

CMDA-3654

Homework 6

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Due as a .pdf upload

Instructions:

Delete the Instructions section from your write-up!!

I have given you this assignment as an .Rmd (R Markdown) file.

- Change the name of the file to: `Lastname_Firstname_CMDA_3654_HW6.Rmd`, and your output should therefore match but with a `.pdf` extension.
- You need to edit the R Markdown file by filling in the chunks appropriately with your code. Output will be generated automatically when you compile the document.
- You also need to add your own text before and after the chunks to explain what you are doing or to interpret the output.
- Feel free to add additional chunks if needed. I **will not** be providing assignments to you like this for the entire semester, just long enough for you to learn how to do it for yourself.

Required: The final product that you turn in must be a .pdf file.

- You can Knit this document directly to a PDF if you have LaTeX installed (which is preferred).
- If you absolutely can't get LaTeX installed and/or working, then you can compile to a .html first, by clicking on the arrow button next to knit and selecting Knit to HTML.
- You must then print you .html file to a .pdf by using first opening it in a web browser and then printing to a .pdf

Problem 1: [30 pts] Exploring Relationships between variables.

Load the `DatasaurusDozen.tsv` file into R.

This data consists of x and y observations for **13 sub-datasets** that have the following names:

`dino`, `away`, `h_lines`, `v_lines`, `x_shape`, `star`, `high_lines`, `dots`, `circle`, `bullseye`, `slant_up`, `slant_down`, `wide_lines`

- a. Use `dplyr` functions to summarize each dataset in the following way: Compute the mean for x, mean for y, sd for x, sd for y, and the correlation coefficient between x and y. **Please round your answers to 2 decimal places.** The answers should be returned automatically in a tibble. Use `kable()` or `pandoc.table()` (use `results='asis'` in chunk definition if using `pandoc.table()`) or some other function to make nicely formatted table of your results.

```
library(tidyverse)
library(dplyr)
#Reading the file
dino<-read.table("/Users/eduardosalvador/Desktop/FINAL\ Spring\ Semester\ 2021/CMDA\ /Assignments/HW6/Datasau
#Looking at the first 6 rows of dataframe
head(dino)
```

| | dataset | x | y |
|---|---------|---------|---------|
| 1 | dino | 55.3846 | 97.1795 |
| 2 | dino | 51.5385 | 96.0256 |
| 3 | dino | 46.1538 | 94.4872 |
| 4 | dino | 42.8205 | 91.4103 |
| 5 | dino | 40.7692 | 88.3333 |
| 6 | dino | 38.7179 | 84.8718 |

```
#Using dplyr to summerize mean,sd and correlation
dinoof %>%
  group_by(dinoof$dataset) %>%
  summarise("x_mean"=mean(x),"y_mean"=mean(y),"x_sd"=sd(x),"y_sd"=sd(y),"correlation_coef"=round(cor(x,y),2))

# A tibble: 13 x 6
  'dinoof$dataset' x_mean y_mean x_sd y_sd correlation_coef
  <chr>           <dbl> <dbl> <dbl> <dbl>          <dbl>
1 away           54.3  47.8  16.8  26.9         -0.06
2 bullseye       54.3  47.8  16.8  26.9         -0.07
3 circle         54.3  47.8  16.8  26.9         -0.07
4 dino           54.3  47.8  16.8  26.9         -0.06
5 dots           54.3  47.8  16.8  26.9         -0.06
6 h_lines        54.3  47.8  16.8  26.9         -0.06
7 high_lines     54.3  47.8  16.8  26.9         -0.07
8 slant_down     54.3  47.8  16.8  26.9         -0.07
9 slant_up       54.3  47.8  16.8  26.9         -0.07
10 star          54.3  47.8  16.8  26.9         -0.06
11 v_lines       54.3  47.8  16.8  26.9         -0.07
12 wide_lines    54.3  47.8  16.8  26.9         -0.07
13 x_shape       54.3  47.8  16.8  26.9         -0.07
```

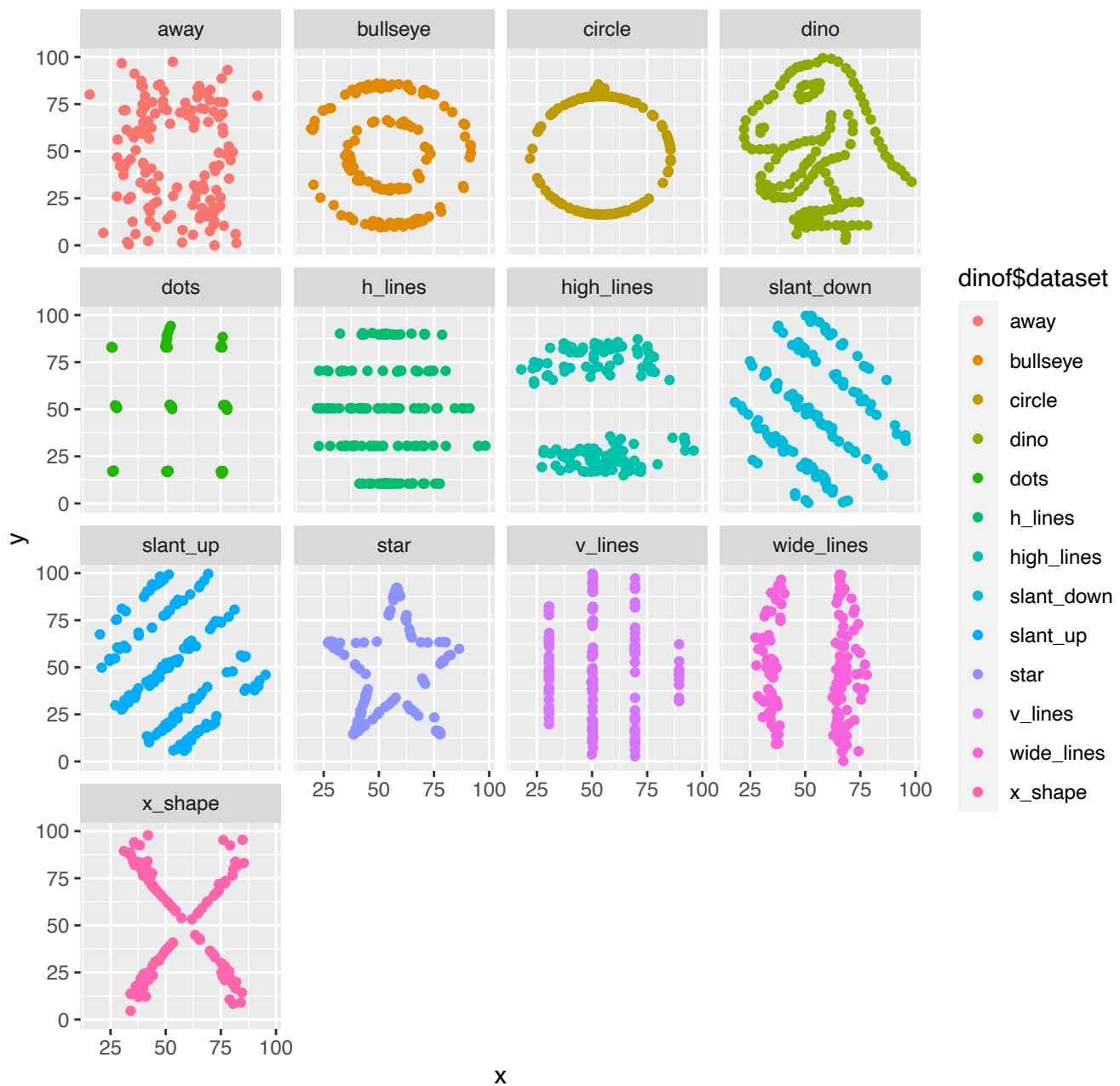
- b. What does the numerical summaries tell you about the data in the 12 different data sets? In particular, does the correlation coefficient provide you with much information about the relationship between x and y?

The 12 different data sets have the same mean for x and y. Furthermore, they also have the same standard di

- c. Now make a basic scatterplot of x and y for the 13 different datasets. Use a different color for each dataset. My best advice is to simply use `ggplot()` with `facet_wrap()`, as this can be done in a single line.

```
#Created scatterplot using ggplot (to color everything, use color function)
ggplot(dinoof,aes(x,y,color=dinoof$dataset))+facet_wrap(vars(dinoof$dataset))+ggtitle("Basic Scatterplot for 13 d
```

Basic Scatterplot for 13 datasets



d. How does your interpretation about the relationships between x and y change after seeing the plots?

##The points on each of the graphs appears to match the tittle of the dataset.

e. What lesson can be learned here?

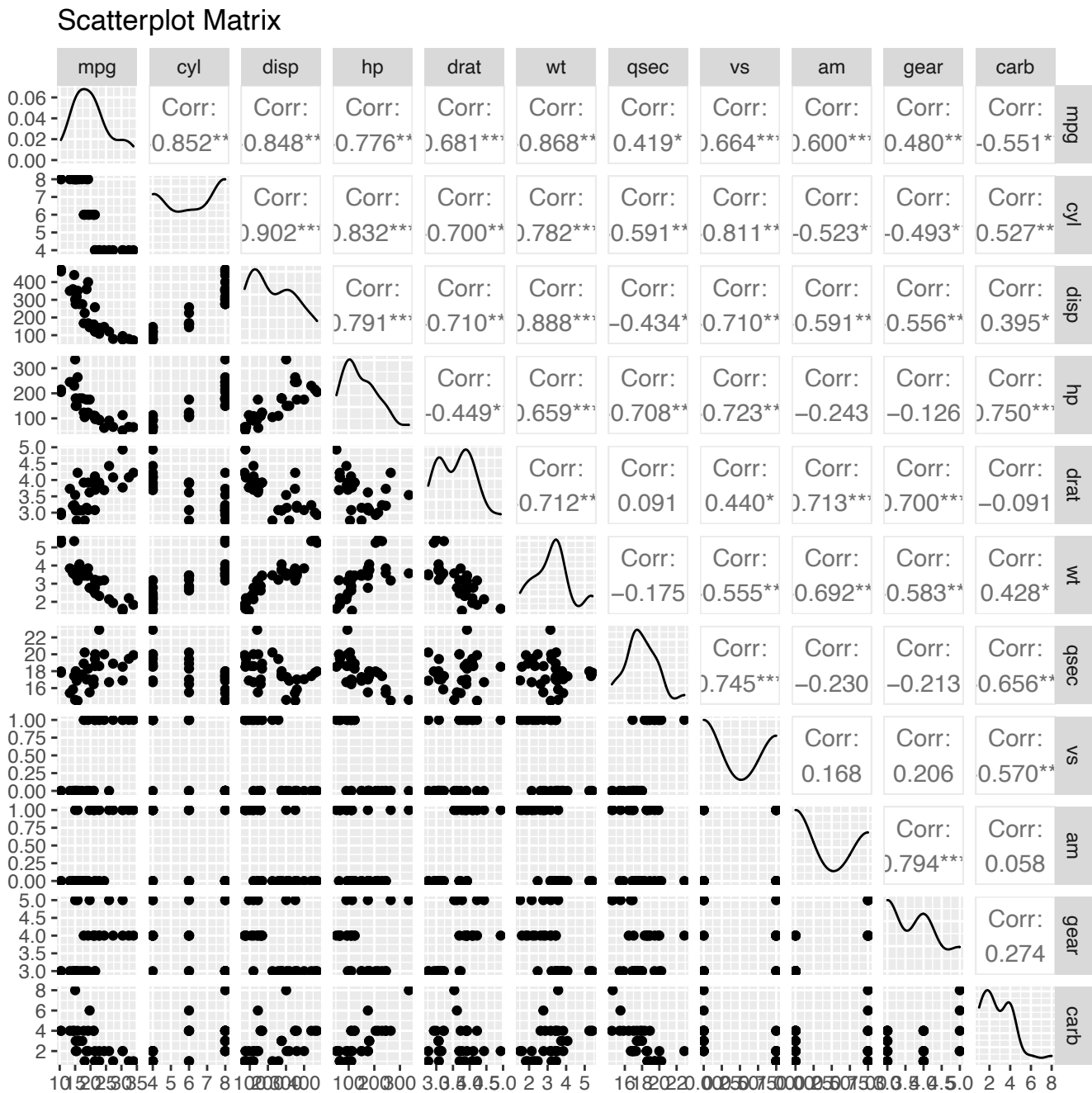
##That some graphs are useless in the sense of making sense of the data since, even tho the mean and sd are th

