Stochastic programming & Robust optimisation

Lecture 2/4

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Outline of this lecture

Introduction

Scenario trees

Generating scenario trees

Scenario (tree) generation methods

Sample Average Approximation (SAA)

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Mathematical programming models in which some of the parameters are assumed to be random variables.

It comprises the following parts:

- 1. A mathematical programming model
- 2. Deterministic parameter values
- 3. Description of the stochasticity, e.g.,
 - a known probability distribution;
 - historical data;
 - distribution properties (average, standard deviation, i.e., moments)

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The most widespread use of stochastic programs relies on scenarios:

- Lead to tractable deterministic equivalents;
- Are approximations of the original stochastic process

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A scenario tree ξ comprises sequentially observed realisations of ξ^t , for $t=1,\ldots,H$:

- \blacktriangleright $\xi = (\xi^t)_{t \in [H]}$, where (\cdot) denotes a sequence and $\xi^t \in \Xi_t$;
- lacktriangle a scenario is denoted $\xi_s=(\xi_s^t)_{t\in[H]}$ forming a "path" through ξ ;
- ▶ Thus, $\xi = \{\xi_s\}_{s \in [S]}$, where S is the number of scenarios.

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Example:

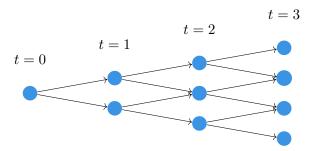


Figure: A 4-stage (lattice) scenario tree with 2 scenarios per stage. $\xi = (\xi^1, \xi^2, \xi^3)$;

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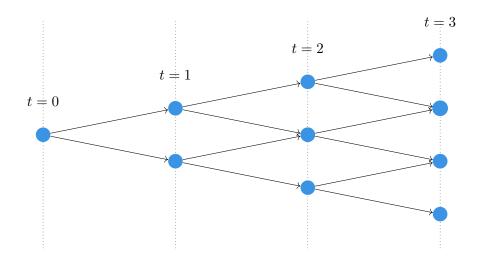
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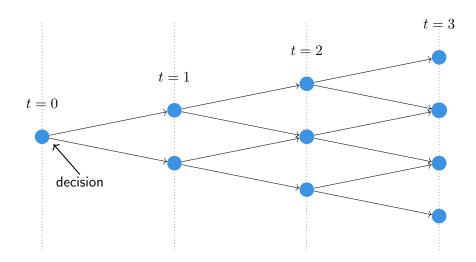
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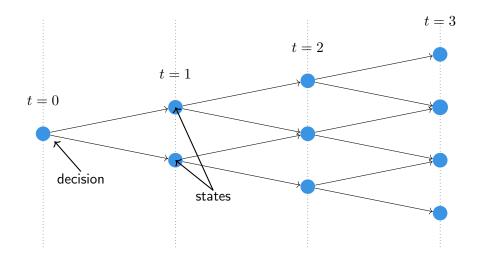
Terminology



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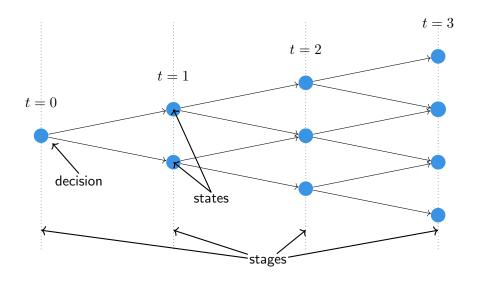


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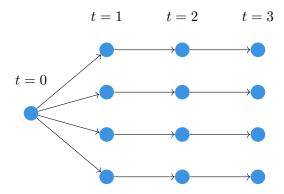


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Terminology



Taxonomy of scenario trees



Branching indicates a decision upon arrival of new information

- No branching, no additional information;
- ► Fan trees represent multi-period 2-stage problems.

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Trade-off approximation quality vs. tractability

Two parameters govern the geometry of a scenario tree:

- **Depth:** number of stages H
- **Breadth (or width):** number of realisations per stage $|\xi^t|$

Trade-off approximation quality vs. tractability

Two parameters govern the geometry of a scenario tree:

- **Depth:** number of stages *H*
- **Breadth (or width):** number of realisations per stage $|\xi^t|$

The total of scenarios is $O(N^H)$ (assuming $|\xi_t| = N$ for $t \in [H]$)

- ► Larger *H* convey more adaptability to revealed information;
- ► Larger S convey a more precise description of the uncertainty;
- Computational tractability issues pressure them to be as small as possible.

