

# 概率统计第五章习题课

1. 设  $X_1, \dots, X_n$  为来自总体  $X \sim N(\mu, \sigma^2)$  的样本, 其中  $\mu, \sigma^2$  未知.

(1) 求样本的样本空间和联合分布密度.

$$\Omega = \{ (x_1, \dots, x_n) \mid x_i \in R, i = 1, 2, \dots, n \}$$

$$\begin{aligned} f(x_1, \dots, x_n) &= \prod_{i=1}^n f(x_i) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \\ &= \frac{1}{(2\pi)^{\frac{n}{2}} \sigma^n} \cdot e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2} \end{aligned}$$

(2) 问下列随机变量中哪些是统计量

$$T_1 = \frac{1}{n-1} \sum_{i=1}^n X_i, \quad \text{😊}$$

$$T_2 = X_n - EX_1;$$

$$T_3 = 2X_2 + X_3; \quad \text{😊}$$

$$T_4 = \max(X_1, X_2, \dots, X_n); \quad \text{😊}$$

$$T_5 = \frac{X_1 - \mu}{\sigma};$$

$$T_6 = \sum_{i=1}^n \left( \frac{X_i}{\sigma} \right)^2.$$

2. 设 $X_1, X_2, \dots, X_n$ 是来自泊松分布 $P(\lambda)$ 的样本,  $\bar{X}, S^2$ 分别为样本均值和样本方差, 求  
 $E(\bar{X}), D(\bar{X}), E(S^2)$

解: 由 $X \sim P(\lambda)$ 知 $E(X) = \lambda$   $D(X) = \lambda$

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = E(X) = \lambda$$

$$D(\bar{X}) = \frac{1}{n} D(X) = \frac{\lambda}{n}$$

$$E(S^2) = D(X) = \lambda$$



**1)关系式** 
$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} (\sum_{i=1}^n X_i^2 - n\bar{X}^2)$$

**推导** 
$$\sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n (X_i^2 - 2X_i\bar{X} + \bar{X}^2)$$

$$= \sum_{i=1}^n X_i^2 - 2\bar{X} \sum_{i=1}^n X_i + \sum_{i=1}^n \bar{X}^2 = \sum_{i=1}^n X_i^2 - 2n\bar{X}^2 + n\bar{X}^2$$

$$= \sum_{i=1}^n X_i^2 - n\bar{X}^2$$

## 2) $E(S^2) = D(X)$

**推导**  $E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = E(X) \quad D(\bar{X}) = \frac{1}{n} D(X)$

$$E(S^2) = E\left(\frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - n\bar{X}^2\right)\right) = \frac{E\left(\sum_{i=1}^n X_i^2\right) - n\left[D(\bar{X}) + E^2(\bar{X})\right]}{n-1}$$

$$= \frac{\sum_{i=1}^n E(X_i^2) - DX - nE^2 X}{n-1} = \frac{\sum_{i=1}^n (DX + E^2 X) - DX - nE^2 X}{n-1}$$

$$= \frac{nDX + nE^2 X - DX - nE^2 X}{n-1} = DX$$

3. 在总体 $N(10, 4)$ 中随机抽容量为5的样本

$X_1, X_2, X_3, X_4, X_5$ . 求

(1)  $P(|\bar{X} - 10| > 2)$

(2)  $P\{\max(X_1, X_2, X_3, X_4, X_5) > 12\}$

(3)  $P\{\min(X_1, X_2, X_3, X_4, X_5) > 8\}$

$$(1) P(|\bar{X} - 10| > 2) = P\left(\left|\frac{\bar{X} - 10}{2/\sqrt{5}}\right| > \sqrt{5}\right)$$

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

$$= 1 - P\left(\left|\frac{\bar{X} - 10}{2/\sqrt{5}}\right| \leq \sqrt{5}\right)$$

$$= 2[1 - \Phi(\sqrt{5})] = 0.026$$

$$\begin{aligned}
 (2) \quad & P \{ \max (X_1, X_2, X_3, X_4, X_5) > 12 \} \\
 &= 1 - P \{ \max (X_1, X_2, X_3, X_4, X_5) \leq 12 \} \\
 &= 1 - \prod_{i=1}^5 P(X_i \leq 12) = 1 - [\Phi(\frac{12-10}{2})]^5 = 0.5785.
 \end{aligned}$$

$$\begin{aligned}
 (3) \quad & P \{ \min (X_1, X_2, X_3, X_4, X_5) > 8 \} \\
 &= \prod_{i=1}^5 P(X_i > 8) = [1 - \Phi(\frac{8-10}{2})]^5 = 0.4215.
 \end{aligned}$$



