EXERCISE CLASS 2 (Part 3/3)

Hypothesis testing in the presence of two samples

Assumptions:

- $X_{11}, X_{12}, ..., X_{1n1}$ is a random sample of size n_1 from population 1;
- $X_{21}, X_{22}, ..., X_{2n2}$ is a random sample of size n_2 from population 2;
- The two populations are **independent**;
- Both populations are **normal** (or central limit theorem applies);
- The variances of the populations are known.

Under those assumptions, the quantity:

$$Z = \frac{\overline{X_1} - \overline{X_2} - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Follows a standard normal distribution, N(0,1)

Testing Hypotheses on the Difference in Means, Variances Known

Null hypothesis:
$$H_0$$
: $\mu_1 - \mu_2 = \Delta_0$

Test statistic:
$$Z_0 = \frac{\overline{X}_1 - \overline{X}_2 - \Delta_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Rejection Criterion

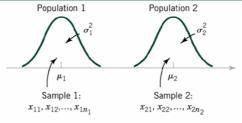
$$H_1$$
: $\mu_1 - \mu_2 \neq \Delta_0$
 H_1 : $\mu_1 - \mu_2 > \Delta_0$

$$z_0 > z_{\alpha/2}$$
 or $z_0 < -z_{\alpha/2}$
 $z_0 > z_{\alpha}$

$$H_1$$
: $\mu_1 - \mu_2 > \Delta_0$

$$z_0 < -z_\alpha$$

 H_1 : $\mu_1 - \mu_2 < \Delta_0$



Two independent populations.

Two-sample z test - confidence interval

The $100(1-\alpha)\%$ confidence interval on the difference between the population means (known variances) is given by:

$$\overline{X_1} - \overline{X_2} - Z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \leq \mu_1 - \mu_2 \leq \overline{X_1} - \overline{X_2} + Z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

Two-sample t-test

The case of equal but unknown variances

Assumptions (case 1: $\sigma_1^2 = \sigma_1^2 = \sigma^2$):

- $X_{11}, X_{12}, ..., X_{1n1}$ is a random sample of size n_1 from population 1;
- $X_{21}, X_{22}, ..., X_{2n2}$ is a random sample of size n_2 from population 2;
- The two populations are **independent**;
- Both populations are **normal** (or central limit theorem applies);
- The variances of the populations are unknown and equal.

Under those assumptions, the quantity:

$$T = \frac{\overline{X_1} - \overline{X_2} - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

Follows a student-t distribution with $n_1 + n_2 - 2$ degrees of fredom. **Pooled variance**:

$$S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

Testing Hypotheses on the Difference in Means of Two Normal Distributions, Variances Unknown and Equal I

Null hypothesis:
$$H_0$$
: $\mu_1 - \mu_2 = \Delta_0$

Test statistic: $T_0 = \frac{\overline{X}_1 - \overline{X}_2 - \Delta_0}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$

Alternative Hypothesis Rejection Criterion

 H_1 : $\mu_1 - \mu_2 \neq \Delta_0$ $t_0 > t_{\alpha/2, n_1 + n_2 - 2}$ or $t_0 < -t_{\alpha/2, n_1 + n_2 - 2}$ $t_0 > t_{\alpha, n_1 + n_2 - 2}$ $t_0 > t_{\alpha, n_1 + n_2 - 2}$ $t_0 < -t_{\alpha, n_1 + n_2 - 2}$

Two-sample t-test

The case of non-equal and unknown variances

Assumptions (case 2: $\sigma_1^2 \neq \sigma_1^2$):

- $X_{11}, X_{12}, ..., X_{1n}$ is a random sample of size n_1 from population 1;
- $X_{21}, X_{22}, ..., X_{2n}$ is a random sample of size n_2 from population 2;
- The two populations are **independent**;
- Both populations are **normal** (or central limit theorem applies);
- The variances of the populations are unknown and not equal.

Under those assumptions, the quantity:

$$T = \frac{\overline{X_1} - \overline{X_2} - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}} \qquad \qquad \nu = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)^2}{\frac{\left(S_1^2/n_1\right)^2}{n_1 - 1} + \frac{\left(S_2^2/n_2\right)^2}{n_2 - 1}}$$

Follows a student-t distribution with ν degrees of fredom

Paired t-test

Assumptions:

- $X_{11}, X_{12}, ..., X_{1n}$ is a random sample of size n from population 1;
- $X_{21}, X_{22}, ..., X_{2n}$ is a random sample of size n from population 2;
- The differences between pairs, $D_j = X_{1j} X_{2j}$, are **normal** (or central limit theorem applies);
- The variance of the differences between pairs is unknown

Under those assumptions, the quantity:

$$T = \frac{\overline{D} - \mu_D}{S_D / \sqrt{n}}$$

Follows a student-t distribution with n-1 degrees of freedom, where: $D_j = X_{1j} - X_{2j} \sim N(\mu_D, \sigma_D^2)$

(NOTICE: we are assuming differences to be random normal var.)

