

Universality of the Uniform.

Let R be a cont. strictly increasing CDF, then

$$X = R^{-1}(u) \sim R \text{ if } u \sim \text{Unif}(0,1).$$

Also, if $X \sim R$, then $R(X) \sim \text{Unif}(0,1)$

$$\hookrightarrow = R(R^{-1}(u))$$

Warning: Don't make this mistake.

$$R(u) = P(X \leq u) \quad \text{Doesn't mean } \cancel{P(X) = P(X \leq X)} = 1 \quad \text{Wrong.}$$

Example: $R(x) = 1 - e^{-x}$ for $x \geq 0$. $\Rightarrow R(x) = 1 - e^{-x}$

Inversing the Function :

$$u = 1 - e^{-x} \Rightarrow x = -\ln(1-u) \Rightarrow R^{-1}(u) = -\underbrace{\ln(1-u)}_{\downarrow} \sim R.$$

$1-u \sim \text{Unif}(0,1)$ if $u \sim \text{Unif}(0,1)$

Scaling:

If $u \sim \text{Unif}(0,1)$, then $\underbrace{a+bu}_{\downarrow}$ is Unif. on some interval.

↓
Linear transform

A non-linear transform on a Uniform \rightarrow Non-Uniform (e.g. u^2 is non-uniform)

Independence of Random Variables

Defn. X_1, \dots, X_n indep. if for all x_1, \dots, x_n

$$\underbrace{P(X_1 \leq x_1, \dots, X_n \leq x_n)}_{\text{Joint CDF}} = P(X_1 \leq x_1) \cdots P(X_n \leq x_n)$$

Discrete,

$$\underbrace{P(X_1 = x_1, \dots, X_n = x_n)}_{\text{Joint PMF}} = P(X_1 = x_1) \cdots P(X_n = x_n)$$

Note: Remember that for the indep. of 3 events you had to check

the triple intersection AND all the pairwise ones.

In that case, we had to write "finitely many" equations while here,

Since we say: For all $x_1, \dots, x_n \Rightarrow$ infinitely many equations.

Full independence means that knowing any collection of them tells you **Nothing** about the ones you don't know about.

$$P(X_1 = x_1, \dots, X_n = x_n) = \begin{cases} 1 & \text{if } X_1 = X_2 \\ 0 & \text{otherwise} \end{cases}$$

Example: $X_1, X_2 \sim \text{Bern}(\frac{1}{2})$ i.i.d ; $X_3 = \begin{cases} 1 & \text{if } X_1 = X_2 \\ 0 & \text{otherwise} \end{cases}$

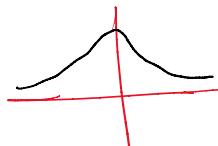
These 3 are pairwise indep, but not indep.

↓ ↓
If you know X_1 and X_2 , you know X_3

Knowing X_1 gives you nothing about X_2 or X_3 , etc

(Gaussian)

Normal Distribution:



Reason of importance:

Central Limit Theorem: Sum of a lot of i.i.d. R.V's looks normal.

Consider $Z \sim N(0, 1)$ has PDF: $f(z) = c \cdot e^{-\frac{z^2}{2}}$
↳ Normalizing Const.

$$\int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz = \text{Gaussian Integral} = \sqrt{2\pi} \Rightarrow f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

Expectation Value and Variance:

$Z \sim N(0, 1)$ Called the Standard Normal

mean Variance

$$\mathbb{E}(Z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} \underbrace{z}_{\text{odd}} \cdot \underbrace{e^{-z^2/2}}_{\text{even}} dz = 0$$

Integration by Parts

$$\mathbb{E}(Z^2) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} \underbrace{z^2}_{\text{even}} \cdot \underbrace{e^{-z^2/2}}_{\text{even}} dz = \frac{2}{\sqrt{2\pi}} \int_0^{+\infty} \underbrace{z \cdot z}_{u} \cdot \underbrace{e^{-z^2/2}}_{dv} dz$$

$$= \frac{2}{\sqrt{2\pi}} \left(\underbrace{(z \cdot e^{-z^2/2})}_{\downarrow} \Big|_0^{+\infty} - \int_0^{+\infty} e^{-z^2/2} dz \right) = \frac{2}{\sqrt{2\pi}} \times \frac{\sqrt{2\pi}}{2} = 1$$

$\frac{1}{2}$, Gaussian

$$\text{Var}(Z) = \mathbb{E}(Z^2) - \mathbb{E}^2(Z) = 1 - 0 = 1$$

Notation:

Φ is the Standard Normal CDF:

$$\Phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt$$

No analytical solution
Computer-calculated

$$\bar{\Phi}(z) = 1 - \Phi(z) \quad \text{by symmetry.}$$