#### **Lecture 11 ANALYSIS OF VARIANCE**

#### **Objective:**

- ANOVA technique used for Regression
- One way ANOVA for categorical explanatory variables
- Two way ANOVA/Multiway for two/more categorical variable.
- ANCOVA used for dummy variable model.

#### Introduction

- The technique of ANOVA was used in conjunction with Least square Regression to interpret the total variation of the data from its mean.
- Traditionally ANOVA technique has also been used for fitting Linear models which are composed of only categorical explanatory variables.
- If the model is composed of only one factor then it is named "One way ANOVA" wheras a model composed of two factors is named "Two way ANOVA".
- An alternative model ANCOVA is composed of quantitative as well as qualitative variables.
   The acronym "ANCOVA" expands to Analysis of Covariance. The dummy variables model was an interpretation and implementation of this model.

#### One Way Analysis of Variance

This is the model that is composed of one categorical/qualitative regressor (factor). The factor could have more than one category.

This classification (one factor with 3 categories) can be represented as:

 $Y_i = \alpha + \gamma_1 D_{i1} + \gamma_2 D_{i2} + \varepsilon_i$  Remember if there are 3 categories we need 2 dummy variables.

Group	D1	D2
1	1	0
2	0	1
3	0	0

The expectation of each response variable is the population mean. Recollect the expectation of the error term is zero. The three equation can be reformulated as follows:

Group 1 
$$\mu_1 = \alpha + \gamma_1(1) + \gamma_2(0)$$
  
 $\mu_1 = \alpha + \gamma_1$   
Group 2  $\mu_2 = \alpha + \gamma_1(0) + \gamma_2(1)$   
 $\mu_2 = \alpha + \gamma_2$   
Group 3  $\mu_3 = \alpha + \gamma_1(0) + \gamma_2(0)$ 

$$\mu_3 = \alpha$$

Solving the equations gives us:

 $\alpha = \mu_3$  Alpha captures the mean of the baseline group.

$$\gamma_1 = \mu_1 - \mu_3$$

$$\gamma_2 = \mu_2 - \mu_3$$

 $\gamma_{1}$  and  $\,\gamma_{2}$  are differences of means of other groups.

The One way Analysis of Variance tests the difference of means amongst the groups of the factor.

#### The omnibus F statistics tests

Null Hypothesis  $H_0:\mu_1=\mu_2=\mu_3$  No difference of the population means for the three groups of the factor

This hypothesis corresponds to H<sub>0</sub>:  $\gamma_1 = \gamma_2 = \gamma_3$ 

## **ALTERNATE MODEL**

The ANOVA model can be formulated as follows:

Analyzing data using a single factor randomized experiment

# One Way ANOVA

<b>Treatment</b> Factor	Observations								
GROUPS								Totals	Average
	1	<b>y</b> 11	<b>y</b> 21	<b>y</b> 31			<b>y</b> j1	y <sub>•1</sub>	$\overline{y_1}$
	2	<b>y</b> 12	<b>y</b> 22	<b>y</b> 32			<b>y</b> j2	$y_{\bullet 2}$	$\overline{y_2}$
	m	y <sub>1m</sub>	<b>y</b> 2m	<b>y</b> 3m			$\mathbf{Y}_{jm}$	$y_{\bullet m}$	$\overline{\mathcal{Y}}_m$
								y <b></b>	$\overline{y_{\bullet \bullet}}$

Suppose we have m levels (treatments/groups) of one factor.

 $Y_{ij}$  Is the ith observation of the jth group where total groups are m. This response  $Y_{ij}$  treatment is a random variable.

 $n_i$  is the total number of observations in the jth group.

$$n = \sum_{i=1}^{m} n_{j}$$
 Overall Total number of observations.

$$\mu_{i} = E(Y_{ij})$$

The observation in the above table can be represented by the following linear statistical model:

$$Y_{ij} = \mu + \alpha_j + \varepsilon_{ij} \quad i = 1.....m$$

$$j = 1.....J$$

 $Y_{ij}$  is a response variable denoting (ij th) observation.

μ is the population mean

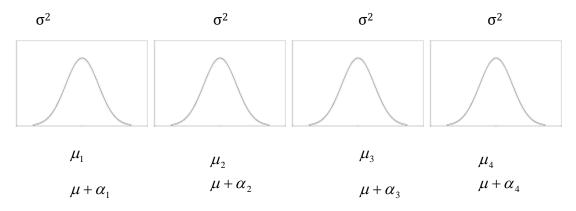
 $\alpha_j$  is the effect of the jth group

 $\in_{ij}$  is the random error component

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$

Where  $\mu_j = \mu + \alpha_j$  the mean of the jth treatment equals the overall mean plus the treatment effect for j th treatment.

