

Group 1 – Ridge Regression and Principal Component Analysis

Time Series Analysis and Machine Learning (HS-2025)

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This report presents the submission of **Group 1** for the **Time Series Analysis and Machine Learning** course (HS-2025). There are two exercises in this assignment. **First**, we replicate both panels in *James — An Introduction to Statistical Learning with Applications in R* (Figure 6.4), which display the standardized ridge regression coefficients of the credit data set. **Second**, we explain how the two principal component lines in Figure 6.14 are derived, including their meaning and why they are orthogonal to each other.

Exercise 1 - Ridge Regression

1. Data Preparation

We start by importing the necessary libraries used throughout the project.

```
# for lm.ridge
library(MASS)
```

The raw credit_data.csv file is read and the binary Yes/No indicators for the predictors “Own”, “Student”, and “Married” are converted into numeric binary variables. This conversion ensures that the model correctly interprets these predictors as binary inputs, with “Yes” corresponding to 1 and “No” corresponding to 0.

```
# Load data and convert categorical Yes/No flags to 0/1 dummies
data <- read.csv("credit_data.csv")

data$Student <- ifelse(data$Student == "Yes", 1, 0)
data$Own      <- ifelse(data$Own      == "Yes", 1, 0)
data$Married  <- ifelse(data$Married == "Yes", 1, 0)
```

The categorical predictor “Region” is transformed into three dummy variables representing the regions South, West, and East. These dummy variables are created without including an intercept (in order to avoid perfect multicollinearity), allowing each region to be treated as a separate category in the design matrix. The new indicator variables are appended to the dataset for subsequent analysis.

```
# Build Region indicators and bind them to the main data frame
region_dummies <- model.matrix(~ Region - 1, data = data)
data <- cbind(data[, !(names(data) %in% "Region")], region_dummies)
```

2. Design Matrix Construction

In this step, we identify the relevant predictors and the response variable to be used in our regression model. The predictor set includes both continuous variables (such as income, limit, and rating) and binary or categorical variables (such as married and region).

To enhance the interpretability and comparability of the model, we standardize all predictors so that each has a mean of zero and a standard deviation of one. Standardization ensures that no variable dominates the estimation simply because it is measured in larger numeric units. It also allows ridge regression to apply the same degree of shrinkage to all coefficients, which improves the stability and comparability of the results.

Without standardization, the penalization term in ridge regression would depend not only on the regularization parameter λ but also on the scale of the j th predictor.

```
# Define predictor matrix X and response vector Y
X <- dplyr::select(
  data, Income, Limit, Rating, Cards, Age, Education, Own, Student,
  Married, RegionEast, RegionWest, RegionSouth
)
Y <- data$Balance

# Standardize predictors to mean 0, variance 1
```

```
X_standardized <- scale(X)
X_std <- as.data.frame(X_standardized)
```

Note: We work with the already build in scale function instead of the equation 6.6 given in the chapter six. Hence, we also mean adjust the variables, while the book only divides by the standard deviation.

3. Helper Functions

To make the analysis easier and more consistent, we define two short helper functions that perform the key calculations used throughout the report.

The first function computes the ridge regression coefficients for a given value of the regularization parameter λ following the model:

$$\hat{\beta}_\lambda = \arg \min_{\beta} \left\{ \sum_{i=1}^n (y_i - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda \sum_{j=1}^p \beta_j^2 \right\}$$

The second function computes the ordinary least squares (OLS) coefficients without an intercept following the model:

$$\hat{\beta}_{OLS} = \arg \min_{\beta} \left\{ \sum_{i=1}^n (y_i - \sum_{j=1}^p \beta_j x_{ij})^2 \right\}$$

```
# Ridge coefficients for a chosen penalty lambda with no intercept (keep all dummies)
compute_coefficients <- function(lambda, X, y) {
  ridge_fit <- lm.ridge(y ~ as.matrix(X) - 1, lambda = lambda)
  coef(ridge_fit)
}

# OLS coefficients with no intercept (keep all dummies)
compute_ols <- function(X, y) {
  df <- data.frame(y = y, X)
  fit <- lm(y ~ . - 1, data = df)
  b <- coef(fit)
  b[colnames(X)]
}
```

Compute the OLS coefficients once for later comparisons.

```
ols_coefficients <- compute_ols(X_std, Y)
```

4. Ridge Paths Across Lambda Grid

In this section, we examine how the estimated coefficients change as the regularization strength λ increases. To do this, we define a sequence of 100 values for λ that span a wide range, from very small values (close to ordinary least squares) to very large ones (strong shrinkage).

