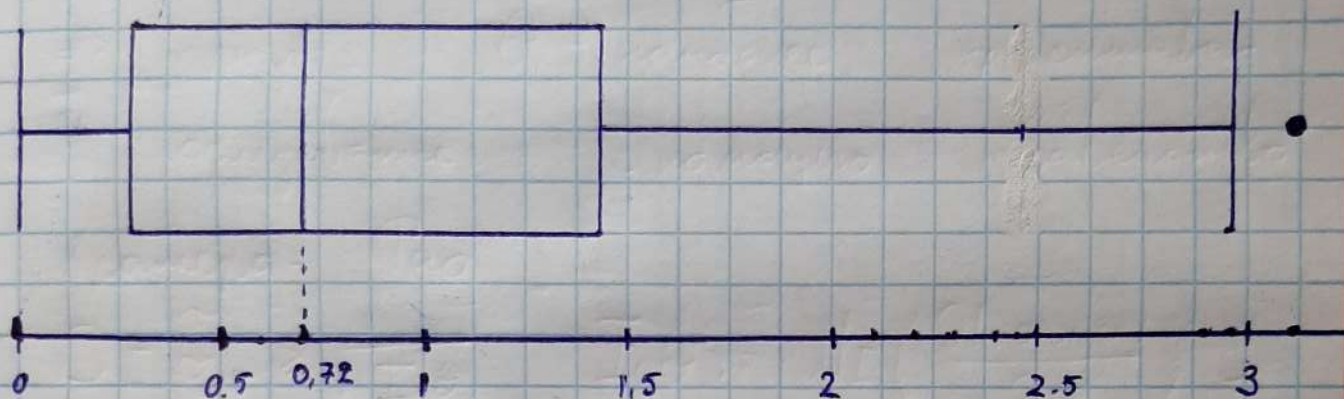


III-1.

Boxplot:



Совместное и маргинальное распределение  $i$  и  $j$  элементов:

$$p(x_i, x_j) = p(x, y) = p(x)p(y) =$$

$$= \left[ 25 C_{24}^{i-1} e^{-x} (1 - e^{-x})^{i-1} (e^{-x})^{25-i} \cdot \right.$$

$$\left. \cdot 25 C_{24}^{j-1} e^{-x} (1 - e^{-x})^{j-1} (e^{-x})^{25-j} \right]$$



III-2

$$f \sim R(0; \theta) \Rightarrow p(x) = \frac{1}{\theta} \cdot \theta$$

$$M[f] = \int_{-\infty}^{+\infty} \frac{1}{\theta} x dx = \int_0^{\theta} \frac{1}{\theta} x dx = \frac{x^2}{2\theta} \Big|_0^{\theta} = \frac{\theta}{2}$$

$$M[f^2] = \int_{-\infty}^{+\infty} \frac{1}{\theta} x^2 dx = \int_0^{\theta} \frac{1}{\theta} x^2 dx = \frac{x^3}{3\theta} \Big|_0^{\theta} = \frac{\theta^2}{3}$$

$$D[f] = M[f^2] - M^2[f] = \frac{\theta^2}{3} - \frac{\theta^2}{4} = \frac{\theta^2}{12}$$

Первая оценка:  $\tilde{\theta}_1 = \frac{2}{n} \sum_{i=1}^n x_i$

а) Несмещенность:

$$M[\tilde{\theta}_1] = M\left[\frac{2}{n} \sum_{i=1}^n x_i\right] = \frac{2}{n} \sum_{i=1}^n M[f] = \frac{2}{n} n M[f] =$$

$$= 2 M[f] = 2 \cdot \frac{\theta}{2} = \theta \Rightarrow M[\tilde{\theta}_1] = \theta \Rightarrow$$

$\Rightarrow \tilde{\theta}_1$  - несмещенная

б) Составленность:

Проверка гомогенного уандина

$$D[\tilde{\theta}_1] = D\left[\frac{2}{n} \sum_{i=1}^n x_i\right] = \frac{4}{n^2} \sum_{i=1}^n D[f] =$$

$$= \frac{4}{n^2} n D[f] = \frac{4}{n} D[f] = \frac{4}{n} \frac{\theta^2}{12} = \frac{\theta^2}{3n} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow$$

