## REGRESSIONS

Correlations for the Advanced?



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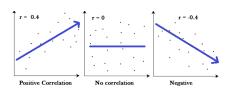
01/04/2020

- 1 The Basics
  - Correlation Tests
  - Regression Models
  - Least Squares vs. Maximum Likelihood
- 2 Methods & Models
  - Single Linear Regression
  - Mixed Effect Models
  - Generalised Linear Models
- 3 Choosing the Right Method

## Correlation is **not** necessarily **causation** (spurious correlations).

Correlation tests yield two

- r value (measure of correlation)
  - ightharpoonup r pprox 1 (strong, positive correlation)
  - $r \approx 0$  (no correlation)
  - $r \approx -1$  (strong, negative correlation)
- p value (measure of statistica significance)



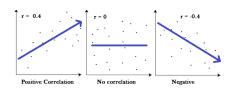
→ Get a feeling for it on Guess The Correlation.

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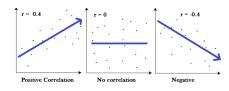
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# Types of Correlations

These approaches are extremely useful in data exploration and for preliminary analyses!

Prominent correlation tests include

- Contingency Coefficient
- Kendall's Tar
- Spearman Correlation
- Pearson Correlation
- Cramer's \
- ANalysis Of VAriance (ANOVA)
- . . . .

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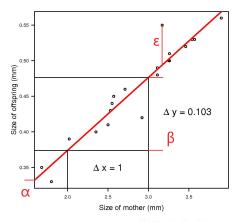
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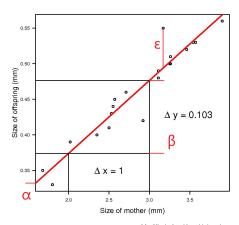


- $\alpha$  The **Intercept**. The value of y when x = 0 (also referred to as  $\beta_0$ ).
- $\beta_i$  The **Correlation Coefficient** The increase in y for a one-unit increase in dependent variable i (usually, x if only one dependent variable).
- $\epsilon$  The **Random Error**. The deviation of data points from the regression line. Usually assumed to follow  $\epsilon \sim N(0, \sigma^2)$



Modified after Knut Helge Jensen.

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Modified after Knut Helge Jensen.

# Assumptions in Theory

Linear regression models need to be inspected for violations of assumptions after regressing:

### Residuals vs. Fitted values

Non-linear patterns identify a non-linear relationship between dependent and independent variables.

## Normal Q-Q plot

Non-normal distribution of residuals shows that the assumption of  $\epsilon \sim N(0, \sigma^2)$  is violated.

### Scale Location

Non-constant variance identifies show that the assumption of homoscedasticity (invoked by least squares fitting).

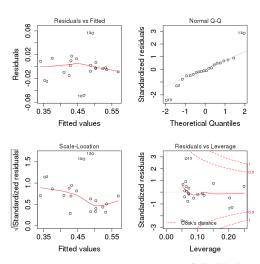
## Residuals vs. Leverage

A non-zero trend identifies the presence of influential outliers.

6 / 29

# Assumptions in R

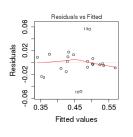
- Simply type 'plot(...)' with '...' denoting your regression model.
- You can also target individual plots by writing:
  - 'plot(..., 1)' for Residuals vs. Fitted values
  - 'plot(..., 2)' for Normal Q-Q plot
  - 'plot(..., 3)' for Scale Location
  - 'plot(..., 4)' for Residuals vs. Leverage

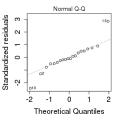


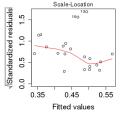
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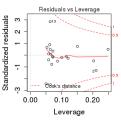
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# Types of Regressions

Less model variables result in a more interpretable model!

# Prominent regression approaches include the following:

- Single Linear Regression
- Multiple Linear Regression
- Linear Mixed Effect Models
- Generalized Linear Models

- Polynomial Regressions
- Generalized Additive Models
- Regression Splines
- Smoothing Splines
- Local Regressions
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# Least Squares vs. Maximum Likelihood

These methods refer to parameter estimation.

