

# PT

## Solution to Assignment 12

1. First, observe that

$$\text{Var}[(X + Y)/2] = (0.5)^2 \text{Var}(X + Y) = 0.25[\text{Var}(X) + \text{Var}(Y) + 2 \text{Cov}(X, Y)]$$

Then,

$$E(X) = \int_0^2 \int_0^2 x \frac{x+y}{8} dx dy = \frac{1}{8} \int_0^2 \frac{8}{3} + 2y dy = \frac{1}{8} \left( \frac{16}{3} + 4 \right) = \frac{2}{3} + \frac{1}{2} = \frac{7}{6}$$

$$E(X^2) = \int_0^2 \int_0^2 x^2 \frac{x+y}{8} dx dy = \frac{1}{8} \int_0^2 4 + \frac{8}{3} y dy = \frac{1}{8} \left( 8 + \frac{16}{3} \right) = 1 + \frac{4}{6} = \frac{10}{6}$$

$$\text{Var}(X) = 10/6 - (7/6)^2 = 11/36$$

By symmetry, the mean and the variance of  $Y$  are the same. Next,

$$E(XY) = \int_0^2 \int_0^2 xy \frac{x+y}{8} dx dy = \frac{1}{8} \int_0^2 \frac{8}{3} y + 2y^2 dy = \frac{1}{8} \left( \frac{16}{3} + \frac{16}{3} \right) = \frac{8}{6}$$

$$\text{Cov}(X, Y) = 8/6 - (7/6)(7/6) = -1/36$$

Finally,

$$\text{Var}(X + Y) = 0.25[11/36 + 11/36 + 2(-1/36)] = 5/36 = 10/72$$

2. We have  $(x, y) = (wz, z)$  so that

$$\frac{\partial(x, y)}{\partial(w, z)} = \det \begin{bmatrix} \frac{\partial x}{\partial w} & \frac{\partial x}{\partial z} \\ \frac{\partial y}{\partial w} & \frac{\partial y}{\partial z} \end{bmatrix} = \det \begin{bmatrix} z & w \\ 0 & 1 \end{bmatrix} = z.$$

Therefore for  $0 < wz < z < 1$ , or  $0 < w < 1$  and  $0 < z < 1$ ,

$$f_{W,z}(w, z) = f_{X,Y}(x(w, z), y(w, z)) \cdot \left| \frac{\partial(x, y)}{\partial(w, z)} \right| = 8wz \cdot z \cdot z = 8wz^3,$$

and  $f_{w,z}(w, z) = 0$  otherwise.

3. Here  $(X, Y)$  are jointly continuous and are related to  $(R, \Theta)$  by a one-to-one relationship. We use the method of transformations (Theorem 5.1). The function  $h(r, \theta)$  is given by

$$\begin{cases} x = h_1(r, \theta) = r \cos \theta \\ y = h_2(r, \theta) = r \sin \theta \end{cases}$$

Thus, we have

$$\begin{aligned} f_{R\Theta}(r, \theta) &= f_{XY}(h_1(r, \theta), h_2(r, \theta)) |J| \\ &= f_{XY}(r \cos \theta, r \sin \theta) |J| \end{aligned}$$

where

$$J = \det \begin{bmatrix} \frac{\partial h_1}{\partial r} & \frac{\partial h_1}{\partial \theta} \\ \frac{\partial h_2}{\partial r} & \frac{\partial h_2}{\partial \theta} \end{bmatrix} = \det \begin{bmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{bmatrix} = r \cos^2 \theta + r \sin^2 \theta = r.$$

We conclude that

$$\begin{aligned} f_{R\Theta}(r, \theta) &= f_{XY}(r \cos \theta, r \sin \theta) |J| \\ &= \begin{cases} \frac{r}{\pi} & r \in [0, 1], \theta \in (-\pi, \pi) \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

Note that from above we can write

$$f_{R\Theta}(r, \theta) = f_R(r) f_{\Theta}(\theta)$$

4.

