

Meta-Analysis Reproduction Summary Excluding HUM251

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1. Summary

This document reports the process taken in the model fitting stage of the meta-analysis in thermal reproduction.

2. Setup

We first read in our data and select all of the effect sizes related to reproduction. We do this using the following code.

```
### Read in effect size data
effectdata <- read.csv("Data/Survival project all pairwise.es.csv")
repdata_warm <- subset(effectdata, Trait.category == "Reproduction" &
  warm.cool == "Warm")
repdata_cool <- subset(effectdata, Trait.category == "Reproduction" &
  warm.cool == "Cool")

allrep <- rbind(repdata_warm, repdata_cool)

### select data for analysis
rdata <- allrep
```

Next we create new columns in our dataframe which will serve as random factors in our multi-level meta analysis models. The following initialises four new columns, namely “obs”, “study_code”, “Species.phylo” and “species”. Lastly, we create a column name “precision” which is equal to the inverse standard error.

```
### Create random factors into data frame
rdata$obs <- factor(c(1:nrow(rdata))) # Unique observation code
rdata$study_code <- factor(rdata$Paper.code) # Model requires column names study_code
rdata$Species.phylo <- factor(rdata$Species.latin) # Species names for phylo matrix
rdata$species <- factor(rdata$Species.latin) # Another species column for random factor

precision <- sqrt(1/rdata$v) # inverse standard error
rdata[, "precision"] <- precision
```

The number of species and total number of studies present in the data are as follows.

```
nlevels(rdata$species) # Check number of species

## [1] 306

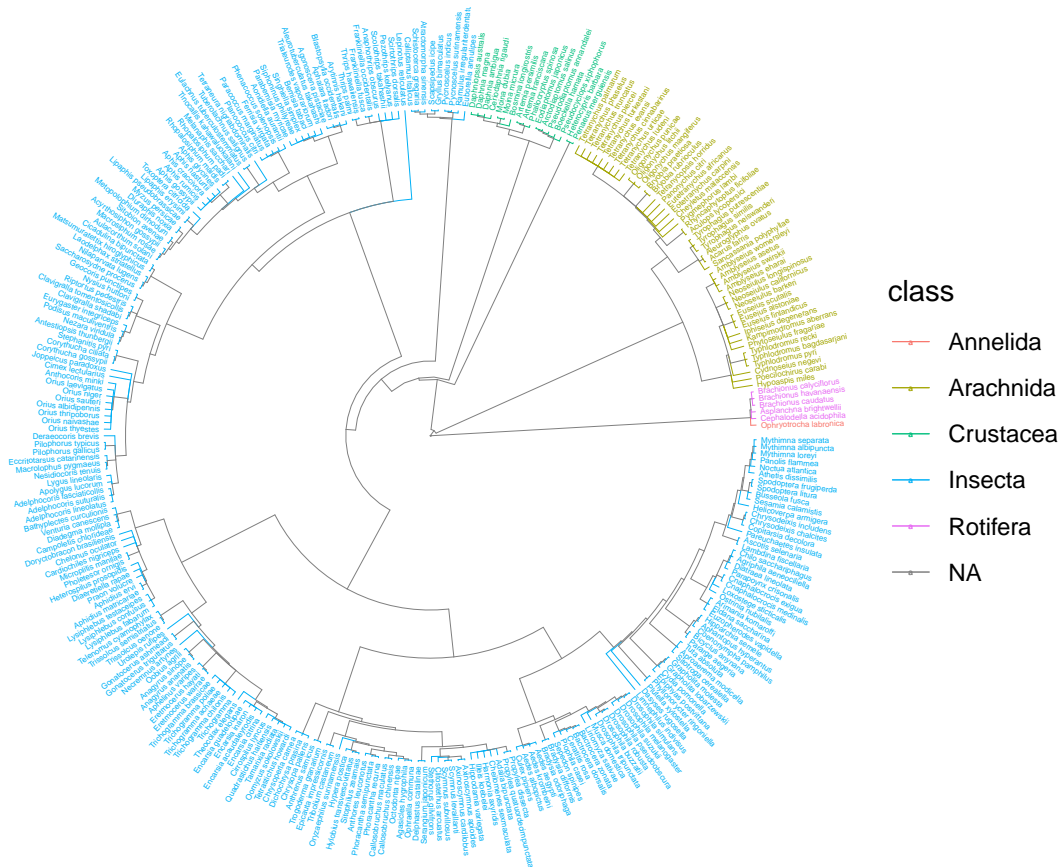
nlevels(rdata$study_code) # Check number of studies

## [1] 339
```

The final stage in the setup is to import a phylogentic tree of the data. Below is the code used to produce the tree and a plot of the tree itself.

```
## import tree from map
tree1 <- read.nexus("all_reproduction_exchUM251_tree.nex")
tree_grafen = compute.brlen(tree1, method = "Grafen", power = 1)
phylo_matrix <- vcv(tree_grafen, cor = TRUE, model = "Brownian") # Make phylogenetic matrix

## character(0)
```



3. Random effects models

In this section we determine which random effects to include in our model. For each model I have provided the code used to specify the structure of the model and a summary of the results. We begin with a model that includes all of the random factors we created earlier.

```
# Adding four random factors
meta2 <- rma.mv(es, v, random = list(~1 | Species.phylo, ~1 |
  species, ~1 | study_code, ~1 | obs), test = "t", dfs = "contain",
  R = list(Species.phylo = phylo_matrix), data = rdata, method = "REML")

summary(meta2)

##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -4048.7191    8097.4382    8107.4382    8133.6090    8107.4816
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor      R
## sigma^2.1  0.0000  0.0000   306    no Species.phylo  yes
## sigma^2.2  0.0000  0.0003   306    no      species    no
## sigma^2.3 10.4420  3.2314   339    no  study_code    no
## sigma^2.4  5.2505  2.2914  1387    no      obs      no
##
## Test for Heterogeneity:
## Q(df = 1386) = 37155.0413, p-val < .0001
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## -1.5543  0.1907  -8.1485  <.0001  -1.9282  -1.1805  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

i2_ml(meta2, method = c("ratio")) # Heterogeneity at each random factor level

##      I2_Total I2_Species.phylo      I2_species      I2_study_code      I2_obs
##      9.946178e+01      7.043695e-10      7.683821e-07      6.618307e+01      3.327870e+01
```

