

Optimal Feature-Based Market Segmentation and Pricing

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In this work, we study semi-personalized pricing strategies where a seller uses features about their customers to segment the market, and customers are offered segment-specific prices. In general, finding jointly optimal market segmentation and pricing policies is computationally intractable, with practitioners often resorting to heuristic segment-then-price strategies. In response, we study how to optimize and analyze feature-based market segmentation and pricing under the assumption that the seller has a trained (noisy) regression model mapping features to valuations. First, we establish novel hardness and approximation results in the case when model noise is independent. Second, in the common cases when the noise in the model is log-concave, we show the joint segmentation and pricing problem can be efficiently solved, and characterize a number of attractive structural properties of the optimal feature-based market segmentation and pricing. Finally, we conduct a case study using home mortgage data, and show that compared to heuristic approaches, our optimal feature-based market segmentation and pricing policies can achieve nearly all of the available revenue with only a few segments.

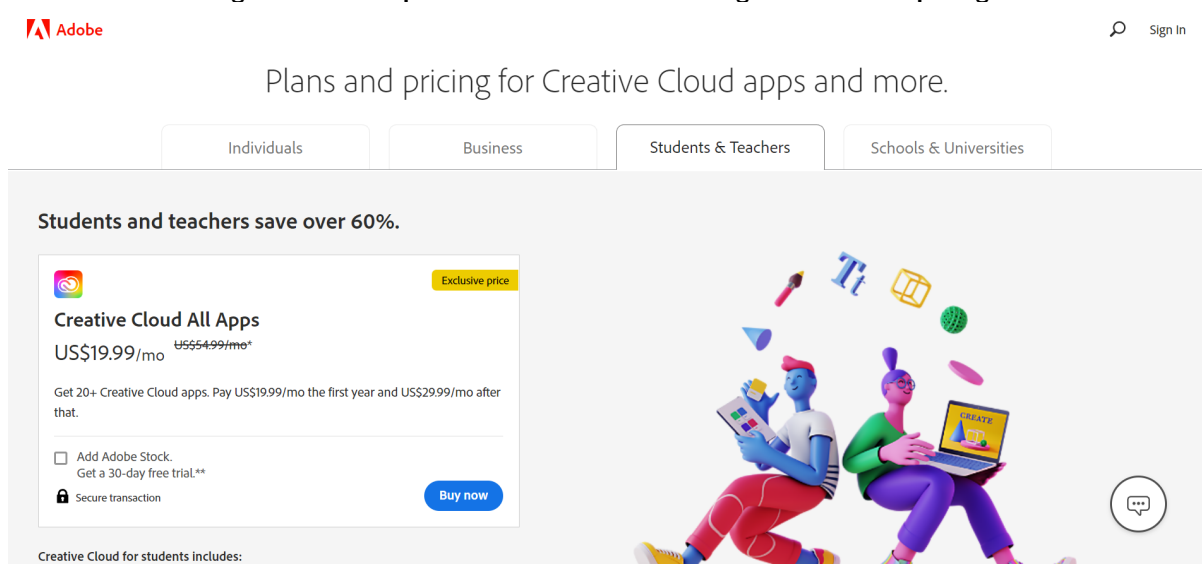
Key words: market segmentation, personalized pricing, third degree price discrimination, regression

History: July 7th, 2022.

1. Introduction

Third-degree price discrimination concerns the now ubiquitous practice of selling a good at different prices for different consumers (Varian 1989). For example, in the sale of proprietary software licenses, prices may differ based on whether the customer is a student, professional, or corporate user (see Fig. 1 for an example and Lehmann and Buxmann (2009) for an extended discussion). In insurance markets, firms gather extremely rich and nuanced feature information about their customers, ostensibly to estimate risk, but which is also leveraged to price discrimination on the basis of demographic and/or geographic information (Buzzacchi and Valletti 2005, Thomas 2012). In commercial markets facing walk-in customers, firms are comparatively limited in their information about their customers but can still profitably leverage feature information to price discriminate, for instance in movie theater ticket sales the customers' age (child, adult, senior) and the time of screening (weekend, matinee) can be used to issue semi-personalized discounts (Dubé et al. 2017).

Figure 1 Example of feature-based market segmentation and pricing.



Note. An example of feature-based market segmentation and pricing for Adobe Creative Cloud products (see <https://www.adobe.com/creativecloud/plans.html>). Here customers are segmented based on their attributes (i.e. student versus professional) and prices vary based on the segment customers are in, with the student rate being 60% of the professional rate (note discounts are enforced by requiring a valid .edu email). Further, note the the number of segments, k , is only four. It will be informative to think of k as ≈ 4 in this work.

Each of these markets vary in both the quality and descriptive power of the information they have about their customers, as well as in the operational difficulty of setting and changing their prices, and thus implementing price discrimination. When the information

about customers in the market is of low quality or is largely censored, fully personalized price discrimination where each customer is charged a personalized price may be undesirable, however as mentioned above, that does not preclude the use of some modest price discrimination via market segmentation. In fact, even when information is richly textured and pricing is largely unconstrained by legal and/or operational considerations, a small static set of prices based on customer features is still often preferred to fully personalized pricing. A small set of market segments and prices is conceptually and operationally simple to implement, and a surprisingly small number of options is often sufficient to achieve strong revenue (Courty and Pagliero 2012). In this paper, we propose and study a general framework for implementing limited price discrimination capturing these trade-offs between predictive power and operational flexibility we term feature-based market segmentation and pricing (FBMSP).

Finding and optimizing generic market segmentation and pricing policies is a well-studied problem in industry with academic roots in operations research/management, marketing, economics, and computer science. However, archetypal formulations of the segmentation and pricing problem are well known to be intractably hard (Kleinberg et al. 1998, 2004). To deal with this hardness, much of the literature has taken a heuristic approach to the problem (Claycamp and Massy 1968, Assael and Roscoe Jr 1976, Chen 2001, Liu et al. 2010, Li and Qiu 2014), separating the segmentation and pricing components. Segmentation then pricing procedures use tools from unsupervised learning to first identify consumer segments/clusters with similar features, and then identify revenue optimal pricing for those chosen market segments. In this paper, we advance the study of segmentation and pricing by finding *jointly optimal* segmentation and pricing's under some realistic assumptions about how firms leverage feature information to predict customers' valuations. Specifically, in practice a firm's valuation model i.e. the model that maps features to valuation or a proxy for willingness-to-pay, is built using regression. Regression models come with their own theory and standard set of assumptions that we profitably utilize to study market segmentation and pricing as well. We show that by leveraging the assumptions of independence and log-concavity of residuals in the regression model, the resulting revenue-maximizing feature-based market segmentation and pricing (FBMSP) enjoys a simple, intuitive structure and can be computed efficiently. Further, our structural results allow us to analyze optimal FBMSP and derive new insights operational and managerial insights about such

policies, including guidance for choosing the number of segments and conditions for when a segmentation and pricing is near-optimal.

1.1. Our Contributions

To summarize our contributions:

1. We first study the algorithmic problem of finding the optimal FBMSPP. In general, the problem is intractably hard, so we focus our attention on the case when valuations are predicted according to a regression model with independent residuals. We show that with no additional assumptions, while we can prove some promising structural properties (c.f. Lemma 1) and provide a $(1 - 1/e)$ approximation algorithm for the optimal segmentation and pricing (c.f. Remark 1), unfortunately finding the optimal FBMSPP is still NP-Hard to compute (c.f. Theorem 1). However, when we further assume the residuals are log-concave, as is often the case in practice, we are able to evade our hardness result. Specifically, when residuals are independent and log-concave, we prove the optimal policy has a simple *interval* structure which allows us to compute in it cubic time via dynamic programming (c.f. Theorem 2).
2. We next turn our attention to analyzing the performance of optimal feature-based market segmentation and pricing. Specifically, we consider the practical operational question of how to choose the number of segments k so as to guarantee minimal loss against a fully personalized pricing benchmark. We show three results that can help guide practitioners in choosing k . First, we show that an upper bound on the loss against personalized pricing can be achieved by simply examining the loss in the model, ignoring the noise term (c.f. Theorem 3). Next, we tightly upper bound the optimal rate at which FBMSPP tends to personalized pricing as a function of the number of segments k and some valuation parameters (c.f. Theorem 4). Finally, we show that the revenue of FBMSPP is concave in k (c.f. Theorem 5). Taken together, these three results allow a practitioner to use their regression model (without reference to the complicating error!) to find k via a simple elbow method, and feel confident that the results of such a heuristic are provably close to optimal.
3. Finally in Section 5, we demonstrate our method on real housing loan data collected in Pennsylvania in 2020, and compare its performance against standard segment-then-price methodologies. We find our approach significantly outperforms heuristic methods, especially when the number of segments is small and the variation in the

valuations comes primarily from variation in $\mu(\cdot)$, as opposed to variation from the prediction error ϵ . We also note that the segmentations found by our approach are qualitatively different than those in segment-then-price, with our approach quickly isolating key differences between groups whereas heuristic approaches can get bogged down in pointless discrimination between groups until it discovers the important differences for the revenue.

1.2. Literature Review

Our work is influenced by, and contributes to, several streams of literature across operations management, marketing, and computer science. Here we overview some of these streams and connect them to our work.

Theory of Price Discrimination There is a deep extensive literature on the theory of pricing discrimination beginning in economics and spanning operations management, marketing, and computer science. Much of the classic literature in this area (Schmalensee (1981), Narasimhan (1984), Katz (1984), Varian (1985), Shih et al. (1988), Bergemann et al. (2015), Cowan (2016), Xu and Dukes (2016)) focuses on the impact of price discrimination on social welfare, or the effects of price discrimination on the resultant equilibrium prices. In this paper, we investigate market segmentation and pricing exclusively from the perspective of a revenue-maximizing monopolist, focusing on computational/practical implementation of such policies.

Specifically, in the language of (Varian 1985) we study third degree price discrimination. In practice, firms engage most often in third-degree price discrimination, which can use additional information about consumer features to offer different prices to different implicit or segments in a multitude of ways (Su (2007), Jerath et al. (2010), Besbes and Lobel (2015), Chen et al. (2005), Cohen et al. (2017), Elmachoub and Hamilton (2021)). Several papers have analyzed the value of such price-discrimination tactics compared to uniform pricing (Huang et al. (2019), Elmachoub et al. (2021)). In contrast, we investigate the value of the optimal feature-based market segmentation and pricing in this paper, and compare these semi-personalized pricing’s against a fully personalized benchmark.

Regression Based Price Discrimination In recent years, data-driven pricing strategies have become increasingly common (Chen et al. (2015), Ferreira et al. (2016), Shukla et al. (2019), Aouad et al. (2019), Elmachoub et al. (2020), Niu et al. (2020), Biggs et al. (2021),

Elmachtoub and Grigas (2022)). In these works, customers are offered a personalized price based on features that are predictive of their valuation of the product, especially by tree-based prescriptive approaches (Athey and Imbens (2016), Kallus (2017), Bertsimas et al. (2019), Biggs et al. (2021)). Unlike most data-driven pricing literature, here, we ignore how the regression model is found and instead take the prediction of customer’s valuation as input, and analyze how it may be profitable leveraged to compute and analyze optimal FBMSP.

Algorithms for Market Segmentation and Pricing Our paper contributes to a line of literature studying the market segmentation and pricing from a algorithmic/computational complexity perspective. Indeed many models of joint market segmentation and pricing are known to be intractably hard to compute going back at least to the pioneering work of Kleinberg et al. (1998, 2004), restricting their applicability in practice. Often in marketing to evade these hardness results the segmentation and pricing decisions are made sequentially instead of being evaluated together (Dolgui and Proth (2010)), and at first blush it seems that Theorem 1 implies our model, for all the structure gained through independence, is ultimately no better. Fortunately, we will see for almost all regression models in practice our model makes jointly optimal segmentation and pricing tractable and well structured.

If the regression error is log-concave, as we assume in Section 3.2, computing the optimal feature-based segmentation is structurally similar to the *1D Clustering* problem for which dynamic programming approaches have been employed (see Gronlund et al. (2017) for a modern overview), and can be solved in polynomial time. Other algorithmic approaches for feature-based pricing can be seen in Cohen et al. (2016), Qiang and Bayati (2016), Javanmard and Nazerzadeh (2016), albeit in different models.

1.3. Paper Outline

The remainder of this paper is organized as follows. In Section 2 we introduce our model for feature-based market segmentation and pricing and provide some preliminary structural results. In Section 3 we study the problem of computing the revenue-optimal FMBSP. In Section 4 we analyze the structure of revenue-optimal FBMSP and provide some theory to guide practitioners in choosing the number of segments, k . In Section 5 we demonstrate our approach on a well known Home Mortgage Disclosure Act dataset. Finally, in Section 6

we provide concluding remarks and highlight future directions. All examples and proofs referenced in the main body can be found in Section A and Section B in the Appendix, respectively.