- (a) From  $m$  and  $c$  we have  $X_2 \sim N(1, 2)$ . Thus

$$\begin{aligned} P(0 \leq X_2 \leq 1) &= \Phi\left(\frac{1-1}{\sqrt{2}}\right) - \Phi\left(\frac{0-1}{\sqrt{2}}\right) \\ &= \Phi(0) - \Phi\left(\frac{-1}{\sqrt{2}}\right) = 0.2602 \end{aligned}$$

- (b)

$$\begin{aligned} m_Y &= EY = AEX + b \\ &= \begin{bmatrix} 1 & 2 \\ 2 & 1 \\ 1 & 1 \end{bmatrix} \cdot \begin{bmatrix} 0 \\ 1 \end{bmatrix} + \begin{bmatrix} 0 \\ 1 \\ 2 \end{bmatrix} \\ &= \begin{bmatrix} 2 \\ 2 \\ 3 \end{bmatrix}. \end{aligned}$$

(c)

$$\begin{aligned} C_Y &= AC_X A^T \\ &= \begin{bmatrix} 1 & 2 \\ 2 & 1 \\ 1 & 1 \end{bmatrix} \cdot \begin{bmatrix} 1 & -1 \\ -1 & 2 \end{bmatrix} \cdot \begin{bmatrix} 1 & 2 & 1 \\ 2 & 1 & 1 \end{bmatrix} \\ &= \begin{bmatrix} 5 & 1 & 2 \\ 1 & 2 & 1 \\ 2 & 1 & 1 \end{bmatrix}. \end{aligned}$$

(d) From  $m_Y$  and  $c_Y$  we have  $Y_3 \sim N(3, 1)$ , thus

$$P(Y_3 \leq 4) = \Phi\left(\frac{4-3}{1}\right) = \Phi(1) = 0.8413$$

5. Note that

$$\begin{aligned} \begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} &= \begin{pmatrix} X_1 + X_2 + X_3 \\ X_1 - X_2 \end{pmatrix} \\ &= \begin{pmatrix} 1 & 1 & 1 \\ 1 & -1 & 0 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} = A \cdot \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix}. \end{aligned}$$

This implies that

$$\begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} \sim N(A\mu, A\Sigma A^T)$$

Where

$$A\mu = \begin{pmatrix} 1 & 1 & 1 \\ 1 & -1 & 0 \end{pmatrix} \begin{pmatrix} 2 \\ 1 \\ 2 \end{pmatrix} = \begin{pmatrix} 5 \\ 1 \end{pmatrix}$$

and

$$A\Sigma A^T = \begin{pmatrix} 1 & 1 & 1 \\ 1 & -1 & 0 \end{pmatrix} \begin{pmatrix} 2 & 1 & 1 \\ 1 & 3 & 0 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & 1 \\ 1 & -1 \\ 1 & 0 \end{pmatrix} = \begin{pmatrix} 10 & 0 \\ 0 & 3 \end{pmatrix}$$

So

$$\begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} \sim N\left(\begin{pmatrix} 5 \\ 1 \end{pmatrix}, \begin{pmatrix} 10 & 0 \\ 0 & 3 \end{pmatrix}\right).$$

6. We know that if

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}\right).$$

then

$$f(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)} \cdot \left[\frac{(x_1-\mu_1)^2}{\sigma_1^2} - \frac{2\rho(x_1-\mu_1)(x_2-\mu_2)}{\sigma_1\sigma_2} + \frac{(x_2-\mu_2)^2}{\sigma_2^2}\right]\right\}.$$

