Text as Data

Justin Grimmer

Associate Professor Department of Political Science University of Chicago

February 12th, 2018

Three categories of documents

Hand labeled

- Training set (what we'll use to estimate model)
- Validation set (what we'll use to assess model)

Unlabeled

- Test set (what we'll use the model to categorize)

Label more documents than necessary to train model

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \leadsto \{\text{not, credit claiming}\}$

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \leadsto \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} (y_i - \beta' \boldsymbol{x}_i)^2$$

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\boldsymbol{\beta}} = \arg \min_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

$$= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

$$= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Problem:

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

$$= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Problem:

- J will likely be large (perhaps J > N)

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

$$= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Problem:

- J will likely be large (perhaps J > N)
- There many correlated variables

Suppose we have N documents, with each document i having label $y_i \in \{-1,1\} \rightsquigarrow \{\text{not, credit claiming}\}$ We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2$$

$$\widehat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

$$= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Problem:

- J will likely be large (perhaps J > N)
- There many correlated variables

Predictions will be variable

Suppose $\boldsymbol{\theta}$ is some value of the true parameter

Suppose θ is some value of the true parameter Bias:

Suppose θ is some value of the true parameter Bias:

$$\mathsf{Bias} \ = \ \mathsf{E}[\widehat{\theta} - \theta]$$

Suppose θ is some value of the true parameter Bias:

Bias =
$$E[\widehat{\theta} - \theta]$$

Suppose θ is some value of the true parameter Bias:

Bias =
$$E[\widehat{\theta} - \theta]$$

$$\mathsf{E}[(\hat{\theta}-\theta)^2]$$

Suppose θ is some value of the true parameter Bias:

Bias =
$$E[\widehat{\theta} - \theta]$$

$$\mathsf{E}[(\hat{\theta} - \theta)^2] = \mathsf{E}[\hat{\theta}^2] - 2\theta \mathsf{E}[\hat{\theta}] + \theta^2$$

Suppose θ is some value of the true parameter Bias:

Bias =
$$E[\widehat{\theta} - \theta]$$

$$E[(\hat{\theta} - \theta)^2] = E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2$$
$$= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2$$

Suppose θ is some value of the true parameter Bias:

Bias
$$= E[\widehat{\theta} - \theta]$$

$$E[(\hat{\theta} - \theta)^2] = E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2$$

$$= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2$$

$$= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + (E[\hat{\theta} - \theta])^2$$

Suppose θ is some value of the true parameter Bias:

Bias
$$= E[\widehat{\theta} - \theta]$$

$$E[(\hat{\theta} - \theta)^{2}] = E[\hat{\theta}^{2}] - 2\theta E[\hat{\theta}] + \theta^{2}$$

$$= E[\hat{\theta}^{2}] - E[\hat{\theta}]^{2} + E[\hat{\theta}]^{2} - 2\theta E[\hat{\theta}] + \theta^{2}$$

$$= E[\hat{\theta}^{2}] - E[\hat{\theta}]^{2} + (E[\hat{\theta} - \theta])^{2}$$

$$= Var(\hat{\theta}) + Bias^{2}$$

Suppose θ is some value of the true parameter Bias:

Bias
$$= E[\widehat{\theta} - \theta]$$

We may care about average distance from truth

$$E[(\hat{\theta} - \theta)^{2}] = E[\hat{\theta}^{2}] - 2\theta E[\hat{\theta}] + \theta^{2}$$

$$= E[\hat{\theta}^{2}] - E[\hat{\theta}]^{2} + E[\hat{\theta}]^{2} - 2\theta E[\hat{\theta}] + \theta^{2}$$

$$= E[\hat{\theta}^{2}] - E[\hat{\theta}]^{2} + (E[\hat{\theta} - \theta])^{2}$$

$$= Var(\hat{\theta}) + Bias^{2}$$

To reduce MSE, we are willing to induce bias to decrease variance we methods that shrink coefficients toward zero

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y})$$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2$$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \beta_j^2$$
Penalty

Penalty for model complexity

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \beta_j^2$$
Penalty

where:

Penalty for model complexity

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \beta_j^2$$
Penalty

where:

- $\beta_0 \rightsquigarrow \text{intercept}$

Penalty for model complexity

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^{J} \beta_j^2}_{\text{Penalty}}$$

where:

- $\beta_0 \rightsquigarrow \text{intercept}$
- $\lambda \leadsto$ penalty parameter

Penalty for model complexity

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^{J} \beta_j^2}_{\text{Penalty}}$$

where:

- $\beta_0 \rightsquigarrow \text{intercept}$
- $\lambda \leadsto$ penalty parameter
- Standardized **X** (coefficients on same scale)

$$oldsymbol{eta}^{\mathsf{Ridge}} \ = \ \arg \, \min_{oldsymbol{eta}} \left\{ f(oldsymbol{eta}, oldsymbol{X}, oldsymbol{Y})
ight\}$$

$$\begin{split} \boldsymbol{\beta}^{\mathsf{Ridge}} &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) \right\} \\ &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \beta_j^2 \right\} \end{split}$$

$$\begin{split} \boldsymbol{\beta}^{\mathsf{Ridge}} &= & \arg \min_{\boldsymbol{\beta}} \left\{ f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) \right\} \\ &= & \arg \min_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \beta_j^2 \right\} \\ &= & \arg \min_{\boldsymbol{\beta}} \left\{ (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta})' (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta}) + \lambda \boldsymbol{\beta}' \boldsymbol{\beta} \right\} \end{split}$$

Demean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

$$\begin{split} \boldsymbol{\beta}^{\mathsf{Ridge}} &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) \right\} \\ &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^{N} \left(y_{i} - \beta_{0} - \sum_{j=1}^{J} \beta_{j} x_{ij} \right)^{2} + \lambda \sum_{j=1}^{J} \beta_{j}^{2} \right\} \\ &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta})' (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta}) + \lambda \boldsymbol{\beta}' \boldsymbol{\beta} \right\} \\ &= \left(\boldsymbol{X}' \boldsymbol{X} + \lambda \boldsymbol{I}_{J} \right)^{-1} \boldsymbol{X}' \boldsymbol{Y} \end{split}$$

Demean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

$$\begin{split} \boldsymbol{\beta}^{\mathsf{Ridge}} &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) \right\} \\ &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^{N} \left(y_{i} - \beta_{0} - \sum_{j=1}^{J} \beta_{j} x_{ij} \right)^{2} + \lambda \sum_{j=1}^{J} \beta_{j}^{2} \right\} \\ &= \operatorname{arg\ min}_{\boldsymbol{\beta}} \left\{ (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta})' (\boldsymbol{Y} - \boldsymbol{X}' \boldsymbol{\beta}) + \lambda \boldsymbol{\beta}' \boldsymbol{\beta} \right\} \\ &= \left(\boldsymbol{X}' \boldsymbol{X} + \lambda \boldsymbol{I}_{J} \right)^{-1} \boldsymbol{X}' \boldsymbol{Y} \end{split}$$

Demean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

Ridge Regression → Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

Ridge Regression → Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

Suppose
$$\mathbf{X}'\mathbf{X} = \mathbf{I}_J$$
.

