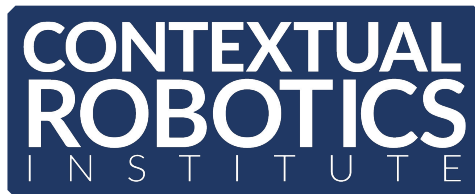


Localization and Mapping

Arash Asgharivaskasi Nikolay Atanasov

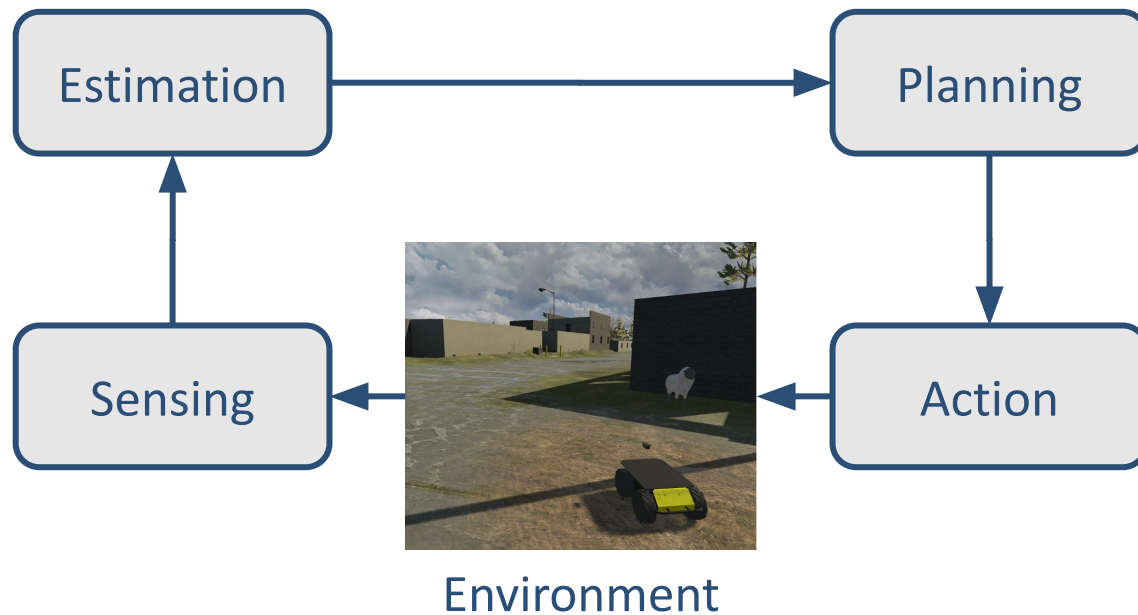
Existential Robotics Laboratory

University of California, San Diego

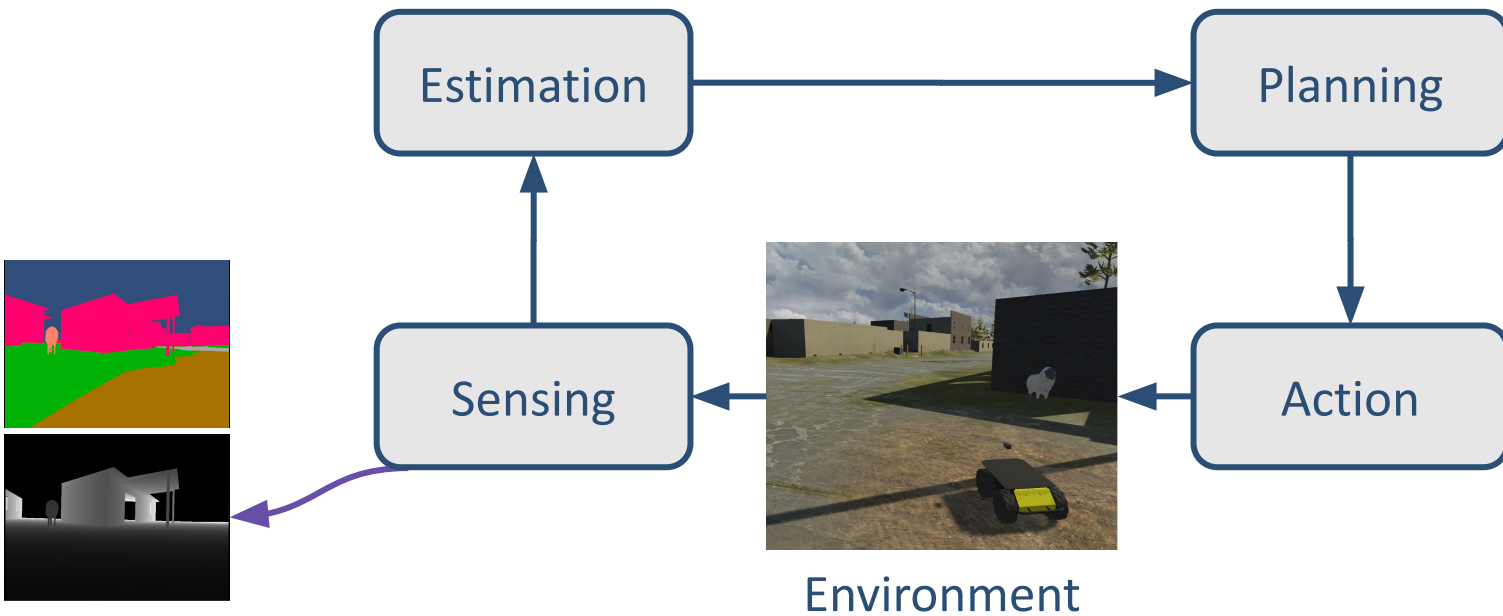


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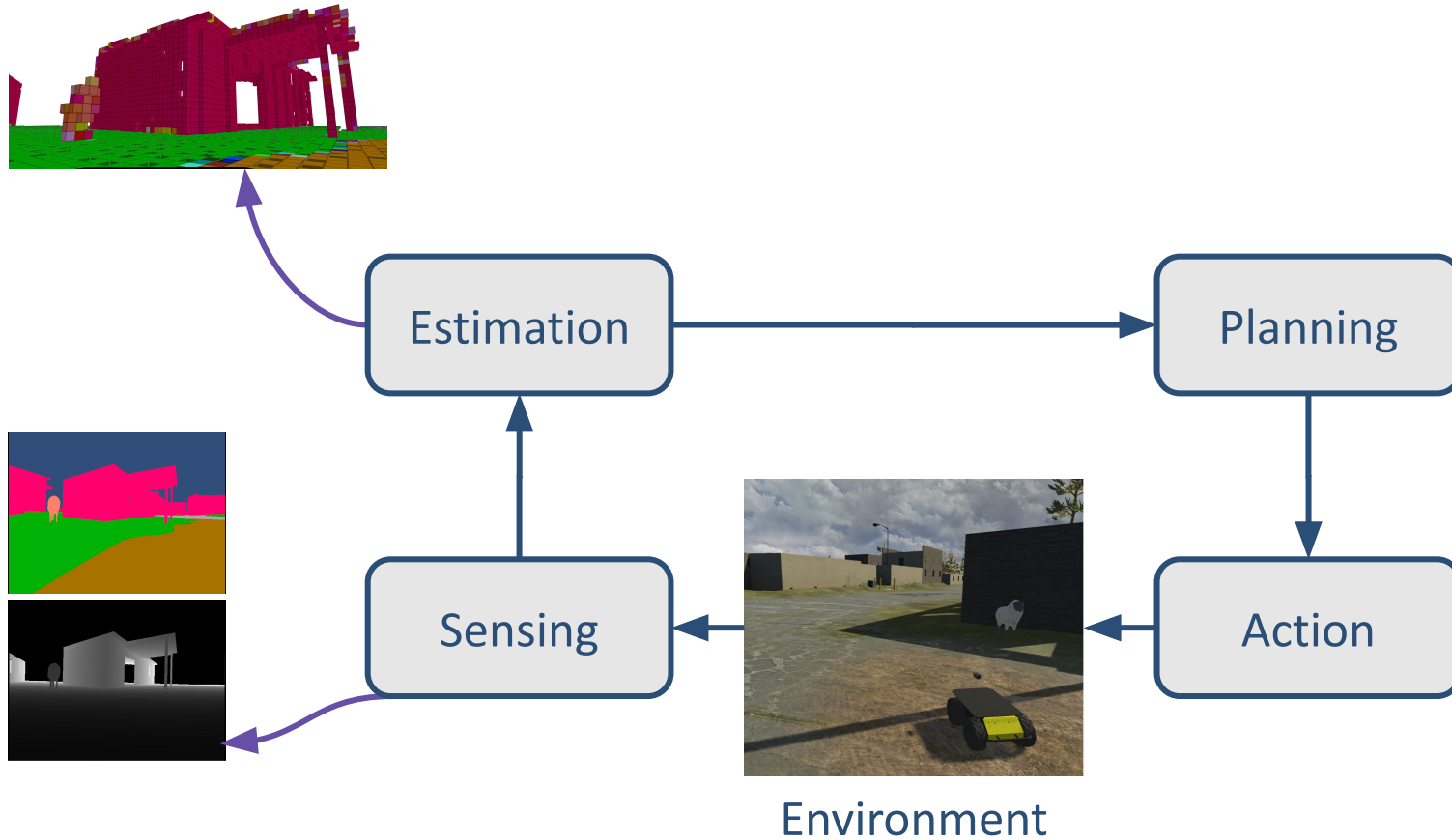
Robot Autonomy



