```
% TPPE32, Lab2
% William Eriksson & Stina Östklint
assets = flip(readmatrix('timeSeries.xlsx', 'Sheet', 'Problem 1 and 2', 'Range', 'C3:G1512'));
% Task 1 a), Variance-Covariance Method
numColumns = size(assets, 2);
numRow = size(assets, 1);
w = 1 / numColumns;
w_matrix = repmat(w, numColumns,1);
returns = zeros(numRow-1,numColumns);
for row = 1: (numRow-1)
  for col = 1:(numColumns)
      returns(row,col) = (assets(row+1, col) - assets(row,col)) / assets(row,col);
   end
end
mu = mean(returns);
cov_matrix = cov(returns);
sigma = sqrt(transpose(w_matrix) * cov_matrix * w_matrix);
port_value = assets(numRow, 1:numColumns)*w_matrix;
VaR95 = (-mu + norminv(0.95)*sigma)*port value *w matrix
VaR95 = 5.7401
VaR99 = (-mu + norminv(0.99)*sigma)*port_value *w_matrix
VaR99 = 8.1705
```

The Value at Risk can be showed above.

```
L = - returns(numRow-1, 1:numColumns) * w_matrix * port_value

L = 0.1188

ES95 = mean(L(L > VaR95))

ES95 = NaN
```

```
ES99 = mean(L(L > VaR99))
```

ES99 = NaN

Since L is smaller than both VaR95 and VaR99 the answers to Expected Shortfall is not defined.

```
% Task 1 b), Monte Carlo Simulation
%i)
log returns = zeros(numRow-1,numColumns);
for row = 1: (numRow-1)
   for col = 1:(numColumns)
       log_returns(row,col) = log(assets(row+1, col)/ assets(row,col));
   end
end
% Optimization of GARCH with fmincon, without variance targeting
% Initial parameter guesses
initial_params = [0.1, 0.1, 0.8];
% Constraints: (0 <= omega, alpha, beta <= 1 and alpha + beta <= 1)</pre>
1b = [0, 0, 0];
ub = [1, 1, 1];
A = [];
b = [];
Aeq = [];
beq = [];
nonlcon = @(params) deal([], params(1) + params(2) - 1); % alpha + beta <= 1
% Optimization options
options = optimoptions('fmincon', 'Display', 'off');
parameters = zeros(3,numColumns);
variance garch = zeros(numRow-1,numColumns);
for i = 1:numColumns
   [params_opt, neg_LL] = fmincon(@(params) garch_likelihood(params, log_returns(1:(numRow-1))
   parameters(1:length(params_opt),i) = params_opt;
   [11 ,variance_GARCH] = garch_likelihood(params_opt, log_returns(1:(numRow-1),i), "false");
   variance garch(1:numRow-1,i) = variance GARCH;
end
parameters;
variance_garch;
```

```
%ii)
standard_returns = zeros(numRow-1,numColumns);
for col = 1:numColumns
    for row = 1:numRow-1
        standard_returns(row,col) = log_returns(row, col) / sqrt(variance_garch(row, col));
    end
end
standard_returns; %Standard returns equals the normally distirbuted residuals
U = zeros(numRow-1,numColumns);
for col = 1:numColumns
    U(1:numRow-1, col) = normcdf(standard_returns(1:numRow-1,col));
end
U;
[Rho_t, df] = copulafit('t', U);
%iii)
noOfSimulations = 1000;
rand_data = copularnd("t", Rho_t, df, noOfSimulations);
%iv)
residuals = norminv(rand_data);
rand_returns = residuals * transpose(sqrt(variance_garch(numRow-1,1:numColumns)));
rand_mu = mean(rand_returns);
rand sigma= sqrt(var(rand returns));
VaR95_rand = (-rand_mu + norminv(0.95)*sigma)*port_value
VaR95\_rand = 6.9701
VaR99_rand = (-rand_mu + norminv(0.99)*sigma)*port_value
VaR99 rand = 9.4006
L rand = - rand returns(1:noOfSimulations,1)*port value;
ES95_rand = mean(L_rand(L_rand > VaR95_rand))
ES95_rand = 15.6224
ES99_rand = mean(L_rand(L_rand > VaR99_rand))
```

