Regression

Victor Kitov v.v.kitov@yandex.ru

Yandex School of Data Analysis



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- Madaraya-Watson regression

Linear regression

- Linear model $f(x, \beta) = x^T \beta = \sum_{i=1}^D \beta_i x^i$
- Define $X \in \mathbb{R}^{N \times D}$, $\{X\}_{ij}$ defines the j-th feature of i-th object, $Y \in \mathbb{R}^n$, $\{Y\}_i$ target value for i-th object.
- Ordinary least squares (OLS) method:

$$\sum_{n=1}^{N} \left(x^{T} \beta - y_{n} \right)^{2} = \sum_{n=1}^{N} \left(\sum_{d=1}^{D} \beta_{d} x_{n}^{d} - y_{n} \right)^{2} \rightarrow \min_{\beta}$$

Solution

Stationarity condition:

$$2\sum_{n=1}^{N} x_n \left(\sum_{d=1}^{D} \beta_d x_n^d - y_n\right) = \mathbf{0}$$

In matrix form:

$$2X^T(X\beta - Y) = \mathbf{0}$$

so

$$\widehat{\beta} = (X^T X)^{-1} X^T Y$$

This is the global minimum, because the optimized criteria is convex.

Linearly dependent features

- Solution $\widehat{\beta} = (X^TX)^{-1}X^TY$ exists when X^TX is non-degenerate
- Problem occurs when one of the features is a linear combination of the other
 - because of the property $\forall X$: $rank(X) = rank(X^TX)$
 - example: constant unity feature c and one-hot-encoding $e_1, e_2, ... e_K$, because $\sum_k e_k \equiv c$
 - interpretation: non-identifiability of $\widehat{\beta}$ for linearly dependent features:
 - linear dependence: $\exists \alpha : x^T \alpha = 0 \ \forall x$
 - suppose β solves linear regression $y = x^T \beta$
 - then $x^T \beta \equiv x^T \beta + k x^T \alpha \equiv x^T (\beta + k \alpha)$, so $\beta + k \alpha$ is also a solution!

Linearly dependent features

- Problem may be solved by:
 - feature selection
 - dimensionality reduction
 - imposing additional requirements on the solution (regularization)

Analysis of linear regression

Advantages:

- single optimum, which is global (for non-singular matrix)
- analytical solution
- interpretable solution and algorithm

Drawbacks:

- too simple model assumptions (may not be satisfied)
- X^TX should be non-degenerate (and well-conditioned)

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Generalization by nonlinear transformations

Nonlinearity by x in linear regression may be achieved by applying non-linear transformations to the features:

$$x \to [\phi_0(x), \phi_1(x), \phi_2(x), \dots \phi_M(x)]$$

$$f(x) = \phi(x)^T \beta = \sum_{m=0}^{M} \beta_m \phi_m(x)$$

The model remains to be linear in β , so all advantages of linear regression remain.

Typical transformations

$\phi_k(x)$	comments
$x^i \in [a, b]$	binarization of feature
$x^i x^j$	interaction of features
$= \left\{ -\gamma \left\ x - \tilde{x} \right\ ^2 \right\}$	closeness to reference point $ ilde{x}$
$\ln x_k$	the alignment of the distribution with heavy tails
$F^{-1}(x_k)$	conversion of atypical continious distribution to uniform ¹

¹why?

Non-linear regression

• Alternatively we can model $\mathcal{X} o \mathcal{Y}$ with arbitrary non-linear function $\widehat{y} = f(x|\theta)$

$$L(\theta|X,Y) = \sum_{n=1}^{N} (f(x_n|\theta) - y_n)^2$$

$$\widehat{\theta} = \arg\min_{\theta} L(\theta|X,Y)$$

- ullet No analytical solution for $\widehat{ heta}$ will exist in general
 - need numeric optimization methods.

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Regularization

ullet Insert additional requirement for regularizer R(eta) to be small:

$$\sum_{n=1}^{N} \left(x_n^T \beta - y_n \right)^2 + \lambda R(\beta) \to \min_{\beta}$$

- $\lambda > 0$ hyperparameter.
- $R(\beta)$ penalizes complexity of models.

$$R(\beta) = ||\beta||_1$$
 Lasso regression $R(\beta) = ||\beta||_2^2$ Ridge regression

- Not only accuracy matters for the solution but also model simplicity!
- λ controls complexity of the model:

Regularization

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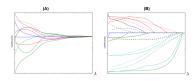
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- Not only accuracy matters for the solution but also model simplicity!
- λ controls complexity of the model: $\uparrow \lambda \Leftrightarrow \text{complexity} \downarrow$.

Comments

• Dependency of β from λ for ridge (A) and LASSO (B):



- LASSO can be used for automatic feature selection.
- λ is usually found using cross-validation on exponential grid, e.g. $[10^{-6}, 10^{-5}, ... 10^{5}, 10^{6}]$.
- It's always recommended to use regularization because
 - it gives smooth control over model complexity.
 - reduces ambiguity for multiple solutions case.