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$
$$= \boldsymbol{X}'\boldsymbol{Y}$$

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= \boldsymbol{X}'\boldsymbol{Y}$$

$$\boldsymbol{\beta}^{\mathsf{ridge}} = (\boldsymbol{X}'\boldsymbol{X} + \lambda \boldsymbol{I}_{J})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= \boldsymbol{X}'\boldsymbol{Y}$$

$$\boldsymbol{\beta}^{\text{ridge}} = (\boldsymbol{X}'\boldsymbol{X} + \lambda \boldsymbol{I}_J)^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= (\boldsymbol{I}_J + \lambda \boldsymbol{I}_J)^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= \boldsymbol{X}'\boldsymbol{Y}$$

$$\boldsymbol{\beta}^{\text{ridge}} = (\boldsymbol{X}'\boldsymbol{X} + \lambda \boldsymbol{I}_J)^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= (\boldsymbol{I}_J + \lambda \boldsymbol{I}_J)^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

$$= (\boldsymbol{I}_J + \lambda \boldsymbol{I}_J)^{-1}\widehat{\boldsymbol{\beta}}$$

$$\widehat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}
= \boldsymbol{X}'\boldsymbol{Y}
\boldsymbol{\beta}^{\text{ridge}} = (\boldsymbol{X}'\boldsymbol{X} + \lambda \boldsymbol{I}_{J})^{-1}\boldsymbol{X}'\boldsymbol{Y}
= (\boldsymbol{I}_{j} + \lambda \boldsymbol{I}_{j})^{-1}\boldsymbol{X}'\boldsymbol{Y}
= (\boldsymbol{I}_{j} + \lambda \boldsymbol{I}_{j})^{-1}\widehat{\boldsymbol{\beta}}
\boldsymbol{\beta}_{j}^{\text{Ridge}} = \frac{\widehat{\boldsymbol{\beta}}_{j}}{1 + \lambda}$$

$$eta_{j} \sim \operatorname{Normal}(0, \tau^{2})$$
 $y_{i} \sim \operatorname{Normal}(eta_{0} + \mathbf{x}_{i}^{'} oldsymbol{eta}, \sigma^{2})$

$$oldsymbol{eta}_{j} \sim \operatorname{Normal}(0, au^{2})$$
 $y_{i} \sim \operatorname{Normal}(eta_{0} + \mathbf{x}_{i}^{'} oldsymbol{eta}, \sigma^{2})$

$$p(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) \propto \prod_{j=1}^{J} p(\beta_j) \prod_{i=1}^{N} p(y_i|\boldsymbol{x}_i,\boldsymbol{\beta})$$

$$oldsymbol{eta}_{j} \sim \operatorname{Normal}(0, au^{2})$$
 $y_{i} \sim \operatorname{Normal}(eta_{0} + \mathbf{x}_{i}^{'} oldsymbol{eta}, \sigma^{2})$

$$p(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) \propto \prod_{j=1}^{J} p(\beta_{j}) \prod_{i=1}^{N} p(y_{i}|\boldsymbol{x}_{i},\boldsymbol{\beta})$$

$$\propto \prod_{j=1}^{J} \frac{1}{\sqrt{2\pi}\tau} \exp\left(-\frac{\beta_{j}^{2}}{2\tau^{2}}\right) \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_{i}-\beta_{0}-\boldsymbol{x}_{i}'\boldsymbol{\beta})^{2}}{2\sigma^{2}}\right)$$

$$\log p(\beta|X,Y) = -\sum_{j=1}^{J} \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^{N} \frac{(y_i - \beta_0 - x'\beta)^2}{2\sigma^2}$$

$$\log p(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) = -\sum_{j=1}^{J} \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^{N} \frac{(y_i - \beta_0 - \boldsymbol{x}'\boldsymbol{\beta})^2}{2\sigma^2}$$
$$-2\sigma^2 \log p(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} (y_i - \beta_0 - \boldsymbol{x}'\boldsymbol{\beta})^2 + \sum_{i=1}^{J} \frac{\sigma^2}{\tau^2} \beta_j^2$$

$$\log p(\beta | \mathbf{X}, \mathbf{Y}) = -\sum_{j=1}^{J} \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^{N} \frac{(y_i - \beta_0 - \mathbf{x}'\beta)^2}{2\sigma^2}$$
$$-2\sigma^2 \log p(\beta | \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta_0 - \mathbf{x}'\beta)^2 + \sum_{j=1}^{J} \frac{\sigma^2}{\tau^2} \beta_j^2$$

where:

$$\log p(\beta | \mathbf{X}, \mathbf{Y}) = -\sum_{j=1}^{J} \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^{N} \frac{(y_i - \beta_0 - \mathbf{x}'\beta)^2}{2\sigma^2}$$
$$-2\sigma^2 \log p(\beta | \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^{N} (y_i - \beta_0 - \mathbf{x}'\beta)^2 + \sum_{j=1}^{J} \frac{\sigma^2}{\tau^2} \beta_j^2$$

where:

$$- \lambda = \frac{\sigma^2}{\tau^2}$$

Definition

Suppose \boldsymbol{X} is an $N \times J$ matrix. Then \boldsymbol{X} can be written as:

$$X = \underbrace{U}_{N \times N} \underbrace{S}_{N \times J} \underbrace{V'}_{J \times J}$$

Where:

$$U'U = I_N$$

 $V'V = VV' = I_J$

S contains min(N, J) singular values, $\sqrt{\lambda_j} \geq 0$ down the diagonal and then 0's for the remaining entries

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvector}}$$

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \mathbf{V} \mathbf{S}' \underbrace{\left(\mathbf{U}' \mathbf{U} \right)}_{I_{I}} \mathbf{S} \mathbf{V}'$$

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

$$\frac{1}{N}X'X = VS'\underbrace{(U'U)}_{I_J}SV'$$
$$= \frac{1}{N}VS'SV'$$

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \mathbf{V} \mathbf{S}' \underbrace{\left(\mathbf{U}' \mathbf{U}\right)}_{I_{J}} \mathbf{S} \mathbf{V}'$$

$$= \frac{1}{N} \mathbf{V} \mathbf{S}' \mathbf{S} \mathbf{V}'$$

$$= \underbrace{\mathbf{V}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_{1} & 0 & \dots & 0 \\ 0 & \lambda_{2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_{J} \end{pmatrix} \underbrace{\mathbf{V}'}_{\text{eigenvectors}}$$

We can write the predicted values for a regular regression as

$$\hat{Y} = X\hat{\beta}
= X (X'X)^{-1} X'Y
= UU'Y = \sum_{j=1}^{J} u_j u'_j Y$$

We can write the predicted values for a regular regression as

$$\hat{Y} = X\hat{\beta}
= X(X'X)^{-1}X'Y
= UU'Y = \sum_{j=1}^{J} u_j u'_j Y$$

We can write β^{ridge} as

We can write the predicted values for a regular regression as

$$\hat{Y} = X\hat{\beta}
= X (X'X)^{-1} X'Y
= UU'Y = \sum_{j=1}^{J} u_j u'_j Y$$

We can write β^{ridge} as

$$\hat{\mathbf{Y}}^{\mathsf{ridge}} = \mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