Comparing with the given form of  $f(x_1, x_2)$ , especially the coefficients in front of  $x_1^2$ ,  $x_2^2$  and  $x_1x_2$ , we get

$$\frac{1}{(1 - \rho^2) \sigma_1^2} = 1 \quad (1)$$

$$\frac{-2\rho}{(1 - \rho^2) \sigma_1 \sigma_2} = -1 \quad (2)$$

$$\frac{1}{(1 - \rho^2) \sigma_2^2} = 2 \quad (3)$$

$(1) \times (3) \div ((2)^2) \Rightarrow \frac{1}{4\rho^2} = 2 \Rightarrow \rho^2 = \frac{1}{8}$ . By (2), we know  $\rho > 0$ . So  $\rho = \frac{\sqrt{2}}{4}$ . By (1) and (3), we get  $\sigma_1^2 = \frac{8}{7}$  and  $\sigma_2^2 = \frac{4}{7}$ . So

$$\text{Cov}(\mathbf{X}) = \begin{pmatrix} \frac{8}{7} & \frac{2}{7} \\ \frac{2}{7} & \frac{4}{7} \end{pmatrix}.$$

Set  $\Sigma = \text{Cov}(\mathbf{X})$ . Then

$$\Sigma^{-1} = \begin{pmatrix} 1 & -\frac{1}{2} \\ -\frac{1}{2} & 2 \end{pmatrix}.$$

So

$$\begin{aligned} (\mathbf{x} - \boldsymbol{\mu})^\top \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) &= (x_1 - \mu_1, x_2 - \mu_2) \begin{pmatrix} 1 & -\frac{1}{2} \\ -\frac{1}{2} & 2 \end{pmatrix} \begin{pmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{pmatrix} \\ &= (x_1 - \mu_1)^2 + 2(x_2 - \mu_2)^2 - (x_1 - \mu_1)(x_2 - \mu_2). \end{aligned}$$

But

$$(\mathbf{x} - \boldsymbol{\mu})^\top \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) = x_1^2 + 2x_2^2 - x_1x_2 - 3x_1 - 2x_2 + 4.$$

So, comparing the Coefficients in front of  $x_1$  and  $x_2$ ,

$$\Rightarrow \begin{cases} -2\mu_1 + \mu_2 = -3 \\ -4\mu_2 + \mu_1 = -2 \end{cases} \Rightarrow \begin{cases} \mu_1 = 2 \\ \mu_2 = 1 \end{cases}$$

Therefore,

$$E(\mathbf{X}) = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} = \begin{pmatrix} 2 \\ 1 \end{pmatrix}$$

$$\text{and } \text{Cov}(\mathbf{X}) = \begin{pmatrix} \frac{8}{7} & \frac{2}{7} \\ \frac{2}{7} & \frac{4}{7} \end{pmatrix}.$$

7. If we define  $U = X + Y$  and  $V = 2X - Y$ , then note that  $U$  and  $V$  are jointly normal.

We have

$$E[U] = 3, \text{Var}[U] = 7, E[V] = 3, \text{Var}[V] = 37$$

and

$$\begin{aligned} \text{Cov}(U, V) &= \text{Cov}(X + Y, 2X - Y) = 2\text{Cov}(X, X) - \text{Cov}(X, Y) + 2\text{Cov}(Y, X) - \text{Cov}(Y, Y) \\ &= 2\text{Var}(X) + \text{Cov}(X, Y) - \text{Var}(Y) = 8 - 3 - 9 = -4. \end{aligned}$$

Thus,

$$\rho(U, V) = \frac{\text{Cov}(U, V)}{\sigma_U \sigma_V} = \frac{-4}{\sqrt{7 \times 37}}$$

We conclude that given  $V = 0$ ,  $U$  is normally distributed with

$$E[U | V = 0] = \mu_U + \rho(U, V) \sigma_U \frac{0 - \mu_V}{\sigma_V} = 3 + \frac{4}{\sqrt{7 \times 37}} \sqrt{7} \frac{3}{\sqrt{37}} = 3.324$$

$$\text{Var}[U | V = 0] = (1 - \rho(U, V)^2) \sigma_U^2 = \left(1 - \frac{16}{259}\right) \times 7 = 6.568$$

