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Problem 1.

We are trying hire the best candidate. We have n candidates $\pi_1, \pi_2, ..., \pi_n$ but we can only hire one. Our interview process is as follows: We interview and reject the first $\pi_1, ..., \pi_{n/e}$ candidates, and then hire the first candidate from the $\pi_{n/e+1}, ..., \pi_n$ order that outperforms all of the first n/e candidates.

(a) For an index i > n/e, let E_i be the event that π_i is the best candidate. Find $P(\pi_i$ is hired $|E_i|$. In this scenario, π_i is hired if no one is hired from the candidates of $\pi_1, ..., \pi_{i-1}$. Let π_k be the best candidate in this pool. Since we are conditioning on E_i , we can assume that $\pi_i > \pi_k$. The first question is where can π_k go? We can break up the pool as follows:

$$\underbrace{\pi_1, \pi_2, ..., \pi_{n/e}, \pi_{n/e+1}, ... \pi_{i-1}}_{\text{always rejected}}$$
 always hired

Since π_k is the best candidate in this pool, the only place it can go is in the first n/e spots. Otherwise, π_k will be hired instead of π_i . Since we have i-1 spots, total, we can express this as:

$$P(\pi_k \text{ not hired}) = \frac{n/e}{i-1}$$

The second question is how many choices of π_k we can have. The only condition we know is the $\pi_i > \pi_k$, but whether π_k is 2nd best, 3rd best, we don't really know. But, this is also easy to figure out, because we have a limited pool of candidates. The way to think of it is π_k is put into the first i-1 pool. But the candidates better than π_k that is not π_i must go after π_i , because otherwise they would come before, thereby replacing π_k . Thus we have the following:

of choices for
$$\pi_k = n - (i - 1)$$

Combine the two parts to get the probability of hiring π_i :

$$P(\pi_i \text{ is hired } | E_i) = \left(\frac{n/e}{i-1}\right)(n-(i-1))$$

(b) If π^* is the best candidate, find an approximation for $P(\pi^*$ is hired).

We can express this is a summation over all choices of i:

$$P(\pi^* \text{ is hired}) = \sum_{i=1}^n P(\pi_i \text{ is hired}|E_i)P(E_i).$$

Since π_i won't be hired if it falls into the first n/e candidates, we can reduce the expression:

$$P(\pi^* \text{ is hired}) = \sum_{i=n/e+1}^{n} P(\pi_i \text{ is hired}|E_i)P(E_i).$$

Expanding this, and then letting j = i - 1 gets us:

$$P(\pi^* \text{ is hired}) = \sum_{i=n/e}^n P(\pi_i \text{ is hired}|E_i)P(E_i)$$

$$= \sum_{i=n/e+1}^n \left(\frac{(n/e)(n-(i-1))}{(i-1)}\right) \left(\frac{1}{n}\right)$$

$$= \sum_{j=n/e}^n \left(\frac{(n/e)(n-j)}{j}\right) \left(\frac{1}{n}\right)$$

$$= \sum_{j=n/e}^n \frac{(n-j)}{ej}$$

$$= \sum_{j=n/e}^n \frac{n}{ej} - \frac{j}{ej}$$

$$= \sum_{j=n/e}^n \frac{n}{ej} - \sum_{j=n/e}^n \frac{1}{e}$$

$$= \sum_{j=n/e}^n \frac{n}{ej} - \left((n-n/e+1)\frac{1}{e}\right)$$

We still have to deal with the summation term. Here, we can use the approximation $\int_a^{b+1} \frac{dx}{x} \le \sum_{i=a}^b \frac{1}{i} \le \int_{a-1}^b \frac{dx}{x}$. Then, we can express the summation as follows:

$$\int_{n/e}^{n+1} \frac{dx}{x} \le \sum_{i=n/e}^{n} \frac{n}{ej} \qquad \le \int_{n/e-1}^{n} \frac{dx}{x} \tag{1}$$

$$\left((n - n/e + 1) \frac{n}{e} \right) \le \sum_{j=n/e}^{n} \frac{n}{ej} \le \left((n - n/e + 1) \frac{n}{e} \right) \tag{2}$$

Thus, our final probability is:

$$P(\pi^* \text{ is hired}) = \left((n - n/e + 1) \frac{n - 1}{e} \right)$$

Problem 2.

