## Principal Component Analysis with external information on both subjects and variables

Analisi Statistica dei Dati Multidimensionali<sup>1</sup>

<sup>1</sup>Corso di Laurea in Scienze Statistiche e Attuariali

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- Im(X) denotes the subspace spanned by the column vectors of X.
- Let  $Ker(\mathbf{X}^T)$  be the kernel (null space) of  $\mathbf{X}^T$ .
- The data matrix X is often accompanied by external information concerning the rows and columns of data matrix.
- In order to consider several hypotheses and evaluate different effects, we consider several external information H and Z for the units and variables of the matrix X.
- Let  $\mathbf{H}_{\mathbf{X}}$  be the external  $n \times c$  row (units) information matrices on  $\mathbf{X}$ , respectively, with c < n.
- We denote a p × I column information matrix on X by Z<sub>X</sub> with I < p.</li>
- We assume that Z<sub>X</sub> is also columnwise centered.

- if the rows of X represents subjects then we can collect the subject's demographic information in H<sub>X</sub> and will try to explore how they are related to the set of variables of the main data;
- if we set H<sub>X</sub> to 1<sub>n</sub> then we highlight the mean tendency across the subjects;
- we can consider for H<sub>X</sub> matrix of dummy variables providing the subject's membership in some prespecified groups in order to analyze the differences among the groups;
- the external information Z<sub>X</sub> for the variables collected in the matrix X can be unitary vectors with p component in order to capture the relationships among the columns of the matrix:

- if variables represent stimuli in a preference judgment study then Z<sub>X</sub> can be a matrix of the descriptor variables of the stimuli;
- if we have different within-subjects experimental conditions then Z<sub>X</sub> could be a matrix of contrasts; in this sense, orthogonal contrasts used in Analysis of Variance, by which different linear principal mean effects can be highlighted, are a particular case of external information;
- finally, if variables are repeated observations then Z<sub>X</sub> could be a matrix of coefficients of orthogonal polynomials.

• Let  $\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}} = \mathbf{Z}_{\mathbf{X}}(\mathbf{Z}_{\mathbf{X}}^{T}\mathbf{Q}_{\mathbf{X}}\mathbf{Z}_{\mathbf{X}})^{-}\mathbf{Z}_{\mathbf{X}}^{T}\mathbf{Q}_{\mathbf{X}}$  be the not symmetric oblique projection operator onto  $Im(\mathbf{Z}_{\mathbf{X}})$  along  $Ker(\mathbf{Z}_{\mathbf{X}}^{T}\mathbf{Q}_{\mathbf{X}})$  (Takane and Shibayama, 1991).

"External Analysis" of the Principal Component Analysis with external information on both subjects and variables (Takane and Shibayama, 1991; Takane and Hunter, 2001) considers the additive data decomposition of a single quantitative matrix X according to the row and column external information matrices H<sub>X</sub> and Z<sub>X</sub> under row and column metrics L and Q<sub>X</sub>, respectively:

• The four submatrices in (1) are mutually trace-orthogonal in their respective metric matrices  $\bf L$  and  $\bf Q_X$  (Takane and Shibayama, 1991). For example, the first and the second terms in (1) are trace-orthogonal since we have simultaneously

$$\begin{aligned} &\textit{trace}[\underbrace{(\textbf{P}_{\textbf{Z}_{\textbf{X}}/\textbf{Q}_{\textbf{X}}}\textbf{X}^{T}\textbf{P}_{\textbf{H}_{\textbf{X}}/\textbf{L}}^{T})\textbf{L}}_{=\textbf{A}^{T}}\underbrace{(\textbf{P}_{\textbf{H}_{\textbf{X}}/\textbf{L}}^{\perp}\textbf{X}\textbf{P}_{\textbf{Z}_{\textbf{X}}/\textbf{Q}_{\textbf{X}}}^{T})\textbf{Q}_{\textbf{X}}] = 0 \\ &\textit{trace}[\underbrace{\textbf{L}(\textbf{P}_{\textbf{H}_{\textbf{X}}/\textbf{L}}\textbf{X}\textbf{P}_{\textbf{Z}_{\textbf{X}}/\textbf{Q}_{\textbf{X}}}^{T})}_{=\textbf{A}}\underbrace{\textbf{Q}_{\textbf{X}}(\textbf{P}_{\textbf{Z}_{\textbf{X}}/\textbf{Q}_{\textbf{X}}}\textbf{X}^{T}\textbf{P}_{\textbf{H}_{\textbf{X}}/\textbf{L}}^{\perp T})}_{=\textbf{B}^{T}}] = 0 \end{aligned}$$

where two matrices of a same size, **A** and **B**, are said to be trace-orthogonal when  $tr(\mathbf{A}^T\mathbf{B}) = tr(\mathbf{A}\mathbf{B}^T) = 0$ .