Most scenario generation methods seek to find trees with minimal $|\xi|$ such that representation quality requirements are observed.

Data source

Typical sources for scenarios include:

- 1. Historical data: past observations as possible future observations;
- 2. **Simulation models:** Monte Carlo, systems dynamics, agent-based and discrete event simulation;
- 3. **Expert elicitation:** typically a small number of scenarios with no possible (out-of-sample) testing.

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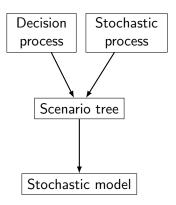
Often, a combination of the above is used:

- 1. Start from the data;
- 2. Define and fit a parametric model;
- 3. Generate observations from the model.

Scenario generation and modelling

Scenario generation must be part of the modelling process

- Problem dependent;
- The method for generating scenarios is a modelling decision;
- Often overlooked in applications;
- Quality of scenarios majorly influences quality of solution ("garbage in = garbage out").



Apart from epistemic error questions, two measures must be considered when generating scenario trees:

1. Error

- Error introduced for using an approximation of the real stochastic process;
- Unlikely to be measurable, but possible to be approximated.

2. Stability

- Scenario-trees approximating the same stochastic process should yield the same solution;
- Likewise, objective function values should be stable.

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- Scenario-trees approximating the same stochastic process should yield the same solution;
- Likewise, objective function values should be stable.

Let ξ be a scenario tree representing the original stochastic process η , and $\mathcal{F}(x,\xi)=\mathbb{E}_{\xi}\left[F(x,\xi)\right]$. We are interested in understanding how well

$$\min_{x} \mathcal{F}(x,\xi)$$
 approximates $\min_{x} \mathcal{F}(x,\eta)$

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Let ξ_k , for $k=1,\ldots,n$, be a collection of alternative scenario trees generated (e.g., by sampling) to represent η . We have that

$$x_k^{\star} = \arg\min_{x} \mathcal{F}(x, \xi_k).$$

The approximation error [Pflug, 2001] is defined as

$$e(\eta, \xi_k) = \mathcal{F}(\arg\min_{x} \mathcal{F}(x, \xi_k), \eta) - \mathcal{F}(\arg\min_{x} \mathcal{F}(x, \eta), \eta)$$
$$= \mathcal{F}(x_k^{\star}, \eta) - \min_{x} \mathcal{F}(x, \eta).$$

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$$= \mathcal{F}(x_k^{\star}, \eta) - \min_{x} \mathcal{F}(x, \eta).$$

- lacktriangle Calculating $\mathcal{F}(x_k^\star,\eta)$ requires evaluating the "true" objective function;
- ▶ Alternatively, Monte Carlo simulation is often employed to approximate $\mathcal{F}(x_k^{\star}, \eta)$;
- ▶ Clearly, there is no way to evaluate $\min_x \mathcal{F}(x, \eta)$.

Out-of-sample stability

Assume that we can approximate $\mathcal{F}(x_k^{\star}, \eta)$. This allows us to

- ightharpoonup compare solutions x_1^{\star} and x_2^{\star} ;
- compare alternative scenario generation methods;
- perform out-of-sample stability test:
 - 1. Generate a set of scenario trees $\{\xi_1, \ldots, \xi_n\}$;
 - 2. Obtain solutions x_k , $k = 1, \ldots, n$;
 - 3. Test whether $\mathcal{F}(x_k^{\star}, \eta) \approx \mathcal{F}(x_l^{\star}, \eta)$, for $k, l = 1, \dots, n : k \neq l$.

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Remarks:

- \bullet $e(\eta, \xi_k) \approx 0 \Rightarrow e(\eta, \xi_k) \approx e(\eta, \xi_l) \equiv \mathcal{F}(x_k^{\star}, \eta) \approx \mathcal{F}(x_l^{\star}, \eta);$
- ► The procedure above can also be used to assess scenario tree width (scenarios per stage).

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In some contexts, can also be defined as

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- No direct connection to out-of-sample stability;
- Useful for assessing the internal stability of a random scenario generation method;
- Translates into confidence in the objective function value reported.

Final considerations

Some practical advice:

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- No stability implies dependence on the scenario tree. To improve stability one can
 - 1. Consider alternative scenario generation methods
 - 2. Increase the number of scenarios
- In case approximating $\mathcal{F}(x_k^\star,\eta)$ is not feasible, cross-testing can be employed. Let

$$\overline{\mathcal{F}} = \{ \mathcal{F}(x_k^{\star}, \xi_l) \}_{k, l=1, \dots, n: k \neq l} .$$

Out-of-sample stability implies that the standard deviation of $\overline{\mathcal{F}}$ is close to 0.