Consider the following linear program

$$\min 2x + 5y - 3z$$
subject to
$$2x - y + 2z \ge 3$$

$$x - y - 2z \ge -1$$

$$-x + y + 5z \ge 1$$

$$x + y \ge 2$$

$$x, y, z \ge 0$$

(a) Find the dual of the linear program.

The dual formulation can be expressed from the original linear program with the following transpose:

Linear	Dual
$minc^T x$	$\max b^T x$
$Ax \ge b$	$A^T x \le c$
$x \ge 0$	$x \ge 0$

As such, the values for A, b, c^T can be pulled from the original linear program:

$$A = \begin{bmatrix} 2 & -1 & 2 \\ 1 & -1 & -2 \\ -1 & 1 & 5 \\ 2 & 1 & 0 \end{bmatrix} \quad b = \begin{bmatrix} 3 \\ -1 \\ 1 \\ 2 \end{bmatrix} \quad c^T = \begin{bmatrix} 2 & 5 & -3 \end{bmatrix}$$

Applying the transpose, and formulated as a dual, we get this:

$$\max 3w - x + y + 2z$$
subject to
$$2w + x - y + z \le 2$$

$$-w - x + y + z \le 5$$

$$2w - 2x + 5y \le -3$$

$$w, x, y, z \ge 0$$

(b) At (x, y, z) = (2, 0, 3/2), the linear program has objective value -1/2. Show this is the optimal solution.

We can show this via linear combination. First, we can organize the program as matrix vectors:

$$v_1 = 2x - y + 2z \ge 3$$

 $v_2 = x - y - 2z \ge -1$
 $v_3 = -x + y + 5z \ge 1$
 $v_4 = x + y \ge 2$

Then, we need to find a linear combination B such that expressing 2x + 5y - 3z in this basis is solvable. (Note, this was reversed engineered from the objective value -1/2. But for the proof, we can *magically* find such a basis):

$$b_1 = v_2 + v_4 \qquad = 2x - 2z \ge 1 \tag{3}$$

$$b_2 = v_3 - v_1 = -3x + 2y + 3z \ge -2 \tag{4}$$

$$b_3 = v_4 - 2v_1 = -3x + 3y - 4z \ge -4 \tag{5}$$

With a basis $\mathcal{B} = \text{span}\left\{\begin{bmatrix} 2\\0\\-2 \end{bmatrix}, \begin{bmatrix} -3\\2\\3\\-4 \end{bmatrix}\right\}$ we would like to express $\begin{bmatrix} 2\\5\\-3 \end{bmatrix}$ in terms of this basis.

This can be thought of as the following:

$$\begin{bmatrix} 2 & -3 & -3 \\ 0 & 2 & 3 \\ -2 & 3 & -4 \end{bmatrix} \begin{bmatrix} a \\ b \\ c \end{bmatrix} = \begin{bmatrix} 2 \\ 5 \\ -3 \end{bmatrix}$$

Solving for a, b, c nets us $a = \frac{65}{14}$, $b = \frac{16}{7}$, $c = \frac{1}{7}$. Plugging into our basis vectors, and we get the following:

$$ab_1 + bb_2 + cb_3$$

$$\frac{65}{14}(2x - 2z) + \frac{16}{7}(-3x + 2y + 3z) + \frac{1}{7}(-3x + 3y - 4z) \ge \frac{65}{14}(1) + \frac{16}{7}(-2) + \frac{1}{7}(-4)$$

$$\frac{1}{14}[65(2x - 2z) + 32(-3x + 2y + 3z) + 2(-3x + 3y - 4z)] \ge \frac{1}{14}[65(1) + 32(-2) + 2(-4)]$$

$$\frac{1}{14}[130x - 130z - 96x + 64y + 96z + 6x + 6y - 8z] \ge \frac{1}{14}[65(1) + 32(-2) + 2(-4)]$$

$$\frac{1}{14}[28x + 70y - 42z] \ge \frac{1}{14}[-7]$$

$$2x + 5y - 3z \ge -1/2$$

Since the minimum value for 2x + 5y - 3z is -1/2, it must be the optimal solution.

Problem 3.

