CA3

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Question 1:

Let X $N_p(\mu, \Sigma)$ with $|\Sigma| > 0$. Using the spectral decomposition of the covariance matrix, prove that

$$(X-\mu)'\Sigma^{-1}(X-\mu)$$
 χ_p^2

1. Start with $\Sigma e_i = \lambda_i e_i$ and note that Σ is positive definite.

$$\begin{split} \Sigma e_p &= \lambda_p e_p \qquad (1) \\ \Sigma e_p &= [\Sigma e_1 \Sigma e_p] \\ &[\lambda_i e_1 \lambda_p e_p] \\ &[\lambda_i e_1 \lambda_p e_p] = e_p D_p \end{split}$$

$$\begin{split} \Sigma e_p &= e_p D_p \\ \Sigma^{-1} &= e_p \ D_p^{-1} \ e_p^{-1} \\ \Sigma^{-1} &= e_p \ D_p^{-1} \ e_p' \\ \Sigma^{-1} &= \sum_{i=1}^p \ e_i \ \lambda_i^{-1} \ e_i'(2) \end{split}$$

2. Using this, rewrite $(X-\mu)'\Sigma^{-1}$ $(X-\mu)$ as the sum of p squared variables

$$(X - \mu)' \Sigma^{-1} (X - \mu)$$

$$= (X - \mu)' \left(\sum_{i=1}^{p} e_i \lambda_i^{-1} e_i' \right) (X - \mu) \qquad (2)$$

$$= \sum_{i=1}^{p} \lambda_i^{-1} (X - \mu)' e_i e_i' (X - \mu)$$

$$= \sum_{i=1}^{p} \lambda_i^{-1} [e_i' (X - \mu)]^2$$

$$= \sum_{i=1}^{p} \left[\frac{1}{\sqrt{\lambda_i}} e_i' (X - \mu) \right]^2$$

3. Showing that this vector i.e. $\left[\frac{1}{\sqrt{\lambda_i}} e_i'(X-\mu)\right]$ is a multivariate normal with zero mean and identity covariance

$$A \ X \sim N_q(A\mu, \ A\Sigma A')$$

$$[\frac{1}{\sqrt{\lambda_i}}e_i']~[X-\mu] \sim N[0,[(\frac{1}{\sqrt{\lambda_i}}e_i')~\Sigma~(\frac{1}{\sqrt{\lambda_i}}e_i')']]$$

Firstly, since X- μ is centered, we know that the <u>Mean</u> of this vector is equal to zero. Secondly, to show that the Variance of this vector is equal to the identity matrix ...

$$\begin{split} &(\frac{1}{\sqrt{\lambda_i}}e_i') \; \Sigma \; (\frac{1}{\sqrt{\lambda_i}}e_i')' \\ &= \frac{1}{\sqrt{\lambda_i}}e_i' \; \Sigma \; \frac{1}{\sqrt{\lambda_i}}e_i \\ &= \frac{1}{\lambda_i}e_i' \; \Sigma \; e_i \\ &= \frac{1}{\lambda_i}e_i' \; \lambda_i \; e_i \quad (1) \\ &= I \end{split}$$

Remembering that e_i' $e_i = I$ because the eigen vectors are orthogonal to each other. Therefore

$$[\frac{1}{\sqrt{\lambda_i}}e_i'] \; [X-\mu] = Z_i \sim N[0,I]$$

4. Finally showing Q.E.D

$$(X-\mu)'\Sigma^{-1}\ (X-\mu) \quad = \sum_{i=1}^p [\frac{1}{\sqrt{\lambda_i}}\ e_i'\ (X-\mu)]^2 \quad = \sum_{i=1}^p Z_i^2 \sim \chi_p^2$$

Question 2

$$\begin{bmatrix} x1\\x2\\x3\\x4\\x5 \end{bmatrix} \sim N_5 (\begin{bmatrix} 5\\0\\-2\\6\\2 \end{bmatrix}, \begin{bmatrix} 8&3&-1&0&5\\3&12&2&2&-2\\-1&2&9&0&1\\0&2&0&8&2\\5&-2&1&2&10 \end{bmatrix})$$

Now define $X1 = \begin{bmatrix} x1 \\ x2 \\ x4 \end{bmatrix}$ and $X2 = \begin{bmatrix} x3 \\ x5 \end{bmatrix}$. Find the conditional (joint) distribution of $X1|X2 = \begin{bmatrix} -1 \\ 2 \end{bmatrix}$

1. First change the matrix around

$$\begin{bmatrix} x1\\x2\\x4\\x3\\x5 \end{bmatrix} \sim N_5 (\begin{bmatrix} 5\\0\\6\\-2\\2 \end{bmatrix}, \begin{bmatrix} 8&3&0&|&-1&5\\3&12&2&|&2&-2\\0&2&8&|&0&2\\-1&2&0&|&9&1\\5&-2&2&|&1&10 \end{bmatrix})$$

$$X1|X2 = \begin{bmatrix} -1 \\ 2 \end{bmatrix} \sim N(E[X1|X2], cov(X1, x2))$$

$$\begin{split} \mu_1 &= \begin{bmatrix} 5 \\ 0 \\ 6 \end{bmatrix} \\ \mu_2 &= \begin{bmatrix} -2 \\ 2 \end{bmatrix} \\ \Sigma_{11} &= \begin{bmatrix} 8 & 3 & 0 \\ 3 & 12 & 2 \\ 0 & 2 & 8 \end{bmatrix} \\ \Sigma_{12} &= \begin{bmatrix} -1 & 5 \\ 2 & -2 \\ 0 & 2 \end{bmatrix} \\ \Sigma_{21} &= \begin{bmatrix} -1 & 2 & 0 \\ 5 & -2 & 2 \end{bmatrix} \end{split}$$

$$\begin{split} \Sigma_{22} &= \begin{bmatrix} 9 & 1 \\ 1 & 10 \end{bmatrix} \\ \Sigma_{21} &= \begin{bmatrix} -1 & 2 & 0 \\ 5 & -2 & 2 \end{bmatrix} \\ X_2 &= \begin{bmatrix} -1 \\ 2 \end{bmatrix} \end{split}$$

2. Calculate the E[X1|X2]

$$E[X1|X2] = \mu_1 + \Sigma_{12} \ \Sigma_{22}^{-1}(X_2 - \ \mu_2)$$

[,1]

[1,] 4.831

[2,] 0.247

[3,] 5.978

3. Calculate cov[X1, X2]

$$Cov(X1,X2) = \Sigma_{11} - \Sigma 12 \ \Sigma_{22}^{-1} \ \Sigma_{21}$$