$\nearrow$  1.(P172, 4)设X为一个连续型随机变量, 试对下列各种情形, 计算Var(X).

(1) 若X的密度函数为

$$f(x)=rac{x}{\sigma^2}{
m exp}\left\{-rac{x^2}{2\sigma^2}
ight\}, x>0,$$

其中 $\sigma > 0$ 为常数,则称X服从瑞利(Rayleigh)分布;

(2) 若X的密度函数为

$$f(x) = rac{\Gamma(lpha + eta)}{\Gamma(lpha)\Gamma(eta)} x^{lpha - 1} (1 - x)^{eta - 1}, 0 < x < 1,$$

其中 $\alpha, \beta > 0$ 为常数,  $\Gamma(x)$ 为 $\Gamma$ 函数, 则称X服从 $\beta$ 分布;

(3) 若X的密度函数为

$$f(x) = \frac{k}{\lambda} \Big(\frac{x}{\lambda}\Big)^{k-1} \exp{\left\{-\left(\frac{x}{\lambda}\right)^k\right\}}, x > 0,$$

其中 $k, \lambda > 0$ 为常数,则称X服从韦布尔分布.

解:

$$(1)EX = \frac{\sqrt{2\pi}\sigma}{2}$$

$$\begin{split} EX^2 &= \int_0^\infty x^2 f(x) \mathrm{d}x \\ &= \int_0^\infty \frac{x^3}{\sigma^2} \mathrm{exp} \left\{ -\frac{x^2}{2\sigma^2} \right\} \mathrm{d}x \\ &= \int_0^\infty \frac{t}{2\sigma^2} \mathrm{exp} \left\{ -\frac{t}{2\sigma^2} \right\} \mathrm{d}t \\ &= t \mathrm{exp} \left\{ -\frac{t}{2\sigma^2} \right\} \Big|_\infty^0 + \int_0^\infty \mathrm{exp} \left\{ -\frac{t}{2\sigma^2} \right\} \mathrm{d}t \\ &= 2\sigma^2 \mathrm{exp} \left\{ -\frac{t}{2\sigma^2} \right\} \Big|_\infty^0 \end{split}$$

$$Var(X) = EX^2 - (EX)^2 = 2\sigma^2 - rac{\pi}{2}\sigma^2$$

(2) $EX = \frac{\alpha}{\alpha + \beta}$ 

$$\begin{split} EX^2 &= \int_0^1 x^2 f(x) \mathrm{d}x \\ &= \int_0^1 \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha + 1} (1 - x)^{\beta - 1} \mathrm{d}x \\ &= \int_0^1 \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha + 2 - 1} (1 - x)^{\beta - 1} \mathrm{d}x \\ &= \frac{B(\alpha + 2, \beta)}{B(\alpha, \beta)} \\ &= \frac{(\alpha + 1)\alpha}{(\alpha + \beta + 1)(\alpha + \beta)} \end{split}$$

$$Var(X) = EX^2 - (EX)^2 = rac{lphaeta}{(lpha + eta + 1)(lpha + eta)^2}$$

$$(3)EX = \lambda\Gamma\left(\frac{1}{k}+1\right) \diamondsuit t = \left(\frac{x}{\lambda}\right)^k$$
,  $\mathbb{M}dt = \frac{kx^{k-1}}{\lambda^k}dx$ ,  $x = \lambda t^{1/k}$ 

$$\begin{split} EX^2 &= \int_0^\infty x^2 f(x) \mathrm{d}x \\ &= \int_0^\infty k \lambda \Big(\frac{x}{\lambda}\Big)^{k+1} \exp\Big\{-\Big(\frac{x}{\lambda}\Big)^k\Big\} \mathrm{d}x \\ &= \int_0^\infty \lambda^2 t^{2/k+1-1} e^{-t} \mathrm{d}t \\ &= \lambda^2 \Gamma(\frac{2}{k}+1) \end{split}$$

$$Var(X) = EX^2 - (EX)^2 = \lambda^2 \left[\Gamma\left(rac{2}{k} + 1
ight) - \Gamma^2\left(rac{1}{k} + 1
ight)
ight]$$

$$sgn(x) = egin{cases} 0, & X = 0, \\ 1, & X > 0, \\ -1, & X < 0. \end{cases}$$

(1)若X服从U(-2,1), 试求Var(sgn(X));

(2)若X服从标准正态分布, 试求E[sgn(X)X].

