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Stat 202C - Project 2

Metropolis-Hastings: Versus Accept-Reject and Utilized in Logistic Regression

Comparison of Metropolis-Hastings with Accept-Reject

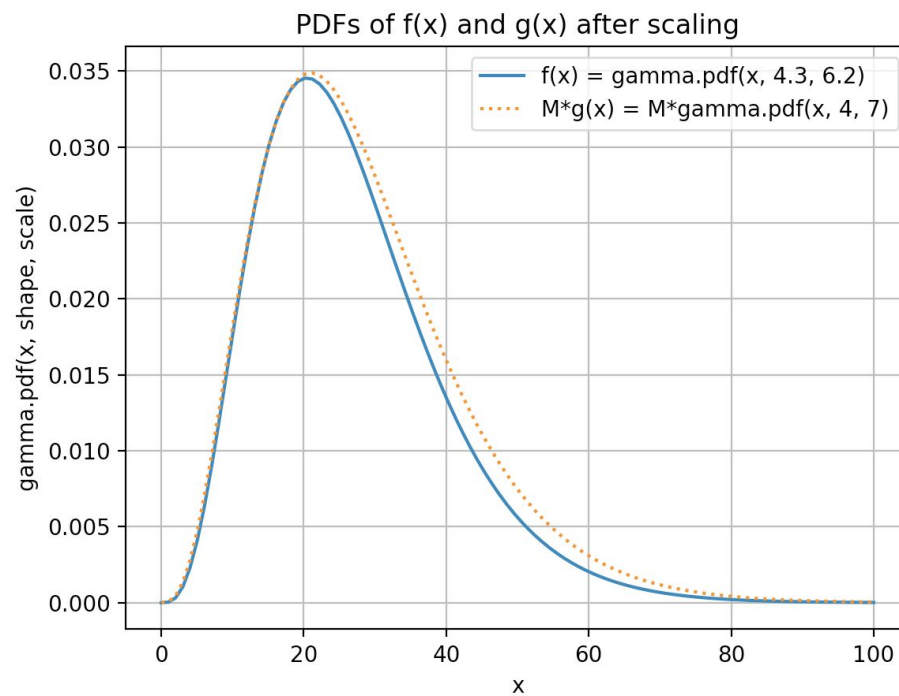
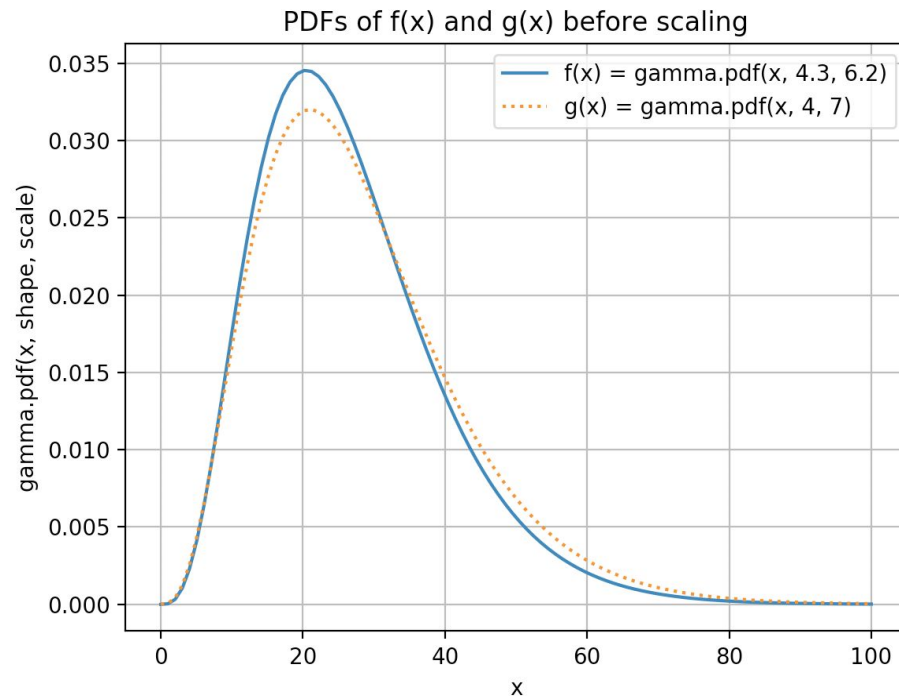
Problem 7.1: Here we will estimate the mean of a $\text{Gamma}(4.3, 6.2)$ random variable using the following algorithms and gamma distribution parameters: **a)** Accept-Reject with a $\text{Gamma}(4, 7)$ candidate, **b)** Metropolis-Hastings with a $\text{Gamma}(4, 7)$ candidate, and **c)** Metropolis-Hastings with a $\text{Gamma}(5, 6)$ candidate. For each algorithm and given Gamma parameters, we will monitor the convergence to the mean of the $\text{Gamma}(4.3, 6.2)$ random variable.

