

Hw6 Biostat 250B

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- 1 Suppose  $X$  has a Binomial distribution with parameter  $n$  and  $p$ . Find the posterior density of  $p$  if we have a Beta( $a, b$ ) prior distribution for  $p$ . Calculate the mean and variance of the posterior density and describe their relationships with those of the Beta( $a, b$ ). Does it make sense to use method of moment estimators to estimate the mean and variance of this posterior density? Explain.
- 2 Suppose  $X$  has an inverted gamma with density given by  $f(x, a, b) = b^a x^{-a-1} \exp(-b/x)/\Gamma(a)$ , where  $a > 0$  and  $b > 0$  and  $\Gamma$  is the Gamma function.
  - (i) Describe how values of  $a$  and  $b$  affect the density of  $X$  in terms of shape, skewness and symmetry.
  - (ii) Find its mean and describe how this density is typically used in estimating parameters in a linear model using a Bayesian paradigm.
- 3 Suppose  $Z$  has a density proportional to  $z^{n/2-1} \exp(-nz/2)$  and conditional on  $Z$ ,  $X|Z=z$  is multivariate normal with mean  $0$  and covariance  $(1/z)I$ , where  $I$  is the  $k \times k$  identity matrix. What is the density of  $X$ ? Determine its mean and covariance matrix of  $X$ .
- 4 Ex. 3m, #2 from Seber and Lee's book.
- 5 Ex. 3m, #3 from Seber and Lee's book.
- 6 Last week, your supervisor came to you with the paper on Bayes Estimation of Two-Phase Linear Regression Model posted on Week 8 on the class website. She was very interested in the paper and want to apply the methodology in the paper to solve a public health problem. But she cannot understand much of the paper and have no idea how to apply the methodology to a public health problem. Your tasks are (i) explain to her in simple English in not more than  $\frac{3}{4}$  page what the contents and methodology of the paper are about, and (ii) tell her a good example how the methodology can be applied to solve a public health problem.

# Q1 Next page for re-worded Q1

1 Suppose  $X$  has a Binomial distribution with parameter  $n$  and  $p$ . Find the posterior density of  $p$  if we have a Beta( $a, b$ ) prior distribution for  $p$ . Calculate the mean and variance of the posterior density and describe their relationships with those of the Beta( $a, b$ ). Does it make sense to use method of moment estimators to estimate the mean and variance of this posterior density? Explain.

Sol.  $X \sim \text{Bin}(n, p)$ ;  $P \sim \text{Beta}(a, b)$ .

(i) posterior of  $p$ :

$$\begin{aligned} f(p|x) &\propto f(x|p)g(p) \\ &= \left[ \binom{n}{x} p^x (1-p)^{n-x} \right] \left[ \frac{1}{B(a,b)} p^{a-1} (1-p)^{b-1} \right] \\ &\propto p^{x+a-1} (1-p)^{n+b-x-1} \\ &\propto \text{density of a Beta}(x+a, n+b-x) \text{ random variable.} \end{aligned}$$

$$\Rightarrow f(p|x) = \frac{1}{B(x+a, n+b-x)} p^{x+a-1} (1-p)^{n+b-x-1}$$

$$\begin{aligned} (ii) \mathbb{E}(p|x) &= \int_0^1 p f(p|x) dp \\ &= \int_0^1 p^{x+a-1} (1-p)^{n+b-x-1} \frac{1}{B(x+a, n+b-x)} dp \\ &= \frac{B(x+a+1, n+b-x)}{B(x+a, n+b-x)} = \frac{x+a}{n+a+b} \quad (\text{a}) \end{aligned}$$

$$\text{Similarly, } \mathbb{E}(p^2|x) = \frac{(x+a)(x+a+1)}{(n+a+b+1)(n+a+b)}$$

Thus,

$$\begin{aligned} \text{Var}(p|x) &= \mathbb{E}(p^2|x) - (\mathbb{E}(p|x))^2 \\ &= \frac{(n+b-x)(x+a)}{(n+a+b+1)(n+a+b)^2} \quad (\text{b}) \end{aligned}$$

(ii): MOM:

$$\begin{aligned} \widehat{\mathbb{E}(p|x)} &= \int x d\bar{F}_n(x)/n \\ &= \frac{\bar{x}}{n} \quad (\Delta) \end{aligned}$$

where  $\bar{F}_n(\cdot)$  is the empirical dist of  $X$ , and in this case,  
 $\bar{F}_n(t) = I(t \geq x)$ .

