

```

library(latex2exp)
library(tidyverse)

## -- Attaching packages ----- tidyverse 1.3.1 --
## v ggplot2 3.3.6      v purrr  0.3.4
## v tibble  3.1.7      v dplyr  1.0.9
## v tidyr   1.2.0      v stringr 1.4.0
## v readr   2.1.2      v forcats 0.5.1

## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()    masks stats::lag()

library(rstan)

## Loading required package: StanHeaders
## rstan (Version 2.21.5, GitRev: 2e1f913d3ca3)
## For execution on a local, multicore CPU with excess RAM we recommend calling
## options(mc.cores = parallel::detectCores()).
## To avoid recompilation of unchanged Stan programs, we recommend calling
## rstan_options(auto_write = TRUE)

##
## Attaching package: 'rstan'

## The following object is masked from 'package:tidyr':
##
##      extract

library(doParallel)

## Loading required package: foreach
##
## Attaching package: 'foreach'

## The following objects are masked from 'package:purrr':
##
##      accumulate, when

## Loading required package: iterators
## Loading required package: parallel

registerDoParallel()

rstan_options(auto_write=TRUE)
options(mc.cores=parallel::detectCores())

tumor_experiments <- read.csv("../01_data/data.csv")

tumor_experiments$percent = tumor_experiments$tumors / tumor_experiments$n
tumor_experiments$y <- tumor_experiments$tumors
# n <- tumor_experiments$n
y <- c(0,0,0,0,0,0,0,0,0,0,0,0,0,1,1,1,1,1,1,1,2,2,2,2,2,2,2,2,
      2,1,5,2,5,3,2,7,7,3,3,2,9,10,4,4,4,4,4,4,10,4,4,4,5,11,12,
      5,5,6,5,6,6,6,6,16,15,15,9,4)

```

```
n <- c(20,20,20,20,20,20,20,19,19,19,19,18,18,17,20,20,20,20,19,19,18,18,25,24,
      23,20,20,20,20,20,20,10,49,19,46,27,17,49,47,20,20,13,48,50,20,20,20,20,
      20,20,20,48,19,19,19,22,46,49,20,20,23,19,22,20,20,20,52,46,47,24,14)
j <- length(y)
```

### Marginal posterior distribution (5.8) and helper functions

$$P(\alpha, \beta | y) \propto P(\alpha, \beta) \prod_j \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{\Gamma(\alpha + c_j)\Gamma(\beta + n_j - c_j)}{\Gamma(\alpha + \beta + n_j)}$$

, We will compute this log-transformed.

```
model.prior = prior = function(alpha, beta) {
  (alpha + beta) ^ (-5 / 2)
}
par.beta = function(x, y) {
  exp(y) / (exp(x) + 1)
}
par.alpha = function(x, y) {
  exp(x) * par.beta(x, y)
}

marg_post = function(alpha, beta){
  ldens <- 0
  for (i in 1:length(y)) ldens <- ldens +
    lgamma(alpha + beta) + lgamma(alpha + y[i]) + lgamma(beta + n[i] - y[i]) -
    (lgamma(alpha) + lgamma(beta) + lgamma(alpha + beta + n[i]))
  ldens + log(alpha*beta) + log(alpha+beta)*(-5/2) }
}
```

figure 5.2

Here we compute the values of the log-transformed posterior density on the grid. Following the example of Gelman BA, we also subtract the maximum value from every point on the grid and exponentiate so that all values lie between 0 and 1. We then show contours at 0.1, 0.2 ... 1 times the density at the mode.

```
xmin <- -2.5
xmax = -1
ymin<- 1.5
ymax <- 3
xcord <- seq(xmin, xmax, length.out = 100)
ycord <- seq(ymin, ymax, length.out = 100)

grid <- expand.grid(x=xcord, y=ycord) %>%
  mutate(a=par.alpha(x,y),
         b=par.beta(x,y)) %>% mutate(z=marg_post(a,b)) %>% mutate(z_rescale = exp(z - max(z)))

ggplot(grid, aes(x=x, y=y)) +
  geom_contour(aes(z=z_rescale), binwidth=0.05) +
  xlab("log(a / b)") +
  ylab("log(a + b)") + metR::geom_text_contour(aes(z=z_rescale)) + theme(aspect.ratio=1) + xlim(xmin
```