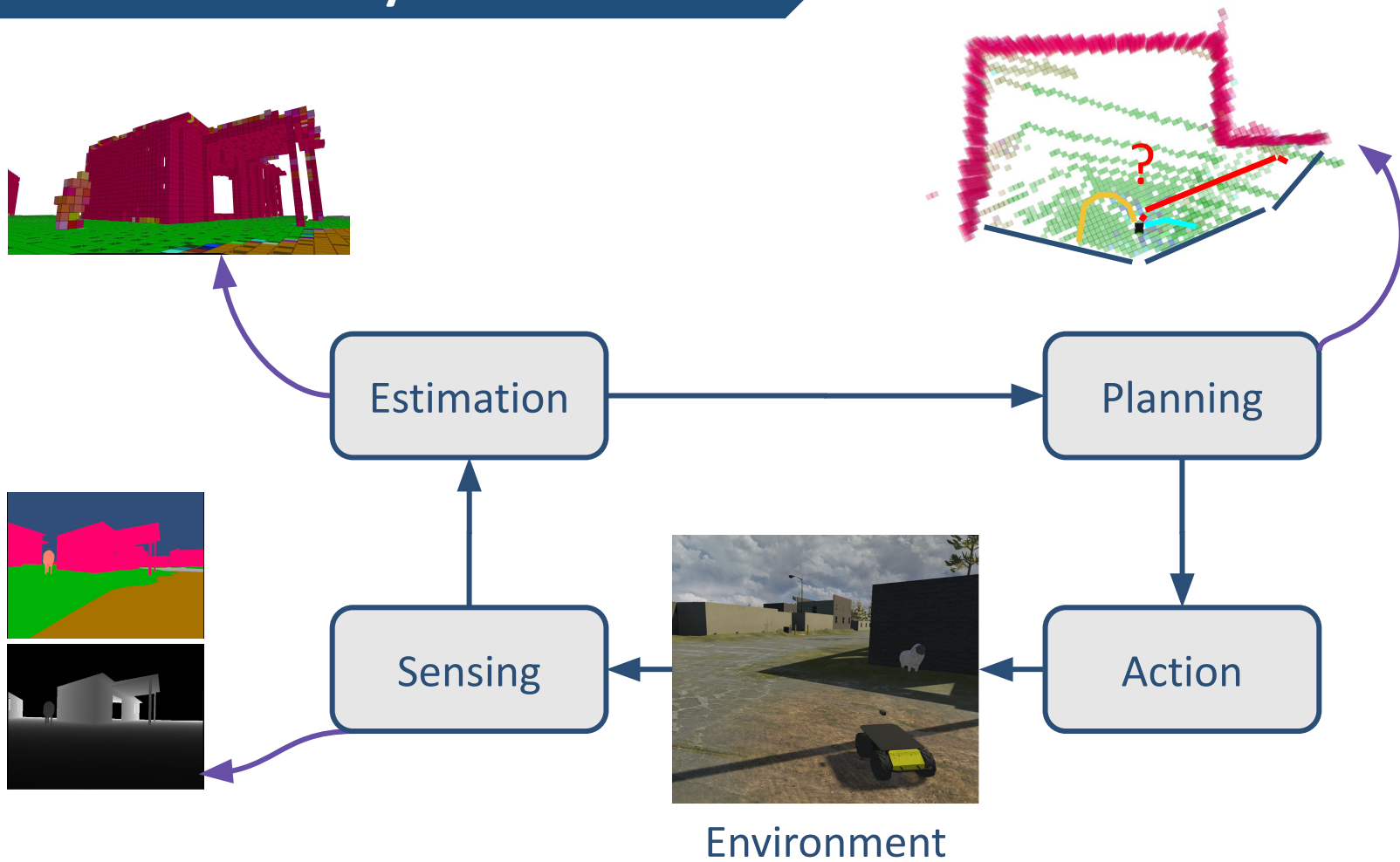
Robot Autonomy



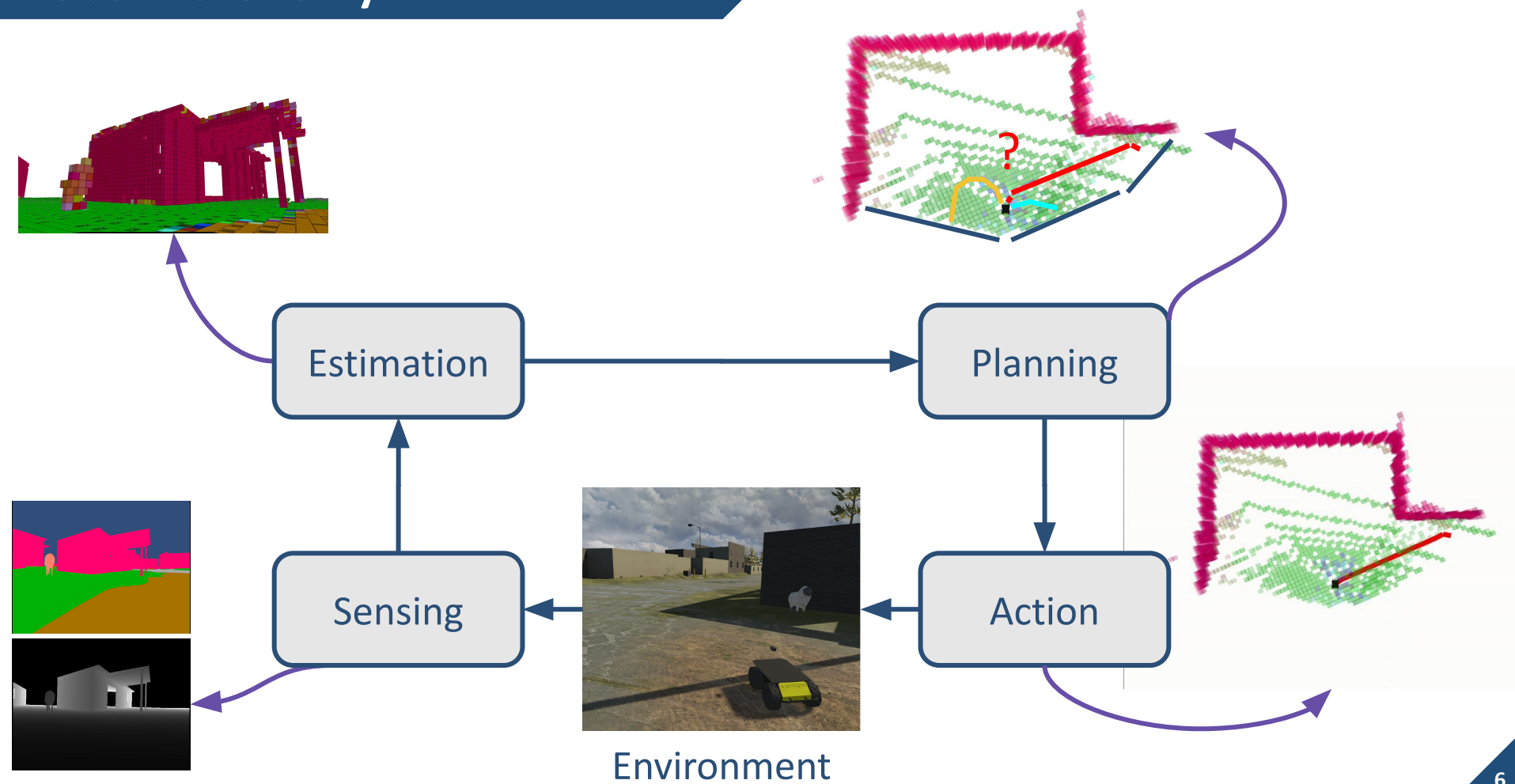
Robot Autonomy



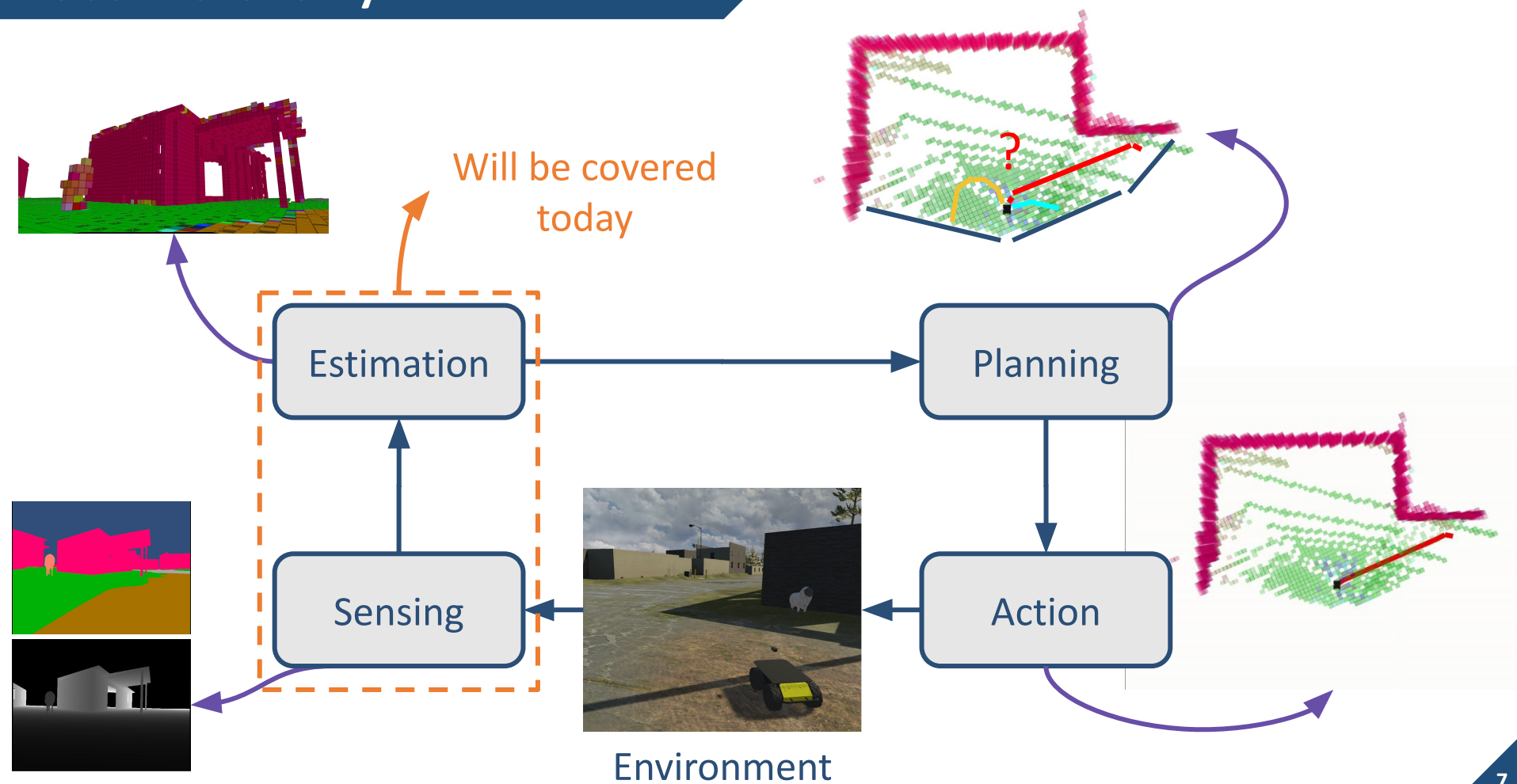
Robot Autonomy



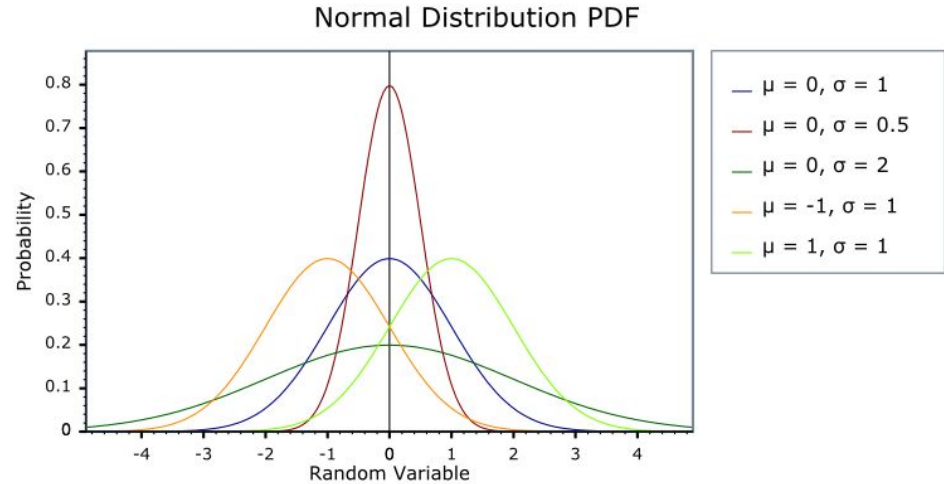
Robot Autonomy



Robot Autonomy



Probability Review



Probability Axioms

- For an event A in sample space Ω , we assume that the following is always true:
 1. $P(A) \geq 0$
 2. $P(\Omega) = 1$
 3. If $A_i \cap A_j = \emptyset, \forall i \neq j$, then $P\left(\bigcup_i A_i\right) = \sum_i P(A_i)$
- Corollary:
 1. $P(\emptyset) = 0$
 2. $\max\{P(A), P(B)\} \leq P(A \cup B) = P(A) + P(B) - P(A \cap B) \leq P(A) + P(B)$
 3. $A \subseteq B \Rightarrow P(A) \leq P(B)$

Probability Formulas

- Total Probability:

$$P(B) = \sum_{i=1}^n P(B \cap A_i), \text{ if } \Omega = \bigcup_i A_i \text{ and } A_i \cap A_j = \emptyset, \forall i \neq j$$

- Conditional Probability:

$$P(A \cap B) = P(A|B)P(B)$$

- Bayes Theorem:

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}, P(B) \neq 0$$

- Corollary:

$$P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_{i=1}^n P(B|A_i)P(A_i)}, \text{ if } \Omega = \bigcup_i A_i \text{ and } A_i \cap A_j = \emptyset, \forall i \neq j$$

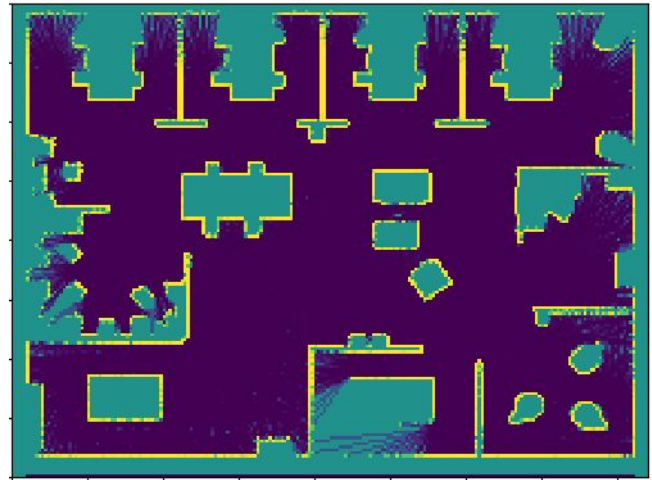
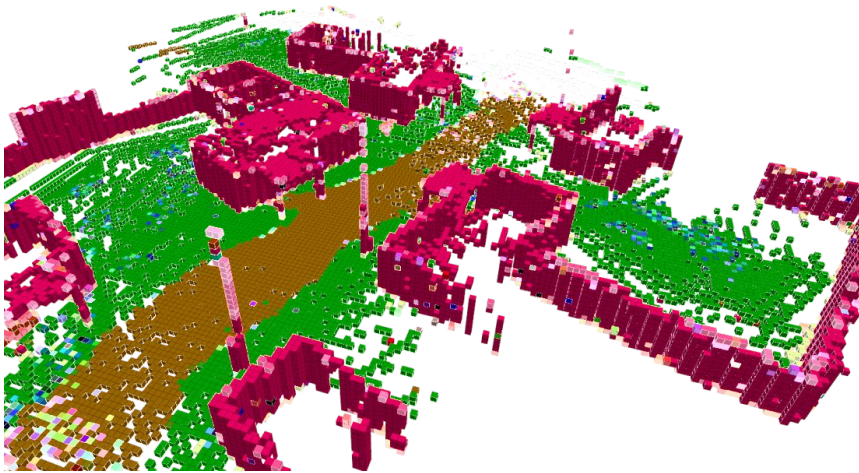
Independent Events

- Events A and B are independent if and only if:

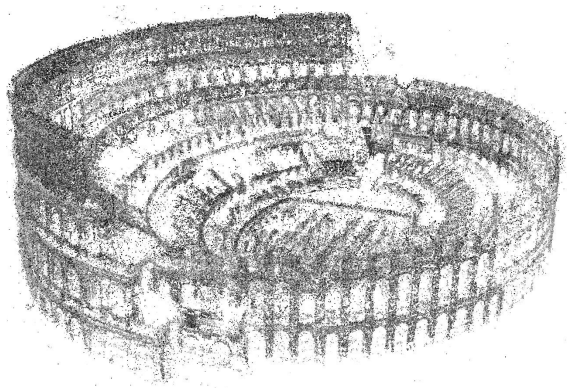
$$P(A|B) = P(A) \Leftrightarrow P(A \cap B) = P(A)P(B)$$

- Observing one does not give any information about another
- Disjoint events are always dependent (why?)