We can write the predicted values for a regular regression as

$$\hat{Y} = X\hat{\beta}
= X (X'X)^{-1} X'Y
= UU'Y = \sum_{j=1}^{J} u_j u'_j Y$$

We can write β^{ridge} as

$$\hat{\mathbf{Y}}^{\text{ridge}} = \mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

$$= \mathbf{U} \tilde{\mathbf{S}} \mathbf{U}' \mathbf{Y}$$

Where

$$\tilde{\mathbf{S}} = \left[\mathbf{S} (\mathbf{S}' \mathbf{S} + \lambda \mathbf{I}_{J})^{-1} \mathbf{S} \right]$$

We can write the predicted values for a regular regression as

$$\hat{Y} = X\hat{\beta}
= X (X'X)^{-1} X'Y
= UU'Y = \sum_{j=1}^{J} u_j u'_j Y$$

We can write β^{ridge} as

$$\hat{Y}^{\text{ridge}} = X (X'X + \lambda I_J)^{-1} X'Y$$

$$= U\tilde{S}U'Y$$

Where

$$\tilde{\mathbf{S}} = \left[\mathbf{S} (\mathbf{S}' \mathbf{S} + \lambda \mathbf{I}_{J})^{-1} \mathbf{S} \right]$$

Which we can write as:

$$\hat{\mathbf{Y}}^{ ext{ridge}} = \sum_{i=1}^J \mathbf{\textit{u}}_j rac{\lambda_j}{\lambda_j + \lambda} \mathbf{\textit{u}}_j^{'} \mathbf{\textit{Y}}$$

Degrees of Freedom for Ridge

We will say that the degrees of freedom for Ridge regression with penalty λ is

$$dof(\lambda) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \lambda}$$

Lasso Regression Objective Function

Different Penalty for Model Complexity

$$f(\beta, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \underbrace{|\beta_j|}_{\text{Penalty}}$$

Lasso Regression Objective Function

Different Penalty for Model Complexity

$$f(\beta, \boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \underbrace{|\beta_j|}_{\text{Penalty}}$$

Lasso Regression Optimization

Definition

Coordinate Descent Algorithms:

Consider $g: \Re^J \to \Re$. Our goal is to find $\mathbf{x}^* \in \Re^J$ such that $g(\mathbf{x}^*) \leq g(\mathbf{x})$ for all $\mathbf{x} \in \Re$.

To find x^* :

Until convergence: for each iteration t and each coordinate j

$$\mathbf{x}_{j}^{t+1} \ = \ \arg \min_{\mathbf{x}_{j} \in \Re} g\big(\mathbf{x}_{1}^{t+1}, \mathbf{x}_{2}^{t+1}, \dots, \mathbf{x}_{j-1}^{t+1}, \mathbf{x}_{j}, \mathbf{x}_{j+1}^{t}, \dots, \mathbf{x}_{J}^{t}\big)$$

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$
- Case 2: If $\beta_j > (<)0 \leadsto$ differentiable \leadsto differentiate and solve for β_j

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$
- Case 2: If $\beta_j > (<)0 \leadsto$ differentiable \leadsto differentiate and solve for β_j

Define
$$\tilde{y}_{i}^{j} = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$$

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$
- Case 2: If $\beta_j > (<)0 \leadsto$ differentiable \leadsto differentiate and solve for β_j

Define
$$\tilde{y}_i^j = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$$

 $r^j \equiv \frac{1}{N} \sum_{i=1}^{N} x_{ij} (y_i - \tilde{y}_i^j)$

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$
- Case 2: If $\beta_j > (<)0 \leadsto$ differentiable \leadsto differentiate and solve for β_j

Define
$$\tilde{y}_{i}^{j} = \beta_{0} + \sum_{l \neq j} x_{il} \beta_{l}$$

$$r^{j} \equiv \frac{1}{N} \sum_{i=1}^{N} x_{ij} (y_{i} - \tilde{y}_{i}^{j})$$
Update step for β_{j} is

$$\tilde{f}(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} |\beta_j|$$

- Case 1: If $\beta_j=0 \leadsto$ not differentiable. But $\beta_j=0$
- Case 2: If $\beta_j > (<)0 \leadsto$ differentiable \leadsto differentiate and solve for β_j

Define
$$\tilde{y}_{i}^{j} = \beta_{0} + \sum_{l \neq j} x_{il} \beta_{l}$$

$$r^{j} \equiv \frac{1}{N} \sum_{i=1}^{N} x_{ij} (y_{i} - \tilde{y}_{i}^{j})$$
Update step for β_{j} is

$$\beta_j \leftarrow \operatorname{sign}(r^j) \max(|r^j| - \lambda, 0)$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Lasso Regression \leadsto Intuition 1, Soft Thresholding Suppose again $\textbf{X}'\textbf{X} = \textbf{I}_J$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = (Y - \boldsymbol{X}\boldsymbol{\beta})'(Y - \boldsymbol{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^{J} |\beta_{j}|$$

Lasso Regression \leadsto Intuition 1, Soft Thresholding Suppose again $\textbf{X}'\textbf{X} = \textbf{I}_J$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = (Y - \boldsymbol{X}\boldsymbol{\beta})'(Y - \boldsymbol{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^{J} |\beta_{j}|$$
$$= -2\boldsymbol{X}'\boldsymbol{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^{J} |\beta_{j}|$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$f(\beta, \mathbf{X}, \mathbf{Y}) = (Y - \mathbf{X}\beta)'(Y - \mathbf{X}\beta) + \lambda \sum_{j=1}^{J} |\beta_j|$$
$$= -2\mathbf{X}'\mathbf{Y}\beta + \beta'\beta + \lambda \sum_{j=1}^{J} |\beta_j|$$

The coefficient is

Lasso Regression \leadsto Intuition 1, Soft Thresholding Suppose again $\textbf{X}'\textbf{X} = \textbf{I}_J$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = (Y - \boldsymbol{X}\boldsymbol{\beta})'(Y - \boldsymbol{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^{J} |\beta_{j}|$$
$$= -2\boldsymbol{X}'\boldsymbol{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^{J} |\beta_{j}|$$

The coefficient is

$$eta_j^{\mathrm{LASSO}} = \mathrm{sign}\left(\widehat{eta}_j\right) \left(|\widehat{eta}_j| - \lambda\right)_+$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = (Y - \boldsymbol{X}\boldsymbol{\beta})'(Y - \boldsymbol{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^{J} |\beta_{j}|$$
$$= -2\boldsymbol{X}'\boldsymbol{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^{J} |\beta_{j}|$$

The coefficient is

$$\beta_j^{\rm LASSO} \ = \ {\rm sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

- $sign(\cdot) \rightsquigarrow 1 \text{ or } -1$

Lasso Regression \leadsto Intuition 1, Soft Thresholding Suppose again $\textbf{X}'\textbf{X} = \textbf{I}_J$

$$f(\boldsymbol{\beta}, \boldsymbol{X}, \boldsymbol{Y}) = (Y - \boldsymbol{X}\boldsymbol{\beta})'(Y - \boldsymbol{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^{J} |\beta_{j}|$$
$$= -2\boldsymbol{X}'\boldsymbol{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^{J} |\beta_{j}|$$

The coefficient is

$$\beta_j^{\rm LASSO} \ = \ {\rm sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

-
$$sign(\cdot) \rightsquigarrow 1 \text{ or } -1$$

$$- \ \left(|\widehat{\beta}_j| - \lambda \right)_+ = \max(|\widehat{\beta}_j| - \lambda, 0)$$