Problem 2: [30 pts] Linear Regression

Consider the `mtcars` dataset. Say we want to build a linear regression model that predicts `mpg`, using any subset of the other variables as predictors.

- a. Begin by creating a scatterplot matrix between mpg and all other predictors. Report the correlations as well in either the upper or lower half of the scatterplot matrix.

```
library(GGally)
#Create scatterplot with ggpairs
ggpairs(mtcars)+ggtitle("Scatterplot Matrix")
```



- b. What are the three variables most highly correlated with mpg?

##The variables most highly correlated with mpg are wt (.868), then cyl (.852) and displacement (.848)

- c. Fit three simple linear regression models using your previous three variables/predictors. Report summaries for the models. Which model would you choose and why?

```
#Simple linear regression with mpg~wt
slrmwt<-lm(mpg ~ wt,data=mtcars)
summary(slrmt)
```

```
Call:
lm(formula = mpg ~ wt, data = mtcars)

Residuals:
    Min       1Q   Median       3Q      Max
-4.5432 -2.3647 -0.1252  1.4096  6.8727

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  37.2851     1.8776  19.858 < 2e-16 ***
wt          -5.3445     0.5591  -9.559 1.29e-10 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.046 on 30 degrees of freedom
Multiple R-squared:  0.7528,    Adjusted R-squared:  0.7446
F-statistic: 91.38 on 1 and 30 DF,  p-value: 1.294e-10
```

```
#Simple linear regression with mpg~cyl
slrmcyl<-lm(mpg ~ cyl ,data=mtcars)
summary(slrmcyl)
```

```
Call:
lm(formula = mpg ~ cyl, data = mtcars)

Residuals:
    Min       1Q   Median       3Q      Max
-4.9814 -2.1185  0.2217  1.0717  7.5186

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  37.8846     2.0738  18.27 < 2e-16 ***
cyl          -2.8758     0.3224  -8.92 6.11e-10 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.206 on 30 degrees of freedom
Multiple R-squared:  0.7262,    Adjusted R-squared:  0.7171
F-statistic: 79.56 on 1 and 30 DF,  p-value: 6.113e-10
```

```
#Simple linear regression with mpg~disp
slrmdisp<-lm(mpg ~ disp,data=mtcars)
summary(slrmdisp)
```

```
Call:
lm(formula = mpg ~ disp, data = mtcars)

Residuals:
    Min       1Q   Median       3Q      Max
-4.8922 -2.2022 -0.9631  1.6272  7.2305

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 29.599855   1.229720  24.070 < 2e-16 ***
disp        -0.041215   0.004712  -8.747 9.38e-10 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.251 on 30 degrees of freedom
```

Multiple R-squared: 0.7183, Adjusted R-squared: 0.709
F-statistic: 76.51 on 1 and 30 DF, p-value: 9.38e-10

##I would use the simple linear regression model with the variable wt since it gave the largest R² (0.7528) o

- d. Create a multiple linear regression (MLR) model using `stepAIC()` to identify the best subset of predictors from all of the variables in `mtcars` (obviously `mpg` is still the response variable). Report these predictors, and a summary of the model these predictors produced.

```
library(MASS)
#Null model
modfit0<-lm(mpg~1,data=mtcars)
#Model with all of the predictors variables (a . calls for all of the variables in the dataframe)
modfit_full<-lm(mpg~.,data=mtcars)
#Using stepwise regression
modelfit_best<-stepAIC(modfit0,trace=0,scope = list(lower=modfit0,upper=modfit_full),direction = "both")
summary(modelfit_best)
```

Call:
`lm(formula = mpg ~ wt + cyl + hp, data = mtcars)`

Residuals:

| Min | 1Q | Median | 3Q | Max |
|---------|---------|---------|--------|--------|
| -3.9290 | -1.5598 | -0.5311 | 1.1850 | 5.8986 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|----------|------------|---------|--------------|
| (Intercept) | 38.75179 | 1.78686 | 21.687 | < 2e-16 *** |
| wt | -3.16697 | 0.74058 | -4.276 | 0.000199 *** |
| cyl | -0.94162 | 0.55092 | -1.709 | 0.098480 . |
| hp | -0.01804 | 0.01188 | -1.519 | 0.140015 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.512 on 28 degrees of freedom
Multiple R-squared: 0.8431, Adjusted R-squared: 0.8263
F-statistic: 50.17 on 3 and 28 DF, p-value: 2.184e-11

##The best subset of predictors are wt, cyl and hp with a Multiple R² of 0.8431 and Adjusted R² of 0.8263 wh

- e. Compare your MLR model to your three simple linear regression models earlier. Are any of those predictors in your MLR model? Are the coefficients the same for those predictors? If not, explain what may have caused the change.

##The predictors/variables wt and cyl were in my simple linear regression model but, hp wasn't, instead I thou

Problem 3: [30 pts] More Linear Regression

Sometimes your dataset is rather small, but you see that a simple linear regression is not appropriate so you try harder to fit a more complicated model. This is an example of such a situation.

A poultry scientist was studying various dietary additives to increase the rate at which chickens gain weight. One of the potential additives was studied by creating a new diet that consisted of a standard basal diet supplemented with varying amounts of the additive (0, 20, 40, 60, 80, and 100 grams). There were 60 chicks available for the study. Each of the six

diets was randomly assigned to 10 chicks. At the end of 4 weeks, the feed efficiency ratio, feed consumed (gm) to weight gain (gm), was obtained for the 60 chicks. The experiment was also concerned with the effects of high levels of copper in the chick feed. Five of the 10 chicks in each level of the feed additive received 400 ppm of copper, while the remaining five chicks received no copper.

The data is contained in the `chicken.csv` data file.

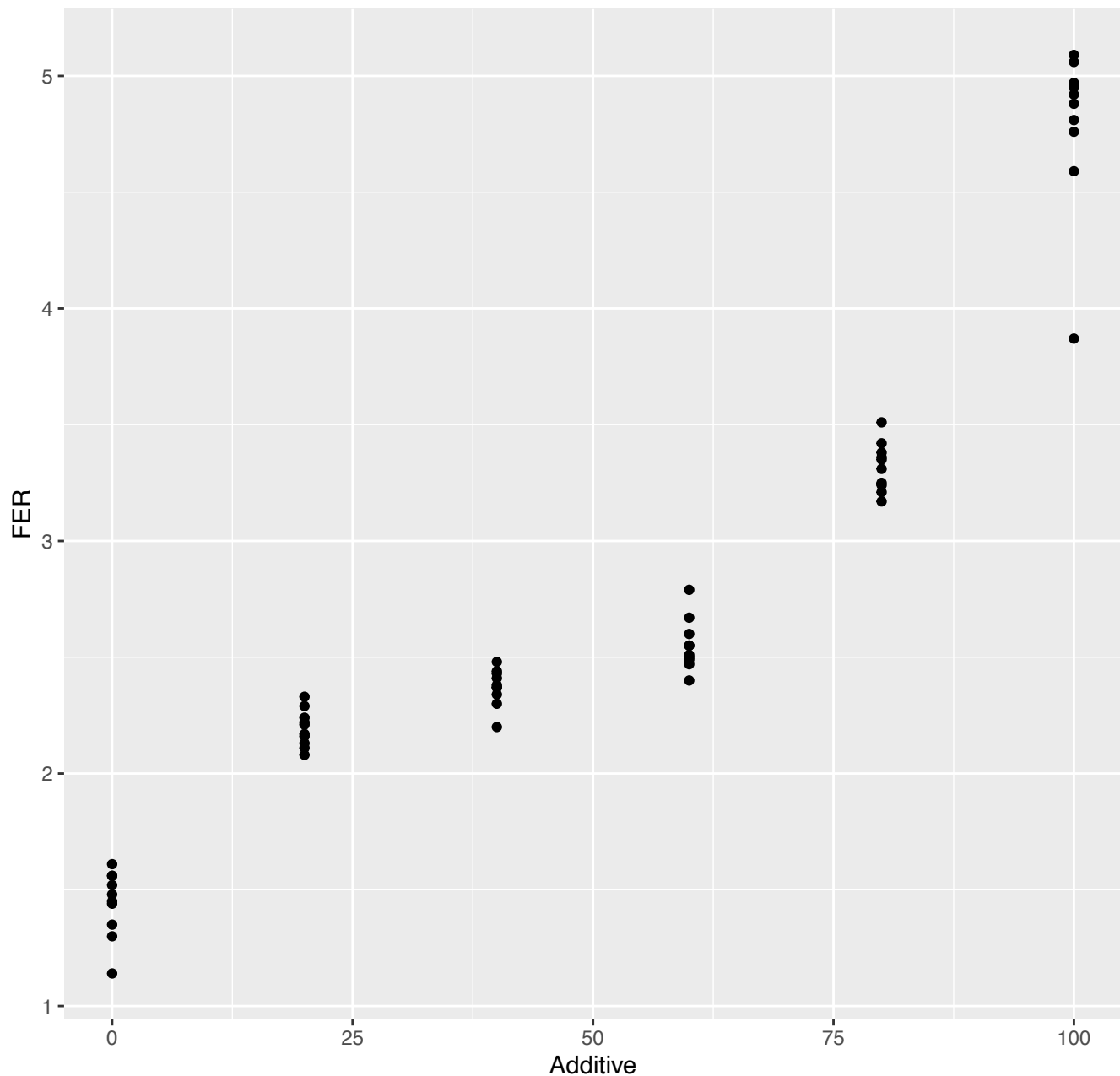
- a. In order to explore the relationship between feed efficiency ratio (FER) and feed additive (A), plot the FER versus A.

```
#Read the file
chicken_df<-read.csv("/Users/eduardosalvador/Desktop/FINAL\ Spring\ Semester\ 2021/CMDA\ /Assignments/HW6/chic
#Look at table
head(chicken_df)
```

| | FER | Additive | Copper |
|---|------|----------|--------|
| 1 | 1.30 | 0 | 0 |
| 2 | 1.35 | 0 | 0 |
| 3 | 1.44 | 0 | 0 |
| 4 | 1.52 | 0 | 0 |
| 5 | 1.56 | 0 | 0 |
| 6 | 1.61 | 0 | 400 |

```
#Used ggplot to plot FER vs Additive
ggplot(chicken_df,aes(Additive,FER)) + ggtitle("FER versus Additive")+ geom_point()
```


FER versus Additive



b. What type of regression appears most appropriate?