4. 设总体  $X \sim N(0, 0.2^2)$ ,  $(X_1, X_2, \dots, X_8)$  为其样本,

求  $a$ , 使  $P(\sum_{i=1}^8 X_i^2 < a) = 0.95$ .

$$\frac{X - \mu}{\sigma} \sim N(0, 1)$$

$$\frac{X_i}{0.2} \sim N(0, 1) \quad \frac{1}{0.2^2} \sum_{i=1}^8 X_i^2 \sim \chi^2(8)$$

$$P(\sum_{i=1}^8 X_i^2 < a) = P(\frac{1}{0.2^2} \sum_{i=1}^8 X_i^2 < \frac{a}{0.2^2})$$

$$= P(\chi^2(8) < \frac{a}{0.2^2}) = 0.95.$$

$$P(\chi^2(8) > \frac{a}{0.2^2}) = 0.05.$$

$$a = 0.62$$



5.某厂灯泡的寿命  $X \sim N(2500, 250^2)$ , 为使灯泡的平均寿命大于2450的概率超过99%,至少应检查多少灯泡?

设应检查 $n$ 只灯泡,则

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$P(\bar{X} > 2450) = 1 - P(\bar{X} \leq 2450)$$

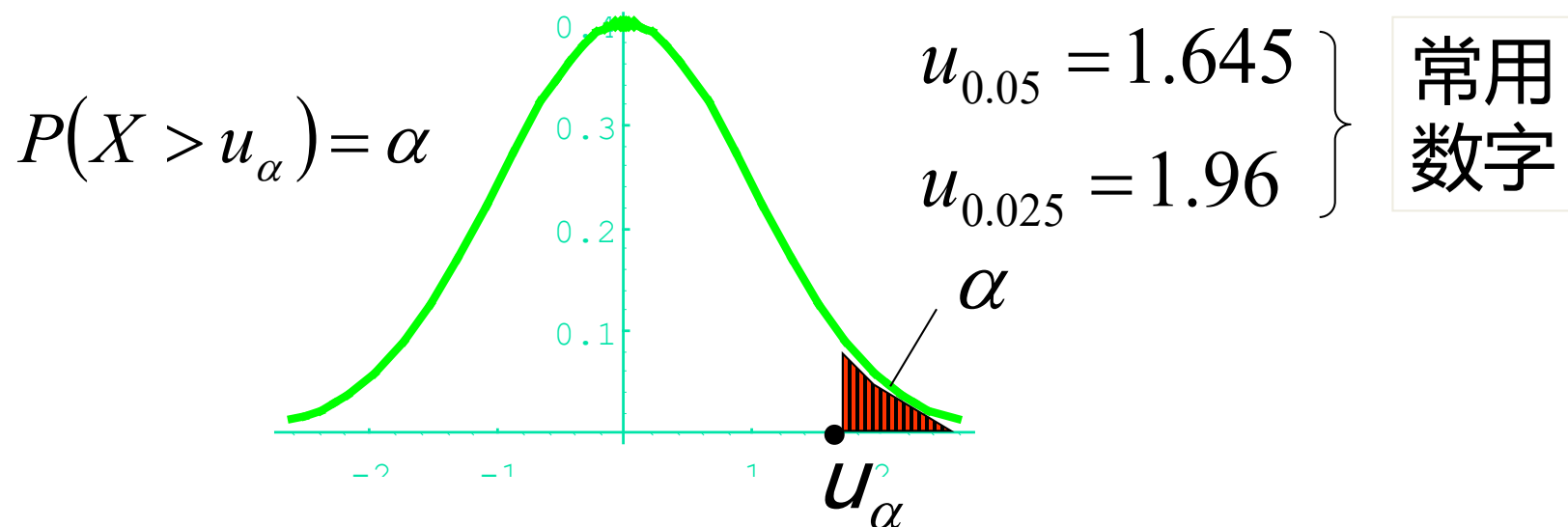
$$= 1 - \Phi\left(\frac{2450 - 2500}{250/\sqrt{n}}\right) = \Phi\left(\frac{\sqrt{n}}{5}\right) > 0.99$$

$$\therefore \frac{\sqrt{n}}{5} > 2.33$$

$$n \geq 136$$

## 7.分位数查表

### 标准正态分布的上 $\alpha$ 分位数



$$t_{0.95}(20) = t_{1-0.05}(20) = -t_{0.05}(20) = -1.7247$$

$$F_{0.95}(4,6) = \frac{1}{F_{0.05}(6,4)} = \frac{1}{6.16} = 0.1623$$

## 8.证明

$$F_{1-\alpha}(m, n) = \frac{1}{F_{\alpha}(n, m)}$$

$$P(F(m, n) \geq F_{1-\alpha}(m, n)) = 1 - \alpha$$

$$P(F(n, m) \geq F_{\alpha}(n, m)) = \alpha$$

**证** 
$$P(F(m, n) \geq F_{1-\alpha}(m, n)) = P\left(\frac{1}{F(m, n)} \leq \frac{1}{F_{1-\alpha}(m, n)}\right)$$

$$= 1 - P\left(\frac{1}{F(m, n)} \geq \frac{1}{F_{1-\alpha}(m, n)}\right) = 1 - \alpha$$