Accounting for non-independence of data points from the same experiment

The data has a nested structure. Each study (study_code) may have a number of experiments (effect.size.code) which share a common control temperature. Each effect size has its own unique code, obs. Effect sizes from the same experiment which share a control temperature are thought to be non-independent. The following code create a covariance matrix “VCV_shared” which assumes a correlation of 0.5 between effect sizes from the same experiment. We include this structure in our proceeding models.

```
rdata$shared_control <- factor(rdata$Effect.size.code)
VCV_shared <- impute_covariance_matrix(vi = rdata$v, cluster = rdata$shared_control,
  r = 0.5)

# Add new variance matrix into the mixed-effects
# meta-analysis model
```

```
meta3 <- rma.mv(es, VCV_shared, random = list(~1 | Species.phylo,
~1 | species, ~1 | study_code, ~1 | obs), test = "t", dfs = "contain",
R = list(Species.phylo = phylo_matrix), data = rdata, method = "REML")
```

```
summary(meta3)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3770.1104    7540.2208    7550.2208    7576.3917    7550.2643
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor      R
## sigma^2.1  0.0000  0.0000   306    no Species.phylo  yes
## sigma^2.2  0.0000  0.0002   306    no      species    no
## sigma^2.3  5.1474  2.2688   339    no  study_code    no
## sigma^2.4  4.4602  2.1119  1387    no      obs      no
##
## Test for Heterogeneity:
## Q(df = 1386) = 38656.3912, p-val < .0001
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## -1.3401  0.1414  -9.4771  <.0001  -1.6172  -1.0629  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
i2_ml(meta3, method = c("ratio")) # Heterogeneity at each random factor level
```

```
##      I2_Total I2_Species.phylo      I2_species      I2_study_code      I2_obs
##      9.912389e+01      1.238319e-09      6.017065e-07      5.310681e+01      4.601707e+01
```

Model without phylogeny

We also originally included a variance-covariance matrix for phylogenetic relatedness of the included species as a random effect in the model (Chamberlain et al., 2012), but this has now been removed as its inclusion did not improve model fit and the phylogenetic signal was very weak.

```
## without phylogeny but with shared control
meta5 <- rma.mv(es, VCV_shared, random = list(~1 | species, ~1 |
  study_code, ~1 | obs), test = "t", dfs = "contain", data = rdata,
  method = "REML")

summary(meta5)

##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3770.1104    7540.2208    7548.2208    7569.1575    7548.2497
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  0.0000  0.0002   306     no      species
## sigma^2.2  5.1474  2.2688   339     no  study_code
## sigma^2.3  4.4602  2.1119  1387     no        obs
##
## Test for Heterogeneity:
## Q(df = 1386) = 38656.3912, p-val < .0001
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## -1.3401  0.1414  -9.4771  <.0001  -1.6172  -1.0629  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

i2_ml(meta5, method = c("ratio")) # Heterogeneity at each random factor level

##      I2_Total      I2_species I2_study_code      I2_obs
## 9.912389e+01  3.936585e-07  5.310681e+01  4.601707e+01
```

Model without phylogeny or species

```
## without phylogeny or species
meta4 <- rma.mv(es, VCV_shared, random = list(~1 | study_code,
~1 | obs), test = "t", dfs = "contain", data = rdata, method = "REML")

summary(meta4)

##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3770.1104    7540.2208    7546.2208    7561.9233    7546.2382
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.1474  2.2688   339      no  study_code
## sigma^2.2  4.4602  2.1119  1387      no      obs
##
## Test for Heterogeneity:
## Q(df = 1386) = 38656.3912, p-val < .0001
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## -1.3401  0.1414  -9.4771  <.0001  -1.6172  -1.0629  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

i2_ml(meta4, method = c("ratio")) # Heterogeneity at each random factor level

##      I2_Total I2_study_code      I2_obs
##      99.12389      53.10681      46.01708
```

Model without phylogeny, species or study_code

```
## without phylogeny, species or study_code
meta7 <- rma.mv(es, VCV_shared, random = list(~1 | obs), test = "t",
  dfs = "contain", data = rdata, method = "REML")
```

```
summary(meta7)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3926.3572    7852.7144    7856.7144    7867.1827    7856.7230
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed  factor
## sigma^2    7.0479  2.6548   1387     no     obs
##
## Test for Heterogeneity:
## Q(df = 1386) = 38656.3912, p-val < .0001
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## -1.2769  0.0743  -17.1774  <.0001  -1.4226  -1.1312  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

i2_ml(meta7, method = c("ratio")) # Heterogeneity at each random factor level

## I2_Total  I2_obs
## 98.80947 98.80947
```

We can see from the above that the best fitting model according to AIC is “meta4” which includes only the study code and the unique effect size code, obs. There is a AIC difference of 4 between the model meta4 and the next best model meta5 . We continue our analysis using meta4 as our base model.

4. Meta-regressions

Starting with the best fitting random-effect model from Section 3, “meta8” we now include single factors as a fixed effect. We initially explore the fixed factors

- **reftemp**: The experiment’s control (reference) temperature.
- **treattemp**: The treatment temperature
- **warm.cool**: A categorical variable indicating whether treatment is warmer or cooler than the reference temperature
- **diff**: The difference between the reference and treatment temperature.

Reference temperature

```
meta_trait_ref <- rma.mv(es, VCV_shared, mod = ~reftemp, random = list(~1 |  
  study_code, ~1 | obs), test = "t", dfs = "contain", data = rdata,  
  method = "REML")
```

```
summary(meta_trait_ref)
```

```
##  
## Multivariate Meta-Analysis Model (k = 1387; method: REML)  
##  
##      logLik      Deviance      AIC      BIC      AICc  
## -3766.7483    7533.4967    7541.4967    7562.4305    7541.5257  
##  
## Variance Components:  
##  
##      estim      sqrt  nlvls  fixed      factor  
## sigma^2.1  5.0944  2.2571   339     no  study_code  
## sigma^2.2  4.4713  2.1145  1387     no           obs  
##  
## Test for Residual Heterogeneity:  
## QE(df = 1385) = 38565.0269, p-val < .0001  
##  
## Test of Moderators (coefficient 2):  
## QM(df = 1) = 1.8362, p-val = 0.1754  
##  
## Model Results:  
##  
##      estimate      se      zval      pval      ci.lb      ci.ub  
## intrcpt      0.2198  1.1597   0.1896  0.8497   -2.0531   2.4927  
## reftemp     -0.0631  0.0466  -1.3551  0.1754   -0.1545   0.0282  
##  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```


Treatment temperature

```
meta_trait_treattemp <- rma.mv(es, VCV_shared, mod = ~c_treattemp,  
  random = list(~1 | study_code, ~1 | obs), test = "t", dfs = "contain",  
  data = rdata, method = "REML")
```