Problem 7.1a: In this subproblem we evaluate the Accept-Reject algorithm with a $\text{Gamma}(4, 7)$ candidate. The Accept-Reject algorithm is shown below, taken from *Monte Carlo Statistical Methods* by Robert and Casella.

Algorithm A.4 –Accept–Reject Method–

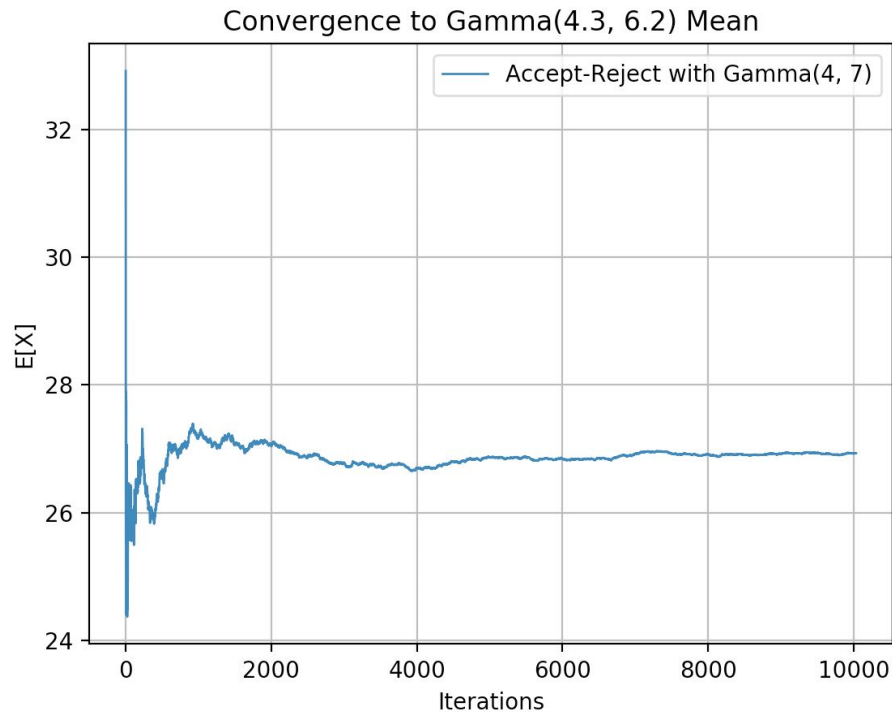
1. Generate $X \sim g$, $U \sim \mathcal{U}_{[0,1]}$;
 2. Accept $Y = X$ if $U \leq f(X)/Mg(X)$;
 3. Return to 1. otherwise.
- [A.4]

The two plots below show the target distribution $f(x)$ and trial distribution $g(x)$ before and after scaling $g(x)$ by the value M . M satisfies the inequality $\text{Gamma}(x; 4.3, 6.2) \leq M \cdot \text{Gamma}(x; 4, 7)$ and is the smallest value such that $f(x)/g(x) < M$. In this case, as we can evaluate a parameterized $\text{Gamma}(a, b)$ distribution using the python scipy.stats method `gamma.pdf(x, a, b)`, the maximized value $M = 1.0895$ was found using the scipy.optimize function `minimize`.



We can see the trial distribution $g(x)$ successfully encompasses or encloses the target distribution $f(x)$ after $g(x)$ is scaled by the optimized value M .

Below we plot the convergence of the Accept-Reject algorithm to the true estimated mean value of a Gamma(4.3, 6.2) random variable over 10,000 iterations. The mean of a Gamma distribution is $E[X] = k\theta$, where k and θ are the shape and scale parameters. Hence the true estimated mean value of a Gamma(4.3, 6.2) random variable is **26.66**. The final value for the Accept-Reject estimator with a Gamma(4, 7) candidate is **26.9331**.