Compare soft assignment

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \operatorname{sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \operatorname{sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

With hard assignment, selecting M biggest components

Compare soft assignment

$$\beta_j^{\rm LASSO} \ = \ {\rm sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

With hard assignment, selecting M biggest components

$$\beta_j^{\text{subset}} = \widehat{\beta}_j \cdot I\left(|\widehat{\beta}_j| \ge |\widehat{\beta}_{(M)}|\right)$$

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \operatorname{sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

With hard assignment, selecting M biggest components

$$\beta_j^{\text{subset}} = \widehat{\beta}_j \cdot I\left(|\widehat{\beta}_j| \ge |\widehat{\beta}_{(M)}|\right)$$

Intuition 2: Prior on coefficients → Laplace "The Bayesian LASSO"

Compare soft assignment

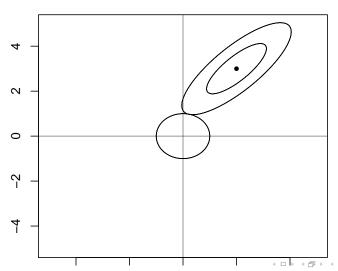
$$\beta_j^{\text{LASSO}} = \operatorname{sign}\left(\widehat{\beta}_j\right) \left(|\widehat{\beta}_j| - \lambda\right)_+$$

With hard assignment, selecting M biggest components

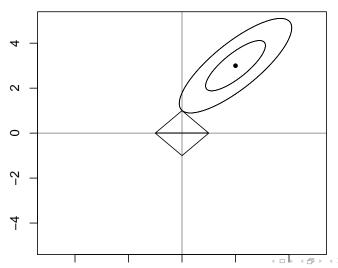
$$\beta_j^{\text{subset}} = \widehat{\beta}_j \cdot I\left(|\widehat{\beta}_j| \ge |\widehat{\beta}_{(M)}|\right)$$

Intuition 2: Prior on coefficients \leadsto Laplace "The Bayesian LASSO" Why does LASSO induce sparsity?

Ridge Regression



LASSO Regression



Contrast
$$\beta=(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}})$$
 and $\tilde{\beta}=(1,0)$

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$ Under ridge:

Contrast $\beta=(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}})$ and $\tilde{\beta}=(1,0)$ Under ridge:

$$\sum_{j=1}^{2} \beta_{j}^{2} = \frac{1}{2} + \frac{1}{2} = 1$$

Contrast $\beta=(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}})$ and $\tilde{\beta}=(1,0)$ Under ridge:

$$\sum_{j=1}^{2} \beta_{j}^{2} = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^{2} \tilde{\beta}_{j}^{2} = 1 + 0 = 1$$

Contrast $\beta=(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}})$ and $\tilde{\beta}=(1,0)$ Under ridge:

$$\sum_{j=1}^{2} \beta_{j}^{2} = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^{2} \tilde{\beta}_{j}^{2} = 1 + 0 = 1$$

Under LASSO

Contrast $\beta=\left(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}}\right)$ and $\tilde{\beta}=\left(1,0\right)$

Under ridge:

$$\sum_{j=1}^{2} \beta_{j}^{2} = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^{2} \tilde{\beta}_{j}^{2} = 1 + 0 = 1$$

Under LASSO

$$\sum_{j=1}^{2} |\beta_j| = \frac{1}{\sqrt{2}} + \frac{1}{\sqrt{2}} = \sqrt{2}$$

Contrast $\beta=(\frac{1}{\sqrt{2}},\frac{1}{\sqrt{2}})$ and $\tilde{\beta}=(1,0)$ Under ridge:

 $\sum_{i=1}^{2}$

$$\sum_{j=1}^{2} \beta_{j}^{2} = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^{2} \tilde{\beta}_{j}^{2} = 1 + 0 = 1$$

Under LASSO

$$\sum_{j=1}^{2} |\beta_j| = \frac{1}{\sqrt{2}} + \frac{1}{\sqrt{2}} = \sqrt{2}$$

$$\sum_{j=1}^{2} |\tilde{\beta}_j| = 1 + 0 = 1$$

Ridge and LASSO: The Elastic-Net

Combining the two criteria → Elastic-Net

$$f(\beta, X, Y) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

Ridge and LASSO: The Elastic-Net

Combining the two criteria → Elastic-Net

$$f(\beta, X, Y) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

The new update step (for coordinate descent:)

Ridge and LASSO: The Elastic-Net

Combining the two criteria → Elastic-Net

$$f(\beta, X, Y) = \frac{1}{2N} \sum_{i=1}^{N} \left(y_i - \beta_0 - \sum_{j=1}^{J} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{J} \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

The new update step (for coordinate descent:)

$$\beta_j \leftarrow \frac{\operatorname{sign}(r^j)\operatorname{max}(|r^j| - \lambda\alpha, 0)}{1 + \lambda(1 - \alpha)}$$

Selecting λ

How do we determine λ ? \leadsto Cross validation

Selecting λ

How do we determine λ ? \leadsto Cross validation Applying models gives score (probability) of document belong to class \leadsto threshold to classify

Selecting λ

How do we determine λ ? \leadsto Cross validation Applying models gives score (probability) of document belong to class \leadsto threshold to classify

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$. Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

Use MLE to obtain $\hat{\beta}$.

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$. Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$. Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

$$L\left(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i)\right)$$

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

$$L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i)) = (Y_i - f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))^2$$

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

Use MLE to obtain $\hat{\beta}$. Potential loss functions:

Assume:

$$L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i)) = (Y_i - f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))^2$$
$$= |Y_i - f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i)|$$

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

 $\mu_i = f(\beta, \mathbf{x}_i)$

$$L(Y_{i}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i})) = (Y_{i} - f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i}))^{2}$$

$$= |Y_{i} - f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i})|$$

$$= I(Y_{i} = I(f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i}) > \tau))$$

Training and Test Sets

The useful "fiction" of training and test sets:

Training and Test Sets

The useful "fiction" of training and test sets:

- Training set: data set used to fit the model

Training and Test Sets

The useful "fiction" of training and test sets:

- Training set: data set used to fit the model
- Test set: data used to evaluate fit of the model

Training and Test Sets

The useful "fiction" of training and test sets:

- Training set: data set used to fit the model
- Test set: data used to evaluate fit of the model

Even if no division, useful to think about systematic components of data.

Suppose that we have:

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = \textit{N}_{\mathsf{train}}$

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = \textit{N}_{\text{test}}$

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\mathsf{test}}$

Training (in-sample) error is:

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = \textit{N}_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\mathsf{test}}$

Training (in-sample) error is:

Error_{in} =

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\mathsf{test}}$

Training (in-sample) error is:

Error_{in} =
$$\sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))$$

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = \textit{N}_{\mathsf{test}}$

Training (in-sample) error is:

Error_{in} =
$$\sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))$$

We'd like to estimate out of sample performance with

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

Error_{in} =
$$\sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))$$

We'd like to estimate out of sample performance with

$$\mathsf{Error}_{\mathsf{out}} = \mathsf{E}[L(\boldsymbol{Y}_{i \in \mathcal{O}}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\mathsf{test}}$

Training (in-sample) error is:

Error_{in} =
$$\sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))$$

We'd like to estimate out of sample performance with

$$\mathsf{Error}_{\mathsf{out}} = \mathsf{E}[L(\boldsymbol{Y}_{i \in \mathcal{O}}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

where the expectation is taken over samples for test sets and supposes we have a training set.

