```
assets = flip(readmatrix('timeSeries.xlsx', 'Sheet', 'Problem 1 and 2', 'Range', 'C3:G1512'));
% Task 1 a), Variance-Covariance Method
numColumns = size(assets, 2);
numRow = size(assets, 1);
w = 1 / numColumns;
w_matrix = repmat(w, numColumns,1);
returns = zeros(numRow-1, numColumns);
for row = 1: (numRow-1)
   for col = 1:(numColumns)
      returns(row,col) = (assets(row+1, col) - assets(row,col)) / assets(row,col);
   end
end
mu = w_matrix' * mean(returns)';
cov_matrix = cov(returns);
sigma = sqrt(transpose(w_matrix) * cov_matrix * w_matrix);
VaR95 = (-mu + norminv(0.95)*sigma)
```

```
VaR95 = 0.0325
```

```
VaR99 = (-mu + norminv(0.99)*sigma)
```

VaR99 = 0.0463

The Value at Risk can be showed above.

```
ES95 = mu + sigma*(normpdf(norminv(0.95))/0.05)

ES95 = 0.0424
```

```
ES99 = mu + sigma*(normpdf(norminv(0.99))/0.01)
```

ES99 = 0.0545

Since L is smaller than both VaR95 and VaR99 the answers to Expected Shortfall is not defined.

```
%i)
log_returns = zeros(numRow-1,numColumns);
for row = 1: (numRow-1)
    for col = 1:(numColumns)
        log_returns(row,col) = log(assets(row+1, col)/ assets(row,col));
    end
end
% Optimization of GARCH with fmincon, without variance targeting
% Initial parameter guesses
initial_params = [0.1, 0.1, 0.8];
% Constraints: (0 <= omega, alpha, beta <= 1 and alpha + beta <= 1)</pre>
1b = [0, 0, 0];
ub = [1, 1, 1];
A = [];
b = [];
Aeq = [];
beq = [];
nonlcon = @(params) deal([], params(1) + params(2) - 1); % alpha + beta <= 1
% Optimization options
options = optimoptions('fmincon', 'Display', 'off');
parameters = zeros(3,numColumns);
variance_garch = zeros(numRow-1,numColumns);
for i = 1:numColumns
    [params_opt, neg_LL] = fmincon(@(params) garch_likelihood(params, log_returns(1:(numRow-1))
    parameters(1:length(params_opt),i) = params_opt;
    [11 ,variance_GARCH] = garch_likelihood(params_opt, log_returns(1:(numRow-1),i), "false");
    variance_garch(1:numRow-1,i) = variance_GARCH;
end
parameters;
variance_garch;
%ii)
standard_returns = zeros(numRow-1,numColumns);
```

```
%ii)
standard_returns = zeros(numRow-1,numColumns);

for col = 1:numColumns
    for row = 1:numRow-1
        standard_returns(row,col) = log_returns(row, col) / sqrt(variance_garch(row, col));
    end
end

standard_returns; %Standard_returns_equals_the_normally_distirbuted_residuals
```

```
U = zeros(numRow-1,numColumns);
for col = 1:numColumns
   U(1:numRow-1, col) = normcdf(standard_returns(1:numRow-1,col));
end
U;
[Rho t, df] = copulafit('t', U);
%iii)
noOfSimulations = 1000;
rand_data = copularnd("t", Rho_t, df, noOfSimulations);
%iv)
residuals = norminv(rand data);
sim returns = residuals * transpose(sqrt(variance_garch(numRow-1,1:numColumns)));
VaR95_sim= -prctile(sim_returns, 5)
VaR95_sim = 0.1275
VaR99_sim= -prctile(sim_returns, 1)
VaR99_sim = 0.1700
ES95_sim= -mean(sim_returns(sim_returns < -VaR95_sim))</pre>
ES95 sim = 0.1545
ES99_sim= -mean(sim_returns(sim_returns < -VaR99_sim))</pre>
ES99_sim = 0.1971
% Task 1 c), Extreme Value Theory
%i) MLE
portfolio_returns = log_returns * w_matrix;
threshold = -quantile(portfolio_returns, 0.05)
threshold = 0.0284
excess_losses = -portfolio_returns(portfolio_returns < -threshold) - threshold;</pre>
```