```
summary(meta_trait_treattemp)
```

```
##  
## Multivariate Meta-Analysis Model (k = 1387; method: REML)  
##  
##      logLik      Deviance      AIC      BIC      AICc  
## -3759.9690    7519.9380    7527.9380    7548.8718    7527.9670  
##  
## Variance Components:  
##  
##      estim      sqrt  nlvls  fixed      factor  
## sigma^2.1  5.1557  2.2706    339      no  study_code  
## sigma^2.2  4.3805  2.0930   1387      no      obs  
##  
## Test for Residual Heterogeneity:  
## QE(df = 1385) = 38615.4347, p-val < .0001  
##  
## Test of Moderators (coefficient 2):  
## QM(df = 1) = 16.5055, p-val < .0001  
##  
## Model Results:  
##  
##      estimate      se      zval      pval      ci.lb      ci.ub  
## intrcpt      -1.3584  0.1413  -9.6123  <.0001  -1.6354  -1.0814  ***  
## c_treattemp  -0.0354  0.0087  -4.0627  <.0001  -0.0524  -0.0183  ***  
##  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Warm vs Cool

```
meta_trait_warm <- rma.mv(es, VCV_shared, mod = ~warm.cool, random = list(~1 |  
  study_code, ~1 | obs), test = "t", dfs = "contain", data = rdata,  
  method = "REML")
```

```
summary(meta_trait_warm)
```

```
##  
## Multivariate Meta-Analysis Model (k = 1387; method: REML)  
##  
##      logLik      Deviance      AIC      BIC      AICc  
## -3760.9334    7521.8668    7529.8668    7550.8006    7529.8958  
##  
## Variance Components:  
##  
##      estim      sqrt  nlvls  fixed      factor  
## sigma^2.1  5.1402  2.2672   339     no  study_code  
## sigma^2.2  4.3888  2.0950  1387     no      obs  
##  
## Test for Residual Heterogeneity:  
## QE(df = 1385) = 38400.3073, p-val < .0001  
##  
## Test of Moderators (coefficient 2):  
## QM(df = 1) = 14.8227, p-val = 0.0001  
##  
## Model Results:  
##  
##      estimate      se      zval      pval      ci.lb      ci.ub  
## intrcpt      -1.0819  0.1560  -6.9363  <.0001  -1.3876  -0.7762  ***  
## warm.coolWarm  -0.5061  0.1315  -3.8500  0.0001  -0.7638  -0.2485  ***  
##  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We model warm versus cool without and intercept so we can visualise the estimates easier.

```
meta_trait_warm_nointer <- rma.mv(es, VCV_shared, mod = ~warm.cool -
  1, random = list(~1 | study_code, ~1 | obs), data = rdata,
  test = "t", dfs = "contain", method = "REML")

summary(meta_trait_warm_nointer)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3760.9334    7521.8668    7529.8668    7550.8006    7529.8958
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.1402  2.2672   339     no  study_code
## sigma^2.2  4.3888  2.0950  1387     no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1385) = 38400.3073, p-val < .0001
##
## Test of Moderators (coefficients 1:2):
## QM(df = 2) = 104.7301, p-val < .0001
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## warm.coolCool   -1.0819  0.1560   -6.9363  <.0001   -1.3876   -0.7762  ***
## warm.coolWarm   -1.5880  0.1553  -10.2230  <.0001   -1.8924   -1.2835  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Difference

```
meta_trait_diff <- rma.mv(es, VCV_shared, mod = ~diff, random = list(~1 |
  study_code, ~1 | obs), test = "t", dfs = "contain", data = rdata,
  method = "REML")
```

```
summary(meta_trait_diff)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3760.7161    7521.4323    7529.4323    7550.3661    7529.4612
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.2122  2.2830   339      no  study_code
## sigma^2.2  4.3759  2.0919  1387      no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1385) = 38574.5945, p-val < .0001
##
## Test of Moderators (coefficient 2):
## QM(df = 1) = 15.0667, p-val = 0.0001
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## intrcpt    -1.3478  0.1418  -9.5026  <.0001   -1.6258   -1.0698   ***
## diff       -0.0343  0.0088  -3.8816  0.0001   -0.0517   -0.0170   ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Treatment temperature as a quadratic effect

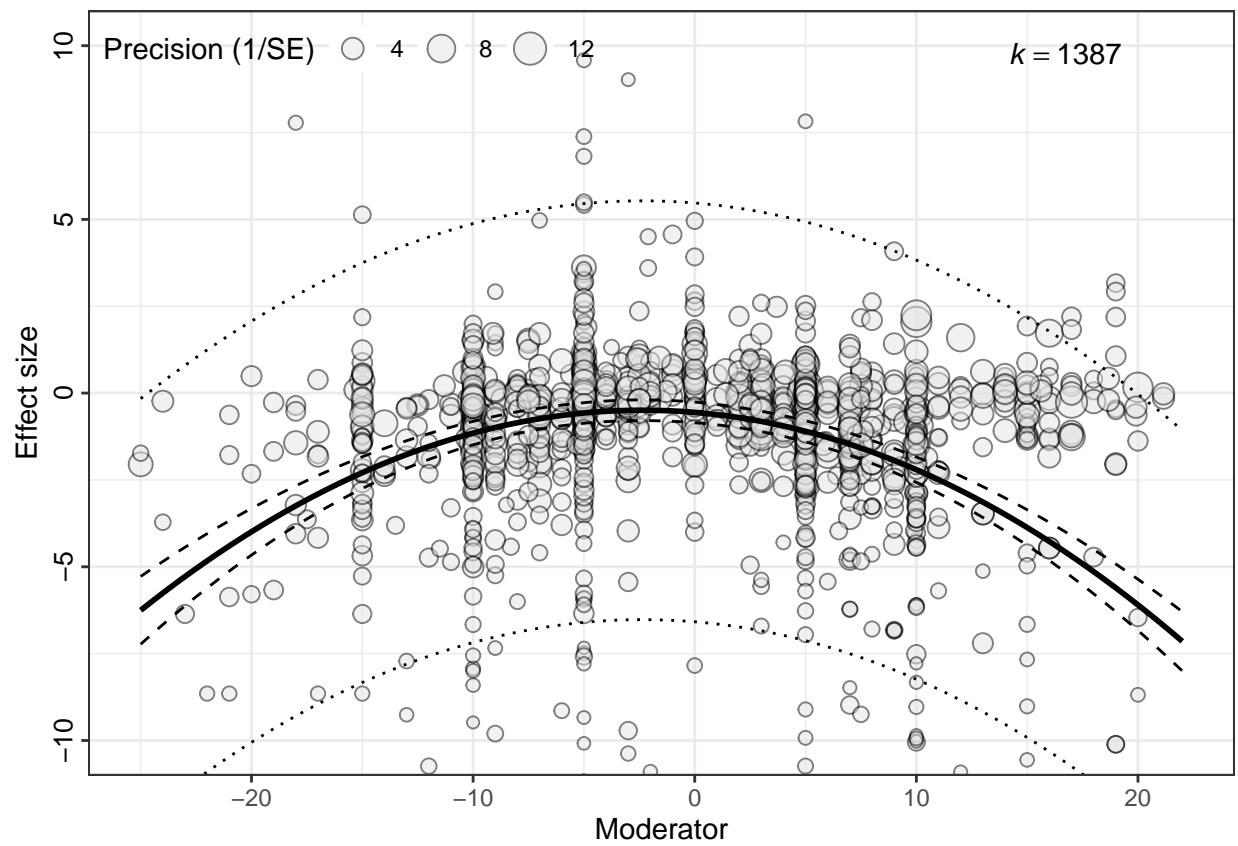
We expect that treatment temperature will have a quadratic effect on reproduction. In other words, we expect that reproduction will decrease either side of an optimum temperature for each species. We also assume that this optimum temperature will be close to the control temperature. That said, we investigate **treattemp** as a quadratic fixed effect.

