

BDA - Project

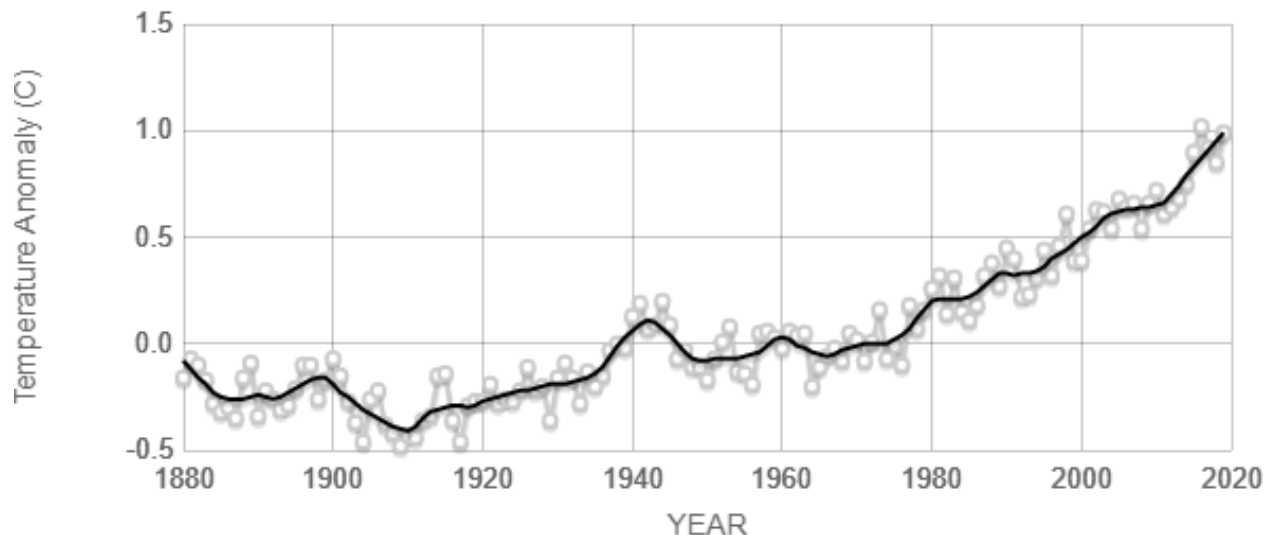
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1. Introduction

One of the biggest challenges of humankind in the 2020s is figuring out ways to slow down the growth of greenhouse gas emissions and stop global warming (due to human activities) under 2 °C. The increasing trend of global temperature is easily seen in Figure 1 [cite NASA] in which the global surface temperature is illustrated relative to 1951-1980 average temperatures. Warming can also be seen with one's own eyes by observing the winters that are warming year by year, by noticing that the number of devastating hurricanes is increased, and by finding out the increased rate of ice melting in glaciers during summer.



Source: climate.nasa.gov

Figure 1: Global Land-Ocean Temperature Index

In response to that warming, many countries have declared a climate emergency to emphasize the criticality of the situation. In addition, young people have organized climate demonstrations around the world, politicians are talking more and more about climate change, and presidents and prime ministers are negotiating agreements and commitments to solve this, one of humanity's greatest, problem. But what if, despite attempts of negotiation, the necessary CO₂ reduction decisions are not achieved?

In this project, our goal is to model the historical emission trends of selected countries as well as attempts to model their future emissions. We are examining a scenario in which emissions continue to develop at a historical rate, and the necessary reductions are not achieved. In our modeling, the other parameters e.g. population growth and technical conditions, are similar to historical data in our modeling.

2. Data description

Our CO₂ data was obtained from *Our World in Data* (OWID) web page [cite OWID_net] and the actual CSV file from OWID GitHub page [cite OWID_git]. As mentioned earlier, climate change is a hot topic in the daily news, and there is a lot of studies and research concerning how CO₂ emissions are influencing global warming. The data set was also used, for example, when researchers studied the climate impact of the different policy recommendation which targeted to reduce greenhouse gases from the atmosphere.