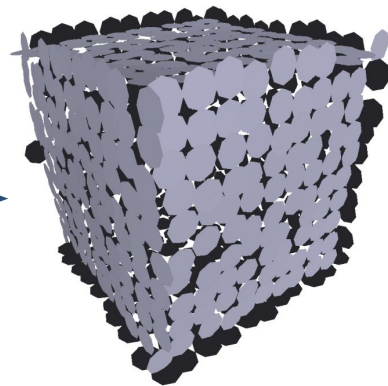
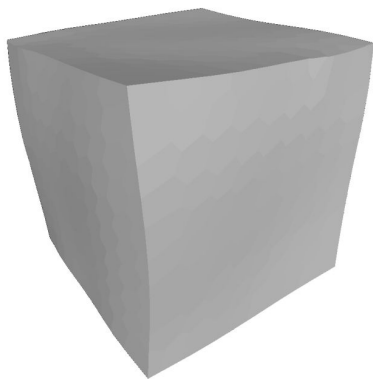
Mapping



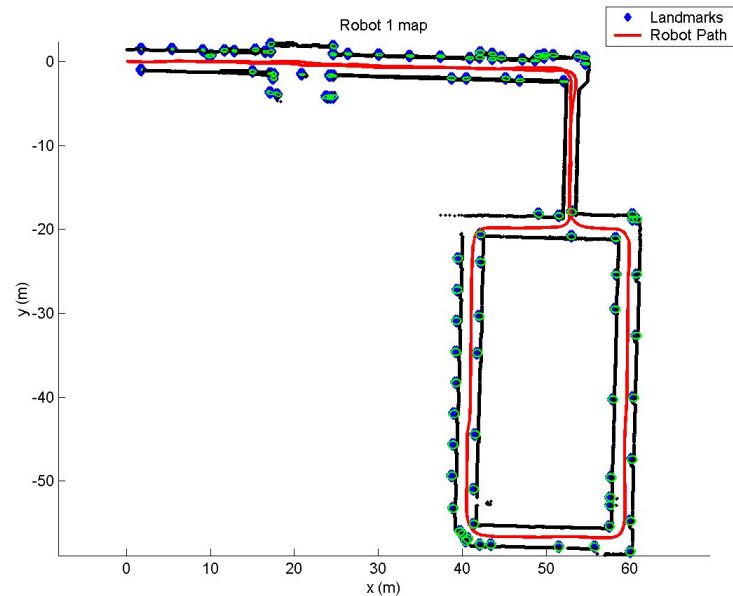
Map Representations (Sparse)



Point Cloud

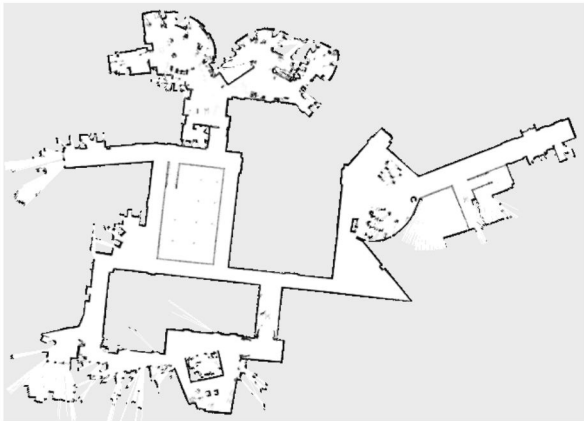


Surfels

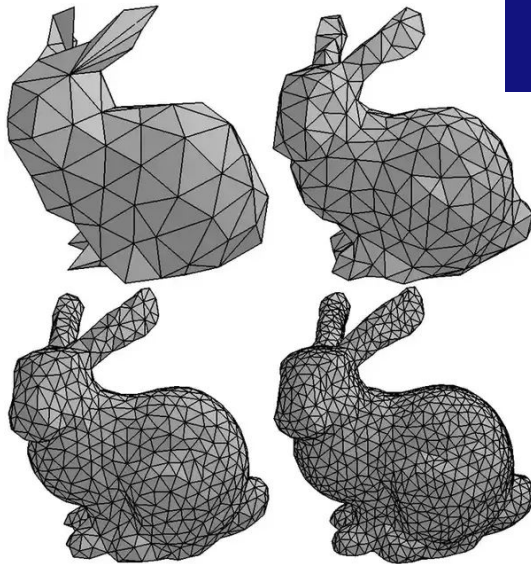


Landmark-based

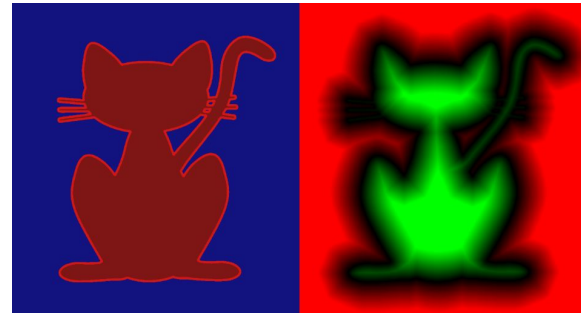
Map Representations (Dense)



Occupancy Grid Map

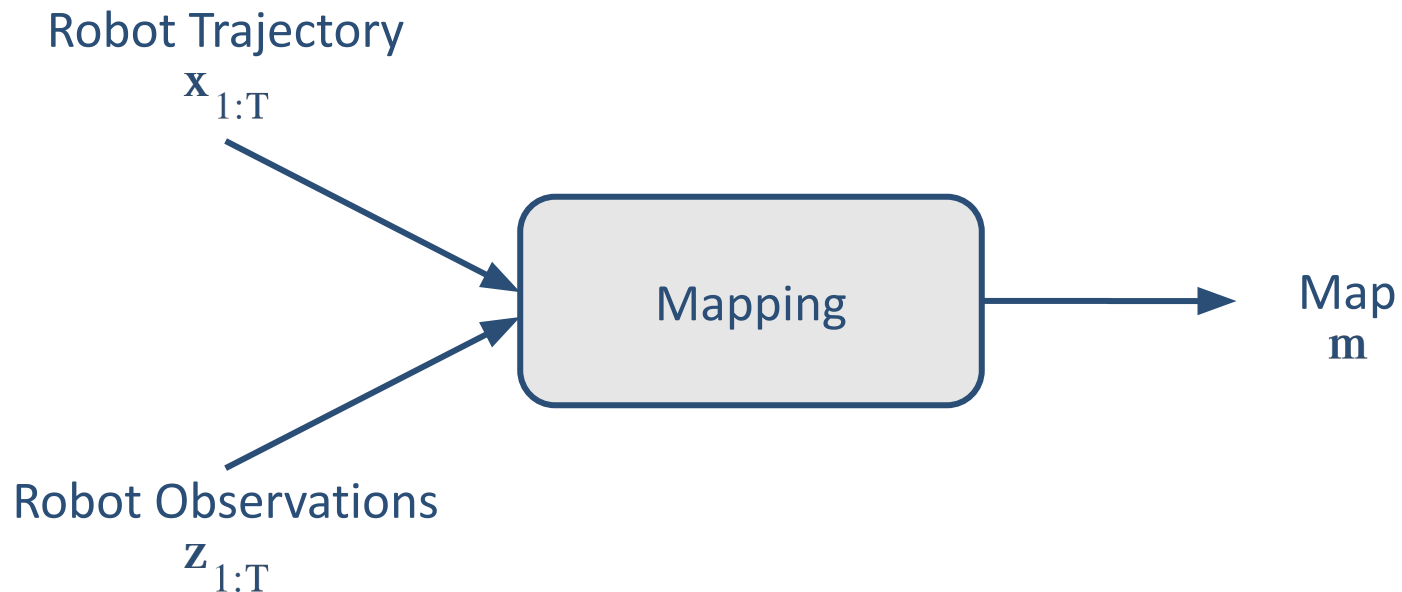


Polygon Mesh



Signed Distance Field

Problem Formulation



Problem Formulation

Robot 2-D Trajectory
 $(x, y, \theta)_{1:T}$

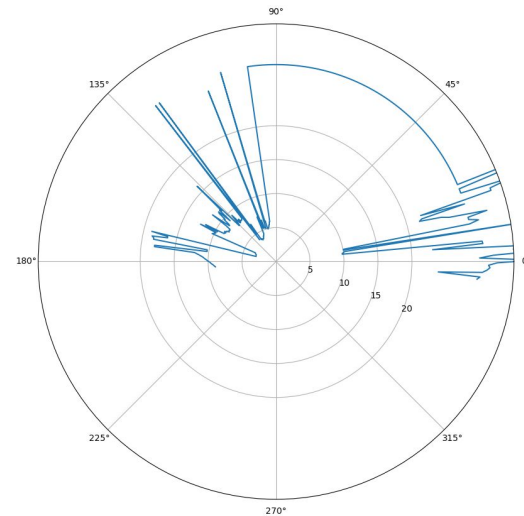
LiDAR Scans
 $(r, \varphi)_{1:T}$

Log-odds Mapping

Probabilistic
Occupancy Grid
 $P(m)$

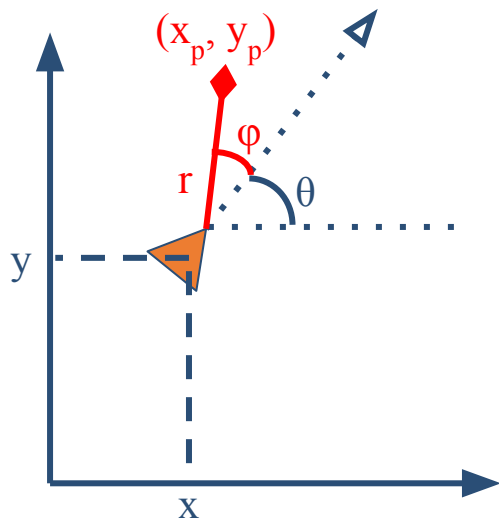
LiDAR Scanner

- LiDAR: Light Detection And Ranging
- Illuminates the scene with pulsed laser light and measures the return times and wavelengths of the reflected pulses
- Each LiDAR scan provides a collection of range measurements $\{r_1, \dots, r_B\}$ over a set of known angles $\{\varphi_1, \dots, \varphi_B\}$ where B is the number of rays



