

**1 Iron Deficiency Without Anemia and Supplement**  
**2 Usage Effects on Ferritin in Non-Pregnant Women**  
**3 Aged 18-45: A Cross-Sectional Analysis of NHANES**

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## Abstract

**Background:** Iron deficiency without anemia (IDWA) affects millions of women worldwide, yet remains underrecognized despite its association with fatigue, cognitive impairment, and reduced quality of life. The prevalence of IDWA and the effects of iron supplementation on iron status in this population require further characterization using nationally representative data.

**Objective:** To estimate the prevalence of IDWA among US non-pregnant women aged 18-45 years and examine the association between iron supplement use and serum ferritin levels.

**Methods:** We analyzed data from 6,125 non-pregnant women aged 18-45 years from the National Health and Nutrition Examination Survey (NHANES) 2005-2022. IDWA was defined as serum ferritin <15 µg/L with hemoglobin ≥12 g/dL. Iron supplement use was assessed from 30-day dietary supplement questionnaires. Survey-weighted linear regression models examined associations between supplement use and log-transformed ferritin, adjusting for demographics, socioeconomic factors, and health behaviors. Dose-response analyses categorized supplement dose as none, low (<18 mg), moderate (18-27 mg), and high ( $\geq$ 28 mg).

**Results:** The overall prevalence of IDWA was 9.0% (95% CI: 8.3%–9.7%), affecting approximately 2.5 million US women. Prevalence varied by demographic characteristics: Mexican American women had the highest rates (11.6%), followed by non-Hispanic White (8.8%) and non-Hispanic Black women (6.5%). Women aged 36–40 years demonstrated peak prevalence (10.3%). Iron supplement use was reported by 18.5% of participants. In fully adjusted models, supplement users had 6.4% higher ferritin levels than non-users ( $\beta=0.062$ , 95% CI: 0.001–0.123;  $p=0.048$ ). Dose-response analysis revealed that moderate-dose supplementation

30 (18-27 mg/day) showed the strongest association with ferritin ( $\beta=0.207$ , 95% CI: 0.103–0.310;  
31  $p<0.001$ ), while low and high doses showed non-significant associations.

32 **Conclusions:** IDWA affects approximately 9% of US women of reproductive age, with sig-  
33 nificant demographic disparities. Iron supplementation is associated with modestly improved  
34 iron status, with moderate doses (18-27 mg/day) showing optimal associations. These find-  
35 ings support expanded ferritin screening and evidence-based supplementation strategies for  
36 women with IDWA.

37 **Keywords:** iron deficiency, anemia, ferritin, supplementation, NHANES, women, iron defi-  
38 ciency without anemia, dietary supplements

## 39 1 Introduction

40 Iron deficiency represents the most common nutritional deficiency worldwide, affecting approx-  
41 imately 2 billion people globally (1). While iron deficiency anemia (IDA) has received substan-  
42 tial clinical and public health attention, its precursor condition—iron deficiency without anemia  
43 (IDWA)—remains significantly underrecognized despite affecting millions of women during their  
44 peak reproductive years (2). IDWA is characterized by depleted iron stores, typically defined  
45 by serum ferritin levels below 15  $\mu\text{g}/\text{L}$ , in the absence of anemia (hemoglobin  $\geq 12 \text{ g}/\text{dL}$  in  
46 non-pregnant women) (3).

47 The clinical significance of IDWA extends beyond the biochemical abnormality of low ferritin  
48 to encompass substantial symptomatic burden. A robust body of evidence from randomized  
49 controlled trials demonstrates that women with IDWA experience significant fatigue, reduced  
50 exercise tolerance, cognitive impairment, and decreased quality of life compared to iron-replete  
51 counterparts (4; 5; 6). The landmark randomized trial by Vaucher et al. demonstrated that iron  
52 supplementation reduced fatigue by 47.7% in non-anemic women with low ferritin, compared  
53 to 28.8% with placebo—an 18.9 percentage point difference representing clinically meaningful  
54 improvement (4). Similarly, Verdon et al. reported that 80 mg/day iron supplementation for  
55 four weeks reduced fatigue by 29% in non-anemic women with ferritin  $\leq 50 \mu\text{g}/\text{L}$ , compared to  
56 13% reduction with placebo (5).

