```
library(latex2exp)
library(tidyverse)
## -- Attaching packages ------ tidyverse 1.3.1 --
## v ggplot2 3.3.6
                    v purrr
                              0.3.4
## v tibble 3.1.7
                   v dplyr
                              1.0.9
## v tidyr 1.2.0 v stringr 1.4.0
## v readr 2.1.2
                   v forcats 0.5.1
## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()
                masks stats::lag()
library(rstan)
## Loading required package: StanHeaders
## rstan (Version 2.21.5, GitRev: 2e1f913d3ca3)
## For execution on a local, multicore CPU with excess RAM we recommend calling
## options(mc.cores = parallel::detectCores()).
## To avoid recompilation of unchanged Stan programs, we recommend calling
## rstan_options(auto_write = TRUE)
##
## Attaching package: 'rstan'
## The following object is masked from 'package:tidyr':
##
##
      extract
library(doParallel)
## Loading required package: foreach
##
## Attaching package: 'foreach'
## The following objects are masked from 'package:purrr':
##
      accumulate, when
## Loading required package: iterators
## Loading required package: parallel
registerDoParallel()
rstan_options(auto_write=TRUE)
options(mc.cores=parallel::detectCores())
tumor_experiments <- read.csv("../01_data/data.csv")</pre>
tumor_experiments$percent = tumor_experiments$tumors / tumor_experiments$n
# y <- tumor_experiments$tumors</pre>
# n <- tumor_experiments$n
2,1,5,2,5,3,2,7,7,3,3,2,9,10,4,4,4,4,4,4,4,10,4,4,4,5,11,12,
       5,5,6,5,6,6,6,6,16,15,15,9,4)
```

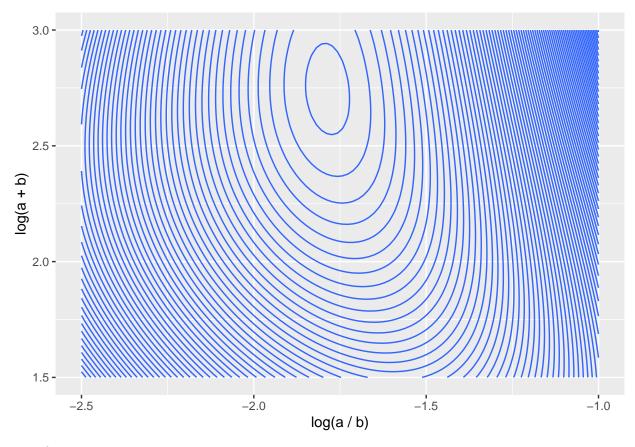
Marginal posterior distribution (5.8) and helper functions

$$P(\alpha, \beta | \mathbf{y}) \propto P(\alpha, \beta) \prod_{j} \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{\Gamma(\alpha + c_{j})\Gamma(\beta + n_{j} - c_{j})}{\Gamma(\alpha + \beta + n_{j})}$$

, We will compute this log-transformed.

Below are the point-wise estimates on the grid.