## Ordinary Least Squares (OSL):

- Used for most basic linear regressions
- obtain coefficient estimates  $\hat{\beta}_0$  and  $\hat{\beta}_1$  such that the linear model fits the available data well that is, so that  $\hat{y}_i \approx \hat{\beta}_0 + \hat{\beta}_1 x_i$  for i = 1, ..., n.

#### Minimize:

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$
 (1)

with  $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ 

## Maximum Likelihood Estimation (MLE):

- Used in logistic regressions and generalized linear models
- estimates for  $\beta_0$  and  $\beta_1$  such that the predicted probability  $\hat{Pr}(x_j)$  corresponds to the observed response variable status.

## Maximize:

$$\ell(\beta_0,\beta_1) = \prod_{i:y_i=1} Pr(x_i) \prod_{i':y_{i'}=1} (1 - Pr(x_{i'})) \quad \ (2)$$

# Purpose & Assumptions

## Single linear regression

lm() in base R

## Purpose: Identify whether and how two variables are related.

- Down to *Study-Design*:
  - Predictor variable is continuous (ratio or interval scale)
  - Response variable is continuous (ratio or interval scale)
  - Variable values are **independent** (not paired)
- Need for Post-Hoc Tests
  - Variable values follow homoscedasticity (equal variance across entire data range)
  - Residuals follow normal distribution (normality)
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Assumptions:

# Example - The Data

```
# measures of Diameter (labelled as Girth), Height, and Volume of Timber
data("trees")
head(trees)
## Girth Height Volume
```

```
## 1 8.3 70 10.3

## 2 8.6 65 10.3

## 3 8.8 63 10.2

## 4 10.5 72 16.4

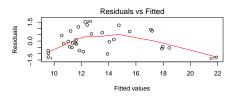
## 5 10.7 81 18.8

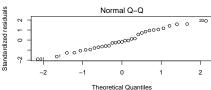
## 6 10.8 83 19.7
```

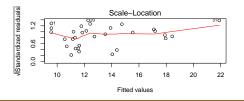
→ Let's see if there is a good regression to be had between *Girth* and *Volume*.

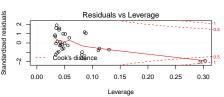
# Example - The Model

```
SingleLin_Mod <- with(trees, lm(Girth ~ Volume))
par(mfrow=c(2,2))
plot(SingleLin_Mod)</pre>
```





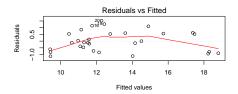


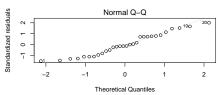


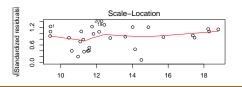
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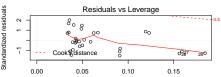
# Example - Refining The Model

```
trees <- trees[-31,] # removing the influential otulier in row 31
SingleLin_Mod <- with(trees, lm(Girth ~ Volume))
par(mfrow=c(2,2))
plot(SingleLin_Mod)</pre>
```









# Example - Model Output

summary (SingleLin\_Mod)

```
##
## Call:
## lm(formula = Girth ~ Volume)
##
## Residuals:
     Min 10 Median 30
                             Max
## -1.126 -0.699 -0.109 0.557 1.521
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 7.4141 0.3217 23.0 <2e-16 ***
## Volume
          0.1954 0.0101 19.3 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.772 on 28 degrees of freedom
## Multiple R-squared: 0.93, Adjusted R-squared: 0.928
## F-statistic: 374 on 1 and 28 DF, p-value: <2e-16
```

At a Volume of 0, Girth is predicted to be 7.4141 (of course that doesn't make sense, not only is a volume of 0 biological nonsense, height also plays a part here for sure). For a one-unit increase in Volume, Girth is predicted to go up by 0.1954 inches (yes, they recorded in inches). Both estimates are statistically significant.

# Purpose & Assumptions

## Linear mixed effect model

lme() in base nlme package

Purpose: Identify whether and how two variables are related.

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Assumptions:

# Fixed effects and random effects are also referred to as fixed effect factors and random effect factors.

## Fixed Effects

- Informative factor levels regarding hypothesis.
- Want to study these levels and their effects.
- Factor levels are deliberate part of the study-design.
- Higher sample size ≠ higher number of levels.