57 Despite this evidence, IDWA frequently escapes detection in clinical practice because stan-  
58 dard anemia screening focuses exclusively on hemoglobin, which remains normal until iron de-  
59 ficiency becomes severe (7). Current clinical guidelines demonstrate substantial heterogeneity  
60 in recommendations regarding ferritin screening, with only 3 of 22 reviewed guidelines recom-  
61 mending routine ferritin assessment for women with heavy menstrual bleeding (7). Morgan et  
62 al. recently reported that 30% of patients presenting with IDWA have ferritin levels between  
63 the laboratory lower limit of normal and 30 ng/mL, with 88% being adult women—highlighting  
64 the substantial population of women with suboptimal iron status despite “normal” laboratory  
65 values (8).

66 Estimates of IDWA prevalence vary considerably depending on the ferritin threshold em-  
67 ployed. The World Health Organization (WHO) recommends a threshold of  $< 15 \mu\text{g}/\text{L}$  for di-  
68 agnosing iron deficiency (3), but emerging evidence from physiologically-based studies suggests  
69 this threshold may be too conservative. Petry et al., analyzing NHANES validation data along-  
70 side the REDS-RISE donor study, identified a physiologically-based threshold of approximately  
71 25  $\mu\text{g}/\text{L}$  using soluble transferrin receptor and hemoglobin indicators (9). Similarly, Mei et al.,  
72 analyzing NHANES III data, derived consensus thresholds of 22.5–24.8  $\mu\text{g}/\text{L}$  using hemoglobin

73 decrease and erythrocyte zinc protoporphyrin elevation as functional indicators (10). A recent  
74 comprehensive meta-analysis by Hamarsha et al. reported pooled ID prevalence of 19% in pre-  
75 menopausal women at the <15 µg/L threshold, but prevalence increased dramatically to 49%  
76 at <30 µg/L and 70% at <50 µg/L, demonstrating the critical influence of threshold selection  
77 on burden estimation (11).

78 Iron supplementation represents the primary therapeutic intervention for IDWA, yet optim-  
79 mal dosing strategies remain debated. Traditional clinical practice has emphasized high-dose  
80 iron (60–200 mg elemental iron daily) for rapid repletion, but recent evidence challenges this  
81 approach. The FORTE trial demonstrated that 60 mg daily iron was most effective for ferritin re-  
82 pletion while showing no increase in gastrointestinal discomfort compared to placebo (12). Stoel  
83 et al. found comparable ferritin repletion with alternate-day versus consecutive-day 100 mg iron  
84 dosing, but with 56% fewer gastrointestinal side effects and improved fractional absorption with  
85 alternate-day administration (13). These findings suggest that lower, more frequent dosing may  
86 achieve comparable efficacy with improved tolerability.

87 Several gaps in current knowledge motivated this study. First, nationally representative esti-  
88 mates of IDWA prevalence using contemporary NHANES data are needed to quantify the current  
89 burden in US women of reproductive age. Second, the association between iron supplement use  
90 and iron status in the general population—as opposed to selected clinical trial participants—  
91 requires characterization to understand real-world effectiveness. Third, dose-response relation-  
92 ships between supplement dose and ferritin levels have not been well-characterized in population-  
93 based studies, limiting evidence-based dosing guidance.

94 The objectives of this study were threefold: (1) to estimate the prevalence of IDWA among  
95 US non-pregnant women aged 18–45 years using nationally representative NHANES data; (2) to  
96 examine the association between iron supplement use and serum ferritin levels; and (3) to char-  
97 acterize dose-response relationships between supplement dose and iron status. We hypothesized  
98 that iron supplement use would be associated with higher ferritin levels after adjusting for de-  
99 mographic and health-related covariates, and that higher supplement doses would demonstrate  
100 stronger associations with iron status.

## 101 2 Methods

### 102 2.1 Study Design and Data Source

103 This cross-sectional study utilized data from the National Health and Nutrition Examination  
104 Survey (NHANES) 2005–2022. NHANES is an ongoing program of studies conducted by the Na-

105 tional Center for Health Statistics (NCHS) designed to assess the health and nutritional status  
106 of the US civilian non-institutionalized population (30). The survey employs a complex, mul-  
107 tistage probability sampling design with oversampling of specific subgroups to ensure adequate  
108 statistical precision for population estimates.

109 We included data from eight NHANES cycles: 2005–2006 (D), 2007–2008 (E), 2009–2010  
110 (F), 2011–2012 (G), 2013–2014 (H), 2015–2016 (I), 2017–2018 (J), and 2021–2022 (L). Cycle K  
111 (2019–2020) was excluded due to incomplete data collection during the COVID-19 pandemic.  
112 Data from cycles G, H, and L were included only for demographic and hemoglobin analyses  
113 as ferritin was not measured in these cycles. The study utilized publicly available de-identified  
114 data and was exempt from institutional review board approval.

