

Stochastic programming and robust optimisation

Lecture 3/5

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Outline of this lecture

Introduction

Chance constraints

Risk measures

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Risk measures

Beyond expected values

Many settings require that a **risk profile** is imposed:

- ▶ Expected values assume a **risk neutral** stance;
- ▶ Risk neutral means that the product **probability** \times **outcome** is the single factor under consideration.

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- ▶ Risk neutral means that the product **probability** \times **outcome** is the single factor under consideration.

In many settings, the decision-maker may have other implicit priorities:

- ▶ Impose (probabilistic) guarantees on **feasibility**
 \Rightarrow **chance constraints**;
- ▶ Avoid being exposed to the possibility of a **high-loss**
 \Rightarrow **risk measures**.

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Chance constraints

Risk measures

Feasibility guarantees

A solution may **reveal itself infeasible** once the uncertainty unveils

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- ▶ Appealing in settings in which **safety and resilience** are requirements.

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Let us first consider the case **without** recourse. Our problem is thus

$$\begin{aligned} \min. \quad & c^\top x \\ \text{s.t.:} \quad & Ax = b \\ & T(\xi)x = h(\xi), \quad \forall \xi \in \Xi \\ & x \geq 0. \end{aligned}$$

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Notice the **static** nature of the problem

- ▶ once a decision x is made, we observe the realisation of the uncertainty ξ ;
- ▶ no correction decision is allowed.

Feasibility guarantees

There are two main paradigms on how to model feasibility requirements for the constraint $T(\xi)x = h(\xi)$, $\forall \xi \in \Xi$:

1. Impose that $T(\xi)x = h(\xi)$ holds for a set of realisations $U \subseteq \Xi$ (robust optimisation)
2. Impose that the probability of $T(\xi)x = h(\xi)$ holding is at least a given threshold (chance constraints)

As one may suspect, the two approaches are interrelated.

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Imposing a chance constraint means that a solution is deemed acceptable if, for a given confidence level α , we have that

$$\mathbb{P}(T(\xi)x = h(\xi)) \geq \alpha.$$

Types of chance constraints

Let $T(\xi)$ be a $m_2 \times n_1$ matrix, where $T(\xi)_i$ represents its i^{th} -row, and $h(\xi)$ a m_2 vector with components $h(\xi)_i$.

There are two types of chance constraints:

1. Individual chance constraints (ICC):

$$p_i(x) := \mathbb{P}((T(\xi)_i)^\top x = h(\xi)_i) \geq \alpha_i, \quad \forall i \in [m_2],$$

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JCCs are typically more challenging from a tractability standpoint

- ▶ ICCs can be used to **approximate** a JCC;
- ▶ **Bonferroni inequality**: for a given x , if $p_i(x) > \alpha_i, \forall i \in [m_2]$, where $\alpha_i = 1 - \frac{(1-\alpha)}{m_2}$, then $p(x) \geq \alpha$.

Tractability of chance constraints

Tractability issues stem from the properties of the **feasible sets** generated by chance constraints. Let

$$C(\alpha_1, \dots, \alpha_{m_2}) := \bigcap_{i \in [m_2]} C_i(\alpha_i), \text{ where}$$

$$C_i(\alpha_i) = \{x \in \mathbb{R}^n : p_i(x) \geq \alpha_i\}.$$

- ▶ No general result that guarantees the **convexity** of $C_i(\alpha_i)$;
- ▶ **Particular cases** do exist for important distributions which lead to the convexity of $C(\alpha_1, \dots, \alpha_{m_2})$.

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- ▶ **Particular cases** do exist for important distributions which lead to the convexity of $C(\alpha_1, \dots, \alpha_{m_2})$.

For example, assume that $T(\xi) = T, \forall \xi \in \Xi$, and that $h(\xi) = \xi$. For the univariate case, we have that

$$C(\alpha) = \{x \in \mathbb{R}^{n_1} : Tx \geq F^{-1}(\alpha)\}.$$

Tractability of chance constraints

One general result that is known and can be informative in the multivariate case is the following:

Theorem 1

Let $T(\xi) = T$, $\forall \xi \in \Xi$, and $h(\xi) = \xi$, where $\xi \in \mathbb{R}^{m_2}$ is a random vector with density function f . If $\log(f)$ is concave (assuming $\log(0) = -\infty$), then $C(\alpha)$ is closed and convex for all $\alpha \in [0, 1]$.

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- ▶ **Main case:** $\xi \sim \text{Normal}(\mu, \Sigma)$ with vector mean μ and covariance matrix Σ ;
- ▶ Uniform case also “trivial” to show that holds;
- ▶ For a list of “other distributions”: [Nemirovski and Shapiro, 2007]

Tractability of chance constraints

Another important known case is this: $T(\xi)$ is a $1 \times n$ random vector and $h(\xi) = h, \forall \xi \in \Xi$.