Point Cloud Position

- Given robot pose (x, y, θ) and a range measurement (r, ϕ) , the location (x_p, y_p) of the observed point can be found by polar to Cartesian transformation:



$$x_p = x + r \cos(\phi + \theta)$$

$$y_p = y + r \sin(\phi + \theta)$$

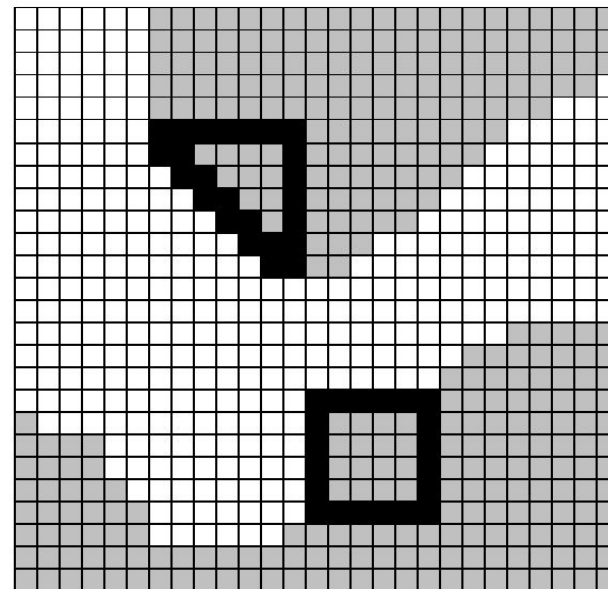
1. Visit the wiki page of the workshop
2. Navigate to the [Localization and Mapping](#) page
3. Run [Demo 1](#)

Point Cloud Representation

- Pros:
 - Does not require ray-tracing
 - Light-weight computation
- Cons:
 - Memory inefficient: map size grows with the number of observations
 - Not suitable for collision avoidance

Occupancy Grid Mapping

- One of the simplest yet widely used map representations
- The environment is divided into a regular grid of N cells
- The value of each cell indicates its occupancy (1=occupied and 0=free)



Log-odds Mapping

- We represent the map \mathbf{m} as a vector where an element m_i is a cell
- Due to the measurement noise and the error in robot location, we need to maintain a probability distribution over the map as $P(\mathbf{m}|\mathbf{x}_{1:T}, \mathbf{z}_{1:T})$

- For simplicity, we assume that map cells are independent:

$$P(\mathbf{m}|\mathbf{x}_{1:T}, \mathbf{z}_{1:T}) = \prod_{i=1}^N P(m_i|\mathbf{x}_{1:T}, \mathbf{z}_{1:T})$$

- Therefore, we can compute occupancy probability for each cell m_i separately

Log-odds Mapping

- Given robot trajectory $\mathbf{x}_{1:T}$ and observations $\mathbf{z}_{1:T}$, let $\gamma_{i,T}$ be the probability that m_i is occupied, i.e.:

$$\gamma_{i,T} = P(m_i = 1 | \mathbf{x}_{1:T}, \mathbf{z}_{1:T})$$

$$\begin{aligned} \text{Bayes Rule} \quad & \frac{1}{\eta_T} P(\mathbf{z}_T | m_i = 1, \mathbf{x}_T) P(m_i = 1 | \mathbf{x}_{1:T-1}, \mathbf{z}_{1:T-1}) = \frac{1}{\eta_T} P(\mathbf{z}_T | m_i = 1, \mathbf{x}_T) \gamma_{i,T-1} \end{aligned}$$

$$1 - \gamma_{i,T} = P(m_i = 0 | \mathbf{x}_{1:T}, \mathbf{z}_{1:T})$$

$$\begin{aligned} \text{Bayes Rule} \quad & \frac{1}{\eta_T} P(\mathbf{z}_T | m_i = 0, \mathbf{x}_T) (1 - \gamma_{i,T-1}) \end{aligned}$$

$$\Rightarrow \frac{\gamma_{i,T}}{1 - \gamma_{i,T}} = \frac{P(\mathbf{z}_T | m_i = 1, \mathbf{x}_T)}{P(\mathbf{z}_T | m_i = 0, \mathbf{x}_T)} \frac{\gamma_{i,T-1}}{1 - \gamma_{i,T-1}}$$

Odds Ratio

- Using Bayes rule once more, we have:

$$\frac{P(\mathbf{z}_T | m_i = 1, \mathbf{x}_T)}{P(\mathbf{z}_T | m_i = 0, \mathbf{x}_T)} = \underbrace{\frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)}}_{\text{inverse observation model}} \underbrace{\frac{P(m_i = 0)}{P(m_i = 1)}}_{\text{map prior}}$$

- The inverse observation model simply encodes our trust in the sensors:

$$\frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)} = \begin{cases} \text{TP/FP} & \mathbf{z}_T \text{ indicates occupied} \\ \text{FN/TN} & \mathbf{z}_T \text{ indicates free} \end{cases}$$

Inverse Observation Model

- A “good” sensor has high TP and TN, and small FP and FN

$$\frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)} = \begin{cases} \text{TP/FP} & \mathbf{z}_T \text{ indicates occupied} \\ \text{FN/TN} & \mathbf{z}_T \text{ indicates free} \end{cases}$$

$$\frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)} = \begin{cases} \left\{ \begin{array}{l} 10 \\ 0.1 \end{array} \right\} & \text{Good Sensor} \\ \left\{ \begin{array}{l} 0.4 \\ 0.4 \end{array} \right\} & \text{Bad Sensor} \\ \left\{ \begin{array}{l} 1 \\ 1 \end{array} \right\} & \text{Bad Sensor} \\ \left\{ \begin{array}{l} 10 \\ 10 \end{array} \right\} & \text{Bad Sensor} \end{cases}$$

\mathbf{z}_T indicates occupied
 \mathbf{z}_T indicates free

Log-odds Mapping

- Going back to the odds ratio formula:

$$\frac{\gamma_{i,T}}{1 - \gamma_{i,T}} = \frac{P(\mathbf{z}_T | m_i = 1, \mathbf{x}_T)}{P(\mathbf{z}_T | m_i = 0, \mathbf{x}_T)} \frac{\gamma_{i,T-1}}{1 - \gamma_{i,T-1}} = \frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)} \frac{P(m_i = 0)}{P(m_i = 1)} \frac{\gamma_{i,T-1}}{1 - \gamma_{i,T-1}}$$

- Taking log of both sides yields:

$$\underbrace{\log\left(\frac{\gamma_{i,T}}{1 - \gamma_{i,T}}\right)}_{\lambda_{i,T}} = \underbrace{\log\left(\frac{P(m_i = 1 | \mathbf{z}_T, \mathbf{x}_T)}{P(m_i = 0 | \mathbf{z}_T, \mathbf{x}_T)}\right)}_{\Delta\lambda_{i,T}(\mathbf{z}_T)} + \underbrace{\log\left(\frac{\gamma_{i,T-1}}{1 - \gamma_{i,T-1}}\right)}_{\lambda_{i,T-1}} - \underbrace{\log\left(\frac{P(m_i = 1)}{P(m_i = 0)}\right)}_{\lambda_{i,0}}$$

- Therefore, the map update can be expressed as accumulating log of odds ratios over time; hence the name log-odds mapping:

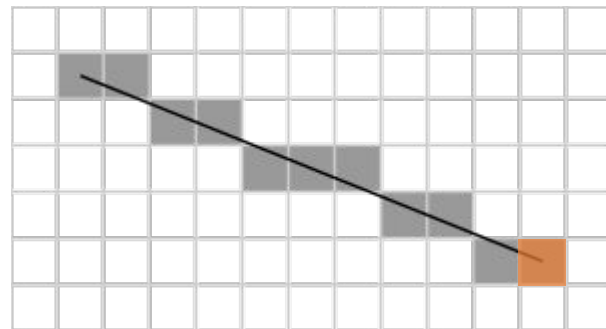
$$\boxed{\lambda_{i,T} = \Delta\lambda_{i,T}(\mathbf{z}_T) + \lambda_{i,T-1} - \lambda_{i,0}}$$

Log-odds Mapping

- In order to recover cell occupancy probability from log-odds value, we use the logistic sigmoid function:

$$P(m_i = 1 | \mathbf{x}_{1:T}, \mathbf{z}_{1:T}) = \gamma_{i,T} = \sigma(\lambda_{i,T}) = \frac{\exp(\lambda_{i,T})}{1 + \exp(\lambda_{i,T})}$$

- For occupancy grid mapping using LiDAR, we need to determine the cells visited by each ray
- This process is called line rasterization



Log-odds Mapping (Summary)

- Log-odds occupancy grid mapping is equivalent to maintaining a grid of log-odds values $\lambda_{i,T}$ for each cell
- There is a one-to-one relationship between log-odds value $\lambda_{i,T}$ and occupancy probability $\gamma_{i,T}$, namely the logistic sigmoid function
- Upon each LiDAR measurement \mathbf{z}_T , follow the below steps:
 - For each ray, find the set of visited cells using line rasterization
 - Use the log-odds formula in order to update the map:
 - If cell i is observed free:

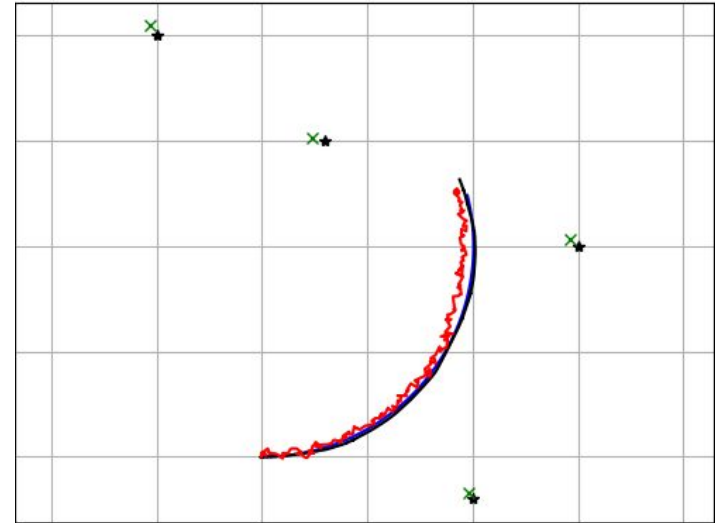
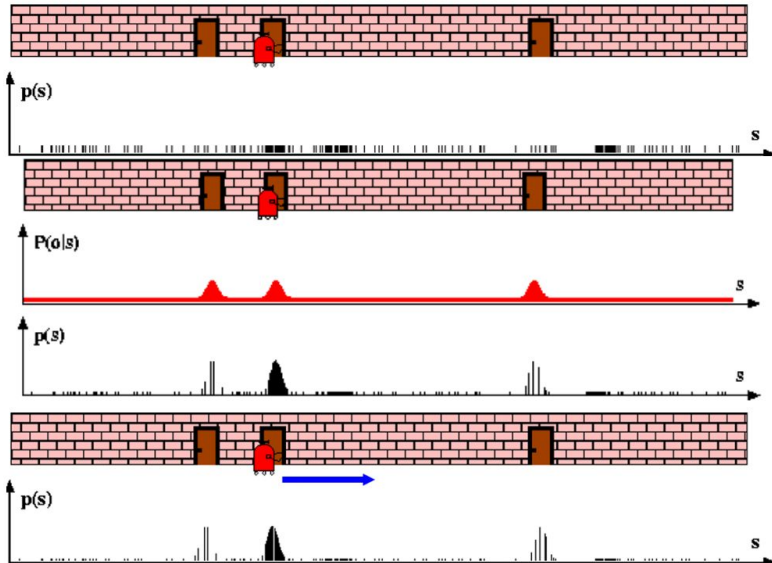
$$\lambda_{i,T} = \Delta\lambda_{i,T}(\text{free}) + \lambda_{i,T-1} - \lambda_{i,0}$$

- If cell i is observed occupied:

$$\lambda_{i,T} = \Delta\lambda_{i,T}(\text{occupied}) + \lambda_{i,T-1} - \lambda_{i,0}$$

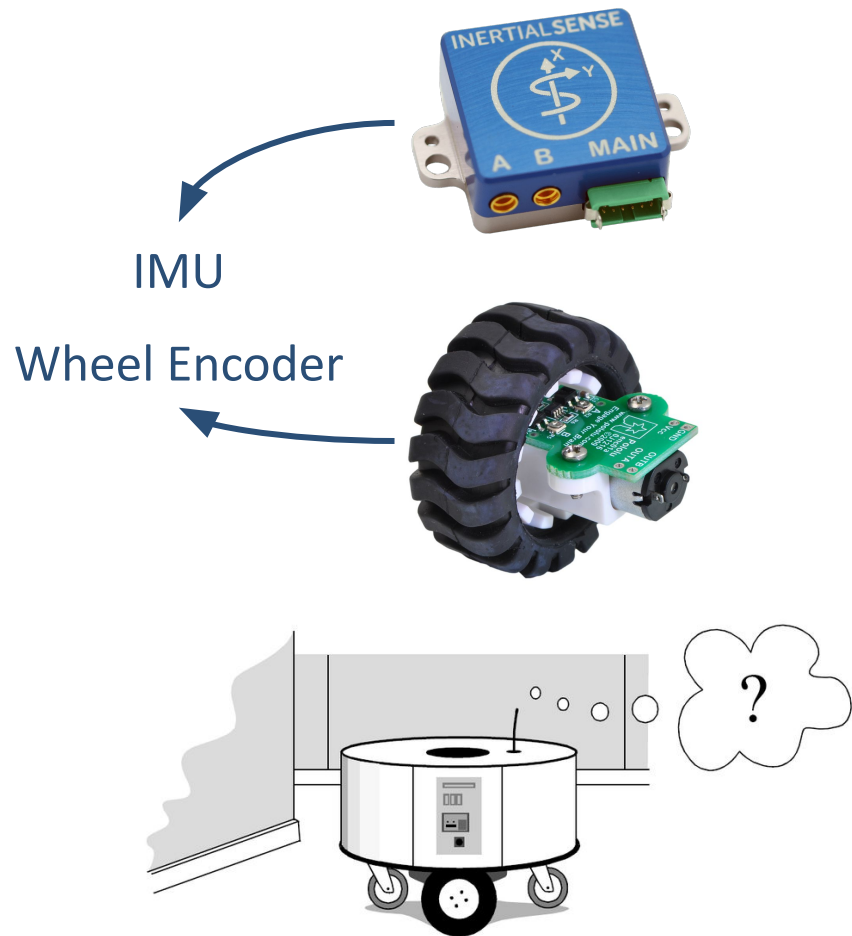
1. Visit the wiki page of the workshop
2. Navigate to the [Localization and Mapping](#) page
3. Run [Demo 2](#)

Localization



Robot Localization

- Odometry: The use of motion sensor data to estimate change of robot pose
- Motion sensors usually have small measurement noise:
 - Accurate between t and $t+1$
 - Error accumulates during long time horizon
- Therefore, simply relying on odometry leads to large error in robot pose estimation



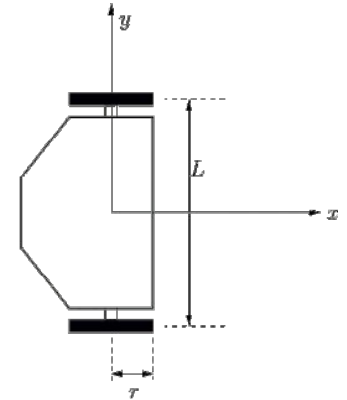
Differential-drive Model

- Consider the differential-drive motion model
- The velocity of left and right wheels control the motion of the robot:
 - Sum of wheels velocities: Robot linear velocity
 - Difference of wheels velocity: Robot angular velocity
- Therefore, the control inputs are linear velocity v and angular velocity ω

$$x_{t+1} = x_t + \tau v_t \cos(\theta_t)$$

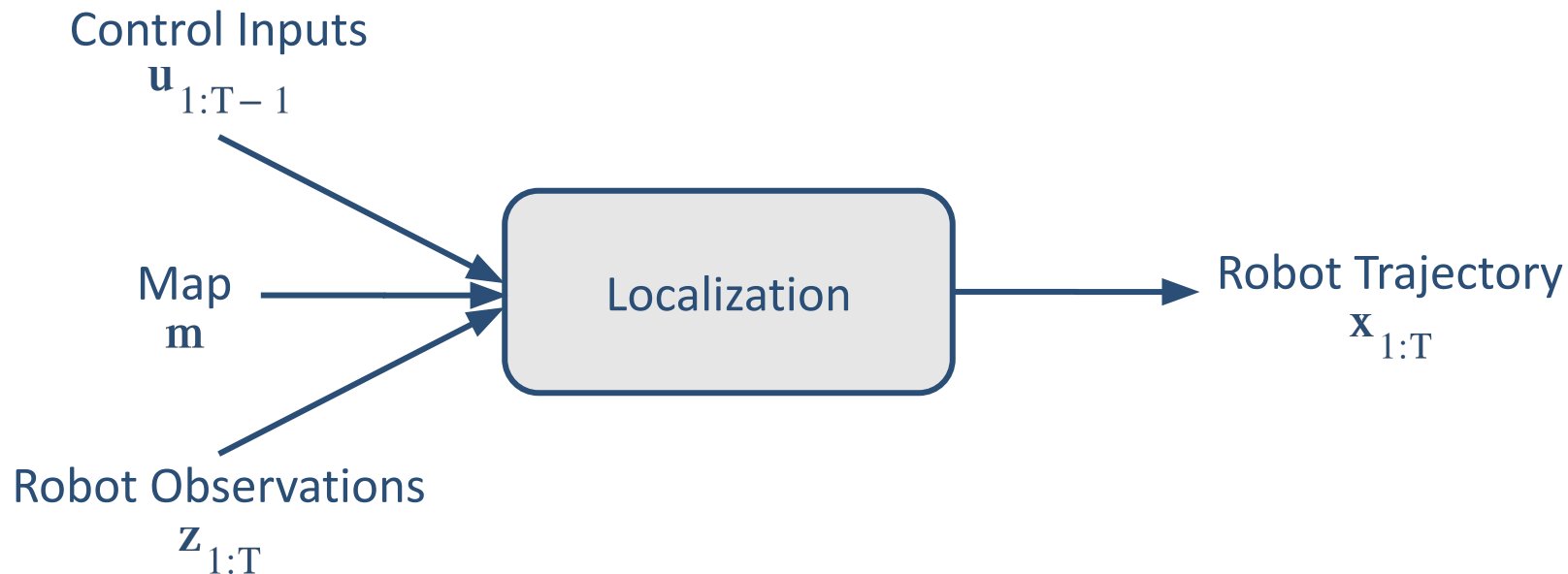
$$y_{t+1} = y_t + \tau v_t \sin(\theta_t)$$