$\Rightarrow \tilde{\theta}_1$  - составленная



Вторая оценка:  $\tilde{\theta}_2 = x_{\min}$

$$\varphi(x) = 1 - (1 - F(x))^n = 1 - \left(1 - \int_0^x \frac{1}{\theta} dt\right)^n - \text{функция}$$

распределения

min.

$$\varphi(x) = \varphi'(x) = -n \left(1 - \int_0^x \frac{1}{\theta} dt\right)^{n-1} \left(-\frac{1}{\theta}\right) =$$

$$= \frac{n}{\theta} \left(1 - \frac{x}{\theta}\right)^{n-1} (0; \theta) - \text{плотность распределения}$$

а) Несмещенность:

$$M[\tilde{\theta}_2] = M[x_{\min}] = \int_{-\infty}^{+\infty} n \frac{x}{\theta} \left(1 - \frac{x}{\theta}\right)^{n-1} dx = \{x = \theta t\}$$

$$= \int_0^1 n t (1-t)^{n-1} \theta dt = \int_0^1 n \theta t (1-t)^{n-1} dt =$$

$$= B(2, n) \cdot \theta_n = \theta_n \frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)} = \frac{\theta_n \Gamma(n)}{(n+1)(n) \Gamma(n)} =$$

$$= \frac{\theta}{n+1} \Rightarrow M[\tilde{\theta}_2] = \frac{\theta}{n+1} \Rightarrow \tilde{\theta}_2 - \text{смещена}$$

Коррекция:  $\tilde{\theta}_2' = (n+1) \tilde{\theta}_2 = (n+1) x_{\min}$

б) Составляющие:

Проверка достаточного условия:

$$D[\tilde{\theta}_2] = M[\tilde{\theta}_2^2] - M^2[\tilde{\theta}_2]$$

$$M[\tilde{\theta}_2^2] = \int_{-\infty}^{+\infty} n \frac{x^2}{\theta} \left(1 - \frac{x}{\theta}\right)^{n-1} dx = \{t = \frac{x}{\theta}\} =$$

$$= \int_0^1 n \theta^2 t^2 (1-t)^{n-1} dt = B(3, n) n \theta^2 =$$



$$\theta_n^2 \frac{\Gamma(3)\Gamma(n)}{\Gamma(n+3)} = \frac{2\theta_n^2 \Gamma(n)}{(n+2)(n+1)n\Gamma(n)} = \frac{2\theta^2}{(n+2)(n+1)}$$

$$\begin{aligned} D[\tilde{\theta}_2] &= \frac{2\theta^2}{(n+2)(n+1)} - \frac{\theta^2}{(n+1)^2} = \\ &= \frac{2(n+1)\theta^2 - (n+2)\theta^2}{(n+2)(n+1)^2} = \frac{2n\theta^2 + 2\theta^2 - n\theta^2 - 2\theta^2}{(n+2)(n+1)^2} = \\ &= \frac{n\theta^2}{(n+2)(n+1)^2} \end{aligned}$$

Дисперсия несмещенной оценки

$$\begin{aligned} D[\tilde{\theta}_2'] &= D[(n+1)\tilde{\theta}_2] = (n+1)^2 D[\tilde{\theta}_2] = \\ &= (n+1)^2 \frac{n\theta^2}{(n+2)(n+1)^2} = \frac{n\theta^2}{n+2} \xrightarrow{n \rightarrow \infty} 0 \end{aligned}$$

Оценочное выражение не биномиально.

По определению:

$$\tilde{\theta}_2' \xrightarrow{P} \theta$$

$$P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \geq P(\tilde{\theta}_2' \geq \theta + \varepsilon) =$$

$$= P((n+1)x_{\min} \geq \theta + \varepsilon) =$$

$$P(x_1 \geq \frac{\theta + \varepsilon}{n+1}, \dots, x_n \geq \frac{\theta + \varepsilon}{n+1}) = \prod_{i=1}^n P(x_i \geq \frac{\theta + \varepsilon}{n+1}) =$$

$$= (1 - P(x < \frac{\theta + \varepsilon}{n+1}))^n = (1 - F(x < \frac{\theta + \varepsilon}{n+1}))^n =$$

$$= (1 - \frac{1}{\theta} \frac{\theta + \varepsilon}{n+1})^n \xrightarrow{n \rightarrow \infty} e^{-\frac{\theta + \varepsilon}{\theta}} \Rightarrow$$