We are given two $n \times d$ matrices A and B where $n \gg d$. Let $a_1,...,a_n$ be the column vectors corresponding to the rows of A, and similarly for $b_1,...,b_n$. We would like to compute $P = A^T B$ quickly. In this case, we will pick a $r \in 1,2,...,n$ with probability $p_r = \frac{\|a_r\|\cdot\|b_r\|}{\sum_{j=1}^n \|a_j\|\cdot\|b_j\|}$. Our approximation is then $\hat{P} = \frac{1}{p_r} a_r b_r^T$.

(a) Our error can be expressed in terms of $||P - \hat{P}||_F$. Show that

$$\mathbb{E}\left[\|P - \hat{P}\|_F^2\right] \le \left(\sum_{i=1}^n \|a_i\| \cdot \|b_i\|\right)^2$$

The Frobenius norm is the sum of the squares of the entries. $P = A^T B$ is a $d \times d$ matrix:

$$\mathbb{E}\left[\|P - \hat{P}\|_{F}^{2}\right] = \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} (P_{uv} - \hat{P}_{uv})^{2}\right]$$

$$\leq \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} (\hat{P}_{uv})^{2}\right]$$

$$\leq \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} (\frac{1}{p_{r}} ([a_{r}b_{r}^{T}]_{uv}))^{2}\right]$$

$$\leq \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} \left(\frac{\sum_{i=1}^{n} ||a_{i}|| \cdot ||b_{i}||}{||a_{r}|| \cdot ||b_{r}||}\right)^{2} ([a_{r}b_{r}^{T}]_{uv})^{2}\right]$$

$$\leq \mathbb{E}\left[\left(\frac{\sum_{i=1}^{n} ||a_{i}|| \cdot ||b_{i}||}{||a_{r}|| \cdot ||b_{r}||}\right)^{2} \sum_{u=1}^{d} \sum_{v=1}^{d} ([a_{r}b_{r}^{T}]_{uv})^{2}\right]$$

The expression $\sum_{u=1}^{d} \sum_{v=1}^{d} ([a_r b_r^T]_{uv})^2$ can be reduced by examining the matrix $a_r b_r^T$:

$$a_r b_r^T = \begin{bmatrix} a_r^{(1)} b_r^{(1)} & a_r^{(1)} b_r^{(2)} & \dots & a_r^{(1)} b_r^{(d)} \\ a_r^{(2)} b_r^{(1)} & a_r^{(2)} b_r^{(2)} & \dots & a_r^{(2)} b_r^{(d)} \\ \vdots & \vdots & \ddots & \vdots \\ a_r^{(d)} b_r^{(1)} & a_r^{(d)} b_r^{(2)} & \dots & a_r^{(d)} b_r^{(d)} \end{bmatrix}$$

If we square all the entries, and then apply the summation, we get the following:

$$\begin{split} \sum_{u=1}^{d} \sum_{v=1}^{d} ([a_{r}b_{r}^{T}]_{uv})^{2} &= \begin{pmatrix} (a_{r}^{(1)}b_{r}^{(1)})^{2} &+ (a_{r}^{(1)}b_{r}^{(2)})^{2} &+ \dots &+ (a_{r}^{(1)}b_{r}^{(d)})^{2} \\ + (a_{r}^{(2)}b_{r}^{(1)})^{2} &+ (a_{r}^{(2)}b_{r}^{(2)})^{2} &+ \dots &+ (a_{r}^{(2)}b_{r}^{(d)})^{2} \\ + & \dots &+ (a_{r}^{(d)}b_{r}^{(1)})^{2} &+ (a_{r}^{(1)}b_{r}^{(2)})^{2} &+ \dots &+ (a_{r}^{(d)}b_{r}^{(d)})^{2} \end{pmatrix} \\ &= \begin{pmatrix} (a_{r}^{(1)})^{2} & \left[& (b_{r}^{(1)})^{2} &+ (b_{r}^{(2)})^{2} &+ \dots &+ (b_{r}^{(d)})^{2} & \right] \\ + & (a_{r}^{(2)})^{2} & \left[& (b_{r}^{(1)})^{2} &+ (b_{r}^{(2)})^{2} &+ \dots &+ (b_{r}^{(d)})^{2} & \right] \\ + & \dots &+ (a_{r}^{(d)})^{2} & \left[& (b_{r}^{(1)})^{2} &+ (b_{r}^{(2)})^{2} &+ \dots &+ (b_{r}^{(d)})^{2} & \right] \end{pmatrix} \\ &= \|a_{r}\|^{2} \cdot \|b_{r}\|^{2} \\ &= (\|a_{r}\| \cdot \|b_{r}\|)^{2} \end{split}$$