Final considerations

Some practical advice:

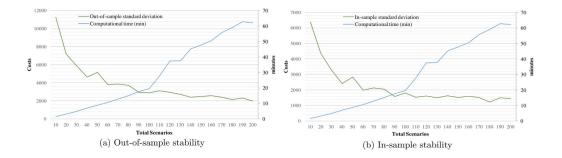
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For a rigorous treatment of stability, check [Dupačová, 1990, Schultz, 2000, Heitsch et al., 2006].

Stability for scenario trees [Dillon et al., 2017]



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Scenario generation methods

The main types of scenario-generation methods are:

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- 1. **Sampling:** Monte-Carlo sampling, or quasi Monte-Carlo sampling using variance reduction techniques (e.g., Sobol sequences). Combined with Sample Average Approximation (SAA).
- Moment matching: artificially generates a set of scenarios with the same (four plus correlation, usually) moments as the desired distribution;
- Metric-based: form smaller scenario sets whilst minimising some probabilistic distance metric. Includes clustering (k-means and related methods) and scenario reduction.

Moment matching

Build a scenario tree $\xi=\{(z_s,p_s)\}_{s\in[S]}$ that has statistical moments $f_m(z,p)$ matching M_m^{VAL} target values.

- Moments extracted from the original distribution, or data;
- ▶ The following problem must be solved ([Høyland and Wallace, 2001]):

$$\begin{split} & \min_{z,p \geq 0} \sum_{m \in M} w_m (f_m(z,p) - M_m^{\text{VAL}})^2 \\ & \text{s.t.: } \sum_{j=1}^S p_j = 1, \end{split}$$

where w_m are weights.

Remark: [Høyland et al., 2003] show how the above problem can be heuristically solved.

Metric-based methods

Probability-metric based methods use the following result [Pflug, 2001]

$$e(\eta, \xi_k) \le Kd(\eta, \xi_k)$$

where K is a (Lipschitz-related) constant and d is a Wasserstein distance between η and ξ_k . Thus, the focus is on obtaining trees that minimise d.

Metric-based methods

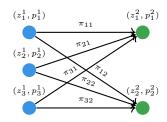
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Let $\xi^l=(z^l,p^l)\in\Xi^l$. The (p-order) Wasserstein distance $d(\xi^1,\xi^2)$ is given by:

$$\begin{split} & \underset{\pi}{\text{min.}} & \sum_{i \in \xi^1, j \in \xi^2} ||z_i^1 - z_j^2||_p \pi_{ij} \\ & \text{s.t.:} & \sum_{j \in \xi^2} \pi_{ij} = p_i^1, \ \forall i \in \xi_1 \\ & \sum_{i \in \xi^1} \pi_{ij} = p_j^2, \ \forall j \in \xi_2. \end{split}$$



Metric-based methods

1. "Clustering-like" methods:

- k-means, and variants incorporating Wasserstein distance as the metric [Condeixa et al., 2020]
- ▶ Work well in case scenarios are generated from data [Kaut, 2021];
- ▶ [Löhndorf, 2016]: Learning-based algorithms such as competitive learning and Voronoi cell sampling as alternatives to k-means.

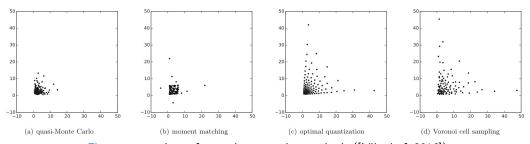


Figure: comparison of scenario generation methods ([Löhndorf, 2016])

Metric-based methods

- 2. Scenario reduction methods: Obtain ξ^2 from ξ^1 where $|\xi^2| < |\xi^1|$.
 - Based on the theory of stability of stochastic programs [Römisch, 2003]
 - Changes in the solution can be approximated using a Fortet-Mourier-type metric
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 - Calculation leads to a Monge-Kantorovich mass transportation problem
 - "Historical" chronology:
 - 1. [Dupačová et al., 2003, Heitsch and Römisch, 2003]: first backward reduction and forward selection methods;
 - 2. [Heitsch and Römisch, 2007] improved versions of the heuristics;
 - 3. [Heitsch and Römisch, 2009] The above does not work for multi-stage problems. Provides a method that does.