### Random Effects

- Uninformative factor levels regarding hypothesis.
- Do not want to study these levels and their effects.
- Factor levels are imposed by nature/type of study.
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# Example - The Data

## 5 76 8 1 93

1.0

```
# measures of Weight, Diet, Time, and Chicks
data ("ChickWeight")
head (ChickWeight)
##
    weight Time Chick Diet
## 1
       42
    51 2 1
## 3 59 4 1
```

→ Let's see if there is a good regression to be had between *weight* and *Time* while accounting for random effects belonging to Chick, and fixed effects of Diet.

## Example - The Model

We now have our model. However, we know that time is a component and we likely have repeated samples here. In these cases, we need to account for auto-correlation by defining a correlation structure.

Let's see which model (basic or the one with auto-correlative structure) performs better:

```
## MultiLin_Mod 2 11 4457 4505 -2217 1 vs 2 1032 <.0001
```

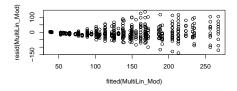
We clearly prefer the more sophisticated, auto-correlative model and want to see which of its parameters are informative:

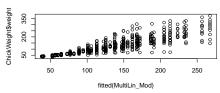
anova (MultiLin Mod) # all parameters should be kept

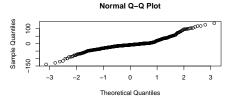
We keep all parameters. Although the inclusion of Diet is not significant, the interaction of Diet and Time is, therefore, both Time and Diet need to stay irrespective of their significance.

# Example - Assessing the Model

```
par(mfrow=c(2,2))
plot(fitted(MultiLin_Mod), resid(MultiLin_Mod)) # values around 0 -> good
plot(fitted(MultiLin_Mod), ChickWeight$weight) # pattern fuzzy, but linear -> good
qqnorm(resid(MultiLin_Mod)) # residuals are not normal dsitributed -> bad
```







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# Example - Model Output

```
summary (MultiLin Mod)
## Linear mixed-effects model fit by REML
  Data: ChickWeight
   AIC BIC logLik
  4457 4505 -2217
##
## Random effects:
  Formula: ~+1 | Chick
         (Intercept) Residual
## StdDev: 0.006581
                      42.46
##
## Correlation Structure: AR(1)
## Formula: ~1 | Chick
  Parameter estimate(s):
   Phi
## 0.9706
## Fixed effects: weight ~ Time * Diet
##
            Value Std.Error DF t-value p-value
## (Intercept) 40.42 9.487 524
                               4.260 0.0000
## Time
            6.06 0.350 524
                               17.347 0.0000
## Diet2 -0.88 16.430 46 -0.054 0.9575
        -2.19 16.430 46 -0.133 0.8944
## Diet3
        -1.01 16.431 46 -0.062 0.9512
## Diet4
## Time:Diet2 2.21 0.589 524 3.749 0.0002
## Time:Diet3 4.86 0.589 524 8.250 0.0000
## Time:Diet4 3.13 0.592 524 5.297 0.0000
## Correlation:
##
            (Intr) Time Diet2 Diet3 Diet4 Tm:Dt2 Tm:Dt3
```

## Time

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# Example - Model Output Explained

Variance between chicks (0.006581) than residual variance (42.46). This is good!

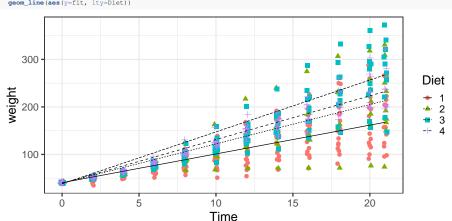
A chick is expected to have a weight of 40.4165 at Time = 0 and Diet = 1. Per time-step, weight is expected to increase by 6.0637.

Mean chick weight is different to (read: "Diet1 weights are smaller by") Diet = 1 by -0.8794, -2.1932, and -1.0109 for Diet = 2, Diet = 3, and Diet = 4 respectively (take note that these differences aren't statistically significant).

weight of chicks increases, on average, increases by 2.2077 units more per one-unit increase in time when compared to Diet=1. The same logic applies to Diet=3, and Diet=4.