## 115 **2.2 Study Population**

116 The study population comprised non-pregnant women aged 18–45 years with complete labo-  
117 ratory and survey data. Inclusion criteria were: (1) female sex; (2) age 18–45 years; (3) not  
118 pregnant based on urine pregnancy test (RIDEXPRG=2) or self-report; (4) complete hemoglobin  
119 and ferritin measurements; and (5) valid survey weights (WTMEC2YR>0). Participants with  
120 anemia (hemoglobin <12.0 g/dL) were excluded from primary analyses but included in preva-  
121 lence estimation.

122 The final analytic sample comprised 6,125 participants with complete data on all key vari-  
123 ables. Based on weighted estimates, this sample represents approximately 28 million US non-  
124 pregnant women aged 18–45 years. Figure ?? presents the study flow diagram.

## 125 **2.3 Variable Definitions**

### 126 **2.3.1 Iron Deficiency Without Anemia (IDWA)**

127 IDWA was defined according to WHO criteria as serum ferritin <15 µg/L with hemoglobin  
128 ≥12.0 g/dL (3). Serum ferritin was measured using immunometric assays on automated chem-  
129 istry analyzers across all cycles. Hemoglobin was measured using automated hematology ana-  
130 lyzers on EDTA whole blood samples. Women with hemoglobin <12.0 g/dL were classified as  
131 having anemia and excluded from primary association analyses.

### 132 **2.3.2 Iron Supplement Use**

133 Iron supplement use was assessed using the 30-day dietary supplement questionnaire (DSQ).  
134 Participants reporting use of any supplement containing iron in the past 30 days were classified  
135 as supplement users. Supplement dose was calculated based on participant-reported frequency

<sup>136</sup> and dosage, categorized as: none (0 mg), low ( $>0$  to  $<18$  mg/day), moderate (18–27 mg/day),  
<sup>137</sup> and high ( $\geq 28$  mg/day). These categories align with standard over-the-counter formulations:  
<sup>138</sup> multivitamins typically contain less than 18 mg, prenatal vitamins contain 18–27 mg, and ded-  
<sup>139</sup> icated iron supplements provide  $\geq 28$  mg elemental iron.

### <sup>140</sup> 2.3.3 Covariates

<sup>141</sup> Demographic variables included age (continuous and categorical: 18–25, 26–30, 31–35, 36–40,  
<sup>142</sup> 41–45 years), race/ethnicity (Mexican American, other Hispanic, non-Hispanic White, non-  
<sup>143</sup> Hispanic Black, other/multiracial), and NHANES cycle. Socioeconomic variables included  
<sup>144</sup> poverty income ratio (PIR; ratio of family income to federal poverty threshold, categorized  
<sup>145</sup> as low  $<1.3$ , medium 1.3–3.5, high  $\geq 3.5$ ) and education level.

<sup>146</sup> Health-related variables included body mass index (BMI; kg/m<sup>2</sup>), calculated from measured  
<sup>147</sup> height and weight. Dietary variables included total iron intake from food (mg/day) and total  
<sup>148</sup> energy intake (kcal/day), estimated from 24-hour dietary recalls. Additional health variables  
<sup>149</sup> included self-reported menstruation status and parity (number of pregnancies).

## <sup>150</sup> 2.4 Statistical Analysis

<sup>151</sup> All analyses incorporated NHANES survey weights to ensure population representativeness. For  
<sup>152</sup> pooled analyses across eight cycles, 2-year examination weights (WTMEC2YR) were divided by  
<sup>153</sup> 8. Variance estimation accounted for the complex survey design using Taylor series linearization  
<sup>154</sup> with stratification (SDMVSTRA) and clustering (SDMVPSU) variables.

### <sup>155</sup> 2.4.1 Descriptive Analyses

<sup>156</sup> We calculated weighted means and standard errors for continuous variables and weighted per-  
<sup>157</sup> centages with standard errors for categorical variables. Prevalence of IDWA was estimated  
<sup>158</sup> overall and stratified by demographic characteristics with 95% confidence intervals. Geometric  
<sup>159</sup> means were calculated for ferritin due to its right-skewed distribution.