$\Rightarrow \tilde{\theta}_2$  - не состоятельная оценка

Третья оценка:  $\tilde{\theta}_3 = x_{\max}$

$$\varphi(x) = (F(x))^n = \left(\frac{x}{\theta}\right)^n$$

$$\varphi(x) = \varphi'(x) = n \left(\frac{x}{\theta}\right)^{n-1} \frac{1}{\theta} = \frac{nx^{n-1}}{\theta^n}$$

а) Несмещенность:

$$M[\tilde{\theta}_3] = \int_0^{\theta} \frac{nx^n}{\theta^n} dx = \frac{n}{n+1} \frac{x^{n+1}}{\theta^n} \Big|_0^{\theta} = \frac{n}{n+1} \theta \Rightarrow$$

$\Rightarrow \tilde{\theta}_3$  - смещена

Коррекция:  $\tilde{\theta}_3' = \frac{n+1}{n} \tilde{\theta}_3 = \frac{n+1}{n} x_{\max}$

б) Состоятельность

Проверка самостоятельно уверенно

$$M[\tilde{\theta}_3^2] = \int_0^{\theta} \frac{nx^{n+1}}{\theta^n} = \frac{n}{n+2} \frac{x^{n+2}}{\theta^n} \Big|_0^{\theta} = \frac{n}{n+2} \theta^2$$

$$D[\tilde{\theta}_3] = M[\tilde{\theta}_3^2] - M^2[\tilde{\theta}_3] = \frac{n}{n+2} \theta^2 - \frac{n^2}{(n+1)^2} \theta^2 =$$

$$= n \left( \frac{(n+1)^2 \theta^2 - n(n+2) \theta^2}{(n+2)(n+1)^2} \right) = \frac{n \theta^2}{(n+2)(n+1)^2}$$

$$= n \left( \frac{n^2 \theta^2 + 2n \theta^2 + \theta^2 - n^2 \theta^2 - 2n \theta^2}{(n+2)(n+1)^2} \right) = \frac{n \theta^2}{(n+2)(n+1)^2}$$



Дисперсия несмещённой оценки

$$D[\tilde{\theta}_3'] = D\left[\frac{n+1}{n} X_{\max}\right] =$$

$$= \frac{(n+1)^2}{n^2} D[\tilde{\theta}_3] = \frac{\theta^2}{n(n+2)} \xrightarrow{n \rightarrow \infty} 0$$

$\Rightarrow \tilde{\theta}_3'$  — состоятельна

Чемберская оценка:  $\tilde{\theta}_4 = X_{\min} + X_{\max}$

а) Несмещённость

$$M[\tilde{\theta}_4] = M[\tilde{\theta}_2 + \tilde{\theta}_3] = M[\tilde{\theta}_2] + M[\tilde{\theta}_3] =$$

$$\frac{\theta}{n+1} + \frac{n}{n+1} \theta = \theta \Rightarrow \tilde{\theta}_4 \text{ — не смещена}$$

б) Состоятельность

Проверка гомогенного условия:

$$D[\tilde{\theta}_4] = D[\tilde{\theta}_2] + D[\tilde{\theta}_3] + 2 \operatorname{cov}(\tilde{\theta}_2, \tilde{\theta}_3)$$

Плотность распределения

$$g(y, z) = n(n-1) \left(\frac{z}{\theta} - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta^2} (z \geq y, z, y \in [0, \theta])$$

$$M[\tilde{\theta}_2 \tilde{\theta}_3] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} yz g(y, z) = \int_0^\theta dz \int_0^z yz n(n-1) \left(\frac{z}{\theta} - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta^2} dy$$

Внутренний интеграл

$$\int_0^z yz n(n-1) \left(\frac{z}{\theta} - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta^2} dy = \left\{ t = \frac{y}{z} \right\} =$$



$$\begin{aligned}
 &= \int_0^1 z^2 t^{n(n-1)} (1-t)^{n-2} \frac{1}{\theta^n} z^{n-2} \cdot z dt = \\
 &= z^{n+1} \frac{n(n-1)}{\theta^n} B(2, n-1) = \frac{z^{n+1}}{\theta^n} n(n-1) \frac{\Gamma(2) \Gamma(n-1)}{\Gamma(n+1)} = \\
 &= \frac{z^{n+1}}{\theta^n} n(n-1) \frac{\Gamma(n-1)}{n \Gamma(n)} = \frac{z^{n+1}}{\theta^n} n(n-1) \frac{\Gamma(n-1)}{n(n-1) \Gamma(n-1)} = \\
 &= \frac{z^{n+1}}{\theta^n}
 \end{aligned}$$