Ridge regression solution

Ridge regression criterion

$$\sum_{n=1}^{N} \left(x_n^T \beta - y_n \right)^2 + \lambda \beta^T \beta \to \min_{\beta}$$

Stationarity condition can be written as:

$$2\sum_{n=1}^{N} x_n \left(x_n^T \beta - y_n \right) + 2\lambda \beta = 0$$
$$2X^T (X\beta - Y) + \lambda \beta = 0$$
$$\left(X^T X + \lambda I \right) \beta = X^T Y$$

so

$$\widehat{\beta} = (X^T X + \lambda I)^{-1} X^T Y$$

Comments

- $X^TX + \lambda I$ is always non-degenerate as a sum of:
 - non-negative definite X^TX
 - ullet positive definite λI
- Intuition:
 - out of all valid solutions select one giving simplest model
- Other regularizations also restrict the set of solutions.

Different account for different features

• Traditional approach regularizes all features uniformly:

$$\sum_{n=1}^{N} \left(x_n^T \beta - y_n \right)^2 + \lambda R(\beta) \to \min_{w}$$

Suppose we have K groups of features with indices:

$$I_1, I_2, ... I_K$$

• We may control the impact of each group on the model by:

$$\sum_{n=1}^{N} \left(x_n^T \beta - y_n \right)^2 + \lambda_1 R(\{\beta_i | i \in I_1\}) + \ldots + \lambda_K R(\{\beta_i | i \in I_K\}) \to \min_{w}$$

- $\lambda_1, \lambda_2, ... \lambda_K$ can be set using cross-validation
- In practice use common regularizer but with different feature scaling.

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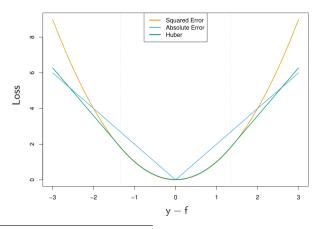
Idea

• Generalize quadratic to arbitrary loss:

$$\sum_{n=1}^{N} \left(x^{T} \beta - y_{n} \right)^{2} \to \min_{\beta} \qquad \Longrightarrow \qquad \sum_{n=1}^{N} \mathcal{L}(x_{n}^{T} \beta - y_{n}) \to \min_{\beta}$$

• Robust means solution is robust to outliers in the training set.

Non-quadratic loss functions23



²What is the value of constant prediction, minimizing sum of squared errors?

³What is the value of constant prediction, minimizing sum of absolute errors?

• For $y_1, ... y_N \in \mathbb{R}$ constant minimizers $\widehat{\mu}$:

$$rg \min_{\mu} \sum_{n=1}^{N} (y_n - \mu)^2 = rg \min_{\mu} \sum_{n=1}^{N} |y_n - \mu| = rg \min_{\mu} |y_n -$$

• For $y_1,...y_N \in \mathbb{R}$ constant minimizers $\widehat{\mu}$:

$$\arg\min_{\mu} \sum_{n=1}^{N} (y_n - \mu)^2 = \frac{1}{N} \sum_{n=1}^{N} y_n$$

$$\arg\min_{\mu} \sum_{n=1}^{N} |y_n - \mu| = \text{median}\{y_1, ... y_N\}$$

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• For $x, y \sim P(x, y)$ and functional minimizers f(x):

$$\arg \min_{f(x)} \mathbb{E} \left\{ (f(x) - y)^2 \middle| x \right\} =$$

$$\arg \min_{f(x)} \mathbb{E} \left\{ |f(x) - y| \middle| x \right\} =$$

• For $y_1,...y_N \in \mathbb{R}$ constant minimizers $\widehat{\mu}$:

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• For $x, y \sim P(x, y)$ and functional minimizers f(x):

$$\arg\min_{f(x)} \mathbb{E}\left\{ \left. (f(x) - y)^2 \right| x \right\} = \mathbb{E}[y|x]$$

$$\arg\min_{f(x)} \mathbb{E}\left\{ \left| f(x) - y \right| |x \right\} = \mathrm{median}[y|x]$$

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Weighted account for observations

Weighted account for observations

$$\sum_{n=1}^{N} w_n (x_n^T \beta - y_n)^2$$

- Weights may be:
 - increased for incorrectly predicted objects
 - algorithm becomes more oriented on error correction
 - decreased for incorrectly predicted objects
 - they may be considered outliers that break our model

⁴Derive solution for weighted regression.

Robust regression

- Initialize $w_1 = ... = w_N = 1/N$
- Repeat:
 - estimate regression $\hat{y}(x)$ using observations (x_i, y_i) with weights w_i .
 - for each i = 1, 2, ...N:
 - re-estimate $\varepsilon_i = \widehat{y}(x_i) y_i$
 - recalculate $w_i = K(|\varepsilon_i|)$
 - normalize weights $w_i = \frac{w_i}{\sum_{n=1}^N w_n}$

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Comments: $K(\cdot)$ is some decreasing function, repetition may be

- predefined number of times
- until convergence of model parameters.