For each λ , the ridge regression is re-estimated, and the resulting coefficients are stored. This creates a set of “ridge paths” that show how each variable’s influence gradually shrinks toward zero as the penalty term $\lambda \sum_{j=1}^p \beta_j^2$ grows.

```
# Log-spaced lambda grid
lambdas <- 10^seq(-2, 5, length.out = 100)

# Collect ridge coefficients for every lambda
coefficients <- sapply(lambdas, function(lambda) {
  compute_coefficients(lambda, X_std, Y)
})
```

```
# Clean up into a data frame
df <- t(coefficients)
colnames(df) <- gsub("as\\\.matrix\\(X\\)", "", colnames(df))
df <- cbind(lambda = lambdas, df)
df <- as.data.frame(df)
```

5. Norm Calculations

In this step, we compare the ℓ_2 norm of the ridge coefficients to the ℓ_2 norm of the OLS coefficients which measure the square root of the sum of the squared coefficients. As the regularization strength λ increases, the ridge coefficients shrink, and their norm becomes smaller relative to the OLS solution.

The resulting ratio (ridge norm divided by OLS norm) provides a convenient way to express the degree of shrinkage on a standardized 0–1 scale. This ratio will later serve as a rescaled x-axis when visualizing how coefficients change with increasing regularization.

The ratio is given by:

$$\frac{\|\hat{\beta}_\lambda^R\|_2}{\|\hat{\beta}_{OLS}\|_2} = \frac{\sqrt{\sum_{j=1}^p (\hat{\beta}_{j,\lambda}^R)^2}}{\sqrt{\sum_{j=1}^p (\hat{\beta}_{j,OLS})^2}}$$

```
# Ridge 2 norm per lambda
ridge_norms <- apply(df[, !(names(df) %in% "lambda")], 1,
                      function(row) sqrt(sum(row^2)))

# OLS 2 norm (drop NA caused by omitted dummy)
ols_norm <- sqrt(sum(ols_coefficients^2, na.rm = TRUE))

# Attach the ridge/OLS norm ratio to the data frame
df$norm_ratio <- ridge_norms / ols_norm
```

6. Visualization

6.1 Lambda-on-Log-Scale Plot

This plot shows how each standardized coefficient changes as the regularization strength λ increases, using a logarithmic x-axis. At small values of λ , the model behaves like OLS with little shrinkage, while larger values gradually pull the coefficients toward zero. The highlighted variables (Income, Limit, Rating, and Student) represent the most influential predictors, while the others are shown in gray for context. Overall, the plot illustrates how ridge regression balances model complexity and stability by progressively reducing the size of the coefficients.

```
plot(df$lambda, df$Income, type = "l", log = "x", lwd = 2,
      ylim = c(-300, 450), col = "black",
      xlab = expression(lambda),
      ylab = "Standardized Coefficients",
      xaxt = "n", yaxt = "n")
axis(side = 1, at = c(0.01, 1, 100, 10000),
      labels = c("1e-02", "1e+00", "1e+02", "1e+04"))
axis(side = 2, at = seq(-300, 400, by = 100),
      labels = seq(-300, 400, by = 100))
```

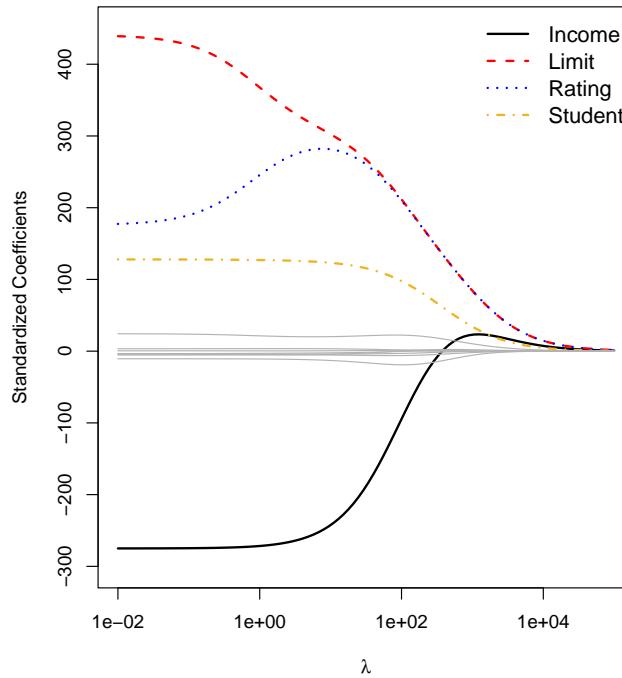
```

# Highlighted series
lines(df$lambda, df$Limit, col = "red", lty = 2, lwd = 2)
lines(df$lambda, df$Rating, col = "blue", lty = 3, lwd = 2)
lines(df$lambda, df$Student, col = "goldenrod2", lty = 4, lwd = 2)

# Remaining predictors in gray
lines(df$lambda, df$Cards, col = "grey70")
lines(df$lambda, df$Age, col = "grey70")
lines(df$lambda, df$Education, col = "grey70")
lines(df$lambda, df$Own, col = "grey70")
lines(df$lambda, df$Married, col = "grey70")
lines(df$lambda, df$RegionEast, col = "grey70")
lines(df$lambda, df$RegionWest, col = "grey70")
lines(df$lambda, df$RegionSouth, col = "grey70")

legend("topright",
       legend = c("Income", "Limit", "Rating", "Student"),
       col = c("black", "red", "blue", "goldenrod2"),
       lty = c(1, 2, 3, 4),
       lwd = 2,
       bty = "n",
       cex = 1.2)

