(\beta - \bar{\beta})^T X(\beta - \bar{\beta}))}{(n-p)S^2} \right)^{-\frac{n}{2}}$$

Multivariate t-distribution

Page 475 (Lee & Seber)

Write  $Y \sim t_m(v, \mu, \Sigma)$  if

$$f_Y(y) = \frac{\Gamma(\frac{v+m}{2})}{(\pi v)^{\frac{m}{2}} \Gamma(\frac{v}{2})} \frac{1}{|\Sigma|^{\frac{1}{2}}} \left( 1 + \frac{(y-\mu)^T \Sigma^{-1} (y-\mu)}{v} \right)^{-\frac{v+m}{2}}$$

$$\mathbb{E} Y = \mu, \text{ Cov } Y = \Sigma.$$

$$\mathbb{E}(\mathbb{E}(Y|X)) = Y; \text{ Var } Y = \mathbb{E}(\text{Var}(Y|X)) + \text{Var}(\mathbb{E}(Y|X))$$

Before we worked with improper prior densities. Now work with proper priors

$$\Theta^T = (\beta^T, \sigma^2)$$

$$f(\beta, \sigma^2) = \underbrace{f_1(\beta|\sigma^2)}_{\mathcal{N}_p(m, \sigma^2 V)} f_2(\sigma^2)$$

**IG**  $\mathcal{N}_p(m, \sigma^2 V)$  Inverted Gamma

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x}$$

$$f(\theta|y) \propto f(y|\theta)g(\theta)$$

$$\propto (\sigma^2)^{-\frac{n+d+2+p}{2}} \exp\left(-\frac{1}{2\sigma^2} [(y - X\beta)^T (y - X\beta) + (\beta - m)^T V(\beta - m)]\right)$$

where  $f(\sigma^2) \propto \frac{1}{(\sigma^2)^{d+2}} e^{-\frac{Q}{2\sigma^2} + a}$

$$\text{Let } Q = (y - X\beta)^T (y - X\beta) + (\beta - m)^T V^{-1}(\beta - m)$$

$$f(\beta|y) \propto \int_0^\infty (\sigma^2)^{-\frac{n+d+2+p}{2}} e^{-\frac{1}{2\sigma^2}(Q+a)} d\sigma^2$$

$$\propto (Q+a)^{-\frac{d+n+p}{2}} \quad (\text{verify})$$

$$\propto \left(1 + \frac{1}{n+d} (\beta - m_*)^T W_*^{-1} (\beta - m_*)\right)^{-\frac{d+n+p}{2}}$$

(Page 76)

$$\beta|y \sim t_p(n+d, m_*, W_*) \quad \square$$