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\mathsf{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

Error_{in} =
$$\sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_i))$$

We'd like to estimate out of sample performance with

$$\mathsf{Error}_{\mathsf{out}} = \mathsf{E}[L(\boldsymbol{Y}_{i \in \mathcal{O}}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

where the expectation is taken over samples for test sets and supposes we have a training set.

Error =
$$E\left[E[L(\boldsymbol{Y}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{X}))|\mathcal{T}]\right]$$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$
 $var(\epsilon_i) = \sigma_{\epsilon}^2$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$
 $var(\epsilon_i) = \sigma_{\epsilon}^2$
Define $f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})$

```
Suppose Y_i = f(\mathbf{x}_i) + \epsilon_i
Where E[\epsilon_i] = 0
var(\epsilon_i) = \sigma_{\epsilon}^2
Define f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})
With squared error loss:
```

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$
 $var(\epsilon_i) = \sigma_{\epsilon}^2$
Define $f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})$
With squared error loss:

$$Error(x_0) = E[(Y_i - \hat{f}(x_i))^2 | x_i = x_0]$$

```
Suppose Y_i = f(\mathbf{x}_i) + \epsilon_i
Where E[\epsilon_i] = 0
var(\epsilon_i) = \sigma_{\epsilon}^2
Define f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})
With squared error loss:
```

$$Error(\mathbf{x}_0) = E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$
$$= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$
 $\text{var}(\epsilon_i) = \sigma_{\epsilon}^2$
Define $f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})$
With squared error loss:

$$Error(\mathbf{x}_0) = E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

$$= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

$$= \sigma_{\epsilon}^2 + \left[f(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right]^2 + E[(\hat{f}(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)])^2]$$

Suppose
$$Y_i = f(\mathbf{x}_i) + \epsilon_i$$

Where $E[\epsilon_i] = 0$
 $var(\epsilon_i) = \sigma_{\epsilon}^2$
Define $f(\hat{\boldsymbol{\beta}}, \mathbf{x}) = \hat{f}(\mathbf{x})$
With squared error loss:

$$Error(\mathbf{x}_0) = E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

$$= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

$$= \sigma_{\epsilon}^2 + \left[f(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right]^2 + E[\left(\hat{f}(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)]\right)^2]$$

$$= Irreducible error + Bias^2 + Variance$$

Probit Regression (for motivational purposes)

Suppose:

$$Y_i \sim \text{Bernoulli}(\pi_i)$$

 $\pi_i = \Phi(\beta' \mathbf{x}_i)$

where $\Phi(\cdot)$ is the cumulative normal distribution. Implies log-likelihood

$$\log L(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = \sum_{i=1}^{N} \left[Y_i \log \Phi(\boldsymbol{\beta}'\boldsymbol{x}_i) + (1-Y_i) \log(1-\Phi(\boldsymbol{\beta}'\boldsymbol{x}_i)) \right]$$

Log-likelihood is a loss function → overly optimistic: improves with more parameters

There are many ways to fit models And many choices made when performing model fit How do we choose?

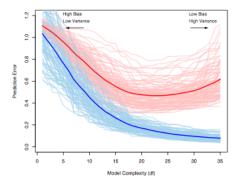


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error err, while the light red curves show the conditional test error Err for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error E[err].

There are many ways to fit models And many choices made when performing model fit How do we choose? Bad way to choose:

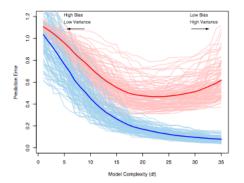


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error err, while the light red curves show the conditional test error Err for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error E[err].

There are many ways to fit models

And many choices made when performing model fit

How do we choose?

Bad way to choose: within sample model fit (HTF Figure 7.1)

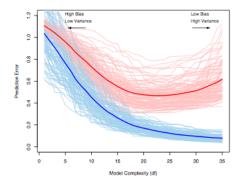


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error err, while the light red curves show the conditional test error Err for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error E[err].

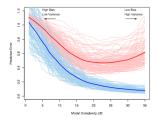


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error ēπ, while the light red curves show the conditional test error Ēπ-γ for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Ēπ and the expected training error Ēπ̄T.

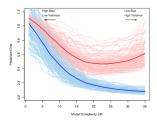


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error ēπ, while the light red curves show the conditional test error Ēπ-γ for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Ēπ and the expected training error Ēḡπ̄¬].

Model overfit → in sample error is optimistic:

- Some model complexity captures systematic features of the data

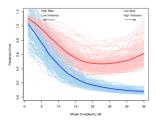


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error ēπ, while the light red curves show the conditional test error Ēπ-τ for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Ēπ and the expected training error Ēḡπ̄T].

- Some model complexity captures systematic features of the data
- Characteristics found in both training and test set

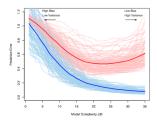


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error ēπ, while the light red curves show the conditional test error Ēπ-γ for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Ēπ and the expected training error Ēπ̄T.

- Some model complexity captures systematic features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set

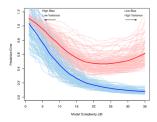


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error ēπ, while the light red curves show the conditional test error Ēπ-γ for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Ēπ and the expected training error Ēḡπ̄¬].

- Some model complexity captures systematic features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set
- Additional model complexity: idiosyncratic features of the training set

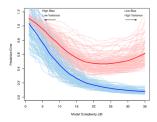


FIGURE 7.1. Rehavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error err, while the light red curves show the conditional test error Err. for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error EBETT!

- Some model complexity captures systematic features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set
- Additional model complexity: idiosyncratic features of the training set
- Reduces error in training set, increases error in test set

How Do We Choose Covariates?

Best model depends on task

- Causal inference observational study: make treatment assignment ignorable
- Prediction: improve predictive performance

Suppose we have P covariates. 2^P potential models

Suppose we have P covariates. 2^P potential models Stepwise procedures

Suppose we have P covariates.

2^P potential models

Stepwise procedures

- 1) Forward selection
 - a) No variables in model.
 - b) Check all variables p-value if include, include lowest p-value
 - c) Repeat until included p-value is above some threshold

Suppose we have P covariates.

2^P potential models

Stepwise procedures

- 1) Forward selection
 - a) No variables in model.
 - b) Check all variables p-value if include, include lowest p-value
 - c) Repeat until included p-value is above some threshold
- 2) Backward elimination
 - a) Fit model with all variables (if possible)
 - b) Remove variable with largest p-value
 - Repeat until potentially excluded p-value is below some threshold

Suppose we have P covariates.

2^P potential models

Stepwise procedures

- 1) Forward selection
 - a) No variables in model.
 - b) Check all variables p-value if include, include lowest p-value
 - c) Repeat until included p-value is above some threshold
- 2) Backward elimination
 - a) Fit model with all variables (if possible)
 - b) Remove variable with largest p-value
 - c) Repeat until potentially excluded p-value is below some threshold

Problematic:

- 1) Not optimal model selection (path dependent)
- 2) P-value \neq objective of model

Approximate optimism and compensate in loss function.

