# CS 475 Machine Learning: Homework 1 Analytical (35 points)

Assigned: Monday, September 13, 2021 Due: Wednesday, September 22, 2021, 11:59 pm US/Eastern

Partner 1: NAME (JHED), Partner 2: NAME (JHED)

#### Instructions

We have provided this LATEX document for turning in this homework. We give you one or more boxes to answer each question. The question to answer for each box will be noted in the title of the box. You can change the size of the box if you need more space.

Other than your name, do not type anything outside the boxes. Leave the rest of the document unchanged.

Do not change any formatting in this document, or we may be unable to grade your work. This includes, but is not limited to, the height of textboxes, font sizes, and the spacing of text and tables. Additionally, do not add text outside of the answer boxes. Entering your answers are the only changes allowed.

We strongly recommend you review your answers in the generated PDF to ensure they appear correct. We will grade what appears in the answer boxes in the submitted PDF, NOT the original latex file.

# 1 Probability and Linear Algebra: Diagnostic

This section is ungraded and intended for diagnostic purposes only. While answers to these questions are easy to compute with access to a statistical language interpreter, or look up on the internet, we advise you not to do so. These questions are an opportunity to verify that you feel comfortable with the prerequisite topics for this class. If you don't know/remember everything, that doesn't mean you can't still do well, but you would need to put in extra effort reviewing the relevant background.

#### **Probability**

- 1. Recall that variance is defined as  $Var(X) = E[(X E[X])^2]$ . Prove that  $Var(X) = E[X^2] E[X]^2$ .
- 2. Let X be a random variable such that X = YZ, where  $Y \sim \mathcal{N}(0, \sigma^2)$  and  $Z \sim \text{Bernoulli}(p)$ . Find the mean and variance of X.

#### Linear Algebra

- 1. Show that the vector w is orthogonal to the hyperplane  $w^Tx + b = 0$ .
- 2. Consider the matrix A below:

$$A = \begin{bmatrix} 1 & 2 & 5 \\ 2 & 4 & 3 \\ 4 & 5 & 8 \end{bmatrix}$$

- (a) What is the rank of A?
- (b) Compute the determinant of A.

#### 2 Likelihood

Given n data points  $\{x_1, x_2, \dots, x_n\}$  and the following linear model

$$y_i = \omega^T x_i + \epsilon_i$$

where  $\epsilon_i$  is a random variable representing the noise and is independent of **x**.

(a) Assume  $\epsilon_i$  comes from a standard Gaussian distribution, i.e.

$$P(\epsilon_i) = \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-\epsilon_i^2}{2}\right)$$

Compute the conditional log-likelihood of  $\mathbf{y}$  given  $\mathbf{x}$  and  $\omega$ . Give the simplest function of  $y_i$  and  $\omega$  such that minimizing this function is equivalent to maximizing the conditional log-likelihood.

$$L_{[D]}(\omega) = \prod_{i=1}^{n} p(y_i \mid x_i; \omega)$$

$$= \prod_{i=1}^{n} p(\omega^T x_i + \epsilon_i \mid x_i; \omega)$$

$$= \prod_{i=1}^{n} p(\epsilon_i \mid x_i; \omega)$$

$$= \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y_i - \omega^T x_i)^2}{2}\right)$$

$$logL_{[D]}(\omega) = \sum_{i=1}^{n} log \frac{1}{\sqrt{2\pi}} + \sum_{i=1}^{n} -\frac{(y_i - \omega^T x_i)^2}{2}$$

Maximizing  $log L_{[D]}(\omega)$  is equivalent to minimizing:

$$\sum_{i=1}^{n} (y_i - \omega^T x_i)^2$$

(b) Assume  $\epsilon_i$  comes from a Laplace distribution, i.e.

$$P(\epsilon_i) = \frac{1}{2} \exp\left(-|\epsilon_i|\right)$$

Compute the conditional log-likelihood of  $\mathbf{y}$  given  $\mathbf{x}$  and  $\omega$ . Give the simplest function of  $y_i$  and  $\omega$  such that minimizing this function is equivalent to maximizing the conditional log-likelihood

$$L_{[D]}(\omega) = \prod_{i=1}^{n} p(\omega^{T} x_{i} + \epsilon_{i} \mid x_{i}; \omega)$$

$$= \prod_{i=1}^{n} p(\epsilon_{i} \mid x_{i}; \omega)$$

$$= \prod_{i=1}^{n} \frac{1}{2} \exp\left(-|\epsilon_{i}|\right)$$

$$= \prod_{i=1}^{n} \frac{1}{2} \exp\left(-|y_{i} - \omega^{T} x_{i}|\right)$$

$$log L_{[D]}(\omega) = \sum_{i=1}^{n} log \frac{1}{2} + \sum_{i=1}^{n} (-|y_i - \omega^T x_i|)$$