Problem 7.1b, 7.1c: Now we evaluate the Metropolis-Hastings algorithm with two sets of Gamma parameters: the same Gamma(4, 7) candidate given to the Accept-Reject algorithm and a Gamma(5, 6) candidate. The Metropolis-Hastings algorithm is shown below, taken again from *Monte Carlo Statistical Methods* by Robert and Casella.

Algorithm A.24 –Metropolis–Hastings–

Given $x^{(t)}$,

1. Generate $Y_t \sim q(y|x^{(t)})$.
2. Take

$$X^{(t+1)} = \begin{cases} Y_t & \text{with probability } \rho(x^{(t)}, Y_t), \\ x^{(t)} & \text{with probability } 1 - \rho(x^{(t)}, Y_t), \end{cases}$$

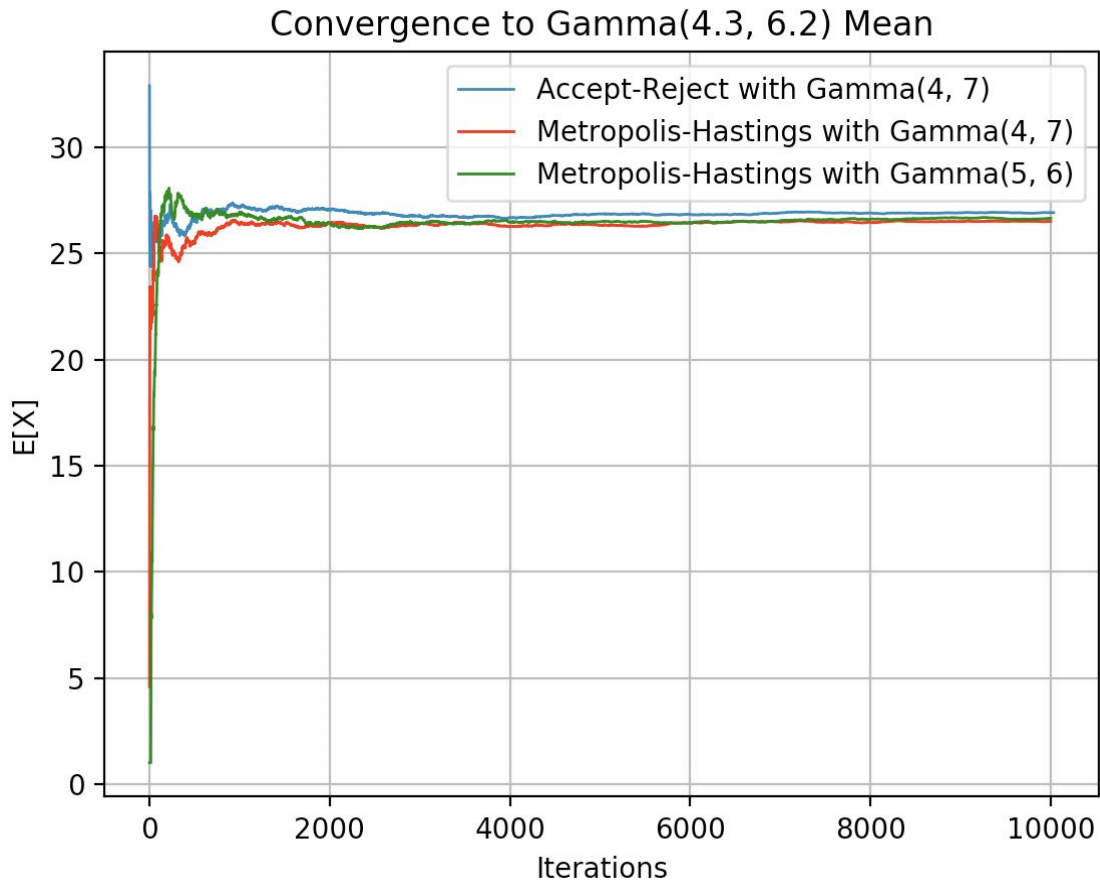
where

[A.24]

$$\rho(x, y) = \min \left\{ \frac{f(y)}{f(x)} \frac{q(x|y)}{q(y|x)}, 1 \right\}.$$

Recall the final value for the Accept-Reject estimator with a Gamma(4, 7) candidate is **26.9331**. The final value for the Metropolis-Hastings estimator with a Gamma(4, 7) candidate is **26.5114**. The final value for the Metropolis-Hastings estimator with a Gamma(5, 6) candidate is **26.6621**. Note that these final estimated values are after 10,000 iterations. With more iterations, they continually better approximate the true value **26.66**, with the Metropolis-Hastings estimator with a Gamma(5, 6) candidate generally producing the closest approximation.