Plug this back into our original problem and we get the following:

$$\mathbb{E}\left[\|P - \hat{P}\|_{F}^{2}\right] \leq \mathbb{E}\left[\left(\frac{\sum_{i=1}^{n} \|a_{i}\| \cdot \|b_{i}\|}{\|a_{r}\| \cdot \|b_{r}\|}\right)^{2} \sum_{u=1}^{d} \sum_{v=1}^{d} ([a_{r}b_{r}^{T}]_{uv})^{2}\right]$$

$$= \mathbb{E}\left[\left(\frac{\sum_{i=1}^{n} \|a_{i}\| \cdot \|b_{i}\|}{\|a_{r}\| \cdot \|b_{r}\|}\right)^{2} (\|a_{r}\| \cdot \|b_{r}\|)^{2}\right]$$

$$= \left(\sum_{i=1}^{n} \|a_{i}\| \cdot \|b_{i}\|\right)^{2} \checkmark$$

(b) Instead of using 1 sample r, we use m i.i.d. samples $r_1, ..., r_m$ and then take the average result. The new estimate is $\hat{P} = \frac{1}{m} \sum_{i=1}^m \frac{1}{p_{r_i}} a_{r_i} b_{r_i}^T$. What happens to the bound of $\mathbb{E} \left[||P - \hat{P}||_F^2 \right]$? We can use the same logic as above to manipulate terms:

$$\mathbb{E}\left[\|P - \hat{P}\|_{F}^{2}\right] = \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} (P_{uv} - \hat{P}_{uv})^{2}\right]$$

$$\leq \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} (\hat{P}_{uv})^{2}\right]$$

$$\leq \mathbb{E}\left[\sum_{u=1}^{d} \sum_{v=1}^{d} \left(\frac{1}{m} \sum_{i=1}^{m} \frac{1}{p_{r_{i}}} ([a_{r_{i}} b_{r_{i}}^{T}]_{uv})\right)^{2}\right]$$

$$\leq \mathbb{E}\left[\left(\frac{1}{m}\right)^{2} \sum_{i=1}^{m} \left(\frac{1}{p_{r_{i}}}\right)^{2} \sum_{u=1}^{d} \sum_{v=1}^{d} \left([a_{r_{i}} b_{r_{i}}^{T}]_{uv}\right)^{2}\right]$$

$$\leq \mathbb{E}\left[\left(\frac{1}{m}\right)^{2} \sum_{i=1}^{m} \left(\sum_{i=1}^{n} ||a_{i}|| \cdot ||b_{i}||\right)^{2}\right]$$

$$\leq \frac{1}{m} \left(\sum_{i=1}^{n} ||a_{i}|| \cdot ||b_{i}||\right)^{2}$$

The error gets reduced by a factor of $\frac{1}{m}$

Problem 4.

Consider a function $f(x): \mathbb{R}^2 \to \mathbb{R}$ with

$$f(x) = \frac{7}{2}x_1^2 + 4x_1x_2 + \frac{13}{2}x_2^2 = \frac{1}{2}x^T \begin{bmatrix} 7 & 4\\ 4 & 13 \end{bmatrix} x$$

(a) Show f(x) is β -smooth for as small β as you can. β -smoothness is defined by:

$$\|\nabla f(x) - \nabla f(y)\| \le \beta \|x - y\|$$

We can find $\nabla f(x)$ using the Jacobian matrix:

$$\nabla f(x) = \begin{bmatrix} \frac{\partial f}{\partial x_1} \\ \frac{\partial f}{\partial x_2} \end{bmatrix} = \begin{bmatrix} 7x_1 + 4x_2 \\ 4x_1 + 13x_2 \end{bmatrix} = \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} x$$

We can apply this to the definition of β -smoothness:

$$\|\nabla f(x) - \nabla f(y)\| \le \beta \|x - y\|$$

$$\left\| \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} x - \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} y \right\| \le \beta \|x - y\|$$

$$\left\| \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} \right\| \cdot \|x - y\| \le \beta \|x - y\|$$

$$\left\| \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} \right\| \le \beta$$

$$\left\| \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} \right\|_{2} \le \beta$$

$$\sigma_{1} \le \beta$$

Computing the SVD of $\begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix}$, we have that $\sigma_1 = 15$. Thus, $\beta = 15$.