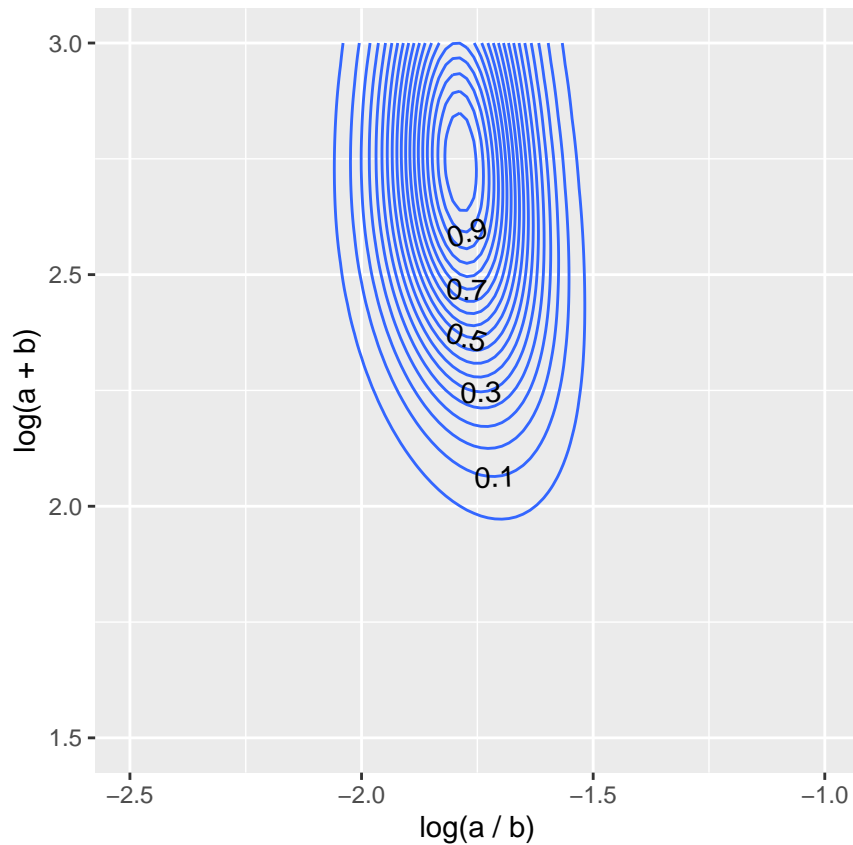


figure 5.3a

The same explanation for figure 5.2 holds but here we zoom in on the density by choosing a narrower grid.

```
xmin <- -2.3
xmax <- -1.3
ymin <- 1
ymax <- 5
xcord <- seq(xmin, xmax, length.out = 100)
ycord <- seq(ymin, ymax, length.out = 100)

grid <- expand.grid(x=xcord, y=ycord) %>%
  mutate(alpha=par.alpha(x,y),
           beta=par.beta(x,y)) %>% mutate(z=marg_post(alpha,beta)) %>% mutate(z_rescale = exp(z - max(z)))

ggplot(grid, aes(x=x, y=y)) +
  geom_contour(aes(z=z_rescale), binwidth=0.05) +
  xlab("log(a / b)") +
  ylab("log(a + b)") + metR::geom_text_contour(aes(z=z_rescale)) + theme(aspect.ratio=1) + xlim(xmin
```

