概统作业 (Week 10)

May 13, 2023

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(0) 选 D.

(1)

$$\begin{split} E(Y_1) &= \int_{-\infty}^{+\infty} y \cdot \frac{f_1(y) + f_2(y)}{2} dy = \frac{1}{2} \cdot \left(\int_{-\infty}^{+\infty} y \cdot f_1(y) dy + \int_{-\infty}^{+\infty} y \cdot f_2(y) dy \right) = \frac{E(X_1) + E(X_2)}{2}. \\ E(Y_2) &= E\left(\frac{X_1 + X_2}{2} \right) = \frac{1}{2} \cdot \left(E(X_1 + X_2) \right) = \frac{E(X_1) + E(X_2)}{2} = E(Y_1). \end{split}$$

(2)

$$\begin{split} E(Y_1^2) &= \int_{-\infty}^{+\infty} y^2 \cdot \frac{f_1(y) + f_2(y)}{2} dy = \frac{1}{2} \cdot \left(\int_{-\infty}^{+\infty} y^2 \cdot f_1(y) dy + \int_{-\infty}^{+\infty} y^2 \cdot f_2(y) dy \right) = \frac{E(X_1^2) + E(X_2^2)}{2}. \\ & Var(Y_1) = E(Y_1^2) - E(Y_1)^2 = \frac{E(X_1^2) + E(X_2^2)}{2} - \left(\frac{E(X_1) + E(X_2)}{2} \right)^2 \\ & \mathbb{Z} \ Var(X_1) = E(X_1^2) - E(X_1)^2, \ Var(X_2) = E(X_2^2) - E(X_2)^2, \ \not\equiv \\ & Var(Y_1) = \frac{Var(X_1) + Var(X_2)}{4} + \frac{E(X_1^2) + E(X_2^2) - 2E(X_1)E(X_2)}{4} \end{split}$$

对于 Y_2 ,有

$$Var(Y_2) = Var\left(\frac{X_1 + X_2}{2}\right) = \frac{Var(X_1) + Var(X_2)}{4}$$

又由方差 $Var(X) = E(X^2) - E(X)^2 \ge 0$,有

$$\begin{cases} E(X_1^2) & \geq E(X_1)^2 \\ E(X_2^2) & \geq E(X_2)^2 \end{cases} \Rightarrow E(X_1^2) + E(X_2^2) - 2E(X_1)E(X_2) \geq E(X_1)^2 + E(X_2)^2 - 2E(X_1)E(X_2) \geq 0.$$

故

$$Var(Y_1) - Var(Y_2) \geq 0 \ \Rightarrow \ Var(Y_1) \geq Var(Y_2).$$

2 (P174 T27)

$$\begin{split} P(X=k) &= \binom{n}{k} p^k (1-p)^{n-k} \ (k=0,1,\dots,n) \\ M_X(s) &= E(e^{sX}) = \sum_{k=0}^n e^{sk} \binom{n}{k} p^k (1-p)^{n-k} = \sum_{k=0}^n \binom{n}{k} (pe^s)^k (1-p)^{n-k} = (pe^s+1-p)^n. \end{split}$$

(2)

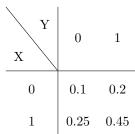
$$\begin{split} P(X=k) &= \frac{e^{-\lambda}\lambda^k}{k!} \ (k=0,1,2,\ldots) \\ M_X(s) &= \sum_{k=0}^{\infty} \frac{e^{-\lambda}\lambda^k}{k!} e^{sk} = e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda e^{s^k}}{k!} = e^{-\lambda} e^{\lambda e^s} = e^{\lambda(e^s-1)}. \end{split}$$

3 (P175 T30)

(1)

$$\begin{split} E(X) &= (0.05 + 0.05 + 0.1) - (0.1 + 0.2 + 0.2) = -0.3. \\ E(Y) &= (0.2 + 0.15 + 0.1) - (0.1 + 0.05 + 0.05) = 0.25. \\ E(XY) &= (0.1 + 0.1) - (0.2 + 0.05) = -0.05. \\ Cov(X, Y) &= E(XY) - E(X)E(Y) = -0.05 + 0.075 = 0.025. \end{split}$$

(2)



$$E(X^2) = (0.05 + 0.05 + 0.1) + (0.1 + 0.2 + 0.2) = 0.7.$$

$$E(Y^2) = (0.2 + 0.15 + 0.1) + (0.1 + 0.05 + 0.05) = 0.65$$

$$E(X^2Y^2) = 0.45.$$

$$Cov(X^2, Y^2) = E(X^2Y^2) - E(X^2)E(Y^2) = 0.45 - 0.7 \times 0.65 = -0.005.$$

4 (P175 T32)

(1)

$$Cov(\alpha X + \beta Y, \alpha X - \beta Y) = \alpha^2 Var(X) - \beta^2 Var(Y) = (\alpha^2 - \beta^2)\sigma^2.$$

(2) 又 $(\alpha X + \beta Y, \alpha X - \beta Y)$ 服从二维正态分布, 故当 $\alpha = \pm \beta$ 时, $\alpha X + \beta Y, \alpha X - \beta Y$ 相互独立.

5 (P175 T34)

(1)

$$\begin{split} E(X) &= \iint_{(x,y) \in G} x dx dy = 0. \qquad E(Y) = \iint_{(x,y) \in G} y dx dy = 0. \\ E(XY) &= \iint_{(x,y) \in G} x y dx dy = 0. \end{split}$$

故

$$Cov(X,Y) = E(XY) - E(X)E(Y) = 0.$$

$$\begin{split} f(x,y) &= \frac{1}{S} = \frac{1}{2}. \\ f_X(x) &= \int_{-\infty}^{+\infty} dy \int_{|x|+|y| \le 1} f(x,y) dx = \begin{cases} 1+x, & (-1 \le x \le 0) \\ 1-x, & (0 < x \le 1) \end{cases} \\ f_Y(y) &= \int_{-\infty}^{+\infty} dx \int_{|x|+|y| \le 1} f(x,y) dy = \begin{cases} 1+y, & (-1 \le y \le 0) \\ 1-y, & (0 < y \le 1) \end{cases} \\ f(x,y) \neq f_X(x) f_Y(y) \implies X, Y \text{ π 相 五独立}. \end{split}$$