2.1 Choosing the sample and estimating it's resemblance

In our modeling, we selected 19 different countries from the OWID data set and examined CO₂ data between the years 1950-2018. We decided to not take all countries into the modeling as there are holes and missing

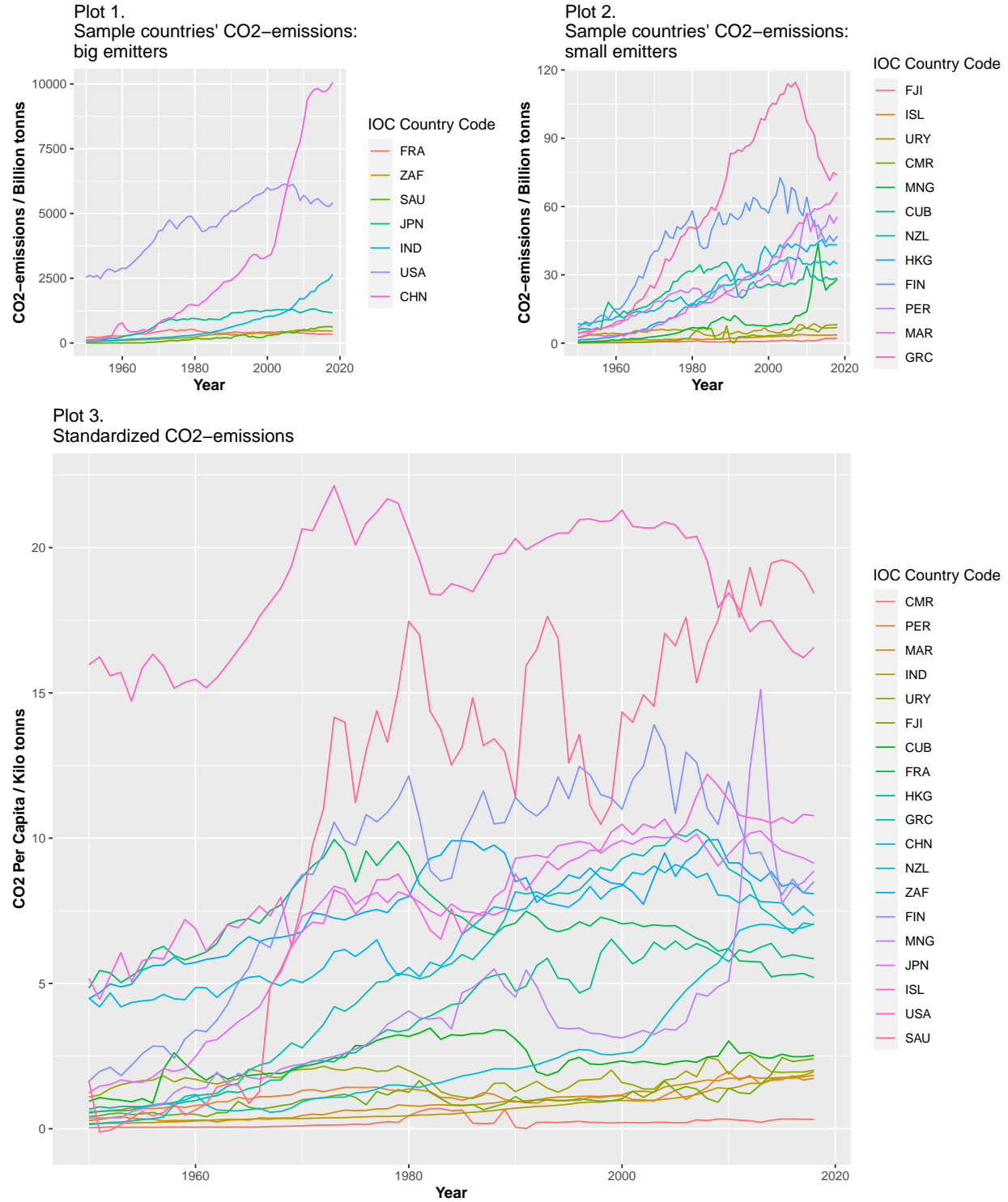
information in the dataset. The countries we chose cover the whole globe and are roughly evenly distributed across continents. However, we estimated that the data is probably more reliable in the western countries and thus were more open-minded in selecting them. Even that said, we think that the geographical distribution covers the whole world pretty well. Another important aspect of division is the division between large and small emitters. Even though it is quite difficult to perform such division, we tried to take countries from both ends pretty evenly. However, it is worth noting that this division was performed intuitively and it does not rely on any actual metrics. Lastly, we thought that the division between developing and western countries is extremely important to consider too. Therefore, this aspect was taken into account when considering the sample countries, too. We estimated that the number of developing countries in the world exceeds the number of western countries and thus tried to choose developing countries a bit more into the sample set.