Theorem 2

Assume that $T(\xi) = \xi = (\xi_i)_i = 1^{n_1}$ is the only random parameter, where $\xi \sim \text{Normal}(\mu, \Sigma)$ with $\mu = (\mu_i)_{i=1}^{n_1}$ a vector of means and Σ the covariance matrix. Then

$$C(\alpha) = \left\{ x \in \mathbb{R}^{n_1} : \mu^\top x \geq h + \Phi^{-1}(\alpha) \sqrt{x^\top \Sigma x} \right\}$$

Tractability of chance constraints

Proof.

The random variable $\xi^\top x$ is a multivariate normal with mean $\mu^\top x$ and variance $x^\top \Sigma x$. Letting Z follow a standard normal, we have that

$$\begin{aligned}\mathbb{P}(\xi^\top x \geq h) \geq \alpha &\Leftrightarrow \mathbb{P}\left(\frac{\xi^\top x - \mu^\top x}{\sqrt{x^\top \Sigma x}} \geq \frac{h^\top x - \mu^\top x}{\sqrt{x^\top \Sigma x}}\right) \geq \alpha \\&\Leftrightarrow 1 - \mathbb{P}\left(Z \leq \frac{h - \mu^\top x}{\sqrt{x^\top \Sigma x}}\right) \geq \alpha \\&\Leftrightarrow 1 - \Phi\left(\frac{h - \mu^\top x}{\sqrt{x^\top \Sigma x}}\right) \geq \alpha \\&\Leftrightarrow \Phi\left(\frac{\mu^\top x - h}{\sqrt{x^\top \Sigma x}}\right) \geq \alpha \\&\Leftrightarrow \frac{\mu^\top x - h}{\sqrt{x^\top \Sigma x}} \geq \Phi^{-1}(\alpha) \Leftrightarrow \mu^\top x \geq h + \Phi^{-1}(\alpha)\sqrt{x^\top \Sigma x} \quad \square\end{aligned}$$

Notice that the constraint is convex if $\Phi^{-1}(\alpha) \geq 0$, i.e., $\alpha \in [1/2, 1]$.

Discretisation of chance constraints

An alternative way of handling chance constraints is using **scenarios**:

- ▶ Allows for general (discrete) distributions;
- ▶ **Convex** problems by construction (for an originally convex problem);
- ▶ Allows for recourse decisions;
- ▶ Requires **binary variables** per scenario, which may be an issue computationally.

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Let us consider again our scenario-based deterministic equivalent 2SSP:

$$\begin{aligned} \min. \quad & c^\top x + \sum_{s \in S} P_s q_s^\top y_s \\ \text{s.t.:} \quad & Ax = b, x \geq 0 \\ & T_s x + W_s y_s = h_s, \forall s \in S \\ & y_s \geq 0, \forall s \in S. \end{aligned}$$

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Let $v_s \in \{0, 1\}$, $u_s \in \mathbb{R}$, $\forall s \in S$, and M be a sufficiently large (big-M) value. Then, we can reformulate our chance-constrained problem as

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$$\min. c^\top x + \sum_{s \in S} P_s q_s^\top y_s \quad (1a)$$

$$\text{s.t.: } Ax = b \quad (1b)$$

$$T_s x + W_s y_s = h_s + u_s, \quad \forall s \in S \quad (1c)$$

$$|u_s| \leq M v_s, \quad \forall s \in S \quad (1d)$$

$$\sum_{s \in S} P_s v_s \leq 1 - \alpha \quad (1e)$$

$$x \geq 0 \quad (1f)$$

$$y_s \geq 0, u_s \in \mathbb{R}, v_s \in \{0, 1\} \quad \forall s \in S, \quad (1g)$$

where α be our **feasibility likelihood**.

Discretisation of chance constraints

Some final remarks:

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- ▶ $\sum_{s \in S} P_s v_s$ gives the **infeasibility probability**;
- ▶ Notice that one **binary variable** per scenario is required;
- ▶ The above can be circumvented using **integrated chance constraints**.
 - Alternative, one can impose limits on the **expected infeasibility** (variables $u_s, \forall s \in S$)
 - This is achieved by replacing (1d) and (1e) with

$$\sum_{s \in S} P_s u_s \leq \beta$$

where β is a limit on the expected amount of infeasibility;

Chance constraints

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Chance constraints

Risk measures

Measuring risk

Recall that we seek to find a solution x that optimises

$$\min_x \mathbb{E}_\xi [F(x, \xi)],$$

where:

- ▶ $F(x, \xi) = \{c^\top x + Q(x, \xi) : x \in X\};$
- ▶ $Q(x, \xi) = \min_y \{q(\xi)^\top y : W(\xi)y = h(\xi) - T(\xi)x, y \geq 0\};$
- ▶ $X = \{x \in \mathbb{R}^n : Ax = b, x \geq 0\}.$

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That is, we choose $x^\star = \operatorname{argmin}_x \mathbb{E}_\xi [F(x, \xi)]$.