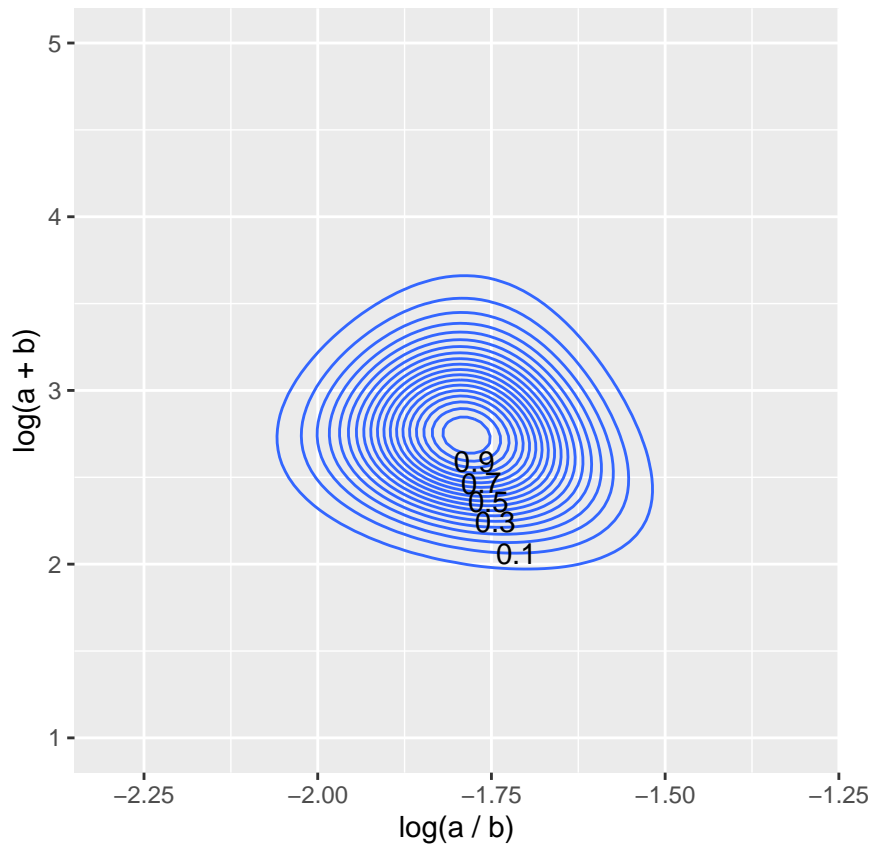


figure 5.3b

Below, we sample our log-transformed  $\alpha$ 's and  $\beta$ 's from the normalized grid. We then plot these samples in a similar fashion to figure 5.3a

```
nsamp <- 1000
p <- grid$z_rescale

samp_indices <- sample(length(p), size = nsamp,
                        replace = T, prob = p/sum(p))

samp_x <- grid$x[samp_indices[1:nsamp]]
samp_y <- grid$y[samp_indices[1:nsamp]]

samp_x <- samp_x + runif(nsamp, (xcord[1] - xcord[2])/2, (xcord[2] - xcord[1])/2)
samp_y <- samp_y + runif(nsamp, (ycord[1] - ycord[2])/2, (ycord[2] - ycord[1])/2)
samps <- tibble(ind = 1:nsamp, x = samp_x, y = samp_y)

sam <- ggplot(data = samps) +
  geom_point(aes(x, y), color = 'blue') +
  labs(title = 'Posterior draws', x = "log(a / b)", y = "log(a + b)") +
  theme(aspect.ratio=1) + xlim(xmin, xmax) + ylim(ymin, ymax)
sam
```