```

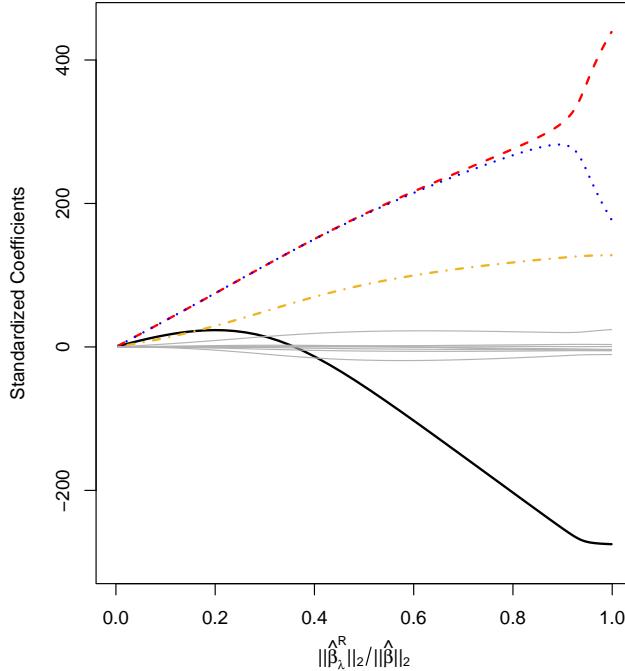


The ridge paths illustrate how each standardized coefficient shrinks toward zero as λ grows, although Rating and Income also increase partially before asymptotically approaching zero. On the left side of the plot, the coefficients correspond closely to the OLS estimates. However, this graph only depicts the coefficients based on a full training set. Thus, we can not draw any conclusions about the optimal value of λ and the coefficients. To make this possible, we would need to calculate the MSPE for multiple values of λ using 10-fold cross-validation and choose the minimum MSPE.

6.2 Norm-Ratio Plot

This plot presents the same coefficient paths as before but uses the ratio of the ridge to OLS ℓ_2 norms on the x-axis instead of λ . This ratio provides a standardized measure of shrinkage that ranges from 1 (no regularization, same as OLS) to 0 (maximum shrinkage). By expressing shrinkage as a relative measure rather than a raw penalty value, this visualization makes it easier to compare the degree of regularization across different models or datasets.

```
df_ratio <- df[order(df$norm_ratio), ]  
  
plot(df_ratio$norm_ratio, df_ratio$Income, type = "l", lwd = 2,  
      xlim = c(0, 1), ylim = c(-300, 450), col = "black",  
      xlab = expression("||" * hat(beta)[lambda]^R * "||"[2] / "||" * hat(beta) * "||"[2]),  
      ylab = "Standardized Coefficients")  
  
lines(df_ratio$norm_ratio, df_ratio$Limit, col = "red", lty = 2, lwd = 2)  
lines(df_ratio$norm_ratio, df_ratio$Rating, col = "blue", lty = 3, lwd = 2)  
lines(df_ratio$norm_ratio, df_ratio$Student, col = "goldenrod2", lty = 4, lwd = 2)  
  
lines(df_ratio$norm_ratio, df_ratio$Cards, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$Age, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$Education, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$Own, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$Married, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$RegionEast, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$RegionWest, col = "grey70")  
lines(df_ratio$norm_ratio, df_ratio$RegionSouth, col = "grey70")
```