Scenario reduction

Types of reduction algorithms. Let K be a target value for $|\xi^2|$

- **Backward reduction:** repeat until $|\xi^2| = K$. Start from ξ^1
 - 1. Find the scenario whose removal causes the smallest error increase
 - 2. Remove the scenario and redistribute its probability
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Some final practical remarks:

- ▶ In [Heitsch and Römisch, 2003], their results indicate:
 - 50% of the scenarios gives 90% relative accuracy
 - 1% of the scenarios gives 50% accuracy
- **Forward selection** gives better results, but is slow for large $|\xi^1|$ and K.
- ► Scenred2 (GAMS) is an available implementation.

Some of my own experience

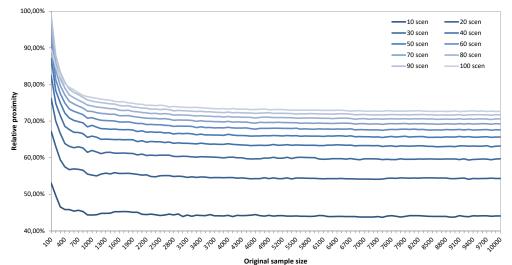


Figure: Relative accuracy for scenario reduction; x-axis is $|\xi^1|$, lines are different $|\xi^2|$. [Oliveira et al., 2016]

Some of my own experience

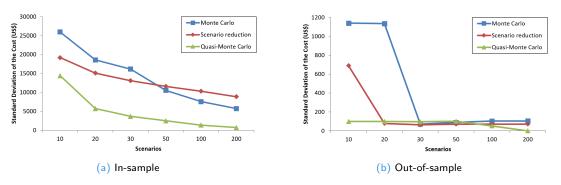


Figure: Objective function standard deviation comparing 3 alternative scenario reduction methods. Original sample had 1000 scenarios [Fernández Pérez et al., 2018]

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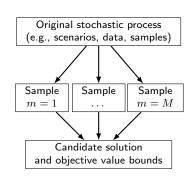
Scenario (tree) generation methods

Sample Average Approximation (SAA)

What is SAA

SAA [Shapiro and Homem-de Mello, 1998] is an alternative to generating scenario trees in the context of stochastic programming

- Purely based on sampling;
- Monte Carlo simulation for estimating objective function bounds;
- Useful for handling large scenario sets;
- The sample m scenario tree size N is such that $N << |\xi|$ or $|\eta|$;
- Requires solving M problems.



SAA is based on the law of large numbers (LLN) and the central limit theorem (CLT). As such, we can

- Estimate bounds using mean values;
- Estimate confidence intervals.

 $^{^{1}}f(x)$ is a shorthand for $\mathcal{F}(x,\xi)$.

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First, let us define our notation for 2SSPs

$$z = \min_{x} f(x),$$

where:

- $f(x) = \mathbb{E}_{\xi} [F(x,\xi)]^{1}$
- $F(x,\xi) = \{c^{\top}x + Q(x,\xi) : x \in X\};$
- $Q(x,\xi) = \min_{y} \{ q(\xi)^{\top} y : W(\xi) y = h(\xi) T(\xi) x, y \ge 0 \};$
- $X = \{x \in \mathbb{R}^n : Ax = b, x \ge 0\}.$

 $^{^{1}}f(x)$ is a shorthand for $\mathcal{F}(x,\xi)$.

Calculating a lower bounds for z

Let N be the number of samples we draw from our original stochastic process, forming the scenario tree $\xi = \{\xi_1, \dots, \xi_N\}$.

Then, we can solve the sample-based approximation problem

$$\hat{z}_N = \min_{x} \left\{ \tilde{f}_N(x) = \frac{1}{N} \sum_{n=1}^N F(x, \xi_n) \right\}.$$
 (1)

 $^{^2}$ LLN: $\lim_{N o\infty}\mathbb{E}\left[rac{\sum_{n=1}^N X_n}{N}
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 (1)

First, notice that $\tilde{f}_N(x)$ is an unbiased estimator² for f(x):

$$\mathbb{E}_{\xi} \left[\tilde{f}_{N}(x) \right] = \frac{1}{N} \mathbb{E}_{\xi} \left[\sum_{n=1}^{N} F(x, \xi_{n}) \right] \xrightarrow{LLN} \frac{1}{N} (Nf(x)) = f(x). \quad \Box$$

 2 LLN: $\lim_{N \to \infty} \mathbb{E}\left[\frac{\sum_{n=1}^N X_n}{N}\right] = \frac{N\overline{X}}{N} = \overline{X}$ for i.i.d. random variable X_n with mean value \overline{X} .