Below we plot the convergence for each of the three algorithms and Gamma distributions: **a)** Accept-Reject with a Gamma(4, 7) candidate, **b)** Metropolis-Hastings with a Gamma(4, 7) candidate, and **c)** Metropolis-Hastings with a Gamma(5, 6) candidate.



As we can see, the Metropolis-Hastings algorithm with a Gamma(5, 6) candidate, resulting in a final estimate of **26.9331**, best approximated the true estimated mean value **26.66** of a Gamma(4.3, 6.2) random variable over the first 10,000 iterations.

Metropolis-Hastings for Logistic Regression

Problem 7.20: Here we use data from the Institute for Child Health Policy at the University of Florida to answer a question of interest: whether the status of the HMO affects the emergency room choice. An appropriate model is the logistic regression model,

$$\text{logit}(p_{ij}) = a + bx_i + cz_{ij}, \quad i = 1, \dots, k, \quad j = 1, \dots, n_i$$

where x_i is the HMO type, z_{ij} is the health status of the child, and p_{ij} is the probability of using an emergency room.

Problem 7.20a: First we verify the following likelihood function for the logistic regression model:

$$\prod_{i=1}^k \prod_{j=1}^{n_i} \left(\frac{\exp(a + bx_i + cz_{ij})}{1 + \exp(a + bx_i + cz_{ij})} \right)^{y_{ij}} \left(\frac{1}{1 + \exp(a + bx_i + cz_{ij})} \right)^{1-y_{ij}}$$

Sketch:

7.20a) Logistic model: $p(y=1) = \text{prob} \frac{\exp(a + bx_i + cz_{ij})}{1 + \exp(a + bx_i + cz_{ij})}$. $\text{logit}(p) = \log \frac{p}{1-p} = a + bx + cz$

$Y_{ij} \sim \text{Bernoulli}(p_{ij})$

logistic function: $\sigma(x) = \frac{e^x}{1 + e^x} = \frac{1}{1 + e^{-(a + bx + cz)}}$

times input x and outputs y as 0,1

$\text{logit}(\text{log odds}) = \text{yfp}(x) = \text{logit}(p(x)) = \log \frac{p(x)}{1-p(x)} = a + bx + cz_{ij}$

known as logit p_{ij} equivalent to linear regression expression

After exponentiating both sides: $\exp(\text{logit} \frac{p(x)}{1-p(x)}) = \exp(a + bx + cz_{ij})$

$\frac{p(x)}{1-p(x)} = e^{a + bx + cz_{ij}}$

$[\text{odds} = e^{\text{linear regression}}]$

So our generalized linear model, parameterized by $\beta = (a, b, c)$, is: $\text{logit}(p) = \beta^T (1, x, z) = \frac{1}{1 + e^{-\beta^T x}}$

Therefore, $P(Y=1 | X, Z) = 1 - p_0(X, Z)$

And probability that $Y=0$ or $1 = P_Y(Y | \beta) = p_0(X)^Y (1 - p_0(X))^{1-Y}$

Likelihood function (assuming observations in sample are iid Bernoulli distributed):

$$L(\beta | z) = P_Y(Y | \beta)$$

$$= \prod_{i=1}^k \prod_{j=1}^{n_i} p_{ij}^{y_{ij}} (1 - p_{ij})^{1-y_{ij}}$$

$$= \prod_{i=1}^k \prod_{j=1}^{n_i} \left(\frac{\exp(a + bx_i + cz_{ij})}{1 + \exp(a + bx_i + cz_{ij})} \right)^{y_{ij}} \left(\frac{1}{1 + \exp(a + bx_i + cz_{ij})} \right)^{1-y_{ij}}$$

Summing over all data points

Requires summing over each j data point in both x and z classes

Problem 7.20b: We can run a standard GLM on the data and get the estimated mean, standard error, and variance of a , b , and c .

The mean, standard error, and variance of a are **-1.9739**, **0.221**, **0.0489**, respectively.

The mean, standard error, and variance of b are **0.1622**, **0.080**, **0.0064**, respectively.

The mean, standard error, and variance of c are **0.2844**, **0.093**, **0.0086**, respectively.

Problem 7.20c: Now we employ normal candidate densities, using the GLM estimates for mean and variance, in a Metropolis-Hastings algorithm that samples from the likelihood function. Below are histograms and kernel density estimates of the parameter values using this method.