Similarly,

$$\begin{aligned} \widehat{\text{Var}(p|x)} &= \frac{1}{n} \frac{\bar{x}}{n} \left( 1 - \frac{\bar{x}}{n} \right) \\ &= \frac{\bar{x}(n-\bar{x})}{n^3} \quad (\times) \end{aligned}$$

Note that

$(\Delta)$  & (a), (b) & ( $\times$ )

are equivalent in Large sample sense

Thus, it is reasonable to use  
 $(\Delta)$  & ( $\times$ ).

# Q1 Reworded.

- 1 Suppose  $X|p$  has the Binomial distribution with parameter  $n$  and  $p$  and  $p$  has a Beta( $a, b$ ) density.
- Find the mean and variance of  $p$ .
  - Find the posterior of  $p|X$  and identify the density.
  - The predictive distribution of  $X$  is obtained by first writing down the joint density of  $X$  and  $p$  and integrating out  $p$ . Show that the predictive distribution of  $X$  is

$$P(X=x) = {}^n C_x B(a+x, b+n-x)/B(a, b), x=0, 1, 2, \dots, n.$$

This is the Beta-Binomial distribution with parameter  $n$ ,  $a$  and  $b$  and is denoted by  $X \sim \text{BetaBin}(n, a, b)$ .

- Use your result in (i) and compute the mean and variance of the Beta-Binomial random variable with parameters  $n, a$  and  $b$ .
- Let  $s=a/(a+b)$  and show that results in (iv) can be, respectively, expressed as  $ns$  and  $ns(1-s)(a+b+n)/(a+b+1)$ , thus showing that the variance of the Beta-Binomial variate has variance larger than the corresponding Binomial variate and they have the same mean.

$$\text{Sol. (i): } f(p) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} p^{a-1} (1-p)^{b-1}$$

$$\mathbb{E} P = \int_0^1 p f(p) dp = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \int_0^1 p^{a+1-1} (1-p)^{b-1} dp \\ = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \frac{\Gamma(a+1)\Gamma(b)}{\Gamma(a+b+1)} = \frac{a}{a+b}$$

$$\mathbb{E} P^2 = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \frac{\Gamma(a+2)\Gamma(b)}{\Gamma(a+b+2)} = \frac{a(a+1)}{(a+b)(a+b+1)}$$

$$\Rightarrow \text{Var } P = \mathbb{E} P^2 - (\mathbb{E} P)^2 = \frac{ab}{(a+b)^2(a+b+1)}$$

$$(ii) f(p|x) \propto f(x|p) f(p)$$

$$\propto p^{x+a-1} (1-p)^{n+b-x-1}$$

$$\Rightarrow \mathbb{E}(p|x) = \frac{x+a}{n+a+b} \quad \&$$

$$f(p|x) = \frac{\Gamma(n+a+b+1)}{\Gamma(x+a+1)\Gamma(n-x+b)} p^{x+a-1} (1-p)^{n+b-x-1}$$

$$(iii): f(x) = \int f(x|p) f(p) dp$$

$$= \int \binom{n}{x} p^x (1-p)^{n-x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} p^{a-1} (1-p)^{b-1} dp$$

$$= \binom{n}{x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \int_0^1 p^{a+x-1} (1-p)^{n+b-x-1} dp$$

$$= \binom{n}{x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \frac{\Gamma(a+x)\Gamma(n+b-x)}{\Gamma(n+a+b-1)}$$

$$= \binom{n}{x} B(a+x, n+b-x) / B(a, b)$$

(iv): Let  $X \sim \text{BetaBin}(n, a, b)$

Then

$$\mathbb{E} X = \mathbb{E}(\mathbb{E}(X|P))$$

$$= \mathbb{E}(nP)$$

$$= \frac{na}{a+b}$$

$$\text{Var } X = \mathbb{E}(\text{Var}(X|P)) + \text{Var}(\mathbb{E}(X|P))$$

$$= \mathbb{E}(np-np^2) + \text{Var}(nP)$$

$$= \frac{na}{a+b} - \frac{na(a+1)}{(a+b)(a+b+1)} + \frac{n^2 ab}{(a+b)^2(a+b+1)}$$

$$= \frac{nab(a+b+n)}{(a+b)^2(a+b+1)}$$

$$(v) \mathbb{E} X = ns. \quad 1-S = \frac{b}{a+b}$$

$$\text{Var } X = \frac{ns(1-S)(a+b+n)}{a+b+1}$$

$$\text{Since } \frac{a+b+n}{a+b+1} > 1.$$

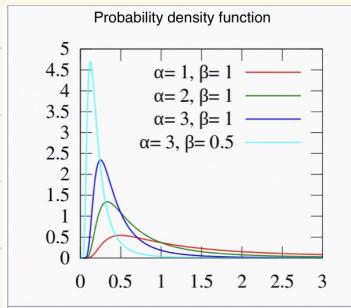
The variance is larger.