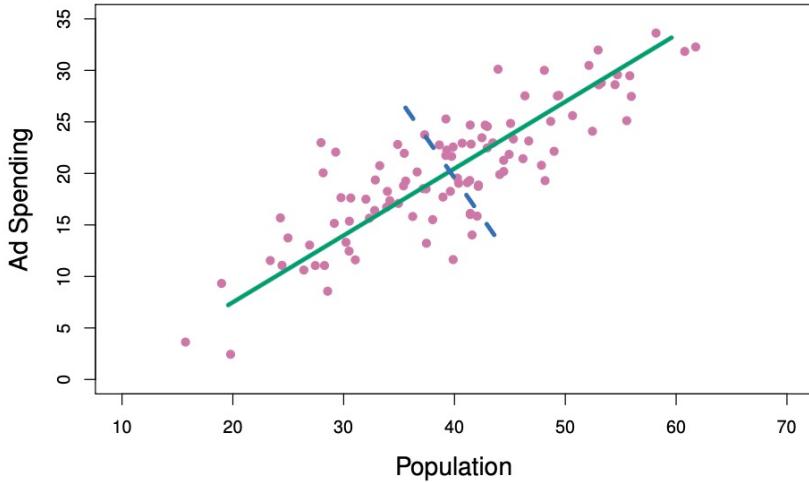


This plot visualizes how the standardized ridge coefficients change relative to the ratio of shrinkage (the ratio of the ridge to OLS ℓ_2 norms). On the right-hand side of the plot, the coefficients correspond to the OLS estimates when λ is zero. Thus, the ratio is equivalent to one. Moving from right to left, the magnitude of the ℓ_2 norm of the ridge regression declines, which decreases the ratio from one to zero. Again, we can not draw any conclusions about the optimal value of λ and the coefficients.

Exercise 2 - Principal Component Analysis

Exercise 2 is split in three parts:

1. Explain in detail how the two lines in Figure 6.14 are derived.
2. Why are they orthogonal to each other?
3. What do they mean?



1. Overview

The two predictors population size ($X_1 = \text{pop}$) and ad spending ($X_2 = \text{ad}$) for 100 different cities are shown as purple circles. The green solid line indicates the first principal component PC_1 , and the blue dashed line indicates the second principal component PC_2 derived from the data.

Principal-Component-Analysis (PCA) is a dimension reduction technique to reduce the predictor space $P = \{X_1, X_2, \dots, X_p\}$ to M principal components (PC), denoted as Z_1, Z_2, \dots, Z_M which are linear combinations of the original predictors X_1, X_2, \dots, X_p . The overarching topic of this exercise is to understand how PCA uses the variables pop and ad to constructs those components.

2. Derivation of PC_1

In our case, we can, at most, derive two principal components as we have two variables. Our PCs take the form $PC_i = Z_i = a_i \cdot (\text{pop} - \bar{\text{pop}}) + b_i \cdot (\text{ad} - \bar{\text{ad}})$ with a_i, b_i being the *principal component loadings*.

To derive the first principal component, we want to find a **linear combination of pop and ad** that captures the **maximum variance** of those input variables. This is equivalent to minimizing the sum of the squared perpendicular distances between each data point and the line constructed by the linear combination, i.e. the principal component.

To proceed, it is vital to restrict a and b because they would otherwise blow up the variance arbitrarily. Therefore, we can only consider linear combination of the form $a^2 + b^2 = 1$.

We are now maximizing this variance using the Lagrange method:

$$\max_{\text{pop}, \text{ad}} \text{Var}(a \cdot (\text{pop} - \bar{\text{pop}}) + b \cdot (\text{ad} - \bar{\text{ad}})) \quad \text{s.t.} \quad a^2 + b^2 = 1$$

For two general predictors X_1 and X_2 , the Lagrangian function thus results in:

$$(*) \quad \mathcal{L}(a, b, \lambda) = a^2 \cdot \text{Var}(X_1) + b^2 \cdot \text{Var}(X_2) + 2ab \cdot \text{Cov}(X_1, X_2) + \lambda(a^2 + b^2 - 1)$$

For the case that X_1 and X_2 are **standardized variables**, there is a simple, analytical solution to the maximization problem. As the variance (and thus also the standard deviation) equals one, we can drop the variance term and set the covariance equal to the correlation. Hence, the equation simplifies as follows:

$$(**) \quad \mathcal{L}(a, b, \lambda) = a^2 + b^2 + 2ab \cdot \rho_{X_1, X_2} + \lambda(a^2 + b^2 - 1)$$

To continue, we are defining the three conditions for the maximization problem. First, we take the two partial derivatives of the Lagrangian with respect to a and b and then include our constraint as the third equation:

$$\begin{aligned} (1) \quad \frac{\partial \mathcal{L}}{\partial a} &= 2a + 2b \cdot \rho_{X_1, X_2} + \lambda \cdot 2a \stackrel{!}{=} 0 \\ (2) \quad \frac{\partial \mathcal{L}}{\partial b} &= 2b + 2a \cdot \rho_{X_1, X_2} + \lambda \cdot 2b \stackrel{!}{=} 0 \\ (3) \quad a^2 + b^2 &= 1 \end{aligned}$$