$$\theta_{t+1} = \theta_t + \tau \omega_t$$

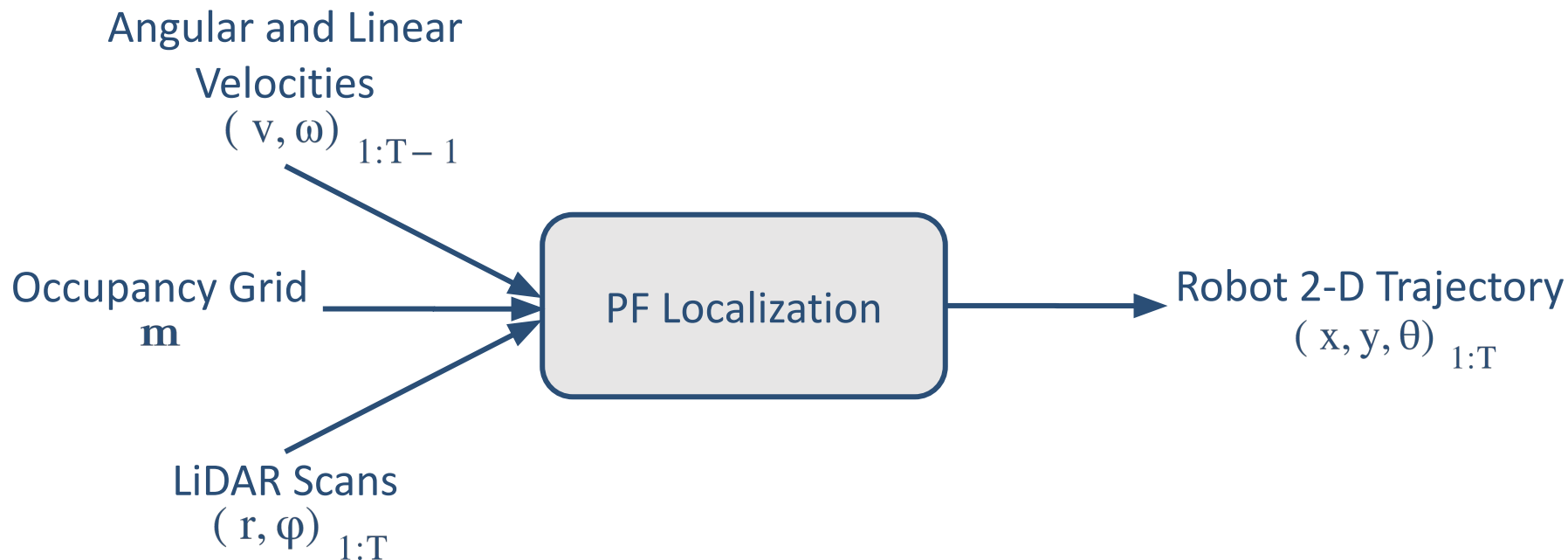


1. Visit the wiki page of the workshop
2. Navigate to the [Localization and Mapping](#) page
3. Run [Demo 3](#)

Problem Formulation



Problem Formulation



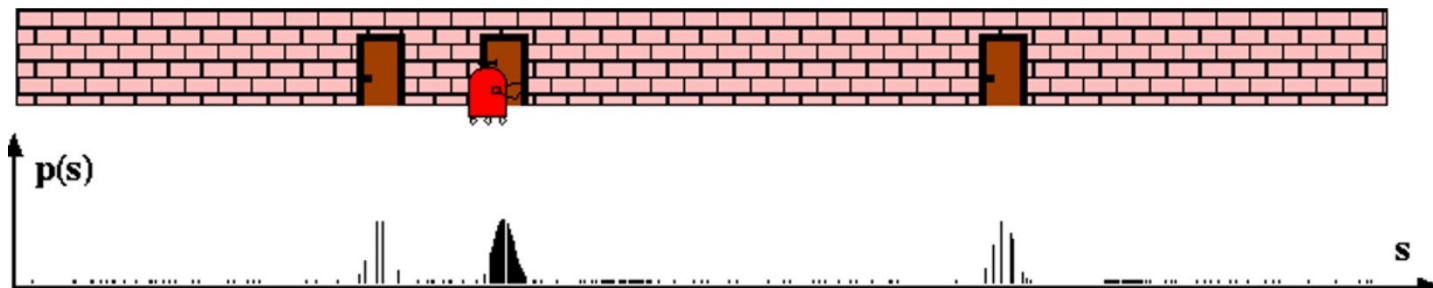
Particle Filtering

- We maintain a probability distribution over robot pose via a set of hypotheses (particles) with pose $\{\boldsymbol{\mu}^{(k)}\}_k$ and their respective weight $\{\alpha^{(k)}\}_k$:

$$P(\mathbf{x}_T | \mathbf{z}_{1:T}, \mathbf{u}_{1:T-1}) = \begin{cases} \alpha^{(k)} & \mathbf{x}_T = \boldsymbol{\mu}^{(k)} \\ 0 & \text{otherwise} \end{cases}$$

- The weights $\alpha^{(k)}$ represent the probability of each particle, hence:

$$\sum_{k=1}^N \alpha^{(k)} = 1$$



Prediction Step

- Given a particle set $\{(\boldsymbol{\mu}_{T|T}^{(k)}, \alpha_{T|T}^{(k)})\}_k$ at time T , we are interested to obtain the predicted particle set $\{(\boldsymbol{\mu}_{T+1|T}^{(k)}, \alpha_{T+1|T}^{(k)})\}_k$ after applying the control \mathbf{u}_T
- We can simply plug each $\boldsymbol{\mu}_{T|T}^{(k)}$ to the motion model
- However, we also want to model the motion sensor noise
- In other words, we want our particles be diverse, since each of them represent a hypothesis over robot pose
- Therefore, we should add a small noise to our motion model when we apply the motion model to each particle

Prediction Step

- For a differential-drive motion model, each particle represents a 2-D robot pose:

$$\mathbf{\mu}_{T|T}^{(k)} = \left(\mu_{x, T|T}^{(k)}, \mu_{y, T|T}^{(k)}, \mu_{\theta, T|T}^{(k)} \right)$$

- Prediction step:

$$\mu_{x, T+1|T}^{(k)} = \mu_{x, T|T}^{(k)} + \tau \left(v_T + n_v^{(k)} \right) \cos\left(\mu_{\theta, T|T}^{(k)}\right)$$

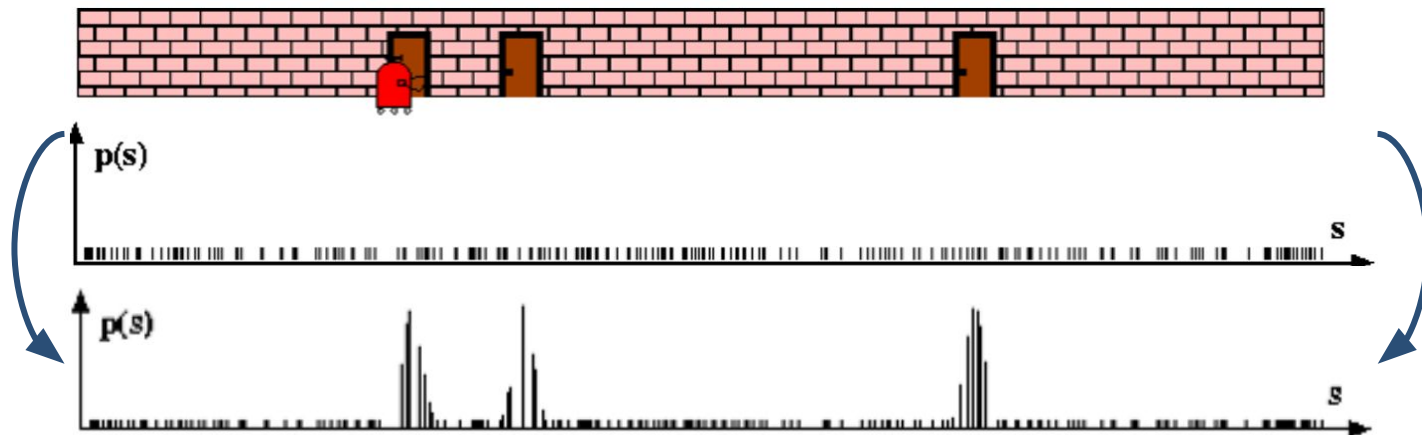
$$\mu_{y, T+1|T}^{(k)} = \mu_{y, T|T}^{(k)} + \tau \left(v_T + n_v^{(k)} \right) \sin\left(\mu_{\theta, T|T}^{(k)}\right)$$

$$\mu_{\theta, T+1|T}^{(k)} = \mu_{\theta, T|T}^{(k)} + \tau \left(\omega_T + n_{\omega}^{(k)} \right)$$

- Note that the particle weights $\{\alpha^{(k)}\}_k$ do not change during the prediction step
- We denote the predicted particle set as $\{(\mathbf{\mu}_{T+1|T}^{(k)}, \alpha_{T+1|T}^{(k)})\}_k$

Update Step

- Given a predicted particle set $\{(\boldsymbol{\mu}_{T+1|T}^{(k)}, \alpha_{T+1|T}^{(k)})\}_k$, we want to update the weight of each particle according to the sensor observation \mathbf{z}_{T+1}
- Here, we are interested to know if a hypothesis $\boldsymbol{\mu}_{T+1|T}^{(k)}$ agrees with an observation \mathbf{z}_{T+1}
- If it agrees (disagrees), increase (decrease) the particle's weight $\alpha_{T+1|T}^{(k)}$