The $100(1-\alpha)\%$ confidence interval on the difference between the population means is given by:

$$\overline{D} - t_{\frac{\alpha}{2}, n-1} S_D / \sqrt{n} \leq \mu_D \leq \overline{D} + t_{\frac{\alpha}{2}, n-1} S_D / \sqrt{n}$$

 S_D is lower than the pooled standard deviation: paired test is more precise (smaller c.i.)

EXERCISE 3

Ten people are involved in a diet program. The weights before and after the program is reported in the table (expressed in pounds, 1 lb =0.454). The data is stored in the file $ESE2_ex3.csv$.

- a) Is there statistical evidence (95%) to state that the diet program was effective?
- b) Can we state that (95%) the program yielded a mean weigth reduction higher than 10 lb?
- c) Design a two-sided confidence interval at 95% on the weight difference

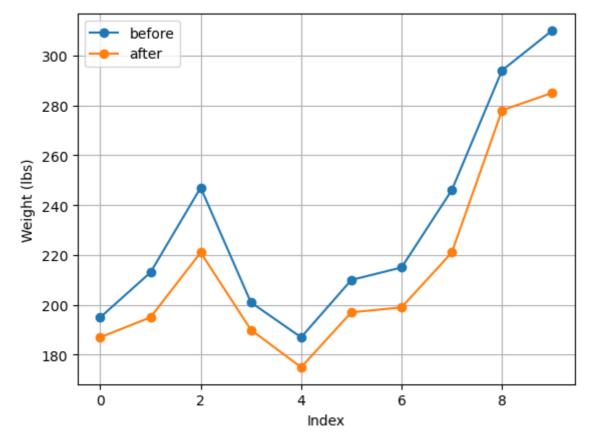
```
In []: # Import the necessary libraries
   import numpy as np
   import matplotlib.pyplot as plt
   import pandas as pd
   from scipy import stats

# Import the dataset
   data = pd.read_csv('ESE2_ex3.csv')

# Inspect the dataset
   data.head()
```

```
Out[]:
            before after
                           d
               195
                     187
                           8
         1
               213
                     195 18
         2
               247
                     221 26
         3
               201
                     190 11
         4
               187
                     175 12
```

```
In [ ]: # Let's plot the data first
   plt.plot(data['before'], 'o-', label='before')
   plt.plot(data['after'], 'o-', label='after')
   plt.xlabel('Index')
   plt.ylabel('Weight (lbs)')
   plt.legend()
   plt.grid()
   plt.show()
```



Point a

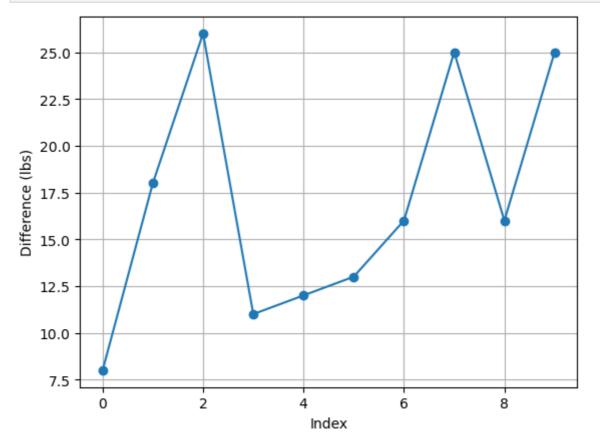
Is there statistical evidence (95%) to state that the diet program was effective?

Solution

Data are PAIRED, so we are interested in the difference between the two.

```
In [ ]: # Plot the difference d
plt.plot(data['d'], 'o-')
plt.xlabel('Index')
plt.ylabel('Difference (lbs)')
```

```
plt.grid()
plt.show()
```



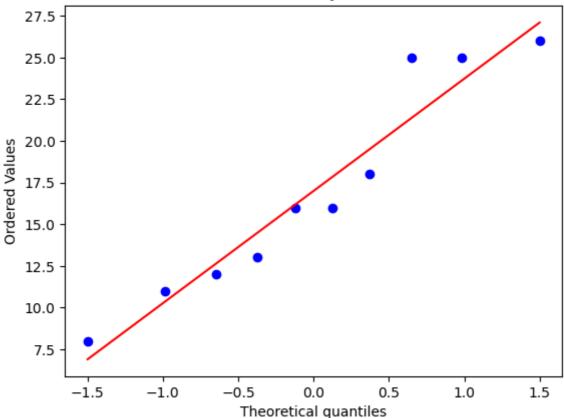
Let's check the normality of the data with the Shapiro-Wilk test.

```
In []: # Check the normality of the difference
# We can use the Shapiro-Wilk test
_, p_value_SW = stats.shapiro(data['d'])
print('p-value of the Shapiro-Wilk test: %.3f' % p_value_SW)

# QQ-plot
stats.probplot(data['d'], dist="norm", plot=plt)
plt.show()
```