initial_params = [0.1, 1]; % Initial guesses for xi and beta

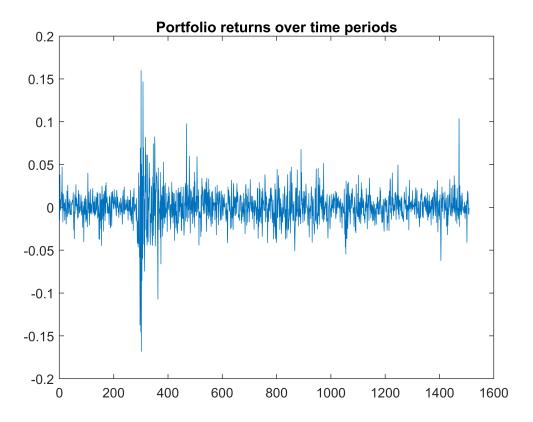
```
lb = [-Inf, 0]; % Lower bounds (beta > 0)
 ub = [Inf, Inf]; % Upper bounds
 A = [];
 b = [];
 Aeq = [];
 beq = [];
 nonlcon = [];
 options = optimoptions('fmincon', 'Display', 'off');
 [params_opt_EVT] = fmincon(@(params) EVT_log_likelihood(params, excess_losses), initial_params
 epsilon = params_opt_EVT(1)
 epsilon = 0.3075
 beta = params_opt_EVT(2)
 beta = 0.0120
 VaR_99_1d_EVT = (threshold + (beta/epsilon) * (((1 - 0.99)/(1-0.95))^(-epsilon) - 1))
 VaR_{99}1d_{EVT} = 0.0534
 ES_99_1d_EVT = (VaR_99_1d_EVT + beta - epsilon * threshold) / (1 - epsilon)
 ES_{99}1d_{EVT} = 0.0819
The value of VaR and the constituent quantities can be shows in the display.
 %ii) 500 observations
 newdata = portfolio_returns(250:750);
```

```
newdata = portfolio_returns(250:750);
threshold = -quantile(newdata, 0.05)

threshold = 0.0369

New_excess_losses = -newdata(newdata < -threshold) - threshold;
figure;
plot(portfolio_returns) %Seem to be large fluctuations in returns from data points 250 to 750</pre>
```

title("Portfolio returns over time periods")



```
% data_points = portfolio_returns(250:750);
initial_params = [0.1, 1]; % Initial guesses for xi and beta
lb = [-Inf, 0]; % Lower bounds (beta > 0)
ub = [Inf, Inf]; % Upper bounds
A = [];
b = [];
Aeq = [];
beq = [];
nonlcon = [];
options = optimoptions('fmincon', 'Display', 'off');
[params_opt_EVT_2, neg_LL_EVT] = fmincon(@(params) EVT_log_likelihood(params, New_excess_losses
epsilon_500 = params_opt_EVT_2(1)
epsilon_500 = 0.3882
```

```
beta_500 = params_opt_EVT_2(2)
```