##The regression type most appropriate for FER versus Additive appears to be the Polynomial Regression

c. Fit first-order, quadratic, and cubic regression models to the data. Which regression equation provides the best fit to the data? Justify your answer using evidence based upon plots and relevant summaries.

```
library(pander)
library(olsrr)
#First-order regression model
f_model<-lm(Additive~FER,data=chicken_df)
summary(f_model)
```

Call:
lm(formula = Additive ~ FER, data = chicken_df)

```

Residuals:
    Min       1Q   Median       3Q      Max
-18.5200 -12.0502  -0.5643  13.0942  21.2142

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  -32.352     4.679   -6.914 4.08e-09 ***
FER           29.641     1.572   18.856 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 13.01 on 58 degrees of freedom
Multiple R-squared:  0.8598,    Adjusted R-squared:  0.8573
F-statistic: 355.6 on 1 and 58 DF,  p-value: < 2.2e-16

f_model%>%summary%>% pander()

```

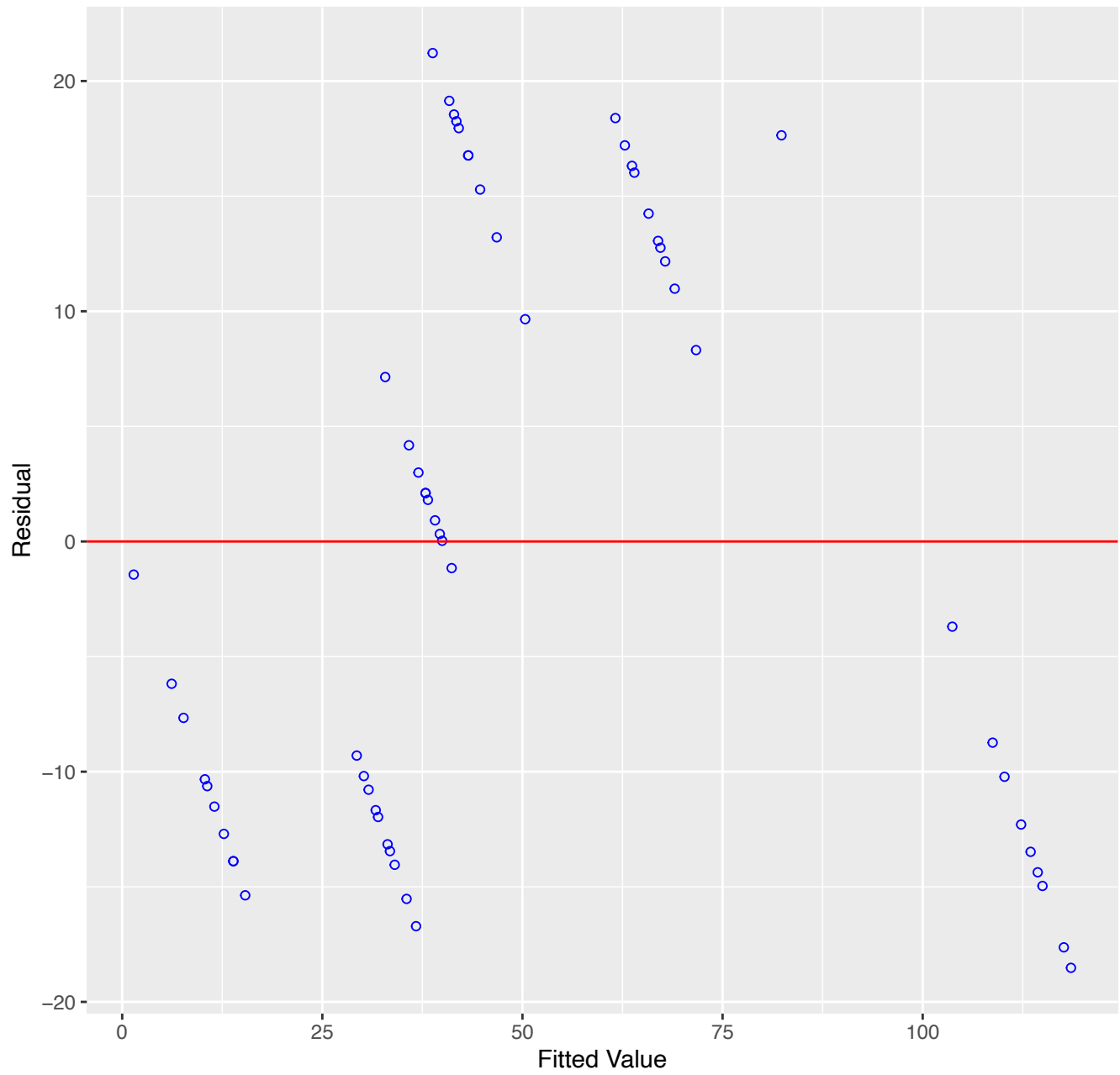
| | Estimate | Std. Error | t value | Pr(> t) |
|--------------------|----------|------------|---------|-----------|
| (Intercept) | -32.35 | 4.679 | -6.914 | 4.084e-09 |
| FER | 29.64 | 1.572 | 18.86 | 2.04e-26 |

Table 2: Fitting linear model: Additive ~ FER

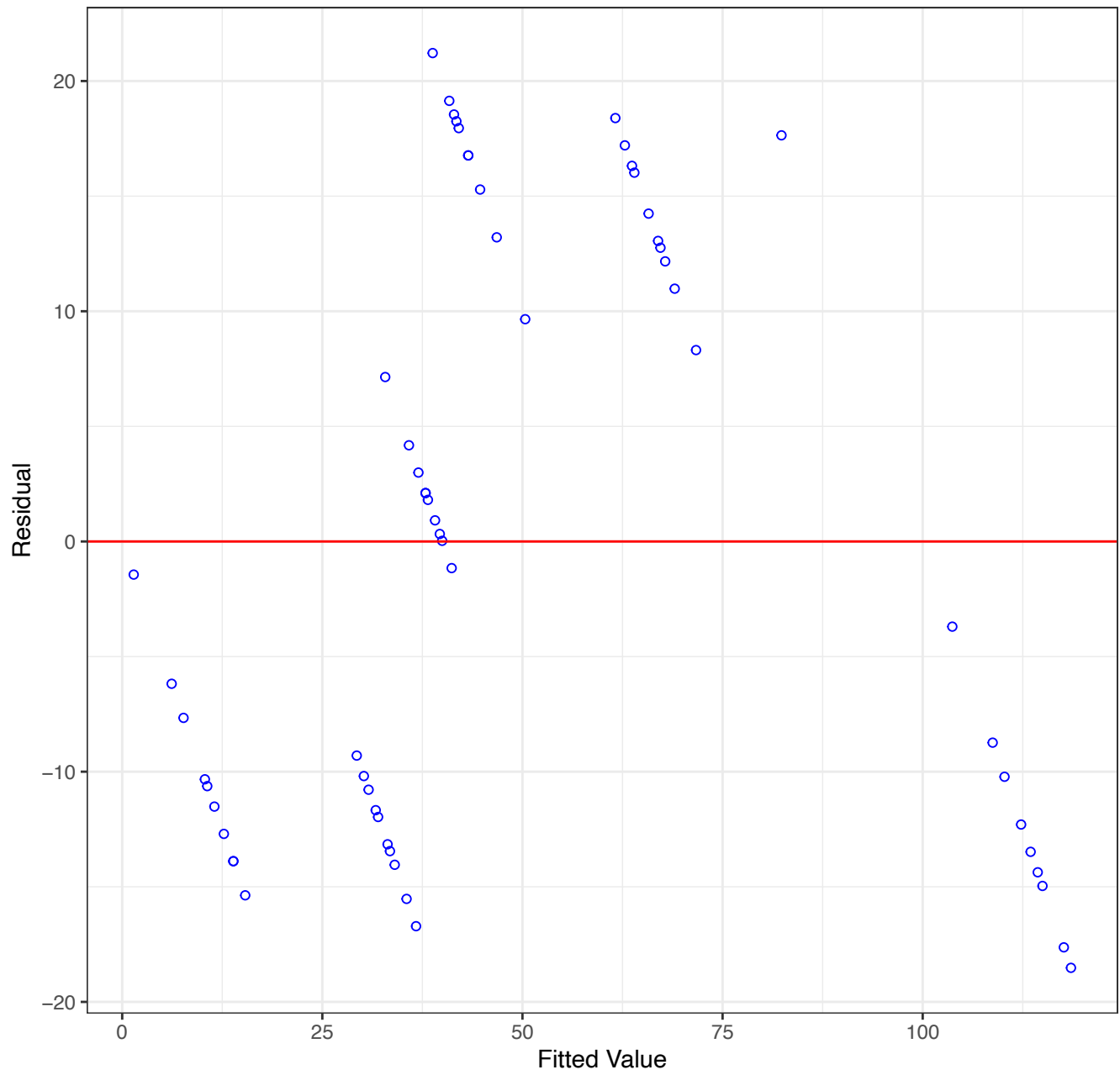
| Observations | Residual Std. Error | R^2 | Adjusted R^2 |
|--------------|---------------------|--------|----------------|
| 60 | 13.01 | 0.8598 | 0.8573 |

```
ols_plot_resid_fit(f_model)+theme_bw()
```