p-value of the Shapiro-Wilk test: 0.270





Now that we know that the data are normally distributed, we can use the ttest to evaluate the following hypothesis:

$$H_0: \mu_d = 0$$

$$H_1: \mu_d > 0$$

The t-test statistic is:

$$t_0=rac{ar{d}}{s_d/\sqrt{n}}$$

where \bar{d} is the sample mean of the differences, s_d is the sample standard deviation of the differences and n is the number of observations.

```
In []: # Compute the t-statistic and the corresponding p-value
    n = len(data['d']) # number of samples
    df = n - 1 # degrees of freedom

t0 = np.mean(data['d']) / (data['d'].std() / np.sqrt(n))
    print('t-statistic: %.3f' % t0)

# Compute the p-value
    p_value_t0 = 1 - stats.t.cdf(np.abs(t0), df)
    print('p-value: %.3f' % p_value_t0)
```

t-statistic: 8.384 p-value: 0.000

Alternatively, we can use the ttest_1samp function from the scipy.stats module.

```
In [ ]: # Perform the t-test on the difference using the stats.ttest_1samp function
    t0_stats, p_value_t0_stats = stats.ttest_1samp(data['d'], popmean = 0, alternatives
    print('t-statistic from stats.ttest_1samp: %.3f' % t0_stats)
    print('p-value from stats.ttest_1samp: %.3f' % p_value_t0_stats)

t-statistic from stats.ttest_1samp: 8.384
    p-value from stats.ttest_1samp: 0.000
```

Alternatively, we can directly use the ttest_rel function from the scipy.stats module on the original data (without computing the differences).

```
In [ ]: # Alternatively, you can perform a paired t-test using the stats.ttest_rel function
t0_stats_trel, p_value_t0_stats_trel = stats.ttest_rel(data['before'], data['after
print('t-statistic from stats.ttest_rel: %.3f' % t0_stats_trel)
print('p-value from stats.ttest_rel: %.3f' % p_value_t0_stats_trel)

t-statistic from stats.ttest_rel: 8.384
p-value from stats.ttest_rel: 0.000
```

Point b

Can we state that (95%) the program yielded a mean weigth reduction higher than 10 lbs?

Solution

We can use the same t-test as in point a, but with the following hypothesis:

$$H_0: \mu_d = \Delta_0$$

$$H_1: \mu_d > \Delta_0$$

The t-test statistic is still:

$$t_0 = rac{ar{d} - \Delta_0}{s_d/\sqrt{n}}$$

where \bar{d} is the sample mean of the differences, s_d is the sample standard deviation of the differences, n is the number of observations and Δ_0 is the hypothesized value of the mean difference.

```
In [ ]: # Answer to point b
CL = 0.95  # confidence level
alpha = 1 - CL # significance level

delta0 = 10  # null hypothesis
t0_delta0_stats, p_value_t0_delta0_stats = stats.ttest_1samp(data['d'], popmean = print('t-statistic from stats.ttest_1samp: %.3f' % t0_delta0_stats)
print('p-value from stats.ttest_1samp: %.4f' % p_value_t0_delta0_stats)
```

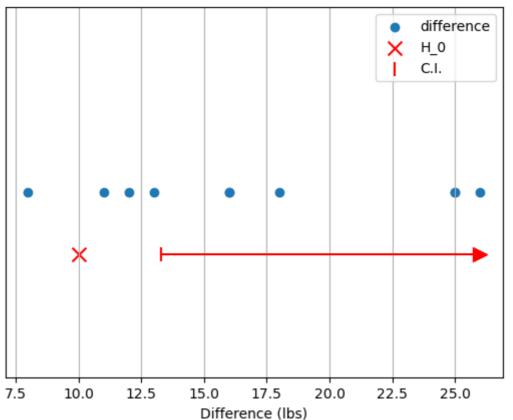
```
t-statistic from stats.ttest_1samp: 3.452
p-value from stats.ttest_1samp: 0.0036
```

Let's plot the confidence interval for the mean difference.

```
In [ ]: # Calculate the lower bound of the one-sided confidence interval
        t_alpha = stats.t.ppf(1 - alpha, df)
        CI_lower = data['d'].mean() - t_alpha * data['d'].std() / np.sqrt(n)
        print('Lower bound of the one-sided confidence interval: %.3f' % CI_lower)
        # Visualize the confidence interval on a dot plot
        plt.title('One-sided confidence interval for the mean with CL = %.2f' % CL)
        plt.scatter(data['d'], np.zeros(n), label='difference')
        # plot H0
        plt.scatter(delta0, -0.01, label='H_0', color='r', marker='x', s=100)
        # plot the confidence interval
        plt.scatter(CI_lower, -0.01, label='C.I.', color='r', marker='|', s=100)
        plt.plot([CI_lower, np.max(data['d'])], [-0.01, -0.01], color='r')
        plt.scatter(np.max(data['d']), -0.01, color='r', marker='>', s=100)
        # Add Labels and Legend
        plt.ylim(-0.03, 0.03)
        plt.xlabel('Difference (lbs)')
        plt.yticks([])
        plt.legend()
        plt.grid()
        plt.show()
```

Lower bound of the one-sided confidence interval: 13.283

One-sided confidence interval for the mean with CL = 0.95



Point c

Design a two-sided confidence interval at 95% on the weight difference.

