


Review:

$$Y = X\beta + \varepsilon, \quad r(X) = P, \\ n \times p \quad E\varepsilon = 0, \quad \text{Var } \varepsilon = \sigma^2 V.$$

$$\Rightarrow \hat{\beta}_w = (X^T V^{-1} X)^{-1} X^T V^{-1} Y$$

$$\begin{aligned} Y &= X\beta + Z\gamma + \varepsilon \\ &= (X \ Z) \begin{pmatrix} \beta \\ \gamma \end{pmatrix} + \varepsilon \end{aligned}$$

γ components are random.

$$\text{Cov } \gamma = D; \quad \text{Cov } \varepsilon = R$$

$$\text{Cov}(\gamma, \varepsilon) = 0.$$

(Ex) 1-way Anova with 3 total & 2 reps:

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}$$

$$Y = \begin{pmatrix} Y_1 \\ \vdots \\ Y_6 \end{pmatrix}, \quad X\mu = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}, \quad \mu = J_6 \mu, \quad \beta = \mu.$$

$$Z = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}; \quad \gamma = \begin{pmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{pmatrix}$$

We have 1-way balanced ANOVA random effects model in a mat. form.

$$\begin{aligned} f(y, r) &= f_{Y|r} f_r \\ &= \frac{1}{|R|^{\frac{1}{2}}} \frac{1}{(2\pi)^p} e^{-\frac{1}{2}(y - X\beta - Z\gamma)^T R^{-1} (y - X\beta - Z\gamma)} \\ &\quad \cdot e^{-\frac{1}{2}(r - D)^T D^{-1} (r - D)} \frac{1}{|D|^{\frac{1}{2}} (2\pi)^q} \end{aligned}$$

- Verify the MLEs for β & γ are

$$(1) X^T R^{-1} X \beta + X^T R^{-1} Z \gamma = X^T R^{-1} Y$$

$$(2) Z^T R^{-1} X \beta + (Z^T R^{-1} Z + D^{-1}) \gamma = Z^T R^{-1} Y$$

Henderson's equations.

Re-parametrize as follows (when appropriate):

$$\begin{aligned} X^T R^{-1} X \beta + X^T R^{-1} Z D (D^{-1} \gamma) &= X^T R^{-1} Y \\ Z^T R^{-1} X \beta + (I + Z^T R^{-1} Z D) D^{-1} \gamma &= Z^T R^{-1} Y \end{aligned}$$

Sols provide est. for β & $\alpha = D^{-1} \gamma$.

What are solns $\hat{\beta}$ & $\hat{\gamma}$?

$$-\hat{\beta}_w = (X^T V^{-1} X)^{-1} X^T V^{-1} Y, \quad V = Z D Z^T + R \quad (\text{cov}(Y))$$

use

$$V^{-1} = R^{-1} Z (D^{-1} + Z^T R^{-1} Z)^{-1} Z^T R^{-1}$$

$$-\hat{\gamma} = D Z^T V^{-1} (Y - X\hat{\beta})$$

Note: $\hat{\beta}_w$ is unbiased for β & $E\hat{\gamma} = 0$.

• Suppose u is a RV with mean 0 & finite variance.

• A linear predictor $d + a^T y$ of $C^T \beta + u$ is unbiased iff

$$E(d + a^T y) = E(C^T \beta + u) = C^T \beta,$$

or $C^T \beta + u$ is predictable iff

\exists an unbiased linear predictor of $C^T \beta + u$

Claim: $C^T \beta + u$ is predictable iff

$$\exists \text{ s.t. } C = X^T a$$

This implies $d + a^T y$ is a LUP of $C^T \beta + u$ iff $d = 0$, $C = X^T a$.

Recall (MSE) of $\hat{\theta}$ of θ is

$$\mathbb{E}(\hat{\theta} - \theta)^2.$$

$$MSE(\hat{\theta}) = Var(\hat{\theta}) + (Bias(\hat{\theta}))^2$$

Claim: $C^T \hat{\beta}_w + \hat{u}$ has the smallest MSE among all LUE of $C^T \beta + u$,

where $\hat{u} = \underbrace{Cov(\varepsilon, u)}_{\text{vector}} \underbrace{(Var \varepsilon)^{-1}}_{\text{matrix}} (y - X \hat{\beta})$

$$\hat{\beta}_w = (X^T V^{-1} X)^{-1} X^T V^{-1} y$$

Assume that $Cov(\varepsilon, u) = \sigma^2 K$.