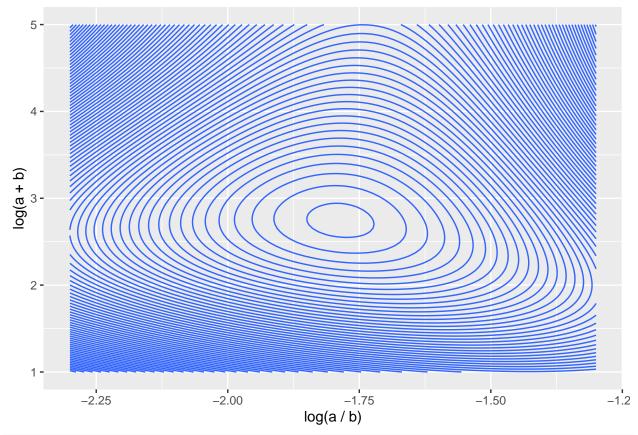
figure 5.2



### unused

```
# model.loglike = function(alpha, beta, cs, ns) {
      base = lgamma(alpha + beta) - lgamma(alpha) - lgamma(beta)
      data = mapply(a=alpha, b=beta, FUN=function(a, b) {
#
#
          sum(lgamma(a + cs) +
#
                  lgamma(b + ns - cs) -
#
                  lqamma(a + b + ns))
      })
#
#
      nrow(tumor_experiments) * base + data
# }
# xydensity = function(x, y, counts, totals, prior=model.prior, loglike=model.loglike) {
#
      expand.grid(x=x, y=y) \%
#
          mutate(alpha = par.alpha(x, y),
#
                 beta = par.beta(x, y)) \%%
#
          mutate(logPrior = log(prior(alpha, beta)),
#
                 logLike = loglike(alpha, beta, counts, totals),
#
                 rawLogPost = logPrior + logLike) %>%
#
          mutate(logJacobian = log(alpha) + log(beta),
#
                 logPost = rawLogPost + logJacobian)
# }
\# \ xcord \leftarrow seq(-2.5, -1, length.out = 100)
# ycord <- seq(1.5, 3, length.out = 100)
# # xcord < -seq(-2.3, -1.3, length.out = 100)
# # ycord <- seq(1, 5, length.out = 100)
# dens.points = xydensity(x=xcord,
```

## figure 5.3



```
cal.expvals = function(dens) {
  normPost = dens$logPost - max(dens$logPost)
  alpha = sum(dens$alpha * exp(normPost)) / sum(exp(normPost))
  beta = sum(dens$beta * exp(normPost)) / sum(exp(normPost))
  x = log(alpha / beta)
```

```
y = log(alpha + beta)
mean = alpha / (alpha + beta)
data.frame(alpha=alpha, beta=beta, x=x, y=y, mean=mean)
}
```

Expected values for  $\alpha, \beta, x, y$ , and  $\theta$  based on pointwise estimates of  $\alpha$  and  $\beta$ 

```
exp.vals <- cal.expvals(grid)</pre>
exp.alpha <- exp.vals$alpha</pre>
exp.beta <- exp.vals$beta
exp.vals
        alpha
                 beta
## 1 2.402568 14.3195 -1.785085 2.81673 0.1436764
# theta.post.point = data.frame(Theta=seq(0,1,0.001))
# theta.post.point$PostDens = dbeta(theta.post.point$Theta, exp.alpha, exp.beta)
# ggplot(theta.post.point) + aes(x=Theta, y=PostDens) + geom_line()
\# model.map = optim(c(1, 1), function(pars) {
#
      alpha = pars[1]
#
      beta = pars[2]
      log(model.prior(alpha, beta)) + model.loglike(alpha, beta, y, n)
# }, control=list(fnscale=-1))
# model.map
```

#### Stan

Fit the model

```
writeLines(readLines("ratmodel.stan"))
## data {
```

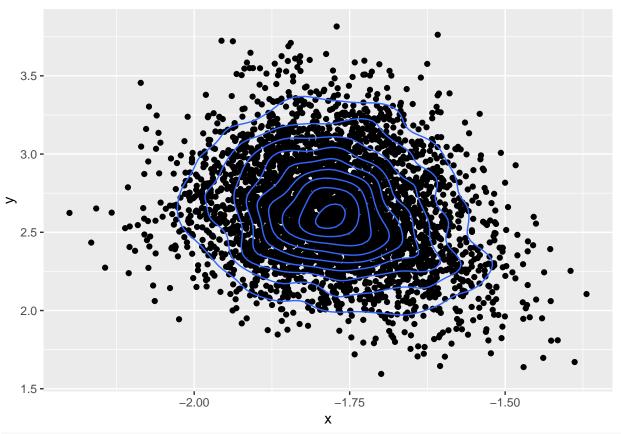
```
##
       int<lower=0> J;
##
       int<lower=0> n[J];
##
       int<lower=0> y[J];
## }
## parameters {
##
       real<lower=0,upper=1> phi;
##
       real<lower=0.1> lambda;
##
       real<lower=0,upper=1> theta[J];
## }
## transformed parameters {
       real<lower=0> alpha;
##
##
       real<lower=0> beta;
##
       alpha = lambda * phi;
       beta = lambda * (1 - phi);
##
## }
## model {
##
       phi ~ beta(1,1);
##
       lambda ~ pareto(0.1, 1.5);
       theta ~ beta(alpha, beta);
##
       y ~ binomial(n, theta);
```

```
## }
```

# ## [1] 4000

Simulated contour and points in figure 5.3(a and b)

```
a <- rat_sim$alpha
b <- rat_sim$beta
ggplot(data.frame(x=log(a/b), y=log(a+b), a=a, b=b)) +
   geom_point(aes(x=x, y=y)) +
   geom_density_2d(aes(x=x, y=y))</pre>
```



## # contour(kde2d(log(a/b), log(a+b)))

```
theta_sims = data.frame(alpha=rat_sim$alpha, beta=rat_sim$beta) %>%
    mutate(Theta=rbeta(n(), alpha+y[1], beta + n[1] - y[1]))
```