Solution

We can use the same t-test as in point a, but with the following hypothesis:

$$H_0: \mu_d = 0$$

$$H_1: \mu_d
eq 0$$

The t-test statistic is still:

$$t_0=rac{ar{d}}{s_d/\sqrt{n}}$$

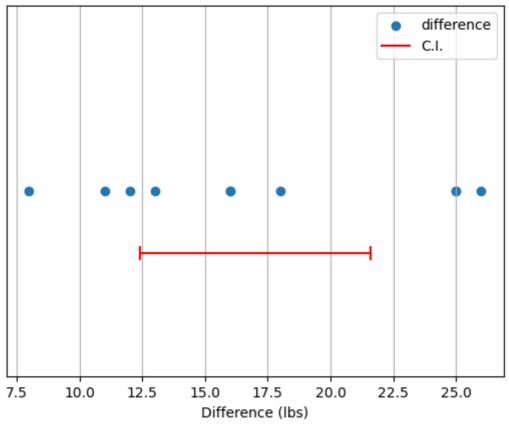
And the corresponding confidence interval is:

$$ar{d} - t_{lpha/2,n-1} rac{s_d}{\sqrt{n}} \leq \mu_d \leq ar{d} \, + t_{lpha/2,n-1} rac{s_d}{\sqrt{n}}$$

Let's plot the confidence interval for the mean difference.

```
In []: # Visualize the confidence interval on a dot plot
plt.title('Confidence interval for the mean with CL = %.2f' % CL)
plt.scatter(data['d'], np.zeros(n), label='difference')
# plot the confidence interval alonside the dot plot
plt.scatter(CI[0], -0.01, color='r', marker='|', s=100)
plt.plot([CI[0], CI[1]], [-0.01, -0.01], color='r', label='C.I.')
plt.scatter(CI[1], -0.01, color='r', marker='|', s=100)
# Add Labels and Legend
plt.ylim(-0.03, 0.03)
plt.xlabel('Difference (lbs)')
plt.yticks([])
plt.legend()
plt.grid()
plt.show()
```

Confidence interval for the mean with CL = 0.95



F-test (equality of variances)

A brief note on the Fisher F distribution

Let W and Y be **independent chi-squared random variables** with u and v degrees of freedom, respectively. Then, the ratio:

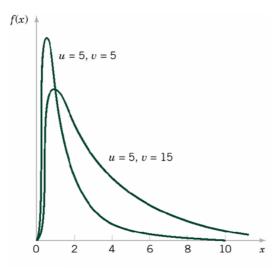
$$F = \frac{W/u}{Y/v}$$

Follows an F distribution with u and v degrees of freedom ($F_{u,v}$). Its probability density function is:

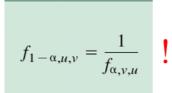
$$f(x) = \frac{\Gamma\left(\frac{u+v}{2}\right)\left(\frac{u}{v}\right)^{u/2}x^{(u/2)-1}}{\Gamma\left(\frac{u}{2}\right)\Gamma\left(\frac{v}{2}\right)\left[\left(\frac{u}{v}\right)x+1\right]^{(u+v)/2}}, \ 0 < x < \infty$$

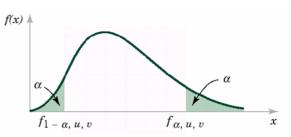
Remind: $X^2 = \frac{(n-1)S^2}{\sigma^2}$ follows a *chi-squared* (χ^2) *distribution* with n-1 degrees of freedom, thus: $S^2 \sim [\sigma^2/(n-1)]\chi^2(n-1)$

A brief note on the Fisher F distribution



Probability density functions of two ${\cal F}$ distributions.





Upper and lower percentage points of the F distribution.

The F-test

Assumptions:

- $X_{11}, X_{12}, ..., X_{1n1}$ is a random sample of size n_1 from population 1;
- $X_{21}, X_{22}, ..., X_{2n2}$ is a random sample of size n_2 from population 2;
- The two populations are **normal** (or central limit theorem applies);
- The two populations are independent;
- The variances of the populations are unknown (obvious)

Under those assumptions, the quantity:

$$F = \frac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2}$$

Follows an F distribution with $n_1 - 1$ and $n_2 - 1$ degrees of freedom.