## Q2:

2 Suppose  $X$  has an inverted gamma with density given by  $f(x, a, b) = b^a x^{a-1} \exp(-bx)/\Gamma(a)$ , where  $a > 0$  and  $b > 0$  and  $\Gamma$  is the Gamma function.

- (i) Describe how values of  $a$  and  $b$  affect the density of  $X$  in terms of shape, skewness and symmetry. Find its mean and describe how this density is typically used in estimating parameters in a linear model using a Bayesian paradigm.

Sol. (i)



skewness:

$$\frac{4\alpha-2}{\alpha-3} \text{ for } \alpha > 3$$

kurtosis:

$$\frac{300\alpha-66}{(\alpha-3)(\alpha-4)} \text{ for } \alpha > 4$$

Thus,  $b$  has nothing to do with skewness & kurtosis.

$$\begin{aligned} \text{(ii): } \mathbb{E}X &= \int_0^\infty x f(x|a,b) dx \\ &= \int_0^\infty b^a x^{-a+1-1} e^{-\frac{b}{x}} \frac{1}{\Gamma(a)} dx \\ &= \frac{b^a}{\Gamma(a)} \frac{\Gamma(a-1)}{b^{a-1}} \\ &= \frac{b}{a-1} \text{ for } a > 1. \end{aligned}$$

LM:  $Y = X\beta + \varepsilon$ ,  $\varepsilon \sim \mathcal{N}(0, \sigma^2 I)$ .

We would like to put a prior on  $(\beta, \sigma^2)$ .

$$f(\beta, \sigma^2) = f(\beta | \sigma^2) f(\sigma^2)$$

&  $f(\sigma^2)$  is an inverted gamma density, &  $\beta | \sigma^2 \sim \mathcal{N}(\mu, \sigma^2 V)$

## Q3:

3 Suppose  $Z$  has a density proportional to  $z^{n/2-1} \exp(-nz/2)$  and conditional on  $Z$ ,  $X|Z=z$  is multivariate normal with mean 0 and covariance  $(1/z)I$ , where  $I$  is the  $k \times k$  identity matrix. What is the density of  $X$ ? Determine its mean and covariance matrix of  $X$ .

Sol.  $X|Z=z \sim \mathcal{N}(0, \frac{1}{z} I)$

$$f_Z(z) = \frac{(\frac{n}{2})^{\frac{n}{2}}}{\Gamma(\frac{n}{2})} z^{n/2-1} e^{-\frac{n}{2}z}$$

(i) Density of  $X$ :

$$f_X(x) = \int_0^\infty f_{X|Z}(x) f_Z(z) dz$$

$$= \int_0^\infty (2\pi)^{-\frac{n}{2}} z^{\frac{n}{2}} \exp\left\{-\frac{z}{2} x^T x\right\} \frac{(\frac{n}{2})^{\frac{n}{2}}}{\Gamma(\frac{n}{2})} z^{\frac{n}{2}-1} e^{-\frac{n}{2}z} dz$$

$$\propto \int_0^\infty z^{n-1} e^{-\frac{1}{2}(x^T x + n)z} dz$$

$$\propto \left(1 + \frac{1}{n} x^T x\right)^{-n} \quad (\Delta)$$

$(\Delta)$  is proportional to a multivariate t-dist.

$$t_n(0, I)$$

with  $df=n$ ,  $\mu=0$ ,  $\Sigma=I$ .

Thus,  $X \sim t_n(0, I)$ , density is

$$f_X(x) = \frac{\Gamma(n)}{(2\pi)^{\frac{n}{2}} \Gamma(\frac{n}{2})} \left(1 + \frac{1}{n} x^T x\right)^{-\frac{n+n}{2}}$$

(ii): mean & variance

Mean is trivial:  $g(x) = x f_X(x)$  is an odd function, thus if exists,  $\int g(x) dx = 0 \Rightarrow \mathbb{E}X = 0$ , provided  $n \geq 2$ .

Another method:

$$\mathbb{E}X = \mathbb{E}(\mathbb{E}(X|Z))$$

$$= \mathbb{E}(0)$$

$$= 0.$$

For variance, See next page:

## Q4:

Let  $X = (X_1, \dots, X_n)$ .