Update Step

- In order to incorporate a new observation \mathbf{z}_{T+1} to the particle set, we need to use the Bayes rule:

$$\begin{aligned} P(\mathbf{x}_{T+1} = \mathbf{\mu}_{T+1|T}^{(k)} | \mathbf{z}_{1:T+1}, \mathbf{u}_{1:T}) &= \alpha_{T+1|T+1}^{(k)} \\ &\stackrel{\text{Bayes Rule}}{=} \frac{1}{\eta_{T+1}} P(\mathbf{z}_{T+1} | \mathbf{x}_{T+1} = \mathbf{\mu}_{T+1|T}^{(k)}) P(\mathbf{x}_{T+1} = \mathbf{\mu}_{T+1|T}^{(k)} | \mathbf{z}_{1:T}, \mathbf{u}_{1:T}) \\ &= \frac{1}{\eta_{T+1}} \underbrace{P(\mathbf{z}_{T+1} | \mathbf{x}_{T+1} = \mathbf{\mu}_{T+1|T}^{(k)})}_{\text{forward observation model}} \alpha_{T+1|T}^{(k)} \end{aligned}$$

- We still need expressions for the forward observation model and the denominator η_{T+1}

Laser Correlation Model

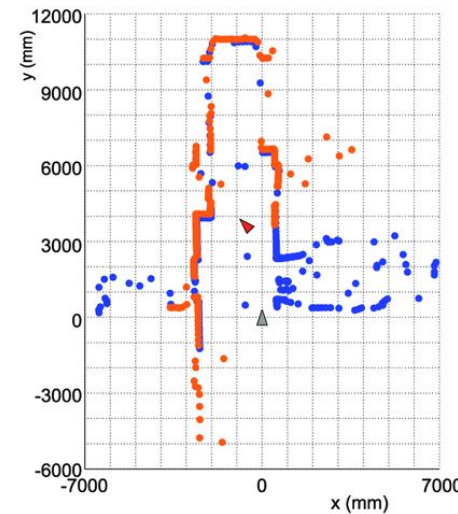
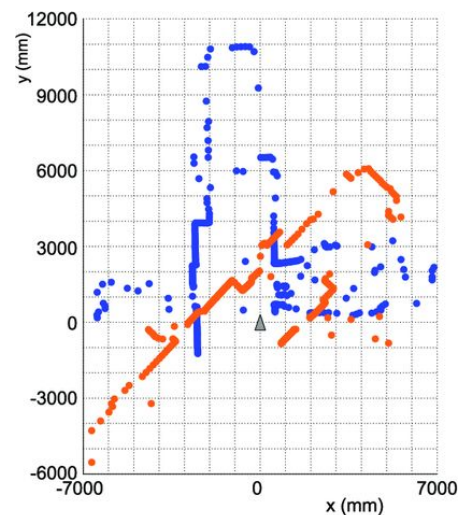
- We define the forward observation model as a function of the correlation between map \mathbf{m} and LiDAR scan \mathbf{z} obtained from robot pose \mathbf{x} :

$$P(\mathbf{z}|\mathbf{x}) \propto \exp(\text{corr}(r(\mathbf{z}, \mathbf{x}), \mathbf{m}))$$

where $r(\mathbf{z}, \mathbf{x})$ is the LiDAR point cloud from pose \mathbf{x} and:

$$\text{corr}(\mathbf{r}, \mathbf{m}) = \sum_i \mathbb{1} \{r_i = m_i\}$$

- The correlation is large if $r(\mathbf{z}, \mathbf{x})$ and the map \mathbf{m} agree



- Going back to the weight update:

$$\alpha_{T+1|T+1}^{(k)} = \frac{1}{\eta_{T+1}} \exp\left(\text{corr}\left(r\left(\mathbf{z}_{T+1}, \boldsymbol{\mu}_{T+1|T}^{(k)}\right), \mathbf{m}\right)\right) \alpha_{T+1|T}^{(k)}$$

- Furthermore, the denominator η_{T+1} is constant among all particles
- Therefore, we can eliminate it by normalization:

for $k = 1 \dots N$:

$$\alpha_{T+1|T+1}^{(k)} \leftarrow \exp\left(\text{corr}\left(r\left(\mathbf{z}_{T+1}, \boldsymbol{\mu}_{T+1|T}^{(k)}\right), \mathbf{m}\right)\right) \alpha_{T+1|T}^{(k)}$$

$$\eta_{T+1} = \sum_{k=1}^N \alpha_{T+1|T+1}^{(k)}$$

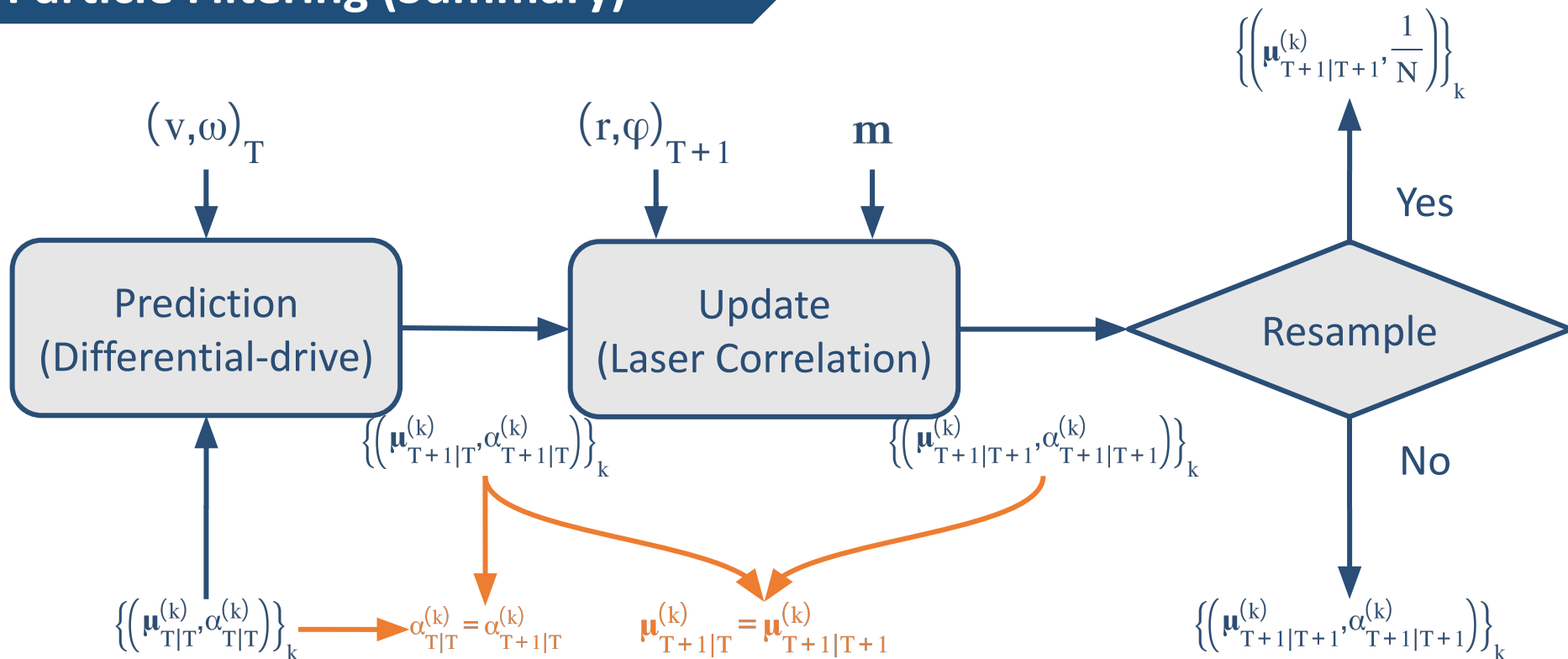
for $k = 1 \dots N$:

$$\alpha_{T+1|T+1}^{(k)} \leftarrow \frac{\alpha_{T+1|T+1}^{(k)}}{\eta_{T+1}}$$

- Particle depletion: a situation in which most of the particle weights become close to zero
- This means that only a minority of the particles can be good candidates for the true robot pose while the rest are irrelevant
- Resampling is a procedure to avoid particle depletion
- Given a particle set $\{(\boldsymbol{\mu}_{T|T}^{(k)}, \alpha_{T|T}^{(k)})\}_k$, create a new particle set with equal weights $\alpha_{T|T}^{(k)} = 1/N$ by adding many particles to the locations with high weights and few particles to the locations with small weights
- Resampling can be triggered whenever the effective number of particles N_{eff} is less than a threshold:

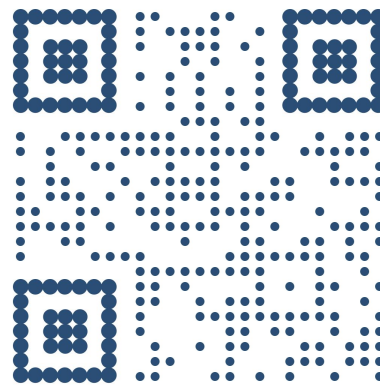
$$N_{\text{eff}} = \frac{1}{\sum_k \left(\alpha_{T|T}^{(k)}\right)^2} < N_{\text{thr}}$$

Particle Filtering (Summary)



1. Visit the wiki page of the workshop
2. Navigate to the [Localization and Mapping](#) page
3. Run [Demo 4](#)

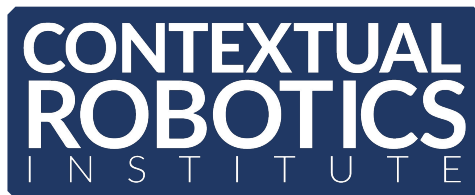
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