Testing Hypotheses on the Equality of Variances of Two Normal Distributions Null hypothesis: H_0 : $\sigma_1^2 = \sigma_2^2$ Test statistic: $F_0 = \frac{S_1^2}{S_2^2}$ (5-21) Alternative Hypotheses H_1 : $\sigma_1^2 \neq \sigma_2^2$ Rejection Criterion $f_0 > f_{\alpha/2,n_1-1,n_2-1} \text{ or } f_0 < f_{1-\alpha/2,n_1-1,n_2-1}$ H_1 : $\sigma_1^2 > \sigma_2^2$ $f_0 > f_{\alpha,n_1-1,n_2-1}$ H_1 : $\sigma_1^2 < \sigma_2^2$ $f_0 < f_{1-\alpha,n_1-1,n_2-1}$

The $100(1-\alpha)\%$ confidence interval on the difference between the population means is given by:

$$\frac{S_1^2}{S_2^2} f_{1-\alpha/2, \mathbf{n_2} - \mathbf{1}, \mathbf{n_1} - \mathbf{1}} \le \frac{\sigma_1^2}{\sigma_2^2} \le \frac{S_1^2}{S_2^2} f_{\alpha/2, \mathbf{n_2} - \mathbf{1}, \mathbf{n_1} - \mathbf{1}}$$

EXERCISE 4

We want to evaluate the resistance of resistors provided by two different suppliers. The data is stored in the file ESE2_ex4.csv.

- a) Design a boxplot to compare the two samples and estimate the major descriptive statistics
- b) What can we infer about the mean resistance of the resistors provided by the two different suppliers?
- c) Compute the Type II error expression in the variance equality test and compute the test power when the true variance of the first supplier is 1.5 times larger than the one of the second supplier

```
In []: # Import the necessary libraries
import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
from scipy import stats

# Import the dataset
data = pd.read_csv('ESE2_ex4.csv')

# Inspect the dataset
data.head()
```

```
Out[]: supp1 supp2

0 96.8 106.8

1 100.0 103.7

2 99.9 104.0

3 98.6 102.8

4 101.2 107.2
```

Let's inspect the data.

The dataset contains NaN values because the number of observations is different for each supplier. We need to remove them.

```
In [ ]: # Let's split the dataset into two dataframes, one for each supplier
    data1 = data['supp1']
    data2 = data['supp2']

# and remove the nan values from data1
    data1 = data1.dropna()
```

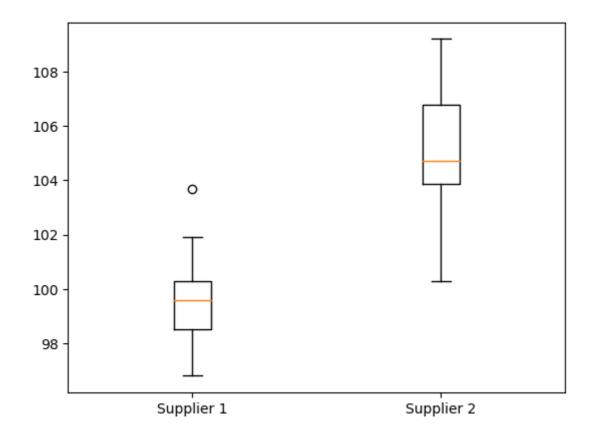
Point a

Design a boxplot to compare the two samples and estimate the major descriptive statistics.

Solution

Use the boxplot function from matplotlib.pyplot to plot the boxplot.

```
In [ ]: # Answer to point a
  plt.boxplot([data1, data2], labels=['Supplier 1', 'Supplier 2'])
  plt.show()
```



Use the describe function from pandas to compute the descriptive statistics.

In []:	data.	describe()	
Out[]:		supp1	supp2
	count	25.000000	35.000000
	mean	99.576000	105.068571
	std	1.528965	1.962557
	min	96.800000	100.300000
	25%	98.500000	103.850000
	50%	99.600000	104.700000
	75%	100.300000	106.800000
	max	103.700000	109.200000

Point b

What can we infer about the mean resistance of the resistors provided by the two different suppliers?

Solution

First we need to check the assumptions:

Normality

- Independence (we assume it is ok, for now) within and between
- Other? Outliers?

Let's check the normality with the Shapiro-Wilk test.

```
# Answer to point b
n1 = len(data1) # number of observations for supplier 1
n2 = len(data2) # number of observations for supplier 2
# Let's check the assumptions
# Normality
# Shapiro-Wilk test
_, p_value_SW_1 = stats.shapiro(data1)
_, p_value_SW_2 = stats.shapiro(data2)
print('p-value for Shapiro-Wilk test for supplier 1: %.3f' % p_value_SW_1)
print('p-value for Shapiro-Wilk test for supplier 2: %.3f' % p_value_SW_2)
# QQ-plot
fig, ax = plt.subplots(1, 2, figsize=(10, 5))
stats.probplot(data1, plot=ax[0])
ax[0].set_title('Probability plot of Supplier 1')
stats.probplot(data2, plot=ax[1])
ax[1].set_title('Probability plot of Supplier 2')
plt.show()
p-value for Shapiro-Wilk test for supplier 1: 0.523
p-value for Shapiro-Wilk test for supplier 2: 0.722
            Probability plot of Supplier 1
                                                         Probability plot of Supplier 2
  104
  103
                                               108
  102
                                               106
                                             Ordered Values
Ordered Values
  101
  100
                                               104
   99
   98
                                               102
   97
                                               100
               -1
                        0
                                 1
                                                            -1
                                                                     0
```