For the reasons presented above, we believe that the sample we use in this project, resembles the situation in the world quite well. However, we estimated that it is possible that the sample is slightly biased towards western countries. It is important to note this since we examine results where the CO₂-emissions data is standardized with the countries' population. As the CO₂-emissions are standardized, the importance of correct ratio (number) of countries between different division-aspects increases. As the sample may be a bit biased, the results may propose higher numbers of CO₂-emissions per capita in the world than what they actually are.

2.2 Plotting the sample

Below is plotted three graphs. On the first row, we investigate our sample countries' CO₂-emissions by country. Please note the y-axis difference between large and small emitters in the graphs. It is worth noting that the CO₂-emissions development of China is very concerning as it has almost doubled its CO₂-emissions during the last 15 years. In addition, India, Greece, Morocco, Peru and Mongolia has been showing a bit concerning trend during last decades.

On the second row, we plotted the sample countries' emission standardized with the population of the country. Thus, we obtained a "CO₂ per capita" -estimate for each country. This is the data that we used later in our models. Especially between 1950s and 70s, western countries play significant role as the big emitters. However, during 2000s, the situation has changed as western countries have systematically been able to lower their emissions per capita. At the same time, developing countries have been increasing their emissions and thus the situation has tied.



3 Model description

In this chapter, we will present our model structures and stan codes of our implementation of a non-hierarchical pooled model and a hierarchical model. Before this, we briefly introduce the mathematical structure behind the models.

3.1 Pooled model

A pooled model is one of the most straightforward model structure to understand. In the pooled model, all data points are used as one “pool” without considering groups or particular features different pieces of data could have. The whole dataset is used as a one, and modeling is done based on that collection of data. If we assume that priors of the mean and standard deviation follow standard normal distributions, we can present the mathematical structure of the pooled model in the form

$$\mu \sim N(0, 1) \tag{1}$$

$$\sigma \sim N(0, 1) \tag{2}$$

$$y_i \sim N(\mu, \sigma) \tag{3}$$

We used these standardized normal priors just for illustration purposes, and the correct choice of priors we utilized when modeling is presented in chapter 4. Respectively, the pooled model’s implementation with probabilistic programming language *Stan* is presented in chapter 5.

3.2 Hierarchical normal model

Unlike the previous model, the hierarchical model takes into account the possibility that some of the subgroups of the whole dataset have similar properties. Due to this observation, the hierarchical model presents a “hyper-prior” that is common to each group. For each group, its posterior distribution of mean is calculated using that hyper-prior, taking into account only all samples belonging to that group. This property can be illustrated in Figure 2. In Figure 2, τ is a hyper-parameter and θ_i s presents modeled parameter of each group.

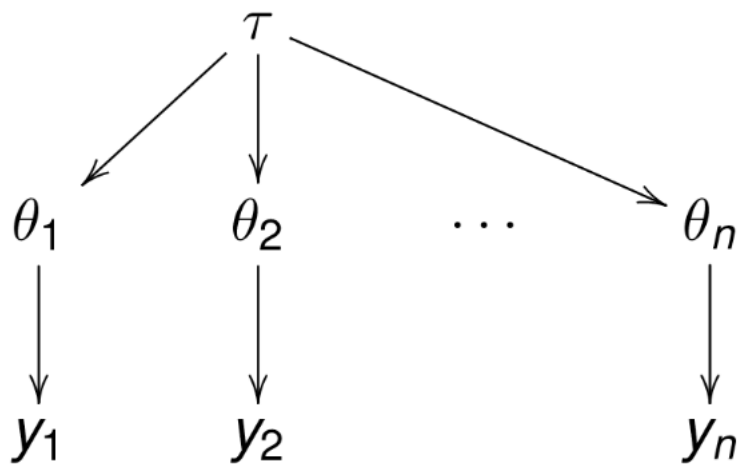


Figure 2: A hierarchical model –

We can summarize the hierarchical model mathematically as

$$\mu_0 \sim N(0, 1) \tag{4}$$

$$\sigma_0 \sim N(0, 1) \tag{5}$$

$$\mu_i \sim N(\mu_0, \sigma_0) \tag{6}$$

$$\sigma \sim N(0, 1) \tag{7}$$

$$y_{ji} \sim N(\mu_i, \sigma) \tag{8}$$

where μ_0 and σ_0 are hyper-priors for mean and standard deviation. Again, we used these normal distributions just for illustration purposes. In addition, we assumed through the project that all the groups have a common variance (σ in [4]).