```
meta_trait_treat2 <- rma.mv(es, VCV_shared, mod = ~poly(c_treattemp,
  degree = 2, raw = TRUE), random = list(~1 | study_code, ~1 |
  obs), test = "t", dfs = "contain", data = rdata, method = "REML")
```

```
summary(meta_trait_treat2)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3682.3878    7364.7755    7374.7755    7400.9392    7374.8191
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.8064  2.4096    339      no  study_code
## sigma^2.2  3.6290  1.9050   1387      no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1384) = 36417.9019, p-val < .0001
##
## Test of Moderators (coefficients 2:3):
## QM(df = 2) = 185.2754, p-val < .0001
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## intrcpt          -0.5552  0.1579   -3.5172  0.0004   -0.8646   -0.2458
## poly(c_treattemp, degree = 2, raw = TRUE)1   -0.0526  0.0082   -6.4356 <.0001   -0.0687   -0.0366
## poly(c_treattemp, degree = 2, raw = TRUE)2   -0.0112  0.0009  -12.8809 <.0001   -0.0129   -0.0095
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Below is a bubble plot of the fitted quadratic model.



Modelling response with binned treatment temperatures

The last model we will try is one with categorised or ‘binned’ treatment temperatures. We categories our effect sizes into one seven bins depending on the treatment temperature. The bins and the number of effect sizes in each bin are given below

```
## [1] "<15" ">40" "15-20" "20-25" "25-30" "30-35" "35-40"
```

```
table(rdata$bin.temp)
```

```
##
```

```
## <15 >40 15-20 20-25 25-30 30-35 35-40
```

```
## 107 77 249 307 180 339 128
```

```
meta_trait_bintemp <- rma.mv(es, VCV_shared, mod = ~bin.temp -  
1, random = list(~1 | study_code, ~1 | obs), data = rdata,  
method = "REML")
```

```
summary(meta_trait_bintemp)
```

```
##
```

```
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
```

```
##
```

```
## logLik Deviance AIC BIC AICc
```

```
## -3663.6641 7327.3282 7345.3282 7392.3967 7345.4595
```

```
##
```

```
## Variance Components:
```

```
##
```

```
## estim sqrt nlvls fixed factor
```

```
## sigma^2.1 5.5364 2.3530 339 no study_code
```

```
## sigma^2.2 3.6126 1.9007 1387 no obs
```

```
##
```

```
## Test for Residual Heterogeneity:
```

```
## QE(df = 1380) = 36020.9128, p-val < .0001
```

```
##
```

```
## Test of Moderators (coefficients 1:7):
```

```
## QM(df = 7) = 294.0378, p-val < .0001
```

```
##
```

```
## Model Results:
```

```
##
```

```
## estimate se zval pval ci.lb ci.ub
```

```
## bin.temp<15 -2.2382 0.2727 -8.2077 <.0001 -2.7727 -1.7037 ***
```

```
## bin.temp>40 -4.4663 0.3662 -12.1967 <.0001 -5.1840 -3.7486 ***
```

```
## bin.temp15-20 -1.3672 0.1891 -7.2292 <.0001 -1.7378 -0.9965 ***
```

```
## bin.temp20-25 -0.5414 0.1783 -3.0357 0.0024 -0.8909 -0.1918 **
```

```
## bin.temp25-30 -0.2847 0.2094 -1.3599 0.1739 -0.6951 0.1256
```

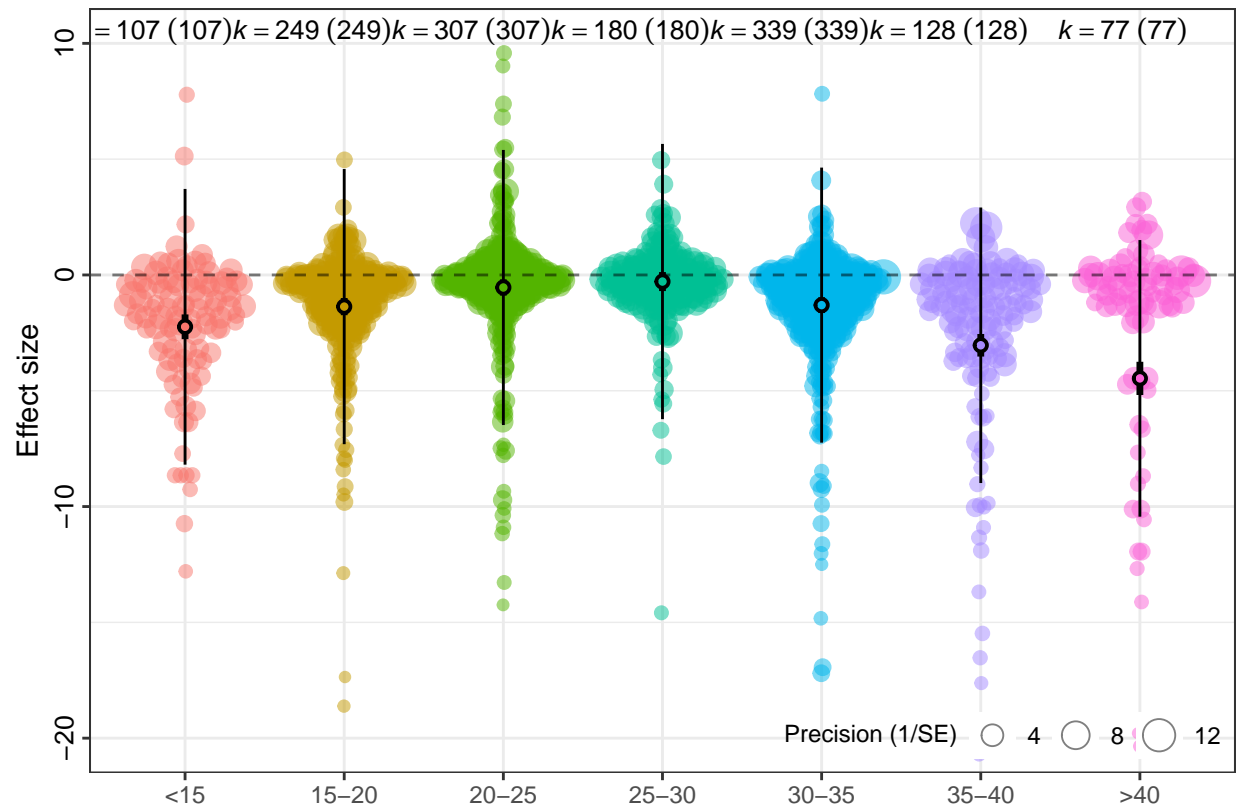
```
## bin.temp30-35 -1.3030 0.1744 -7.4720 <.0001 -1.6447 -0.9612 ***
```

```
## bin.temp35-40 -3.0364 0.2484 -12.2229 <.0001 -3.5233 -2.5495 ***
```

```
##
```

```
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



Other fixed effects

Here are tabled number of counts for each level of our moderators.