# Example - Model Output Visualised

```
library(ggplot2)
ChickWeightSfit <- predict(MultiLin_Mod, level=0)
ggplot(ChickWeight, aes(Time, weight)) +
geom_jitter(aes(colour=Diet, shape=Diet), width=0.1, size=3) +
theme_bw(base_size=20) +
geom_line(aes(y=fit, lty=Diet))</pre>
```



# Purpose & Assumptions

### **Generalized Linear Models**

glm() in base R

Purpose: Identify whether and how two variables are related.

- Down to Study-Design:
  - Variable values are **independent** (not paired)
- Assumptions: Need for Post-Hoc Tests
  - Absence of influential outliers
  - Response and Predictor are related in a linear fashion
  - → Allow for non-normal distributions and heteroscedasticity

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# Linear Predictor, Link Function, and Variance Function

## **Components** of a Generalized Linear Model:

- **Linear predictor** e.g.:  $y_i = \alpha + \beta_1 + x_i$
- Link function  $g(\hat{y_i}) = y_i$ Relationship between predictor value and estimated value.
- **3** Variance function  $var(y_i) = \phi V_i(\overline{x})$ Variance depends on predictor mean, dispersion parameter  $\phi$  is a constant

Error	Link function	Variance function	Typical type of data
normal	identity	1 (constant)	Textbook examples
Poisson	log		Count data
binomial	$logit, log(\overline{x}/(1-\overline{x}))$	$var = \overline{x}(1 - \overline{x})/n$	Binary data

Aarhus University Biostatistics - Why? What? How? 24 / 29

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## Which combinations of components do I use?

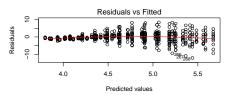
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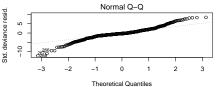
# Example - The Data

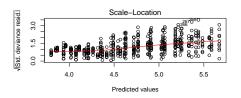
→ Let's reassess our earlier analysis of chick weight as a function of time and diet. This time, we will forego the random effect of chicks since that would create a generalized linear mixed effect model.

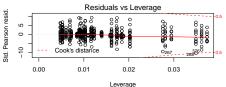
# Example - The Model

```
GeneralLin_Mod <- glm(weight ~ Time*Diet, family = poisson, data = ChickWe
par(mfrow=c(2,2))
plot(GeneralLin_Mod)</pre>
```









# Example - Model Output

```
summary (GeneralLin_Mod)
##
## Call:
## qlm(formula = vs ~ wt + disp, family = binomial, data = mtcars)
##
## Deviance Residuals:
     Min
             10 Median 30
                                 Max
## -1.675 -0.284 -0.084 0.573 2.082
##
## Coefficients.
##
             Estimate Std. Error z value Pr(>|z|)
## (Intercept) 1.6086 2.4390 0.66
                                         0.510
              1.6264 1.4907 1.09 0.275
        -0.0344 0.0154 -2.24 0.025 *
## disp
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
      Null deviance: 43.86 on 31 degrees of freedom
## Residual deviance: 21.40 on 29 degrees of freedom
## AIC: 27.4
##
## Number of Fisher Scoring iterations: 6
```

As you can see, the estimates to earlier have changed drastically.

# Example - Model Output Explained

A chick is now expected to have a weight of 3.7988 at Time = 0 and Diet = 1. Per time-step, weight is expected to increase by 0.0692.

Mean chick weight of Diet=1 is smaller by 0.0686, 0.0506, and 0.1616 for Diet=2, Diet=3, and Diet=4 respectively (these differences are now significant).

weight of chicks increases, on average, increases by 0.0056 units more per one-unit increase in time when compared to Diet=1. The same logic applies to Diet=3, and Diet=4.

# Choices, Choices, Choices...

- Linear Model lm. When all assumptions are met (i.e.: homoscedasticity, normality, independence).
- Linear Mixed Effect Model lme. When the assumption of independence is violated.
- Generalized Linear Model glm. When the assumptions of homoscedasticity and homoscedasticity are violated.
- Generalized Linear Mixed Effect Model glmmPQL from MASS, or glmer from lme4. When the assumptions of homoscedasticity, homoscedasticity, and independence are violated.