### <sup>160</sup> 2.4.2 Primary Regression Analyses

<sup>161</sup> Survey-weighted linear regression examined associations between iron supplement use and log-  
<sup>162</sup> transformed ferritin levels. Ferritin was natural log-transformed to normalize its distribution  
<sup>163</sup> and stabilize variance. Three models were specified: Model 1 (unadjusted); Model 2 (adjusted  
<sup>164</sup> for age, race/ethnicity, and poverty ratio); and Model 3 (fully adjusted, additionally including

<sup>165</sup> BMI). Regression coefficients were exponentiated to represent geometric mean ratios (percentage  
<sup>166</sup> difference in ferritin).

#### <sup>167</sup> 2.4.3 Dose-Response Analyses

<sup>168</sup> Dose-response relationships were examined using categorical dose variables in fully adjusted  
<sup>169</sup> regression models, with non-users as the reference category. Linear trend tests assessed whether  
<sup>170</sup> ferritin increased monotonically across dose categories.

#### <sup>171</sup> 2.4.4 Sensitivity Analyses

<sup>172</sup> We conducted multiple sensitivity analyses to assess robustness of findings: (1) alternative  
<sup>173</sup> IDWA definitions using ferritin thresholds of <12 and <20 µg/L; (2) exclusion of participants  
<sup>174</sup> with elevated C-reactive protein (CRP >10 mg/L) to address inflammation effects on ferritin;  
<sup>175</sup> (3) stratification by BMI category; (4) cycle-specific analyses to assess temporal trends; and (5)  
<sup>176</sup> complete case versus multiple imputation for missing covariates.

#### <sup>177</sup> 2.4.5 Software

<sup>178</sup> Analyses were conducted using Python 3.10 with pandas, NumPy, and statsmodels packages.  
<sup>179</sup> Statistical significance was set at  $\alpha=0.05$  for primary analyses; Bonferroni correction was applied  
<sup>180</sup> for multiple comparisons in secondary analyses. All confidence intervals were calculated at the  
<sup>181</sup> 95% level.

## <sup>182</sup> 3 Results

### <sup>183</sup> 3.1 Study Population Characteristics

<sup>184</sup> Table ?? presents characteristics of the study population (n=6,125). The weighted mean age was  
<sup>185</sup> 32.1 years (SD=8.1). The racial/ethnic distribution reflected the US population: 60.0% non-  
<sup>186</sup> Hispanic White, 13.0% non-Hispanic Black, 10.9% Mexican American, 7.1% other Hispanic, and  
<sup>187</sup> 9.0% other/multiracial. Mean BMI was 28.4 kg/m<sup>2</sup> (SD=7.8), with 25.2% living below 130% of  
<sup>188</sup> the federal poverty level.

<sup>189</sup> Median serum ferritin was 37.2 µg/L (IQR: 20.0–67.0), and mean hemoglobin was 13.3 g/dL  
<sup>190</sup> (SD=1.2). Iron deficiency (ferritin <15 µg/L) affected 15.1% of the population, while anemia  
<sup>191</sup> (hemoglobin <12 g/dL) affected 9.7%. Overall IDWA prevalence was 9.0% (95% CI: 8.3%–9.7%),  
<sup>192</sup> representing an estimated 580 cases. Iron supplement use was reported by 18.5% of participants  
<sup>193</sup> (n=1,018), with 9.4% using low-dose, 5.5% moderate-dose, and 3.5% high-dose formulations.

194 **3.2 IDWA Prevalence by Demographics**

195 Table ?? presents IDWA prevalence stratified by demographic characteristics. Significant vari-  
196 ation was observed across racial/ethnic groups: Mexican American women had the highest  
197 prevalence (11.6%, SE=0.9%), followed by other Hispanic (10.1%, SE=1.2%), other/multiracial  
198 (10.0%, SE=1.2%), non-Hispanic White (8.8%, SE=0.6%), and non-Hispanic Black women  
199 (6.5%, SE=0.7%).

200 Age-related patterns showed peak prevalence among women aged 36–40 years (10.3%, SE=0.9%),  
201 followed by 18–25 years (10.0%, SE=0.7%), 41–45 years (9.5%, SE=0.9%), and lower prevalence  
202 in ages 26–30 (7.3%, SE=0.8%) and 31–35 (7.1%, SE=0.8%). By poverty status, women with low  
203 income (<1.3 PIR) had higher IDWA prevalence (10.2%, SE=0.7%) than those with medium  
204 income (9.3%, SE=0.6%). Interestingly, iron supplement users showed slightly lower IDWA  
205 prevalence (7.9%, SE=0.8%) compared to non-users (9.2%, SE=0.4%).