- ▶ Each $x' \in X$ has an associated a **probability distribution** $f_x(\xi)$ which maps a cost $F(x', \xi)$ to the probability of scenario ξ ;
- ▶ Thus, x' is **preferred** over x'' if $\mathbb{E}_\xi [F(x', \xi)] < \mathbb{E}_\xi [F(x'', \xi)]$

Measuring risk

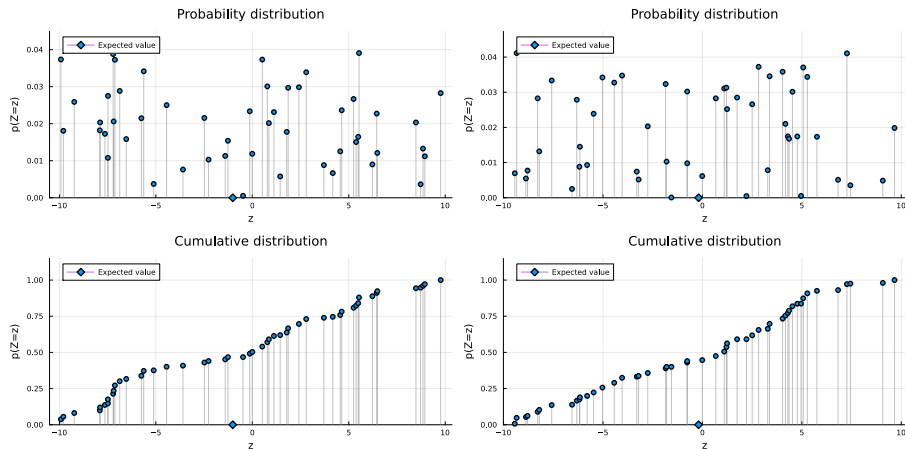


Figure: Comparing two solutions: the solution generating the cost distribution on the left is preferred, as it has a lower expected value

Measuring risk

However, choosing between distributions using their expected values neglects **information about the dispersion**:

- ▶ Higher-order statistical moments are disregarded
- ▶ **Tails** of the cost distribution are often relevant from a decision-making standpoint.

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- ▶ **Tails** of the cost distribution are often relevant from a decision-making standpoint.

To capture more information about such tails we can define **risk measures** $r : X \rightarrow \mathbb{R}$ such that

- ▶ r associates the random variable $F(x, \xi)$ generated by the solution x with a **real-valued risk** $r_\xi(x)$
- ▶ Analogously, x' can be chosen over x'' if $r_\xi[F(x', \xi)] < r_\xi[F(x'', \xi)]$

Trading off risk and expected return

Being **two conflicting objectives**, risk and return are typically considered under a bi-objective standpoint, e.g., using

1. **Weighted terms** in the objective function:

$$\min_x \mathbb{E}_\xi [F(x, \xi)] + \beta r_\xi [F(x, \xi)],$$

where $\beta = 0$ represents a risk-neutral stance and risk aversion increases as $\beta \rightarrow \infty$;

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2. A **risk exposition budget** δ :

$$\begin{aligned} \min_x \quad & \mathbb{E}_\xi [F(x, \xi)] \\ \text{s.t.:} \quad & r_\xi [F(x, \xi)] \leq \delta. \end{aligned}$$

Coherent risk measures

[Artzner et al., 1999] provides axiomatic definitions for coherent risk measures:

1. **Translation invariance:** $r_\xi [F(x, \xi) + a] = r_\xi [F(x, \xi)] + a$ for $a \in \mathbb{R}$.
2. **Subadditivity:** $r_\xi [F(x', \xi) + F(x'', \xi)] \leq r_\xi [F(x', \xi)] + r_\xi [F(x'', \xi)]$
3. **Positive homogeneity:** $r_\xi [F(x, \xi) \times a] = r_\xi [F(x, \xi)] \times a$ for $a \in \mathbb{R}$.
4. **Monotonicity:** if for every ξ , we have that $F(x', \xi) \leq F(x'', \xi)$, then $r_\xi [F(x', \xi)] \leq r_\xi [F(x'', \xi)]$

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This has been further developed by [Rockafellar, 2007], who establishes the role of **coherence** in optimisation problems. A coherent risk measure:

- ▶ Preserves **convexity**;
- ▶ Preserves **certainty**;
- ▶ Is insensitive to **scaling**.