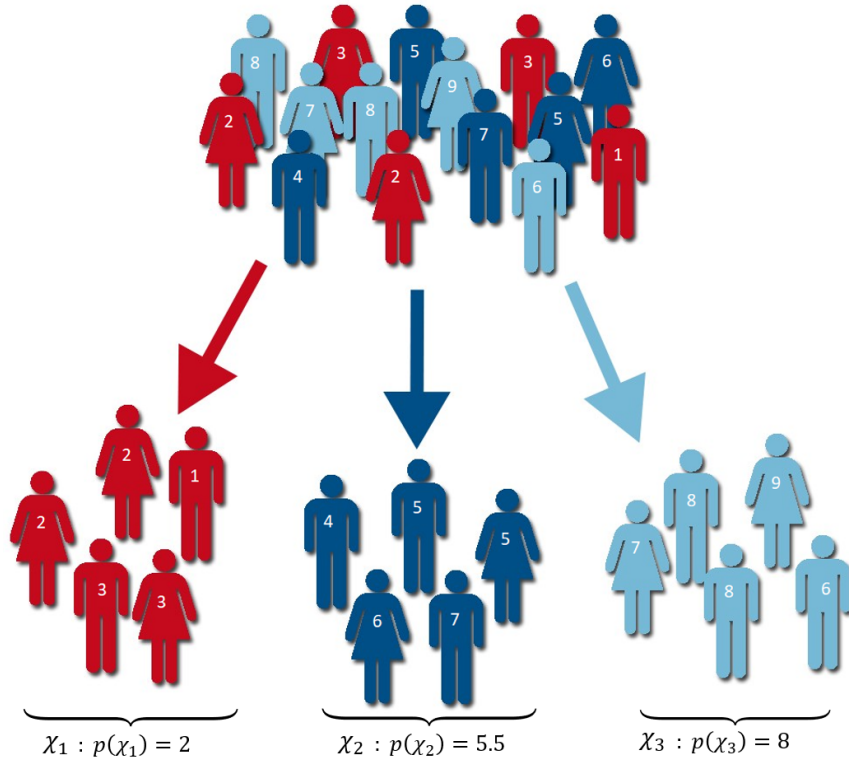
2. Model and Preliminaries

We consider a revenue-maximizing seller offering a good in unlimited supply. For simplicity of presentation, we will assume the good is produced costlessly and so revenue and profit are equivalent (we note the model presented in this paper and all results do however easily extend to the case when each good has a per unit cost c). We will further assume each customer in the market is described by some feature vector \mathbf{x} of their observable characteristics, and has some valuation for the good which depends on their feature vector, $V|\mathbf{x}$. The market characteristics as a whole can be described as a distribution over the feature vectors $\mathbf{X} \sim F_{\mathbf{X}}$, which is supported on some feature space $\mathcal{X} := \text{supp}(\mathbf{X})$. These features vectors can consist of any information about the customers, including demographic information like gender, household status, income etc.

In line with modern practice, we model the seller as having trained some regression model $\mu : \mathcal{X} \rightarrow \mathbb{R}^+$ to predict a customers valuation for a good from their feature vector. We assume the regression model has residual error ϵ but is correct in expectation, so that the predicted valuation for a customer with features \mathbf{x} is $\mu(\mathbf{x}) := \mathbb{E}[V|\mathbf{x}]$, and the valuation model is $V = \mu(\mathbf{X}) + \epsilon$. We will use F to be the distribution of the valuations V , $F_{\mathbf{X}}$ to be the distribution of the feature vectors, F_{ϵ} to be the distribution of the error term ϵ , and $f_{\mathbf{X}}$, f_{ϵ} , and f to be the densities, respectively. We will use \bar{F} to denote the survival function, i.e., $\bar{F}(x) := 1 - F(x)$.

For a seller with a valuation model $\mu(\cdot)$, we will study the revenue achievable by selling strategies where the feature space of the market, \mathcal{X} , is partitioned into k segments $\{\mathcal{X}_i\}_{i=1}^k$, $\cap \mathcal{X}_i = \emptyset$, $\cup \mathcal{X}_i = \mathcal{X}$, such that on each segment the seller offers a distinct price $p_{\epsilon}(\mathcal{X}_i)$. Now we are ready to define feature-based market segmentation and pricing strategies, which is the main object of this study.

Feature-Based Market Segmentation and Pricing: In feature-based market segmentation and pricing the seller partitions the feature space into k segments $\{\mathcal{X}_i\}_{i=1}^k$, and

Figure 2 An example of FBMSP.

Note. Depicted is an example of feature-based market segmentation. For each customer some numeric prediction of their valuation is given. The feature space \mathcal{X} consists of all combinations of the color and gender for the customer, and the feature-based market segmentation leverages color (not necessarily optimally) to sort them into three segments \mathcal{X}_i , $i \in [3]$, each with a distinct segment level price, $p(\mathcal{X}_i)$.

on each segment offers a single price $p(\mathcal{X}_i)$ (see Fig. 2 for example). The expected profit of such a segmentation is,

$$\mathcal{R}_{kXP}(\{\mathcal{X}_i\}_{i=1}^k, \{p_\epsilon(\mathcal{X}_i)\}_{i=1}^k) := \sum_{i=1}^k p_\epsilon(\mathcal{X}_i) \int_{\mathbf{x} \in \mathcal{X}_i} \Pr(\mu(\mathbf{x}) + \epsilon \geq p_\epsilon(\mathcal{X}_i)) f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}. \quad (1)$$

Given a segmentation it will often be convenient to think of the prices as the revenue optimal ones for that segment. To that end we denote the optimal price on segment \mathcal{X}_i by $p_\epsilon(\mathcal{X}_i)$ i.e.,

$$p_\epsilon(\mathcal{X}_i) := \arg \max_p \int_{\mathbf{x} \in \mathcal{X}_i} \Pr(\mu(\mathbf{x}) + \epsilon \geq p) f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}.$$

We will use $\mathcal{R}_{kXP} := \max_{\mathcal{X}_1, \dots, \mathcal{X}_k} \sum_{i=1}^k \mathcal{R}_{kXP}(\{\mathcal{X}_i\}_{i=1}^k, \{p_\epsilon(\mathcal{X}_i)\}_{i=1}^k)$ denote the optimal profit for a feature-based market segmentation and pricing strategy.

Note that our framework for feature-based market segmentation and pricing is very flexible, and captures many well studied models as a special case. For instance when $k = 1$,

FBMSP is just the revenue generated by a single static price for the good (sometimes referred to as the revenue of the monopoly price, or posted price, or single price). Similarly when the number of segments is very large, i.e. $k \rightarrow \infty$, FBMSP becomes the revenue of feature-based personalized pricing, where each customer receives an appropriately personalized price. The revenue of feature-based personalized pricing is a useful upper bound to compare with the revenue of optimal FBMSP for some fixed number of segments k . In Section 4 we consider the question of how large does k need to be, in general, to approach the revenue of personalized pricing, with the hope that a reasonably small k should suffice (see Elmachetoub et al. (2021) for a detailed discussion of when feature-based personalized pricing is provably close or far from the revenue of a single price).

Similarly, when the error distribution ϵ is 0 almost surely (a.s.), our model represents the achievable revenue in the world of the prediction model μ without regard to the models potential error. We term this optimistic case *model market segmentation*, and in Section 4 will show that reasoning about the profit in a world of perfect prediction can provide a useful upper bound on the loss of FBMSP with error.

2.1. Key assumptions and preliminaries.

As mentioned in the introduction, the optimal feature-based market segmentation and pricing is generally hard to compute. To ensure tractability, in our work we will carry through the common regression assumption that the model error, ϵ , is independent across features i.e. $\mathbf{X} \perp \epsilon$. We consider this assumption to be quite mild, as it underlies many predictive models used in practice, including for example, the well known logit model where a customer's valuation is a linear combination of that customer's features, the offered price, and an idiosyncratic error following a logistic distribution which is *independent* of \mathbf{X} . Similar remarks hold for other regression-based models with independent errors. The upshot will be that this necessary assumption for regression is also quite harmonious with pricing and gives considerable structure and control for analyzing pricing models.

In the next section we will delve into the structure of optimal FBMSP, but first we will illustrate how the independence assumption smooths our problem by considering some related objectives. To this end, consider three auxiliary functions that will later be helpful in analysis of FBMSP, and also are of independent interest.

$$\textbf{Price: } p_\epsilon(x) := \inf_p \{\arg \max_p p \overline{F}_\epsilon(p - x)\}, \quad \textbf{Margin: } \theta_\epsilon(x) := p_\epsilon(x) - x,$$

$$\textbf{Revenue: } \mathcal{R}_\epsilon(x) := \max_p p \Pr(x + \epsilon \geq p) = p_\epsilon(x) \bar{F}_\epsilon(\theta_\epsilon(x)).$$

The price, margin, and revenue functions all serve to model a seller pricing a good for a customer after having predicted their valuation as $x \in \mathbb{R}$, up to some stochastic error ϵ . $p_\epsilon(x)$ is the optimal price to offer a customer with predicted valuation x , $\mathcal{R}_\epsilon(x)$ is the revenue of the optimal monopoly price when the valuation distribution is $x + \epsilon$, and $\theta_\epsilon(x)$ is the difference or *margin* between the predicted valuation x and the offered price $p_\epsilon(x)$. Note, $p_\epsilon(x)$ is uniquely defined to be the minimum price that achieves the maximum revenue, such a minimum is necessary since for some distributions ϵ , there may be many prices that maximize the revenue (see Example EC.1, for an extensive discussion on when the optimal price is unique, or equivalently when the revenue function is strictly unimodal, see Ziya et al. (2004)).

In the following lemma we summarize some of the structure we observe in these functions simply from assuming independence of residuals.

LEMMA 1 (General Properties of $p_\epsilon(\cdot), \theta_\epsilon(\cdot), \mathcal{R}_\epsilon(\cdot)$). *Suppose that $V = \mu(\mathbf{X}) + \epsilon$ where $\mathbf{X} \perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Then the following properties hold:*

- (a) $\theta_\epsilon(x)$ is an decreasing function.
- (b) For any $0 < x_1 < x_2$, we have

$$\bar{F}_\epsilon(\theta_\epsilon(x_1))(x_2 - x_1) \leq \mathcal{R}_\epsilon(x_2) - \mathcal{R}_\epsilon(x_1) \leq \bar{F}_\epsilon(\theta_\epsilon(x_2))(x_2 - x_1).$$

Moreover, for all x such that $p_\epsilon(x)$ is continuous (i.e. $p_\epsilon(x^-) = p_\epsilon(x^+)$), the derivative of $\mathcal{R}_\epsilon(x)$ exists and $\frac{d}{dx} \mathcal{R}_\epsilon(x) = \bar{F}_\epsilon(\theta_\epsilon(x))$.

- (c) $\mathcal{R}_\epsilon(x)$ is increasing, continuous, and convex.

Lemma 1 states that independence alone induces prices that result in a monotone increasing sales probabilities via $\theta_\epsilon(x)$, and a convex revenue function $\mathcal{R}_\epsilon(x)$ with interpretable, bounded derivatives. All three parts of the lemma are proved by examining the induced optimal prices, $p_\epsilon(x)$, and noting that $p_\epsilon(x)$ cannot increase very quickly (i.e. super linearly). Unfortunately, our control is not perfect as $p_\epsilon(x)$ can otherwise be quite poorly behaved; there can be many optimal prices for a given valuation, and worse $p_\epsilon(x)$ can be discontinuous in x at arbitrarily many points x , see Example EC.1 for an example. The jump discontinuities in $p_\epsilon(x)$ translate directly to non-differentiable points in the revenue

function. As we will see in Section 3, the structure provided by independence is not quite enough to enable the efficient computation of the optimal FBMS, however it will be critical in our analysis of such policies.

3. Computing Optimal Feature-Based Market Segmentation and Pricing

In this section we study the problem of finding the jointly optimal FBMS, culminating with conditions and an algorithm under which the optimal policy can be computed efficiently. We will first show that when valuations are drawn from a regression model with general independent residuals, the optimal policy is NP-Hard to compute. We then identify that the hardness stems from some pathological segmentation properties, and define a natural property to characterize nice segmentations which we call *interval*. Our main positive result of this section is to show that when the residuals are log-concave the optimal segmentations are interval, and further, the optimal interval segmentations can be found in cubic time via dynamic programming. Thus, for realistic valuation models under standard regression assumptions, the jointly optimal segmentation and pricing can be directly computed instead of having to resort to heuristic segment-then-price approaches.

3.1. Hardness of FBMS

To understand some of the difficulty of FBMS, in this subsection we will review some standard hardness results and show that even under the assumption of independent residuals, the problem remains intractable. Our proof of this hardness result will yield guiding intuition for how an additional condition of log-concavity on the residuals should be computationally useful.

In general, the hardness of market segmentation and pricing problems typically follows from a reduction to hitting set (Kleinberg et al. 1998, 2004). The correspondence between the two problems often works as follows: imagine you have n customers each with valuations described by an independent distribution F_i such that $p\bar{F}_i(p)$ is maximized by some discrete set of optimal prices S_i . Then the best case market segmentation and pricing problem is simply to find a partition of customers in k segments such that on each segment, the intersection of each customer's optimal price set is non-empty. This implies then that there is an obvious optimal price for each segment that clearly maximizes the revenue by construction, and so the hardness is merely to find the partition with such a property. This is exactly the difficulty in k -hitting set, and so follows the reduction.

Note, in the above that the assumption of unique valuation distributions for each customer, F_i , is crucially important in setting up the correspondence with hitting set. In our model, where valuations are described by a common regression function with independent noise, it is no longer clear if such a construction is possible. That is, now for customer i with features \mathbf{x}_i , their valuation distribution is $V|\mathbf{x}_i = \mu(\mathbf{x}_i) + \epsilon$. As we show in the following theorem, the problem remains NP-hard in this case, although the proof requires a significantly more intricate construction of the common error distribution ϵ .

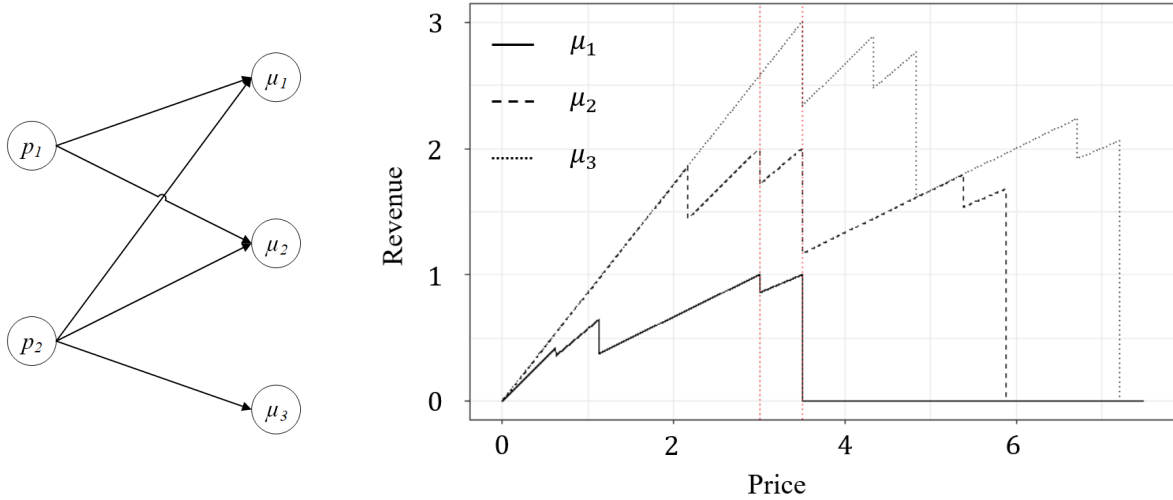
THEOREM 1 (Hardness of FBMSP). *Suppose that $V = \mu(\mathbf{X}) + \epsilon$, $\mathbf{X} \perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Then finding the optimal FBMSP policy is NP-hard.*

Sketch of Proof of Theorem 1 The proof of Theorem 1 follows by reduction to hitting set, as in the general case. For every instance of the hitting set problem, we show that there exist estimations of customers' valuation and prediction error's, such that deciding if there are k (or less) elements that hits all the subsets is equivalent to deciding whether there are k segments and prices such that the total revenue is $\frac{n(n+1)}{2}$, where n is the number of subsets in the hitting set problem (equivalently, the number of customers in the market). Our construction follows by designing a error distribution ϵ which results in a number claw like functions for each customer, that are then spread by translation to encode a set of optimal prices for each valuation level $\mu(\mathbf{x}_i)$. Fig. 3 gives an example of our hardness reduction for a small instance. \square

Theorem 1 implies it's impossible to solve general FBMSP efficiently if $P \neq NP$. This leaves us two options, either to look for approximate solutions for general FBMSP or to enrich the structure of our model by imposing additional assumptions. We briefly explore the former in Remark 1, but will focus mainly on the later.