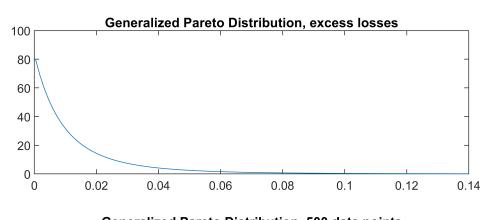
 $beta_500 = 0.0186$

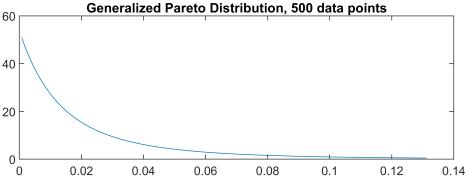
```
%Plot the distribution functions
x1 = linspace(min(excess_losses), max(excess_losses), 100);
density1 = gpd_pdf(x1, epsilon, beta);
```

```
% Calculate the density for the 500 data points using the estimated parameters
x2 = linspace(min(New_excess_losses), max(New_excess_losses), 100);
density2 = gpd_pdf(x2, epsilon_500, beta_500);

figure;
subplot(2, 1, 1);
plot(x1, density1)
title('Generalized Pareto Distribution, excess losses')

subplot(2, 1, 2);
plot(x2, density2);
title('Generalized Pareto Distribution, 500 data points')
```





```
%VaR99 is normally distributed, VaR99_rand is Monte Carlo Simulated
Relative_VaR = VaR_99_1d_EVT/VaR99
```

Relative_VaR = 1.1547

The distribution functions shows that for VaR 99% the inital excess losses are larger for a lower percentage but it decreases faster than for the 500 data points. This may be due to the 500 data point interval being chosen as a volatile one.

```
window size = 500;
N = length(log_returns);
confidence_levels = [0.95, 0.99];
portfolio_weights = zeros(N, numColumns);
portfolio_weights(1:window_size+1 , 1:numColumns) = 1/numColumns;
sorted returns = zeros(1, noOfSimulations);
VaR_RW = zeros(1, length(confidence_levels));
ES_RW = zeros(1, length(confidence_levels));
failures = zeros(1, length(confidence_levels));
port_return = zeros(N, 1);
hist return = zeros(window size,1);
for t = window_size+2:(N)
    port_return(t,1) = log_returns(t, :) * portfolio_weights(t-1,:)';
    %Update portfolio weights
   for i = 1:numColumns
       portfolio_weights(t,i) = portfolio_weights(t-1,i) * (1 + log_returns(t, i)) / (1+port_
   end
end
for t = window_size+2:(N)
   hist_return = log_returns(t-window_size:t-1,:) * portfolio_weights(t-1,:)';
   % Calculate VaR and ES
   sorted returns = sort(hist return);
   for cl = 1:length(confidence_levels)
       alpha = 1 - confidence_levels(cl);
       VaR_RW(t-window_size, cl) = -prctile(sorted_returns , alpha * 100);
       ES RW(t-window size, cl) = -mean(sorted returns(sorted returns <= -VaR RW(t-window size
   end
   % Compare realized returns with VaR
   for cl = 1:length(confidence_levels)
       failures(t-window_size, cl) = port_return(t,1) < -VaR_RW(t-window_size, cl);</pre>
   end
end
```

```
% Count failures for 95% VaR
failures_95 = failures(:, 1); % First column corresponds to 95% VaR
observed_failure_rate = mean(failures_95);

% Perform binomial test, because the value can either be higher or lower
% than VaR
p_value = binocdf(sum(failures_95), N - window_size, expected_failure_rate);

% Two-sided test: Check if observed failure rate is significantly different from expected
if p_value < 0.025 || p_value > 0.975
    disp('Reject null hypothesis: Failure rate is significantly different from expected.');
else
    disp('Fail to reject null hypothesis: Failure rate is as expected.');
end
```

Reject null hypothesis: Failure rate is significantly different from expected.

```
disp(['Observed failure rate: ', num2str(observed_failure_rate)]);
```

Observed failure rate: 0.030723

```
disp(['Expected failure rate: ', num2str(expected_failure_rate)]);
```

```
Expected failure rate: 0.05
```

```
disp(['p-value: ', num2str(p_value)]);
```

p-value: 0.0018135

The p-value for the binomial cdf for the failure rate with 95% confidence is within the p-value interval. Therefore we reject the null hypothesis, in other words the failure rate is significantly different from the expected one. This implies that the VaR model is underestimating risk.

```
%Lnull is the likelihood under the null hypothesis (independent breaches),
%Lalternative is the likelihood of dependent breaches.

LR = -2 * log((1 - observed_failure_rate)^(n00 + n10) * observed_failure_rate^(n01 + n11)) + 2
```

