

## MS: Solution to Test 2

1. (a) By computing the first moment, we yield  $\theta = 3/2 - E(X)/2$ , since

$$E(X) = 1 \cdot P(X = 1) + 3 \cdot P(X = 3) = 1 \cdot \theta + 3 \cdot (1 - \theta) = 3 - 2\theta$$

Replacing  $E(X)$  by the sample mean  $\bar{X}$ , we have the MOM estimator of  $\theta$  as

$$\hat{\theta}_{MOM} = \frac{3}{2} - \frac{1}{2}\bar{X}$$

where

$$\bar{X} = \frac{1 + 3 + 1 + 1}{4} = \frac{3}{2}.$$

Then the MOM estimate of  $\theta$  is

$$\hat{\theta}_{MOM} = \frac{3}{2} - \frac{1}{2} \cdot \frac{3}{2} = \frac{3}{4}.$$

- (b) The likelihood function is

$$\begin{aligned}\mathcal{L}(\theta) &= f(x_1, \dots, x_n | \theta) = f(x_1 | \theta) \cdots f(x_n | \theta) \\ &= P(X = x_1 | \theta) \cdots P(X = x_n | \theta)\end{aligned}$$

Substituting  $n = 4$  and the realizations of  $X$  yields

$$\mathcal{L}(\theta) = P(X = 3 | \theta) \cdot P(X = 1 | \theta)^3 = (1 - \theta) \cdot \theta^3$$

- (c) To find the MLE of  $\theta$ , it's easier to work with the natural logarithm of  $\mathcal{L}(\theta)$ ,

$$l(\theta) = \ln(\mathcal{L}(\theta)) = 3 \ln(\theta) + \ln(1 - \theta),$$

$$l'(\theta) = \frac{3}{\theta} - \frac{1}{1 - \theta} = \frac{3 - 4 \cdot \theta}{\theta \cdot (1 - \theta)}.$$

So

$$l'(\theta) = \frac{3 - 4 \cdot \theta}{\theta \cdot (1 - \theta)} = 0 \iff 3 - 4 \cdot \theta = 0 \iff \hat{\theta}_{MLE} = \frac{3}{4}.$$

2. (a) Let  $n$  be the sample size, and let  $X_1, \dots, X_n \sim N(0, \theta)$  be independent identically distributed random variables with the same density function. Remember that the density function of  $X \sim N(0, \theta)$  is

$$f(x) = \frac{1}{\sqrt{\theta} \cdot \sqrt{2\pi}} \cdot e^{-\frac{x^2}{2\theta}}, \quad x \in \mathbb{R}.$$

To find the MLE of  $\theta$ , we first define the likelihood function:

$$\mathcal{L}(\theta) = f(x_1, \dots, x_n | \theta) = f(x_1 | \theta) \cdots f(x_n | \theta).$$

Substituting the definition of the density function of  $X$  yields

$$\begin{aligned}\mathcal{L}(\theta) &= \frac{1}{\sqrt{\theta} \cdot \sqrt{2\pi}} \cdot e^{-\frac{x_1^2}{2\theta}} \cdots \frac{1}{\sqrt{\theta} \cdot \sqrt{2\pi}} \cdot e^{-\frac{x_n^2}{2\theta}} \\ &= \left( \frac{1}{\sqrt{\theta} \cdot \sqrt{2\pi}} \right)^n \cdot e^{-\frac{x_1^2 + \dots + x_n^2}{2\theta}}\end{aligned}$$

It's easier to work with the natural logarithm of the given expression, so we define

$$l(\theta) = \ln(\mathcal{L}(\theta)) = n \cdot \ln\left(\frac{1}{\sqrt{\theta} \cdot \sqrt{2\pi}}\right) - \frac{1}{2\theta} \cdot \sum_{i=1}^n x_i^2,$$

and we need to find its global maximum on the interval  $(0, +\infty)$  (where  $\theta$  can take on values).

The derivative of  $l$  (with respect to  $\theta$ ) is

$$l'(\theta) = -\frac{n}{2\theta} + \frac{1}{2\theta^2} \cdot \sum_{i=1}^n x_i^2.$$

So

$$\begin{aligned} l'(\theta) = 0 &\iff -\frac{n}{2\theta} + \frac{1}{2\theta^2} \cdot \sum_{i=1}^n x_i^2 = 0 \iff \sum_{i=1}^n x_i^2 = n \cdot \theta \\ &\iff \theta = \frac{1}{n} \cdot \sum_{i=1}^n x_i^2. \end{aligned}$$

Therefore, the MLE for  $\theta$  is

$$\hat{\theta} = \frac{1}{n} \cdot \sum_{i=1}^n x_i^2.$$

(b) The log-density function is

$$\log f(X|\theta) = -\log(\sqrt{2\pi\theta}) - \frac{X^2}{2\theta}$$

Then the score function is

$$\begin{aligned} \frac{\partial}{\partial \theta} \log f(X|\theta) &= -\frac{1}{2\theta} + \frac{X^2}{2\theta^2} \quad \text{and} \quad \frac{\partial^2}{\partial \theta^2} \log f(X|\theta) = \frac{1}{2\theta^2} - \frac{X^2}{\theta^3} \\ I(\theta) &= -E\left(\frac{\partial^2}{\partial \theta^2} \log f(X|\theta)\right) = -\frac{1}{2\theta^2} + \frac{E(X^2)}{\theta^3} = -\frac{1}{2\theta^2} + \frac{\text{Var}(X) + E(X)^2}{\theta^3} = \frac{1}{2\theta^2} \end{aligned}$$

The asymptotic variance of the mle  $\hat{\theta}_{MLE}$  is  $1/nI(\theta)$ , that is,  $2\theta^2/n$ .

