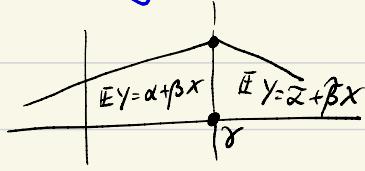



Two phase regression & Fieller's theorem



$$\mathbb{E} Y = \begin{cases} \alpha_1 + \beta_1 x, & x \leq r \\ \alpha_2 + \beta_2 x, & x \geq r \end{cases}$$

Continuity assumption.

$$\alpha_1 + \beta_1 r = \alpha_2 + \beta_2 r$$

$$\hat{\beta} = \frac{\hat{\alpha}_1 - \hat{\alpha}_2}{\hat{\beta}_2 - \hat{\beta}_1} \quad \text{point estimate}$$

$y = X\beta + \varepsilon$ for a two phase reg.

Calibration: $y, X, \hat{\beta}$ are known.

- Observe y_0 but no info. on (x_0) was provided. Q: How to estimate x_0 ?

$$y_0 = \hat{\alpha} + \hat{\beta}(x_0 - \bar{x})$$

$$\Rightarrow x_0 = \frac{y_0 - \hat{\alpha}}{\hat{\beta}} + \bar{x} = \frac{y_0 - \bar{y}}{\hat{\beta}} + \bar{x}$$

Fieller's Theorem Suppose U, V

are uncorrelated r.v.s,

$$U \sim N(\mu_U, \sigma_U^2), \sigma_U^2 = \sigma^2 \alpha_1$$

$$V \sim N(\mu_V, \sigma_V^2), \sigma_V^2 = \sigma^2 \alpha_2$$

$$\hat{\sigma}_U = S \alpha_1, \hat{\sigma}_V = S \alpha_2.$$

Define $g = t_{m, \alpha/2}^2 S^2 \alpha_2^2 / \sigma^2$, m : df associated with S^2 for estimating σ^2 .

$$\sigma_U^2 = \sigma^2 \alpha_1, \sigma_V^2 = \sigma^2 \alpha_2$$

If $g < 1$, a $100(1-\alpha)\%$ C.I. for

$$\frac{\mu_U}{\mu_V}$$
 is:

$$\frac{1}{1-g} \left(\frac{U}{V} \pm \frac{t_{m, \alpha/2} S}{\sqrt{V}} \sqrt{(1-g) \alpha_1^2 + \frac{U^2 \alpha_2^2}{V^2}} \right)$$

$$\text{Pf: } \theta = \frac{\mu_U}{\mu_V} \text{ and } W = U - \theta V$$

$$\mathbb{E} W = \mu_U - \theta \mu_V = 0$$

$$\text{Var } W = \sigma^2 \alpha_1 + \theta^2 \sigma^2 \alpha_2$$

$$\widehat{\text{Var}} W = S^2 \alpha_1 + S^2 \alpha_2 \frac{U}{V}$$

$$U - \theta V \sim N(0, \sigma^2 \alpha_1^2 + \theta^2 \sigma^2 \alpha_2^2)$$

$$\text{Also } \frac{mS^2}{\sigma^2} \sim \chi_m^2$$

$$\frac{(U - \theta V) / \sqrt{\sigma^2 \alpha_1^2 + \theta^2 \alpha_2^2}}{\sqrt{\frac{mS^2}{\sigma^2}/m}} \sim t_m$$

$$(U - \theta V)^2 = t_{m, \alpha/2}^2 S^2 (\alpha_1^2 + \theta^2 \alpha_2^2)$$

Verify equation simplifies to

$$(1-g)\theta^2 - \frac{2U}{V}\theta + \frac{U^2}{V^2} - g \frac{\alpha_1^2}{\alpha_2^2} = 0$$

$$\theta = \frac{1}{2a} (-b \pm \sqrt{b^2 - 4ac})$$

Verify

$$\theta \in \frac{1}{1-g} \left(\frac{U}{V} \pm \frac{t_{m, \alpha/2} S}{\sqrt{V}} \sqrt{(1-g) \alpha_1^2 + \frac{U^2 \alpha_2^2}{V^2}} \right)$$

$$U \leftarrow y_0 - \bar{y}; V \leftarrow \hat{\beta}$$

$$\text{Var } U = \sigma^2 + \frac{\sigma^2}{n} = \sigma^2 (1 + \frac{1}{n})$$

$$\text{Var } V = \sigma^2 \left(\frac{1}{\sum (x_i - \bar{x})^2} \right)$$

Verify $100(1-\alpha)\%$ C.I. for x_0 is

$$\bar{x} + \frac{y_0 - \bar{y}}{\hat{\beta}(1-g)} \pm \frac{t_{n-2, \alpha/2} S}{\hat{\beta}(1-g)} \sqrt{\frac{1}{n} (1+n) + \frac{(y_0 - \bar{y})^2}{\hat{\beta}^2 \sum (x_i - \bar{x})^2}}$$

$$\text{Ex } \bar{E}Y = \beta_0 + \beta_1(X - \bar{X})$$

$$\text{If } Y=0: X - \bar{X} = -\frac{\beta_0}{\beta_1}$$

$\hat{x}_0 = \hat{\beta}_0 / \hat{\beta}_1$. Fieller's thm applies.