REMARK 1. While an optimal policy for general error distribution cannot be found in polynomial time, we note that a constant factor approximation to the optimal feature-based market segmentation and pricing is obtainable. Specifically, a $(1 - 1/e)$ factor approximate policy can be found in polynomial time since the objective function is positive valued, monotone, and submodular. We formalize this observation in Section D in the appendix.

\square

Figure 3 Example of hitting set and FBMSP.

Note. In the left panel is a graph representation of a small instance of a hitting set problem. To illustrate the translation of a hitting set problem to a FBMSP problem, assume we have 3 customers with predicted valuation $\mu_1 = 3$, $\mu_2 = \frac{43}{8}$, $\mu_3 = \frac{161}{24}$. Further, let the estimation error ϵ be supported on $\{-\frac{161}{24}, -\frac{77}{5}, -\frac{19}{8}, -\frac{15}{8}, 0, \frac{1}{2}\}$ with probability mass $\{\frac{1}{7}, \frac{4}{21}, \frac{2}{21}, \frac{5}{21}, \frac{1}{21}, \frac{2}{7}\}$. Finally define $p_1 = 3$ and $p_2 = 3.5$. In the right panel we plot the revenue curves for each valuation. We can see that the revenue for each customer is maximized only at either p_1 or p_2 which represent the connections between price nodes and valuation nodes in left panel, i.e., revenue from customer 3 is maximized at price $p_2 = 3.5$, revenues from customer 2 and 1 are both maximized at $p_1 = 3$ and $p_2 = 3.5$.

3.2. Feature-Based Market Segmentation and Pricing with Log-Concave Residuals

In the previous subsection we studied FBMSP in a setting where the underlying error in the valuation model was arbitrary, and in the proof of Theorem 1 we leveraged this freedom to construct a general error distribution such that it induced jagged, delicately overlapped revenue functions that made the problem intractable. In this section we will consider assumptions that evade such pathological constructions. To that end, recall in Lemma 1 we were able to characterize many things about the revenue function $\mathcal{R}_\epsilon(x)$, but less about the structure of the pricing function which noted could vary (drop) dramatically between similar valuations. It is precisely these discontinuities in $p_\epsilon(x)$ that enable our construction in Fig. 3, and give it its discrete quality that makes it difficult to optimize. A natural question then to ask is, for suitably smooth error distributions/revenue functions, is it still hard to compute the optimal FBMSP?

We will make one additional assumption about the error distribution that enforces such a notion of smoothness. Namely, we will assume the distribution is *log-concave*, a canonical assumption in the pricing and revenue management literature. Note many standard distributions are log-concave including normal, exponential, uniform distributions, etc.

DEFINITION 1 (LOG-CONCAVE ERROR). A random variable ϵ with density f_ϵ is **log-concave** if $\log(f_\epsilon(x))$ is a concave function.

To understand how log-concavity in the error function translates into tractability for FBMS, we will first show that it implies a continuous, increasing price function $p_\epsilon(x)$, precluding behaviour like in Fig. 3. Leveraging this continuity in the prices, we can then show the optimal segmentation must be well structured in the sense that the segmentation groups together customers with similar predicted valuations for the good. Such segmentations are natural, easy to interpret as low/medium/high/etc. type segments, and as we will show, easy to optimize and analyze. We will call segmentations that group together customers with similar valuations *interval*, and define them as follows.

DEFINITION 2 (INTERVAL SEGMENTATION). We will call an segmentation, $\{\mathcal{X}_i\}_1^k$, an **interval** segmentation if there exists real numbers $0 < s_0 < s_1 < \dots < s_k = \sup_{\mathbf{x}} \mu(\mathbf{x})$ such that each segment \mathcal{X}_i can be written as $\mathcal{X}_i = \{\mathbf{x} | \mu(\mathbf{x}) \in [s_{i-1}, s_i)\}$.

When describing interval segmentations, we will often describe the segmentation by just the end points of the intervals in the valuation space that define them, $\{s_i\}_0^k$. We emphasize that not all optimal market segmentations are interval, certainly the ones induced by the construction in Theorem 1 are not, but also even simple error distributions can have more complicated structure as we demonstrate in Example EC.2. Thankfully, it turns out the smooth notion of error captured by log-concavity, and the intuitive structure of interval segmentations are harmonious notions. In the following lemma we show that log-concavity in the error removes any jump discontinuities from the price function, which in turn allows us to prove that the revenue optimal FBMS is interval.

LEMMA 2 (**Properties of Log-Concave Error**). Suppose that $V = \mu(\mathbf{X}) + \epsilon$ where ϵ is log-concave with density f_ϵ , $\mathbf{X} \perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Then,

- (a) $p_\epsilon(x)$ is an increasing and continuous function.
- (b) The optimal segmentation is interval.
- (c) The price on each segment $p_\epsilon(\mathcal{X}_i) = p_\epsilon(\mu(\mathbf{x}))$ for some $\mathbf{x} \in \mathcal{X}_i$.

Lemma 2 shows that, by assuming the error in the regression model is log-concave, all the previously mentioned pathologies vanish. First, we show that $p_\epsilon(x)$ becomes a strictly increasing function which, combined with Lemma 1, implies that the revenue function $\mathcal{R}_\epsilon(x)$ is differentiable everywhere, and its derivative is simply the sale probability. We

then show in (b) that the upshot of this additional smoothness for FBMSP is that the segmentation policy becomes interval. Moreover, the optimal price to offer on each segment is contained *in the segment*, as the optimal price for some feature vector inside the segment. This locality of the price and segment then enables fast computation of the optimal policy via dynamic programming, as we describe in the main theorem for this section.

THEOREM 2 (Computing Feature-Based Market Segmentation). *Suppose that $V = \mu(\mathbf{X}) + \epsilon$ where ϵ is log-concave with density f_ϵ , $\mathbf{X} \perp\!\!\!\perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Let $n = |\text{supp}(\mu(\mathbf{X}))|$ and suppose $\mathcal{R}_\epsilon(x)$ for fixed x can be computed in time m_ϵ . Then the optimal feature-based market segmentation can be computed in $O(kn^3m_\epsilon)$.*

Theorem 2 is our main result, and states that by leveraging the structural properties in Lemma 2 the optimal policy can be computed quickly and efficiently in terms of the size of the support of the regression model. Note, we assume $\mu(\mathbf{X})$ is finitely supported, and believe this is natural and corresponds to simply running your regression model over a sample of customers. Further, we assume the running time to compute $\mathcal{R}_\epsilon(x)$ as a subroutine is bounded by some number that is related only to ϵ . Again, we believe this assumption is natural since when ϵ is log-concave, the revenue function $p\Pr(x + \epsilon \geq p)$ is concave in p and can be computed simply by checking first order conditions. We also note that when the error distribution is discrete, there is a corresponding notion of *discrete log-concavity* Saumard and Wellner (2014), under which our results continue hold without modification.

In this section, we have characterized when and how we can compute FBMSP optimally, in the subsequent sections we turn our attention to it's efficacy as a revenue management strategy.

4. Analyzing Feature-Based Market Segmentation and Pricing

In the previous section, we studied how to compute the optimal FBMSP under some assumptions, for a given number of segments/prices k . We showed that while, in general, it is hard to do so, in the important and realistic case when error is log-concave, the optimal policy has an intuitive structure that allows for easy computation. In this section, we continue to build on the structural insights of the last section, and show that beyond just computation, optimal FBMSP inherits a number of attractive properties and performance guarantees that may help guide practitioners in implementing such policies, and particularly in deciding how many segments, k , to use.

Throughout this section, when the underlying valuation model varies we will use a superscript to explicitly identify the valuation distribution with which the revenue is computed i.e., $\mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon}$ is the standard revenue of FBMS, $\mathcal{R}_{kXP}^{\mu(\mathbf{X})}$ is the revenue of FBMS with no error in valuation model, and so on.

4.1. FBMS vs. Feature-Based Personalized Pricing

In this subsection, we study the relative gaps between the optimal FBMS and the natural upper bound of feature-based personalized pricing, paying close attention to how this gap informs a good choice of k . As mentioned in the introduction, FBMS closely resembles real-world data-driven semi-personalized pricing strategies where sellers are constrained in the number of the segments/prices they can offer. Specifically, in FBMS the number of prices and segments is capped at k whereas feature-based personalized pricing is equivalent to FBMS when $k \rightarrow \infty$. In fact, for any market where valuations are distributed according $V = \mu(\mathbf{X}) + \epsilon$, the revenue of a seller implementing feature-based personalized pricing can be succinctly described as an expectation over the revenue function i.e., $\lim_{k \rightarrow \infty} \mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon} := \mathbb{E}_{x \sim F_{\mathbf{X}}}[\mathcal{R}_{\epsilon}(\mu(\mathbf{X}))]$, since the seller offers the optimal price for each context \mathbf{x} which garners revenue $\mathcal{R}_{\epsilon}(\mu(\mathbf{x}))$.

Intuitively then, a good choice of k should be one that is not too large, so as to be implementable, but one that is still close to the maximum achievable revenue, i.e., one that shrinks the gap,

$$\mathbb{E}_{\mathbf{X} \sim F_{\mathbf{X}}}[\mathcal{R}_{\epsilon}(\mu(\mathbf{X}))] - \mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon}. \quad (2)$$

One difficulty that may be encountered when attempting to choose k to reduce Eq. (2) is that it is sensitive to the error distribution ϵ , which may be hard to know precisely or require extensive market research to obtain. It may be preferable for an analyst attempting to choose k to work with an upper bound on this difference that is agnostic to the true error distribution. Interestingly, by assuming there is no error an analyst can achieve precisely such an upper bound. In the following theorem we show that one can bound the loss between FBMS and feature-based personalized pricing by examining the gap between the two policies when ϵ is assumed to be 0 a.s. We will refer to this loss as *model market loss*, since it depends only on $\mu(\mathbf{X})$ and not the underlying error distribution.

THEOREM 3 (Model Loss vs. True Loss). *Suppose $V = \mu(\mathbf{X}) + \epsilon$, $\mathbf{X} \perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Then,*

$$\underbrace{\mathbb{E}_{\mathbf{X} \sim F_{\mathbf{X}}}[\mathcal{R}_{\epsilon}(\mu(\mathbf{X}))] - \mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon}}_{\text{Actual Market Loss}} \leq \underbrace{\mathbb{E}_{X \sim F_{\mathbf{X}}}[\mu(X)] - \mathcal{R}_{kXP}^{\mu(\mathbf{X})}}_{\text{Model Market Loss}}.$$

Theorem 3 gives a theoretical foundation through which an analyst can analyze the performance of FBMSP for various k directly in the model without worrying about the particular form of the error distribution. Whatever loss is perceived in the model market bounds the true loss in practice automatically.

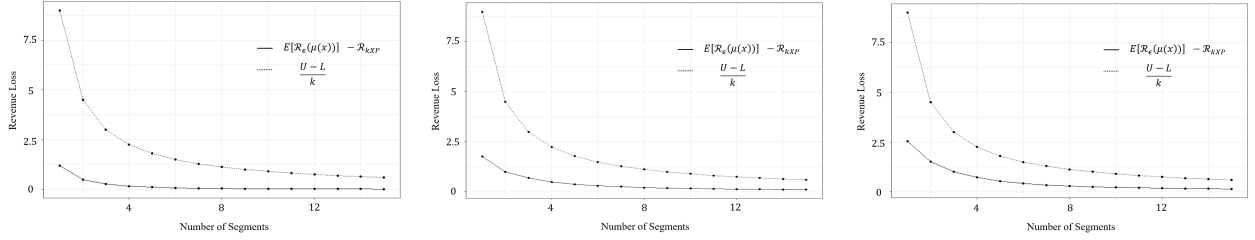
Further, we note that the proof of Theorem 3 is constructive, and implies a simple heuristic for setting feature-based market segmentation and pricing strategies when ϵ is not log-concave, or ϵ is unknown. In these instances, a seller can simply compute the optimal k -FBMSP letting ϵ be 0 a.s. In this situation, the optimal policy is interval and can be described by segmentation end points $\{s_i\}_{i=0}^k$ on the model market $\mu(\mathbf{X})$, which can be used to generate the segments \mathcal{X}_i . From those segments, since ϵ is either unknown or not tractable to work with computationally, the firm can instead perform price experimentation to learn the prices that maximize $p_\epsilon(\mathcal{X}_i) \Pr(s_i + \epsilon \geq p_\epsilon(\mathcal{X}_i))$, and offer that price on each segment. While both the partition into segments $\{\mathcal{X}_i\}_{i=1}^k$, and the prices offered on each segment $\{p_\epsilon(\mathcal{X}_i)\}_{i=1}^k$ may be sub-optimal, such a strategy is guaranteed to earn more than $\mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon} + \mathcal{R}_{kXP}^{\mu(\mathbf{X})} - \mathbb{E}_{X \sim F_{\mathbf{X}}}[\mu(X)]$ by rearranging Theorem 3, and this guarantee smoothly tends to the optimum as the error in the model diminishes.