(c) Since  $\sqrt{nI(\theta)}(\hat{\theta}_{MLE} - \theta)$  approximately follows  $N(0, 1)$ , namely,

$$\begin{aligned} \frac{\sqrt{n}(\hat{\theta}_{MLE} - \theta)}{\sqrt{2\theta^2}} &\sim N(0, 1) \\ -z_{0.05} &\leq \frac{\sqrt{n}(\hat{\theta}_{MLE} - \theta)}{\sqrt{2\theta^2}} \leq z_{0.05} \end{aligned}$$

1. With  $z_{0.05} = 1.645$ , then

$$1 - 1.645 \cdot \sqrt{\frac{2}{n}} \leq \frac{\hat{\theta}_{MLE}}{\theta} \leq 1 + 1.645 \cdot \sqrt{\frac{2}{n}}$$

An approximate 90% interval estimator for  $\theta$ :

$$\left[ \hat{\theta}_{MLE} / \left( 1 + 1.645 \cdot \sqrt{\frac{2}{n}} \right), \hat{\theta}_{MLE} / \left( 1 - 1.645 \cdot \sqrt{\frac{2}{n}} \right) \right]$$

where  $n$  is large enough and  $n \geq 6$ .

2. Replace unknown  $I(\theta)$  by  $I(\hat{\theta}_{MLE})$ , an approximate 90% interval estimator for  $\theta$ :

$$\left[ \hat{\theta}_{MLE} \cdot \left( 1 - 1.645 \cdot \sqrt{\frac{2}{n}} \right), \hat{\theta}_{MLE} \cdot \left( 1 + 1.645 \cdot \sqrt{\frac{2}{n}} \right) \right]$$

(d) The mle of  $\theta$  is unbiased since

$$\hat{\theta}_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i^2 \quad \text{with mean} \quad E(\hat{\theta}_{MLE}) = \frac{n\theta}{n} = \theta$$

The exact distribution of  $\frac{n\hat{\theta}_{MLE}}{\theta}$  is  $\chi^2(n)$ . An exact 90% confidence interval of  $\theta$  can be constructed from the statistics  $n\hat{\theta}_{MLE}/\theta$ :

$$\left[ 10\hat{\theta}_{MLE}/\chi_{0.05}^2(10), 10\hat{\theta}_{MLE}/\chi_{0.95}^2(10) \right]$$

where  $\chi_{0.05}^2(10) = 18.307$  and  $\chi_{0.95}^2(10) = 3.94$ .

3. (a) Accept  $H_0$ , since

$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = -1.739$$

and  $|z| < 1.96 = z_{0.025}$ .

(b) Reject  $H_0$ , since

$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = -1.739$$

and  $z < -1.645 = -z_{0.05}$ .

4. (a) Suppose  $X_1, \dots, X_n$  are samples from  $U(0, \theta)$ . Then for  $0 \leq x \leq \theta$ ,

$$\begin{aligned} F_{X_{\max}}(x) &= P(X_{\max} \leq x) = P(X_1 \leq x, \dots, X_n \leq x) = \prod_{i=1}^n P(X_i \leq x) = \left(\frac{x}{\theta}\right)^n \\ \Rightarrow f_{X_{\max}}(x) &= \frac{nx^{n-1}}{\theta^n}, \quad 0 \leq x \leq \theta. \end{aligned}$$

We set

$$P(X_{\max} \leq 2.5 \mid H_0 \text{ is true}) = \alpha, \tag{1}$$

and the decision rule is "Reject  $H_0$  if  $X_{\max} \leq 2.5$ ".

The pdf of  $X_{\max}$  given that  $H_0$  is true is

$$f_{X_{\max}}(x \mid \theta = 3) = \frac{nx^{n-1}}{3^n}, \quad 0 \leq x \leq 3$$

$$\begin{aligned} P(X_{\max} \leq 2.5 \mid H_0 \text{ is true}) &= \alpha \\ \Rightarrow \int_0^{2.5} \frac{nx^{n-1}}{3^n} dx &= \alpha \\ \Rightarrow \alpha &= \left(\frac{2.5}{3}\right)^n \end{aligned}$$

(b) To make  $\alpha \leq 0.05$ , we have to choose  $n$  such that

$$\left(\frac{2.5}{3}\right)^n \leq 0.05.$$

We thus obtain

$$n \ln(2.5/3) \leq \ln(0.05) \quad \Rightarrow \quad n \geq 16.431.$$

So  $n = 17$ .