Residual vs Fitted Values

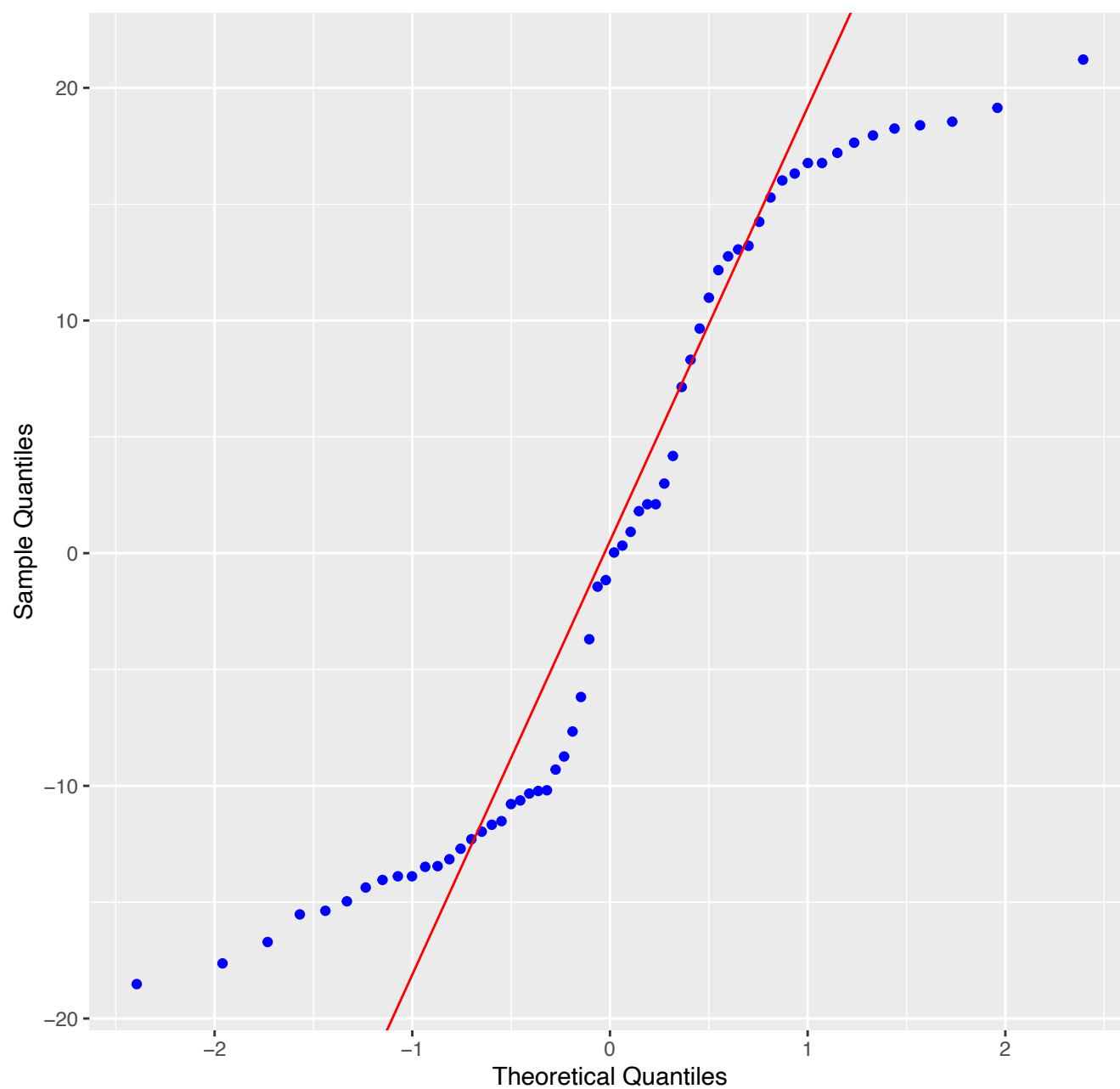


Residual vs Fitted Values



```
ols_plot_resid_qq(f_model)
```

Normal Q-Q Plot



```
ols_test_normality(f_model)
```

| Test | Statistic | pvalue |
|--------------------|-----------|--------|
| Shapiro-Wilk | 0.8902 | 1e-04 |
| Kolmogorov-Smirnov | 0.1686 | 0.0661 |
| Cramer-von Mises | 4.7343 | 0.0000 |
| Anderson-Darling | 2.3759 | 0.0000 |

```
#Quadratic regression model
```

```
q_model<-lm(Additive~FER+I(FER^2),data=chicken_df)
summary(q_model)
```

Call:

```
lm(formula = Additive ~ FER + I(FER^2), data = chicken_df)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|---------|--------|--------|-------|--------|
| -20.756 | -3.341 | -1.026 | 5.986 | 21.251 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|----------|------------|---------|--------------|
| (Intercept) | -100.639 | 9.362 | -10.75 | 2.48e-15 *** |
| FER | 78.217 | 6.333 | 12.35 | < 2e-16 *** |
| I(FER^2) | -7.525 | 0.966 | -7.79 | 1.54e-10 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9.134 on 57 degrees of freedom

Multiple R-squared: 0.9321, Adjusted R-squared: 0.9297

F-statistic: 391 on 2 and 57 DF, p-value: < 2.2e-16

```
q_model%>%summary%>% pander()
```

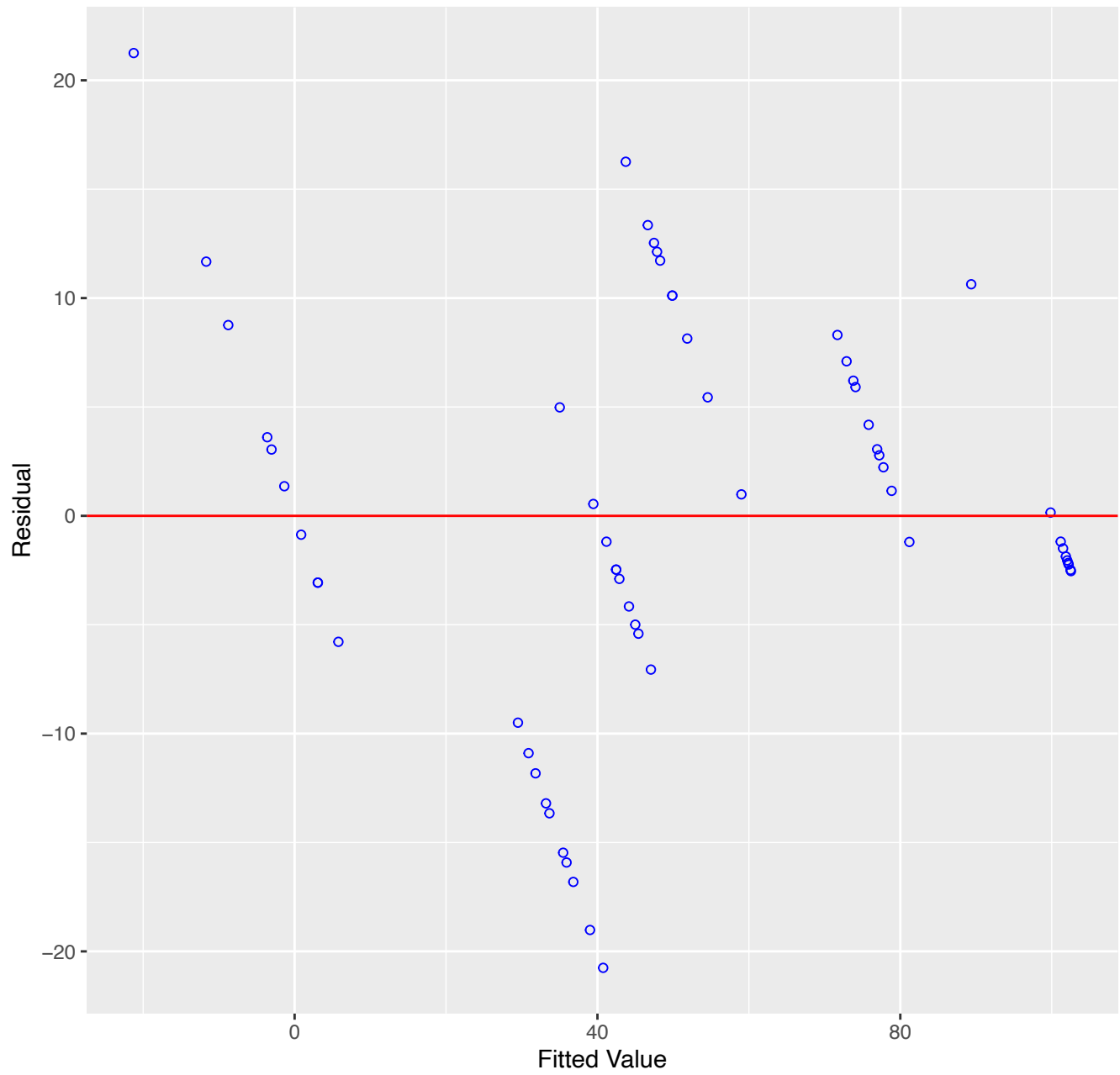
| | Estimate | Std. Error | t value | Pr(> t) |
|--------------------|----------|------------|---------|-----------|
| (Intercept) | -100.6 | 9.362 | -10.75 | 2.483e-15 |
| FER | 78.22 | 6.333 | 12.35 | 9.43e-18 |
| I(FER^2) | -7.525 | 0.966 | -7.79 | 1.538e-10 |

Table 4: Fitting linear model: Additive ~ FER + I(FER^2)

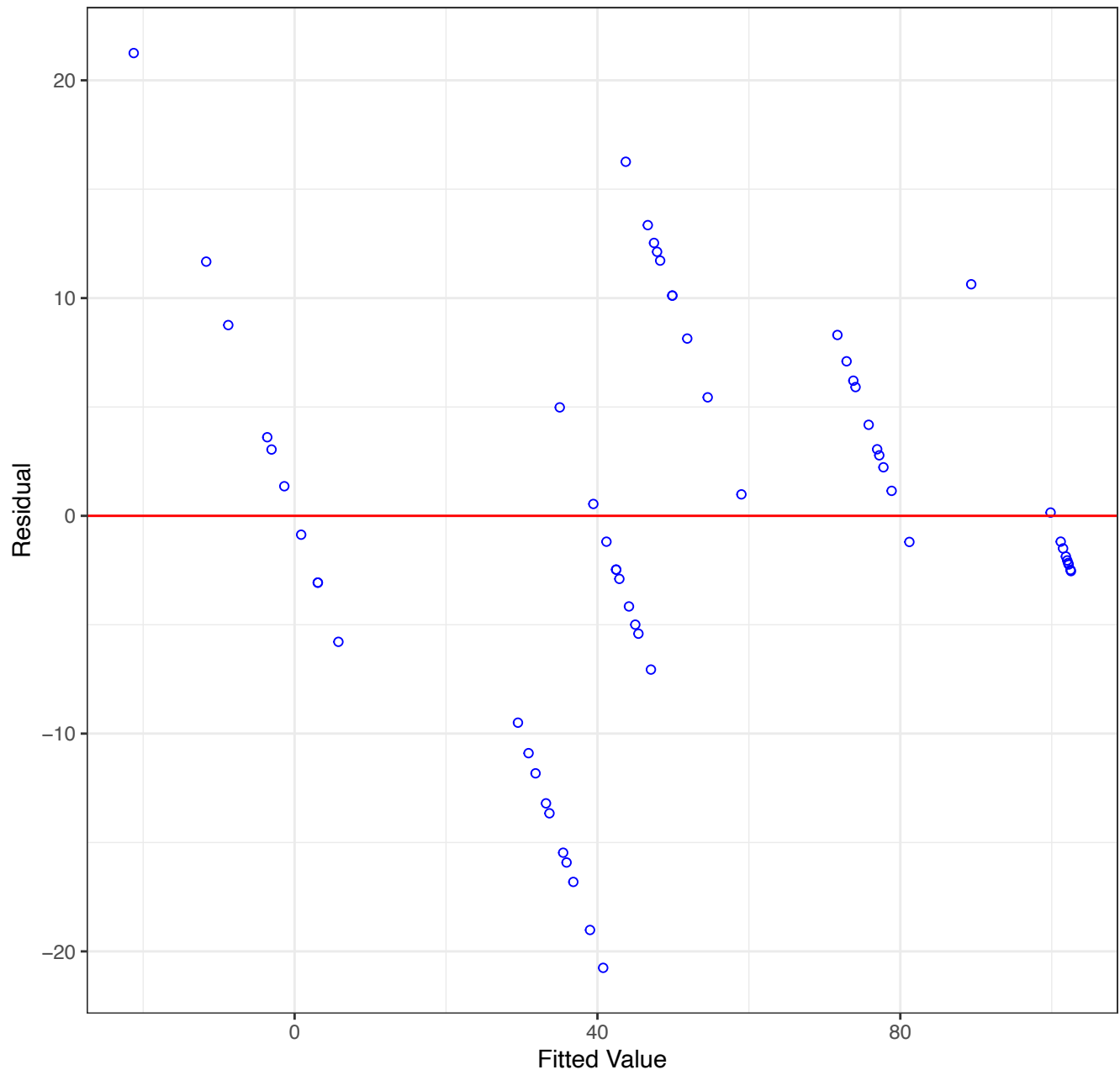
| Observations | Residual Std. Error | R^2 | Adjusted R^2 |
|--------------|---------------------|--------|----------------|
| 60 | 9.134 | 0.9321 | 0.9297 |