Conditional value-at-risk

The **most widespread risk measure** in the context of optimisation is the Conditional Value-at-Risk (CVaR)

- ▶ A **coherent** risk measure widely used in other areas;
- ▶ Empirical results on its efficacy in production planning:
[Alem et al., 2020]

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Let X be a random variable and F_X its cumulative distribution function. Then, for a confidence level α , the **Value-at-Risk** (VaR_α) is defined as

$$VaR_\alpha(X) = \min\{\eta : F_X(\eta) \geq \alpha\}.$$

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The conditional VaR_α represents the **expectation** of X in the **conditional distribution** of its α -upper tail, i.e.,

$$CVaR_\alpha(X) = \min_{\eta \in \mathbb{R}} \left\{ \eta + \frac{\mathbb{E}[X - \eta]^+}{(1 - \alpha)} \right\},$$

where $[\cdot]^+ = \max\{0, \cdot\}$.

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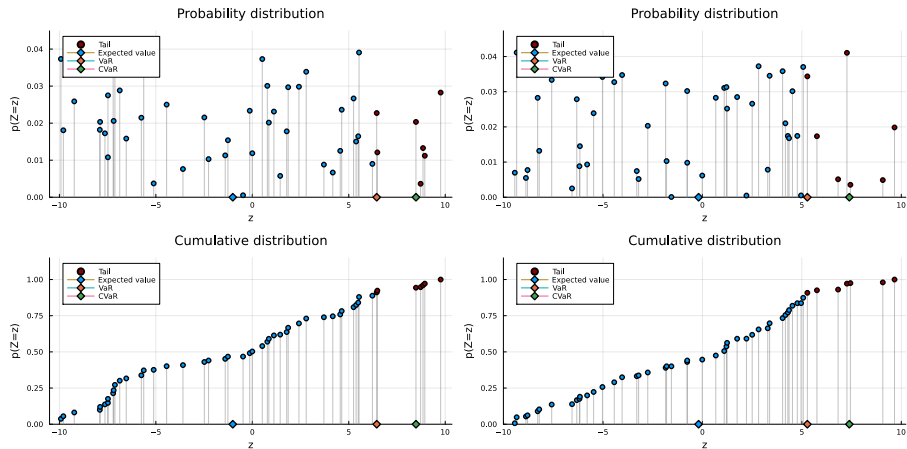


Figure: Comparing two solutions: the solution on the right has better (smaller) $CVaR_{90\%}$

Conditional value-at-risk in SP models

One appealing feature of CVaR is its convexity:

- ▶ Requires discretisation to handle the expected value;
- ▶ In the context of optimisation, this means that **no additional binary variables** are needed;
- ▶ This contrasts with VaR (or chance constraints), which need such variables.

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Recall our **risk-neutral** scenario-based deterministic equivalent 2SSP:

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Conditional value-at-risk in SP models

Let us define the following auxiliary terms:

- ▶ $\beta \in [0, 1]$: weight for the risk term,
- ▶ α : confidence level;
- ▶ $\eta \geq 0$: represent the value at risk (VaR);
- ▶ $\pi_s \geq 0, \forall s \in S$: account for $[X - \eta]^+$. Here $X \equiv c^\top x + q_s^\top y_s$.

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Then, the **risk-averse** scenario-based deterministic equivalent 2SSP is

$$\min. (1 - \beta) \left[c^\top x + \sum_{s \in S} P_s q_s^\top y_s \right] + \beta \left[\eta + \frac{\sum_{s \in S} P_s \pi_s}{1 - \alpha} \right]$$

$$\text{s.t.: } Ax = b, x \geq 0$$

$$T_s x + W_s y_s = h_s, \forall s \in S$$

$$\pi_s \geq c^\top x + q_s^\top y_s - \eta, \forall s \in S$$

$$y_s \geq 0, \pi_s \geq 0, \forall s \in S$$

$$\eta \geq 0.$$

Conditional value-at-risk in SP models

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- ▶ There are other more recent risk measures for multistage settings ([Dowson et al., 2022]) that:

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- ▶ There are other more recent risk measures for multistage settings ([Dowson et al., 2022]) that:
 - Behave better (computationally) in **dynamic** settings

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Some final practical remarks:

- ▶ Scaling is important. $\beta = 0.5$ is not necessarily a **midpoint** between risk-neutral and risk-averse solutions.
- ▶ If maximising, pay attention to the **sign** of the additional terms, as they must change accordingly.
- ▶ CVaR can also be used in **multi-stage** problems (see [Shapiro, 2011])
- ▶ There are other more recent risk measures for multistage settings ([Dowson et al., 2022]) that:
 - Behave better (computationally) in **dynamic** settings
 - Serve as proxy to **other risk-aversion paradigms** (worst-case minimisation or distributional robustness)

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