Ex Prognostic factors interaction assessment in a 2-treatment trial with treatment T & C, along with a prognostic factor.

$$\bar{E}Y_1 = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

$$\bullet x_1 = \begin{cases} 1 & \text{if treatment T} \\ 0 & \text{if C} \end{cases}$$

$$\bullet x_2 \text{ some prognostic factor}$$

$$\bar{E}Y_T = \beta_0 + \beta_1 + \beta_2 x_2 + \beta_3 x_2$$

$$\bar{E}Y_C = \beta_0 + \beta_2 x_2$$

• $\bar{E}(Y_T - Y_C) = \beta_1 + \beta_3 x_2$

Q: what x_2 for which $\bar{E}(Y_T - Y_C) > 0$.

i.e. WT estimate $x_2 = -\frac{\beta_1}{\beta_3}$

$$\hat{\beta} = (X^T X)^{-1} X^T \bar{y} = \begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{\beta}_2 \\ \hat{\beta}_3 \end{pmatrix}; \text{cov}(\hat{\beta}) = \sigma^2 (X^T X)^{-1}$$

$$\tilde{\beta} = \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_3 \end{pmatrix}; \text{cov}(\tilde{\beta}) = V = \begin{pmatrix} V_{11} & V_{13} \\ V_{13} & V_{33} \end{pmatrix}$$

let $w = \tilde{\beta}_1 + \theta \tilde{\beta}_3, \theta = -\frac{\beta_1}{\beta_3}$

then $\bar{E}w = 0$

$$\text{Var } w = (V_{11} + \theta^2 V_{33} + 2\theta V_{13}) \sigma^2$$

So $\frac{(n-4)\hat{\sigma}^2}{\sigma^2} \sim \chi^2_{n-4}; \frac{w}{\sqrt{\text{Var } w}} \sim N(0,1)$

If σ^2 is known,

$$(\hat{\beta}_1 + \theta \hat{\beta}_3)^2 = Z_{\alpha/2}^2 (V_{11} + \theta^2 V_{33} + 2\theta V_{13}) \sigma^2$$

CI for θ are

$$\theta = \frac{-w_2 \pm \sqrt{w_2^2 - 4w_1 w_2}}{2w_2}$$

$$w_1 = \hat{\beta}_1^2 - V_{11} Z_{\alpha/2}^2$$

$$w_2 = 2\hat{\beta}_1 \hat{\beta}_3 - 2V_{13} Z_{\alpha/2}^2$$

$$w_3 = \hat{\beta}_3^2 - V_{33} Z_{\alpha/2}^2$$

Note: $\beta_1 = 0$ iff $\frac{\hat{\beta}_1}{\sqrt{V_{11}}} > Z_{\alpha/2}$
iff $w_1 > 0$.

$\beta_3 = 0$ iff $\frac{\hat{\beta}_3}{\sqrt{V_{33}}} > Z_{\alpha/2}$ iff $w_3 > 0$

Ex Health economics.

cost effectiveness ratio

$$R = \frac{\mu_{CT} - \mu_{CC}}{\mu_{ET} - \mu_{EC}}$$

$$\hat{R} = \frac{\bar{G}_T - \bar{G}_C}{\bar{E}_T - \bar{E}_C}$$

C.I. for R. Define

$$w = (\bar{G}_T - \bar{G}_C) - \theta (\bar{E}_T - \bar{E}_C)$$

Var w (estimate)

Extensions to correlated variables
in the ratio.

$$\begin{pmatrix} X \\ Y \end{pmatrix} \sim N\left(\begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix}, \sigma^2 \begin{pmatrix} V_{11} & V_{12} \\ V_{21} & V_{22} \end{pmatrix}\right)$$

WT find C.I. for $\theta = \mu_X/\mu_Y$

Let $W = X - \theta Y$

$$\mathbb{E} W = 0, \text{Var } W = \sigma^2(V_{11} + V_{22}\theta^2 - 2\theta V_{12}) \\ = \sigma^2 - \sigma_w^2$$

$$\frac{W}{\sqrt{\sigma^2 - \sigma_w^2}} \sim N(0, 1)$$

$$\frac{r S_r^2}{\sigma^2} \sim \chi_r^2 \quad S_r^2: \text{unbiased estimator of } \sigma^2$$

$$\text{Then } \frac{X - \theta Y}{S_r \sqrt{V_{11} + \theta^2 V_{22} - 2\theta V_{12}}} \sim t_r$$

Verify the limits for θ are

$$\frac{1}{1-g} \left\{ \frac{X}{Y} + \frac{V_{12}g}{V_{22}} \pm \frac{t_r}{\sqrt{r}} \sqrt{r} \right\} \text{ where } r \text{ is}$$

$$V_{11} + \frac{2XV_{12}}{Y} + \frac{X^2}{Y^2} V_{22} - g(V_{11} - \frac{V_{12}^2}{V_{22}})$$

$$\& \quad g = \frac{t_r^2 S_r^2 V_{22}}{Y^2} \quad [\text{Wiki...}]$$

(Ex) Two phase reg.

\hat{f} = estimated threshold when

(Seber, 1961) change occurs

$$= \frac{\hat{\alpha}_1 - \hat{\alpha}_2}{\hat{\beta}_1 - \hat{\beta}_2}$$

$$W = (\hat{\alpha}_1 - \hat{\alpha}_2) + \gamma(\hat{\beta}_1 - \hat{\beta}_2) \Rightarrow$$

$$\mathbb{E} W = 0$$