4. Priors

For our modeling, we needed to define hyper priors μ_0 and σ_0 . In addition, common σ was defined for the hierarchical model's standard deviation between data points.

Our first goal was to define the hyper prior μ_0 . To aid this problem, we searched information in the internet about country-wise CO2-emissions per capita. The figure below illustrates the results that we found. The figure is taken from <https://www.economicshelp.org/blog/10296/economics/top-co2-polluters-highest-per-capita/> on the 1st of December, 2020.

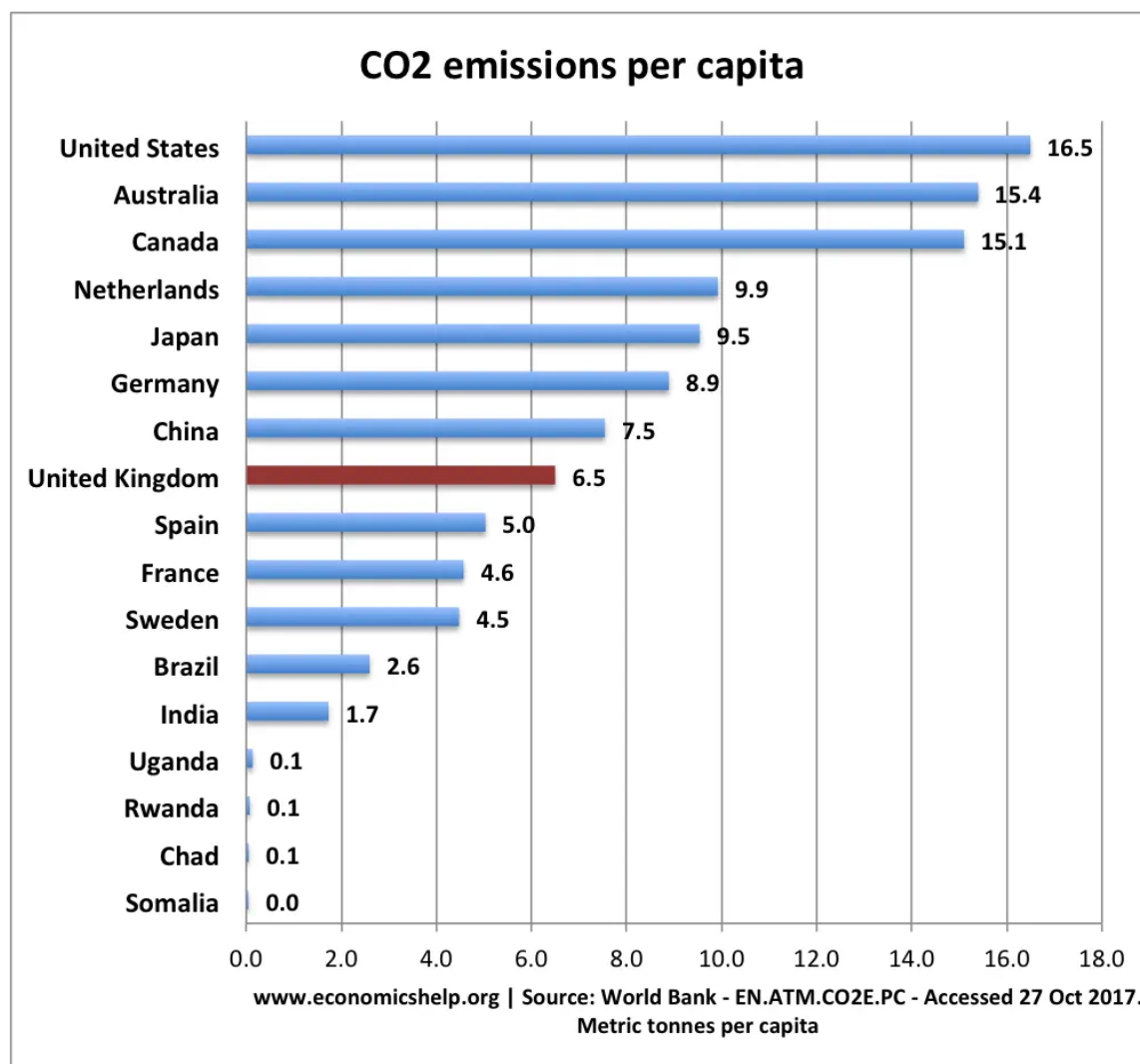


Figure 3: Selected countries CO2 emissions per capita

Source: <https://www.economicshelp.org/blog/10296/economics/top-co2-polluters-highest-per-capita/>
Accessed December 1, 2020.