Внешний интеграл:

$$\int_0^\theta \frac{z^{n+1}}{\theta^n} dz = \frac{1}{n+2} \frac{z^{n+2}}{\theta^n} \Big|_0^\theta = \frac{\theta^2}{n+2} \Rightarrow$$

$$\begin{aligned}
 \Rightarrow \text{cov}(\tilde{\theta}_2, \tilde{\theta}_3) &= M[\tilde{\theta}_2 \tilde{\theta}_3] - M[\tilde{\theta}_2] M[\tilde{\theta}_3] = \\
 &= \frac{\theta^2}{n+2} - \frac{n\theta^2}{(n+1)^2} = \frac{n^2\theta^2 + 2n\theta^2 + \theta^2 - n^2\theta^2 - 2n\theta^2}{(n+2)(n+1)^2} = \\
 &= \frac{\theta^2}{(n+2)(n+1)^2}
 \end{aligned}$$

$$D[\tilde{\theta}_4] = \frac{2n\theta^2}{(n+2)(n+1)^2} + \frac{2\theta^2}{(n+2)(n+1)^2} = \frac{2\theta^2}{(n+2)(n+1)} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow$$

$\Rightarrow \tilde{\theta}_4$  - самосогласованна

Тогда оценка:  $\tilde{\theta}_5 = x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i$

а) Несмещенность

$$\begin{aligned}
 M[\tilde{\theta}_5] &= M\left[x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i\right] = M[f] + \frac{1}{n-1} \sum_{i=2}^n M[f] = \\
 &= M[f] + \frac{1}{n-1} (n-1) M[f] = 2M[f] = \theta \Rightarrow
 \end{aligned}$$



$\Rightarrow \tilde{\theta}_5$  — не смещена

б) Состоятельность

Проверка достаточного условия:

$$D[\tilde{\theta}_5] = D\left[x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i\right] =$$

$$= D[f] + \frac{1}{(n-1)^2} \sum_{i=2}^n D[f] = D[f] + \frac{1}{n-1} D[f] =$$

$$= \frac{n}{n-1} D[f] \xrightarrow{n \rightarrow \infty} 0$$

По определению:

$$x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i \xrightarrow{P} f + M[f] \xrightarrow{n \rightarrow \infty} 0 \Rightarrow$$

$\Rightarrow \tilde{\theta}_5$  — не состоятельна

Определение наиболее эффективной

оценки:

$$D[\tilde{\theta}_1] = \frac{\theta^2}{3n}$$

$$D[\tilde{\theta}_3'] = \frac{\theta^2}{n(n+2)}$$

$$D[\tilde{\theta}_4] = \frac{2\theta^2}{(n+2)(n+1)}$$

$$\frac{1}{n(n+2)} \leq \frac{1 \cdot 2}{(n+2)(n+1)}$$

$$\frac{1}{n} \leq \frac{2}{n+1} ; n+1 \leq 2n ; n \leq 2n-1 \quad \forall n \Rightarrow$$



$\Rightarrow$  прямая оценка самая эффективная

III-3

$$f \sim p(x) = \begin{cases} \frac{1}{\theta} \exp(-\frac{x}{\theta}), & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Требуется оценка:  $\tilde{\theta} = \star$

$$\begin{aligned} M[f] &= \int_0^{+\infty} \frac{x}{\theta} \exp(-\frac{x}{\theta}) dx = -x \exp(-\frac{x}{\theta}) \Big|_0^{+\infty} + \\ &+ \int_0^{+\infty} \exp(-\frac{x}{\theta}) dx = -\theta \exp(-\frac{x}{\theta}) \Big|_0^{+\infty} = \theta \end{aligned}$$

$$D[f] = M[f^2] - M[f]^2$$

$$M[f^2] = \int_0^{+\infty} \frac{x^2}{\theta} \exp(-\frac{x}{\theta}) dx = -x^2 \exp(-\frac{x}{\theta}) \Big|_0^{+\infty} +$$

$$2 \int_0^{+\infty} x \exp(-\frac{x}{\theta}) dx = -2\theta x \exp(-\frac{x}{\theta}) \Big|_0^{+\infty} +$$