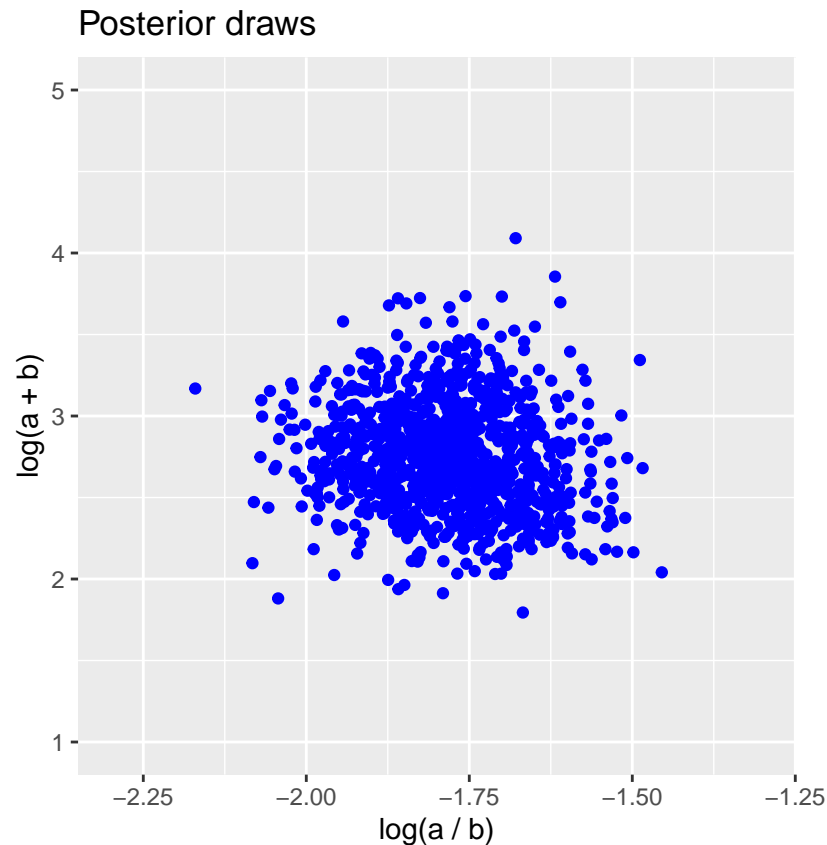


figure 5.4

Here, we sample each  $\theta_j$  from their respective posterior distribution and compute the 95% confidence interval for these  $\theta$  using the simulated  $\alpha$ 's and  $\beta$ 's from above. We then plot these with the medians shown as points. We also apply a small “jitter” as recommended in Gelman BA to all the x coordinates to reduce overlap.

```
samps <- samps %>% mutate(a=par.alpha(x,y),
                          b=par.beta(x,y))

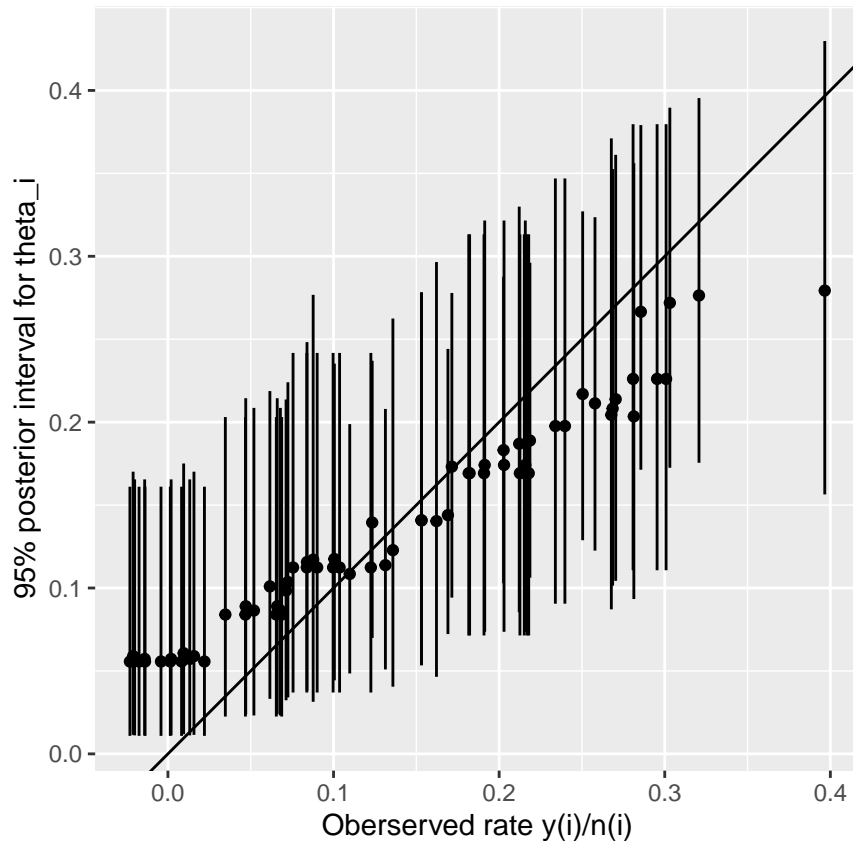
a <- samps$a
b <- samps$b

qh <- function(q, n, y) colMeans(mapply(function(q, n, y, a, b)
  mapply(qbeta, q, a + y, n - y + b), q, n, y, MoreArgs = list(samps$a, samps$b)))

qq_hier <- data.frame(id=1:length(n), n=n, y=y,
                      low=qh(0.025, n, y), high=qh(0.975, n, y), med=qh(0.5, n, y), perc=(y/n))

qq_hier$perc <- jitter(qq_hier$perc, amount=0.025)

ggplot(qq_hier, aes(x=perc, y=med, ymin=low, ymax=high)) +
  geom_point() + geom_abline(intercept=0, slope=1) + geom_linerange() +
  theme(aspect.ratio=1) + labs(x = "Observed rate y(i)/n(i)", y = "95% posterior interval for theta_")
```



```
# qq_hier[order(qq_hier$perc, decreasing = TRUE),]

# cal.expvals = function(dens) {
#   normPost = dens$logPost - max(dens$logPost)
#   alpha = sum(dens$alpha * exp(normPost)) / sum(exp(normPost))
#   beta = sum(dens$beta * exp(normPost)) / sum(exp(normPost))
#   x = log(alpha / beta)
#   y = log(alpha + beta)
#   mean = alpha / (alpha + beta)
#   data.frame(alpha=alpha, beta=beta, x=x, y=y, mean=mean)
# }
```

Expected values for  $\alpha$ ,  $\beta$ ,  $x$ ,  $y$ , and  $\theta$  based on pointwise estimates of  $\alpha$  and  $\beta$

```
# grid
#
# exp.vals <- cal.expvals(grid)
#
# exp.alpha <- exp.vals$alpha
# exp.beta <- exp.vals$beta
#
# exp.vals
```

## Stan

I am no longer using stan to generate the figures from section 5.3 so for now we can ignore the code below.