Results shown in the Figure 3 represent selected countries on year 2016 and are pretty well aligned with the CO2-emissions per capita that we calculated and drew at our data description section. However, it seems

that the underlying dataset is not the same that we decided to use in this project. The idea of searching external information from the internet is to obtain a better understanding about the underlying truth without having to use our own dataset at this point. We want to emphasize that at this point we used our dataset only to estimate whether our own dataset is aligned with the other information found from the internet or not aligned at all.

4.1 Choosing hyper prior mu

With the information that we obtained from investigating the internet more thoroughly, we were able to estimate the order of magnitude for the CO2-emissions per capita. Considering our range of fluctuation, we can first exclude the results below zero. There won't be negative values in CO2-emissions calculations. On the other end, we estimated that the values won't exceed 50 kilo tons per year per country. However, given the information that we were able to obtain, we believe that there is a lot of space for error in value range [0,50]. Thus, we believe that the actual expected mean for hyper prior μ_0 is closer to 0 than 50. We estimated that most of the values should be somewhere between 0 and 30. Therefore, the expected value for μ_0 was placed to 15: $E(\mu_0) = 15$. With the observations presented above, we were finally able to estimate the distribution for μ_0 : $\mu_0 \sim \text{logN}(2.58, 0.5)$ seems to be reasonable. With this distribution and values, μ_0 is strictly restricted to the positive values, has its expected value $E(\mu_0) \approx 15.0$ and the $P(\mu_0 < 50) \approx 1.00$.

4.2 Choosing hyper prior sigma

Next, we had to consider the distribution for the hyper prior σ_0 . Again, restricting the values to positive side only seems reasonable. In our consideration, we prioritized that the probability for σ_0 being zero would be low but on the contrary, the values close to zero would have high probability. Estimating the appropriate tail for the distribution was rather difficult with the information that we were able to obtain. Therefore, we used iterative method to find suitable hyper prior σ_0 . We ended up to a Gamma distribution $\text{Gamma}(\alpha = 2.5, \beta = 0.8)$. This distribution gives $E(\sigma_0) = 1$ and $P(\sigma_0 < 6) \approx 0.99$. This seems reasonable, as we do not want to narrow the μ_j distributions too much with our prior choices.

4.3 Choosing common sigma for observations

Choosing the common sigma for given data points (observations), one can try to answer to a question, how big a jump could the data make between observations n and $n+1$. Given our relatively small value range [0,50] and a fact that the data observations derive from countries' CO2-emissions per capita, we consider the jumps to be only some unit digits at most. For example, with standard deviation of 2, the probability that the "random" jump from n to $n+1$ being greater than 2 is roughly 0.3, which is relatively big probability for such a big jump given circumstances. We believe that the changes within country CO2-emissions per capita tend to be smaller. Therefore, inverse-chi-square distribution with degree of freedom 2.5 seems reasonable. With $\sigma \sim \text{inv} - \text{chi}(2.5)$, all values are positive, expected value is roughly 0.5 and $P(\sigma) < 2 \approx 0.86$. However, this distribution leaves the possibility for the sigma being even higher than 2, which option we want to leave open.

5. Stan

```
data { int <lower=0> N; // number of observations
vector[N] y; // observations }

parameters { real mu; real<lower=0> sigma; }

model { mu ~ lognormal(0, 10); // priors from last week
sigma ~ inv_chi_square(1); // priors from last week
// pooled model likelihood, common mu and sigma for all observations
y ~ normal(mu, sigma); }

generated quantities { real ypred; vector[N] log_lik;
//predictive distribution for any machine
ypred = normal_rng(mu, sigma); }
```



```

for (i in 1:N){
  log_lik[i] = normal_lpdf(y[i] | mu, sigma);
}
}

```

5.2 Hierarchical model

Example code for hierarchical model

```

data { int<lower=0> N; // Number of observations int<lower=0> N_c; // Number of countries vector[N_c]
y[N]; // Observations }

parameters { vector[N_c] mu; // group means real hyper_mu; // prior mean real<lower=0> hyper_sigma;
// prior std constrained to be positive real<lower=0> sigma; // COMMON std constrained to be positive
}

model { hyper_mu ~ normal(0, 100); // weakly informative hyper-prior hyper_sigma ~ inv_chi_square(1);
// weakly informative hyper-prior

mu ~ normal(hyper_mu, hyper_sigma); // population prior with unknown parameters
sigma ~ inv_chi_square(1); // weakly informative prior for group (common) std

for (j in 1:N_c) {
  y[,j] ~ normal(mu[j], sigma); // likelihood
}
}

generated quantities { real y_pred; vector[N_c] log_lik[N];
y_pred = normal_rng(hyper_mu, sigma);

for (j in 1:N_c) {
  for (i in 1:N) {
    log_lik[i, j] = normal_lpdf(y[i,j] | mu[j], sigma);
  }
}
}