```
table(rdata$Class)
```

```
##
##  Annelida Arachnida Crustacea  Insecta  Rotifera
##        2      212      40    1107      26
```

```
table(rdata$Habitat)
```

```
##
##   Aquatic      Both Terrestrial
##      71      41    1275
```

```
table(rdata$Sex.exposed)
```

```
##
##           Both           Female      Male Parthenogenetic      Unsure
##           871           310           6           197           3
```

```
table(rdata$Fertilisation.mode)
```

```
##
## External Internal
##        2      1174
```

```
table(rdata$Agricultural.importance)
```

```
##
## Control agent      No      Pest      Vector
##        422      212      725      28
```

```
table(rdata$Lab.or.field)
```

```
##
##      Field      Lab      Mix Semi-natural
##        8     1356        6        17
```

```
table(rdata$Exposure.duration)
```

```
##
##      < 24 hours      1 to 5 days      Mix More than 5 days Natural variation
##           86           45           22           1226           8
```

```
table(rdata$Life.stage.of.animal)
```

```
##
##  Adult      Egg  Embryo Juvenile  Larvae      Mix      Pupae
##    504      14      3      52      13      773      28
```

Given the imbalance in number of effect sizes in each level, taxonomic class and habitat moderators do not seem like they are going to be useful.

The moderators with the good spreads are (providing we do some recategorisation) are: 1. Sex exposed.

2. Life stage.

Sex exposed

We could lump categories so that we have cases where males are included (Both, Male), versus cases with just females (Female, Parthenogenetic), with Unsure removed. I would predict that the 'Both' category would show the biggest drop for reproduction, but there will be no difference for lifespan

```
table(new_data$Sex.exposed)
```

```
##
##   Both Female
##   877     507
```

```
meta_treat_sex <- rma.mv(es, VCV_shared_sex, mod = ~poly(c_treattemp,
  degree = 2, raw = TRUE) * Sex.exposed, random = list(~1 |
  study_code, ~1 | obs), data = new_data, method = "REML")
```

```
summary(meta_treat_sex)
```

```
##
## Multivariate Meta-Analysis Model (k = 1384; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3655.4228    7310.8456    7326.8456    7368.6727    7326.9508
```

```
## Variance Components:
##
```

```
##           estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.8495  2.4186   338     no  study_code
## sigma^2.2  3.5461  1.8831  1384     no           obs
```

```
## Test for Residual Heterogeneity:
## QE(df = 1378) = 36058.6623, p-val < .0001
```

```
## Test of Moderators (coefficients 2:6):
## QM(df = 5) = 218.3540, p-val < .0001
```

```
## Model Results:
```

```
##
##                                     estimate      se      zval      pval      c
## intrcpt                           -0.6954  0.1951  -3.5645  0.0004  -1.
## poly(c_treattemp, degree = 2, raw = TRUE)1  -0.0242  0.0100  -2.4196  0.0155  -0.
## poly(c_treattemp, degree = 2, raw = TRUE)2  -0.0099  0.0010  -9.5679  <.0001  -0.
## Sex.exposedFemale                    0.3501  0.3154   1.1100  0.2670  -0.
## poly(c_treattemp, degree = 2, raw = TRUE)1:Sex.exposedFemale -0.0785  0.0159  -4.9443  <.0001  -0.
## poly(c_treattemp, degree = 2, raw = TRUE)2:Sex.exposedFemale -0.0040  0.0017  -2.3941  0.0167  -0.
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Life-stage

We could lump categories so that we have cases where only adults were exposed (Adult), versus cases where immature stages were exposed (Juvenile, Larvae, Pupae, Mix)- perhaps after excluding 'Egg' and 'Embryo' because these categories are a bit weird. I would predict that exposure of juveniles is worse than just exposure of adults

```
table(ls_data$Life.stage.of.animal)
```

```
##
##      Adult Immature
##      504      866
```

```
meta_treat_ls <- rma.mv(es, VCV_shared_life, mod = ~poly(c_treattemp,
  degree = 3, raw = TRUE) * Life.stage.of.animal, random = list(~1 |
  study_code, ~1 | obs), data = ls_data, method = "REML")
```

```
summary(meta_treat_ls)
```

```
##
## Multivariate Meta-Analysis Model (k = 1370; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3562.7153    7125.4306    7145.4306    7197.5977    7145.5934
##
```

```
## Variance Components:
```

```
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.7222  2.3921   338     no  study_code
## sigma^2.2  3.2170  1.7936  1370     no      obs
##
```

```
## Test for Residual Heterogeneity:
## QE(df = 1362) = 34641.6909, p-val < .0001
##
```

```
## Test of Moderators (coefficients 2:8):
## QM(df = 7) = 222.4540, p-val < .0001
##
```

```
## Model Results:
```

```
##
##                                     estimate      se      zval
## intrcpt                           -0.7501  0.2504  -2.9953
## poly(c_treattemp, degree = 3, raw = TRUE)1  -0.0005  0.0181  -0.0250
## poly(c_treattemp, degree = 3, raw = TRUE)2  -0.0071  0.0012  -5.7647
## poly(c_treattemp, degree = 3, raw = TRUE)3  -0.0001  0.0001  -1.3180
## Life.stage.of.animalImmature              0.3226  0.3084   1.0459
## poly(c_treattemp, degree = 3, raw = TRUE)1:Life.stage.of.animalImmature -0.0450  0.0229  -1.9700
## poly(c_treattemp, degree = 3, raw = TRUE)2:Life.stage.of.animalImmature -0.0082  0.0017  -4.9611
## poly(c_treattemp, degree = 3, raw = TRUE)3:Life.stage.of.animalImmature -0.0001  0.0001  -1.1485
##
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Sub-analysis on pest species

```
pest_data <- subset(rdata, Agricultural.importance == "Pest")

VCV_shared_pest <- impute_covariance_matrix(vi = pest_data$v,
  cluster = pest_data$shared_control, r = 0.5)

meta_pest <- rma.mv(es, VCV_shared_pest, mod = ~poly(c_treattemp,
  degree = 2, raw = TRUE), random = list(~1 | study_code, ~1 |
  obs), data = pest_data, method = "REML")
```

```
summary(meta_pest)
```

```
##
## Multivariate Meta-Analysis Model (k = 725; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -1636.4749    3272.9498    3282.9498    3305.8599    3283.0336
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  2.0370  1.4273   166     no  study_code
## sigma^2.2  3.0406  1.7437   725     no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 722) = 19002.3449, p-val < .0001
##
## Test of Moderators (coefficients 2:3):
## QM(df = 2) = 53.5416, p-val < .0001
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## intrcpt      -0.5542  0.1529   -3.6246  0.0003   -0.8538   -0.2545 ***
## poly(c_treattemp, degree = 2, raw = TRUE)1   -0.0260  0.0097   -2.6797  0.0074   -0.0450   -0.0070 **
## poly(c_treattemp, degree = 2, raw = TRUE)2   -0.0077  0.0011   -7.1422 <.0001   -0.0098   -0.0056 ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