```
ols_plot_resid_fit(q_model)+theme_bw()
```

Residual vs Fitted Values

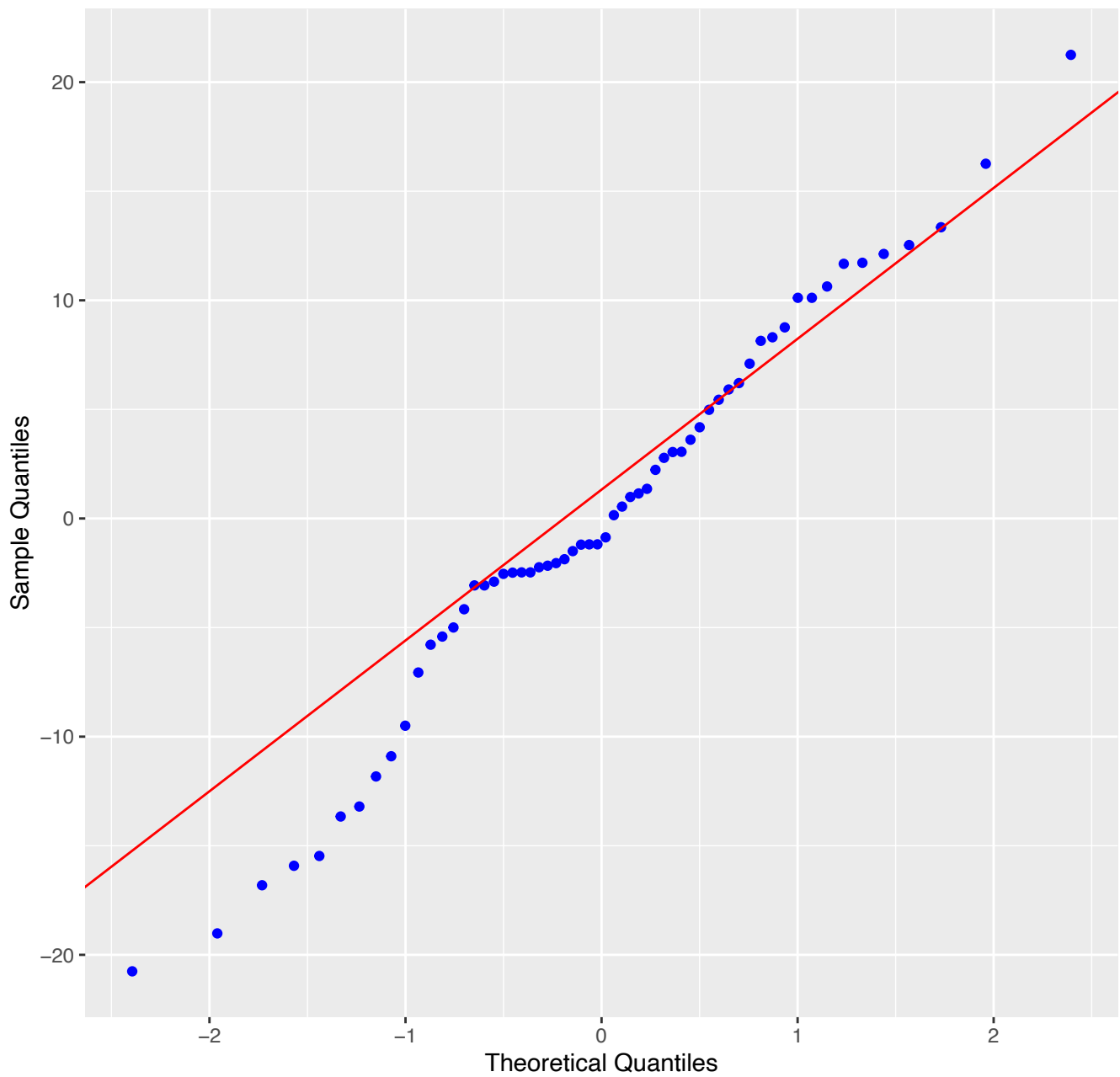


Residual vs Fitted Values



```
ols_plot_resid_qq(q_model)
```


Normal Q-Q Plot



```
ols_test_normality(q_model)
```

| Test | Statistic | pvalue |
|--------------------|-----------|--------|
| Shapiro-Wilk | 0.9808 | 0.4619 |
| Kolmogorov-Smirnov | 0.1163 | 0.3916 |
| Cramer-von Mises | 4.7082 | 0.0000 |
| Anderson-Darling | 0.4826 | 0.2222 |

```
#Cubic regression model
c_model<-lm(Additive~FER+I(FER^2)+I(FER^3),data=chicken_df)
summary(c_model)
```

Call:

```
lm(formula = Additive ~ FER + I(FER^2) + I(FER^3), data = chicken_df)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|----------|---------|--------|--------|---------|
| -17.3941 | -4.2858 | 0.0314 | 4.0400 | 19.2826 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|----------|------------|---------|--------------|
| (Intercept) | -12.7719 | 22.7726 | -0.561 | 0.577139 |
| FER | -26.4251 | 25.8822 | -1.021 | 0.311656 |
| I(FER^2) | 30.0682 | 9.1187 | 3.297 | 0.001698 ** |
| I(FER^3) | -4.0714 | 0.9833 | -4.141 | 0.000118 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8.063 on 56 degrees of freedom

Multiple R-squared: 0.948, Adjusted R-squared: 0.9452

F-statistic: 340.3 on 3 and 56 DF, p-value: < 2.2e-16

```
c_model%>%summary%>% pander()
```

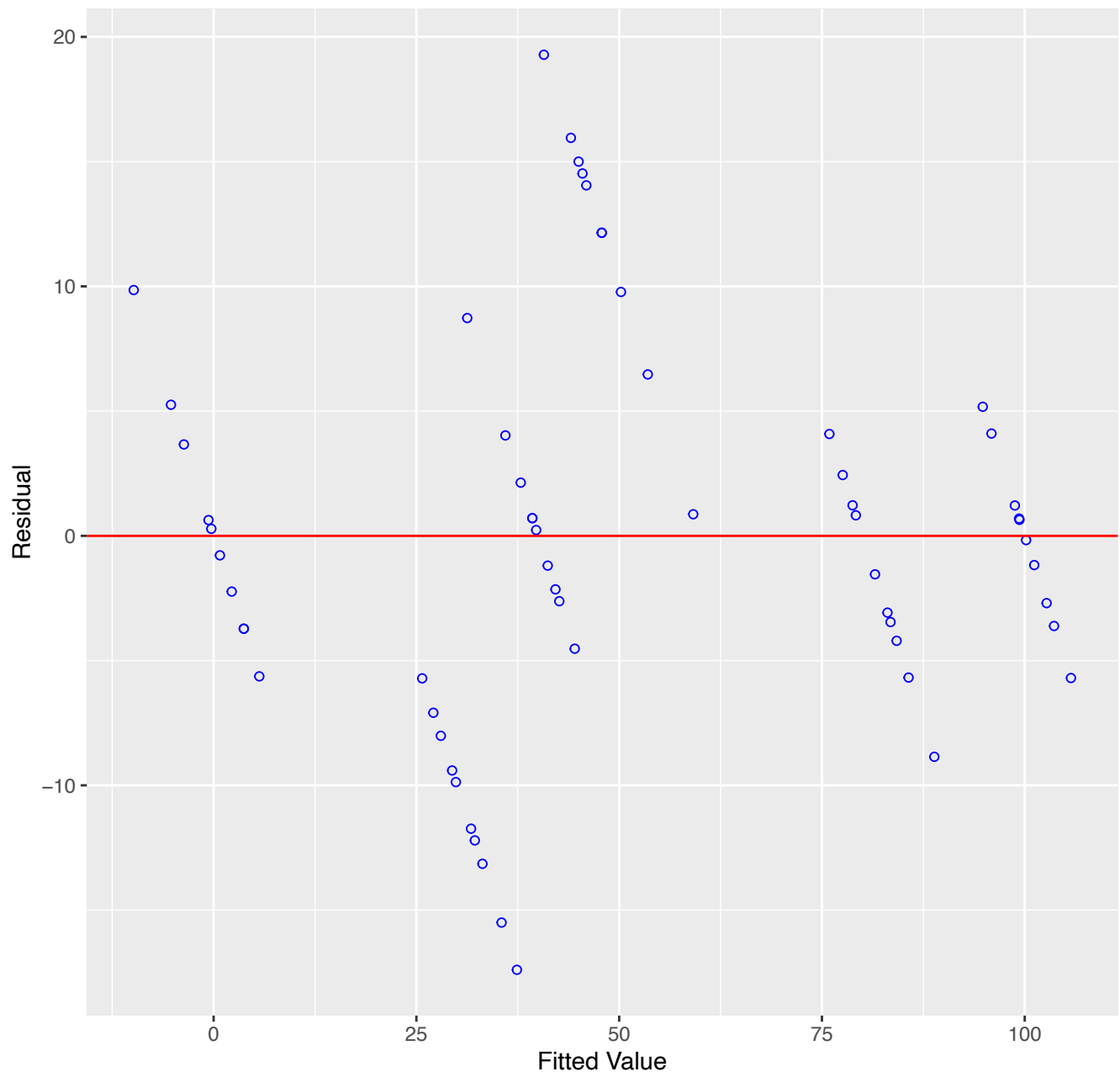
| | Estimate | Std. Error | t value | Pr(> t) |
|--------------------|----------|------------|---------|-----------|
| (Intercept) | -12.77 | 22.77 | -0.5608 | 0.5771 |
| FER | -26.43 | 25.88 | -1.021 | 0.3117 |
| I(FER^2) | 30.07 | 9.119 | 3.297 | 0.001698 |
| I(FER^3) | -4.071 | 0.9833 | -4.141 | 0.0001178 |