Theorem 3 allows an analyst to search for a choice of k without referring to the error distribution, a next natural question to ask is then, how long can this search take? That is, at what rate does \mathcal{R}_{kXP} to converge to the revenue of feature-based personalized pricing? In our next theorem we show this convergence is linear in k , and quite fast when the range of valuations is not too wide.

THEOREM 4 (Bounded Loss with k Segments). *Suppose $V = \mu(\mathbf{X}) + \epsilon$, $\mathbf{X} \perp\!\!\!\perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Let $L = \inf_{\mathbf{x}} \mu(\mathbf{x})$ and $U = \sup_{\mathbf{x}} \mu(\mathbf{x})$, then*

$$\mathbb{E}_{\mathbf{X} \sim F_{\mathbf{X}}}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))] - \mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon} \leq \frac{U - L}{k}.$$

The proof of Theorem 4 constructs a (suboptimal) segmentation strategy by equally partitioning the quantile space. Interestingly, the dependence $O\left(\frac{1}{k}\right)$ appears typical for many valuation distributions, as we plot in Fig. 4. Intuitively, this behavior can be explained in the following way. As we segment into smaller pieces, any distribution with a smooth density appears locally uniform on each segment. Example EC.3 establishes that the convergence rate for a uniform matches Theorem 4 up to constant factors, suggesting that, at least for large k , the rate should also be approximately tight for many distributions.

Figure 4 Revenue loss for standard distributions.

Note. Depicted is the revenue loss vs the number of segments for three standard error distributions, and when predicted customer valuations are drawn uniformly from $[1, 10]$ i.e., $\mu(\mathbf{X}) \sim \text{Uniform}[1, 10]$. For simplicity, in the plots, we use \mathcal{R}_{kXP} and $\mathbb{E}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))]$ to denote $\mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon}$ and $\mathbb{E}_{\mathbf{X} \sim F_{\mathbf{X}}}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))]$, respectively. In the left panel, the prediction error $\epsilon \sim \mathcal{N}(0, 1)$. In the middle panel, the prediction error $\epsilon \sim \text{Uniform}[-1, 1]$. In the right panel, the prediction error follows Weibull distribution where the shape parameter is 5 and scale parameter is 1. In each panel, we plot the revenue loss versus the bound $\frac{U-L}{k}$ in Theorem 4. We note the convergence rate for all three error distributions appears to be $\Theta(1/k)$.

4.2. Revenue Concavity in the Number of Segments

Theorem 3 and Theorem 4 give an analyst insight into how to handle the error when searching for k , and a bound on how large a k may be to achieve a desired level of revenue loss. In the final result of this section we will show a nice structural property of the optimal revenue that an analyst can use to further hone their search for k . Specifically, in Theorem 5, we show that when the residual is log-concave, the revenue of FBMSP is concave in the number of segments.

THEOREM 5 (Segmentation Concavity). *Suppose that $V = \mu(\mathbf{X}) + \epsilon$ where ϵ is log-concave with density f_ϵ , $\mathbf{X} \perp \epsilon$, and $\mathbb{E}[\epsilon] = 0$. Then $\{\mathcal{R}_{(k+1)XP} - \mathcal{R}_{kXP}\}_{k=1}^\infty$ is a non-increasing sequence.*

Concavity in terms of k has the operational interpretation that the revenue garnered by additional segments has diminishing marginal returns. Such a property is not guaranteed in general, and especially not for heuristic segment-then-price approaches, as we will see in Section 5. Moreover, combining Theorems 3 and 4 and Theorem 5 together allows an analyst to search for the best choice of k via an *elbow method* (Bholowalia and Kumar 2014) on the model market. Such a method results in a k with a provable guarantee on the loss (Theorem 3), is likely quite small (Theorem 4), and further the elbow will be unique (Theorem 5).

In the next section we implement our segmentation and pricing strategy on real data and illustrate its advantages versus segment-then-price heuristics.

5. Case Study: Setting Mortgage Interest Rates

In Section 3, we showed how to find jointly optimal FBMSF when a firm has trained a regression-based valuation model with independent, log-concave residuals. Then in Section 4, we provided a set of results to aid in the analysis of FBMSF policies and guide the choice of k , the number of segments/prices. In this section, we perform a case study to highlight some features of our approach, which we compare and contrast with prominent heuristic-based approaches for segment-then-price. Specifically, using a real data set of home mortgage offers and acceptances in Pennsylvania in 2020, we build a probit regression model to predict the probability that an applicant will take a mortgage at a given the interest rate. Then, we transform the probit regression model into a model of customer valuation measured as the maximum interest rate they will accept. We then compare our optimal method for FBMSF with segment-then-price (STP) via a number of different simulations on the data set. All data and code for this section are publicly available at [Blinded for Review].

Variable	Type	Description and Statistics
Action taken	Binary	The action taken on the covered loan or application <ul style="list-style-type: none"> • 1 (accepted), Frequency = 11491, Percent = 77.0% • 0 (rejected), Frequency = 3425, Percent = 23.0%
Interest rate	Continuous	The interest rate for the covered loan or application (%) <ul style="list-style-type: none"> • Mean = 3.4%, Std = 0.9%
Income	Continuous	Applicant's gross annual income (in thousands of dollars) <ul style="list-style-type: none"> • Mean = 110.08, Std = 94.76
Derived race	Binary	Single aggregated race categorization derived from applicant race fields <ul style="list-style-type: none"> • 1 (white), Frequency = 12724, Percent = 85.3% • 0 (not white), Frequency = 2192, Percent = 14.7%
Derived gender	Binary	Single aggregated gender categorization derived from applicant gender fields <ul style="list-style-type: none"> • 1 (joint), Frequency = 5799, Percent = 38.9% • 0 (male or female), Frequency = 9117, Percent = 61.1%

Table 1 Descriptions and summary statistics for explanatory variables.

5.1. Description of data set

Our case study is based on a data set collected in accordance with the Home Mortgage Disclosure Act (HMDA)¹. Specifically, we downloaded the data provided by all financial

¹ The HMDA website where the data is hosted is <https://ffiec.cfbp.gov/>

institutions in Pennsylvania who had offered a loan for the purpose of enabling a home purchase. The data set includes all the information for approved applications in 2020, including demographic information about the applicant, their income level, and the interest rate and loan amount the bank offered. After removing unsuitable rows, there were 14,916 approved applications in total, and 11,491 (77%) of approved applications resulted in a loan accepted at a bank offered interest rate. Table 1 summarizes the variables (feature of customers) we use in our case study.

As a preliminary, note we can think of the interest rate the bank offers on the loan as a take-it-or-leave-it price, and the customers choice whether or not to accept the loan as a decision to purchase or not purchase at a given price. Using the price variation in this data, we will estimate a customer valuation model so that we can train and evaluate market segmentation and pricing models. In the first step, we use a probit regression model to predict the probability that a customer will take the offered interest rate. Table 2 shows the coefficient estimates for the probit regression model, and Fig. EC.5 shows the prediction of the probability a customer will take the approved application for a given interest rate. We then transform our probit model into a linear valuation model of the form $V|\mathbf{x} = \mu(\mathbf{x}) + \mathcal{N}(0, \sigma)$, where \mathbf{x} is a feature vector including the interest rate, income level of the customer, and demographic information, and $\mu(\mathbf{x}) = \mathbf{x}^t \beta$. Our transformation from probit regression model to linear valuation model follows Cameron and James (1987), for a short primer describing such transformations, see Section C in the appendix. All subsequent feature-based segmentation and pricing policies will be based on this derived linear valuation model, $V|\mathbf{x} = \mu(\mathbf{x}) + \mathcal{N}(0, \sigma)$.

5.2. Comparison with Segment-then-Price

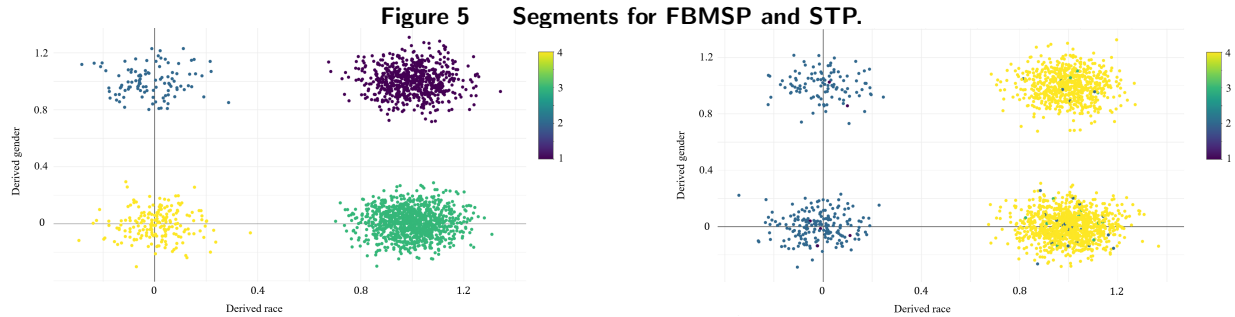
To access the real impact of our optimal FBMSPP policies, we will compare against heuristic segment-then-price (STP) policies. For STP, we will segment customers using the popular

Variable	Estimate	Std. Error	z value	$\Pr(> z)$	Significance
(Intercept)	3.8869	0.1126	34.5314	$< 2.2 \times 10^{-16}$	***
Interest Rate	-0.8704	0.0277	-31.4735	$< 2.2 \times 10^{-16}$	***
Income	-0.0009	0.0002	-4.1142	3.885×10^{-5}	***
Derived race	0.4709	0.0545	8.6374	$< 2.2 \times 10^{-16}$	***
Derived gender	0.1721	0.0441	3.9003	9.606×10^{-5}	***

Table 2 Probit regression coefficients. Significance levels: ***: < 0.001 , **: < 0.01 , *: < 0.05 .

k -medioids algorithm (Reynolds et al. 2006, Schubert and Rousseeuw 2019, 2021) with Gower distance (Gower 1966, 1967). The price optimization is then done over the found segments and can be computed in polynomial time for error with finite support.

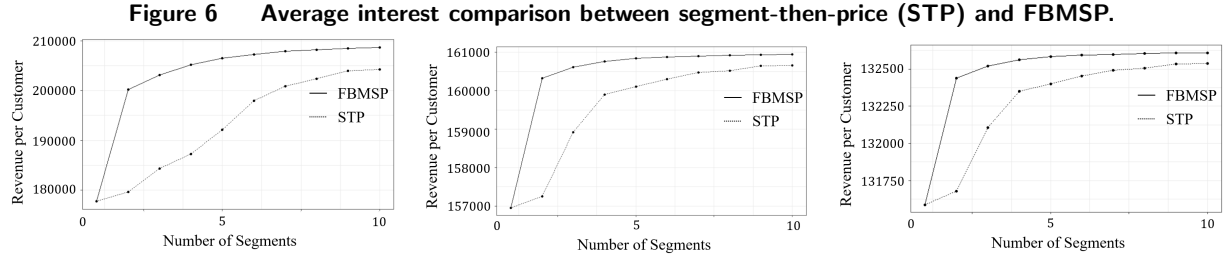
First, we look at the segments generated by segment-then-price and our FBMSF model. Fig. 5 shows that STP will group customers first based on differences in gender and/or race. Gender and race are certainly heterogeneous across our data set, however, these differences are not necessarily the distinctions that are revenue-maximizing to delineate. In comparison, optimal FBMSF will segment customers into different groups based on their valuations, which is only weakly correlated with gender/race in our data. Therefore, compared to STP, FBMSF will not only achieve better revenue but does so in an explainable way by grouping customers with similar predicted valuations, instead of merely similar demographic features which may have negative social or legal ramifications.



Note. Here we plot segments for STP and FBMSF when the number of segments $k = 4$. Since both derived gender and derived race are binary variables, we add some random noise to each point for clarity of presentation (without noise, all points of the same color would be on top of one another in the left panel). In the left panel, the segments are obtained using k -medoids algorithm. In the right panel, we use dynamic programming to do optimal FBMSF.

To compare the difference in revenue garnered by STP and FBMSF, we will examine the difference in total interest a customer will pay on average over the lifetime of the loan (i.e. average revenue per customer), where the interest is calculated using the standard fixed monthly payment formula (Capinski and Zastawniak (2003)). In Fig. 6, we plot the expected revenue the firm can get from one customer on average, against the number of segments. We first note that the revenue per customer is increasing for both FBMSF and STP model. However for FBMSF, the revenue the firm can get from each customer on average, is concave in the number of segments, while the STP model is not. In our case study, STP often gets stuck at small choices of k and requires a 3+ of segments before it

can achieve strong revenue, whereas FBMSPP is guaranteed to get the most revenue out of a small number of segments. We see the differences then between FBMSPP and STP are most pronounced when only a small number of segments are used, which is precisely the case of interest in industry. We further note that for both models while smaller error in the prediction model will yield higher revenue, the gap between the two strategies is also more pronounced when the error is small, suggesting that for sophisticated firms with high quality feature data the benefits of FBMSPP are even greater.



Note. Here we plot the average interest per customer for STP and FBMSPP for different levels of prediction error in our valuation model. In the left panel, the standard deviation of prediction error is $\sigma = 0$, in the middle panel, the standard deviation of prediction error is $\sigma = 0.5$, in the right panel, the standard deviation of prediction error is $\sigma = 1$.

5.3. Finding the Optimal Number of Segments using Regression Model

One additional benefit of the concavity of optimal FBMSPP is it enables us to easily choose the number of segments via the elbow method heuristic. The elbow method is the most commonly used heuristic for finding the optimal number of segments for unsupervised learning. The intuition is that one should choose a number of segments so that adding another segment doesn't give much better modeling of the data (see Bholowalia and Kumar (2014) for more discussion about the elbow method and its applications). To use the elbow method, one prerequisite is that the objective function is monotone in the number of segments. In general, the objective function, revenue per customer, is not necessarily even increasing for the STP. Unlike STP, in Theorem 5 we showed that the revenue is concave in the number of segments k . At some value for k , the revenue increases dramatically, and after that, it reaches a plateau, and increasing the number of segments does not dramatically increase revenue. In Fig. 6, for our FBMSPP model, 2 or 3 is the elbow of the revenue per customer vs. k plot, whereas for STP, the possible elbow is at 8 or 9, a prohibitively large number of segments in practice.