We want to compare the means of two populations. Variances are unknown, thus there are two possible situations:

Theoretical quantiles

- Equal (unknown) variances
- Different (unknown) variances

Theoretical quantiles

First step: hypothesis test on the equality of variances

Null hypothesis: the two variances are equal

$$H_0:\sigma_1^2=\sigma_2^2$$

Alternative hypothesis: the two variances are different

$$H_1:\sigma_1^2
eq\sigma_2^2$$

This hypothesis test is equivalent to:

$$H_0:rac{\sigma_1^2}{\sigma_2^2}=1$$

$$H_1:rac{\sigma_1^2}{\sigma_2^2}
eq 1$$

We can use the F-test to test the equality of variances. The test statistic is:

$$F_0 = rac{s_1^2}{s_2^2}$$

The corresponding confidence interval is:

$$rac{s_1^2}{s_2^2}F_{lpha/2,n_2-1,n_1-1} \leq rac{\sigma_1^2}{\sigma_2^2} \leq rac{s_1^2}{s_2^2}F_{1-lpha/2,n_2-1,n_1-1}$$

```
In [ ]: CL = 0.95  # Confidence level
alpha = 1 - CL # Significance level

# Test the equality of variances
# F-test
F0 = data1.var()/data2.var()
df1 = n1 - 1 # degrees of freedom for supplier 1
df2 = n2 - 1 # degrees of freedom for supplier 2
CI = [F0 * stats.f.ppf(alpha/2, df2, df1), F0 * stats.f.ppf(1-alpha/2, df2, df1)]
print('Confidence interval on the ratio of variances (CL = %.2f): [%.3f, %.3f]' %
```

Confidence interval on the ratio of variances (CL = 0.95): [0.293, 1.323]

There is no statistical evidence to state that the two variances are different (accept null hypothesis), because the interval includes 1.

Let's plot the distribution under the null hypothesis, the rejection region and the test statistic.

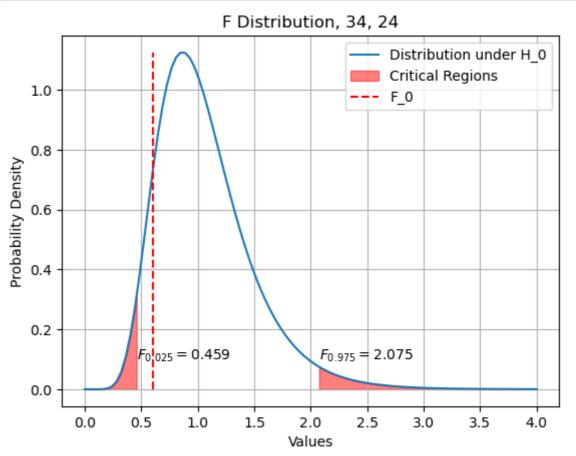
```
In [ ]: # plot the cumulative probability
    x = np.linspace(0, 4, 100)
    plt.plot(x, stats.f.pdf(x, df1, df2), label='Distribution under H_0')

# Adding Title, Labels and Grid
    plt.title("F Distribution, %d, %d" % (df2, df1))
    plt.xlabel("Values")
    plt.ylabel("Probability Density")
    plt.grid(True)

# Filling the Probability Area
    F_1 = stats.f.ppf(alpha/2, df1, df2)
```

```
F_2 = stats.f.ppf(1-alpha/2, df1, df2)
x_fill = np.linspace(0, F_1, 100)
y_fill = stats.f.pdf(x_fill, df1, df2)
plt.fill_between(x_fill, y_fill, color='red', alpha=0.5)
x_fill = np.linspace(F_2, np.max(x), 100)
y_fill = stats.f.pdf(x_fill, df1, df2)
plt.fill_between(x_fill, y_fill, color='red', alpha=0.5, label='Critical Regions')
# Add text to the plot with the chi2 values and centering the text
plt.text(F_1, 0.1, r'$F_{%.3f} = {%.3f}$' % (alpha/2, F_1), fontsize=10)
plt.text(F_2, 0.1, r'$F_{%.3f} = {%.3f}$' % (1-alpha/2, F_2), fontsize=10)

# Plot the test statistic F0
plt.vlines(F0, 0, np.max(stats.f.pdf(x, df1, df2)), color='r', linestyle='--', label*
# Showing Plot
plt.legend()
plt.show()
```



Let's compute the p-value corresponding to the test statistic.

```
In []: # Compute the p-value
p_value_F0 = 2 * stats.f.cdf(F0, df1, df2)

print('p-value for F-test for equal variances: %.3f' % p_value_F0)
```

p-value for F-test for equal variances: 0.205

We can also use the Chi-squared statistic to compute the confidence interval for the two standard deviations.