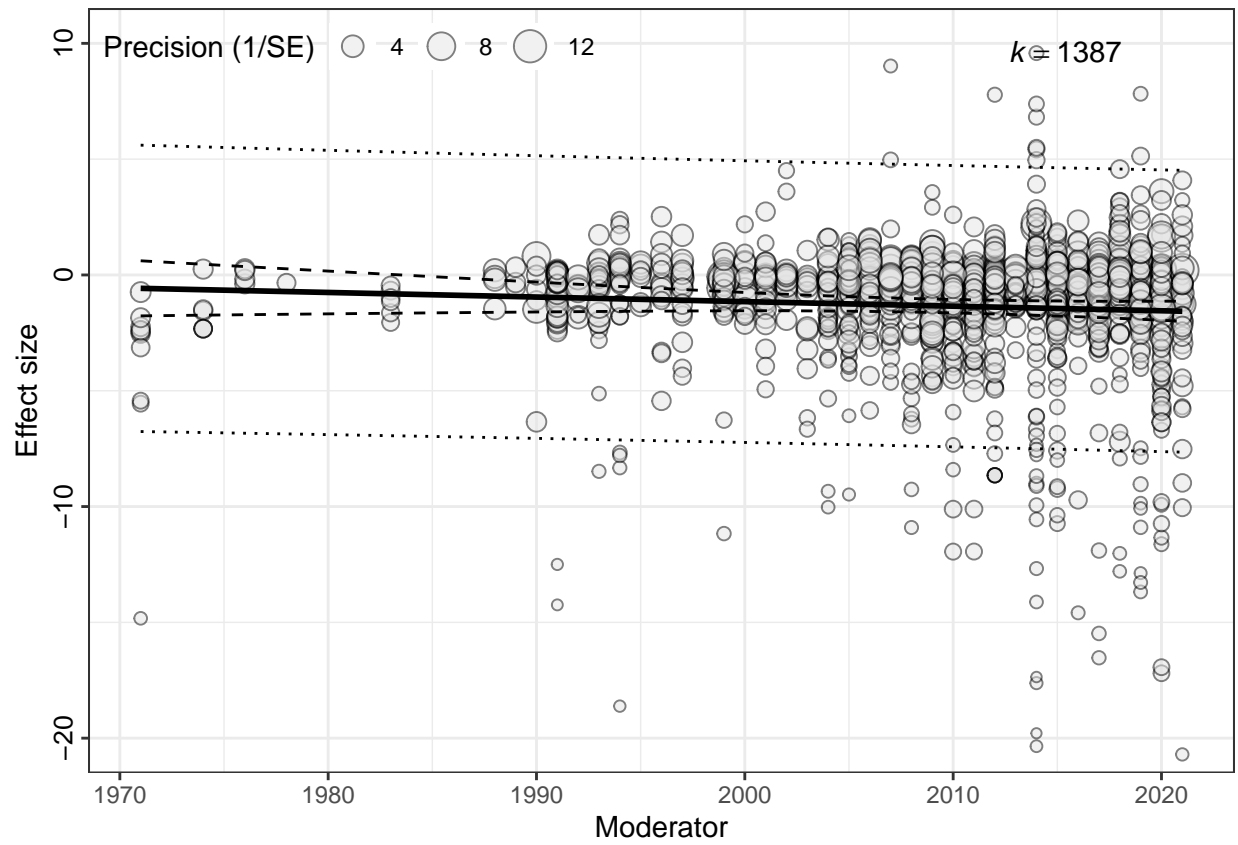
Publication Bias.

We fit meta-regression model with year as a moderator to see whether there is an publication bias. In other words, is effect size strongly correlated with publication year. We found now evidence to suggest that year has an effect on the report effect sizes. A summary of the model is given below.

```
meta_year <- rma.mv(es, VCV_shared, mod = ~Publication.year,
  random = list(~1 | study_code, ~1 | obs), data = rdata, test = "t",
  dfs = "contain", method = "REML")
```

```
summary(meta_year)
```

```
##
## Multivariate Meta-Analysis Model (k = 1387; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3766.6712    7533.3425    7541.3425    7562.2763    7541.3715
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  5.1258  2.2640    339      no  study_code
## sigma^2.2  4.4583  2.1115   1387      no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1385) = 38642.9810, p-val < .0001
##
## Test of Moderators (coefficient 2):
## QM(df = 1) = 1.6764, p-val = 0.1954
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## intrcpt      38.3496  30.6547   1.2510  0.2109  -21.7325   98.4317
## Publication.year -0.0198  0.0153  -1.2947  0.1954   -0.0496    0.0101
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



Sensitivity Analysis

Here, we perform a sensitivity analysis by removing the smallest and largest 2.5% of effect sizes.

The 2.5th and 97.5th percentiles are given below.

```
print(minq)

##      2.5%
## -14.36496

print(maxq)

##      97.5%
##  2.362038
```

The summary of the effective sizes is now

```
summary(sdata$es)

##      Min.   1st Qu.   Median     Mean   3rd Qu.     Max.
## -14.2452  -1.8325  -0.6242  -1.3739   0.0000   2.3584
```

We run the quadratic treatment temperature and the binned temperature models again with our new subsetted data i.e. only with data that is between the 2.5th and 97.5th percentile.

Treatment temperature as a quadratic effect (sensitivity analysis)

We re-create the variance-covariance matrix with our new subsetted data, which we name `sdata`. Then we run the meta analysis model again with treatment temperature as quadratic fixed effect,