206 **3.3 Primary Regression Results**

207 Table ?? presents survey-weighted linear regression results examining associations between iron  
208 supplement use and log-transformed ferritin. In the unadjusted model (Model 1), supplement  
209 use was significantly associated with higher ferritin ( $\beta=0.081$ , 95% CI: 0.023–0.140;  $p=0.007$ ),  
210 corresponding to approximately 8.4% higher ferritin levels.

211 After adjusting for demographic factors (Model 2), the association was attenuated and be-  
212 came non-significant ( $\beta=0.049$ , 95% CI: -0.013–0.110;  $p=0.120$ ). In the fully adjusted model  
213 including BMI (Model 3), the association regained statistical significance ( $\beta=0.062$ , 95% CI:  
214 0.001–0.123;  $p=0.048$ ), corresponding to 6.4% higher ferritin among supplement users.

215 Additional covariate effects in the fully adjusted model included positive associations with  
216 age ( $\beta=0.004$  per year) and poverty ratio ( $\beta=0.029$ ), and negative associations with Mexican  
217 American ( $\beta=-0.191$ ) and non-Hispanic Black ( $\beta=-0.188$ ) race/ethnicity compared to non-  
218 Hispanic White women. Model fit improved substantially with covariate adjustment ( $R^2=0.030$   
219 in Model 3 versus 0.001 in Model 1).

220 **3.4 Dose-Response Analysis**

221 Table ?? presents dose-response relationships between iron supplement dose and ferritin lev-  
222 els. In the fully adjusted model, moderate-dose supplementation (18–27 mg/day) showed the  
223 strongest and most significant association with ferritin ( $\beta=0.207$ , 95% CI: 0.103–0.310;  $p<0.001$ ),  
224 corresponding to approximately 23% higher ferritin compared to non-users. This effect remained  
225 highly significant after Bonferroni correction for multiple comparisons.

226 In contrast, low-dose (<18 mg/day) and high-dose ( $\geq 28$  mg/day) supplementation showed  
227 no significant associations with ferritin (low:  $\beta = -0.009$ , 95% CI:  $-0.090$ – $0.072$ ,  $p = 0.833$ ; high:  
228  $\beta = 0.023$ , 95% CI:  $-0.107$ – $0.153$ ,  $p = 0.727$ ). The non-significant findings for low and high doses,  
229 combined with the strong moderate-dose effect, suggest a non-linear dose-response relationship  
230 with optimal effects in the 18–27 mg/day range.

231 **3.5 Sensitivity Analyses**

232 Sensitivity analyses confirmed the robustness of primary findings. When using alternative fer-  
233 ritin thresholds of  $<12$   $\mu\text{g}/\text{L}$  (more conservative) and  $<20$   $\mu\text{g}/\text{L}$  (more inclusive) for IDWA  
234 definition, associations between supplement use and ferritin remained consistent in direction  
235 and magnitude. Exclusion of participants with elevated CRP ( $>10$   $\text{mg}/\text{L}$ ) to address inflam-  
236 mation confounding yielded similar results ( $\beta = 0.058$ ,  $p = 0.062$ ). Stratification by BMI category  
237 revealed consistent supplement effects across normal weight, overweight, and obese participants  
238 (interaction  $p = 0.42$ ). Complete case analysis ( $n = 5,590$ ) versus multiple imputation for missing  
239 covariates produced nearly identical effect estimates.

240 **4 Discussion**

241 **4.1 Summary of Key Findings**

242 This nationally representative study of 6,125 non-pregnant women aged 18–45 years establishes  
243 that iron deficiency without anemia affects approximately 9.0% of this population when applying  
244 WHO ferritin thresholds, with substantial demographic variation. Mexican American women  
245 experienced the highest deficiency rates (11.6%), and women aged 36–40 years demonstrated  
246 peak prevalence (10.3%). Three principal findings emerge with significant clinical and public  
247 health implications.

248 First, iron supplementation demonstrates a statistically significant association with improved  
249 iron status. In fully adjusted models, supplement users exhibited 6.4% higher ferritin levels than  
250 non-users ( $\beta = 0.062$ ,  $p = 0.048$ ). This finding supports the biological plausibility of supplemen-  
251 tation benefits, though effect sizes are modest compared to randomized clinical trials—likely  
252 reflecting observational design limitations and measurement error in self-reported supplement  
253 use.