We assume by the linear model assumptions  $\in_{ij}$  (random error component) is normally and independently distributed with mean zero and variance  $\sigma^2$ . Each treatment can assumed to be normally distributed with mean  $\mu_i$  and variance  $\sigma^2$ .



Here the treatment have been chosen therefore it is called a fixed effect model (explanatory variable is non random). If the treatment is chosen at random then it is a non fixed effect model whose results can be generalized to treatments in the population.

Here we check the variability of  $\alpha_j$ 

We are creating an analysis of variance for a fixed effect model.

For a fixed effect model treatment effects  $\alpha_j$  are the deviations from the overall mean.

Therefore the sigma constraint or sum to zero constraint(sum will be above the overall eamn some will b below the overall mean):

$$\sum_{i=1}^{m} \alpha_{j} = 0$$

Hypothesis testing: Testing the equality of a treatment means.

Null Hypothesis  $H_0: \mu_1 = \mu_2 = \mu_3$ 

This can be equivalently written as

 $H_0: \alpha_1 = \alpha_2 = \alpha_3 = \dots \alpha_I = 0$ 

 $H_1$ : Atleast one of the  $\alpha_i \neq 0$ 

m=number of categories in a factor n=number of

Sources of	Sum of	Degree	Mean Squared	F	H <sub>0</sub>
Variation	Squares	of			
		freedom			
Treatment	$\sum n_j (\overline{Y_j} - \overline{Y})^2$	m-1	$\frac{\operatorname{Re} gSS}{m-1} = \operatorname{Re} gMS$	Re gMS	$\alpha_1 = \alpha_2 = \dots = \alpha_m = 0$
Groups			m-1	RMS	$\alpha_1 = \alpha_2 = \dots = \alpha_m = 0$ $\mu_1 = \mu_2 = \dots = \mu_m = 0$
Within	$\sum \sum (Y_{ij} - \overline{Y_j})^2$	n-m	$\frac{RSS}{}=RMS$		
groups			n-m		
Error					
Residuals					
Total	$\sum \sum (Y_{ij} - \overline{Y})^2$	n-1			

# **Example**

Ho: Mean of tensile strengths of hardwoods are not significantly different.

Hardwood	Observations							
Concentration								
%(Treatments)	1	2	3	4	5	6	<b>Totals</b>	Average
5	7 y <sub>11</sub>	8 y <sub>12</sub>	15	11	9	10	60	<b>10</b> $\overline{y_1}$
10	12	17	13	18	19	15	94	15.67
15	14	18	19	17	16	18	102	17
20	19	25	22	23	18	20	127	21.17
							383	15.96
								<u>y</u>

Total 
$$\sum \sum (Y_{ij} - \overline{Y})^2 = (7-15.96)^2 + (8-15.96)^2 + \dots (20-15.96)^2 = 512.96$$
  
Between  $\sum n_j (\overline{Y_j} - \overline{Y})^2 = (10-15.96)^2 + (15.67-15.96)^2 + \dots (21.17-15.96)^2 = 382.79$   
Within  $\sum \sum (Y_{ij} - \overline{Y_j})^2 = (7-10)^2 + \dots (10-10)^2 + (12-15.67)^2 + \dots (20-21.17)^2 = 130.17$ 

Sources of	Sum of Squares	Degree of	Mean Squared	F
Variation		freedom		Re gMS
				RMS
				19.60
Treatment	SS <sub>B</sub> 382.79	m-1 4-1 =3	$\frac{\operatorname{Re} gSS}{m-1} = \operatorname{Re} gMS$	
Group	Sum of Square		$\frac{m-1}{m-1}$ = Re gMS	
	Deviation		=382.79/3	
	Between		=127.60	
	groups			
Within	SSw 130.17	n-m=24-4=20	$\frac{RSS}{}=RMS$	
groups	Sum of Square			
Error	Deviation		=130.17/20	
	within factor		=6.51	
Total	SS <sub>T</sub> 512.96	24-1=23		

Taking  $\alpha = .01 \text{ P}(F_{3,20} > 19.60) = 3.59*10^{-6} \text{ which is considerably smaller than .01.}$ 

We have strong evidence to reject  $H_0$ . Therefore our test is significant and we can conclude that Hardwood concentration has an effect on tensile strength.

### TWO WAY ANALYSIS CLASSIFICATION

For a two way ANOVA we have two factors. The classification table can be represented as follows:

The means are the means for the response variables for the level of categorical variable C and that of categorical variable R.