LR = 3.0747

```
% From lecture 5
critical_value = chi2inv(0.95,1)
```

critical value = 3.8415

```
% Compare test statistic to critical value
if LR > critical_value
    disp('Reject null hypothesis: VaR breaches are serially dependent.');
else
    disp('Fail to reject null hypothesis: VaR breaches are serially independent.');
end
```

Fail to reject null hypothesis: VaR breaches are serially independent.

```
disp(['Likelihood ratio test statistic: ', num2str(LR)]);
```

Likelihood ratio test statistic: 3.0747

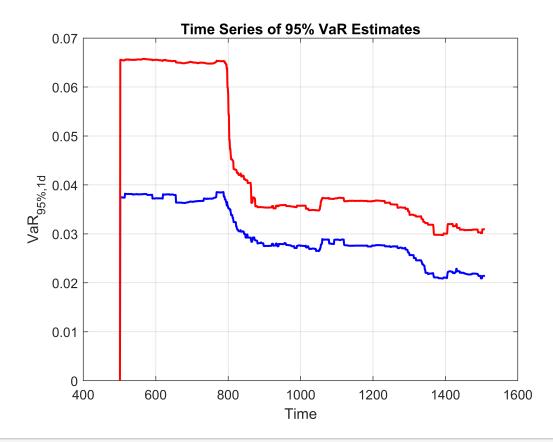
```
disp(['Critical value: ', num2str(critical_value)]);
```

Critical value: 3.8415

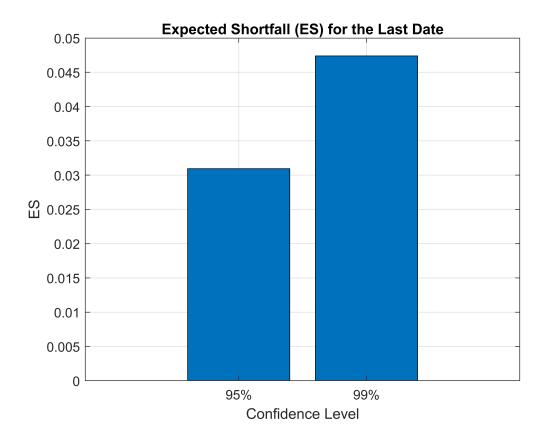
Since the Likelihood ratio is larger than the critical value for 95% confidence interval we fail to reject the null hypothesis.

```
%Task 2, plots

% Plot time series of 95% VaR
figure;
plot(window_size+1:N, VaR_RW(1:N-window_size, 1), 'b', 'LineWidth', 1.5);
hold on;
plot(window_size+1:N, ES_RW(1:N-window_size, 1), 'r', 'LineWidth', 1.5); % Plot 95% ES
xlabel('Time');
ylabel('VaR_{95%,1d}');
title('Time Series of 95% VaR Estimates');
grid on;
```



```
% Plot ES for the last date
figure;
bar(ES_RW(end, :));
xticklabels({'95%', '99%'});
xlabel('Confidence Level');
ylabel('ES');
title('Expected Shortfall (ES) for the Last Date');
grid on;
```



The VaR estimates fluctuate over time, indicating changes in the risk profile of the portfolio. Periods with higher VaR values suggest increased risk, while lower VaR values indicate reduced risk.