$$+ 2\theta \int_0^{+\infty} \exp(-\frac{x}{\theta}) dx = -2\theta^2 \exp(-\frac{x}{\theta}) \Big|_0^{+\infty} = 2\theta^2$$

$$D[f] = 2\theta^2 - \theta^2 = \theta^2$$



Несмещенность:

$$\begin{aligned} M[\tilde{\theta}_1] &= M[\bar{x}] = M\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n} \sum_{i=1}^n M[x_i] = \\ &= \frac{1}{n} \cdot n \cdot \theta = \theta \Rightarrow M[\tilde{\theta}_1] = \theta \Rightarrow \text{несмещенная} \end{aligned}$$

$$D[\tilde{\theta}_1] = D\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n^2} \sum_{i=1}^n D[x_i] = \frac{\theta^2}{n}$$

Вторая оценка:  $\tilde{\theta}_2 = \frac{x_{\min} + x_{\max}}{2}$

$$M[\tilde{\theta}_2] = \frac{1}{2} (M[x_{\min}] + M[x_{\max}])$$

Считаем  $M[x_{\min}]$ :

Функция распределения:  $F_{\min}(x) = 1 - (1 - F(x))^n =$

$$= 1 - \left(1 - \int_0^x \frac{1}{\theta} e^{-\frac{t}{\theta}} dt\right)^n = 1 - \left(1 + e^{-\frac{x}{\theta}}\right)^n =$$

$$= 1 - e^{-\frac{xn}{\theta}}$$

$$p(x) = \left(1 - e^{-\frac{xn}{\theta}}\right)' = \frac{n}{\theta} e^{-\frac{xn}{\theta}} = \frac{\theta}{3} \frac{3}{\theta} e^{-\frac{3x}{\theta}}$$

$$M[x_{\min}] = \int_0^{+\infty} \frac{3x}{\theta} e^{-\frac{3x}{\theta}} dx = \left\{ t = \frac{3x}{\theta} \right\} =$$

$$= \int_0^{+\infty} \frac{\theta}{3} t e^{-t} dt = \frac{\theta}{3} \Gamma(2) = \frac{\theta}{3}$$

Считаем  $M[x_{\max}]$ :

Функция распределения:  $F_{\max}(x) = (1 - e^{-\frac{x}{\theta}})^n$



$$p(x) = n(1 - e^{-\frac{x}{\theta}})^{n-1} \frac{1}{\theta} e^{-\frac{x}{\theta}} = \frac{3}{\theta} e^{-\frac{x}{\theta}} (1 - e^{-\frac{x}{\theta}})^2$$

$$= \frac{3}{\theta} e^{-\frac{x}{\theta}} (1 - 2e^{-\frac{x}{\theta}} + e^{-\frac{2x}{\theta}}) =$$

$$= \frac{3}{\theta} (e^{-\frac{x}{\theta}} - 2e^{-\frac{2x}{\theta}} + e^{-\frac{3x}{\theta}})$$

$$M[X_{\max}] = \int_0^{+\infty} \frac{3x}{\theta} e^{-\frac{x}{\theta}} dx - \frac{3}{\theta} \int_0^{+\infty} \frac{2x}{\theta} e^{-\frac{2x}{\theta}} dx +$$

$$+ \int_0^{+\infty} \frac{3x}{\theta} e^{-\frac{3x}{\theta}} dx = 3 \int_0^{+\infty} t e^{-t} dt - 3 \int_0^{+\infty} \frac{t}{2} e^{-t} dt$$

$$+ \int_0^{+\infty} \frac{t}{3} e^{-t} dt = 3\theta - \frac{3}{2}\theta + \frac{\theta}{3} = \frac{3}{2}\theta + \frac{\theta}{3}$$

$$M[\tilde{\theta}_2] = \frac{1}{2} \left( \frac{3}{2}\theta + \frac{\theta}{3} + \frac{\theta}{3} \right) = \frac{1}{2} \left( \frac{3}{2}\theta + \frac{2}{3}\theta \right) =$$

$$= \frac{13}{12}\theta \Rightarrow \text{оценка — смещенная.}$$

$$\text{Корректировка: } \tilde{\theta}_2' = \frac{12}{13} \tilde{\theta}_2 = \frac{6}{13} (X_{\min} + X_{\max})$$