```
# recreate vcv_shared matrix
VCV_shared_sa <- impute_covariance_matrix(vi = sdata$v, cluster = sdata$shared_control,
r = 0.5)
```

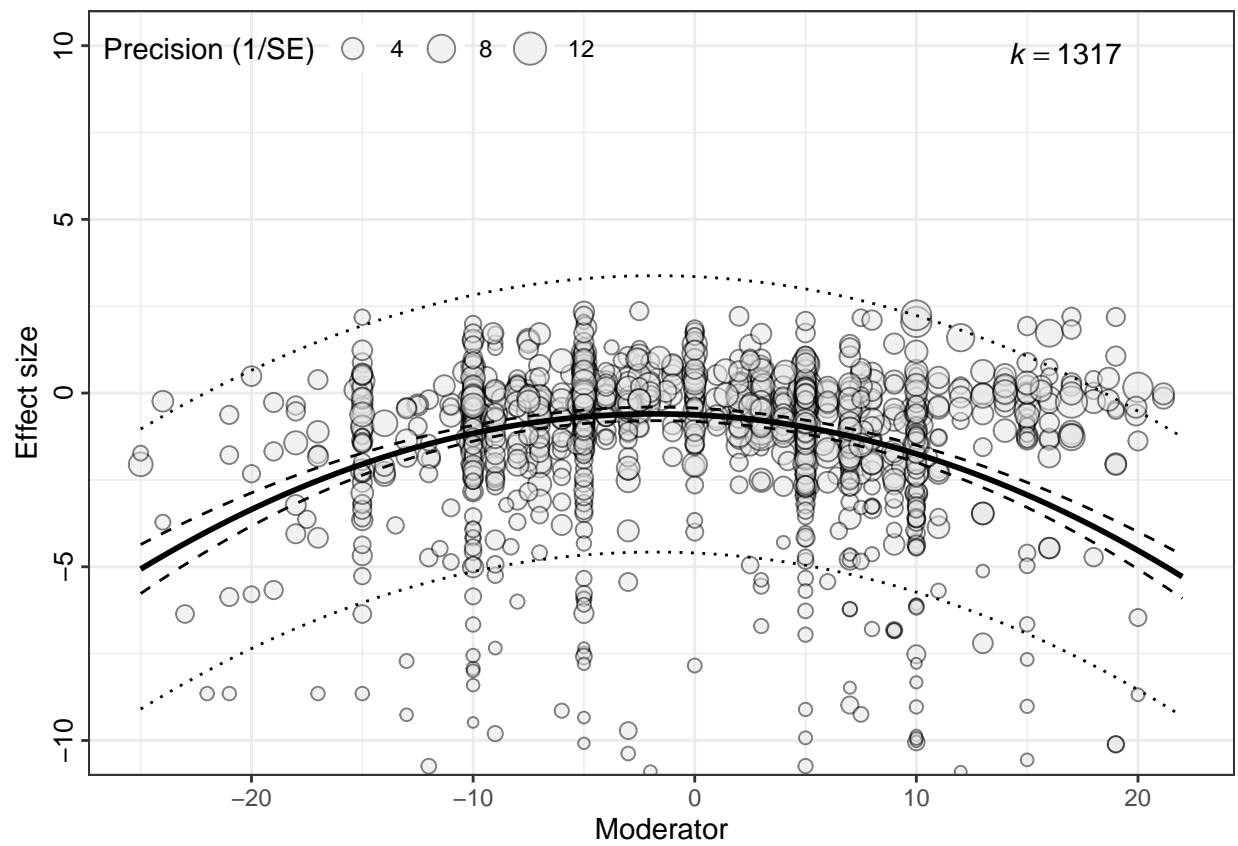
```
meta_sa_treat2 <- rma.mv(es, VCV_shared_sa, mod = ~poly(c_treattemp,
degree = 2, raw = TRUE), random = list(~1 | study_code, ~1 |
obs), test = "t", dfs = "contain", data = sdata, method = "REML")
```

```
summary(meta_sa_treat2)
```

```
##
## Multivariate Meta-Analysis Model (k = 1317; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -2688.7673    5377.5346    5387.5346    5413.4388    5387.5805
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  2.1891  1.4796   333     no  study_code
## sigma^2.2  1.9202  1.3857  1317     no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1314) = 28346.3766, p-val < .0001
##
## Test of Moderators (coefficients 2:3):
## QM(df = 2) = 176.1132, p-val < .0001
##
```

```
## Model Results:
##
##               estimate      se      zval      pval      ci.lb      ci.ub
## intrcpt          -0.6247  0.1043   -5.9910 <.0001   -0.8290  -0.4203 **
## poly(c_treattemp, degree = 2, raw = TRUE)1  -0.0296  0.0061   -4.8815 <.0001   -0.0414  -0.0177 **
## poly(c_treattemp, degree = 2, raw = TRUE)2  -0.0083  0.0006  -13.0113 <.0001   -0.0095  -0.0070 **
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```


Below is a bubble plot of the fitted quadratic model to the subsetting data.



Modelling response with binned treatment temperatures

Lastly, we rerun the model where our effect sizes are categorised into one seven bins depending on the treatment temperature. The bins and the number of effect sizes in each bin of the new subsetting data are given below

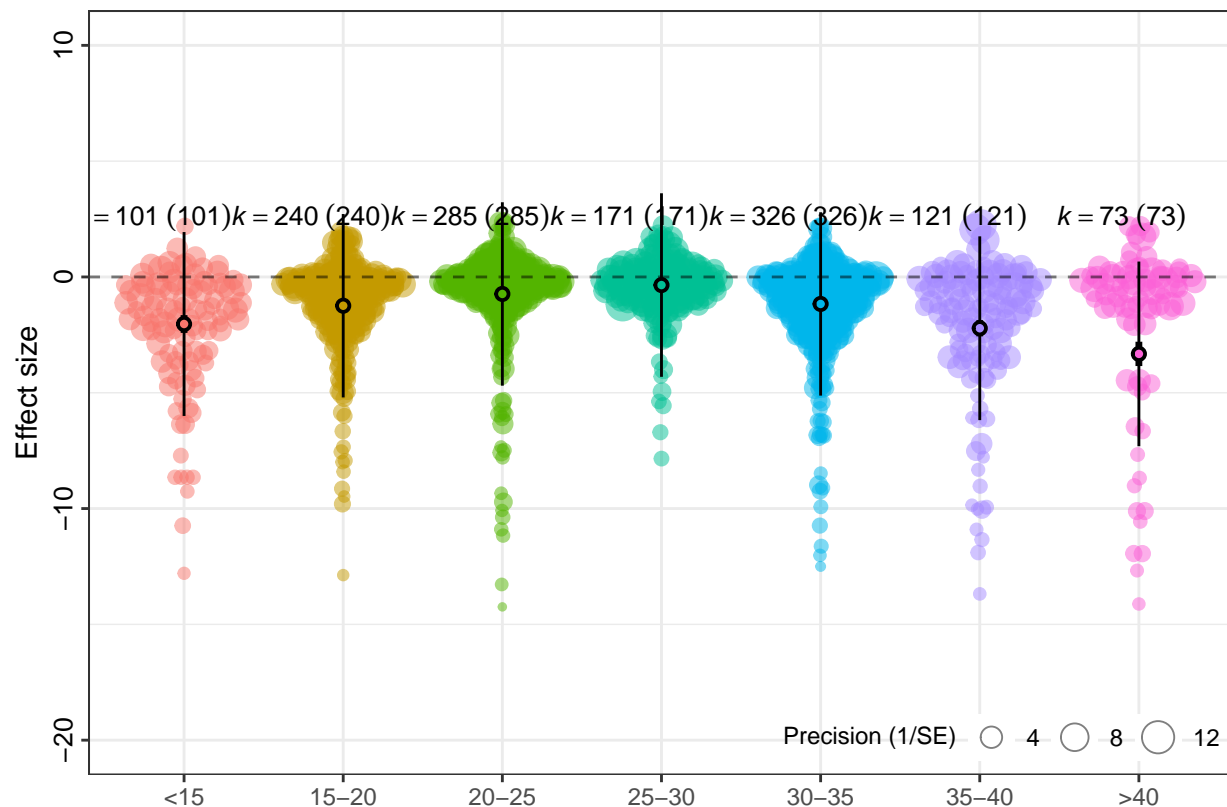
```
table(sdata$bin.temp)
```

```
##
##    <15    >40 15-20 20-25 25-30 30-35 35-40
##    101     73   240   285   171   326   121
```

```
meta_sa_bintemp <- rma.mv(es, VCV_shared_sa, mod = ~bin.temp -
  1, random = list(~1 | study_code, ~1 | obs), data = sdata,
  test = "t", dfs = "contain", method = "REML")
```

```
summary(meta_sa_bintemp)
```

```
##
## Multivariate Meta-Analysis Model (k = 1317; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -2682.1397    5364.2794    5382.2794    5428.8795    5382.4179
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  2.1228  1.4570   333     no  study_code
## sigma^2.2  1.9424  1.3937  1317     no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1310) = 28196.0335, p-val < .0001
##
## Test of Moderators (coefficients 1:7):
## QM(df = 7) = 339.9113, p-val < .0001
##
## Model Results:
##
##      estimate      se      zval      pval      ci.lb      ci.ub
## bin.temp<15    -2.0314  0.1958  -10.3731  <.0001   -2.4153   -1.6476 ***
## bin.temp>40    -3.3215  0.2678  -12.4015  <.0001   -3.8464   -2.7965 ***
## bin.temp15-20  -1.2425  0.1308   -9.4974  <.0001   -1.4989   -0.9861 ***
## bin.temp20-25  -0.7339  0.1241   -5.9157  <.0001   -0.9770   -0.4907 ***
## bin.temp25-30  -0.3498  0.1483   -2.3581  0.0184   -0.6406   -0.0591 *
## bin.temp30-35  -1.1655  0.1194   -9.7619  <.0001   -1.3994   -0.9315 ***
## bin.temp35-40  -2.2142  0.1793  -12.3461  <.0001   -2.5657   -1.8627 ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



Now we completely remove any study that has an effect size in the highest or lowest 2.5%.