The Expected Shortfall tells us how bad the losses could be when they go beyond the VaR threshold. If the ES is high, it means that when losses do exceed the VaR, they are likely to be quite large. The ES at the 99% confidence level is higher than at the 95% level, which is expected since the 99% level captures more extreme outcomes.

```
q = 0.05; % Dividend yield
risk_free_rates = Indexes{:, 4}/100;
S = Indexes{end, 2}; % Underlying asset price (numeric value)
option_prices = zeros(height(mid_iv),1);
Deltas = zeros(height(mid_iv), 1);
Vegas = zeros(height(mid_iv), 1);
Rhos = zeros(height(mid_iv), 1);
for i = 1:length(option_prices)
     [option_prices(i, 1), Deltas(i, 1), Vegas(i, 1), Rhos(i, 1)] = BlackScholes(S, K(i, 1), ri
end
%Calculate changes
% Extract numeric data from the table
sp500_prices = Indexes{:, 2}; % Extract S&P 500 prices as a numeric array
vix prices = Indexes{:, 3}/100; % Extract VIX prices as a numeric array
% Extract risk-free rates as a numeric array
% Calculate changes in risk factors
log_returns_sp_500 = sp500_prices(2:end) - sp500_prices(1:end-1); % Log returns for S&P 500
delta_VIX = vix_prices(2:end) - vix_prices(1:end-1); % Change in VIX
delta_r = risk_free_rates(2:end) - risk_free_rates(1:end-1); % Change in risk-free rate
RF = log_returns_sp_500(end,1)* transpose(h) * Deltas + delta_VIX(end,1) * transpose(h) * Vegas
RF = 4.1650e + 05
risk_factor_changes = [log_returns_sp_500, delta_VIX, delta_r];
% Calculate the covariance matrix of risk factor changes
cov_matrix = cov(risk_factor_changes);
% Gradient vector (g) for the portfolio
g = [Deltas'; % Delta contribution
    Vegas'; % Vega contribution
     Rhos']; % Rho contribution
% Calculate portfolio variance
portfolio_variance = h' * g' * cov_matrix * g * h;
portfolio_std = sqrt(portfolio_variance)
portfolio_std = 1.3676e+06
```

%Value at Risk

z_alpha = norminv(0.99); % For 99% confidence level

```
VaR_portfolio = z_alpha * portfolio_std
```

```
VaR_portfolio = 3.1814e+06
```

Functions:

```
function [LL] = EVT_log_likelihood(params, data)
    epsilon = params(1);
    beta = params(2);
    T = length(data);
    loglikelihoods = zeros(T, 1);
    for t = 1:T
        loglikelihoods(t) = -log(beta) - (1 + (1 / epsilon)) * log(1 + (epsilon * data(t)) / beta
    end
    LL = -sum(loglikelihoods);
end
function [BS, Delta, Vega, Rho] = BlackScholes(S0, K, r, q, t, T, current_sigma, Type)
new_T = days252bus(t, T)/252;
S = S0*exp(-q * (new_T));
d1 = (\log(S/K) + (r - q + (current_sigma^2)/2)*(new_T)) / (current_sigma*sqrt(new_T));
d2 = d1 - current_sigma*sqrt(new_T);
if Type == "Call"
    BS = S*exp(-q* (new_T))* normcdf(d1) - K*exp(-r*(new_T))*normcdf(d2);
    Delta = exp(-q * new_T) * normcdf(d1);
    Rho = K * new_T * exp(-r * new_T) * normcdf(d2);
elseif Type == "Put"
    BS = K*exp(-r*(new_T))*normcdf(-d2) - S*exp(-q*(new_T))* normcdf(-d1);
```

```
Delta = -exp(-q * new_T) * normcdf(-d1);
    Rho = -K * new_T * exp(-r * new_T) * normcdf(-d2);
end
   Vega = S * exp(-q * new_T) * normpdf(d1) * sqrt(new_T);
end
function y = gpd_pdf(x, epsilon, beta)
    y = (1 / beta) * (1 + epsilon * (x / beta)).^(-1/epsilon - 1);
end
function [LL, variance] = garch_likelihood(params, returns, is_fixed)
    if is_fixed == "true"
        omega = var(returns);
        alpha = params(1);
        beta = params(2);
    elseif is fixed == "false"
        omega = params(1);
        alpha = params(2);
        beta = params(3);
    end
   % Initialize variables
    T = length(returns);
    variances = zeros(T,1);
    loglikelihoods = zeros(T,1);
   % Initial variance
    variances(1) = var(returns);
   % Calculate variances and log-likelihoods
    for t = 2:T
        variances(t) = omega * (1-alpha-beta) + alpha * returns(t-1)^2 + beta * variances(t-1)
        loglikelihoods(t) = -0.5 * (log(2*pi) + log(variances(t)) + returns(t)^2 / variances(t)
    end
    variance = variances(1:T,1);
   % Negative log-likelihood
    LL = -sum(loglikelihoods);
end
```