Table 6: Fitting linear model: Additive ~ FER + I(FER^2) + I(FER^3)

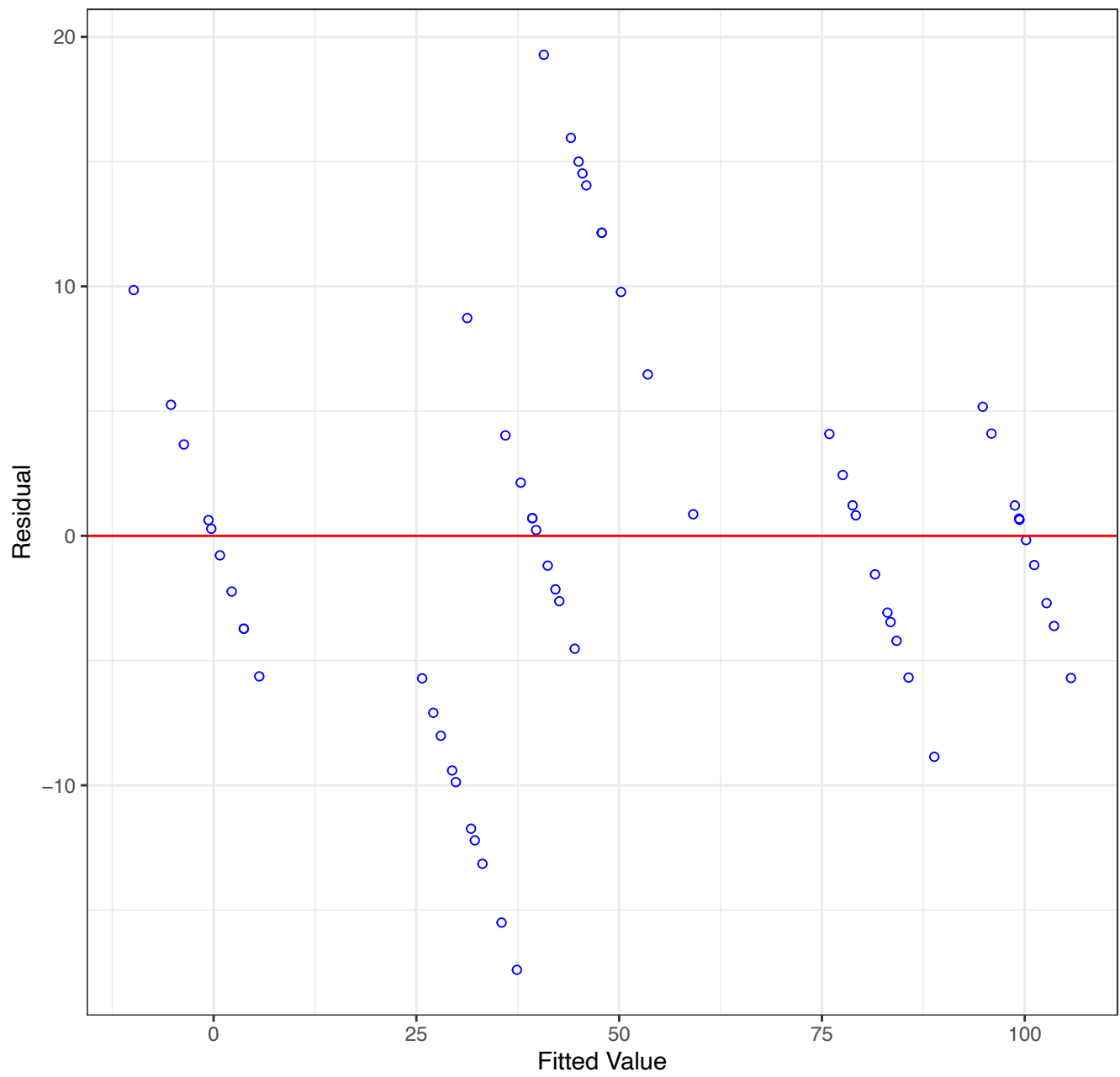
| Observations | Residual Std. Error | R^2 | Adjusted R^2 |
|--------------|---------------------|-------|----------------|
| 60 | 8.063 | 0.948 | 0.9452 |

```
ols_plot_resid_fit(c_model)+theme_bw()
```

Residual vs Fitted Values

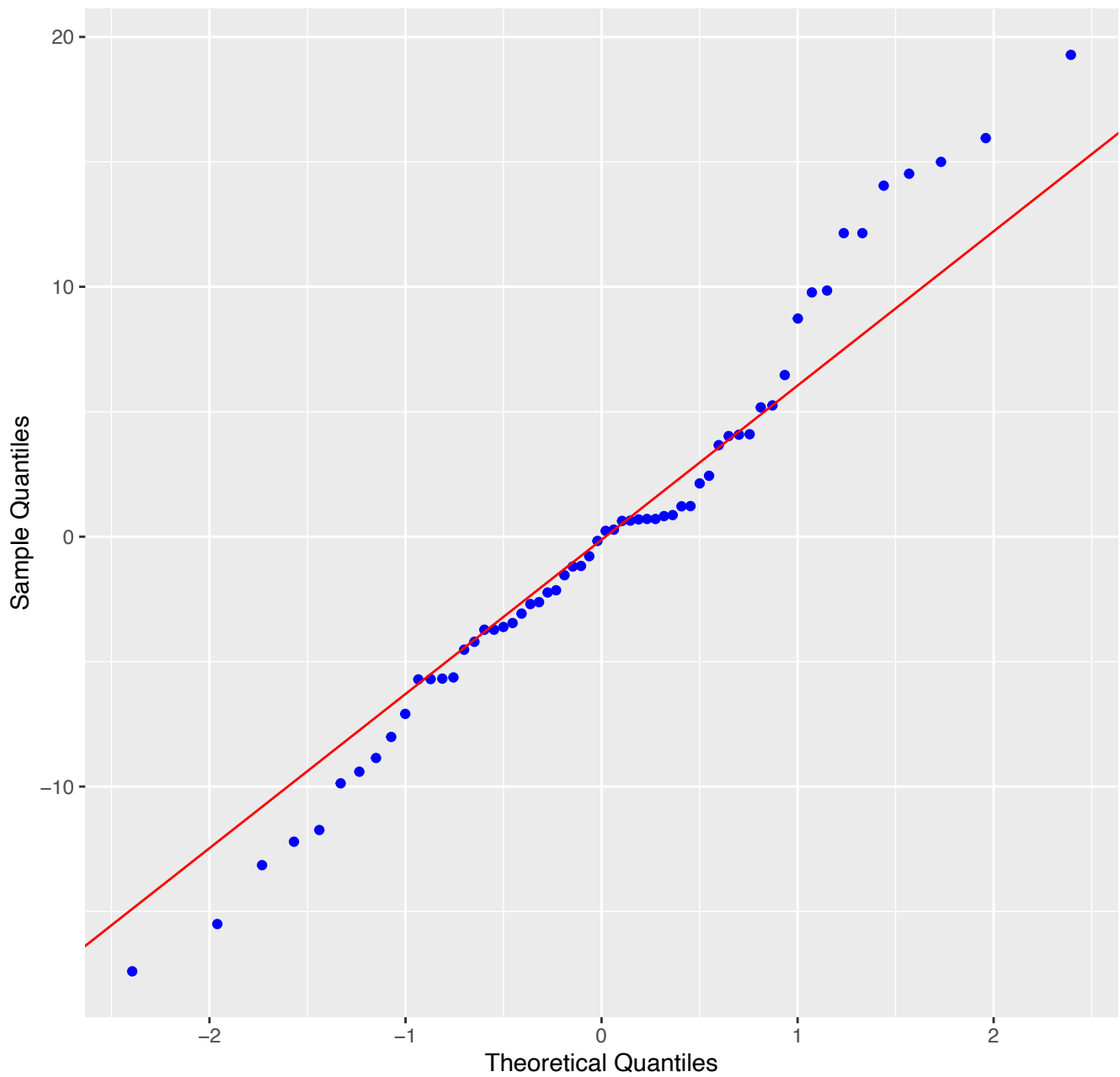


Residual vs Fitted Values



```
ols_plot_resid_qq(c_model)
```

Normal Q-Q Plot



```
ols_test_normality(c_model)
```

| Test | Statistic | pvalue |
|--------------------|-----------|--------|
| Shapiro-Wilk | 0.9773 | 0.3236 |
| Kolmogorov-Smirnov | 0.1214 | 0.3390 |
| Cramer-von Mises | 3.634 | 0.0000 |
| Anderson-Darling | 0.5943 | 0.1180 |

##I think cubic model provides the best fit since it has the least residual error and the line has the closest

- d. Is there anything peculiar about any of the data values? Provide an explanation of what may have happened. (Hint: Look at regression diagnostics like plots of the residuals versus the fitted values (or x), plot the leverages, or plot some measure of influence.)

##There appears to be a pattern aligning in the residuals vs fitted values which is very peculiar since it is

- e. Using your best polynomial model from (b) & (c). Fit a new model that includes the linear addition of copper and display the estimate table. Does Copper provide a significant improvement to the fit? Carry out an F-test that compares the Full model that contains Copper and the reduced model that has your polynomial model fit on the additive only. Discuss the results.