Дисперсия:

$$D[X_{\min}] = M[X_{\min}^2] - M^2[X_{\min}]$$

$$M[X_{\min}^2] = \int_0^{+\infty} \frac{3}{\theta} x^2 e^{-\frac{3x}{\theta}} dx = \left\{ t = \frac{3x}{\theta} \right\} = \int_0^{+\infty} \frac{\theta^2}{9} t^2 e^{-t} dt =$$

$$= \frac{\theta^2}{9} \Gamma(3) = \frac{2}{9}\theta^2$$

$$MD[f] = \frac{2}{9}\theta^2 - \left(\frac{\theta}{3}\right)^2 = \frac{\theta^2}{9}$$



$$D[X_{\max}] = M[X_{\max}^2] - M^2[X_{\max}]$$

$$\begin{aligned} M[X_{\max}^2] &= \int_0^{+\infty} \frac{3x^2}{\theta} e^{-\frac{x}{\theta}} dx - 3 \int_0^{+\infty} \frac{2x^2}{\theta} e^{-\frac{2x}{\theta}} dx + \\ &+ \int_0^{+\infty} \frac{3x^2}{\theta} e^{-\frac{3x}{\theta}} dx = 3\theta^2 \Gamma(3) - \frac{3}{4}\theta^2 \Gamma(3) + \\ &+ \frac{\theta^2}{9} \Gamma(3) = 2 \left( 3\theta^2 - \frac{3}{4}\theta^2 + \frac{1}{9}\theta^2 \right) = \\ &= \frac{9}{2}\theta^2 + \frac{2}{9}\theta^2 \end{aligned}$$

$$\begin{aligned} D[X_{\max}] &= \frac{9}{2}\theta^2 + \frac{2}{9}\theta^2 - \left( \frac{3}{2}\theta^2 + \frac{\theta}{3} \right)^2 = \\ &= \frac{9}{2}\theta^2 + \frac{2}{9}\theta^2 - \frac{9}{4}\theta^2 - \theta^2 - \frac{1}{9}\theta^2 = \\ &= \frac{9}{4}\theta^2 - \frac{4}{9}\theta^2 + \frac{1}{9}\theta^2 = \frac{5}{4}\theta^2 + \frac{1}{9}\theta^2 = \frac{49}{36}\theta^2 \end{aligned}$$

$$D[\tilde{\theta}_2] = \frac{1}{4}D[X_{\min}] + \frac{1}{4}D[X_{\max}] + \frac{2}{4}\text{COV}(X_{\min}, X_{\max})$$

$$\text{COV}(X_{\min}, X_{\max}) = M[X_{\min} \cdot X_{\max}] - M[X_{\min}]M[X_{\max}] =$$

$$M[X_{\min} X_{\max}] = \int_0^{+\infty} \int_0^{+\infty} g(y, z) yz dy dz$$

$$g(y, z) = n(n-1)(F(z) - F(y))^{n-2} p(z)p(y), \quad z \geq y$$

$$g(y, z) = 6(1 - e^{-\frac{z}{\theta}} - 1 + e^{-\frac{y}{\theta}}) \frac{1}{\theta^2} e^{-\frac{z}{\theta}} \cdot e^{-\frac{y}{\theta}} =$$

$$= 6(e^{-\frac{y}{\theta}} - e^{-\frac{z}{\theta}}) \frac{1}{\theta^2} e^{-\frac{z}{\theta}} \cdot e^{-\frac{y}{\theta}}$$



$$M[x_{\min} x_{\max}] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} yz \cdot 6(e^{-\frac{y}{\theta}} - e^{-\frac{z}{\theta}}) \frac{1}{\theta^2} e^{-\frac{z}{\theta}} \cdot e^{-\frac{y}{\theta}} dy dz =$$

$$= \int_0^{+\infty} dz \int_0^z 6yz (e^{-\frac{y}{\theta}} - e^{-\frac{z}{\theta}}) \frac{1}{\theta^2} e^{-\frac{z}{\theta}} \cdot e^{-\frac{y}{\theta}} dy$$