Calculating lower bounds for z

We now show that $\mathbb{E}\left[\hat{z}_{N}\right]$ is a lower bound on z:

$$\hat{z}_{N} = \min_{x} \left\{ \frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n}) \right\} \leq \frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n})$$

$$\mathbb{E}_{\xi} \left[\min_{x} \left\{ \frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n}) \right\} \right] \leq \mathbb{E}_{\xi} \left[\frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n}) \right]$$

$$\mathbb{E}_{\xi} \left[\hat{z}_{N} \right] \leq \mathbb{E}_{\xi} \left[\frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n}) \right]$$

$$\mathbb{E}_{\xi} \left[\hat{z}_{N} \right] \leq \min_{x} \left\{ \mathbb{E}_{\xi} \left[\frac{1}{N} \sum_{n=1}^{N} F(x, \xi_{n}) \right] \right\} \xrightarrow{N \to \infty}$$

$$\min_{x} \left\{ \mathbb{E}_{\xi} \left[F(x, \xi) \right] \right\} = \min_{x} f(x) = z. \quad \Box$$

Calculating lower bounds for z

In turn, we can approximate $\mathbb{E}\left[\hat{z}_{N}\right]$ using a sample estimate.

1. For that, we sample M scenario trees of size N:

$$\{\xi_1^1, \dots, \xi_N^1\}, \dots, \{\xi_1^M, \dots, \xi_N^M\}.$$

³Again an unbiased estimator, see footnote 2.

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2. For each scenario tree, we solve

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2. For each scenario tree, we solve

$$\hat{z}_N^m = \min_{x} \left\{ \frac{1}{N} \sum_{n=1}^N F(x, \xi_n^m) \right\}$$

3. We can then estimate $\mathbb{E}\left[\hat{z}_{N}\right]$ as

$$L_N^M = \frac{1}{M} \sum_{m=1}^M \hat{z}_N^m.$$

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Statistical bounds for ${\cal L}_N^M$

We can use the CLT to provide confidence intervals for L_N^M . A sample-estimate for $\sigma_{L_N^M}^2$ can be obtained as

$$s_{L_N^M}^2 = \frac{1}{M-1} \sum_{m=1}^M (\hat{z}_N^m - L_N^M)^2.$$

Statistical bounds for L_N^M

We can use the CLT to provide confidence intervals for L_N^M . A sample-estimate for $\sigma_{L_N^M}^2$ can be obtained as

$$s_{L_N^M}^2 = \frac{1}{M-1} \sum_{m=1}^M (\hat{z}_N^m - L_N^M)^2.$$

We can use $s^2_{L^M_N}$ to obtain an 1- α confidence interval for L^M_N :

$$\left[L_{N}^{M} - \frac{z_{\alpha/2} s_{L_{N}^{M}}}{\sqrt{M}}, L_{N}^{M} + \frac{z_{\alpha/2} s_{L_{N}^{M}}}{\sqrt{M}} \right]$$

where $z_{\alpha/2}$ is the standard normal $1 - \alpha/2$ quantile.

Calculating upper bounds for z

Let

$$\hat{x}_N^m = \underset{x}{\operatorname{argmin}} \left\{ \frac{1}{N} \sum_{n=1}^N F(x, \xi_n^m) \right\}, \ \forall m \in [M].$$

Under a relatively complete recourse assumption, we have that $f(\hat{x}_N^m) \geq z$, $\forall m \in [M]$.

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Under a relatively complete recourse assumption, we have that $f(\hat{x}_N^m) \geq z$, $\forall m \in [M]$.

We can obtain an unbiased estimate for $f(\hat{x}_N^m)$ by

- 1. Choosing one solution $\hat{x}_N^{m'}$, $m' \in [M]$;
- 2. Sampling T scenario trees of size \overline{N}

$$\left\{\xi_1^1, \dots, \xi_{\overline{N}}^1\right\}, \dots, \left\{\xi_1^T, \dots, \xi_{\overline{N}}^T\right\}$$

3. For each scenario tree t, we evaluate

$$\check{z}_{\overline{N}}^t = \frac{1}{\overline{N}} \sum_{n=1}^N F(\hat{x}_N^{m'}, \xi_n^t)$$

Calculating upper bounds for z

4. We can estimate $f(\hat{x}_N^m)$ as

$$U_{\overline{N}}^{T} = \frac{1}{T} \sum_{t=1}^{T} \check{z}_{\overline{N}}^{t}.$$