6. Conclusions

Increasingly rich consumer profiles and choice models enable retailers to personalize consumers at finer and finer levels. However, building such tools comes at an investment cost in the form of technology, data scientists, marketing, etc. Motivated by this trade-off, and by a desire to improve on common heuristic approaches, we provide a framework to compute and analyze semi-personalized, feature-based market segmentation and pricing policies under realistic assumptions about how firms predict the valuations of customers.

Specifically, we define and study the feature-based market segmentation and pricing problem, where sellers have trained a regression model to predict customers' valuations using their features. We first prove the computation of optimal feature-based market segmentation and pricing is NP-hard for independent residuals, and provide a $(1 - 1/e)$ approximation algorithm by leveraging the greedy algorithm in the submodular maximization problem. With an additional assumption of log-concavity on the prediction error, we show that the optimal policy is cubic-time solvable with dynamic programming.

Further, we analyze the properties of optimal feature-based market segmentation and pricing. We prove a number of results for the revenue of k segmentation against a fully personalized pricing. We show that the loss of k segmentation is upper bounded the loss without prediction error. We also show the revenue of optimal FBMSPP is concave in the number of segments k , practitioners can find the most suitable k by the simple elbow method, and without loss of much revenue.

Overall, our work seeks to deepen our understanding of semi-personalized pricing strategies, and demonstrate that they are computable, and effective when compared to complicated fully personalized pricing strategies. There are many interesting and important directions to consider for future work, we highlight three of them here. First, this paper assumes the production cost of the good is uniform over all segments. Follow-up work may consider heterogeneous production costs among different segments, and ask whether the optimal FBMSPP in this case still uses interval segments when the residuals are log-concave. Second, it may also be interesting to consider the approximation ratio for interval segmentations facing general error distributions. Example EC.2 demonstrates that interval segmentation is not optimal for some error distributions, but how far it is from the optimal segmentation in the worst case is unknown. Finally, we assume the firm can charge customers in different segments any segment level price. In practice, the firm may only be

able to offer a price menu for customers to choose from it. One may also consider models similar to FBMSP in the price menu setting.

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Appendix A: Omitted Examples

EXAMPLE EC.1 ($p_\epsilon(x)$ CAN BE DISCONTINUOUS). Suppose $\mu(\mathbf{X}) \sim \text{Uniform}[1, 2]$ (or any other continuous distribution on $[1, 2]$) and ϵ is either $-.5$ or $.5$ with probability $\frac{1}{2}$. Then for every $x \leq 1.5$, the optimal price is $p_\epsilon(x) = x + .5$, and $p_\epsilon(x) = x - .5$ otherwise. Thus at 1.5 $p_\epsilon(x)$ is discontinuous, and by Lemma 1 the revenue function $\mathcal{R}_\epsilon(x)$ is non-differentiable. \square

EXAMPLE EC.2 (OPTIMAL SEGMENTATIONS NEED NOT BE INTERVAL). In this example we give $\mu(\mathbf{X})$ and ϵ such that the optimal segmentation and pricing is non-interval. Specifically, for any number $k \geq 2$, assume $\mu(\mathbf{X})$ is uniformly distributed on the set $\{k, k+2, k+4, \dots, 3k\}$, and ϵ is either k or $-k$ with probability $\frac{1}{2}$, respectively. It can be computed without too much difficulty that the unique optimal k -market segmentation and pricing uses segments,

$$\mathcal{X}_1 = \{k, 3k\}, \mathcal{X}_2 = \{k+2\}, \dots, \mathcal{X}_k = \{k+2(k-1)\},$$

with corresponding price for each segment,

$$p(\mathcal{X}_1) = 2k, p(\mathcal{X}_2) = 2k+2, \dots, p(\mathcal{X}_k) = 2k+2(k-1).$$

As the first segment is not interval, the optimal segmentation needs not to be interval for any k . \square

EXAMPLE EC.3 (TIGHTNESS OF THEOREM 4). Suppose the regression model has no error, i.e. $V|\mathbf{x} = \mu(\mathbf{x})$, and let $\mu(\mathbf{X}) \sim \text{Uniform}[0, t]$ for some $t > 0$. Then, $\mathbb{E}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))] = \mathbb{E}[V] = \frac{t}{2}$. To compute \mathcal{R}_{kXP} for some k , note by Theorem 2 the optimal segmentation here is interval and further each segment can be described by a left and right endpoint in the space of predicted valuations. Let $0 < s_0 < \dots < s_k = t$ describe those segments (i.e. $\mathcal{X}_i = \{x | \mu(x) \in [s_{i-1}, s_i]\}$) with corresponding prices p_1, \dots, p_k . It is easy to see since $\epsilon = 0$, the optimal price and segmentations must satisfy $s_{i-1} = p_i$ for $i = 1, \dots, k$ since, if not, increasing the segment interval s_{i-1} up to p_i only increases revenue.

Now, on segment \mathcal{X}_i , the conditional distribution of V is still uniform, so the contribution of that segment to $\mathbb{E}[V]$ is $\frac{s_i + s_{i-1}}{2} \cdot \frac{s_i - s_{i-1}}{t}$ for all i , since only $\frac{s_i - s_{i-1}}{t}$ fraction of the market is in this interval. By contrast, for $i = 1, \dots, k$, the k -market segmentation strategy on segment i earns revenue $p(\mathcal{X}_i) \Pr(\mu(\mathbf{X}) \geq p(\mathcal{X}_i) | \mathcal{X} \in [s_{i-1}, s_i]) \Pr(\mathbf{X} \in [s_{i-1}, s_i]) = s_{i-1} \frac{s_i - s_{i-1}}{t}$ since $p_i = s_{i-1}$ and thus, all customers in the segment buy. The difference in revenue is then

$$\begin{aligned} \mathbb{E}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))] - \mathcal{R}_{kXP} &= \frac{s_1^2}{2t} + \sum_{i=1}^k \frac{s_i + s_{i-1}}{2} \cdot \frac{s_i - s_{i-1}}{t} - s_{i-1} \frac{s_i - s_{i-1}}{t} \\ &= \frac{s_1^2}{2t} + \frac{1}{2t} \sum_{i=1}^k (s_{i+1} - s_i)^2 = \frac{1}{2t} \sum_{i=0}^k (s_{i+1} - s_i)^2. \end{aligned}$$

By inspection, for a fixed s_0 , the segmentation which minimizes this difference is equispaced, i.e., $s_i = s_{i-1} + \frac{t}{k+1}$ for $i = 1, \dots, k$. Plugging in gives $\mathbb{E}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))] - \mathcal{R}_{kXP} = \frac{t}{2(k+1)} = \frac{\mathbb{E}[V]}{k+1}$. \square

Appendix B: Omitted Proofs

B.1. Omitted Proofs from Section 2

Proof of Lemma 1. (a) Fix some ϵ and positive real numbers x_1, x_2 such that $x_1 < x_2$ and recall $\theta_\epsilon(x) := p_\epsilon(x) - x$ is the difference between the price and x . Further recall $p_\epsilon(x_1), p_\epsilon(x_2)$ are prices that maximize $p\bar{F}_\epsilon(p - x_1)$ and $p\bar{F}_\epsilon(p - x_2)$ respectively. Thus, by optimality we have the following two inequalities

$$(x_1 + \theta_\epsilon(x_1)) \bar{F}_\epsilon(\theta_\epsilon(x_1)) \geq (x_1 + \theta_\epsilon(x_2)) \bar{F}_\epsilon(\theta_\epsilon(x_2)), \quad (\text{EC.1})$$

$$(x_2 + \theta_\epsilon(x_2)) \bar{F}_\epsilon(\theta_\epsilon(x_2)) \geq (x_2 + \theta_\epsilon(x_1)) \bar{F}_\epsilon(\theta_\epsilon(x_1)). \quad (\text{EC.2})$$

Rearranging the two inequalities yields,

$$\frac{x_1 + \theta_\epsilon(x_1)}{x_1 + \theta_\epsilon(x_2)} \geq \frac{\bar{F}_\epsilon(\theta_\epsilon(x_2))}{\bar{F}_\epsilon(\theta_\epsilon(x_1))} \geq \frac{x_2 + \theta_\epsilon(x_1)}{x_2 + \theta_\epsilon(x_2)}.$$

Consequently,

$$(x_1 + \theta_\epsilon(x_1))(x_2 + \theta_\epsilon(x_2)) \geq (x_1 + \theta_\epsilon(x_2))(x_2 + \theta_\epsilon(x_1)),$$

Simplifying the expression, we get

$$(x_2 - x_1)\theta_\epsilon(x_1) \geq (x_2 - x_1)\theta_\epsilon(x_2).$$

Finally, noting $x_2 - x_1 > 0$, the inequality is equivalent to $\theta_\epsilon(x_1) \geq \theta_\epsilon(x_2)$ and thus the margin monotone decreasing.

(b) Again, Fix some ϵ and positive real numbers x_1, x_2 such that $x_1 \leq x_2$. Then $x_1 + \epsilon \leq_{\text{st}} x_2 + \epsilon$ in the sense of first order stochastic dominance, and it is well known that stochastic dominance of the valuations implies $\mathcal{R}_\epsilon(x_1) \leq \mathcal{R}_\epsilon(x_2)$ (see for instance Hart and Reny (2015) for an extended discussion). Combining this observation with Eqs. (EC.1) and (EC.2) above yields,

$$\begin{aligned} \mathcal{R}_\epsilon(x_2) - \mathcal{R}_\epsilon(x_1) &\geq (x_2 + \theta_\epsilon(x_1)) \bar{F}_\epsilon(\theta_\epsilon(x_1)) - (x_1 + \theta_\epsilon(x_1)) \bar{F}_\epsilon(\theta_\epsilon(x_1)) = (x_2 - x_1) \bar{F}_\epsilon(\theta_\epsilon(x_1)), \\ \mathcal{R}_\epsilon(x_2) - \mathcal{R}_\epsilon(x_1) &\leq (x_2 + \theta_\epsilon(x_2)) \bar{F}_\epsilon(\theta_\epsilon(x_2)) - (x_1 + \theta_\epsilon(x_2)) \bar{F}_\epsilon(\theta_\epsilon(x_2)) = (x_2 - x_1) \bar{F}_\epsilon(\theta_\epsilon(x_2)). \end{aligned}$$

Dividing through both sides by $x_2 - x_1$ gives,

$$\bar{F}_\epsilon(\theta_\epsilon(x_1)) \leq \frac{\mathcal{R}_\epsilon(x_2) - \mathcal{R}_\epsilon(x_1)}{x_2 - x_1} \leq \bar{F}_\epsilon(\theta_\epsilon(x_2)). \quad (\text{EC.3})$$

When p_ϵ is continuous then θ_ϵ is also continuous, and taking $x_1 \rightarrow x_2$ squeezes the derivative to be $\bar{F}_\epsilon(p_\epsilon(x) - x)$ as desired.

(c) $\mathcal{R}_\epsilon(x)$ was noted to be increasing in the proof of (b). Now to prove continuity, fix ϵ and positive real numbers x_1, x_2 such that $x_1 < x_2$. Then,

$$\mathcal{R}_\epsilon(x_1) \leq \mathcal{R}_\epsilon(x_2) = (x_2 + \theta_\epsilon(x_2) + (x_2 - x_1) - (x_2 - x_1)) \bar{F}_\epsilon(\theta_\epsilon(x_2)) \leq \mathcal{R}_\epsilon(x_1) + (x_2 - x_1),$$

where the last inequality follows from distributing and applying Eq. (EC.2), and the fact that $\bar{F}_\epsilon(\cdot) \leq 1$. Taking $x_1 \rightarrow x_2$ gives us the continuity of $\mathcal{R}_\epsilon(x)$.

For convexity, again fix positive real numbers x_1, x_2 and also $\lambda \in (0, 1)$. Then,

$$\begin{aligned}\mathcal{R}_\epsilon(x_1) &= p_\epsilon(x_1) \bar{F}_\epsilon(p_\epsilon(x_1) - x_1) \\ &\geq (p_\epsilon(\lambda x_1 + (1 - \lambda)x_2) + (1 - \lambda)(x_1 - x_2)) \bar{F}_\epsilon(p_\epsilon(\lambda x_1 + (1 - \lambda)x_2) - (\lambda x_1 + (1 - \lambda)x_2)),\end{aligned}$$

where the inequality follows from noting that $p_\epsilon(x_1)$ is revenue optimal for $x_1 + \epsilon$ and any other price can earn no more. Similarly,

$$\begin{aligned}\mathcal{R}_\epsilon(x_2) &= p_\epsilon(x_2) \bar{F}_\epsilon(p_\epsilon(x_2) - x_2) \\ &\geq (p_\epsilon(\lambda x_1 + (1 - \lambda)x_2) - \lambda(x_1 - x_2)) \bar{F}_\epsilon(p_\epsilon(\lambda x_1 + (1 - \lambda)x_2) - (\lambda x_1 + (1 - \lambda)x_2)).\end{aligned}$$

Combine the two inequalities above, we get

$$\lambda \mathcal{R}_\epsilon(x_1) + (1 - \lambda) \mathcal{R}_\epsilon(x_2) \geq \mathcal{R}_\epsilon(\lambda x_1 + (1 - \lambda)x_2),$$

which means $\mathcal{R}_\epsilon(x)$ is convex in x . \square

B.2. Omitted Proofs from Section 3

Proof of Theorem 1. We will prove hardness by showing the Hitting set problem can be reduced to an instance of k feature-based market segmentation and pricing (kXP). Let \mathcal{X} be the ground set of elements of size, $|\mathcal{X}| = m$, and let $\{H_i\}_{i=1}^n$ be a collection of subsets of \mathcal{X} . Consider the decision version of the hitting set problem, which asks whether there exists a subset of $\mathcal{X}^* \subset \mathcal{X}$, $|\mathcal{X}^*| \leq k$, such that \mathcal{X}^* has non-empty intersection with each H_i . To build a corresponding k -market segmentation and pricing problem, suppose we have n customers such that each customer's valuation is $\mu_i + \epsilon^2$, where $i = 1, 2, \dots, n$. Let $p_j = n + \frac{j-1}{m}$, for $j = 1, 2, \dots, m$, and let $\mu_1 = p_1$, and $\mu_i = \mu_{i-1} + \frac{p_1}{2(i-1)} + \frac{p_m}{2i}$, for $i = 2, \dots, n$. We will now construct an $\epsilon \sim F_\epsilon$ such that, for each μ_i , $\mathcal{R}_\epsilon(\mu_i)$ is maximized at price p_j if and only if in the hitting set problem the subset H_i contains element x_j .