①

$$\text{Cov}(X_i, X_j) = \mathbb{E}(\text{Cov}(X_i, X_j | Z)) - \text{Cov}(\mathbb{E}(X_i | Z), \mathbb{E}(X_j | Z)).$$

②

$$② = \text{Cov}(0, 0) = 0.$$

$$① = \begin{cases} \mathbb{E}(0) & \text{if } i \neq j \\ \mathbb{E}\left(\frac{1}{Z}\right) & \text{if } i=j \end{cases}$$

Thus we have,

$$\begin{aligned} \text{Var } X &= \text{Var}(\mathbb{E}(X | Z)) + \\ &\quad \mathbb{E}(\text{Var}(X | Z)) \\ &= I \mathbb{E}\left(\frac{1}{Z}\right). \end{aligned}$$

$$\mathbb{E}\frac{1}{Z} = \int_0^\infty \frac{\left(\frac{n}{2}\right)^{\frac{n}{2}}}{\Gamma\left(\frac{n}{2}\right)} z^{\frac{n}{2}-1} e^{-\frac{n}{2}z} \cdot z^{-1} dz$$

$$= \frac{\left(\frac{n}{2}\right)^{\frac{n}{2}}}{\Gamma\left(\frac{n}{2}\right)} \frac{\Gamma\left(\frac{n}{2}-1\right)}{\left(\frac{n}{2}\right)^{\frac{n}{2}-1}}$$

$$= \frac{\frac{n}{2}}{\frac{n}{2}-1} = \frac{n}{n-2}$$

$$\Rightarrow \text{Var } X = \frac{n}{n-2} I$$

2. Using the noninformative prior for  $\theta$ , show that the conditional posterior density  $f(\beta | y, \sigma)$  is multivariate normal. Hence deduce that the posterior mean of  $\beta$  is  $\hat{\beta}$ .

Sol.  $Y | \beta, \sigma^2 \sim \mathcal{N}_n(X\beta, \sigma^2 I)$ .

density of  $\tilde{\Theta} = (\tilde{\beta}, \tilde{\sigma}) : f(\Theta) \propto \frac{1}{\sigma^6}$

Then

$$\begin{aligned} f(\beta | y, \sigma) &\propto f(y | \beta, \sigma) f(\beta) \\ &\propto e^{-\frac{1}{2\sigma^2} (y - X\beta)^T (y - X\beta)} \end{aligned}$$

However, let  $\hat{X}\hat{\beta} = P_X y$ , then

$$\begin{aligned} \|y - X\beta\|_2^2 &= \|y - P_X y + P_X y - X\beta\|_2^2 \\ &= \|y - P_X y\|_2^2 + \|X(\hat{\beta} - \beta)\|_2^2 \\ &\quad + 2 \langle y - P_X y, P_X y - X\beta \rangle \\ &\stackrel{(④)}{=} \|y - P_X y\|_2^2 + \|X(\hat{\beta} - \beta)\|_2^2 \end{aligned}$$

$$(⑤) : y - P_X y = (I - P_X) y \in \text{Span}(X)^\perp$$

$$P_X y - X\beta = X(\hat{\beta} - \beta) \in \text{Span}(X)$$

$$\Rightarrow \langle y - P_X y, P_X y - X\beta \rangle = 0.$$

$$\Rightarrow f(\beta | y, \sigma) \propto e^{-\frac{1}{2\sigma^2} (\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)}$$

Thus,  $\beta | y, \sigma \sim \mathcal{N}_p(\hat{\beta}, \sigma^2 (X^T X)^{-1})$

$$\mathbb{E}(\beta | y, \sigma) \stackrel{\text{a.s.}}{=} \hat{\beta}.$$

## Q5:

3. Suppose that we use the noninformative prior for  $\theta$ .

(a) If  $v = \sigma^2$ , show that  $f(v) \propto 1/v$ .

(b) Obtain an expression proportional to  $f(\beta, v | y)$ .

(c) Using (3.62), integrate out  $\beta$  to obtain

$$f(v | y) \propto v^{-(\nu/2+1)} \exp\left(-\frac{a}{v}\right),$$

where  $\nu = n - p$  and  $a = \|y - X\hat{\beta}\|^2/2$ .

(d) Find the posterior mean of  $v$ .