Fit the model

```

writeLines(readLines("ratmodel-old.stan"))

## data {
##   int<lower=0> J;
##   int<lower=0> n[J];
##   int<lower=0> y[J];
## }
## parameters {
##   real<lower=0,upper=1> phi;
##   real<lower=0.1> lambda;
##   real<lower=0,upper=1> theta[J];
## }
## transformed parameters {
##   real<lower=0> alpha;
##   real<lower=0> beta;
##   alpha = lambda * phi;
##   beta = lambda * (1 - phi);
## }
## model {
##   phi ~ beta(1,1);
##   lambda ~ pareto(0.1, 1.5);
##   theta ~ beta(alpha, beta);
##   y ~ binomial(n, theta);
## }

rat_fit = stan(file="ratmodel-old.stan", data=list(J=j, y=y, n=n),
               iter=2000, chains=4, control=list(adapt_delta=0.99))

# pairs(rat_fit, pars=c("alpha", "beta", "lp_"))
rat_sim = rstan::extract(rat_fit, permuted=TRUE)

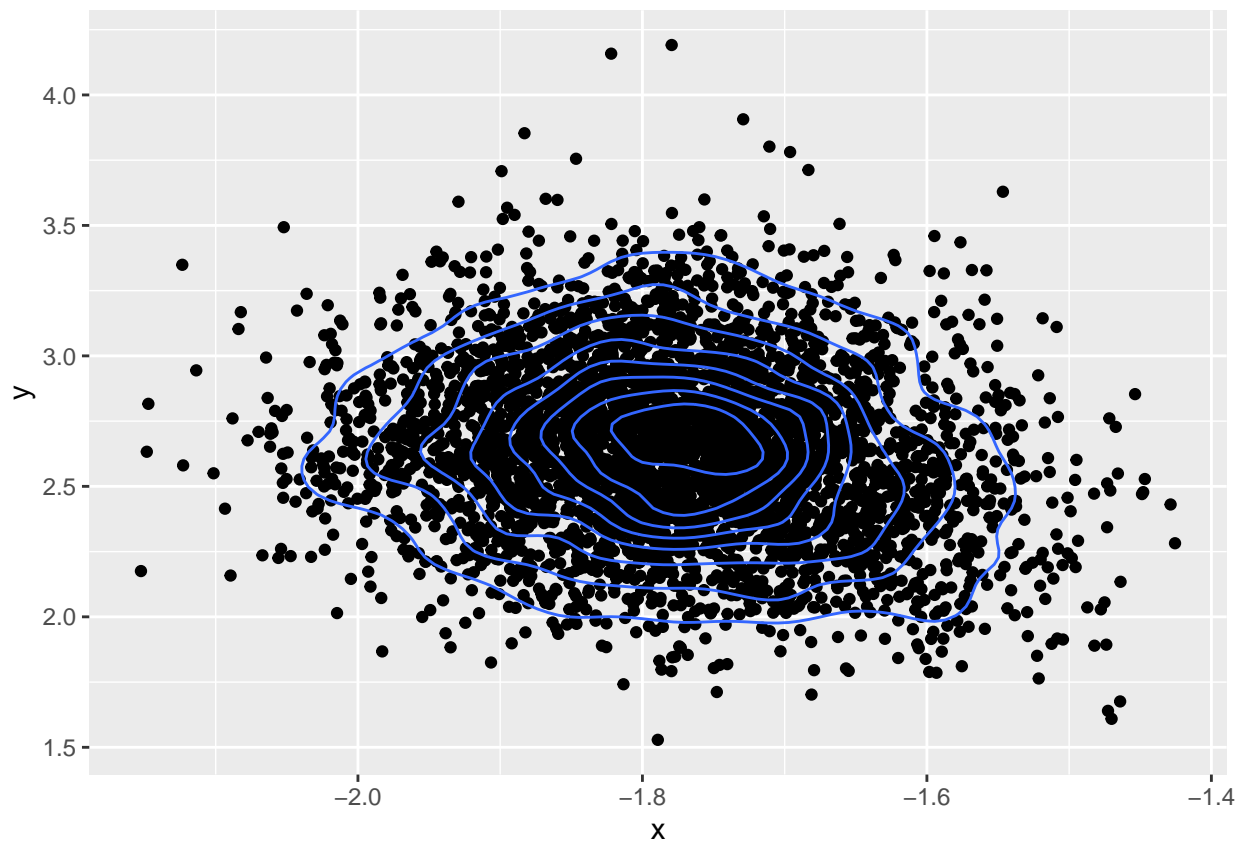
n_sims = length(rat_sim$lp_)
n_sims

## [1] 4000

Simulated contour and points in figure 5.3(a and b)

a <- rat_sim$alpha
b <- rat_sim$beta
ggplot(data.frame(x=log(a/b), y=log(a+b), a=a, b=b)) +
  geom_point(aes(x=x, y=y)) +
  geom_density_2d(aes(x=x, y=y))

```



```
# contour(kde2d(log(a/b), log(a+b)))
```

```
theta_sims = data.frame(alpha=rat_sim$alpha, beta=rat_sim$beta) %>%  
  mutate(Theta=rbeta(n(), alpha+y[1], beta + n[1] - y[1]))
```