This property implies that the sum of squares of **X** can be exactly decomposed according to the sum of sums of squares corresponding to each submatrix in (1):

$$\begin{split} SS(\mathbf{X})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} = & SS(\mathbf{P}_{\mathbf{H}_{\mathbf{X}}/\mathbf{L}}\mathbf{X}\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}}^{T})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} + SS(\mathbf{P}_{\mathbf{H}_{\mathbf{X}}/\mathbf{L}}^{\perp}\mathbf{X}\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}}^{T})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} + \\ & + SS(\mathbf{P}_{\mathbf{H}_{\mathbf{X}}/\mathbf{L}}\mathbf{X}\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}}^{\perp T})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} + SS(\mathbf{P}_{\mathbf{H}_{\mathbf{X}}/\mathbf{L}}^{\perp}\mathbf{X}\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}}^{\perp T})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} \end{split}$$

where  $SS(\mathbf{X})_{\mathbf{L},\mathbf{Q}_{\mathbf{X}}} = tr(\mathbf{X}^T \mathbf{L} \mathbf{X} \mathbf{Q}_{\mathbf{X}}).$ 

Each component in (1) has a specific statistical meaning:

- $\mathbf{P}_{\mathbf{H}_{\mathbf{X}}/\mathbf{L}}\mathbf{X}\mathbf{P}_{\mathbf{Z}_{\mathbf{X}}/\mathbf{Q}_{\mathbf{X}}}^{T}$  represents the row and column constraints effects;
- ullet  $\left| \mathbf{P}_{\mathbf{H_X}/\mathbf{L}}^{\perp} \mathbf{X} \mathbf{P}_{\mathbf{Z_X}/\mathbf{Q_X}}^{T} \right|$  reflects the column constraints effect;
- ullet  $\mathbf{P}_{\mathbf{H_X}/\mathbf{L}}\mathbf{X}\mathbf{P}_{\mathbf{Z_X}/\mathbf{Q_X}}^{\perp T}$  reflects the row constraints effect
- the last term  $\left| \mathbf{P}_{\mathbf{H}_{X}/\mathbf{L}}^{\perp} \mathbf{X} \mathbf{P}_{\mathbf{Z}_{X}/\mathbf{Q}_{X}}^{\perp T} \right|$  pertains to what can be explained by neither row nor column external information.

## Row sides metric matrix **L** can assume several forms:

- if there are no differences in importance among the statistical units then L can be set to the identity matrix;
- if there are differences in importance among the statistical units then special diagonal matrices are used as L in order to differentially weight the rows of X;
- finally, when rows of the data matrix are time points in single-subject multivariate time series data, Escoufier (1987) suggests to use the inverse of the matrix of serial correlations as L.

- By using this approach, external information H<sub>X</sub> and Z<sub>X</sub> are incorporated within a single data set.
- It is evident that the Takane and Shibayama's basic decomposition (1991) is obtained from (1) by setting  $\mathbf{L} = \mathbf{I}_n$  and  $\mathbf{Q}_{\mathbf{X}} = \mathbf{I}_p$ :

$$\label{eq:control_equation} \left| \mathbf{X} = \mathbf{P}_{\mathbf{H}_{\mathbf{X}}} \mathbf{X} \mathbf{P}_{\mathbf{Z}_{\mathbf{X}}} + \mathbf{P}_{\mathbf{H}_{\mathbf{X}}}^{\perp} \mathbf{X} \mathbf{P}_{\mathbf{Z}_{\mathbf{X}}}^{\perp} + \mathbf{P}_{\mathbf{H}_{\mathbf{X}}}^{\perp} \mathbf{X} \mathbf{P}_{\mathbf{Z}_{\mathbf{X}}}^{\perp} \right.$$

 An "Internal analysis" (PCA) is then performed on each component of decomposition. Consider the following model

$$\mathbf{X} = \mathbf{H}_{\mathbf{X}} \mathbf{M} \mathbf{Z}_{\mathbf{X}}^{T} + \mathbf{B} \mathbf{Z}_{\mathbf{X}}^{T} + \mathbf{H}_{\mathbf{X}} \mathbf{C} + \mathbf{E}$$
 (2)

where **M**  $(c \times l)$ , **B**  $(n \times l)$ , and **C**  $(c \times p)$  are matrices of coefficients to be estimated, and **E**  $(n \times p)$  a matrix of error components.

- The four terms in (2) explain portions of the original data matrix, X.
- The first term pertains to what can be explained by both H<sub>X</sub> and Z<sub>X</sub>, the second term by Z<sub>X</sub>, the third term by H<sub>X</sub>, and the fourth term by neither H<sub>X</sub> nor Z<sub>X</sub>.