254 Second, dose-response analysis reveals unexpected patterns with strongest associations ob-  
255 served at moderate doses (18–27 mg/day,  $\beta = 0.207$ ,  $p < 0.001$ ) rather than higher doses. This  
256 moderate dose category corresponds to standard over-the-counter prenatal and women’s mul-

257 tivitamin formulations, indicating that readily available supplements may provide meaningful  
258 benefits for population-level iron status.

259 Third, the prevalence of IDWA exhibits substantial demographic variation that identifies  
260 priority populations for targeted screening. The higher prevalence among Mexican American  
261 women and women in later reproductive years (36–45) suggests cumulative iron depletion effects  
262 that may warrant enhanced clinical attention.

## 263 4.2 Comparison with Literature

264 Our observed IDWA prevalence of 9.0% (95% CI: 8.3%–9.7%) aligns with the lower bound of  
265 estimates reported in recent meta-analyses while highlighting the critical influence of diagnostic  
266 thresholds on prevalence calculations. The landmark meta-analysis by Hamarsha et al. reported  
267 a pooled prevalence of 19% (95% CI: 19–20%) for iron deficiency at the <15 µg/L threshold  
268 among premenopausal women, nearly double our observed estimate (11). This discrepancy  
269 may reflect temporal trends in iron status, population demographic differences, or variations in  
270 laboratory methodologies across studies.

271 Importantly, Hamarsha et al. demonstrated the dramatic sensitivity of prevalence estimates  
272 to threshold selection, reporting 49% prevalence at <30 µg/L and 70% at <50 µg/L thresholds  
273 (11). These findings underscore that our 9.0% estimate using the WHO <15 µg/L threshold  
274 likely represents a conservative lower bound, with the true burden of iron depletion potentially  
275 affecting 35–40% of women when employing physiologically-based thresholds around 25 µg/L  
276 (9; 10).

277 The physiologically-based threshold literature provides crucial context for interpreting our  
278 findings. Petry et al., in their analysis combining the REDS-RISE donor study with NHANES  
279 2003–2018 validation data, identified a threshold of approximately 25 µg/L using soluble trans-  
280 ferrin receptor and hemoglobin indicators (9). Their NHANES validation cohort demonstrated  
281 remarkable consistency with donor populations, supporting generalizability of the ~25 µg/L  
282 threshold. Similarly, Mei et al., analyzing NHANES III data, derived consensus thresholds of  
283 22.5–24.8 µg/L (10). These findings suggest that current WHO guidelines may identify iron  
284 deficiency only at relatively advanced stages of depletion.

285 Our finding of a 6.4% higher ferritin level among supplement users appears modest com-  
286 pared to effect sizes reported in randomized controlled trials. The landmark RCT by Vaucher  
287 et al. demonstrated substantially larger effects: ferritin increased by 6.8 µg/L at 6 weeks and  
288 11.4 µg/L at 12 weeks with 80 mg/day iron compared to placebo (4). Several factors explain this  
289 discrepancy. First, cross-sectional measurement cannot capture temporal dynamics—we cannot

290 determine whether supplement users initiated supplementation due to documented deficiency  
291 or for prophylactic reasons. Second, self-reported supplement use introduces measurement error  
292 including inaccurate recall and uncertainty regarding duration. Third, RCTs typically enroll  
293 participants with documented deficiency (ferritin <50 µg/L), where treatment effects are max-  
294 imized, whereas our population-based sample includes women across the ferritin spectrum.

295 Our dose-response finding of strongest effects at moderate doses (18–27 mg/day) initially  
296 appears paradoxical given the FORTE trial’s findings favoring higher 60 mg doses (12). How-  
297 ever, several explanations reconcile these observations. First, the moderate dose category in  
298 our analysis aligns closely with standard over-the-counter formulations, which may represent  
299 chronic, sustained use patterns. Higher dose supplements may indicate intermittent use, pre-  
300 scription iron for documented deficiency with subsequent discontinuation, or shorter durations  
301 limiting cumulative effect detection. Second, Stoel et al. demonstrated that alternate-day dosing  
302 with high-dose iron achieved comparable ferritin repletion to consecutive-day dosing but with  
303 56% fewer gastrointestinal side effects (13). The moderate-dose group may represent optimal  
304 adherence patterns that maximize long-term bioavailability.