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) → Minimize

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As ${\it N}\to\infty$

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2\mathsf{E}[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[\mathsf{E}[\log \mathsf{L}(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2E[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[E[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$AIC = -2\left[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2E[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[E[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$AIC = -2\left[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2\mathsf{E}[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[\mathsf{E}[\log \mathsf{L}(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$\mathsf{AIC} = -2\left[\log \mathsf{L}(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

where d is the number of parameters in the model

- Intuition: balances model fit with penalty for complexity

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2E[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[E[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$AIC = -2\left[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate optimism in likelihood based models

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2E[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[E[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$AIC = -2\left[\log L(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate optimism in likelihood based models
- Derived from a method to compute similarity between estimated model and true model (under assumptions of course)

Approximate optimism and compensate in loss function. Akaike Information Criterion (AIC) \leadsto Minimize As $N\to\infty$

$$-2\mathsf{E}[\log P_{\hat{\boldsymbol{\beta}}}(\boldsymbol{Y})] = -2\left[\mathsf{E}[\log \mathsf{L}(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y})] - d\right]$$

$$\mathsf{AIC} = -2\left[\log \mathsf{L}(\hat{\boldsymbol{\beta}}|\boldsymbol{X},\boldsymbol{Y}) - d\right]$$

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate optimism in likelihood based models
- Derived from a method to compute similarity between estimated model and true model (under assumptions of course)
- Can be extended to general models, though requires estimate of irresolvable error

Bayesian Information Criterion (BIC) [Schwarz Criterion]

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

where d is again the effective number of parameters

- Intuition: balances model fit with penalty for complexity

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

- Intuition: balances model fit with penalty for complexity
- Derived from Bayesian approach to model selection

Bayesian Information Criterion (BIC) [Schwarz Criterion]

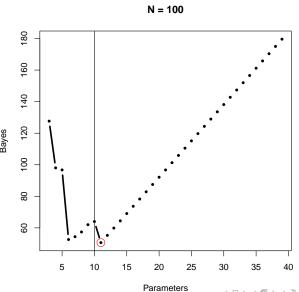
$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

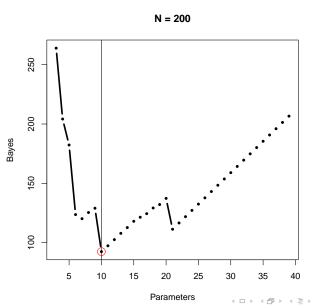
- Intuition: balances model fit with penalty for complexity
- Derived from Bayesian approach to model selection
- Approximation to Bayes' factor

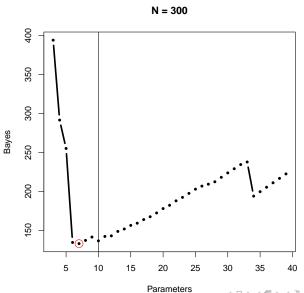
Bayesian Information Criterion (BIC) [Schwarz Criterion]

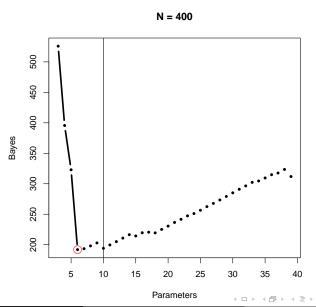
$$BIC = -2 \log L(\widehat{\beta}|\boldsymbol{X}, \boldsymbol{Y}) + (\log N)d$$

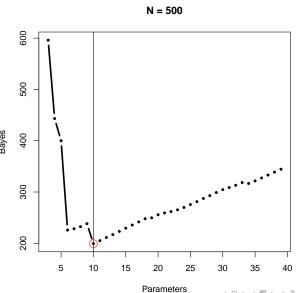
- Intuition: balances model fit with penalty for complexity
- Derived from Bayesian approach to model selection
- Approximation to Bayes' factor
- Penalizes more heavily than AIC

