## Development and calibration of tumor models

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TEXAS ADVANCED COMPUTING CENTER







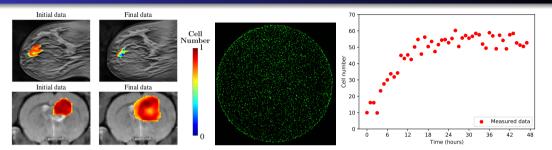




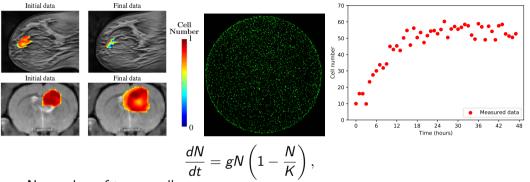
# Model Calibration



## Motivation



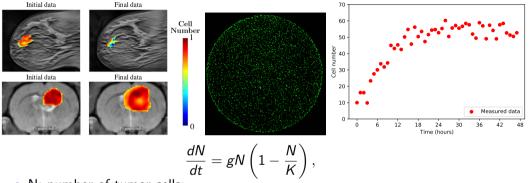
### Motivation



- N: number of tumor cells;
- g: tumor growth rate;
- K: environment carrying capacity.



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#### Model Calibration

It is done to adjust the selected model parameters, such as growth and death rates, to obtain the best fit between the model predicted responses and the available data.

# Model calibration: frequentist and Bayesian statistics<sup>1</sup>

#### Frequentist and maximum likelihood approaches

- the probability of an event is the frequency with which the event is expected to be observed over infinite samples;
- the model parameters remain constant during the repeatable process;
- maximum likelihood approach: find the parameter that enables the model to deliver the best characterization of the true target distribution.

# Model calibration: frequentist and Bayesian statistics<sup>1</sup>

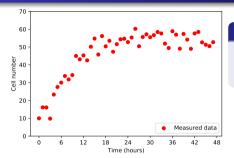
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### Bayesian approach

- transfers information from prior knowledge and observation data to the estimated parameter;
- the estimated parameter is regarded as a random variable, and the true parameter is merely a realization of the random variable.

<sup>&</sup>lt;sup>1</sup>Oden et al., Encyclopedia of Computational Mechanics Second Edition (2017)<sub>3</sub> ,

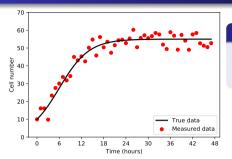


• **D**: measured data;

#### Mathematical model

$$\frac{dN}{dt} = gN\left(1 - \frac{N}{K}\right)$$

- $\theta$ : vector of model parameters,  $\theta = (g, K)$ ;
- g: tumor growth rate;
- K: environmental carrying capacity;
- $Y(\theta)$ : model prediction;

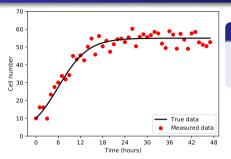


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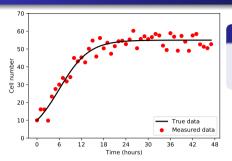
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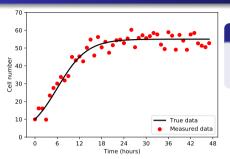
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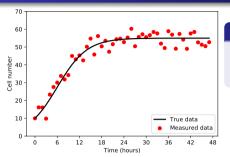
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### Assuming:

- the experimental noise is normally distributed ( $\epsilon \sim \mathcal{N}(0_{N\times 1}, \sigma_{data}^2 I_{N\times N})$ );
- ② the model inadequacy is normally distributed  $(\gamma \sim \mathcal{N}(0_{N\times 1}, \sigma_{model}^2 I_{N\times N}));$
- **1** the variance of the total error  $(\sigma)$  is such as  $\sigma^2 = \sigma_{data}^2 + \sigma_{model}^2$ ;
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$$\pi(oldsymbol{\mathcal{D}}|oldsymbol{ heta}) = \prod_{i=1}^{N_t} rac{1}{\sqrt{2\pi\sigma^2}} e^{-rac{(D_i - Y_i(oldsymbol{ heta}))^2}{2\sigma^2}},$$

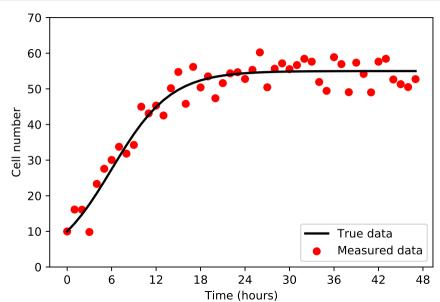
•  $N_t$ : the number of data points.

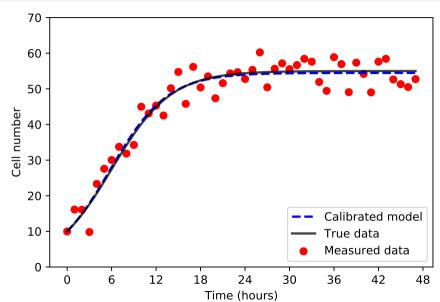


Find  $\hat{\theta} \in \Theta$ , where  $\Theta$  is the parameter space, that maximize the likelihood.

It is customary to work with the more manageable log-likelihood function.

$$\begin{split} \hat{\boldsymbol{\theta}} &= \operatorname*{argmax}[\log \pi(\boldsymbol{D}|\boldsymbol{\theta})]; \\ &= \operatorname*{argmax}_{\boldsymbol{\theta} \in \Theta} \left[ -\frac{N_t}{2} \log(2\pi) - \frac{N_t}{2} \log(\sigma^2) - \sum_{i=1}^{N_t} \frac{1}{2} \frac{(D_i - Y_i(\boldsymbol{\theta}))^2}{2\sigma^2} \right]. \end{split}$$





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Converting to probability densities  $\pi$ , if A represents the parameter  $\theta$  of a model, and B the observational data D:

$$\underbrace{\pi(\boldsymbol{\theta}|\boldsymbol{D})}_{\text{posterior}} = \underbrace{\frac{\pi(\boldsymbol{D}|\boldsymbol{\theta})}{\pi(\boldsymbol{\theta})}}_{\text{likelihood prior}}; \qquad \pi(\boldsymbol{D}) \int_{\boldsymbol{\Theta}} \pi(\boldsymbol{D}|\boldsymbol{\theta}) \pi(\boldsymbol{\theta}) \, \mathrm{d}\boldsymbol{\theta}$$

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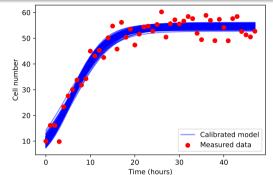
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