305 The concept of hepcidin-mediated absorption blocking provides mechanistic insight into our  
306 observed dose-response pattern. High-dose iron administration acutely elevates hepcidin levels,  
307 suppressing intestinal iron absorption for 24–48 hours. Consequently, moderate daily dosing may  
308 achieve superior cumulative absorption compared to daily high-dose regimens, despite lower per-  
309 dose quantities. Our finding that moderate doses show the strongest association may reflect this  
310 biological reality: 18–27 mg daily may represent an optimal balance that elevates iron status  
311 without triggering substantial hepcidin-mediated absorption inhibition.

312 Demographic patterns in our findings align with established disparities in iron status. Our  
313 finding of highest IDWA prevalence among Mexican American women (11.6%) aligns with prior  
314 NHANES-based studies documenting higher iron deficiency prevalence among Mexican Amer-  
315 ican women (14; 15). The observation of lower IDWA prevalence among non-Hispanic Black  
316 women (6.5%) compared to non-Hispanic White women (8.8%) appears counterintuitive given  
317 prior reports of higher anemia prevalence among Black women. However, this pattern may re-  
318 flect complex interactions between iron deficiency and anemia etiologies, including higher rates of  
319 hemoglobinopathies and chronic inflammation affecting hemoglobin independent of iron status.

320 The age-related pattern showing peak IDWA prevalence among women aged 36–40 years  
321 (10.3%) supports the cumulative iron depletion hypothesis. Mauracher et al. demonstrated that  
322 menstrual blood loss accounts for approximately 8% of explained variance in both hemoglobin  
323 and ferritin levels, with heavy menstrual bleeding associated with three-fold increased odds of

324 anemia (16). Across the reproductive lifespan, cumulative menstrual losses compound with  
325 parity-related iron depletion to produce highest deficiency risk in later reproductive years.

### 326 **4.3 Clinical and Public Health Implications**

327 Applying our observed 9.0% IDWA prevalence to the US population of non-pregnant women  
328 aged 18–45 years suggests that 2.5 million women are affected by IDWA using WHO thresholds.  
329 However, if we apply physiologically-based thresholds around 25 µg/L, the affected population  
330 likely expands to 10–12 million women—a four-fold increase with profound implications for  
331 healthcare resource allocation (2). This expanded estimate aligns with recent estimates that  
332 38% of non-pregnant reproductive-age women have iron deficiency without anemia (2).

333 The symptom burden associated with IDWA amplifies its public health significance. Vaucher  
334 et al. demonstrated a 47.7% reduction in fatigue with iron supplementation compared to 28.8%  
335 with placebo—representing substantial quality-of-life improvement (4). Applied to our estimated  
336 2.5–12 million affected women, these effect sizes suggest that evidence-based iron supplementa-  
337 tion could meaningfully improve symptoms for millions of US women currently without access  
338 to diagnosis or treatment.

339 Our findings support reconsideration of the WHO <15 µg/L ferritin threshold. Women with  
340 ferritin levels between 15–25 µg/L—classified as “normal” by current WHO guidelines—may  
341 experience subtle but clinically meaningful impairments in oxygen transport, exercise toler-  
342 ance, and cognitive function. Clinical guidelines should incorporate this nuance, recommending  
343 heightened attention to women with ferritin in the 15–30 µg/L range.

344 For population-level prevention and treatment, moderate-dose supplementation (18–27 mg/day)  
345 represents a pragmatic approach with several advantages. This dose range is readily available  
346 in over-the-counter formulations, enabling widespread access without prescription requirements.  
347 Moderate doses are associated with fewer gastrointestinal side effects than high-dose regimens,  
348 potentially improving adherence. Daily moderate dosing avoids the hepcidin-mediated absorp-  
349 tion inhibition that may limit bioavailability of high-dose regimens.

### 350 **4.4 Strengths and Limitations**

351 This analysis possesses several notable strengths. First, NHANES provides a nationally rep-  
352 resentative sample with findings directly generalizable to approximately 28 million US women.  
353 The complex multistage probability sampling design, with oversampling of minority populations,  
354 ensures adequate representation of demographic subgroups. Second, the large sample size pro-  
355 vides statistical power to detect modest associations and enables precise prevalence estimation

356 with narrow confidence intervals. Third, objective laboratory measures reduce measurement er-  
357 ror compared to self-reported outcomes. Fourth, inclusion of multiple NHANES cycles enhances  
358 temporal generalizability.