```
summary(meta_sub_bintemp)
```

```
##
## Multivariate Meta-Analysis Model (k = 1217; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -2316.9699    4633.9398    4651.9398    4697.8252    4652.0898
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  1.4339  1.1974   309    no  study_code
## sigma^2.2  1.5158  1.2312  1217    no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1210) = 24351.0817, p-val < .0001
##
## Test of Moderators (coefficients 1:7):
## QM(df = 7) = 305.6725, p-val < .0001
##
## Model Results:
##
```

```
##               estimate      se      zval      pval      ci.lb      ci.ub
## bin.temp<15      -1.9502  0.1740 -11.2057 <.0001 -2.2914 -1.6091 ***
## bin.temp>40      -2.2312  0.2590  -8.6148 <.0001 -2.7388 -1.7236 ***
## bin.temp15-20    -1.1827  0.1162 -10.1757 <.0001 -1.4105 -0.9549 ***
## bin.temp20-25    -0.6228  0.1101  -5.6567 <.0001 -0.8386 -0.4070 ***
## bin.temp25-30    -0.3410  0.1318  -2.5873  0.0097 -0.5993 -0.0827 **
## bin.temp30-35    -1.1553  0.1061 -10.8840 <.0001 -1.3634 -0.9473 ***
## bin.temp35-40    -1.7366  0.1649 -10.5340 <.0001 -2.0597 -1.4135 ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Geary's Test

Here we use Geary's test to determine which effect sizes to remove. According to Hedges et al 1999, effect sizes are deemed valid and accurate when the standardized mean of either the control or the treatment group is ≥ 3 . In other words, if

$$\frac{\bar{X}}{SD}\sqrt{N} \geq 3$$

As a sensitivity analysis we will determine which of our effect sizes do not meet this threshold i.e. are < 3 . We then rerun our best fitting models with these effect sizes removed.

```
# Read in data with calculated standardised mean
gearydata <- read.csv("../Data/Gearys_test_data.csv")

# view the first fews rows of selected columns
gearydata %>%
  dplyr::select(Experiment.code, Trait.category, reftemp, treattemp,
    es, v, gtest) %>%
  head(., 30)

# How many effect sozes have a standardised mean < 3 ?
outliers_lon <- subset(gearydata, Trait.category == "Longevity" &
  gtest < 3) #82
outliers_sur <- subset(gearydata, Trait.category == "Survival" &
  gtest < 3) # 171
outliers_rep <- subset(gearydata, Trait.category == "Reproduction" &
  gtest < 3) # 3
```

There are 256 effect sizes in total that have a $gtest < 3$. 82 of these are longevity effect sizes, 3 are survival and 171 are reproduction.

We remove these effect sizes and rerun the best fitting model from earlier which is the model `meta_trait_treat2`, where temperature is modelled as a quadratic effect on reproduction.

Treatment temperature as a quadratic effect with Geary's outliers removed.

```
# Remove outliers
s.gearydata <- gearydata %>%
  subset(gtest > 3)

# Resubset reproduction dat
gdata_rep_warm <- subset(s.gearydata, Trait.category == "Reproduction" &
```

```

    warm.cool == "Warm")
gdata_rep_cool <- subset(s.gearydata, Trait.category == "Reproduction" &
    warm.cool == "Cool")

all_gdata <- rbind(gdata_rep_warm, gdata_rep_cool)

all_gdata <- all_gdata %>%
  mutate(c_treattemp = treattemp - 25)

### Create random factors into data frame
all_gdata$obs <- factor(c(1:nrow(all_gdata))) # Unique observation code
all_gdata$study_code <- factor(all_gdata$Paper.code) # Model requires column names study_code
all_gdata$Species.phylo <- factor(all_gdata$Species.latin) # Species names for phylo matrix
all_gdata$species <- factor(all_gdata$Species.latin) # Another species column for random factor

precision <- sqrt(1/all_gdata$v) # inverse standard error
all_gdata[, "precision"] <- precision

# recalculate v matrix
all_gdata$shared_control <- factor(all_gdata$Effect.size.code)
VCV_shared_ga <- impute_covariance_matrix(vi = all_gdata$v, cluster = all_gdata$shared_control,
    r = 0.5)

summary(meta_ga_treat2)

```

```

##
## Multivariate Meta-Analysis Model (k = 1198; method: REML)
##
##      logLik      Deviance      AIC      BIC      AICc
## -3224.0379    6448.0758    6458.0758    6483.5053    6458.1263
##
## Variance Components:
##
##      estim      sqrt  nlvls  fixed      factor
## sigma^2.1  6.2549  2.5010   330     no  study_code
## sigma^2.2  4.0534  2.0133   1198     no      obs
##
## Test for Residual Heterogeneity:
## QE(df = 1195) = 33797.9756, p-val < .0001
##
## Test of Moderators (coefficients 2:3):
## F(df1 = 2, df2 = 1195) = 71.8333, p-val < .0001
##
## Model Results:
##
##      estimate      se      tval      df      pval      ci.lb      ci.ub
## intrcpt      -0.4916  0.1698    -2.8947    327  0.0041    -0.8257    -0.1575
## poly(c_treattemp, degree = 2, raw = TRUE)1    -0.0580  0.0097    -5.9797   1195 <.0001    -0.0770    -0.0390
## poly(c_treattemp, degree = 2, raw = TRUE)2    -0.0114  0.0010   -11.2330   1195 <.0001    -0.0134    -0.0094
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Below is a bubble plot of the fitted quadratic model.