 $ES99_rand = 17.5453$

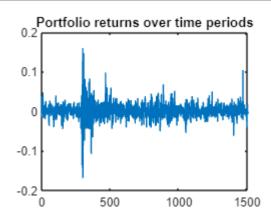
```
% Task 1 c), Extreme Value Theory
%i) MLE
portfolio returns = log returns * w matrix;
threshold = -quantile(portfolio_returns, 0.05)
threshold = 0.0284
excess_losses = -portfolio_returns(portfolio_returns < -threshold) - threshold;</pre>
initial_params = [0.1, 1]; % Initial guesses for xi and beta
lb = [-Inf, 0]; % Lower bounds (beta > 0)
ub = [Inf, Inf]; % Upper bounds
A = [];
b = [];
Aeq = [];
beq = [];
nonlcon = [];
options = optimoptions('fmincon', 'Display', 'off');
[params opt EVT] = fmincon(@(params) EVT log likelihood(params, excess losses), initial params
epsilon = params_opt_EVT(1)
epsilon = 0.3075
beta = params_opt_EVT(2)
beta = 0.0120
excess_rate = length(portfolio_returns)/length(excess_losses)
excess_rate = 20.1200
VaR_99_1d_EVT = (threshold + (beta/ epsilon) * (excess_rate * (1 - 0.99)^(epsilon) - 1))
VaR_{99}1d_{EVT} = 0.1800
ES_99_1d_EVT = (VaR_99_1d_EVT + beta - epsilon * threshold) / (1 - epsilon)
ES_{99_1d_EVT} = 0.2647
```

The value of VaR and the constituent quantities can be shows in the display.

```
%ii) 500 observations
newdata = portfolio_returns(250:750);
threshold = -quantile(newdata, 0.05)
```

threshold = 0.0369

```
New_excess_losses = -newdata(newdata < -threshold) - threshold;
figure;
plot(portfolio_returns) %Seem to be large fluctuations in returns from data points 250 to 750
title("Portfolio returns over time periods")</pre>
```



```
% data_points = portfolio_returns(250:750);
initial_params = [0.1, 1]; % Initial guesses for xi and beta
lb = [-Inf, 0]; % Lower bounds (beta > 0)
ub = [Inf, Inf]; % Upper bounds
A = [];
b = [];
Aeq = [];
beq = [];
nonlcon = [];
options = optimoptions('fmincon', 'Display', 'off');

[params_opt_EVT_2, neg_LL_EVT] = fmincon(@(params) EVT_log_likelihood(params, New_excess_losses
epsilon_500 = params_opt_EVT_2(1)
```

 $epsilon_500 = 0.3882$

```
beta_500 = params_opt_EVT_2(2)
```

 $beta_500 = 0.0186$

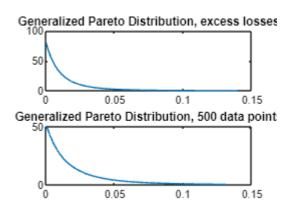
```
%Plot the distribution functions
x1 = linspace(min(excess_losses), max(excess_losses), 100);
```

```
density1 = gpd_pdf(x1, epsilon, beta);

% Calculate the density for the 500 data points using the estimated parameters
x2 = linspace(min(New_excess_losses), max(New_excess_losses), 100);
density2 = gpd_pdf(x2, epsilon_500, beta_500);

figure;
subplot(2, 1, 1);
plot(x1, density1)
title('Generalized Pareto Distribution, excess losses')

subplot(2, 1, 2);
plot(x2, density2);
title('Generalized Pareto Distribution, 500 data points')
```



```
%VaR99 is normally distributed, VaR99_rand is Monte Carlo Simulated Relative_VaR = VaR_99_1d_EVT/VaR99
```

Relative VaR = 0.0220

The distribution functions shows that for VaR 99% the inital excess losses are larger for a lower percentage but it decreases faster than for the 500 data points. This may be due to the 500 data point interval being chosen as a volatile one.

```
ES_RW = zeros(1, length(confidence_levels));
failures = zeros(1, length(confidence_levels));
port_return = zeros(N, 1);
hist_return = zeros(window_size,1);
for t = window size+2:(N)
     port_return(t,1) = log_returns(t, :) * portfolio_weights(t-1,:)';
    %Update portfolio weights
    for i = 1:numColumns
        portfolio_weights(t,i) = portfolio_weights(t-1,i) * (1 + log_returns(t, i)) / (1+port_
    end
end
for t = window_size+2:(N)
    hist return = log_returns(t-window_size:t-1,:) * portfolio_weights(t-1,:)';
    % Calculate VaR and ES
    sorted returns = sort(hist return);
    for cl = 1:length(confidence_levels)
        alpha = 1 - confidence_levels(cl);
        VaR_RW(t-window_size, cl) = -prctile(sorted_returns , alpha * 100);
        ES_RW(t-window_size, cl) = -mean(sorted_returns(sorted_returns <= -VaR_RW(t-window_size
    end
   % Compare realized returns with VaR
    for cl = 1:length(confidence_levels)
        failures(t-window_size, cl) = port_return(t,1) < -VaR_RW(t-window_size, cl);</pre>
    end
end
```

```
else
    disp('Fail to reject null hypothesis: Failure rate is as expected.');
end
```