#First-order regression model using Additive, FER and Copper

```
new_modfit<-lm(Additive~FER+Copper,data=chicken_df)
summary(new_modfit)
```

Call:

```
lm(formula = Additive ~ FER + Copper, data = chicken_df)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|----------|----------|--------|---------|---------|
| -18.9476 | -12.1329 | 0.2289 | 13.5052 | 20.0136 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|------------|------------|---------|--------------|
| (Intercept) | -31.256799 | 4.938201 | -6.330 | 4.13e-08 *** |
| FER | 29.684656 | 1.579631 | 18.792 | < 2e-16 *** |
| Copper | -0.006085 | 0.008439 | -0.721 | 0.474 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 13.06 on 57 degrees of freedom

Multiple R-squared: 0.861, Adjusted R-squared: 0.8561

F-statistic: 176.6 on 2 and 57 DF, p-value: < 2.2e-16

```
summary(f_model)
```

Call:

```
lm(formula = Additive ~ FER, data = chicken_df)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|----------|----------|---------|---------|---------|
| -18.5200 | -12.0502 | -0.5643 | 13.0942 | 21.2142 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|----------|------------|---------|--------------|
| (Intercept) | -32.352 | 4.679 | -6.914 | 4.08e-09 *** |
| FER | 29.641 | 1.572 | 18.856 | < 2e-16 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 13.01 on 58 degrees of freedom

Multiple R-squared: 0.8598, Adjusted R-squared: 0.8573

F-statistic: 355.6 on 1 and 58 DF, p-value: < 2.2e-16

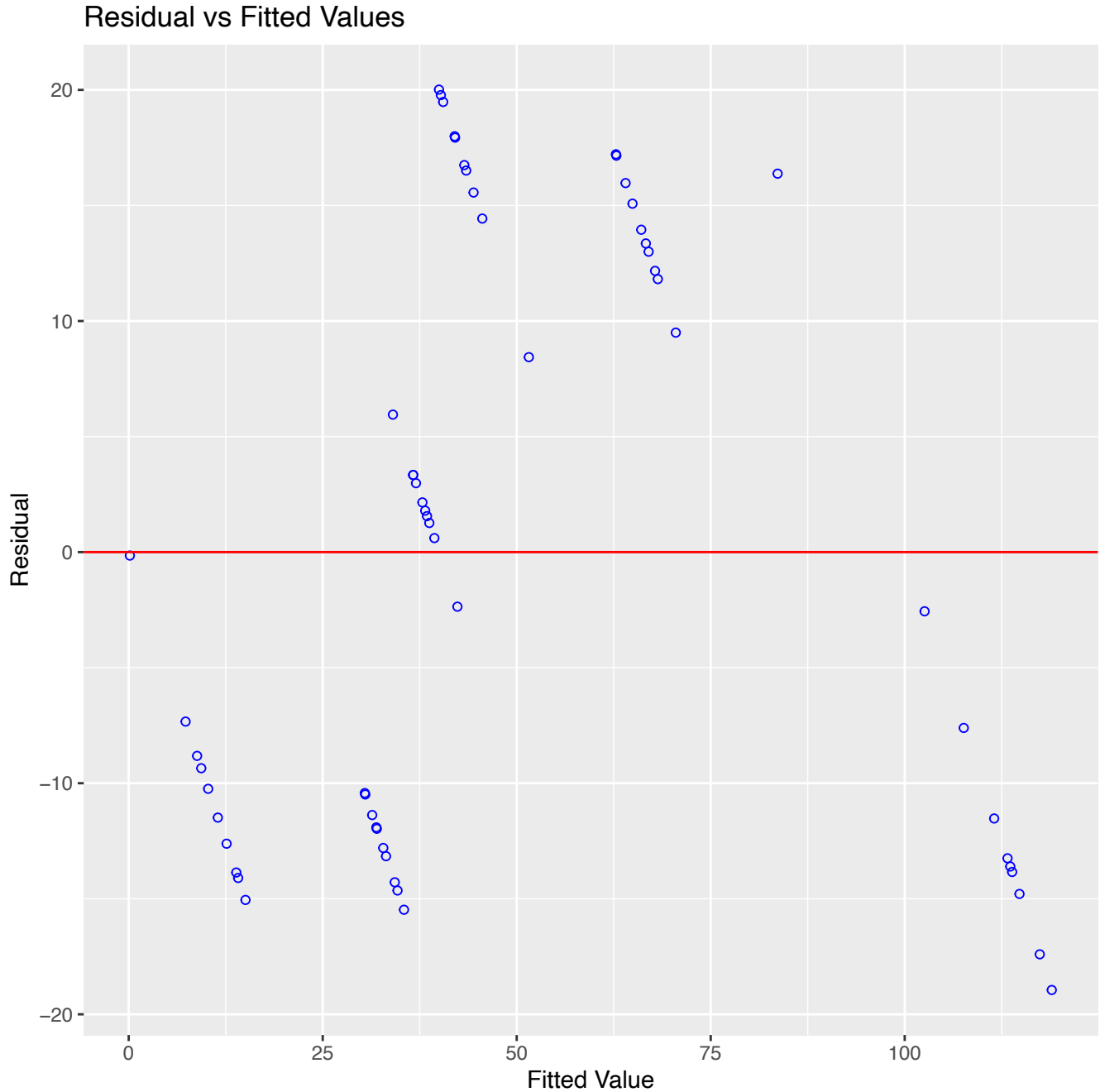
```
new_modfit%>%summary%>% pander()
```

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|-----------|------------|---------|-----------|
| (Intercept) | -31.26 | 4.938 | -6.33 | 4.127e-08 |
| FER | 29.68 | 1.58 | 18.79 | 4.238e-26 |
| Copper | -0.006085 | 0.008439 | -0.7211 | 0.4738 |

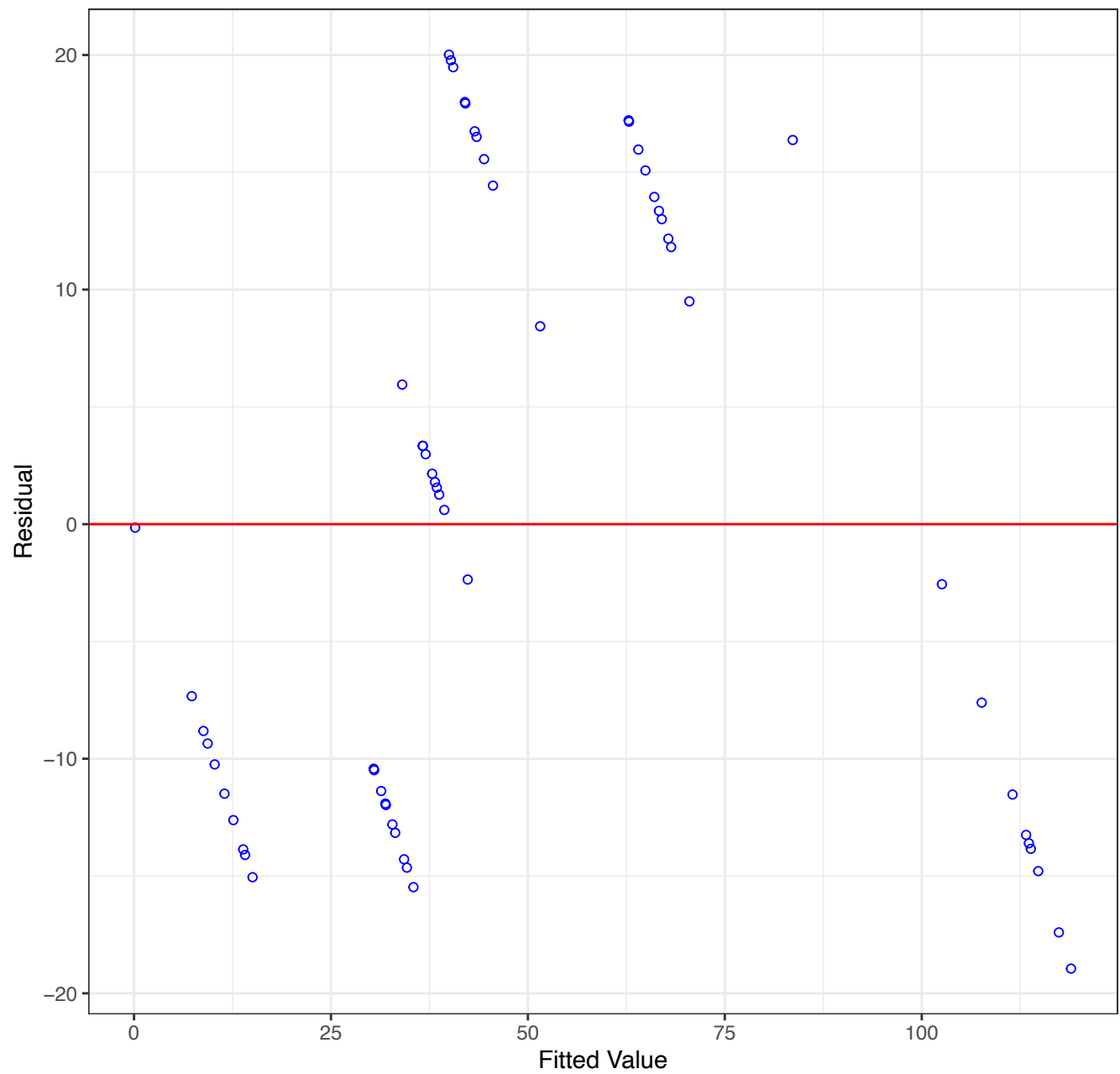
Table 8: Fitting linear model: Additive ~ FER + Copper

| Observations | Residual Std. Error | R^2 | Adjusted R^2 |
|--------------|---------------------|-------|----------------|
| 60 | 13.06 | 0.861 | 0.8561 |

```
ols_plot_resid_fit(new_modfit)+theme_bw()
```