Rember:

$$rac{(n-1)s^2}{\chi^2_{lpha/2,n-1}} \leq \sigma^2 \leq rac{(n-1)s^2}{\chi^2_{1-lpha/2,n-1}}$$

In []: # Compute the 95% confidence interval on the individual standard deviation
 CI_sigma_1 = np.sqrt([(df1 * data1.var())/stats.chi2.ppf(1-alpha/2, df1), (df1 * data1.var())/stats.chi2.ppf(1-alpha/2, df2), (df2 * data2.var())/stats.chi2.ppf(1-alpha/2, df2), (df2 * da

Confidence interval on the standard deviation for supplier 1 (CL = 0.95): [1.194, 2.127]

Confidence interval on the standard deviation for supplier 2 (CL = 0.95): [1.587, 2.571]

Now that we have verified the equality of variances, we can perform the t-test (with equal variances). So we can use the ttest_ind function from scipy.stats and set the equal_var parameter to True.

Remember the formula of the two-sample t-test:

$$T = rac{(ar{X}_1 - ar{X}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{rac{1}{n_1} + rac{1}{n_2}}}$$

where:

$$S_p^2 = rac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1 + n_2 - 2}$$

```
In [ ]: t0, p_value_t0 = stats.ttest_ind(data1, data2, equal_var=True)
print('t-test: t0 = %.3f' % t0)
print('p-value for t-test: %.3f' % p_value_t0)
```

t-test: t0 = -11.680 p-value for t-test: 0.000

Point c

Compute the Type II error expression in the variance equality test and compute the test power when the true variance of the first supplier is 1.5 times larger than the one of the second supplier.

Solution

The Type II error is the probability of accepting the null hypothesis when it is false.

$$\beta = Pr(\text{accept } H_0 \text{ when } H_1 \text{ is true})$$

Let's expand the formula for the F-test:

$$eta = Pr\left(F_{1-lpha/2,n_1-1,n_2-2} \leq rac{s_1^2}{s_2^2} \leq F_{lpha/2,n_1-1,n_2-2} \mid rac{\sigma_1^2}{\sigma_2^2} = \delta
eq 1
ight)$$

If we multiply all the terms by σ_2^2/σ_1^2 we get:

$$rac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} \sim F_{n_1-1,n_2-2}$$

If we substitute σ_2^2/σ_1^2 with the ratio we want to test, we get:

$$eta = Pr\left(rac{F_{1-lpha/2,n_1-1,n_2-2}}{1.5} \leq rac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} \leq rac{F_{lpha/2,n_1-1,n_2-2}}{1.5}
ight)$$

Rearranging the terms we get:

$$eta = Pr\left(rac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} \leq rac{F_{lpha/2,n_1-1,n_2-2}}{1.5}
ight) - Pr\left(rac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} \leq rac{F_{1-lpha/2,n_1-1,n_2-2}}{1.5}
ight)$$

```
In [ ]: # Answer to point c
  ratio = 1.5 # ratio between the variances of the two samples
  beta = stats.f.cdf(stats.f.ppf(1-alpha/2, df1, df2)/ratio, df1, df2) - stats.f.cdf
  print('Power of the test: %.3f' % (1-beta))
```

Power of the test: 0.191

EXERCISE 5

A study is aimed at investigating the microbial spore content in orange jouce pastourized via high pressure processing under two different pressure and temperature condition, namely condition 1 and condition 2. An index that represents the microbial content after the process was measured for the two different conditions (data table in attached file 'ESE2 ex5.csv')

- a) Design a boxplot to compare the two conditions and estimate the major descriptive statistics
- b) Compute the confidence interval on the variance ratio at 95%
- c) Compute a 99% confidence interval on the difference between the two means (highligh the necessary assumptions). What conclusion can we draw?

```
In []: # Import the necessary libraries
import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
from scipy import stats

# Import the dataset
data = pd.read_csv('ESE2_ex5.csv')

# Inspect the dataset
data.head()
```

```
      cond1
      cond2

      0
      19.8
      14.9

      1
      18.5
      12.7

      2
      17.6
      11.9

      3
      16.7
      11.4

      4
      16.7
      10.1
```

```
In []: # The dataset contains nan values because the number of observations is different ;
# Let's split the dataset into two dataframes, one for each supplier
data1 = data['cond1']
data2 = data['cond2']

# and remove the nan values from data2
data2 = data2.dropna()

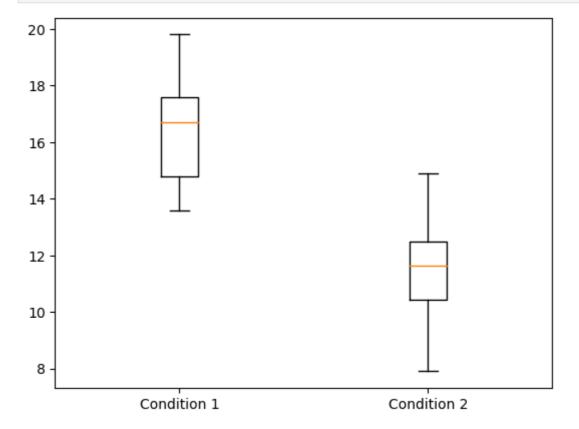
n1 = len(data1) # number of observations for supplier 1
n2 = len(data2) # number of observations for supplier 2

df1 = n1 - 1 # degrees of freedom for supplier 1
df2 = n2 - 1 # degrees of freedom for supplier 2
```

