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Problem 1

- a) Sample \mathbf{x}_i from \tilde{q} . Sample u_i uniformly from $[0, \tilde{q}(\mathbf{x}_i)]$. Accept \mathbf{x}_i as sample if $u_i \leq p(\mathbf{x}_i)$. Else, reject and restart.
- b) True. The previous samples has no influence on the current one.
- c) $w_n = \frac{p(\mathbf{x}_n)}{q(\mathbf{x}_n)}$
- d) $\alpha(\mathbf{x}_{t+1}, \mathbf{x}_t) = \min\left(1, \frac{\tilde{p}(\mathbf{x}_{t+1})q(\mathbf{x}_t)}{\tilde{p}(\mathbf{x}_t)q(\mathbf{x}_{t+1})}\right)$
- e) Even though the proposal distribution of x_{t+1} is independent of x_t , the accept probability is not.
- f) x_1 , x_1 , x_3 , x_4 , x_4
- g) The case where the dimensions are independent is trivial. Now, it is harder to find a tight bounding proposal distribution in a high dimensional space. Since the discussed methods all use a proposal distribution, in general they will not work well.

Problem 2

Using Bishop's 2.140 - 2.142,
$$p(\mu|x,\tau) = \mathcal{N}\left(\mu|\frac{\tau^{-1}}{s_0+\tau^{-1}}\mu_0 + \frac{s_0}{s_0+\tau^{-1}}x, \left(\frac{1}{s_0} + \frac{1}{\tau^{-1}}\right)^{-1}\right).$$
 Using Bishop's 2.149 - 2.151,
$$p(\tau|x,\mu) = \operatorname{Gamma}\left(\tau|a + \frac{1}{2}, b + \frac{(x-\mu)^2}{2}\right).$$

Problem 4

- a) $\mathbb{E}[\mathbf{x}] = [\mathbb{E}[x_1], \dots, \mathbb{E}[x_D]]^T = \boldsymbol{\mu}.$
- b) Note that if $i \neq j$, $x_i \perp x_j$, thus $\Sigma_{ij} = 0$. For $\Sigma_{ii} = \mathbb{V}[x_i] = \mu_i (1 \mu_i)$.
- c) $p(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\pi}) = \sum_k \pi_k p(\mathbf{x}|\boldsymbol{\mu}_k)$. By the linearity of the expectation and part a), $\mathbb{E}[\mathbf{x}] = \sum_k \pi_k \boldsymbol{\mu}_k$.
- d) $\mathcal{L} = \log p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\pi}) = \sum_n \log p(\mathbf{x}_n|\boldsymbol{\mu}, \boldsymbol{\pi}) = \sum_n \log \left\{ \sum_k \pi_k \prod_i \mu_{ki}^{x_{ni}} (1 \mu_{ki})^{(1 x_{ni})} \right\}.$
- e) Since there is a summation inside the logarithm, there is no closed form solution.
- f) $\log p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\pi}) = \sum_{n} \log p(\mathbf{x}_{n}, \mathbf{z}_{n} | \boldsymbol{\mu}, \boldsymbol{\pi}) = \sum_{n} \sum_{k} z_{nk} \{ \log \pi_{k} + \sum_{i} x_{ni} \log \mu_{ki} + (1 x_{ni}) \log (1 \mu_{ki}) \}.$
- g) See below μ and π are vectors in the PGM.

$$\mathsf{h)} \ \mathcal{L} = \sum_{n} \sum_{\mathbf{z}_{n}} q_{n}(\mathbf{z}_{n}) \log p(\mathbf{x}_{n}, \mathbf{z}_{n} | \boldsymbol{\mu}, \boldsymbol{\pi}) - \sum_{n} \sum_{\mathbf{z}_{n}} q_{n}(\mathbf{z}_{n}) \log q_{n}(\mathbf{z}_{n})$$

$$= \sum_{n} \sum_{\mathbf{z}_{n}} q_{n}(\mathbf{z}_{n}) \sum_{l} z_{nk} \left\{ \log \pi_{k} + \sum_{l} x_{ni} \log \mu_{ki} + (1 - x_{ni}) \log (1 - \mu_{ki}) \right\} - \sum_{n} \sum_{\mathbf{z}_{n}} q_{n}(\mathbf{z}_{n}) \log q_{n}(\mathbf{z}_{n})$$