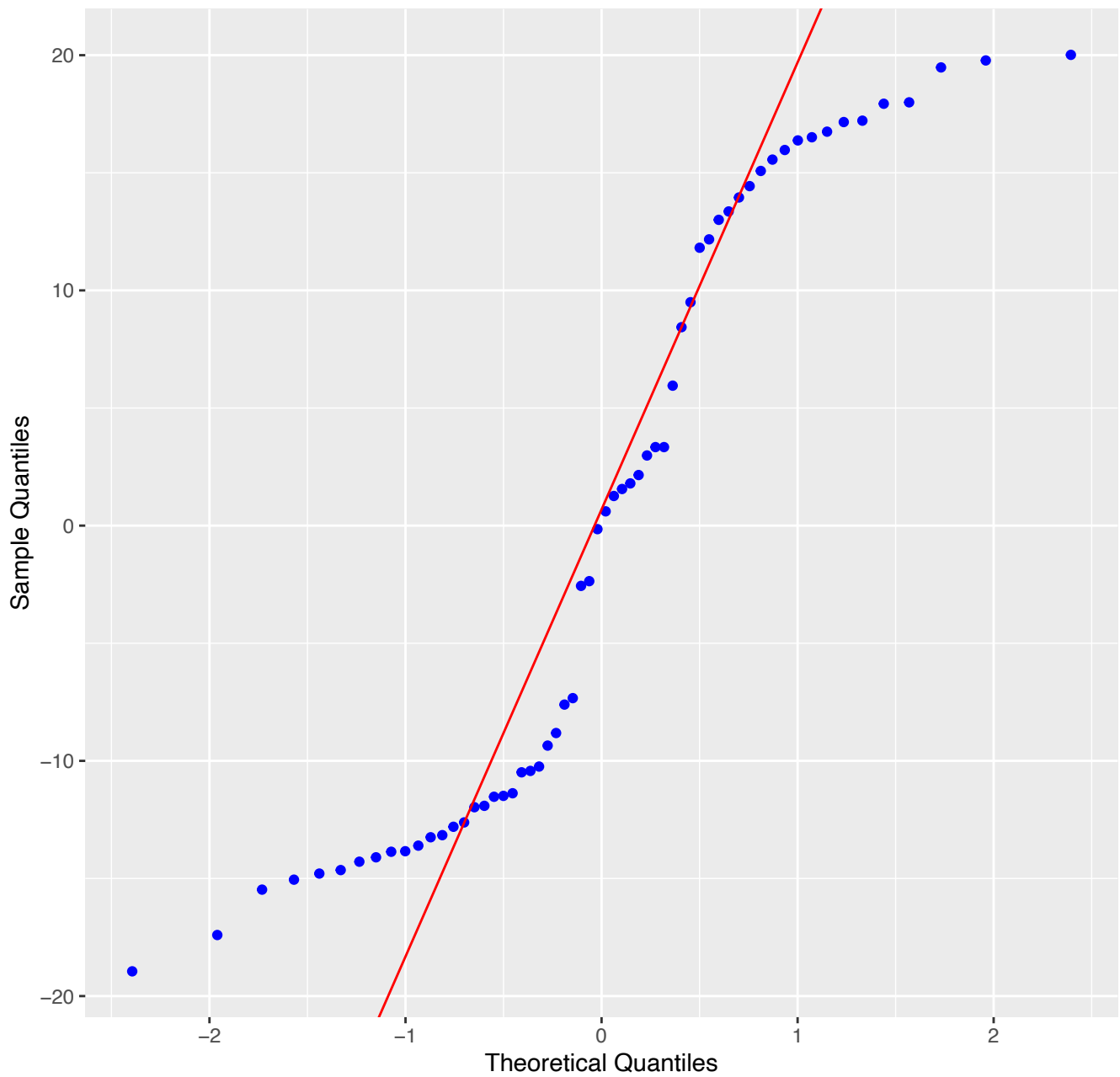


Residual vs Fitted Values



```
ols_plot_resid_qq(new_modfit)
```


Normal Q-Q Plot



```
ols_test_normality(new_modfit)
```

| Test | Statistic | pvalue |
|--------------------|-----------|--------|
| Shapiro-Wilk | 0.8835 | 0.0000 |
| Kolmogorov-Smirnov | 0.1708 | 0.0604 |
| Cramer-von Mises | 4.7681 | 0.0000 |
| Anderson-Darling | 2.5735 | 0.0000 |

#Copper does cluster the points more in the middle and at the ends on the line but statistically speaking, it

Problem 4: [10 pts] Linear Regression with Indicator Variables

Consider the data in `smoking_birthweight.csv`. This data contains 3 variables. The birth weight of a baby (`Weight`), the length of gestation (`Gestation`) in weeks, and the smoking status of the mother (`Smoke`). The smoking status of the mother in this case is coded as `yes` or `no`. This is a categorical variable (aka factor) with 2 categories (a binary variable). We could have coded the levels of this factor as an indicator variable using `TRUE` or `FALSE`, or equivalently 1 or 0, respectively.

- Fit a first-order regression model with birth weight as the response variable and the gestation and smoking status as predictors. Write down the fitted regression model equation and interpret the regression coefficients. If you can do this, you should have no problem handling the extra credit.

```
library(dplyr)
##Read the file
smokebw<-read.csv("/Users/eduardosalvador/Desktop/FINAL\ Spring\ Semester\ 2021/CMDA\ /Assignments/HW6/smoking

#Looking at first six rows
head(smokebw)

  Weight Gestation Smoke
1   2940         38   yes
2   3130         38    no
3   2420         36   yes
4   2450         34    no
5   2760         39   yes
6   2440         35   yes

#Looking for person that is smoker
smokebw<-smokebw %>% mutate(mystatus=if_else(Smoke=="yes",1,0))

#Making a fitted generalized linear model for weight with mystatus and gestation
glm_smoke<-glm(Weight~mystatus+Gestation,data=smokebw)

summary(glm_smoke)

Call:
glm(formula = Weight ~ mystatus + Gestation, data = smokebw)

Deviance Residuals:
    Min       1Q   Median       3Q      Max
-223.693   -92.063    -9.365    79.663   197.507

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -2389.573    349.206  -6.843 1.63e-07 ***
mystatus      -244.544     41.982  -5.825 2.58e-06 ***
Gestation     143.100      9.128  15.677 1.07e-15 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for gaussian family taken to be 13347.24)

    Null deviance: 3735790  on 31  degrees of freedom
Residual deviance:  387070  on 29  degrees of freedom
AIC: 399.63

Number of Fisher Scoring iterations: 2

## My interpretation is that if mother were to smoke, the change in weight decreases by 244.544 with 41.982 er
```

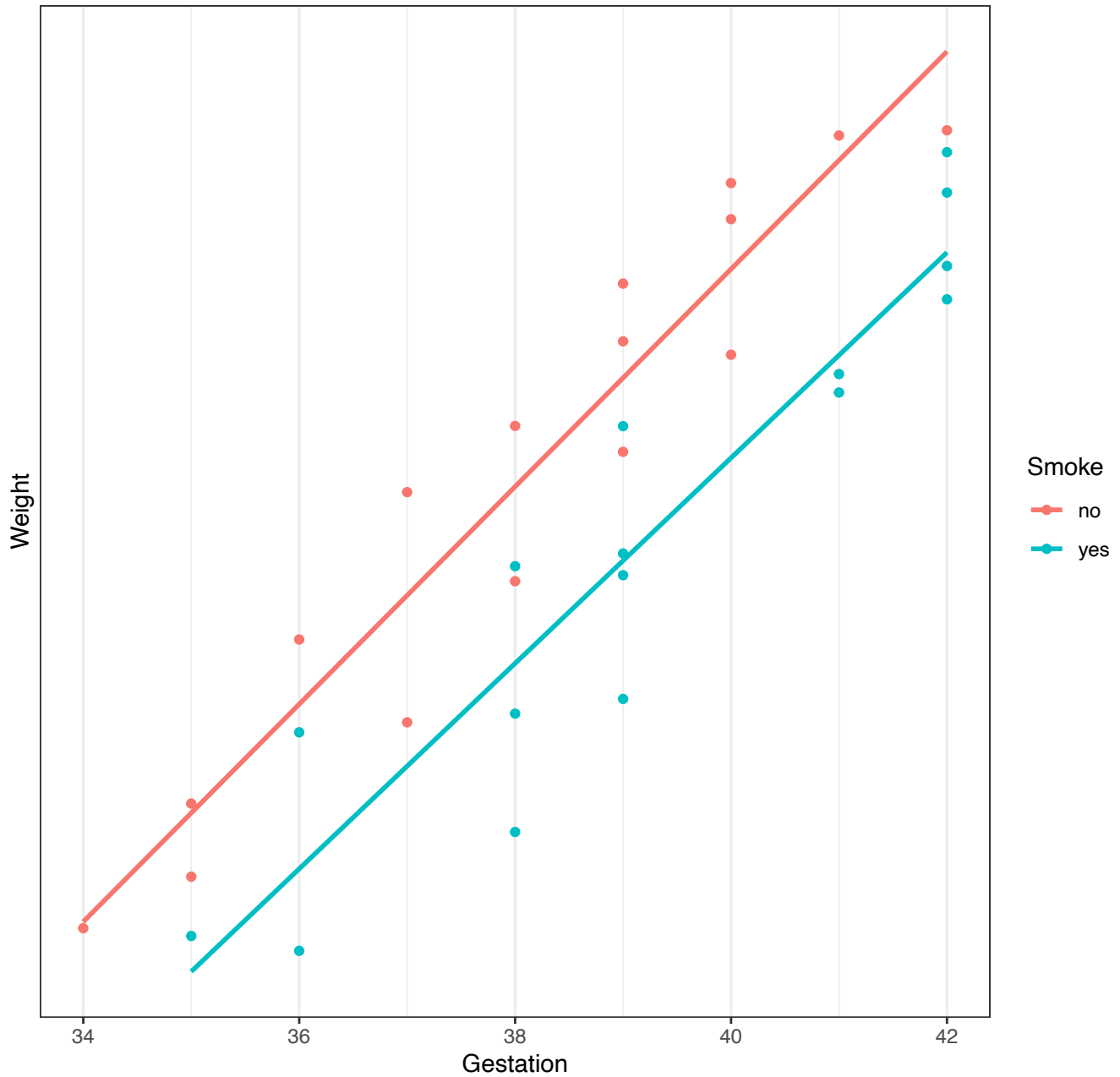
?glm b. Plot the fitted regression lines (yes plural), why are there two?

```
library(ggplot2)
```

```
##GGplot to create fitted regression lines
```

```
ggplot(smokebw,aes(x=Gestation,y=Weight,color=Smoke))+theme_bw()+ggtitle("Fitted Regression Lines")+geom_point
```

Fitted Regression Lines



```
##There are two lines because there is moms that smoke and no smoke
```

Problem 5: [15 pts Extra Credit] Parameter Interpretation with Indicator Variables

Recall that indicator variables, sometimes called “dummy” variables, are binary variables that indicate whether an event is recognized or not (i.e., 1 if TRUE 0 if FALSE). Suppose we have a data set of reported salaries and highest achieved education levels. Suppose the variables are as follows: **salary**, **noHS**, **highSchoolGrad**, **Assoc**, **Bach**, **Masters**, **Doctorate**, where the levels of education are either a 1 or 0 depending on whether that is the given observation’s highest level of achieved education.

- Write down the multiple linear regression model. Specify which β_i are indicator variables.

```
##salary=$\beta_{0}$+$\beta_{1}$noHS+$\beta_{2}$highSchoolGrad+$\beta_{3}$Assoc+$\beta_{4}$Bach+$\beta_{5}$Mas
```

- Write interpretations for all of your model parameters, that is β_i , for $i \in \{0, 1, 2, 3, 4, 5\}$.

```
##$\beta_{i}$ always increases the salary with each increase in education. For example, for highSchoolGrad, the
```

- Now assume we were to add another variable to this data set: an observation’s **gender**. Write down this new model, and now interpret β_0 .

```
##salary=$\beta_{0}$+$\beta_{1}$noHS+$\beta_{2}$highSchoolGrad+$\beta_{3}$Assoc+$\beta_{4}$Bach+$\beta_{5}$Mas
##$\beta_{0}$ is the salary that it doesn't matter your level of education or gender, it is the lowest amount
```