Solving this system of equations, yields the first principal component:

$$PC_1 = \frac{1}{\sqrt{2}}X_1 + \frac{1}{\sqrt{2}}X_2$$

However, this derivation assumes standardization of variables which is not the case in the given example (pop and ad are solely mean adjusted!). Thus, while the initial equation (*) still holds, it does no longer simplify to (**) as the variance is not necessarily equal to one. The idea of the maximization problem is still the same and results in the following:

$$\max_{X_1, X_2} \text{Var}(aX_1 + bX_2) = \max a^2\sigma_{X_1}^2 + b^2\sigma_{X_2}^2 + 2ab \cdot \sigma_{X_1, X_2} \quad \text{s.t.} \quad a^2 + b^2 = 1$$

The corresponding Lagrangian function is:

$$\mathcal{L}(a, b, \lambda) = a^2\sigma_{X_1}^2 + b^2\sigma_{X_2}^2 + 2ab \cdot \sigma_{X_1, X_2} + \lambda(a^2 + b^2 - 1)$$

This gives us the following three conditions:

$$\begin{aligned} (1) \quad \frac{\partial \mathcal{L}}{\partial a} &= 2a \cdot \sigma_{X_1}^2 + 2b \cdot \sigma_{X_1, X_2} + \lambda \cdot 2a \stackrel{!}{=} 0 \\ (2) \quad \frac{\partial \mathcal{L}}{\partial b} &= 2b \cdot \sigma_{X_2}^2 + 2a \cdot \sigma_{X_1, X_2} + \lambda \cdot 2b \stackrel{!}{=} 0 \\ (3) \quad a^2 + b^2 &= 1 \end{aligned}$$

Solving this set of equations for a and b is not possible as we don't have access to the data points used. The book gives us the loadings $a = 0.839$ and $b = 0.544$, resulting in the equation for the first principal component that captures the most variance:

$$PC_1 = Z_1 = 0.839 \times (\text{pop} - \overline{\text{pop}}) + 0.544 \times (\text{ad} - \overline{\text{ad}})$$

3. Orthogonality and derivation of PC_2

As each principal component PC_i **explains additional variance** of the initial predictors, PC_2 explains the complimentary variance in the data that is not captured by PC_1 . Hence PC_2 has to be uncorrelated with PC_1 . We can thus derive PC_2 by writing both principal components in their vector form with $PC_1 = \begin{bmatrix} a_1 \\ b_1 \end{bmatrix}$, $PC_2 = \begin{bmatrix} a_2 \\ b_2 \end{bmatrix}$ and their loadings a_1, b_1 , and a_2, b_2 respectively.

Two principal components are **uncorrelated** when the dot product of their vectors is equal to zero. This is equivalent with **the two principal component lines being orthogonal to each other**.

$$\begin{bmatrix} a_1 \\ b_1 \end{bmatrix} \cdot \begin{bmatrix} a_2 \\ b_2 \end{bmatrix} = 0$$

We can apply the dot product, substitute the loadings of PC_1 and solve for a_2 and b_2 :

$$a_1 \cdot a_2 + b_1 \cdot b_2 = 0$$

$$0.839 \cdot a_2 + 0.544 \cdot b_2 = 0$$

$$\Rightarrow (a_2, b_2) = (0.544, -0.839)$$

The resulting loadings can then be plugged in to yield PC_2 :

$$PC_2 = Z_2 = 0.544 \times (\text{pop} - \overline{\text{pop}}) - 0.839 \times (\text{ad} - \overline{\text{ad}})$$

4. Interpretation and summary

With the previous calculations, we found the linear combinations of predictors that explain the highest variance in the predictors.

The first principal component, depicted as the green line in figure 6.14 is the component that captures the most variance in the scatter plot of X_1 and X_2 . The green line shows the direction in which the data varies the most as it represents the main axis of spread. Projecting the data points onto the green line provides the best one-dimensional representation of the two-dimensional data. In doing so, we are geometrically minimizing or compressing the dimensionality from two to one dimensions.

The blue line, on the other hand, indicates the direction of the remaining variation after removing the part that has already been explained by the first principal component PC_1 . As per the definition of the uncorrelated relationship of PC_1 to PC_2 , their vectors must be orthogonal to each other, with PC_2 representing the direction of second highest variation in the data.

