

## 17.1 Lindeberg's Central Limit Theorem

We have the following set-up:

- A triangular array of random variables  $\{X_{ni}\}$ ,  $1 \leq i \leq k_n$ , such that  $\mathbb{E}X_{ni} = 0$ ,  $\text{Var } X_{ni} = \sigma_{ni}^2 < \infty$ .
- $S_n := \sum_{i=1}^{k_n} X_{ni}$ .
- $s_n^2 := \sum_{i=1}^{k_n} \sigma_{ni}^2$ .
- The random variables in each row are independent of each other, i.e. for every  $n$ ,  $\{X_{ni}\}_{i=1}^{k_n}$  is independent.

Lindeberg's version of the Central Limit Theorem is as follows:

**Theorem 17.1 (Lindeberg's CLT)** *Suppose that the Lindeburg condition holds: i.e. for every  $\varepsilon > 0$ ,*

$$\lim_{n \rightarrow \infty} \frac{1}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|X_{ni}| > \varepsilon s_n\}} X_{ni}^2 dP = 0.$$

*Then for every  $x \in \mathbb{R}$ ,*

$$P \left\{ \frac{S_n}{s_n} \leq x \right\} \rightarrow \Phi(x).$$

To prove this theorem, we will introduce a few lemmas:

**Lemma 17.2** *Let  $\mathcal{C}_b^\infty = \{f : f : \mathbb{R} \rightarrow \mathbb{R}, f \text{ has bounded derivatives of all order}\}$ .*

*If  $F_n, F$  are distribution functions on  $\mathbb{R}$  such that*

$$\int_{-\infty}^{\infty} f dF_n \rightarrow \int_{-\infty}^{\infty} f dF$$

*for all  $f \in \mathcal{C}_b^\infty$ , then  $F_n \Rightarrow F$ , i.e.  $F_n$  converges weakly to  $F$ .*

We proved this lemma last lecture.

**Lemma 17.3** Given  $f \in \mathcal{C}_b^\infty$ , let

$$g(h) := \sup_{x \in \mathbb{R}} \left| f(x+h) - f(x) - hf'(x) - \frac{h^2}{2}f''(x) \right|.$$

Then there exists a constant  $k > 0$  (which depends only on  $f$ ) such that  $g(h) \leq \min(kh^3, kh^2)$ .

This lemma follows directly from Taylor's Theorem with remainder.

**Lemma 17.4** For any  $h_1$  and  $h_2$ ,

$$\left| f(x+h_1) - f(x+h_2) - f'(x)(h_1-h_2) + \frac{f''(x)}{2}(h_1^2-h_2^2) \right| \leq g(h_1) + g(h_2).$$

Let us now prove Lindeberg's CLT.

**Proof:** By Lemma 17.2, it is enough to show that for every  $f \in \mathcal{C}_b^\infty$ ,  $\mathbb{E}f\left(\frac{S_n}{s_n}\right) \rightarrow \mathbb{E}(f(Z))$ , where  $Z$  is a standard normal random variable.

**Central idea of the proof:** Replace  $X_{ni}$  in  $S_n$  one at a time with  $Z_{ni} \sim \mathcal{N}(0, \sigma_{ni}^2)$ .

Let  $T_{ni} := X_{n1} + \dots + X_{n(i-1)} + Z_{n(i+1)} + \dots + Z_{nk_n}$ . Then

$$S_n = T_{nk_n} + X_{nk_n}, \quad \sum_{i=1}^{k_n} Z_i = T_{n1} + Z_{n1}.$$

Note that  $Z := \frac{1}{s_n} \sum_{i=1}^{k_n} Z_i \sim \mathcal{N}(0, 1)$ . Hence,

$$\begin{aligned} \left| \mathbb{E}f\left(\frac{S_n}{s_n}\right) - \mathbb{E}f(Z) \right| &= \left| \sum_{i=1}^{k_n} \mathbb{E}f\left(\frac{T_{ni} + X_{ni}}{s_n}\right) - \mathbb{E}f\left(\frac{T_{ni} + Z_{ni}}{s_n}\right) \right| \\ &\leq \sum_{i=1}^{k_n} \left| \mathbb{E}f\left(\frac{T_{ni} + X_{ni}}{s_n}\right) - \mathbb{E}f\left(\frac{T_{ni} + Z_{ni}}{s_n}\right) \right|. \end{aligned}$$

By independence, for any  $n$  and  $i$ , we have

$$\begin{aligned} \mathbb{E}f'(T_{ni})(X_{ni} - Z_{ni}) &= 0, \\ \mathbb{E}f''(T_{ni})(X_{ni}^2 - Z_{ni}^2) &= 0. \end{aligned}$$