Reject null hypothesis: Failure rate is significantly different from expected.

```
disp(['Observed failure rate: ', num2str(observed_failure_rate)]);
Observed failure rate: 0.030723
```

```
disp(['Expected failure rate: ', num2str(expected_failure_rate)]);
```

Expected failure rate: 0.05

```
disp(['p-value: ', num2str(p_value)]);
```

p-value: 0.0018135

The p-value for the binomial cdf for the failure rate with 95% confidence is within the p-value interval. Therefore we reject the null hypothesis, in other words the failure rate is significantly different from the expected one. This implies that the VaR model is underestimating risk.

```
% Task 2b, VaR Serial dependency (Christoffersen)
% Binary sequence of VaR breaches
VaR breaches = failures 95;
% Christoffersen test for serial independence
%1 = breach
%0 = no breach
n11 = sum(VaR_breaches(1:end-1) == 1 & VaR_breaches(2:end) == 1); %Number of times a breach is
n10 = sum(VaR breaches(1:end-1) == 1 & VaR breaches(2:end) == 0); %Number of times a breach is
n01 = sum(VaR_breaches(1:end-1) == 0 & VaR_breaches(2:end) == 1); %Number of times a no breach
n00 = sum(VaR breaches(1:end-1) == 0 & VaR breaches(2:end) == 0); %Number of times a no breach
% Likelihood ratio test statistic
% Measures Null model (first log) versus Alternative model (second log)
%Lnull is the likelihood under the null hypothesis (independent breaches),
%Lalternative is the likelihood of dependent breaches.
LR = -2 * log((1 - observed_failure_rate)^(n00 + n10) * observed_failure_rate^(n01 + n11)) + 2
```

```
LR = 3.0747
```

```
% From lecture 5
critical_value = 1.64
```

```
critical_value = 1.6400
```

```
% Compare test statistic to critical value
if LR > critical_value
    disp('Reject null hypothesis: VaR breaches are serially dependent.');
else
    disp('Fail to reject null hypothesis: VaR breaches are serially independent.');
end
```

Reject null hypothesis: VaR breaches are serially dependent.

```
disp(['Likelihood ratio test statistic: ', num2str(LR)]);
```

Likelihood ratio test statistic: 3.0747

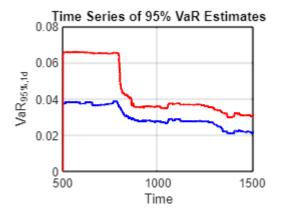
```
disp(['Critical value: ', num2str(critical_value)]);
```

Critical value: 1.64

Since the Likelihood ratio is larger than the critical value for 95% confidence interval we reject the null hypothesis. In other word the VaR breaches are serially dependent.

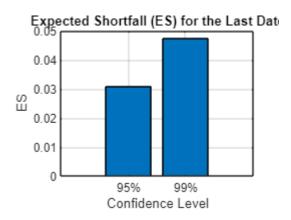
```
%Task 2, plots

% Plot time series of 95% VaR
figure;
plot(window_size+1:N, VaR_RW(1:N-window_size, 1), 'b', 'LineWidth', 1.5);
hold on;
plot(window_size+1:N, ES_RW(1:N-window_size, 1), 'r', 'LineWidth', 1.5); % Plot 95% ES
xlabel('Time');
ylabel('VaR_{95%,1d}');
title('Time Series of 95% VaR Estimates');
grid on;
```



```
% Plot ES for the last date
figure;
bar(ES_RW(end, :));
xticklabels({'95%', '99%'});
xlabel('Confidence Level');
```

```
ylabel('ES');
title('Expected Shortfall (ES) for the Last Date');
grid on;
```



The VaR estimates fluctuate over time, indicating changes in the risk profile of the portfolio. Periods with higher VaR values suggest increased risk, while lower VaR values indicate reduced risk.