Example

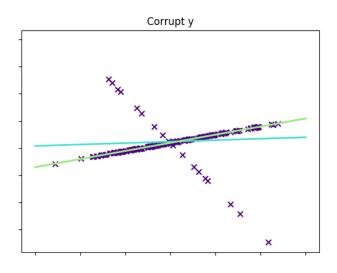


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Minimum squared error estimate

For training sample $(x_1, y_1), ... (x_N, y_N)$ consider finding constant $\hat{y} \in \mathbb{R}$

$$L(\alpha) = \sum_{i=1}^{N} (\widehat{y} - y_i)^2 \to \min_{\widehat{y} \in \mathbb{R}}$$

$$\frac{\partial L}{\partial \alpha} = 2 \sum_{i=1}^{N} (\widehat{y} - y_i) = 0, \text{ so } \widehat{y} = \frac{1}{N} \sum_{i=1}^{N} y_i$$

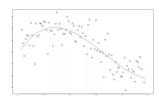
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We need to model general curve y(x):



Minimum squared error estimate

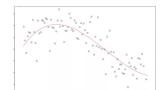
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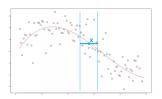
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We need to model general curve y(x):

Nadaraya-Watson regression - localized averaging approach.





Nadaraya-Watson regression

- Equivalent names: local constant regression, kernel regression
- For each x assume $\widehat{y}(x) = \alpha = \alpha(x), \ \alpha \in \mathbb{R}$.

$$Q(\alpha, X_{training}) = \sum_{i=1}^{N} w_i(x)(\alpha - y_i)^2 \to \min_{\alpha \in \mathbb{R}}$$

 Weights depend on the proximity of training objects to the predicted object:

$$w_i(x) = K\left(\frac{\rho(x, x_i)}{h}\right)$$

- K(u) some decreasing function, called kernel.
- h(x) some ≥ 0 function called bandwidth.
 - Intuition: "window width", consider h(x) = h, $K(u) = \mathbb{I}[u \le 1]$.

Parameters

• Typically used $K(u)^5$:

$$K_G(u) = e^{-\frac{1}{2}u^2} - \text{Gaussian kernel}$$

 $K_P(u) = (1-u^2)^2 \mathbb{I}[|u| < 1] - \text{quartic kernel}$

- Typically used h(x):
 - h(x) = const
 - $h(x) = \rho(x, x_{i_K})$, where x_{i_K} K-th nearest neighbour.
 - better for unequal distribution of objects

⁵Compare them in terms of required computation.

Solution

$$Q(\alpha, X_{training}) = \sum_{i=1}^{N} w_i(x)(\alpha - y_i)^2 \to \min_{\alpha \in \mathbb{R}}$$
$$w_i(x) = K\left(\frac{\rho(x, x_i)}{h(x)}\right)$$

• From stationarity condition $\frac{\partial Q}{\partial \alpha}=0$ obtain optimal $\widehat{\alpha}(x)$:

$$f(x,\alpha) = \widehat{\alpha}(x) = \frac{\sum_{i=1}^{N} y_i w_i(x)}{\sum_{i=1}^{N} w_i(x)} = \frac{\sum_{i=1}^{N} y_i K\left(\frac{\rho(x,x_i)}{h(x)}\right)}{\sum_{i=1}^{N} K\left(\frac{\rho(x,x_i)}{h(x)}\right)}$$

Comments

- Under general regularity conditions $g(x, \alpha) \stackrel{P}{\to} E[y|x]$
- The specific form of the kernel function does not affect the accuracy much.
 - but may affect efficiency
- Compared to K-NN: may use all objects, bandwidth controls smoothness.
 - under what selection of K(u) and h(x) it reduces to basic K-NN?

Comments

Insead of optimizing locally with constant lpha

$$Q(\alpha, X_{training}) = \sum_{i=1}^{N} w_i(x)(\alpha - y_i)^2 \to \min_{\alpha \in \mathbb{R}}$$

we could have optimized local linear regression

$$Q(\alpha, X_{training}) = \sum_{i=1}^{N} w_i(x) (x^{\mathsf{T}} \beta - y_i)^2 \to \min_{\alpha \in \mathbb{R}}$$

This better handles approximation on the edges of domain and local extrema.

Summary

- Linear regression gives interpretable analytic solution.
- Non-linear dependencies can be modelled by adding non-linear features.
- When features are correlated, it fails.
- Regularized versions are always preferrable:
 - work in case of correlated features
 - are more robust in close to correlation case
 - ullet λ gives a convenient way to control model complexity
- Robust regression is robust to outliers.
 - we may also use robust loss-functions instead of MSE.