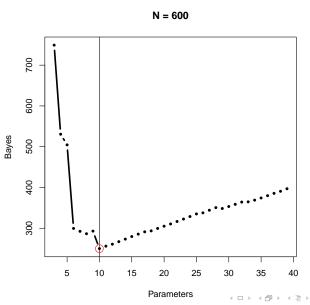


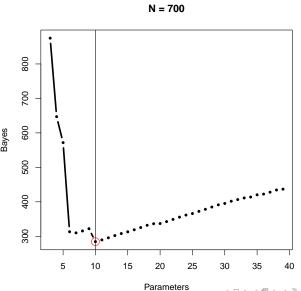


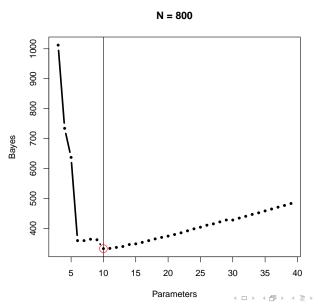


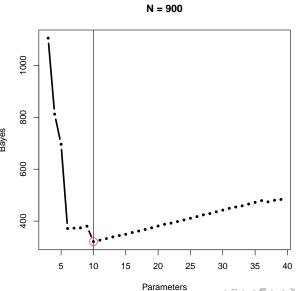


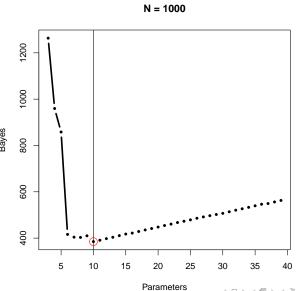


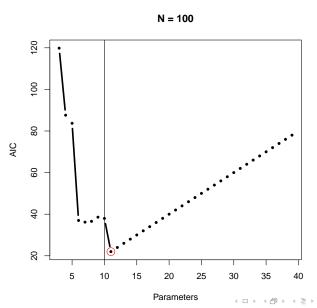


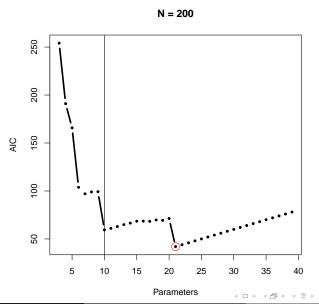


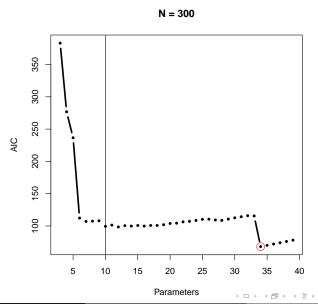


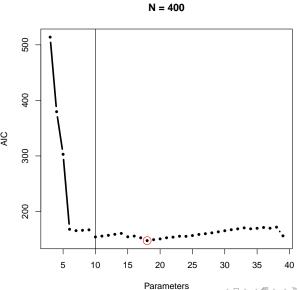


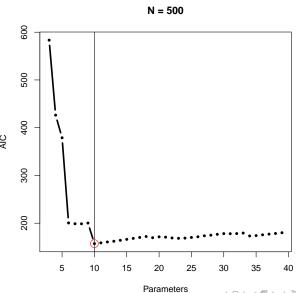


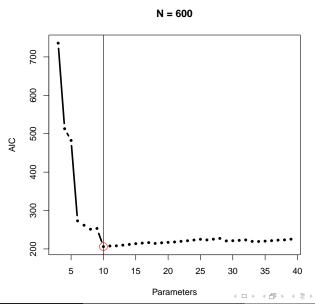


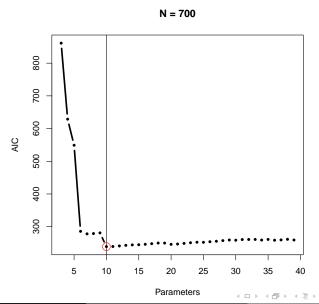


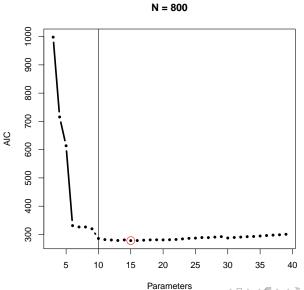


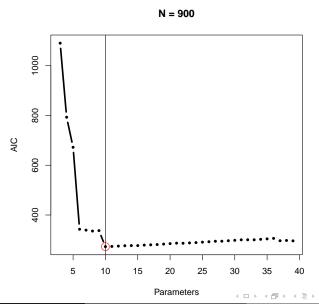




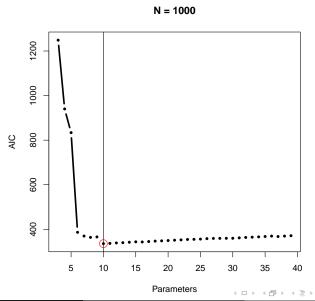








BIC or AIC?



BIC or AIC?

- BIC

- Asymptotically consistent if true model is in choice set
- As $N \to \infty$ will choose correct model with probability 1 (if available)
- Small samples → overpenalize

- AIC

- No asymptotic guarantees → derivation doesn't require truth in set. (KL-criteria)
- In large samples → favors complexity
- Small samples → avoids over penalization

Analytic statistics for selection, include penalty for complexity

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion

Analytic statistics for selection, include penalty for complexity

- AIC: Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Analytic statistics for selection, include penalty for complexity

- AIC: Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function

Analytic statistics for selection, include penalty for complexity

- AIC: Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

- Rely on specific loss function
- Rely on asymptotic argument

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters

Analytic statistics for selection, include penalty for complexity

- AIC: Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters
- Extremely model dependent

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters
- Extremely model dependent

Need: general tool for evaluating models, replicates decision problem

Optimal division of data for prediction:

Optimal division of data for prediction:

- Train: build model

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

- Test: predict remaining data

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets

- Each step: use held out data to evaluate performance

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets
- Each step: use held out data to evaluate performance
- Avoid overfitting and have context specific penalty

Optimal division of data for prediction:

- Train: build model

- Validation: assess model

- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets
- Each step: use held out data to evaluate performance
- Avoid overfitting and have context specific penalty

Estimates:

Error =
$$E\left[E[L(\boldsymbol{Y}, f(\hat{\boldsymbol{\beta}}, \boldsymbol{X}))|\mathcal{T}]\right]$$

Process:

- Randomly partition data into $\ensuremath{\mathsf{K}}$ groups.

Process:

Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Process:

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step Training

Validation ("Test")

Process:

- Randomly partition data into K groups. (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

```
Step
      Training
```

Group2, Group3, Group 4, ..., Group K

Validation ("Test") Group 1

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4,, Group K	Group 1
2	Group 1, Group3, Group 4,, Group K	Group 2

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4,, Group K	Group 1
2	Group 1, Group3, Group 4,, Group K	Group 2
3	Group 1, Group 2, Group 4,, Group K	Group 3

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

: :
```

- Randomly partition data into K groups.
 (Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K
```

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4,, Group K	Group 1
2	Group 1, Group3, Group 4,, Group K	Group 2
3	Group 1, Group 2, Group 4,, Group K	Group 3
:	<u>:</u>	÷
K	Group 1, Group 2, Group 3,, Group K - 1	Group K

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
 - Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
 - Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
 - Predict values for Kth

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
- Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
- Predict values for Kth
- Summarize performance with loss function: $L(\boldsymbol{Y}_i, \hat{f}^{-k}(\boldsymbol{\beta}, \boldsymbol{X}))$

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
 - Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
 - Predict values for Kth
 - Summarize performance with loss function: $L(\boldsymbol{Y}_i, \hat{f}^{-k}(\boldsymbol{\beta}, \boldsymbol{X}))$
 - Mean square error, Absolute error, Prediction error, ...

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
- Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
- Predict values for Kth
- Summarize performance with loss function: $L(\boldsymbol{Y}_i,\hat{f}^{-k}(\boldsymbol{\beta},\boldsymbol{X}))$
 - Mean square error, Absolute error, Prediction error, ...

CV(ind. classification) =
$$\frac{1}{N} \sum_{i=1}^{N} L(\boldsymbol{Y}_i, f^{-k}(\boldsymbol{\beta}, \boldsymbol{X}_i))$$

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
- Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
- Predict values for Kth
- Summarize performance with loss function: $L(\boldsymbol{Y}_i,\hat{f}^{-k}(\boldsymbol{\beta},\boldsymbol{X}))$
 - Mean square error, Absolute error, Prediction error, ...

CV(ind. classification) =
$$\frac{1}{N} \sum_{i=1}^{N} L(\boldsymbol{Y}_i, f^{-k}(\boldsymbol{\beta}, \boldsymbol{X}_i))$$

 $\frac{1}{K}\sum_{j=1}^{K}$ Mean Square Error Proportions from Group j

```
Step Training Validation ("Test")

1 Group2, Group3, Group 4, ..., Group K Group 1

2 Group 1, Group3, Group 4, ..., Group K Group 2

3 Group 1, Group 2, Group 4, ..., Group K Group 3

...

K Group 1, Group 2, Group 3, ..., Group K - 1 Group K

Strategy:
```

- Divide data into K groups
- Train data on K-1 groups. Estimate $\hat{f}^{-K}(oldsymbol{eta},oldsymbol{\mathcal{X}})$
- Predict values for Kth
- Summarize performance with loss function: $L(Y_i, \hat{f}^{-k}(\beta, X))$
 - Mean square error, Absolute error, Prediction error, ...

CV(ind. classification) =
$$\frac{1}{N} \sum_{i=1}^{N} L(\boldsymbol{Y}_i, f^{-k}(\boldsymbol{\beta}, \boldsymbol{X}_i))$$

CV(proportions) =

 $\frac{1}{K}\sum_{i=1}^{K}$ Mean Square Error Proportions from Group j

- Final choice: model with highest CV score

How Do We Select K? (HTF, Section 7.10)

Common values of K

- K = 5: Five fold cross validation
- K = 10: Ten fold cross validation
- K = N: Leave one out cross validation

Considerations:

- How sensitive are inferences to number of coded documents? (HTF, pg 243-244)
- 200 labeled documents
 - $K = N \rightarrow 199$ documents to train,
 - $K=10 \rightarrow 180$ documents to train
 - $K=5 \rightarrow 160$ documents to train
- 50 labeled documents
 - $K = N \rightarrow 49$ documents to train,
 - $K = 10 \rightarrow 45$ documents to train
 - $K = 5 \rightarrow 40$ documents to train
- How long will it take to run models?
 - K-fold cross validation requires $K \times$ One model run
- What is the correct loss function?