- Let X = H<sub>X</sub>MZ<sub>X</sub><sup>T</sup> + E<sub>1</sub>, and consider the problem of estimating M so as to minimize SS(E<sub>1</sub>) = tr(E<sub>1</sub><sup>T</sup>E<sub>1</sub>).
- We obtain

$$\hat{\mathbf{M}} = (\mathbf{H}_{\mathbf{X}}^{T}\mathbf{H}_{\mathbf{X}})^{-}\mathbf{H}_{\mathbf{X}}^{T}\mathbf{X}\mathbf{Z}_{\mathbf{X}}(\mathbf{Z}_{\mathbf{X}}^{T}\mathbf{Z}_{\mathbf{X}})^{-}$$

where  $(\mathbf{H}_{\mathbf{X}}^{\mathsf{T}}\mathbf{H}_{\mathbf{X}})^{-}$  and  $(\mathbf{Z}_{\mathbf{X}}^{\mathsf{T}}\mathbf{Z}_{\mathbf{X}})^{-}$  are g-inverses of  $(\mathbf{H}_{\mathbf{X}}^{\mathsf{T}}\mathbf{H}_{\mathbf{X}})$  and  $(\mathbf{Z}_{\mathbf{X}}^{\mathsf{T}}\mathbf{Z}_{\mathbf{X}})$ , respectively.

The residual from the first term is now equal to

$$\hat{\boldsymbol{E}}_1 = \boldsymbol{X} - \boldsymbol{H}_{\boldsymbol{X}} \hat{\boldsymbol{M}} \boldsymbol{Z}_{\boldsymbol{X}}^{T} = \boldsymbol{X} - \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}} \boldsymbol{X} \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}$$

where  $P_{H_X}$  and  $P_{Z_X}$  are orthogonal projection operators onto spaces spanned by the column vectors of  $H_X$  and  $Z_X$ , respectively.

 $\bullet$  We now separately fit the second and the third terms to  $\hat{\textbf{E}}_1\colon$ 

$$\hat{\mathbf{E}}_1 = \mathbf{B} \mathbf{Z}_{\mathbf{X}}^T + \mathbf{E}_2$$
 $\hat{\mathbf{E}}_1 = \mathbf{H}_{\mathbf{X}} \mathbf{C} + \mathbf{E}_3$ 

 We obtain a least squares estimate of B that minimizes SS(E<sub>2</sub>) by

$$\hat{\mathbf{B}} = \mathbf{P}_{H_{\boldsymbol{X}}}^{\perp} \mathbf{X} \mathbf{Z}_{\boldsymbol{X}} (\mathbf{Z}_{\boldsymbol{X}}^{T} \mathbf{Z}_{\boldsymbol{X}})^{-}$$

where  $\mathbf{P}_{H_X}^{\perp}=(\mathbf{I}-\mathbf{P}_{H_X})$  orthogonal projection operator to  $\mathbf{P}_{H_X}$  (that is  $\mathbf{P}_{H_X}^{\perp}\mathbf{P}_{H_X}=\mathbf{0}$  and  $\mathbf{P}_{H_X}+\mathbf{P}_{H_X}^{\perp}=\mathbf{I}$ ).

Similarly, we obtain

$$\hat{\boldsymbol{C}} = (\boldsymbol{H}_{\boldsymbol{X}}^T\boldsymbol{H}_{\boldsymbol{X}})^-\boldsymbol{H}_{\boldsymbol{X}}^T\boldsymbol{X}\boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}^\perp$$

that minimizes  $SS(\mathbf{E}_3)$ 

Now, the estimate of the fourth term is given by

$$\begin{split} \hat{\textbf{E}} = & \textbf{H}_{\textbf{X}} \hat{\textbf{M}} \textbf{Z}_{\textbf{X}}^{\mathcal{T}} + \hat{\textbf{B}} \textbf{Z}_{\textbf{X}}^{\mathcal{T}} + \textbf{H}_{\textbf{X}} \hat{\textbf{C}} \\ = & \textbf{X} - \textbf{P}_{\textbf{H}_{\textbf{X}}} \textbf{X} \textbf{P}_{\textbf{Z}_{\textbf{X}}} - \textbf{P}_{\textbf{H}_{\textbf{X}}}^{\perp} \textbf{X} \textbf{P}_{\textbf{Z}_{\textbf{X}}} - \textbf{P}_{\textbf{H}_{\textbf{X}}} \textbf{X} \textbf{P}_{\textbf{Z}_{\textbf{X}}}^{\perp} \\ = & \textbf{P}_{\textbf{H}_{\textbf{X}}}^{\perp} \textbf{X} \textbf{P}_{\textbf{Z}_{\textbf{X}}}^{\perp} \end{split}$$

 By substituting the least squares estimates for the corresponding parameters, we obtain the following decomposition of the data matrix, X:

$$\begin{split} \boldsymbol{X} = & (\boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}} + \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}}^{\perp}) \boldsymbol{X} (\boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}} + \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}^{\perp}) \\ = & \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}} \boldsymbol{X} \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}} + \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}}^{\perp} \boldsymbol{X} \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}^{\perp} + \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}}^{\perp} \boldsymbol{X} \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}^{\perp} + \boldsymbol{P}_{\boldsymbol{H}_{\boldsymbol{X}}}^{\perp} \boldsymbol{X} \boldsymbol{P}_{\boldsymbol{Z}_{\boldsymbol{X}}}^{\perp} \end{split}$$