解:

$$\begin{array}{l} (1)E(sgn(X)) = \int_{-\infty}^{\infty} sgn(x)f(x)\mathrm{d}x = \int_{-2}^{0} -\frac{1}{3}\,\mathrm{d}x + \int_{0}^{1}\,\frac{1}{3}\,\mathrm{d}x = -\frac{1}{3}\,, \ E(sgn(X)^{2}) = \int_{-\infty}^{\infty} f(x)\mathrm{d}x = 1, \\ Var(sgn(X)) = E(sgn(X)^{2}) - (E(sgn(X)))^{2} = 1 - \frac{1}{9} = \frac{8}{9} \end{array}$$

(2) 
$$E[sgn(X)X] = \int_{-\infty}^{\infty} sgn(x)xf(x)\mathrm{d}x = \int_{-\infty}^{\infty} |x| \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{x^2}{2}\right\}\mathrm{d}x = \sqrt{\frac{2}{\pi}}$$

**3.(P173,10)**设随机变量X只能取有限个正值 $x_1, x_2, \dots, x_k (k \ge 2)$ , 证明:

0

$$\lim_{n o\infty}rac{E(X^{n+1})}{E(X^n)}=\max_{1\leqslant i\leqslant k}x_i.$$

证明:设 $P(X = x_i) = p_i, i = 1, 2, ..., k$ , 其中 $x_t = \max_{1 \le i \le k} x_i$ , 因此 $\lim_{n \to \infty} (x_i/x_t)^n = 0, i \ne t$ .

则 $E(X^n) = \sum_{i=1}^k p_i x_i^n$ 

$$egin{aligned} \lim_{n o \infty} rac{E(X^{n+1})}{E(X^n)} &= \lim_{n o \infty} rac{\sum_{i=1}^k p_i x_i^{n+1}}{\sum_{i=1}^k p_i x_i^n} \ &= \lim_{n o \infty} x_t rac{\sum_{i=1}^k p_i (x_i/x_t)^{n+1}}{\sum_{i=1}^k p_i (x_i/x_t)^n} \ &= x_t rac{p_t}{p_t} \ &= x_t = \max_{1 \leqslant i \leqslant k} x_i. \end{aligned}$$

4.(P172,5)设随机变量X的密度函数为

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$$f(x) = ax^2 + bx + c, 0 < x < 1,$$

且已知E(X) = 0.5, Var(X) = 0.15, 试求常数a, b, c.

解:

由题,  $EX^2 = Var(X) + (EX)^2 = 0.4$ 

$$\begin{cases} \int_0^1 f(x) dx = 1\\ \int_0^1 x f(x) dx = 0.5\\ \int_0^1 x^2 f(x) dx = 0.4 \end{cases}$$

解得a = 12, b = -12, c = 3.

⊘ 5.证明下面两个广义切比雪夫不等式:

 $\diamondsuit \sigma^2 = Var(X). \ \forall x, a > 0,$ 

$$P(X - EX \geqslant x) \leqslant \frac{\sigma^2 + a^2}{(x+a)^2},$$

$$P(X-EX\geqslant x)\leqslant rac{\sigma^2}{x^2+\sigma^2}.$$

根据这两个不等式, 证明X的中位数m(X)满足:  $|EX - m(X)| \leq \sigma$ .

证明: 令T=X-EX,则 $ET=0,ET^2=\sigma^2$ 取 $x+a\geqslant 0$ ,则利用Markov不等式 $P(Y\geqslant \varepsilon)\leqslant \frac{EY}{\varepsilon}$ ,有

$$P(X - EX \geqslant x) = P(T + a \geqslant x + a \geqslant 0) = P\{(T + a)^2 \geqslant (x + a)^2\} \leqslant \frac{E(T + a)^2}{(x + a)^2} = \frac{ET^2 + 2aET + a^2}{(x + a)^2} = \frac{\sigma^2 + a^2}{(x + a)^2}$$

在上式中,取 $a = \frac{\sigma^2}{x}$ ,得

$$P(X - EX \geqslant x) \leqslant rac{\sigma^2 + \sigma^4/x^2}{(x + \sigma^2/x)^2} = rac{\sigma^2}{x^2 + \sigma^2}$$

取 $x = \sigma$ ,则 $P(X - EX \geqslant \sigma) \leqslant \frac{1}{2}$ ,而 $P\{X \geqslant m(X)\} = \frac{1}{2}$ ,得 $EX - m(X) \geqslant -\sigma$ 

同时, 令Y=-X,有 $P(-Y\leqslant -EY-\sigma)=P(X\leqslant EX-\sigma)\leqslant rac{1}{2}$ ,得 $EX-m(X)\leqslant \sigma$ ,故 $|EX-m(X)|\leqslant \sigma$