If you cross validate, you really need to cross validate (Section 7.10.2, ESL)

- Use CV to estimate prediction error
- All supervised steps performed in cross-validation
- Underestimate prediction error
- Could lead to selecting lower performing model

Example from Facebook Data

What do people say to legislators? (Franco, Grimmer, and Lee 2017)

- 1) Example: estimating classification error
 - a) Accuracy in legislator posts: 75%
 - b) Accuracy in public posts: 66.25%

Credit Claiming (Back to Ridge/Lasso, Grimmer, Westwood, and Messing 2014)

```
library(glmnet)
set.seed(8675309) ##setting seed
folds<- sample(1:10, nrow(dtm), replace=T) ##assigning to fold
out_of_samp<- c() ##collecting the predictions</pre>
```

Credit Claiming (Back to Ridge/Lasso, Grimmer, Westwood, and Messing 2014)

```
for(z in 1:10){
train <- which (folds!=z) ##the observations we will use to train the model
test<- which(folds==z) ##the observations we will use to test the model
part1<- cv.glmnet(x = dtm[train,], y = credit[train], alpha = 1, family =</pre>
binomial) ##fitting the LASSO model on the data.
## alpha = 1 -> LASSO
## alpha = 0 -> RIDGE
## 0<alpha<1 -> Elastic-Net
out_of_samp[test] <- predict(part1, newx= dtm[test,], s = part1$lambda.min,
type =class) ##predicting the labels
print(z) ##printing the labels
conf_table<- table(out_of_samp, credit) ##calculating the confusion table</pre>
> round(sum(diag(conf_table))/len(credit), 3)
[1] 0.844
```

$$\boldsymbol{\beta}^{\mathsf{Ridge}} = \left(\boldsymbol{X}' \boldsymbol{X} + \lambda \boldsymbol{I}_{J} \right)^{-1} \boldsymbol{X}' \boldsymbol{Y}$$

$$eta^{\mathsf{Ridge}} = \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$
 $\widehat{\mathbf{Y}} = \mathbf{X} (\beta)^{\mathsf{Ridge}}$

$$\beta^{\text{Ridge}} = \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

$$\widehat{\mathbf{Y}} = \mathbf{X} (\beta)^{\text{Ridge}}$$

$$= \underbrace{\mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y}$$

$$\beta^{\text{Ridge}} = \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

$$\widehat{\mathbf{Y}} = \mathbf{X}(\beta)^{\text{Ridge}}$$

$$= \underbrace{\mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y}$$

$$\widehat{\mathbf{Y}} = \underbrace{\mathbf{H}}_{\text{Smoother Matrix}} \mathbf{Y}$$

$$\beta^{\text{Ridge}} = \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

$$\widehat{\mathbf{Y}} = \mathbf{X}(\beta)^{\text{Ridge}}$$

$$= \underbrace{\mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{J} \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y}$$

$$\widehat{\mathbf{Y}} = \underbrace{\mathbf{H}}_{\text{Smoother Matrix}} \mathbf{Y}$$

Why do we care?

Why do we care? Leave one out cross validation

Why do we care? Leave one out cross validation

Cross Validation(1) =
$$\frac{1}{N} \sum_{i=1}^{N} (Y_i - f(\mathbf{X}_{-i}, \mathbf{Y}_{-i}, \lambda, \hat{\boldsymbol{\beta}}))^2$$

Why do we care? Leave one out cross validation

Cross Validation(1)
$$= \frac{1}{N} \sum_{i=1}^{N} (Y_i - f(\mathbf{X}_{-i}, \mathbf{Y}_{-i}, \lambda, \hat{\boldsymbol{\beta}}))^2$$
$$= \frac{1}{N} \sum_{i=1}^{N} \left(\frac{Y_i - f(\mathbf{X}, \mathbf{Y}, \lambda, \hat{\boldsymbol{\beta}})}{1 - H_{ii}} \right)^2$$

Generalized Cross Validation and Ridge Regression Calculating **H** can be computationally expensive

Generalized Cross Validation and Ridge Regression Calculating **H** can be computationally expensive

- Trace
$$(\boldsymbol{H}) \equiv \operatorname{Tr}(\boldsymbol{H}) = \sum_{i=1}^{N} H_{ii}$$

Calculating **H** can be computationally expensive

- $\mathsf{Trace}(m{H}) \equiv \mathsf{Tr}(m{H}) = \sum_{i=1}^{N} H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$

Calculating **H** can be computationally expensive

- Trace $(oldsymbol{H}) \equiv \operatorname{Tr}(oldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

Calculating **H** can be computationally expensive

- Trace $(oldsymbol{H}) \equiv \mathrm{Tr}(oldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$\operatorname{Tr}(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\mathsf{Penalty}}}$$

Calculating **H** can be computationally expensive

- Trace $(oldsymbol{H}) \equiv \operatorname{Tr}(oldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables <math>+ 1$)
- For Ridge regression:

$$Tr(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{Penalty}}$$

where λ_j is the j^{th} Eigenvalue from $oldsymbol{\Sigma} = oldsymbol{X}'oldsymbol{X}$

Calculating **H** can be computationally expensive

- Trace $(oldsymbol{H}) \equiv \mathrm{Tr}(oldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$Tr(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{Penalty}}$$

where λ_j is the j^{th} Eigenvalue from $\boldsymbol{\Sigma} = \boldsymbol{X}' \boldsymbol{X}$ (!!!!!)

Calculating **H** can be computationally expensive

- Trace $(oldsymbol{H}) \equiv \operatorname{Tr}(oldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$Tr(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{Penalty}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Define generalized cross validation:

Calculating **H** can be computationally expensive

- Trace $(\boldsymbol{H}) \equiv \operatorname{Tr}(\boldsymbol{H}) = \sum_{i=1}^{N} H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$\operatorname{Tr}(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\mathsf{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\boldsymbol{\Sigma} = \boldsymbol{X}'\boldsymbol{X}$ (!!!!!)

Define generalized cross validation:

$$\mathsf{GCV} = \frac{1}{N} \sum_{i=1}^{N} \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\mathsf{Tr}(\mathbf{H})}{N}} \right)^2$$

Calculating **H** can be computationally expensive

- Trace $(\boldsymbol{H}) \equiv \operatorname{Tr}(\boldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$\operatorname{Tr}(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\mathsf{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}^{'}\mathbf{X}$ (!!!!!)

Define generalized cross validation:

$$\mathsf{GCV} = \frac{1}{N} \sum_{i=1}^{N} \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\mathsf{Tr}(\mathbf{H})}{N}} \right)^2$$

Applicable in any setting where we can write Smoother matrix

Calculating **H** can be computationally expensive

- Trace $(\boldsymbol{H}) \equiv \operatorname{Tr}(\boldsymbol{H}) = \sum_{i=1}^N H_{ii}$
- $Tr(\mathbf{H}) = Effective number of parameters (class regression = number of independent variables + 1)$
- For Ridge regression:

$$\operatorname{Tr}(\boldsymbol{H}) = \sum_{j=1}^{J} \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\mathsf{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}^{'}\mathbf{X}$ (!!!!!)

Define generalized cross validation:

$$\mathsf{GCV} = \frac{1}{N} \sum_{i=1}^{N} \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\mathsf{Tr}(\mathbf{H})}{N}} \right)^2$$

Applicable in any setting where we can write Smoother matrix