Point a

Design a boxplot to compare the two samples and estimate the major descriptive statistics

```
In [ ]: # Answer to point a
  plt.boxplot([data1, data2], labels=['Condition 1', 'Condition 2'])
  plt.show()
```



```
In [ ]: # make a dataframe with the major descriptive statistics and Q1, Q3
descriptive_stats = pd.DataFrame({'Supplier 1': [data1.mean(), data1.std(), data1.
```

```
'Supplier 2': [data2.mean(), data2.std(), data2
index=['Mean', 'StDev', 'Var', 'Min', 'Q1', 'Mo
print(np.round(descriptive_stats,3))
```

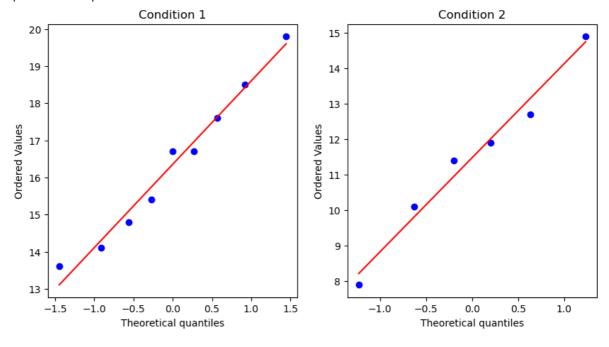
	Supplier 1	Supplier 2
Mean	16.356	11.483
StDev	2.069	2.370
Var	4.283	5.618
Min	13.600	7.900
Q1	14.800	10.425
Median	16.700	11.650
Q3	17.600	12.500
Max	19.800	14.900

Point b

Compute the confidence interval on the variance ratio at 95%.

```
# Answer to point b
In [ ]:
        CL = 0.95
                       # Confidence level
        alpha = 1 - CL # Significance Level
        # Check the assumptions
        _, p_value_SW_1 = stats.shapiro(data1)
        _, p_value_SW_2 = stats.shapiro(data2)
        print('p-value Shapiro-Wilk test for condition 1: %.3f' % p_value_SW_1)
        print('p-value Shapiro-Wilk test for condition 2: %.3f' % p_value_SW_2)
        # QQ-plot
        fig, ax = plt.subplots(1, 2, figsize=(10, 5))
        stats.probplot(data1, plot=ax[0])
        ax[0].set_title('Condition 1')
        stats.probplot(data2, plot=ax[1])
        ax[1].set_title('Condition 2')
        plt.show()
```

p-value Shapiro-Wilk test for condition 1: 0.865 p-value Shapiro-Wilk test for condition 2: 0.991



```
In [ ]: # Compute the CI on the variance ratio
F0 = data1.var() / data2.var()
```

```
CI = [F0 * stats.f.ppf(alpha/2, df2, df1), F0 * stats.f.ppf(1 - alpha/2, df2, df1)
print('The confidence interval on the variance ratio is: [%.3f, %.3f]' % (CI[0], C:

The confidence interval on the variance ratio is: [0.113, 3.673]

In []: # Compute the corresponding p-value
p_value = 2 * stats.f.cdf(F0, df1, df2)
print('The p-value is: %.3f' % p_value)

The p-value is: 0.697
```

Point c

Compute the 99% confidence interval on mean difference.

to conclude...

CONCLUDING SUMMARY

(what you have learned – or just reviewed)

- Type I and Type II error
- Power of a test
- · P-value
- Confidence intervals
- · Checking test assumptions
- · One sample tests:

Test for mean (known variance): one-sample z-test
 Test for mean (unknown variance): one-sample t-test

• Test for variance: chi-squared test (variance)

• Test for proportion: one-sample p-test

• Two sample tests:

Test for mean difference (known var): two-sample z-test
 Test for mean difference (unknown var): two-sample t-test
 Test for mean of paired data (unknown var): paired t-test
 Test for equality of variances: F-test (variances)

• Test for equality of proportions: two-sample p-test