Our construction of ϵ is supported on numbers of the form $p_i - \mu_j$. Before constructing ϵ , note that p_j is strictly increasing in j , and that $p_j - \mu_i < p_{j'} - \mu_i$ as long as $j' > j$. Further note $p_m - \mu_{i+1} < p_1 - \mu_i$ since by the definition of μ_i and μ_{i+1} , $\mu_{i+1} - \mu_i > \frac{p_1}{2n} + \frac{p_m}{2n} > 1$ and since $p_m - p_1 = \frac{m-1}{m} < 1$. Let $t_{1,1} = p_1 - \mu_n, \dots, t_{j,i} = p_j - \mu_{n+1-i}, \dots, t_{m,n} = p_m - \mu_1$, and let $t_{0,n} = -\mu_n$, and $t_{m,n+1} = p_m$. Thus, we have

$$t_{0,n} < t_{1,1} < t_{2,1} < \dots < t_{m,i} < t_{1,i+1} < t_{2,i+1} < \dots < t_{m-1,n} < t_{m,n} < t_{m,n+1}. \quad (\text{EC.4})$$

Now we are ready to define the complementary cumulative distribution function (cCDF) of ϵ . We will let ϵ be such $\bar{F}_\epsilon(t_{0,n}) = 1$, $\bar{F}_\epsilon(t_{m,n+1}) = 0$, and working backwards recursively from $\bar{F}_\epsilon(t_{m,n+1})$ as follows:

$$\bar{F}_\epsilon(t_{j,i}) = \begin{cases} \frac{i}{p_j}, & \text{if } x_j \in H_i \\ \bar{F}_\epsilon(t_{j+1,i}), & \text{if } x_j \notin H_i \text{ and } j < m \\ \bar{F}_\epsilon(t_{1,i+1}) & \text{if } x_j \notin H_i \text{ and } j = m, i < n \\ 0 & \text{otherwise.} \end{cases}$$

Note this construction is well defined and is quadratically supported, an example what \bar{F}_ϵ looks like is provided in Fig. EC.6. Further, for any value t such that $t_{j,i} \leq t < t_{j+1,i}$, $\bar{F}(t) = \bar{F}(t_{j,i})$. Now we need to

² Equivalently, $\mu(\mathbf{X})$ is uniformly supported on these valuations.

check that \bar{F}_ϵ is non-increasing and thus a properly defined cCDF, and also that $p_j \bar{F}_\epsilon(p_j - \mu_i)$ is revenue-maximizing only when j, i are such that $x_j \in H_i$.

To the first point, since $p_j \geq n$ for all $j = 1, 2, \dots, m$, and $\{p_j\}$ is increasing, therefore, $\frac{i}{p_j} < \frac{i}{p_{j+1}}$. Then, to show \bar{F}_ϵ is non-increasing, we only need to show $\frac{i}{p_m} < \frac{i-1}{p_1}$. Note that

$$\begin{aligned} \frac{i}{p_m} - \frac{i-1}{p_1} &= \frac{ip_1 - (i-1)p_m}{p_1 p_m} \\ &= \frac{i(p_1 - p_m) + p_m}{p_1 p_m} > 0, \end{aligned}$$

where the inequality follows from the fact that $p_m - p_1 = \frac{m-1}{m} < 1$ for $1 \leq i \leq n$, and $p_m > n$. Thus \bar{F}_ϵ is non-increasing, i.e., \bar{F}_ϵ is a proper cumulative distribution function.

Next, we show that $p \bar{F}_\epsilon(p - \mu_i) = i$ iff $p = p_j$ and $x_j \in H_i$, and for all other prices the revenue $p \bar{F}_\epsilon(p - \mu_i)$ is strictly less than i . By the definition of $\bar{F}_\epsilon(p_j - \mu_i)$, if $x_j \in H_i$, $\bar{F}_\epsilon(p_j - \mu_i) = \frac{i}{p_j}$, consequently, $p_j \bar{F}_\epsilon(p_j - \mu_i) = i$. So now suppose price p satisfies $\bar{F}_\epsilon(p - \mu_i) = \bar{F}_\epsilon(p_{j'} - \mu_{i'}) = \frac{i'}{p_{j'}}$. To simplify the discussion, we take the largest price p such that $\bar{F}_\epsilon(p - \mu_i) = \bar{F}_\epsilon(p_{j'} - \mu_{i'})$, i.e., $p - \mu_i = p_{j'} - \mu_{i'}$, and by rearranging $p = p_{j'} - \mu_{i'} + \mu_i$. All other prices less than $p_{j'} - \mu_{i'} + \mu_i$ and which satisfies $\bar{F}_\epsilon(p - \mu_i) = \bar{F}_\epsilon(p_{j'} - \mu_{i'}) = \frac{i'}{p_{j'}}$ will give us less revenue. Now we want to show that $p \bar{F}_\epsilon(p - \mu_i) < i$, i.e.,

$$(p_{j'} - \mu_{i'} + \mu_i) \frac{i'}{p_{j'}} < i.$$

If $i' < i$, the inequality is the same as

$$\mu_i - \mu_{i'} \leq \frac{i - i'}{i'} p_{j'}.$$

By the definition of p_j and μ_i ,

$$\mu_{i+1} - \mu_i = \frac{p_1}{2i} + \frac{p_m}{2(i+1)} < \frac{p_1}{i},$$

where the inequality comes from $\frac{p_m}{p_1} < \frac{n+1}{n}$. Therefore,

$$\mu_i - \mu_{i'} \leq \sum_{j=i'}^i \frac{p_1}{j} < \frac{i - i'}{i} p_1 < \frac{i - i'}{i} p_{j'}.$$

Similarly, if $i' > i$, $p \bar{F}_\epsilon(p - \mu_i) < i$ is equivalent to

$$\mu_{i'} - \mu_i > \frac{i' - i}{i'} p_{j'}.$$

Now, by the definition of p_j and μ_i ,

$$\mu_{i+1} - \mu_i = \frac{p_1}{2i} + \frac{p_m}{2(i+1)} > \frac{p_m}{i+1},$$

where the inequality comes from $\frac{p_m}{p_1} < \frac{n+1}{n}$. Therefore,

$$\mu_{i'} - \mu_i \geq \sum_{j=i'}^i \frac{p_m}{j+1} > \frac{i - i'}{i} p_m > \frac{i - i'}{i} p_{j'},$$

as desired.

Finally, to determine whether there exists a subset of $X^* \subset X$, $|X^*| \leq k$, such that X^* has non-empty intersection with each H_i , it is equivalent to determine whether there is a k market segmentation and pricing that yields the maximum revenue $1 + 2 + 3 + \dots + n = \frac{n(n+1)}{2}$. Since the hitting set problem NP-hard, thus FBMSPP is also NP-hard. \square

Proof of Lemma 2. (a) First we will show $p_\epsilon(x)$ is increasing. Since f_ϵ is log-concave, \bar{F}_ϵ is also log-concave (see Bagnoli and Bergstrom (2005) for an extensive overview of the transformation that preserve log-concavity). Further, $\frac{d}{dx} \log(\bar{F}_\epsilon(x)) = \frac{-f_\epsilon(x)}{\bar{F}_\epsilon(x)}$ which by concavity implies the inverse hazard rate, $\frac{\bar{F}_\epsilon(x)}{f_\epsilon(x)}$, is decreasing in x . $p_\epsilon(\cdot)$ is thus unique, satisfies first order conditions for revenue optimality, $\frac{d}{dp} p \bar{F}_\epsilon(p - x)|_{p=p(x)} = 0$, and can be written as $p_\epsilon(x) = \frac{\bar{F}_\epsilon(p_\epsilon(x) - x)}{f_\epsilon(p_\epsilon(x) - x)}$, but by Lemma 1(a) $p_\epsilon(x) - x$ is decreasing thus $p_\epsilon(x)$ it must be an increasing function of x .

(b) Again note if $f(x)$ is log-concave, \bar{F}_ϵ is a log-concave function, thus it has Pólya frequency of order 2 (PF2), which is equivalent to that statement that, for any real numbers x_1, x_2 , and y_1, y_2 , such that $x_1 < x_2$ and $y_1 < y_2$, then $\frac{\bar{F}_\epsilon(x_1 - y_2)}{\bar{F}_\epsilon(x_1 - y_1)} \leq \frac{\bar{F}_\epsilon(x_2 - y_2)}{\bar{F}_\epsilon(x_2 - y_1)}$ (see Saumard and Wellner (2014), Section 11).

Now, let $\{\mathcal{X}_i\}_1^k, \{p_\epsilon(\mathcal{X}_i)\}_1^k$ be the optimal segmentation and pricing and suppose WLOG that the prices are distinct. Further suppose the optimal segmentation was not interval, then there exists $\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3$ such that $\mu(\mathbf{x}_1) < \mu(\mathbf{x}_2) < \mu(\mathbf{x}_3)$, but with $\mathbf{x}_1, \mathbf{x}_3 \in \mathcal{X}_i$, and $\mathbf{x}_2 \in \mathcal{X}_j$ for some $i \neq j$. Suppose $p_\epsilon(\mathcal{X}_i) < p_\epsilon(\mathcal{X}_j)$ (the opposite case when $p_\epsilon(\mathcal{X}_i) > p_\epsilon(\mathcal{X}_j)$ follows by an identical argument, swapping \mathbf{x}_3 with \mathbf{x}_1) and note by optimality of the segmentation,

$$\begin{aligned} p_\epsilon(\mathcal{X}_i) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_3)) &> p_\epsilon(\mathcal{X}_j) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_3)), \\ p_\epsilon(\mathcal{X}_j) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_2)) &> p_\epsilon(\mathcal{X}_i) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_2)). \end{aligned}$$

Combining these two inequalities gives

$$\frac{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_3))}{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_3))} > \frac{p_\epsilon(\mathcal{X}_j)}{p_\epsilon(\mathcal{X}_i)} > \frac{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_2))}{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_2))}.$$

Which can be further rearranged to $\frac{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_3))}{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - \mu(\mathbf{x}_2))} > \frac{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_3))}{\bar{F}_\epsilon(p_\epsilon(\mathcal{X}_j) - \mu(\mathbf{x}_2))}$ which contradicts the PF2 property. Thus the optimal segmentation must be interval and $\mathcal{X}_i = \{\mathbf{x} | \mu(\mathbf{x}) \in [s_{i-1}, s_i]\}$ for some real numbers $s_i < s_{i+1}$.

(c) To show $p_\epsilon(\mathcal{X}_i) = p_\epsilon(\mathbf{x})$ for some $\mathbf{x} \in \mathcal{X}_i$, let $\mathbf{x}' = \arg \min_{\mathbf{x} \in \mathcal{X}_i} \mu(\mathbf{x})$ and recall $p_\epsilon(\mathcal{X}_i) = \arg \max \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} p \bar{F}_\epsilon(p - s) f(\mu^{-1}(s)) ds$. Now suppose $p_\epsilon(\mathcal{X}_i) < \mu(\mathbf{x}')$. By log-concavity, each function $\mathcal{R}_\epsilon(\mu(\mathbf{x}), p) := p \bar{F}_\epsilon(p - \mu(\mathbf{x}))$ is unimodal, and thus increasing in p for $p \leq p_\epsilon(\mu(\mathbf{x}))$,

$$p_\epsilon(\mathcal{X}_i) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - s) \leq p_\epsilon(\mu(\mathbf{x}')) \bar{F}_\epsilon(p_\epsilon(\mu(\mathbf{x}')) - s),$$

for any $s \in [s_i, s_{i+1}]$, which implies

$$\int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} p_\epsilon(\mathcal{X}_i) \bar{F}_\epsilon(p_\epsilon(\mathcal{X}_i) - s) f(\mu^{-1}(s)) ds \leq \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} p_\epsilon(\mu(\mathbf{x}')) \bar{F}_\epsilon(p_\epsilon(\mu(\mathbf{x}')) - s) f(\mu^{-1}(s)) ds,$$

thus $p_\epsilon(\mathcal{X}_i) \geq p_\epsilon(\mathbf{x}')$. A symmetric argument similarly shows $p_\epsilon(\mathcal{X}_i) \leq \arg \max_{\mathbf{x} \in \mathcal{X}_i} p_\epsilon(\mathbf{x})$. \square

Proof of Theorem 2 Suppose the firms prediction model $\mu(\mathbf{X})$ is supported on n values $\{\mu_i\}_{i=1}^n$, occurring with probabilities $\{q_i\}_{i=1}^n$, where $\mu_1 \leq \mu_2 \leq \dots \leq \mu_n$, and let $|\text{supp}(\epsilon)| = m_\epsilon$. By Lemma 2, the optimal segmentation can be index by the sequence $\{s_i\}_{i=0}^k$ is contained in the support of $\mu(\mathbf{X})$. Let the optimal price for segment $[s_{i-1}, s_i]$ be $p_\epsilon([s_{i-1}, s_i])$ i.e.

$$p_\epsilon([s_{i-1}, s_i]) = \arg \max p_i \Pr(\mu(\mathbf{x}) + \epsilon \geq p_i | \mu(\mathbf{x}) \in [s_{i-1}, s_i]) \Pr(\mu(\mathbf{x}) \in [s_{i-1}, s_i]).$$