~~Bruggenwert numerieren:~~

~~$$\int_0^z \frac{6}{\theta^2} yz (e^{-\frac{y}{\theta}} - e^{-\frac{z}{\theta}}) e^{-\frac{z}{\theta}} \cdot e^{-\frac{y}{\theta}} dy =$$

$$= \int_0^z \frac{3}{\theta} z e^{-\frac{z}{\theta}} \frac{2y}{\theta} e^{-\frac{2y}{\theta}} dy - \int_0^z \frac{6}{\theta} z e^{-\frac{2z}{\theta}} \frac{y}{\theta} e^{-\frac{y}{\theta}} dy =$$

$$= \int_0^{\frac{2z}{\theta}} \frac{3}{\theta} \cdot \frac{\theta}{2} z e^{-\frac{z}{\theta}} t e^{-t} dt - \int_0^{\frac{2z}{\theta}} \frac{6}{\theta} \cdot \theta z e^{-\frac{2z}{\theta}} t e^{-t} dt$$

$$= \frac{3}{2} z e^{-\frac{z}{\theta}} \int_0^{\frac{2z}{\theta}} t e^{-t} dt - 6 z e^{-\frac{2z}{\theta}} \int_0^{\frac{2z}{\theta}} t e^{-t} dt$$~~

~~Bruggenwert numerieren:~~

~~$$\frac{6}{\theta^2} z e^{-\frac{z}{\theta}} \int_0^z [y e^{-\frac{2y}{\theta}} - y e^{-\frac{y}{\theta}} \cdot e^{-\frac{z}{\theta}}] dy =$$

$$= \frac{6}{\theta^2} z e^{-\frac{z}{\theta}} \left( \frac{\theta^2}{4} - \theta^2 e^{-\frac{z}{\theta}} + \frac{3}{4} \theta^2 e^{-\frac{2z}{\theta}} + \frac{1}{2} \theta z e^{-\frac{2z}{\theta}} \right)$$~~

~~Bruggenwert numerieren:~~

~~$$\int_0^{+\infty} \frac{6}{\theta^2} z e^{-\frac{z}{\theta}} \left( \frac{\theta^2}{4} - \theta^2 e^{-\frac{z}{\theta}} + \frac{3}{4} \theta^2 e^{-\frac{2z}{\theta}} + \frac{1}{2} \theta z e^{-\frac{2z}{\theta}} \right) dz =$$~~



$$= \frac{13}{18} \theta^2$$

$$M[X_{\min} X_{\max}] = \frac{13}{18} \theta^2$$

$$\text{COV} = \frac{13}{18} \theta^2 - \frac{\theta}{3} \cdot \frac{11\theta}{6} = \frac{2}{18} \theta^2$$

$$\begin{aligned} D[(X_{\min} + X_{\max}) \cdot \frac{1}{2}] &= \frac{1}{4} D[X_{\min}] + \frac{1}{4} D[X_{\max}] + \\ &+ \frac{1}{2} \text{COV}(X_{\min}, X_{\max}) = \frac{1}{4} \frac{\theta^2}{9} + \frac{1}{4} \frac{49\theta^2}{36} + \frac{\theta^2}{18} = \\ &= \frac{61}{144} \theta^2 \end{aligned}$$

$$D[\tilde{\theta}_2'] = \frac{144}{169} \cdot \frac{61}{144} \theta^2 = \frac{61}{169} \theta^2$$

III. третья оценка:  $\tilde{\theta}_3 = X_{(2)}$

$$p(x) = 6 \cdot \frac{1}{\theta} e^{-\frac{x}{\theta}} (1 - e^{-\frac{x}{\theta}}) e^{-\frac{x}{\theta}} = \frac{6}{\theta} (e^{-\frac{2x}{\theta}} - e^{-\frac{3x}{\theta}})$$

$$M[\tilde{\theta}_3] = \int_0^{+\infty} \frac{6x}{\theta} (e^{-\frac{2x}{\theta}} - e^{-\frac{3x}{\theta}}) dx = \frac{5}{6} \theta \Rightarrow \text{смещение}$$

И корректируем:  $\tilde{\theta}_3' = \frac{6}{5} \tilde{\theta}_3$

Дисперсия:

$$D[\tilde{\theta}_3] = M[\tilde{\theta}_3^2] - M^2[\tilde{\theta}_3]$$

$$M[\tilde{\theta}_3^2] = \int_0^{+\infty} \frac{6x^2}{\theta} (e^{-\frac{2x}{\theta}} - e^{-\frac{3x}{\theta}}) dx = \frac{19}{18} \theta^2$$

$$D[\tilde{\theta}_3] = \frac{19}{18} \theta^2 - \frac{25}{36} \theta^2 = \frac{13}{36} \theta^2$$



$$D[\tilde{\Theta}_3'] = \frac{36}{25} \frac{13}{36} \Theta^2 = \frac{13}{25} \Theta^2$$

