

# Statistical Inference

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# 1 Probability Theory

## 1.5(a)

*Solution.* A U.S. birth results in female identical twins. □

## 1.5(b)

*Solution.*

$$P(A \cap B \cap C) = P(A|B \cap C)P(B|C)P(C) = P(A|B)P(B|C)P(C) = \frac{1}{2} \frac{1}{3} \frac{1}{90} = \frac{1}{540}.$$

□

## 1.24(b)

*Solution.* Suppose  $E_i = \{\text{head first appears on } i\text{th toss}\}$ , then

$$P(\text{A wins}) = P\left(\bigcup_{i=k}^{\infty} E_{2k+1}\right) = \sum_{k=0}^{\infty} P(E_{2k+1}) = \sum_{k=0}^{\infty} p(1-p)^{2k+1} = \frac{p}{1 - (1-p)^2}.$$

□

## 1.31(a)

*Proof.* To get the average  $(x_1 + \dots + x_n)/n$ , we need the unordered sample to be  $\{x_1, x_2, \dots, x_n\}$ . The number of ordered samples which results in it is  $n!$  and there are  $n^n$  ordered samples in total. Hence, the probability is  $n!/n^n$ .

For any other resulted average, there will exist some double counting when counting the ordered samples. Therefore, the outcome with average  $(x_1 + \dots + x_n)/n$  is most likely. □

## 1.33

*Solution.*

$$\begin{aligned} P(\text{male}|\text{color-blind}) &= P(\text{color-blind}|\text{male}) \frac{P(\text{male})}{P(\text{color-blind})} \\ &= 0.05 \times \frac{0.5}{0.5 \times 0.05 + 0.5 \times 0.0025} \\ &= \frac{20}{21} = 0.9524. \end{aligned}$$

□

**1.36**

*Solution.* The probabilities of all shots being missed and the target being hit exactly once are  $(4/5)^5 = 0.32768$  and  $5 \times (1/5)(4/5)^4 = 0.4096$  respectively. Hence,

$$P(\text{being hit at least twice}) = 1 - 0.32768 - 0.4096 = 0.26272.$$

And

$$\begin{aligned} & P(\text{being hit at least twice} | \text{being hit at least once}) \\ &= \frac{P(\text{being hit at least twice})}{P(\text{being hit at least once})} = \frac{0.26272}{0.4096} = 0.6414. \end{aligned}$$

□

**1.39(a)**

*Proof.*  $A$  and  $B$  are mutually exclusive means that  $A \cap B = \emptyset$ . Hence,  $P(A \cap B) = 0$ . However,  $P(A), P(B) > 0$ . Therefore,  $P(A \cap B) \neq P(A)P(B)$ . □

**1.39(b)**

*Proof.* As  $A$  and  $B$  are independent,  $P(A \cap B) = P(A)P(B) > 0$ , which implies that  $A \cap B \neq \emptyset$ . □

**Notes on 1.39** An intuitive proof: Since  $A$  and  $B$  are mutually exclusive, if we know that  $A$  did not happen, then the possibility that  $B$  happened will increase. Hence, they are not independent.

**1.52**

*Proof.* Clear that  $g(x) \geq 0$  for all  $x \in \mathbb{R}$  and

$$\begin{aligned} \int_{-\infty}^{\infty} g(x) dx &= \int_{x_0}^{\infty} \frac{f(x)}{1 - F(x_0)} dx \\ &= \frac{1}{1 - F(x_0)} \left( \int_{-\infty}^{\infty} f(x) dx - \int_{-\infty}^{x_0} f(x) dx \right) \\ &= \frac{1}{1 - F(x_0)} (1 - F(x_0)) = 1. \end{aligned}$$

Hence, by Theorem 1.6.5,  $g(x)$  is a pdf. □