We wish to find $\{s_i\}_{i=0}^k \subset \{\mu_i\}_{i=1}^n$ that maximizes

$$\sum_{i=1}^k p_\epsilon([s_{i-1}, s_i]) \Pr(\mu(\mathbf{x}) + \epsilon \geq p_\epsilon([s_{i-1}, s_i]) | \mu(\mathbf{x}) \in [s_{i-1}, s_i]) \Pr(\mu(\mathbf{x}) \in [s_{i-1}, s_i]).$$

where the optimal price $p_\epsilon([s_{i-1}, s_i])$ for each segmentation can be calculated in time $O(nm_\epsilon)$ by simple search over the support. We now give a dynamic programming solution that uses time $O(kn^3m_\epsilon)$ and to populate a table of size kn . Define $D[n', k']$ as the optimal k' -market segmentation that considers only the n' lowest predicted valuations $\{(\mu_i, q_i)\}_{i=1}^{n'}$, our goal is to compute $D[n, k]$ which is the revenue of the optimal FBMSP (the optimal policy can further be reconstructed by standard backward search). Our algorithm depends on the following observation: consider the optimal k -market segmentation and suppose $[s_{k-1}, s_k] = [\mu_{i_k}, \mu_n]$ defines the k^{th} segment. If one considers the market without the customers in the k^{th} segment, the remaining $k-1$ segments must be an optimal $(k-1)$ -market segmentation on $\{(\mu_i, q_i)\}_{i=1}^{i_k-1}$. Formally, we express this observation as the following recursion,

$$D[n', k'] = \max_{l \in [n'-1]} D[l, k'-1] + p_\epsilon([s_{i-1}, s_i]) \sum_{i=l+1}^{n'} \Pr(\mu(\mathbf{x}) + \epsilon \geq p_\epsilon([s_{i-1}, s_i]) | \mu(\mathbf{x}) \in [s_{i-1}, s_i]) q_i, \quad (\text{EC.5})$$

which states that the optimal k' -market segmentation on the lowest n' valuations, is equal to some optimal $(k'-1)$ -segmentation on a smaller market, plus the value of the k'^{th} segment. Using Eq. (EC.5) we may populate a table of size kn , starting at $D[0, 0] = 0$, and computing column-wise. Each computation of $D[n', k']$ requires $O(n^2m_\epsilon)$ operations, thus the table may be populated in $O(kn^3m_\epsilon)$ time. \square

B.3. Omitted Proofs from Section 4

Proof of Theorem 3 Let $\{s_i\}_{i=0}^k \in \mathbb{R}^{k+1}$ denote an optimal interval k -market segmentation for $\mathcal{R}_{kXP}^{\mu(\mathbf{x})}$. Consider the sub-optimal feature-based market segmentation which uses segments $\mathcal{X}_i = \{\mathbf{x} | \mu(\mathbf{x}) \in [s_{i-1}, s_i]\}$ and prices p_i . Note for all $\mathbf{x} \in \mathcal{X}_i$, $p_i \Pr(\mu(\mathbf{x}) + \epsilon \geq p_i) \geq \mathcal{R}_\epsilon(s_{i-1})$ and thus summing over all segments

$$\mathcal{R}_{kXP}^{\mu(\mathbf{x})+\epsilon} \geq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} p_i \Pr(\mu(\mathbf{x}) + \epsilon \geq p_i) f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x} \geq \sum_{i=1}^k \mathcal{R}_\epsilon(s_{i-1}) \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x} \quad (\text{EC.6})$$

Now,

$$\begin{aligned} \mathbb{E}_{\mathbf{x} \sim F_{\mathbf{x}}}[\mathcal{R}_\epsilon(\mu(\mathbf{x}))] - \mathcal{R}_{kXP}^{\mu(\mathbf{x})+\epsilon} &\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} (\mathcal{R}_\epsilon(\mu(\mathbf{x})) - \mathcal{R}_\epsilon(s_{i-1})) f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x} && \text{Eq. (EC.6)} \\ &\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} (\mu(\mathbf{x}) - s_{i-1}) \bar{F}_\epsilon(\theta_\epsilon(\mu(\mathbf{x}))) f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x} && \text{Lemma 1(b)} \\ &\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} (\mu(\mathbf{x}) - s_{i-1}) f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x} \\ &= \mathbb{E}_{\mathbf{x} \sim F_{\mathbf{x}}}[\mu(\mathcal{X}_i)] - \mathcal{R}_{kXP}^{\mu(\mathcal{X}_i)} \end{aligned}$$

as desired. \square

Proof of Theorem 4. Let $L = \inf_{\mathbf{x}} \mu(\mathbf{x})$ and $U = \sup_{\mathbf{x}} \mu(\mathbf{x})$, and recalling the proof of Theorem 3 we have,

$$\mathbb{E}_{V \sim F}[\mathcal{R}_\epsilon(V)] - \mathcal{R}_{kXP}^V \leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i]} (\mathcal{R}_\epsilon(\mu(\mathbf{x})) - \mathcal{R}_\epsilon(s_{i-1})) \bar{F}_\epsilon(\theta_\epsilon(s_{i-1})) f_{\mathbf{x}}(\mathbf{x}) d\mathbf{x}.$$

Now instead we choose the sub-optimal interval segmentation such that it subdivides the quantile space of $F_{\mathbf{X}}$ into k equal regions i.e., $\{s_i\}_{i=0}^k$ such that $\bar{F}(s_{i-1}) - \bar{F}(s_i) = \frac{1}{k}$ for $i = 1, \dots, k$. Then

$$\begin{aligned}
\mathbb{E}_{\mu(\mathbf{X}) \sim F}[\mathcal{R}_\epsilon(\mu(\mathbf{X}))] - \mathcal{R}_{kXP}^{\mu(\mathbf{X})+\epsilon} &\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i)} (\mathcal{R}_\epsilon(\mu(\mathbf{x})) - \mathcal{R}_\epsilon(s_{i-1})) f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} \\
&\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i)} (\mu(\mathbf{x}) - s_{i-1}) \bar{F}_\epsilon(\theta_\epsilon(\mu(\mathbf{x}))) f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} \quad \text{Lemma 1(b)} \\
&\leq \sum_{i=1}^k \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i)} (\mu(\mathbf{x}) - s_{i-1}) f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} \\
&= \sum_{i=0}^k (s_i - s_{i-1}) \int_{\mu(\mathbf{x}) \in [s_{i-1}, s_i)} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} \\
&= \sum_{i=1}^k (s_i - s_{i-1}) (\bar{F}(s_{i-1}) - \bar{F}(s_i)) \\
&= \frac{\sum_{i=1}^k (s_i - s_{i-1})}{k} \\
&= \frac{U - L}{k},
\end{aligned}$$

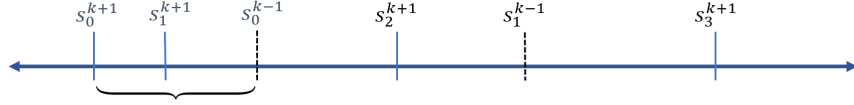
where the third inequality follows $\bar{F}_\epsilon(\theta_\epsilon(\cdot)) \leq 1$, the third equality comes from the choice of $\{s_i\}_{i=0}^k$. Finally, summing $s_i - s_{i-1}$ we get the final equality as desired. \square

Proof of Theorem 5. Fix some $k \geq 2$, we will prove the rearranged inequality $\mathcal{R}_{(k-1)XP} + \mathcal{R}_{(k+1)XP} \leq 2\mathcal{R}_{kXP}$ by explicitly constructing feasible (but not necessarily optimal) size k segmentations. Note, since ϵ is log-concave, by Lemma 2 the optimal segmentation for any k is interval, and can be described by the sequence of numbers $\{s_i^k\}_{i=0}^k$ such that $\mathcal{X}_i^k = \{\mathbf{x} | \mu(\mathbf{x}) \in [s_{i-1}^k, s_i^k]\}$. Further, let $\mathcal{S}_{k-1} := \{s_i^{k-1}\}_{i=0}^{k-1}$, $\mathcal{S}_k := \{s_i^k\}_{i=0}^k$ and $\mathcal{S}_{k+1} := \{s_i^{k+1}\}_{i=0}^{k+1}$ be the optimal segmentations of size $k-1$, k , and $k+1$, respectively, as described by segmentation endpoints. Note by definition $s_{k-1}^{k-1} = s_k^k = s_{k+1}^{k+1} = \sup_{\mathbf{x}} \mu(\mathbf{x})$. Our proof will proceed in two cases.

Case 1: $s_1^{k+1} \leq s_0^{k-1}$. In this case, the first segment of the optimal $(k+1)$ segmentation is before the first segment of the $(k-1)$ segmentation, see Fig. EC.1 for an illustration. Consider two feasible k segmentation $\mathcal{S}'_k = s_0^{k+1} \cup \mathcal{S}_{k-1}$ and $\mathcal{S}''_k = \mathcal{S}_{k+1} \setminus s_0^{k+1}$, see Fig. EC.2 for another illustration. Now note the combined revenue from \mathcal{S}'_k and \mathcal{S}''_k fully covers the revenue from \mathcal{S}_{k-1} (as a subset of \mathcal{S}'_k) and \mathcal{S}_{k+1} except the revenue from the first segment (all other segments of the optimal $(k+1)$ segmentation are covered by \mathcal{S}''_k). Now, in the constructed segmentation \mathcal{S}'_k , the unaccounted for first segment has end points $[s_0^{k+1}, s_0^{k-1})$, which by assumption contains the first segment of the $(k+1)$ segmentation $[s_0^{k+1}, s_1^{k+1})$. Note if one segment subsumes another, it provides more revenue, i.e., if $\mathcal{X}_i \subset \mathcal{X}_j$ then

$$\max_p p \Pr(\mu(\mathbf{x}) + \epsilon \geq p | \mathbf{x} \in \mathcal{X}_i) \Pr(\mathcal{X}_i) \leq \max_p p \Pr(\mu(\mathbf{x}) + \epsilon \geq p | \mathbf{x} \in \mathcal{X}_j) \Pr(\mathcal{X}_j). \quad (\text{EC.7})$$

and thus Eq. (EC.7) implies $\mathcal{R}_{(k-1)XP} + \mathcal{R}_{(k+1)XP} \leq \mathcal{R}_{kXP}(\mathcal{S}'_k) + \mathcal{R}_{kXP}(\mathcal{S}''_k) \leq 2\mathcal{R}_{kXP}$.

Figure EC.1 Case 1 of Theorem 5.

Note. Depicted are the first segment end points of the an interval segmentation of size $k+1$ (solid line) and $k-1$ (dashed line). The first segment $[s_0^{k+1}, s_1^{k+1})$ in the $(k+1)$ segmentation is before the first segment $[s_0^{k-1}, s_1^{k-1})$ of the $(k-1)$ segmentation.

Figure EC.2 New Feasible k -Segmentations for Case 1 of Theorem 5.

Note. Depicted are the constructed size k interval segmentations. In the left panel \mathcal{S}'_k is shown which is equal to the optimal $(k-1)$ segmentation plus a new first segment $[s_0^{k+1}, s_0^{k-1})$. In the right panel \mathcal{S}''_k is shown which is equal to the optimal $(k+1)$ segmentation with the first segment $[s_0^{k+1}, s_1^{k+1})$ removed.

Case 2: There exists $i \in [k-1]$ such that $s_{i-1}^{k-1} \leq s_i^{k+1} \leq s_{i+1}^{k+1} \leq s_i^{k-1}$. In this case, there is an i such that segment $i+1$ of the optimal $(k+1)$ segmentation that is subsumed by segment i of the optimal $(k-1)$ segmentation. As in Case 1, we will construct feasible k segmentations assuming the condition of Case 2 holds. Before constructing the feasible segmentations, we will need two simple facts, both of which follow from Lemma 2.

Fact 1: If $\mathcal{X}_1 = [s_1, s_2)$, $\mathcal{X}_2 = [s_1, s_2 + \Delta)$, where $\Delta \geq 0$, then $p_\epsilon(\mathcal{X}_2) \geq p_\epsilon(\mathcal{X}_1)$, (EC.8)

Fact 2: If $\mathcal{X}_1 = [s_1, s_2)$, $\mathcal{X}_2 = [s_1 - \Delta, s_2)$, where $\Delta \geq 0$, then $p_\epsilon(\mathcal{X}_2) \leq p_\epsilon(\mathcal{X}_1)$. (EC.9)

Fix the i such that $s_{i-1}^{k-1} \leq s_i^{k+1} \leq s_{i+1}^{k+1} \leq s_i^{k-1}$. Such an arrangement of segmentation points is shown in Fig. EC.3. Now define the feasible k -segmentations as

$$\mathcal{S}'_k = \{s_j^{k-1}\}_{j=1}^{i-1} \cup \{s_j^{k+1}\}_{j=i+1}^{k+1},$$

and

$$\mathcal{S}''_k = \{s_j^{k+1}\}_{j=1}^i \cup \{s_j^{k-1}\}_{j=i}^{k-1}.$$

Each new arrangement of segmentation points consists of splicing the beginning of the $(k-1)$ segmentation with the end of the $(k+1)$ segmentation, or vice versa, with the middle segment added or removed. An example of such a segmentation construction is shown in Fig. EC.4. Note compared to the $(k-1)$ and $(k+1)$ segmentations, the new segment in \mathcal{S}'_k is $[s_{i-1}^{k-1}, s_{i+1}^{k+1})$, and the new segment in \mathcal{S}''_k is $[s_i^{k+1}, s_i^{k-1})$, and all the other segments are the same as segments in the optimal $k-1$ or $(k+1)$ segmentation. The only segments in the optimal $(k-1)$ or $(k+1)$ segmentation unaccounted for (i.e. not contained in \mathcal{S}'_k or \mathcal{S}''_k) are $[s_{i-1}^{k-1}, s_i^{k-1})$ and $[s_i^{k+1}, s_{i+1}^{k+1})$. Now let $p_1 := p_\epsilon([s_{i-1}^{k-1}, s_i^{k-1}))$, $p_2 := p_\epsilon([s_i^{k+1}, s_{i+1}^{k+1}))$ be the optimal prices on the those unaccounted for segments, we will need to determine the prices for each new segments in \mathcal{S}'_k and \mathcal{S}''_k such that the combined revenue from them fully covers the revenue from the unaccounted for segments.