Исследование экстремумов:

$$D[\tilde{\Theta}_1] = \frac{\Theta^2}{3}, \quad D[\tilde{\Theta}_2'] = \frac{61}{169} \Theta^2, \quad D[\tilde{\Theta}_3'] = \frac{13}{25} \Theta^2 \Rightarrow$$

$\Rightarrow$  при  $n=3$   $\tilde{\Theta}_1$  - наиболее экстремумов

б) Исследование с помощью неравенства

Крамера-Пао

$$D[\tilde{\Theta}_i] \geq \frac{1}{n I(\Theta)}, \quad I(\Theta) = M\left[\left(\frac{\partial \ln p(x, \Theta)}{\partial \Theta}\right)^2\right]$$

$$\ln p(x, \Theta) = \ln \frac{1}{\Theta} e^{-\frac{x}{\Theta}} = \ln \frac{1}{\Theta} + \ln e^{-\frac{x}{\Theta}} = \ln \frac{1}{\Theta} - \frac{x}{\Theta}$$

$$\frac{\partial \ln p(x, \Theta)}{\partial \Theta} = 1 / \frac{1}{\Theta} \cdot \left(-\frac{1}{\Theta^2}\right) + \frac{x}{\Theta^2} = \frac{x}{\Theta^2} - \frac{1}{\Theta}$$

$$\left(\frac{x}{\Theta^2} - \frac{1}{\Theta}\right)^2 = \frac{x^2}{\Theta^4} - \frac{2x}{\Theta^3} + \frac{1}{\Theta^2}$$

$$M\left[\left(\frac{\partial \ln p(x, \Theta)}{\partial \Theta}\right)^2\right] = M\left[\frac{x^2}{\Theta^4} - \frac{2x}{\Theta^3} + \frac{1}{\Theta^2}\right] =$$

$$= \int_0^{+\infty} \left(\frac{x^2}{\Theta^4} - \frac{2x}{\Theta^3} + \frac{1}{\Theta^2}\right) \frac{1}{\Theta} e^{-\frac{x}{\Theta}} dx =$$

$$= \int_0^{+\infty} \frac{x^2}{\Theta^5} e^{-\frac{x}{\Theta}} dx - \int_0^{+\infty} \frac{2x}{\Theta^4} e^{-\frac{x}{\Theta}} dx + \int_0^{+\infty} \frac{1}{\Theta^3} e^{-\frac{x}{\Theta}} dx$$

$$= \frac{2}{\Theta^2} - \frac{2}{\Theta^2} - \frac{1}{\Theta^2} e^{-\frac{x}{\Theta}} \Big|_0^{+\infty} = \frac{1}{\Theta^2}$$



$$\frac{1}{nI(\theta)} = \frac{1}{n\theta^2} = \frac{\theta^2}{3} = D[\tilde{\theta}_1] \Rightarrow \frac{\partial^2}{\partial \theta^2} \ln L(\theta) = -\frac{2}{\theta^3}$$

$\Rightarrow \tilde{\theta}_1$  эффективна по Крамеру-Рао.

Остальные оценки не эффективны по Крамеру Рао.

Следовательно, оценка  $\tilde{\theta}_1$  является эффективной по Крамеру-Рао.

$$E\left(\frac{\partial \ln L(\theta)}{\partial \theta}\right) = 0; \quad \frac{1}{nI(\theta)} \leq D[\tilde{\theta}_1]$$

$$\frac{\partial}{\partial \theta} \ln L(\theta) = \frac{\partial}{\partial \theta} \ln \left( \frac{1}{\theta^3} \right) = -\frac{3}{\theta^4}$$

$$\frac{\partial^2}{\partial \theta^2} \ln L(\theta) = \frac{\partial}{\partial \theta} \left( -\frac{3}{\theta^4} \right) = \frac{12}{\theta^5}$$

$$\frac{1}{nI(\theta)} = \frac{\theta^5}{12}$$

$$E\left(\frac{\partial^2 \ln L(\theta)}{\partial \theta^2}\right) = \frac{12}{\theta^5}$$

$$= \frac{12}{\theta^5}$$

$$= \frac{12}{\theta^5}$$

$$= \frac{12}{\theta^5}$$