Add these terms in and using the bound in Lemma 17.4, we have

$$\begin{aligned}
\left| \mathbb{E}f\left(\frac{S_n}{s_n}\right) - \mathbb{E}f(Z) \right| &\leq \sum_{i=1}^{k_n} \left| \mathbb{E}f\left(\frac{T_{ni} + X_{ni}}{s_n}\right) - \mathbb{E}f\left(\frac{T_{ni} + Z_{ni}}{s_n}\right) - \mathbb{E}f'\left(\frac{T_{ni}}{s_n}\right) \left(\frac{X_{ni}}{s_n} - \frac{Z_{ni}}{s_n}\right) \right. \\
&\quad \left. + \mathbb{E}\frac{1}{2}f''\left(\frac{T_{ni}}{s_n}\right) \left[ \left(\frac{X_{ni}}{s_n}\right)^2 - \left(\frac{Z_{ni}}{s_n}\right)^2 \right] \right| \\
&\leq \sum_{i=1}^{k_n} \mathbb{E}g\left(\frac{X_{ni}}{s_n}\right) + \sum_{i=1}^{k_n} \mathbb{E}g\left(\frac{Z_{ni}}{s_n}\right) \\
&=: I + II.
\end{aligned}$$

First consider  $I$ . We break the  $i$ th integral into  $\{\omega : \frac{|X_{ni}|}{s_n} \leq \varepsilon\}$  and its complement, and we use  $g(h) \leq kh^3$  on the part and  $g(h) \leq kh^2$  on the second:

$$\begin{aligned}
I &\leq k \sum_{i=1}^{k_n} \int_{\{|X_{ni}|/s_n \leq \varepsilon\}} \frac{|X_{ni}|^3}{s_n^3} dP + k \sum_{i=1}^{k_n} \int_{\{|X_{ni}|/s_n > \varepsilon\}} \frac{|X_{ni}|^2}{s_n^2} dP \\
&\leq k\varepsilon + \frac{k}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|X_{ni}|/s_n > \varepsilon\}} |X_{ni}|^2 dP.
\end{aligned}$$

By the Lindeberg condition, for fixed  $\varepsilon$ , the second term goes to 0 as  $n \rightarrow \infty$ . Hence,  $I$  can be bounded above by a constant multiple of  $\varepsilon$ , with the constant only depending of  $f$ .

Next consider  $II$ . Splitting up the integral in the same way as  $I$ , we obtain

$$\begin{aligned}
II &\leq k\varepsilon + \frac{k}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|Z_{ni}|/s_n > \varepsilon\}} |Z_{ni}|^2 dP \\
&\leq k\varepsilon + \frac{k}{\varepsilon s_n^3} \sum_{i=1}^{k_n} \mathbb{E}|Z_{ni}|^3 \\
&= k\varepsilon + \frac{k\mathbb{E}|Z|^3}{\varepsilon s_n^3} \sum_{i=1}^{k_n} \sigma_{ni}^3.
\end{aligned}$$

Note that for any  $\varepsilon' > 0$ , we have the bound

$$\frac{\sigma_{ni}^2}{s_n^2} \leq (\varepsilon')^2 + \frac{1}{s_n^2} \int_{\{|X_{ni}|/s_n > \varepsilon'\}} |X_{ni}|^2 dP \rightarrow (\varepsilon')^2$$

as  $n \rightarrow \infty$ . Since  $\varepsilon'$  is arbitrary, it means that  $\max \frac{\sigma_{ni}}{s_n} \rightarrow 0$  as  $n \rightarrow \infty$ . Hence,

$$\begin{aligned}
II &\leq k\varepsilon + \frac{k\mathbb{E}|Z|^3}{\varepsilon s_n^3} \sum_{i=1}^{k_n} \sigma_{ni}^3 \\
&\leq k\varepsilon + \frac{k\mathbb{E}|Z|^3}{\varepsilon s_n^3} (\max \sigma_{ni}) \sum_{i=1}^{k_n} \sigma_{ni}^2 \\
&= k\varepsilon + \frac{k\mathbb{E}|Z|^3}{\varepsilon s_n} (\max \sigma_{ni}) \rightarrow 0
\end{aligned}$$

as  $n \rightarrow \infty$ . ■

**Remark:** This is a nice example of a coupling proof, or a probabilistic proof.

## 17.2 CLT Variants

There is a converse of sorts for the theorem, known as the Lindeberg-Feller Theorem. (See Feller Vol II.)