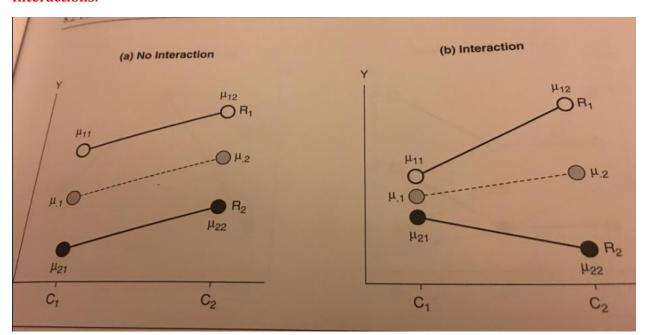
	C <sub>1</sub>	$C_2$	 C <sub>c</sub>	
$R_1$	$\mu_{11}$	$\mu_{12}$	$\mu_{1c}$	$\mu_{1.}$
				Marginal mean R1
$R_2$	$\mu_{21}$	$\mu_{22}$	$\mu_{2c}$	$\mu_{2.}$
R <sub>r</sub>	$\mu_{\rm r1}$	$\mu_{\rm r2}$	$\mu_{rc}$	$\mu_{r.}$
	μ.1	μ.2	 μ <sub>.c</sub>	μ
	Marginal mean C1			

**Interactions**: Informs the change of response variable by the interaction of R and C. If the lines are parallel then R and C are interacting to affect the response variable Y.

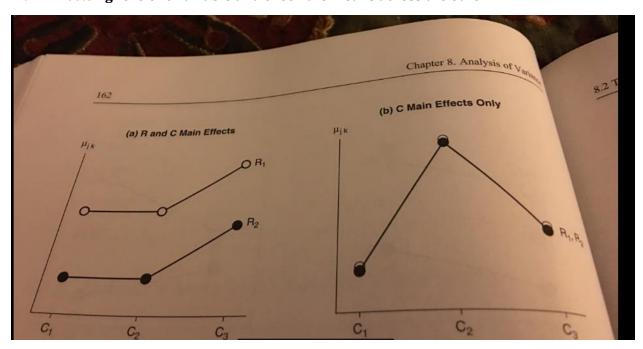
**Main effects**: Are the partial effects of R or C or both. These are the difference of the marginal means and the grand mean.

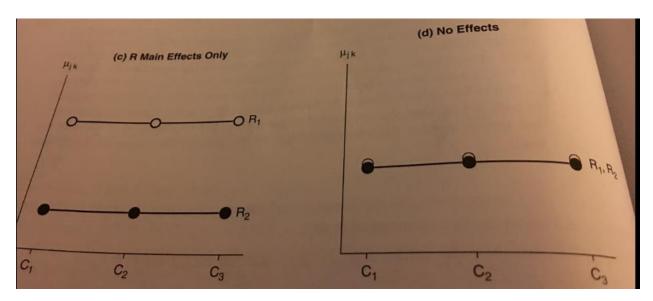
## **Visualizations**

#### **Interactions:**



Main Effects: Ignore one variable and check the means across the other.





Two way Model and Hypothesis Testing

**Interaction Hypothesis** 

$$H_0$$
 :  $\mu_{jk} - \mu_{jk^*} = \mu_{j^*k} - \mu_{j^*k^*}$  No Interactions

Columns effects are invariant across rows

Main Effects Hypothesis. Equality of Marginal Means

#### **Row Classification:**

$$H_0: \mu_{1*} = \mu_{2*}.... = \mu_{r*}$$

#### Column Classification:

$$H_0: \mu_{*_1} = \mu_{*_2}.... = \mu_{*_r}$$

## Generally the main effects hypothesis is of consequence if there is no interactions.

The model can be represented as follows:

$$Y_{ijk} = \mu + \alpha_j + \beta_k + \gamma_{jk} + \varepsilon_{ijk}$$

 $Y_{iik}$  is the ith observation of the jth row and kth column of the RC table.

 $\mu$  is the general mean of Y.

 $\alpha_i, \beta_k$  Are the main effects parameters for row effects and column effects.

 $\gamma_{ik}$  Is the interaction term

 $\varepsilon_{iik}$  Is the random error.

# Three way ANOVA Classification as well as Multiway ANOVA classification can be conducted in a similar manner.

Three Way Hypothesis Formula for factor A (Factor B and C will have similar Hypothesis):

Interaction and Main effects Hypothesis

A Main effects

$$H_0: \alpha_A = 0$$

$$H_{\scriptscriptstyle 0}:\alpha_{\scriptscriptstyle A}=0\,|\,\alpha_{\scriptscriptstyle AB}=\alpha_{\scriptscriptstyle AC}=\alpha_{\scriptscriptstyle ABC}$$

Interactions

AB

$$H_0: \alpha_{AB} = 0$$

$$H_0:\alpha_{AB}=0\,|\,\alpha_{AB}=\alpha_{AC}=\alpha_{ABC}$$

**ABC** 

$$H_0: \alpha_{ABC} = 0$$

Three Way classification has the highest order interaction is equal to the number of factors in the model. It is not mandatory that the model is composed of all higher level factors.

# **ANCOVA**

This model is Analysis of Covariance. This model is identical to the Dummy variable Model. The formulation of the model is different though the interpretation is the same. R implements it by the same command anova().