Maximizing  $log L_{[D]}(\omega)$  is equivalent to minimizing:

$$\sum_{i=1}^{n} (\left| y_i - \omega^T x_i \right|)$$

(c) Which loss is easier to minimize? Which loss is more robust to outliers? Explain in detail.

### 3 Conditional Independence

A large group of people were surveyed on their recent health. Of these, 0.20 had a fever and 0.05 had pneumonia. Among the people who had pneumonia, 0.70 had cough as a symptom and 0.50 had fever as a symptom. Among the people who had a fever, 0.40 had cough as a symptom.

Let us create a probabilistic model where the presence/absence of each of these two symptoms, cough and fever, are conditionally independent given the presence/absence of pneumonia. Using this data for the empirical probabilities of our model, answer the following questions.

1. Find the probability that someone has both a cough and a fever.

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P(cough \cap fever) = P(cough|fever) \cdot P(fever)
= 0.40 \cdot 0.20
= 0.08
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2. Find the probability that someone has pneumonia given that they have a fever but no cough.

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\begin{split} &P(pneumonia|fever \cap no\ cough)\\ &= \frac{P(pneumonia \cap fever \cap no\ cough)}{P(fever \cap no\ cough)}\\ &= \frac{P(fever \cap no\ cough|pneumonia) \cdot P(pneumonia)}{P(no\ cough|fever) \cdot P(fever)}\\ &= \frac{P(no\ cough|fever, pneumonia) \cdot P(fever|pneumonia) \cdot P(pneumonia)}{[1 - P(cough|fever)] \cdot P(fever)}\\ &= \frac{P(no\ cough|pneumonia) \cdot P(fever|pneumonia) \cdot P(pneumonia)}{[1 - P(cough|fever)] \cdot P(fever)}\\ &= \frac{[1 - P(cough|pneumonia)] \cdot P(fever|pneumonia) \cdot P(pneumonia)}{[1 - P(cough|fever)] \cdot P(fever)}\\ &= \frac{(1 - 0.70) \cdot 0.50 \cdot 0.05}{(1 - 0.40) \cdot 0.20}\\ &= 0.0625 \end{split}
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# 4 Conjugate Priors

1. De	efine what a conjugate prior is.
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2. W	hy are conjugate priors useful?
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Th	now that the Gamma distribution is a conjugate prior of the exponential distribution. nat is, show that if $x \sim \text{Exp}(\lambda)$ and $\lambda \sim \text{Gamma}(\alpha, \beta)$ , then $p(\lambda x) \sim \text{Gamma}(\alpha^*, \beta^*)$ is some $\alpha^*$ , $\beta^*$ .

# 5 Gibbs Sampling and the Semi-Graphoid Axioms

- Assume a joint distribution p(x<sub>1</sub>,...,x<sub>k</sub>) over binary random variables X<sub>1</sub>,...,X<sub>k</sub>. What's the size of the joint probability table?
   Assume (X<sub>1</sub>⊥⊥X<sub>3</sub>,...,X<sub>k</sub> | X<sub>2</sub>), (X<sub>k</sub>⊥⊥X<sub>1</sub>,...,X<sub>k-2</sub> | X<sub>k-1</sub>), and (X<sub>i</sub>⊥⊥X<sub>1</sub>,...,X<sub>i-2</sub>,X<sub>i+2</sub>,...,X<sub>k</sub> | X<sub>i-1</sub>, X<sub>i+1</sub>) for each i = 2,...,k − 1. What's the smallest number of parameters we would need to specify to create a Gibbs sampler for p(x<sub>1</sub>,...,x<sub>k</sub>)?
- 3. Assume conditional independences as in the previous question. Use the chain rule of probability and the graphoid axioms to write down the likelihood for the model such that only a polynomial number of parameters (in k) are used.