Thus

$$P(X + Y > 0 | 2X - Y = 0) = P(U > 0 | V = 0) = 1 - \Phi(-1.3) = 0.9032$$

8. Diagonalize  $\mathbf{C}_X$ , we have

$$\mathbf{C}_X = \begin{pmatrix} \frac{54}{49} & -\frac{6}{49} & \frac{24}{49} \\ -\frac{6}{49} & \frac{17}{49} & \frac{30}{49} \\ \frac{24}{49} & \frac{30}{49} & \frac{76}{49} \end{pmatrix} = \begin{pmatrix} \frac{3}{7} & -\frac{6}{7} & \frac{2}{7} \\ \frac{2}{7} & \frac{3}{7} & \frac{6}{7} \\ \frac{6}{7} & \frac{2}{7} & -\frac{3}{7} \end{pmatrix} \begin{pmatrix} 2 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} \frac{3}{7} & -\frac{6}{7} & \frac{2}{7} \\ \frac{2}{7} & \frac{3}{7} & \frac{6}{7} \\ \frac{6}{7} & \frac{2}{7} & -\frac{3}{7} \end{pmatrix}^T$$

Let  $\begin{pmatrix} Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} = \begin{pmatrix} \frac{3}{7} & -\frac{6}{7} & \frac{2}{7} \\ \frac{2}{7} & \frac{3}{7} & \frac{6}{7} \\ \frac{6}{7} & \frac{2}{7} & -\frac{3}{7} \end{pmatrix}^T \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} = \begin{pmatrix} \frac{3}{7}X_1 + \frac{2}{7}X_2 + \frac{6}{7}X_3 \\ -\frac{6}{7}X_1 + \frac{3}{7}X_2 + \frac{2}{7}X_3 \\ \frac{2}{7}X_1 + \frac{6}{7}X_2 - \frac{3}{7}X_3 \end{pmatrix}$ . Then  $\mathbf{C}_Y = \begin{pmatrix} 2 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 0 \end{pmatrix}$ . We obtain

$$0 = \text{Var}(Y_3) = \text{Var}\left(\frac{2}{7}X_1 + \frac{6}{7}X_2 - \frac{3}{7}X_3\right)$$

Hence  $\frac{2}{7}X_1 + \frac{6}{7}X_2 - \frac{3}{7}X_3 = c$  Since  $E[\mathbf{X}] = (1, 2, 3)^T$ , that is  $E[X_1] = 1, E[X_2] = 2, E[X_3] = 0$ , then  $c = \frac{2}{7}E[X_1] + \frac{6}{7}E[X_2] - \frac{3}{7}E[X_3] = 2$  Therefore

$$\frac{2}{7}X_1 + \frac{6}{7}X_2 - \frac{3}{7}X_3 = 2 \quad \text{and} \quad X_3 = \frac{2}{3}X_1 + 2X_2 - \frac{14}{3}$$

By Theorem 4.49,  $\text{Cov}(Y_1, Y_2) = 0$  implies that  $Y_1, Y_2$  are independent.

$$Y_1 = \frac{3}{7}X_1 + \frac{2}{7}X_2 + \frac{6}{7}X_3 = \frac{3}{7}X_1 + \frac{2}{7}X_2 + \frac{6}{7}\left(\frac{2}{3}X_1 + 2X_2 - \frac{14}{3}\right) = X_1 + 2X_2 - 4$$

$$Y_2 = -\frac{6}{7}X_1 + \frac{3}{7}X_2 + \frac{2}{7}X_3 = -\frac{6}{7}X_1 + \frac{3}{7}X_2 + \frac{2}{7}\left(\frac{2}{3}X_1 + 2X_2 - \frac{14}{3}\right) = -\frac{2}{3}X_1 + X_2 - \frac{4}{3}$$

Hence  $X_1 + 2X_2 - 4$  and  $-\frac{2}{3}X_1 + X_2 - \frac{4}{3}$  are independent.