**Theorem 17.5 (Lyapunov)** *Let  $X_{ni}$  be a triangular array such that the  $X_{ni}$ 's have mean 0,  $\mathbb{E}|X_{ni}|^{2+\delta} < \infty$  for some  $\delta > 0$ .*

*If Lyapunov's condition holds:*

$$\lim_{n \rightarrow \infty} \frac{1}{s_n^{2+\delta}} \sum_{i=1}^{k_n} \mathbb{E}|X_{ni}|^{2+\delta} = 0,$$

*then  $P\left\{\frac{S_n}{s_n} \leq x\right\} \rightarrow \Phi(x)$ .*

**Proof:** We just need to show that the Lindeberg condition holds. Given  $\varepsilon > 0$ ,

$$\begin{aligned}
\frac{1}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|X_{ni}|/s_n > \varepsilon\}} |X_{ni}|^2 dP &= \frac{1}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|X_{ni}|^\delta > \varepsilon^\delta s_n^\delta\}} |X_{ni}|^2 dP \\
&\leq \frac{1}{\varepsilon^\delta s_n^{2+\delta}} \sum_{i=1}^{k_n} \mathbb{E}|X_{ni}|^{2+\delta} \rightarrow 0.
\end{aligned}$$
■

**Theorem 17.6 (“Ordinary” CLT)** *Suppose  $X_i$  are iid, mean 0, variance  $\sigma^2$ . Then  $P\left\{\frac{S_n}{\sigma\sqrt{n}} \leq x\right\} \rightarrow \Phi(x)$ .*

**Proof:** Again, we just need to show that the Lindeberg condition holds. Given  $\varepsilon > 0$ ,

$$\begin{aligned} \frac{1}{s_n^2} \sum_{i=1}^{k_n} \int_{\{|X_{ni}|/s_n > \varepsilon\}} |X_{ni}|^2 dP &= \frac{1}{n\sigma^2} \sum_{i=1}^n \int_{\{|X_{ni}| > \varepsilon\sigma\sqrt{n}\}} |X_{ni}|^2 dP \\ &= \frac{1}{n\sigma^2} \cdot n \int_{|X_1| > \varepsilon\sigma\sqrt{n}} |X_1|^2 dP \\ &\rightarrow 0 \end{aligned}$$

as  $n \rightarrow \infty$  by the Dominated Convergence Theorem. ■

**Example:** Say an urn has  $R$  red balls and  $B$  black balls, where  $R + B = N$ . Sample the balls without replacement, and let

$$X_i = \begin{cases} 0 & \text{if ball is red,} \\ 1 & \text{if ball is black.} \end{cases}$$

If  $\frac{n}{N} \rightarrow 0$ , then  $P\left\{\frac{S_n}{\sigma\sqrt{n}} \leq x\right\} \rightarrow \Phi(x)$  for some value of  $\sigma$ .

**General version (Hoeffding Combinatorial CLT):** Let  $M_n = (m^{ij})_{1 \leq i, j \leq n}$  be a square matrix of numbers. Let  $S_n = \sum_{i=1}^n m^{i\pi(i)}$ , where  $\pi$  is a random permutation (i.e. sum of a random diagonal). Under mild conditions on  $M_n$ , the CLT holds.

## 17.3 Bounds on CLT Approximation

Note that the proof for Lindeberg's CLT gives an actual bound for  $\left|\mathbb{E}f\left(\frac{S_n}{s_n}\right) - \mathbb{E}f(Z)\right|$ . The Berry-Esseen Theorem also gives such bounds:

**Theorem 17.7 (Berry-Esseen)** Say  $X_1, \dots, X_n$  have mean 0, variance  $\sigma_i^2$ , and  $\mathbb{E}|X_i|^3 = r_i$  are finite. Then

$$\sup_{-\infty < x < \infty} \left| P\left\{\frac{S_n}{s_n} \leq x\right\} - \Phi(x) \right| \leq \frac{0.78R_n}{s_n^3},$$

where  $R_n = \sum_{i=1}^n r_i$ .

If the  $X_i$ 's are iid, the RHS is  $\frac{0.78\mathbb{E}|X_1|^3}{\sigma^3\sqrt{n}}$ .

Alternatively, one can use Edgeworth expansions for better approximations than  $\Phi(x)$ , e.g.

$$P\left\{\frac{S_n}{s_n} \leq x\right\} = \Phi(x) + \frac{H_1(x)}{\sqrt{n}} + O\left(\frac{1}{n}\right).$$