Note that it is unclear which of p_1 and p_2 is larger, we will argue in two sub-cases based on their ordering. Suppose we have $p_1 \leq p_2$, and let the price for new segments in \mathcal{S}'_k and \mathcal{S}''_k be

$$p_\epsilon([s_{i-1}^{k-1}, s_{i+1}^{k+1}]) = p_1, \quad p_\epsilon([s_i^{k+1}, s_i^{k-1}]) = p_2.$$

and let $\mathcal{R}_\epsilon([s_{j-1}, s_j], p)$ be the revenue from segment $[s_{j-1}, s_j]$ when the price on that segment is p . The difference in revenue between the new and unaccounted for segments is then,

$$\begin{aligned} & \mathcal{R}_\epsilon([s_{i-1}^{k-1}, s_{i+1}^{k+1}], p_1) + \mathcal{R}_\epsilon([s_i^{k+1}, s_i^{k-1}], p_2) - \mathcal{R}_\epsilon([s_{i-1}^{k-1}, s_i^{k-1}], p_1) - \mathcal{R}_\epsilon([s_i^{k+1}, s_{i+1}^{k+1}], p_2) \\ &= (\mathcal{R}_\epsilon([s_{i-1}^{k-1}, s_{i+1}^{k+1}], p_1) - \mathcal{R}_\epsilon([s_{i-1}^{k-1}, s_i^{k-1}], p_1)) + (\mathcal{R}_\epsilon([s_i^{k+1}, s_{i+1}^{k+1}], p_2) - \mathcal{R}_\epsilon([s_i^{k+1}, s_i^{k-1}], p_2)) \\ &= \mathcal{R}_\epsilon([s_{i+1}^{k+1}, s_i^{k-1}], p_2) - \mathcal{R}_\epsilon([s_{i+1}^{k+1}, s_i^{k-1}], p_1), \end{aligned}$$

where the second equality follows from the fact that $[s_{i-1}^{k-1}, s_{i+1}^{k+1}] \subset [s_{i-1}^{k-1}, s_i^{k-1}]$, and $[s_i^{k+1}, s_{i+1}^{k+1}] \subset [s_i^{k+1}, s_i^{k-1}]$. Since every element in $[s_i^{k+1}, s_{i+1}^{k+1}]$ is less than any element in $[s_{i+1}^{k+1}, s_i^{k-1}]$, by Eq. (EC.8) and Eq. (EC.9), we have the price dominance

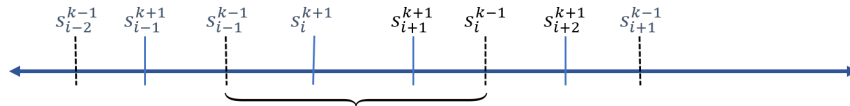
$$p_1 \leq p_2 \leq p_\epsilon([s_{i+1}^{k+1}, s_i^{k-1}]),$$

which implies

$$\mathcal{R}_\epsilon([s_{i+1}^{k+1}, s_i^{k-1}], p_2) \geq \mathcal{R}_\epsilon([s_{i+1}^{k+1}, s_i^{k-1}], p_1).$$

Thus, $\mathcal{R}_{(k-1)XP} + \mathcal{R}_{(k+1)XP} \leq \mathcal{R}_{kXP}(\mathcal{S}'_k) + \mathcal{R}_{kXP}(\mathcal{S}''_k) \leq 2\mathcal{R}_{kXP}$. In the second sub-case when $p_2 \leq p_1$, the proof follows symmetrically now using p_2 for the price of the new segment in \mathcal{S}'_k and p_1 for the price in the new segment for \mathcal{S}''_k , we omit it for brevity.

Figure EC.3 Case 2 of Theorem 5.



Note. Depicted are the $i+1$ segment end points of the an interval segmentation of size $k+1$ (solid line) and i segment of size $k-1$ segmentation (dashed line). The $i+1^{th}$ segment $[s_i^{k+1}, s_{i+1}^{k+1}]$ of $(k+1)$ segmentation is fully contained in the i^{th} segment $[s_{i-1}^{k-1}, s_i^{k-1}]$ of $(k-1)$ segmentation.

Figure EC.4 New Feasible k -Segmentations for Case 2 of Theorem 5.



Note. Depicted are the constructed size k interval segmentations. The new k -segmentations are constructed by crossing over $(k-1)$ segmentation and $(k+1)$ segmentation at s_{i-1}^{k-1} . In the left panel, \mathcal{S}'_k is shown; before s_{i-1}^{k-1} , it contains the segments of \mathcal{S}_{k-1} , after s_{i+1}^{k+1} , it contains segments of \mathcal{S}_{k+1} . In the right panel, \mathcal{S}''_k is shown; before s_i^{k+1} , it contains the segments of \mathcal{S}_{k+1} , after s_i^{k-1} , it contains the segments of \mathcal{S}_{k-1} .

To complete the proof, we now must show that Case 1 and Case 2 are the only cases i.e., if the condition for Case 1 does not hold, the condition for Case 2 must hold for some i . To see this, imagine Case 1 does not hold, then $s_1^{k+1} > s_0^{k-1}$. If then $s_2^{k+1} < s_1^{k-1}$ the proof is complete, so assume $s_2^{k+1} \geq s_1^{k-1}$. If then $s_3^{k+1} < s_2^{k-1}$ the proof is complete and so on. Iterating, since $s_{k+1}^{k+1} = s_{k-1}^{k-1}$ the sequence of deductions must terminate at some i for which Case 2 holds. Thus Case 1 and Case 2 cover all cases, which completes the proof. \square

Appendix C: Transforming a Probit Regression Model into a Valuation Model

In this section we overview how to transform a prediction model for sales into a linear valuation model. Note, in reality we cannot observe customer's valuation for one product directly. Instead, we can see whether the customer will buy the product or not for the offered price p (see Cameron and James (1987) for more details). Assume that the unobserved continuous dependent variable Y is the customer's true valuation or willingness to pay (WTP) for the product. Further, the relation of Y and customer's feature X is

$$Y = \beta_0 + X\beta + \epsilon, \quad (\text{EC.10})$$

where $\epsilon \sim N(0, \sigma)$. Customer i 's decision I_i will be

$$I_i = \begin{cases} 1, & \text{if } Y_i \geq p_i, \\ 0, & \text{otherwise,} \end{cases} \quad (\text{EC.11})$$

where p_i is the price offered to customer i . Then, the probability that customer i will buy the product is

$$\begin{aligned} \Pr(I_i = 1) &= \Pr(Y_i \geq p) = \Pr(+\epsilon_i \geq p) \\ &= \Pr\left(\frac{\epsilon_i}{\sigma} \geq \frac{p - X_i\beta}{\sigma}\right) \\ &= 1 - \Phi\left(\frac{p - \beta_0 - X_i\beta}{\sigma}\right), \end{aligned}$$

where Φ is the cumulative distribution function for the standard normal distribution. Using p and X as explanatory variable, the probit regression model is

$$\Pr(I = 1|X) = 1 - \Phi(\beta'_0 + p\beta_p + X\beta').$$

Therefore, we can use the maximum likelihood estimator (MLE) of probit regression model to recover the regression model of customer's valuation, i.e.,

$$\hat{\beta}_0 = \frac{\hat{\beta}'_0}{\hat{\beta}'_p}, \quad \hat{\beta} = \frac{\hat{\beta}'}{\hat{\beta}'_p}, \quad \hat{\sigma} = \frac{1}{\hat{\beta}'_p},$$

where $\hat{\beta}'_0$, $\hat{\beta}'_p$, $\hat{\beta}'$ are the MLE of β'_0 , β_p , β' . Further, the regression model of customer's valuation recovered from probit regression model is asymptotically unbiased if the price variance is large enough.

Appendix D: Constant Factor Approximation for General Error

In this section we will describe how to obtain a $1-1/e$ approximation of the optimal FBMSPP when the residuals are independent and follow an arbitrary distribution, as sketched in Remark 1.

$(1 - 1/e)$ Approximation Algorithm: Our polynomial time approximation algorithm will leverage the *submodularity* of the objective function for FBMSPP, defined as follows:

DEFINITION EC.1 (SUBMODULARITY). A set function $f : 2^V \rightarrow \mathbb{R}$ is submodular if for every $A, B \subseteq V$,

$$f(A \cap B) + f(A \cup B) \leq f(A) + f(B).$$

An important subclass of submodular functions are those which are monotone, i.e., functions for which enlarging the choice set cannot cause the function value to decrease.

DEFINITION EC.2 (MONOTONICITY). A set function $f : 2^V \rightarrow \mathbb{R}$ is monotone if for every $A \subseteq B \subseteq V$, $f(A) \leq f(B)$.

We will show that the objective function for FBMSP can be expressed as a set function over prices which is monotone and submodular. Note that for n customers with predicted valuations $\{x_i\}_{i=1}^n$ and for error distribution ϵ supported on m points, there are at most $O(nm)$ distinct possible valuation realizations. Further, any optimal price for a segment must correspond to one of these realizations (since if not, raising the price until it reaches a valuation in the support is strictly revenue improving). Thus the set of potential prices is a polynomially sized set equivalent to the set of potential realized valuations, and the revenue objective of FBMSP can be viewed as a set function over that set.

Specifically, if f is the revenue function of FBMSP on price set A , it takes the form,

$$f(A) = \sum_{i=1}^n \max_{p \in A} p \bar{F}(p - \mu(x_i)).$$

Then expressed as a set function over the prices, optimal FBMSP is the solution to

$$\max_{|A| \leq k} \sum_{i=1}^n \max_{p \in A} p \bar{F}(p - \mu(x_i)).$$

The monotonicity of the revenue objective is easy to see since, by definition, enlarging the set of possible prices that can be used for a segment will keep at least the same revenue as for a smaller set of prices. The submodularity comes from the fact, any customer i facing the prices in price set $A \cap B$ will result in less revenue than when facing the prices in price set A or B , i.e.,

$$\max_{p \in A \cap B} p \bar{F}(p - \mu(x_i)) \leq \min \left\{ \max_{p \in A} p \bar{F}(p - \mu(x_i)), \max_{p \in B} p \bar{F}(p - \mu(x_i)) \right\},$$

further,

$$\max_{p \in A \cup B} p \bar{F}(p - \mu(x_i)) = \max \left\{ \max_{p \in A} p \bar{F}(p - \mu(x_i)), \max_{p \in B} p \bar{F}(p - \mu(x_i)) \right\}.$$

Note that

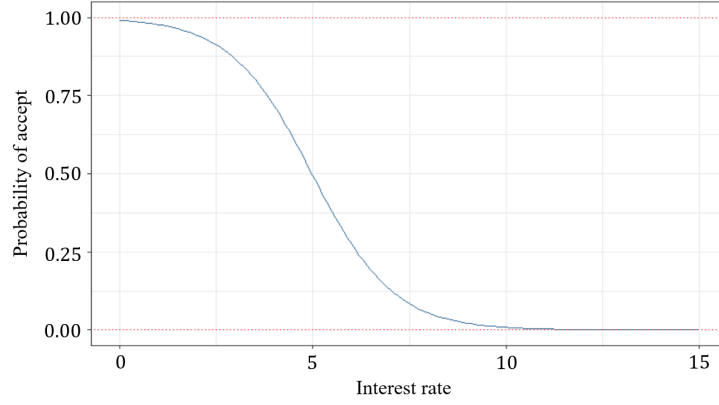
$$\begin{aligned} \max_{p \in A} p \bar{F}(p - \mu(x_i)) + \max_{p \in B} p \bar{F}(p - \mu(x_i)) &= \min \left\{ \max_{p \in A} p \bar{F}(p - \mu(x_i)), \max_{p \in B} p \bar{F}(p - \mu(x_i)) \right\} + \\ &\quad \max \left\{ \max_{p \in A} p \bar{F}(p - \mu(x_i)), \max_{p \in B} p \bar{F}(p - \mu(x_i)) \right\}, \end{aligned}$$

and thus combining these observations and summing over all customers proves submodularity of the objective function.

Note that maximizing positive monotone submodular functions maximization with cardinality constraints is NP-hard in general (see Krause and Golovin (2014)). The cardinality constraint in FBMSP is the number of segments (in some sense, the same as number of prices). Nemhauser et al. (1978) shows that a greedy algorithm can obtain an approximation guarantee of $(1 - 1/e)$ for class of monotone submodular functions maximizing with cardinality constraints. Since FBMSP problem is can be written as a problem of maximizing a monotone submodular function with cardinality constraints, it can be approximated at least within a factor of $(1 - 1/e)$ via the same greedy algorithm.

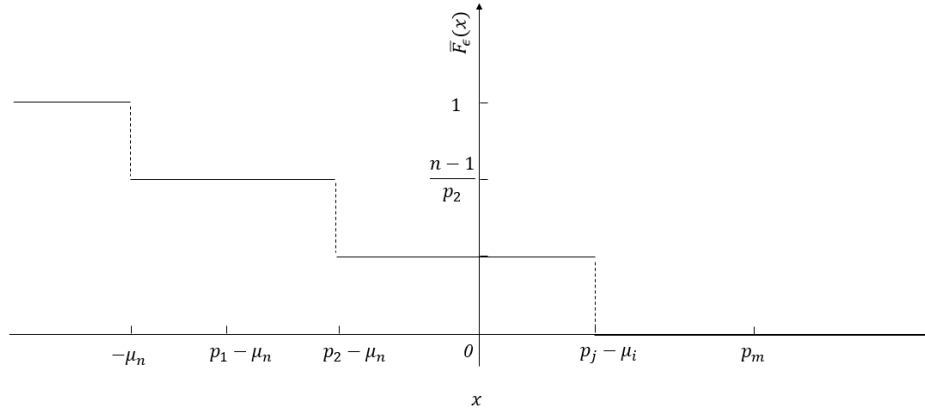
Appendix E: Omitted Figures

Figure EC.5 Prediction of the probit regression model.



Note. Depicted is the output of a probit regression model to predict the probably of mortgage acceptance, our proxy for purchase in the loan setting. The model is trained using features in Table 2.

Figure EC.6 An example of the error distribution \bar{F}_ϵ , constructed for the proof of Theorem 1.



Note. Depicted is an example of the cCDF \bar{F}_ϵ constructed to prove the hardness of FBMSP. Note on the x -axis are the valuation support points, and that the resultant error distribution is a step-function on these supports.