```

6. Model running

The non-hierarchical and hierarchical stan models from chapter 5 are compiled and sampled in this section. We will explain the used parameters as the section proceed.

```

df_data <- data.frame(years=seq(1950,2018), data_co2_population)
df_plot <- melt(data = df_data, id.vars = "years", variable.name = "country")
vectored_data_pop <- data.frame(df_plot[, 'value'])
N <- nrow(vectored_data_pop)

num_of_iter <- 1000
num_of_warmup <- 200

pool_data <- list(N = N,
  y = vectored_data_pop[,1])

```

```

pool_model <- rstan::stan_model(file = "pooled_model_stan_without_loglik.stan");

## Running /usr/lib/R/bin/R CMD SHLIB foo.c
## clang -flto=thin -std=gnu99 -I"/usr/share/R/include" -DNDEBUG -I"/usr/local/lib/R/site-library/Rcpp
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/StanHeaders/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:88:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:1: error: unknown t
## namespace Eigen {
## ^
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:16: error: expected
## namespace Eigen {
## ^
## ;
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/StanHeaders/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:96:10: fatal error: 'complex' file not fo
## #include <complex>
## ^~~~~~
## 3 errors generated.
## make: *** [/usr/lib/R/etc/Makeconf:168: foo.o] Error 1

pool_fit <- rstan::sampling(object = pool_model,
                           data = pool_data,
                           iter = num_of_iter,
                           warmup = num_of_warmup,
                           refresh = 0)

```

We need the total number of observations to be able to run the pooled model, and it's saved to variable `N`. Vectored version of data is also required by the pooled model. We chose 1000 as the number of iterations per chain, as it has also worked relatively reliably in previous work on the course. We used one fifth of the iterations in the warm-up sample to ensure that the chains were close to the maximum probability mass when true iterations start. At this point, the model without logarithmic likelihood is used to make code compiling faster. Lots of more information about function `stan::stan_model` and `stan::sampling` is found from RStan documentation.

```

hier_data <- list(N = nrow(data_co2_population),
                 N_c = ncol(data_co2_population),
                 y = data_co2_population)

hier_model <- rstan::stan_model(file = "hier_model_stan_without_loglik.stan");

## Running /usr/lib/R/bin/R CMD SHLIB foo.c
## clang -flto=thin -std=gnu99 -I"/usr/share/R/include" -DNDEBUG -I"/usr/local/lib/R/site-library/Rcpp
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/StanHeaders/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:88:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:1: error: unknown t
## namespace Eigen {
## ^
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:16: error: expected

```

```
## namespace Eigen {
##      ^
##      ;
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/stan/math/prim/mat/fun/Eigen:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:96:10: fatal error: 'complex' file not found
## #include <complex>
##      ^~~~~~
## 3 errors generated.
## make: *** [/usr/lib/R/etc/Makeconf:168: foo.o] Error 1
```

```
hier_fit <- rstan::sampling(object = hier_model,
                           data = hier_data,
                           iter = num_of_iter,
                           warmup = num_of_warmup,
                           refresh = 0)
```

When running the hierarchical model, the CO₂ data is given in matrix form. The number of iterations and the number of warm-ups is the same as in the pooled model presented above. One group is the one country in this model, so the number of groups is the same as the number of columns in data.