Analogously, we can use the sample-estimate for $\sigma^2_{U_{\overline{N}}^T}$

$$s_{U_{\overline{N}}}^2 = \frac{1}{T-1} \sum_{t=1}^{T} (\check{z}_{\overline{N}}^t - U_{\overline{N}}^T)^2$$

$$\left[U_{\overline{N}}^T - \frac{z_{\alpha/2} s_{U_{\overline{N}}}}{\sqrt{T}}, U_{\overline{N}}^T + \frac{z_{\alpha/2} s_{U_{\overline{N}}}}{\sqrt{T}}\right].$$

In this context, an optimality gap refers to the quantity

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On the other hand, we know that

$$\mathbb{E}\left[\hat{z}_N\right] \le z \le f(\hat{x}_N^{m'}).$$

Since we have estimates for $\mathbb{E}\left[\hat{z}_N\right]$ (L_N^M) and $f(\hat{x}_N^{m'})$ $(U_{\overline{N}}^T)$, we can calculate the optimality gap estimate

$$gap(N,M,\overline{N},T) = U_{\overline{N}}^T - L_N^M.$$

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$$gap(N, M, \overline{N}, T) = U_{\overline{N}}^T - L_N^M.$$

Confidence intervals can also be obtained for $gap(N, M, \overline{N}, T)$ using

$$\sigma_{gap(N,M,\overline{N},T)}^2 = s_{L_N^M}^2 + s_{U_{\overline{N}}}^2.$$

Some remarks on $gap(N, M, \overline{N}, T)$:

 $ightharpoonup gap(N,M,\overline{N},T)$ is a biased estimator, since

$$f(\hat{x}_N^{m'}) - \mathbb{E}\left[\hat{z}_N\right] \ge f(\hat{x}_N^{m'}) - z;$$

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 - 1. $s_{L_N^M}^2$, via increasing N and M: larger N leads to larger problems, but they can be solved as M parallel problems;

On estimating optimality gaps

Some remarks on $gap(N, M, \overline{N}, T)$:

▶ $gap(N, M, \overline{N}, T)$ is a biased estimator, since

$$f(\hat{x}_N^{m'}) - \mathbb{E}\left[\hat{z}_N\right] \ge f(\hat{x}_N^{m'}) - z;$$

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- ▶ Confidence intervals for $gap(N, M, \overline{N}, T)$ can be improved by reducing:
 - 1. $s_{L_N}^2$, via increasing N and M: larger N leads to larger problems, but they can be solved as M parallel problems;
 - 2. $s^2_{U^T_{\overline{N}}}$, via increasing \overline{N} and T; larger \overline{N} leads to more costly evaluation; solvable as T (as $\overline{N} \times T$ for 2SSPs) parallel problems.

Regarding choosing a solution $\hat{x}_N^{m'}$:

- If feasible, evaluate all distinct solutions \hat{x}_N^m for $m \in [M]$ and choose that with best L_N^M , $U_{\overline{N}}^T$ or $qap(N, M, \overline{N}, T)$;
- Too many distinct solutions may indicate that N is too small. Perform stability analysis.
- ➤ SAA holds for non-independent sampling schemes (e.g., Latin hypercube sampling or quasi Monte Carlo). These help keep N small.

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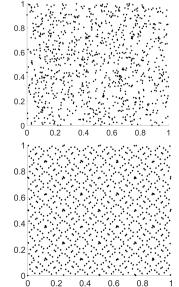


Figure: Monte Carlo (top) and quasi-Monte Carlo (bottom) sampling [Fernández Pérez et al., 2018]

Regarding the choice of N [Oliveira and Hamacher, 2012]:

Notice that \hat{z}_N is the expected value of the random variable

$$z_N(\xi) = F(\hat{x}_N, \xi), \text{ where } \hat{x}_N = \underset{x}{\operatorname{argmin}} \left\{ \frac{1}{N} \sum_{n=1}^N F(x, \xi_n) \right\}$$

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As such, we can estimate its sample-based variance and a $1-\alpha$ confidence interval, given by

$$s_N^2 = rac{1}{N-1} \sum_{n=1}^N (\hat{z}_N - z_N(\xi_n))^2 ext{ and } \hat{z}_N \pm rac{z_{lpha/2} s_N}{\sqrt{N}}.$$

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If we predefine a desired relative width β for the confidence interval, we can infer that

$$N \ge \left(\frac{z_{\alpha/2}s_N}{(\beta/2)\hat{z}_N}\right)^2.$$

Tutorial 3

SAA example

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