Sol.

$$(a): v = 6^2 \Rightarrow \log 6 = \frac{1}{2} \log \nu$$

But  $f(\beta, \log 6) \propto 1$

Jacobian for  $\begin{pmatrix} \beta \\ \log 6 \end{pmatrix} \rightarrow \begin{pmatrix} \beta \\ v \end{pmatrix}$  is

$$J = \begin{bmatrix} I & \\ & \frac{1}{2v} \end{bmatrix}$$

Thus, density for  $\begin{pmatrix} \beta \\ v \end{pmatrix}$  is

$$f(\beta, v) = |J| f(\beta, \log 6)$$

$$= \frac{1}{2v} \propto \frac{1}{v}$$

(b):  $f(\beta, v | y)$

$$\propto f(y | \beta, v) f(\beta, v)$$

$$\propto f(y | \beta, v) \frac{1}{v}$$

$$= \frac{1}{v} \frac{1}{(2\pi)^n} e^{-\frac{1}{2v} (y - X\beta)^T (y - X\beta)}$$

$$\propto v^{-\frac{n}{2}-1} e^{-\frac{1}{2v} \|y - X\beta\|^2} (v^{-1})$$

(c): Note we have

$$\|y - X\beta\|_2^2 = \|y - X\hat{\beta}\|_2^2 + \|X(\hat{\beta} - \beta)\|_2^2$$

Thus,

$$\begin{aligned} & \int_{\mathbb{R}^p} e^{-\frac{1}{2v} \|y - X\beta\|_2^2} d\beta \\ &= \int_{\mathbb{R}^p} e^{-\frac{1}{2v} \|y - X\hat{\beta}\|_2^2} e^{-\frac{1}{2v} (\beta - \hat{\beta})^T X^T X (\beta - \hat{\beta})} d\beta \\ &= \frac{(2\pi)^{\frac{p}{2}} v^{\frac{p}{2}}}{\det(X^T X)} e^{-\frac{1}{2v} \|y - X\hat{\beta}\|_2^2} \end{aligned}$$

$\Rightarrow$

$$\begin{aligned} f(v | y) &= \int_{\mathbb{R}^p} f(\beta, v | y) d\beta \\ &\propto v^{-\frac{n-p}{2}-1} e^{-\frac{1}{2v} \|y - X\hat{\beta}\|_2^2} \\ &\propto v^{-\frac{n}{2}-1} e^{-\frac{a}{v}} \end{aligned}$$

(d): From (c),  $v | y \sim \text{InvGau}(\frac{\nu}{2}, a)$   
From Q2, we have

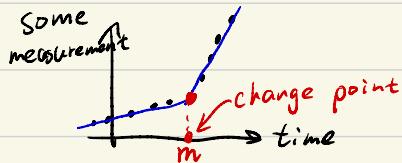
$$E(v | y) = \frac{a}{\frac{\nu}{2} - 1}$$

$$= \frac{2a}{\nu - 2}$$

Q6:

6 Last week, your supervisor came to you with the paper on Bayes Estimation of Two-Phase Linear Regression Model posted on Week 8 on the class website. She was very interested in the paper and want to apply the methodology in the paper to solve a public health problem. But she cannot understand much of the paper and have no idea how to apply the methodology to a public health problem. Your tasks are (i) explain to her in simple English in not more than ¼ page what the contents and methodology of the paper are about, and (ii) tell her a good example how the methodology can be applied to solve a public health problem.

(i): The paper discusses how to detect the change point (following figure) using Bayesian estimation.



The model is ( $m$  is the change point)

$$Y = \begin{cases} \alpha_1 + \beta_1 x_t + \varepsilon_1 & \text{if } t \leq m \\ \alpha_2 + \beta_2 x_t + \varepsilon_2 & \text{else} \end{cases}$$

$$\varepsilon_1, \varepsilon_2 \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$$

Encoding prior information on  $(\alpha_1, \alpha_2, \beta_1, \beta_2, \sigma^2)$  the authors calculate the posterior density of  $t$ :

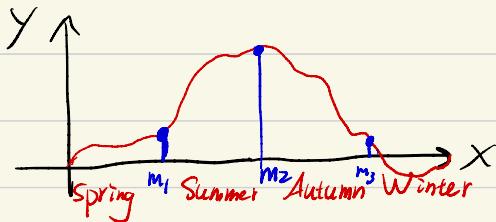
$$(17) f(t | Y) = \int_{\Theta} f(t, \alpha_1, \alpha_2, \beta_1, \beta_2, \sigma^2 | Y) d\theta$$

$$\text{where } \Theta = (\alpha_1, \alpha_2, \beta_1, \beta_2, \sigma^2)$$

Equation (17) & (26) explicitly write down what (17) is.

With  $f(t|Y)$  & different choices of criterion (Aka loss function), one can derive the estimation of the change point  $m$ . See eq (31)(32)(36)(39)(41)

(ii) Assuming  $Y$  represents the temperature &  $X$  represents date.



The above figure shows a typical pattern in a region of North-hemisphere.  $m_1, m_2, m_3$  can be 3 potential "change points" of temperature.

One can apply the estimation of  $m$  in the paper to derive different estimated values of  $m_1, m_2$  and  $m_3$ .