(b) Suppose we run gradient descent from starting point $x^{(0)} = (-1,3)^T$ with constant step size $\eta = \frac{1}{\beta}$. Derive a closed form expression for $x^{(t)}$ and $f(x^{(t)})$.

We can derive a closed form expression for $x^{(t)}$ as follows:

$$\begin{split} x^{(t)} &= x^{(t-1)} - \eta \nabla f(x^{(t-1)}) \\ &= x^{(t-1)} - \eta \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} x^{(t-1)} \\ &= \left(I - \frac{1}{\beta} \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} \right) x^{(t-1)} \\ &= \left(I - \frac{1}{15} \begin{bmatrix} 7 & 4 \\ 4 & 13 \end{bmatrix} \right)^t x^{(0)} \\ &= \left(\frac{1}{15} \right)^t \begin{bmatrix} 8 & 4 \\ 4 & 2 \end{bmatrix}^t x^{(0)} \\ &= \left(\frac{1}{15} \right)^t \begin{bmatrix} \frac{2\sqrt{5}}{5} & -\frac{\sqrt{5}}{5} \\ \frac{\sqrt{5}}{5} & \frac{2\sqrt{5}}{5} \end{bmatrix} \begin{bmatrix} 10 & 0 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} \frac{2\sqrt{5}}{5} & \frac{\sqrt{5}}{5} \\ -\frac{\sqrt{5}}{5} & \frac{2\sqrt{5}}{5} \end{bmatrix} x^{(0)} \\ &= \left(\frac{1}{15} \right)^t \begin{bmatrix} \frac{2\sqrt{5}}{5} & -\frac{\sqrt{5}}{5} \\ \frac{\sqrt{5}}{5} & \frac{2\sqrt{5}}{5} \end{bmatrix} \begin{bmatrix} 10 & 0 \\ 0 & 0 \end{bmatrix}^t \begin{bmatrix} \frac{2\sqrt{5}}{5} & \frac{\sqrt{5}}{5} \\ -\frac{\sqrt{5}}{5} & \frac{2\sqrt{5}}{5} \end{bmatrix} x^{(0)} \\ &= \left(\frac{1}{15} \right)^t \begin{bmatrix} (\frac{4}{5})(10^t) & (\frac{2}{5})(10^t) \\ (\frac{2}{5})(10^t) & (\frac{1}{5})(10^t) \end{bmatrix} \begin{bmatrix} -1 \\ 3 \end{bmatrix} \\ &= \begin{bmatrix} \frac{2}{5}(\frac{2}{3})^t \\ \frac{1}{5}(\frac{2}{3})^t \end{bmatrix} \end{split}$$

As $x^{(t)}$ must be a vector, the final answer is $x^{(t)} = \left(\frac{2}{5}(\frac{2}{3})^t, \frac{1}{5}(\frac{2}{3})^t\right)^T$.

For $f(x^{(t)})$, the process is similar, but also different. We can rely on the fact that $x^{(t)} = \frac{2}{3}x^{(t-1)}$:

$$f(x^{(t)}) = \frac{1}{2}x^{(t)^T} \begin{bmatrix} 7 & 4\\ 4 & 13 \end{bmatrix} x^{(t)}$$

$$= (\frac{2}{3})^2 \cdot \frac{1}{2}x^{(t-1)^T} \begin{bmatrix} 7 & 4\\ 4 & 13 \end{bmatrix} x^{(t-1)}$$

$$= (\frac{2}{3})^2 \cdot f(x^{(t-1)})$$

$$= (\frac{2}{3})^{2t} \cdot f(x^{(0)})$$

Solving for $f(x^0) = 50$ results in $f(x^{(t)}) = (\frac{2}{3})^{2t} \cdot 50$.