The Expected Shortfall tells us how bad the losses could be when they go beyond the VaR threshold. If the ES is high, it means that when losses do exceed the VaR, they are likely to be quite large. The ES at the 99% confidence level is higher than at the 95% level, which is expected since the 99% level captures more extreme outcomes.

```
% Task 3a), Risk Factor
Indexes = flip(readtable('timeSeries.xlsx', 'Sheet', 'Problem 3', 'Range', 'B3:E3429','Variable
% Read the data as a table and preserve the original column headers
portfolio_data = readtable('timeSeries.xlsx', 'Sheet', 'Problem 3', 'Range', 'G3:L6', 'Variable
% Extract columns and convert to datetime
t = datetime(Indexes{end, 1}); % Convert the last row of the first column to datetime
T = datetime(portfolio_data{:, 2}); % Convert the second column to datetime
% Extract other variables as arrays
mid_iv = mean(portfolio_data{:, 4:5}, 2) / 100; % Mean of columns 4 and 5
K = portfolio_data{:, 1}; % Strike prices (numeric array)
Type = portfolio_data{:, 3}; % Option type (cell array of strings)
h = portfolio_data{:, 6}; % Other data (numeric array)
q = 0.05; % Dividend yield
risk_free_rates = Indexes{:, 4}/100;
S = Indexes{end, 2}; % Underlying asset price (numeric value)
option_prices = zeros(height(mid_iv),1);
Deltas = zeros(height(mid_iv), 1);
Vegas = zeros(height(mid_iv), 1);
```

```
Rhos = zeros(height(mid_iv), 1);

for i = 1:length(option_prices)
        [option_prices(i, 1), Deltas(i, 1), Vegas(i, 1), Rhos(i, 1)] = BlackScholes(S, K(i, 1), r: end

%Calculate changes

% Extract numeric data from the table
sp500_prices = Indexes{:, 2}; % Extract S&P 500 prices as a numeric array
vix_prices = Indexes{:, 3}/100; % Extract VIX prices as a numeric array
% Extract risk-free rates as a numeric array

% Calculate changes in risk factors
log_returns_sp_500 = log(sp500_prices(2:end) ./ sp500_prices(1:end-1)); % Log returns for S&P !
delta_VIX = vix_prices(2:end) - vix_prices(1:end-1); % Change in VIX
delta_r = risk_free_rates(2:end) - risk_free_rates(1:end-1); % Change in risk-free rate

RF = log_returns_sp_500(end,1)* transpose(h) * Deltas + delta_VIX(end,1) * transpose(h) * Vegas
```

RF = 3.9999e + 05

portfolio_std = 1.3187e+06

```
%Value at Risk
z_alpha = norminv(0.99); % For 99% confidence level
VaR_portfolio = z_alpha * sqrt(portfolio_variance)
```

VaR_portfolio = 3.0677e+06

Functions:

```
function [LL] = EVT_log_likelihood(params, data)
    epsilon = params(1);
    beta = params(2);
    T = length(data);
    loglikelihoods = zeros(T, 1);
    for t = 1:T
        loglikelihoods(t) = -log(beta) - (1 + (1 / epsilon)) * log(1 + (epsilon * data(t)) / beta
    end
    LL = -sum(loglikelihoods);
end
function [BS, Delta, Vega, Rho] = BlackScholes(S0, K, r, q, t, T, current_sigma, Type)
new_T = days252bus(t, T)/252;
S = S0*exp(-q * (new_T));
d1 = (\log(S/K) + (r - q + (current_sigma^2)/2)*(new_T)) / (current_sigma*sqrt(new_T));
d2 = d1 - current_sigma*sqrt(new_T);
if Type == "Call"
    BS = S*exp(-q* (new_T))* normcdf(d1) - K*exp(-r*(new_T))*normcdf(d2);
    Delta = exp(-q * new_T) * normcdf(d1);
    Rho = K * new_T * exp(-r * new_T) * normcdf(d2);
elseif Type == "Put"
    BS = K*exp(-r*(new_T))*normcdf(-d2) - S*exp(-q*(new_T))* normcdf(-d1);
    Delta = -exp(-q * new_T) * normcdf(-d1);
    Rho = -K * new T * exp(-r * new T) * normcdf(-d2);
end
    Vega = S * exp(-q * new_T) * normpdf(d1) * sqrt(new_T);
end
function y = gpd_pdf(x, epsilon, beta)
    y = (1 / beta) * (1 + epsilon * (x / beta)).^(-1/epsilon - 1);
end
function [LL, variance] = garch_likelihood(params, returns, is_fixed)
    if is_fixed == "true"
```

```
omega = var(returns);
        alpha = params(1);
        beta = params(2);
    elseif is_fixed == "false"
        omega = params(1);
        alpha = params(2);
        beta = params(3);
    end
   % Initialize variables
    T = length(returns);
    variances = zeros(T,1);
    loglikelihoods = zeros(T,1);
   % Initial variance
    variances(1) = var(returns);
   % Calculate variances and log-likelihoods
    for t = 2:T
       variances(t) = omega * (1-alpha-beta) + alpha * returns(t-1)^2 + beta * variances(t-1)
       loglikelihoods(t) = -0.5 * (log(2*pi) + log(variances(t)) + returns(t)^2 / variances(t)
    variance = variances(1:T,1);
   % Negative log-likelihood
    LL = -sum(loglikelihoods);
end
```