7. Convergence diagnostics

We can inspect the convergence of chains using, for example, *potential scale reducing factor* \hat{R} and *effective sample size* (ESS). The first of these, \hat{R} , examines stationarity and mixing of chains. Correspondingly, the effective sample size takes into account the autocorrelation between the samples in a chain. More information about mathematics of these diagnostics can be found in [<https://arxiv.org/pdf/1903.08008.pdf>]. Let's start by monitoring the results with *monitor* function, which also reveals the convergence quantities of chains.

```
monitor(pool_fit)
```

```
## Inference for the input samples (4 chains: each with iter = 1000; warmup = 0):
##
##           Q5      Q50      Q95      Mean  SD   Rhat Bulk_ESS Tail_ESS
## mu         5.0      5.2      5.4      5.2 0.1    1     1901     1284
## sigma      4.9      5.1      5.2      5.1 0.1    1     3033     2148
## ypred     -3.0      5.2     13.3      5.2 5.1    1     3354     3106
## lp__    -2793.0 -2790.9 -2790.3 -2791.2 0.9    1      1454     1699
##
## For each parameter, Bulk_ESS and Tail_ESS are crude measures of
## effective sample size for bulk and tail quantities respectively (an ESS > 100
## per chain is considered good), and Rhat is the potential scale reduction
## factor on rank normalized split chains (at convergence, Rhat <= 1.05).
```

```
monitor(hier_fit)
```

```
## Inference for the input samples (4 chains: each with iter = 1000; warmup = 0):
##
##           Q5      Q50      Q95      Mean  SD   Rhat Bulk_ESS Tail_ESS
## mu[1]         1.3      1.8      2.2      1.8 0.3    1     6708     2402
## mu[2]         0.7      1.2      1.6      1.2 0.3    1     5879     2193
## mu[3]         1.9      2.4      2.8      2.4 0.3    1     4936     2505
## mu[4]        18.2     18.7     19.1     18.7 0.3    1     6758     2549
## mu[5]         7.9      8.3      8.8      8.3 0.3    1     6067     2552
## mu[6]         8.2      8.7      9.2      8.7 0.3    1     6475     2180
```

```
## mu[7]          6.5      7.0      7.5      7.0 0.3      1      3597      2110
## mu[8]          5.0      5.5      5.9      5.5 0.3      1      7951      2490
## mu[9]          0.4      0.9      1.4      0.9 0.3      1      5004      2447
## mu[10]         7.3      7.7      8.2      7.7 0.3      1      5760      2471
## mu[11]        -0.2      0.2      0.7      0.2 0.3      1      3835      2417
## mu[12]        10.5     11.0     11.5     11.0 0.3      1      7063      2637
## mu[13]         0.3      0.7      1.2      0.7 0.3      1      4902      2426
## mu[14]         3.3      3.7      4.2      3.7 0.3      1      5595      2487
## mu[15]         3.4      3.8      4.3      3.8 0.3      1      4923      2115
## mu[16]         6.7      7.2      7.6      7.2 0.3      1      4828      2355
## mu[17]         2.0      2.4      2.9      2.4 0.3      1      4010      2254
## mu[18]         6.1      6.5      7.0      6.5 0.3      1      5045      2499
## mu[19]         0.6      1.1      1.5      1.1 0.3      1      6720      1751
## hyper_mu       4.1      5.7      7.4      5.7 1.0      1      4431      2526
## hyper_sigma    3.6      4.5      6.0      4.7 0.8      1      4325      2520
## sigma          2.3      2.4      2.5      2.4 0.0      1      6670      2200
## y_pred_new_county -2.1      5.9     13.5      5.8 4.8      1      3206      3133
## y_pred_SAU       7.1     11.0     14.8     11.0 2.4      1      3253      2961
## y_pred_FIN       4.7      8.6     12.6      8.7 2.4      1      3075      3195
## lp__           -1847.5 -1841.6 -1837.1 -1841.9 3.3      1      1155      2118
##
## For each parameter, Bulk_ESS and Tail_ESS are crude measures of
## effective sample size for bulk and tail quantities respectively (an ESS > 100
## per chain is considered good), and Rhat is the potential scale reduction
## factor on rank normalized split chains (at convergence, Rhat <= 1.05).
```

First of all, we can see that \hat{R} s for both models are 1 or 1.01, which indicates that the chains are fully converged with a high probability. We can deduce the same fact by inspecting the Tail_ESS, which are over 2000 for all the variables under consideration. So by looking at these convergence diagnostics, we couldn't spot any convergence problems of Monte-Carlo chains.

The convergence of the model simulations was sufficient already on the first try as we used the 2000 iteration with 1000 warm-up samples. With the try-and-error method, we were able to deduce that there was no need for over 1000 iterations, which significantly reduced simulations' execution time. The number of warm-up samples (200) is also experimental. We noticed that even 50 warm-up samples could be sufficient to produce good enough convergence after the warm-up period, but then there could also be an opportunity for the worse convergence. However, the use of half of the samples as warm-ups seemed to be too large in this application.

