

STA409

Answer to Assignment 4

12112627

李乐平

1. Solution.

(1).

$$\begin{aligned} & \sum_{i=1}^a \sum_{j=1}^b n_{ij} E^2(\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y}) \\ &= \sum_{i=1}^a \sum_{j=1}^b n_{ij} (\mu + \alpha_i + \beta_j + \gamma_{ij} - (\mu + \alpha_i) - (\mu + \beta_j)) \\ &= \sum_{i=1}^a \sum_{j=1}^b n_{ij} \gamma_{ij}. \end{aligned}$$

$$\begin{aligned} & \sum_{i=1}^a \sum_{j=1}^b n_{ij} \text{Var}(\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y}) \\ &= \sum_{i=1}^a \sum_{j=1}^b n_{ij} [\\ & \quad \text{Var}(\bar{Y}_{ij}) + \text{Var}(\bar{Y}_{i\cdot}) + \text{Var}(\bar{Y}_{\cdot j}) + \text{Var}(\bar{Y}) \\ & \quad - 2(\text{Cov}(\bar{Y}_{ij}, \bar{Y}_{i\cdot}) + \text{Cov}(\bar{Y}_{ij}, \bar{Y}_{\cdot j}) + \text{Cov}(\bar{Y}_{ij}, \bar{Y})) \\ & \quad + 2(\text{Cov}(\bar{Y}_{i\cdot}, \bar{Y}_{\cdot j}) + \text{Cov}(\bar{Y}_{i\cdot}, \bar{Y}) + \text{Cov}(\bar{Y}_{\cdot j}, \bar{Y})) \\ & \quad] \\ &= \sum_{i=1}^a \sum_{j=1}^b n_{ij} \left[\frac{\sigma^2}{n_{ij}} + \frac{\sigma^2}{n_{i\cdot}} + \frac{\sigma^2}{n_{\cdot j}} + \frac{\sigma^2}{n} - 2\left(\frac{\sigma^2}{n_{i\cdot}} + \frac{\sigma^2}{n} + \frac{\sigma^2}{n_{\cdot j}} + \frac{\sigma^2}{n} - \frac{\sigma^2}{n} - \frac{n_{ij}\sigma^2}{n_{i\cdot}n_{\cdot j}}\right) \right] \\ &= \sum_{i=1}^a \sum_{j=1}^b n_{ij} \left(\frac{\sigma^2}{n_{ij}} - \frac{\sigma^2}{n_{i\cdot}} - \frac{\sigma^2}{n_{\cdot j}} - \frac{\sigma^2}{n} + 2\frac{n_{ij}\sigma^2}{n_{i\cdot}n_{\cdot j}} \right) \\ &= \sigma^2(ab - a - b + 1 + 2\sum_{i=1}^a \sum_{j=1}^b \frac{n_{ij}^2}{n_{i\cdot}n_{\cdot j}}). \end{aligned}$$

Because n_{ij} 's are all the equal,

$$\sum_{i=1}^a \sum_{j=1}^b n_{ij} \text{Var}(\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y}) = (a-1)(b-1)\sigma^2.$$

$$\text{SSAB} = \sum_{i=1}^a \sum_{j=1}^b n_{ij} (\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y})^2.$$

$$E(\text{SSAB}) = \sum_{i=1}^a \sum_{j=1}^b n_{ij} (E^2(\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y}) + \text{Var}(\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y}))$$

$$= (a-1)(b-1)\sigma^2 + \sum_{i=1}^a \sum_{j=1}^b n_{ij} \gamma_{ij}^2.$$

(2).