359 Several limitations must be acknowledged. The fundamental limitation is cross-sectional  
360 design, which precludes establishment of temporal relationships and causal inference. We cannot  
361 determine whether supplement users initiated supplementation due to documented deficiency  
362 or for prophylactic reasons. Self-reported supplement use introduces potential measurement  
363 error from inaccurate recall, variability in supplement composition, and uncertainty regarding  
364 duration and adherence. Our analysis relies on single ferritin measurements, which do not  
365 capture intraindividual biological variability in iron stores. Ferritin fluctuates in response to  
366 acute inflammation, recent illness, and menstrual cycle phase.

367 Unmeasured confounding may influence observed associations. Heavy menstrual bleeding, a  
368 major determinant of iron status, is not directly assessed in NHANES and can only be inferred  
369 from proxy measures. Previous history of iron deficiency anemia—a strong predictor of recurrent  
370 deficiency—is not systematically captured. Dietary iron intake is estimated from 24-hour recalls  
371 with known measurement limitations. Genetic factors affecting iron absorption and metabolism  
372 are not assessed.

#### 373 **4.5 Future Research Directions**

374 Prospective longitudinal studies are urgently needed to establish temporal relationships between  
375 iron supplementation and ferritin changes in the IDWA population. Such studies should enroll  
376 women with documented IDWA, randomize to various supplementation strategies, and follow  
377 ferritin trajectories over 3–6 months with frequent measurements. Randomized trials should  
378 explicitly compare moderate-dose (18–27 mg/day) versus high-dose (60–80 mg/day) strategies,  
379 with attention to both efficacy and tolerability. Investigation of alternative dosing strategies  
380 including alternate-day administration and hepcidin dynamics would inform optimal treatment  
381 protocols.

382 Research on higher ferritin thresholds requires resolution through clinical outcomes research.  
383 Studies should assess symptom prevalence, functional impairment, and treatment response across  
384 the ferritin spectrum from 15–50  $\mu\text{g}/\text{L}$  to determine whether physiologically-based thresholds  
385 identify women with clinically meaningful iron deficiency. Economic analyses are needed to  
386 inform policy decisions regarding expanded IDWA screening, comparing current standard-of-  
387 care against expanded strategies including ferritin testing for women with symptoms or risk  
388 factors.

<sup>389</sup> **5 Conclusion**

<sup>390</sup> This nationally representative analysis demonstrates that iron deficiency without anemia affects  
<sup>391</sup> approximately 9.0% of US non-pregnant women aged 18–45 years when applying WHO fer-  
<sup>392</sup> ritin thresholds, with prevalence estimates expanding substantially when physiologically-based  
<sup>393</sup> thresholds around 25 µg/L are employed. Iron supplementation is associated with modestly  
<sup>394</sup> higher ferritin levels (6.4% increase), with moderate-dose supplementation (18–27 mg/day) show-  
<sup>395</sup> ing the strongest associations.

<sup>396</sup> The clinical significance of IDWA extends beyond the biochemical abnormality of low fer-  
<sup>397</sup> ritin to encompass measurable symptom burden affecting millions of women during their peak  
<sup>398</sup> productive and reproductive years. Fatigue, cognitive impairment, reduced exercise tolerance,  
<sup>399</sup> and decreased quality of life—documented extensively in clinical trial literature—interfere with  
<sup>400</sup> occupational performance, educational attainment, and family responsibilities.

<sup>401</sup> We urge clinicians, healthcare systems, and policymakers to recognize IDWA as a distinct  
<sup>402</sup> clinical entity requiring systematic attention. For clinicians, we recommend screening symp-  
<sup>403</sup> tomatic women with ferritin testing regardless of hemoglobin status, considering higher diagnos-  
<sup>404</sup> tic thresholds (25–30 µg/L), and recommending moderate-dose supplementation (18–27 mg/day)  
<sup>405</sup> for women with documented deficiency. For healthcare systems, we recommend expanding labo-  
<sup>406</sup> ratory panels to include ferritin in standard preventive health assessments and updating clinical  
<sup>407</sup> guidelines to explicitly address IDWA with evidence-based treatment protocols.

<sup>408</sup> By expanding screening, reconsidering diagnostic thresholds, and implementing evidence-  
<sup>409</sup> based supplementation strategies, the healthcare community can meaningfully improve the  
<sup>410</sup> health and quality of life for millions of American women currently suffering from preventable  
<sup>411</sup> iron deficiency. The time has come to move IDWA from an incidental laboratory finding to a  
<sup>412</sup> recognized clinical indication for evaluation, treatment, and prevention.

413 **References**

414 **References**

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