K is a known vector.

Goal: Predict $C^T \beta + u$ (c known) accurately

Note that

$$C^T \hat{\beta} + \hat{u} = C^T \hat{\beta} + K^T V^{-1} (y - X \hat{\beta})$$

has expectation

$$C^T \beta + 0 = \mathbb{E}(C^T \beta + u)$$

Further,

$$\begin{aligned} C^T \hat{\beta} + \hat{u} &= (C^T + K^T V^{-1} X) \hat{\beta} + K^T V^{-1} y \\ &= \underbrace{[C^T - K^T V^{-1} X]}_{= b^T Y} (X^T V^{-1} X)^{-1} X^T V^{-1} y + K^T V^{-1} y \\ &= b^T Y \end{aligned}$$

$\Rightarrow C^T \hat{\beta} + \hat{u}$ is Linear in Y .

It is unbiased for $C^T \beta + u$, we have

$$C = X^T b$$

If $a^T y$ is any other LUP of $C^T \beta + u$,

$$\begin{aligned} MSE(a^T y) &= \mathbb{E}(a^T y - (C^T \beta + u))^2 \\ &= Var(a^T y - u) \\ &= Var(a^T y - b^T y + b^T y - u) \\ &= Var((a-b)^T y) + Var(b^T y - u) + 0 \\ &\quad \text{NOTE: } Cov((a-b)^T y, b^T y - u) = 0. \\ &= Var((a-b)^T y) + Var(b^T y - (C^T \beta + u)) \\ &\hookrightarrow = Var((a-b)^T y) + MSE(b^T y) \geq MSE(b^T y) \end{aligned}$$

- $Cov((a-b)^T y, b^T y - u)$

$$= \sigma^2 (a-b)^T (V b - K)$$

$$\text{But } V b - K = X(X^T V^{-1} X)^{-1} (c - X^T V^{-1} K) \quad (c = X^T b)$$

\therefore Covariance term is Δ

$$\begin{aligned} &\sigma^2 (a-b)^T X (X^T V^{-1} X)^{-1} (c - X^T V^{-1} K) \\ &= \sigma^2 \left\{ \underbrace{\frac{a^T X}{C^T}}_{\frac{a^T X - b^T X}{C^T}} \underbrace{\frac{b^T X}{C}}_{\frac{b^T X - b^T X}{C}} \right\} (\Delta) = 0. \end{aligned}$$

$\bullet C^T \hat{\beta} + \hat{u}$ is BLUP of $C^T \beta + u$ among all LUE of $C^T \beta + u$.

$$Cov(\varepsilon, u) = \sigma^2 K$$

$$\bullet \hat{y} = D Z^T V^{-1} (y - X \hat{\beta})$$

$$\begin{pmatrix} \gamma \\ y \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ X \beta \end{pmatrix}, \begin{pmatrix} D & ? \\ ? & V \end{pmatrix} \right)$$

$$Cov(\gamma, y) = Cov(\gamma, X \beta + Z \gamma + \varepsilon)$$

$$= D Z^T$$

$$\gamma | y \sim \mathcal{N}_2 \left(\begin{pmatrix} D Z^T V^{-1} (y - X \beta) \\ D - D Z^T V^{-1} Z D \end{pmatrix}, \begin{pmatrix} \mu_1 + \Sigma_1 \Sigma_2^{-1} (y - X \beta) \\ \Sigma_2 \end{pmatrix} \right)$$

LX $Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \quad j=1, 2, 3$

$$y = X \beta + Z \gamma + \varepsilon$$

$$R = Cov \varepsilon = \sigma^2 I, \quad J_2 = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix}$$

$$D = 6^2 I_3, \quad V = Z D Z^T + R$$

$$\begin{pmatrix} Y_1 \\ \vdots \\ Y_6 \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \mu + \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{pmatrix} + \varepsilon$$

$$V^{-1} = \frac{1}{6^2} \left(\begin{array}{cc|c} I_2 & \frac{6^2}{26^2 + 6^2} J_2 & 0 \\ \hline 0 & 0 & 0 \end{array} \right)$$

$$\Rightarrow \text{BLUP} = D Z^T V^{-1} (y - X \hat{\beta})$$

$$= \frac{2 \cdot 6^2}{6^2} \left(\frac{\bar{Y}_1 - \bar{Y}}{\bar{Y}_2 - \bar{Y}} \right) / (6^2 + 2 \cdot 6^2)$$