When n_{ij} 's are all equal (say $n_{ij} = n_0, \forall i, j$), we have

$$\begin{aligned}
& \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{i\cdot} - \bar{y})(\bar{y}_{\cdot j} - \bar{y}) = n_0 \sum_{i=1}^a (\bar{y}_{i\cdot} - \bar{y}) \sum_{j=1}^b (\bar{y}_{\cdot j} - \bar{y}) = 0 \\
& \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{i\cdot} - \bar{y})(\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) = n_0 \sum_{i=1}^a (\bar{y}_{i\cdot} - \bar{y}) \sum_{j=1}^b (\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) = 0 \\
& \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{\cdot j} - \bar{y})(\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) = n_0 \sum_{i=1}^a (\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) \sum_{j=1}^b (\bar{y}_{\cdot j} - \bar{y}) = 0 \\
& \text{SSM} = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{ij} - \bar{y})^2 \\
& = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} [(\bar{y}_{i\cdot} - \bar{y}) + (\bar{y}_{\cdot j} - \bar{y}) + (\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y})]^2 \\
& = \text{SSA} + \text{SSB} + \text{SSAB} \\
& + 2 \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{i\cdot} - \bar{y})(\bar{y}_{\cdot j} - \bar{y}) \\
& + 2 \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{i\cdot} - \bar{y})(\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) \\
& + 2 \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^{n_{ij}} (\bar{y}_{\cdot j} - \bar{y})(\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}) \\
& = \text{SSA} + \text{SSB} + \text{SSAB}
\end{aligned}$$

2. Solution.

(1). By simple calculation and we get

$$\bar{y}_A = 43, \bar{y}_B = 31, \bar{y}_C = 43, \bar{y} = 39$$

$$S_A^2 = 12.8, S_B^2 = 8.4, S_C^2 = 51.2$$

$$\text{SSB} = \sum_{\text{Group } g} n_g (\bar{Y}_g - \bar{Y})^2 = 576$$

$$\text{SSW} = \sum_{\text{Group } g} (n_g - 1) S_g^2 = 362$$

ANOVA Table					
Source	DF.	SS.	MS.	F-Value	p-Value
Between	2	576	288	11.934	7.9×10^{-4}
Within	15	362	24.133		
Total	17	938			

Since $p\text{-value} = 7.9 \times 10^{-4} < 0.05$, we have to reject $H_0: \mu_A = \mu_B = \mu_C$ under the significance level of 0.05.

(2). We can easily compute that

Z_{ij}	1	2	3	4	5	6
A	9	16	25	1	4	9
B	16	16	1	4	1	4
C	16	9	25	49	64	121

$$\bar{Z}_A = 10.67, \bar{Z}_B = 7, \bar{Z}_C = 42.67, \bar{Z} = 20.11$$

$$S_{ZA}^2 = 75.47, S_{ZB}^2 = 50.4, S_{ZC}^2 = 1675.47$$

$$SSB_Z = \sum_{\text{Group } g} n_g (\bar{Z}_g - \bar{Z})^2 = 4619.11$$

$$SSW_z = \sum_{\text{Group } g} (n_g - 1) S_{Zg}^2 = 9006.67$$

ANOVA Table of Z_{ij}					
Source	DF.	SS.	MS.	F-Value	p-Value
Between	2	4619.11	2309.56	3.85	4.47×10^{-2}
Within	15	9006.67	600.44		
Total	17	13625.78			

Since $p\text{-value} = 0.0447 < 0.05$, we have to reject $H_0: \sigma_A^2 = \sigma_B^2 = \sigma_C^2$.

(3).

y	27	29	30	32	32	33	35	38	39	40	40	42	45	46	47	48	49	50
R	1	2	3	4.5	4.5	6	7	8	9	10.5	10.5	12	13	14	15	16	17	18
G	B	B	B	B	C	B	B	A	C	A	C	A	A	A	A	C	C	C

$$\bar{r}_A = 12.083, \bar{r}_B = 3.917, \bar{r}_C = 12.5, \bar{r} = 9.5$$

$$KW = \frac{(n-1) \sum_{i=1}^k n_i (\bar{R}_i - \bar{R})^2}{\sum_{i=1}^k \sum_{j=1}^{n_i} (\bar{R}_i - \bar{R})^2} = 9.883$$

The p-value is given by $\chi^2(9.883, 2) = 7.14 \times 10^{-3} < 0.05$, hence $H_0: m_A = m_B = m_C$ must be rejected.

3. Solution.

(1). The null hypothesis H_0 : there is no difference among the mean number of glassware produced by the four workers is accepted since $p\text{-value} = 0.4841 > 0.1$.

源	自由度	平方和	均方	F 值	Pr > F
模型	3	20.9206349	6.9735450	0.84	0.4841
误差	26	215.7460317	8.2979243		
校正合计	29	236.6666667			

(2). The null hypothesis H_0 : there is no difference among the mean number of glassware produced by the four workers is rejected since $p\text{-value} = 0.0056 < 0.1$.