8. Posterior predictive checks

9. Model comparison

```
pool_model_loglik <- rstan::stan_model(file = "pooled_model_stan.stan")

## Trying to compile a simple C file

## Running /usr/lib/R/bin/R CMD SHLIB foo.c
## clang -flto=thin -std=gnu99 -I"/usr/share/R/include" -DNDEBUG -I"/usr/local/lib/R/site-library/RcppEigen/include"
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/stan/math/prim/mat/fun/Eigen:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:88:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:1: error: unknown token
## namespace Eigen {
```

```

## ^
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:16: error: expected
## namespace Eigen {
## ^
## ;
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/stan/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:96:10: fatal error: 'complex' file not fo
## #include <complex>
## ^~~~~~
## 3 errors generated.
## make: *** [/usr/lib/R/etc/Makeconf:168: foo.o] Error 1

pool_fit_loglik <- rstan::sampling(object = pool_model_loglik,
                                data = pool_data,
                                iter = num_of_iter,
                                warmup = num_of_warmup,
                                refresh = 0)

hier_model_loglik <- rstan::stan_model(file = "hierarchical_model_stan.stan")

## Trying to compile a simple C file

## Running /usr/lib/R/bin/R CMD SHLIB foo.c
## clang -flto=thin -std=gnu99 -I"/usr/share/R/include" -DNDEBUG -I"/usr/local/lib/R/site-library/Rcpp
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/stan/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:88:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:1: error: unknown t
## namespace Eigen {
## ^
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/src/Core/util/Macros.h:613:16: error: expected
## namespace Eigen {
## ^
## ;
## In file included from <built-in>:1:
## In file included from /usr/local/lib/R/site-library/StanHeaders/include/stan/math/prim/mat/fun/Eigen
## In file included from /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Dense:1:
## /usr/local/lib/R/site-library/RcppEigen/include/Eigen/Core:96:10: fatal error: 'complex' file not fo
## #include <complex>
## ^~~~~~
## 3 errors generated.
## make: *** [/usr/lib/R/etc/Makeconf:168: foo.o] Error 1

hier_fit_loglik <- rstan::sampling(object = hier_model_loglik,
                                data = hier_data,
                                iter = num_of_iter,
                                warmup = num_of_warmup,
                                refresh = 0)

log_lik_pooled <- extract_log_lik(pool_fit_loglik, merge_chains = FALSE)
r_eff_pooled <- relative_eff(exp(log_lik_pooled), cores=4)
loo_pooled <- loo(log_lik_pooled, r_eff = r_eff_pooled, cores = 4)
print(loo_pooled)

```

```

##
## Computed from 3200 by 1311 log-likelihood matrix
##
##           Estimate   SE
## elpd_loo  -3993.9 32.3
## p_loo      2.7  0.2
## looic      7987.9 64.6
## -----
## Monte Carlo SE of elpd_loo is 0.0.
##
## All Pareto k estimates are good (k < 0.5).
## See help('pareto-k-diagnostic') for details.
pooled_plotters <- function() {
  pooled_df =data.frame(rstan::extract(pool_fit, permuted=T))

  #Histogram
  hist(pooled_df$mu,
        breaks = 100,
        xlim=c(0,22),
        xlab = "Mean of the quality measurements",
        col = "lightyellow",
        main="Posterior distribution of the mean of the sixth machine")
}

hierarchical_plotters <- function() {
  hierarchical_df =data.frame(rstan::extract(hier_fit, permuted=T))

  #MCMC Areas
  mcmc_areas_df <- hierarchical_df %>% select(starts_with('mu')) %>%
    setNames(colnames(data_co2_population))
  mcmc_areas(mcmc_areas_df) + xlab("Testttttttt")

  #Histograms of countries together
  m <- 19
  {
    plot(0,0,type="n",
          xlim=c(0,20), ylim=c(0,1100),
          xlab="x",ylab="freq",
          main="Histograms of each country separately, plotted together")
    for(n in 1:m) {
      var_name <- paste("mu.",n, sep="")
      #hier_matrix[n,] <- unlist(hierarchical_df[var_name])
      plot(
        hist(unlist(hierarchical_df[var_name]), breaks = 12, plot=FALSE),
        col=alpha('blue', 0.25),
        add=T # Add to main plot
      )
    }
  }

  #One histogram for whole data
  one_hist_data <- unlist(hierarchical_df %>% select(starts_with('mu')))

```

```
plot(hist(one_hist_data, breaks = 100, xlim = c(0,22), ylim = c(0, 5000)),  
      col = 'lightblue')  
}
```

References