- i) $\tilde{\mathcal{L}} = \mathcal{L} + \alpha \left(\sum_{k} \pi_{k} 1 \right) + \sum_{n} \lambda_{n} \left(\sum_{\mathbf{z}_{n}} q_{n}(\mathbf{z}_{n}) 1 \right)$
- j) Let \mathbf{e}_k denote the k-th column of the identity matrix \mathbb{I}_K .

$$\frac{\partial \tilde{\mathcal{L}}}{\partial q_n(\mathbf{e}_k)} = \log \pi_k + \log \prod_i \mu_{ki}^{x_{ni}} (1 - \mu_{ki})^{(1 - x_{ni})} - \log q_n(\mathbf{e}_k) - 1 + \lambda_n = 0$$
$$q_n(\mathbf{e}_k) = \exp(\lambda_n - 1) \pi_k \prod_i \mu_{ki}^{x_{ni}} (1 - \mu_{ki})^{(1 - x_{ni})}$$

Note that $\sum_k q_n(\mathbf{e}_k) = 1$, which implies that $\exp(1 - \lambda_n) = \sum_k \pi_k \prod_i \mu_{ki}^{x_{ni}} (1 - \mu_{ki})^{(1 - x_{ni})}$.

$$q_n(\mathbf{e}_k) = \frac{\pi_k \prod_i \mu_{ki}^{x_{ni}} (1 - \mu_{ki})^{(1 - x_{ni})}}{\sum_k \pi_k \prod_i \mu_{ki}^{x_{ni}} (1 - \mu_{ki})^{(1 - x_{ni})}}.$$

This, of course, represents the posterior distribution $p(\mathbf{z}_n|\mathbf{x}_n, \boldsymbol{\mu}, \boldsymbol{\pi})$.

k) Recall that $\sum_k \pi_k = 1$.

$$\frac{\partial \tilde{\mathcal{L}}}{\partial \pi_k} = \frac{1}{\pi_k} \sum_n q_n(\mathbf{e}_k) + \alpha = 0 \Rightarrow -\alpha \pi_k = \sum_n q_n(\mathbf{e}_k) \Rightarrow \pi_k = \frac{1}{-\alpha} \sum_n q_n(\mathbf{e}_k)$$

However, $-\alpha = \sum_k \sum_n q_n(\mathbf{e}_k)$. This implies, $\pi_k = \frac{\sum_n q_n(\mathbf{e}_k)}{\sum_k \sum_n q_n(\mathbf{e}_k)}$.

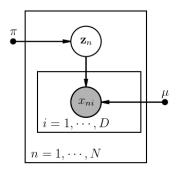


Figure 1: Exercise g)

Problem 5

Consider the discrete-time stochastic process $\{x_t\}_{t>0}$ given by $p(x_t)=0.5\chi_0+0.25\chi_{-1}+0.25\chi_1$ It is easy to see that $\mathbb{E}[x_t]=0$ and $\mathbb{V}[x_t]=\frac{1}{2}$. Note that $x_t \perp \!\!\! \perp x_{t'}$ for $t\neq t'$. Clearly, $z^{(r)}=\sum_{t=1}^r x_t$.

$$\frac{r}{2} = \sum_{t=1}^{r} \mathbb{V}\left[x_{t}\right] = \mathbb{V}\left[\sum_{t=1}^{r} x_{t}\right] = \mathbb{V}[z^{(r)}] = \mathbb{E}\left[\left(z^{(r)}\right)^{2}\right] - \left(\mathbb{E}[z^{(p)}]\right)^{2} = \mathbb{E}\left[\left(z^{(r)}\right)^{2}\right]$$