源	自由度	平方和	均方	F 值	Pr > F
模型	2	75.4303030	37.7151515	6.32	0.0056
误差	27	161.2363636	5.9717172		
校正合计	29	236.6666667			

(3). By the result below, we can obtain

$$\mu_A - \mu_B \in [1.461, 6.273]$$

$$\mu_A - \mu_C \in [-1.270, 3.306]$$

$$\mu_B - \mu_C \in [-5.202, -0.495]$$

That implies differences do exist between tool pairs (A, B) and (B, C).

Comparisons significant at the 0.1 level are indicated by ***.

tool 比较	均值 间 差值	Simultaneous 90% 置信限	
ToolA - ToolC	1.018	-1.270	3.306
ToolA - ToolB	3.867	1.461	6.273 ***
ToolC - ToolA	-1.018	-3.306	1.270
ToolC - ToolB	2.848	0.495	5.202 ***
ToolB - ToolA	-3.867	-6.273	-1.461 ***
ToolB - ToolC	-2.848	-5.202	-0.495 ***

(4). The null hypothesis H_0 : there is no interaction effect between worker and tool is rejected by $p\text{-value} = 0.0874 < 0.1$.

源	自由度	I 型 SS	均方	F 值	Pr > F
worker	3	20.92063492	6.97354497	1.73	0.1964
tool	2	89.32525656	44.66262828	11.09	0.0007
worker*tool	6	53.92077518	8.98679586	2.23	0.0874

(5). The null hypothesis H_0 : the residuals follow a normal distribution is accepted by a Shapiro-Wilk test with $p\text{-value} = 0.6901 > 0.1$.

正态性检验				
检验	统计量	p 值		
Shapiro-Wilk	W	0.97525	Pr < W	0.6901
Kolmogorov-Smirnov	D	0.1	Pr > D	>0.1500
Cramer-von Mises	W-Sq	0.036512	Pr > W-Sq	>0.2500
Anderson-Darling	A-Sq	0.235878	Pr > A-Sq	>0.2500

4. Solution.

(1). The Type I Error Rate is 0.0498.

表: WORK.TASK4_ST1 | 视图: 列名

列	type_I_err_rate
1	0.0498

(2). The Type I Error Rate is $0.0802 > 0.0498$.

表: WORK.TASK4_ST2 | 视图: 列名

列	type_I_err_rate
1	0.0802

From the result, we can see that heteroscedasticity may increase the Type I Error Rate, so homoscedasticity is necessary to be guaranteed if we want to handle an F-test in one-way ANOVA to test the equality of group means.

5. Solution.

(1).

Since the r is known, so the parameter is p . Let

$$h(y) = \frac{(y-r-1)!}{y!(r-1)!}, \eta(p) = \ln p, T(y) = y, A(q) = r \ln \frac{1}{1-p}.$$

Then

$$\begin{aligned} f(y; r, p) &= C_{y+r-1}^y (1-p)^r p^y \\ &= \frac{(y+r-1)!(1-p)^r p^y}{y!(r-1)!} \\ &= \frac{(y+r-1)!}{y!(r-1)!} \exp(y \ln p - r \ln \frac{1}{1-p}) \\ &= h(y) \exp(\eta(p) \cdot T(y) - A(p)) \end{aligned}$$

Hence NB(r, p) belongs to the exponential family when r is known.

(2). Rewrite the distribution of NB(r, p) as a canonical form.

Let

$$q = \ln p, h(y) = \frac{(y-r-1)!}{y!(r-1)!}, \eta_c(q) = q, T(y) = y, A_c(q) = r \ln \frac{1}{1-e^q}$$

then

$$\begin{aligned} f(y; r, p) &= C_{y+r-1}^y (1-p)^r p^y \\ &= \frac{(y+r-1)!(1-p)^r p^y}{y!(r-1)!} \\ &= \frac{(y+r-1)!}{y!(r-1)!} \exp(yq - r \ln \frac{1}{1-e^q}) \\ &= h(y) \exp(\eta_c(q) \cdot T(y) - A_c(q)) \end{aligned}$$

Hence

$$\begin{aligned} E(y) &= A_c'(q) = \frac{re^q}{1-e^q} = \frac{rp}{1-p} \\ \text{Var}(y) &= A_c''(q) = \frac{re^q}{(1-e^q)^2} = \frac{rp}{(1-p)^2} \end{aligned}$$