故 
$$P\left(\frac{1}{F(m, n)} \geq \frac{1}{F_{1-\alpha}(m, n)}\right) = \alpha$$

故 
$$P\left(F(n, m) \geq \frac{1}{F_{1-\alpha}(m, n)}\right) = \alpha$$

由于 
$$\frac{1}{F(m, n)} \sim F(n, m)$$

因而 
$$\frac{1}{F_{1-\alpha}(m, n)} = F_{\alpha}(n, m)$$

9. 设  $X_1, X_2, X_3, X_4$  是来自正态总体  $N(0, 2^2)$  的样本.

$$(1) Y = C[(X_1 - X_2)^2 + (X_3 + X_4)^2] \sim \chi^2(n)$$

$C, n$  为多少?

$$X_1 - X_2 \sim N(0, 8), \quad X_3 + X_4 \sim N(0, 8) \quad \text{▶}$$

$$\frac{X_1 - X_2}{\sqrt{8}} \sim N(0, 1), \quad \frac{X_3 + X_4}{\sqrt{8}} \sim N(0, 1)$$

$$\left( \frac{X_1 - X_2}{\sqrt{8}} \right)^2 + \left( \frac{X_3 + X_4}{\sqrt{8}} \right)^2 \sim \chi^2(2) \quad \text{▶}$$

$$C = \frac{1}{8}, n = 2.$$

## 3.5 正态随机变量的结论

如果随机变量  $X$  与  $Y$  相互独立, 且

$$X \sim N(\mu_1, \sigma_1^2) \quad Y \sim N(\mu_2, \sigma_2^2)$$

$$Z = X + Y,$$

则  $Z \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$



## $\chi^2$ - 分布定义

$\chi^2$  分布是由正态分布派生出来的一种分布.

定义: 设  $X_1, X_2, \dots, X_n$  相互独立, 都服从正态分布  $N(0,1)$ , 则称随机变量:

$$\chi^2 = X_1^2 + X_2^2 + \dots + X_n^2$$

所服从的分布为自由度为  $n$  的  $\chi^2$  分布.

记为  $\chi^2 \sim \chi^2(n)$





$$(2) \text{ 证明 } Z = \frac{(X_1 - X_2)^2}{(X_3 + X_4)^2} \sim F(1,1)$$

$$X_1 - X_2 \sim N(0,8), \quad X_3 + X_4 \sim N(0,8)$$

$$\frac{(X_1 - X_2)^2}{8} \sim \chi^2(1)$$

$$\frac{(X_3 + X_4)^2}{8} \sim \chi^2(1)$$

$$\longrightarrow Z = \frac{(X_1 - X_2)^2}{(X_3 + X_4)^2} \sim F(1,1)$$



# $F$ 分布定义

若  $X \sim \chi^2(n_1)$ ,  $Y \sim \chi^2(n_2)$ ,  $X, Y$  独立,

则称随机变量

$$F = \frac{X / n_1}{Y / n_2}$$

为服从自由度是  $n_1, n_2$  的  $F$ -分布,

记作  $F \sim F(n_1, n_2)$ .



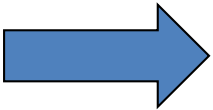

10.  $X_1, X_2, \dots, X_n$  是总体  $N(\mu, \sigma^2)$  的简单样本,

$$X_{n+1} \sim N(\mu, \sigma^2), Y = \frac{X_{n+1} - \bar{X}}{S} \sqrt{\frac{n}{n+1}} \sim \text{——}$$

$$\bar{X} \sim N(\mu, \frac{\sigma^2}{n}), X_{n+1} \sim N(\mu, \sigma^2),$$

$$X_{n+1} - \bar{X} \sim N(0, \frac{n+1}{n} \sigma^2).$$

$$\frac{X_{n+1} - \bar{X}}{\sigma} \cdot \sqrt{\frac{n}{n+1}} \sim N(0, 1). \quad \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$$


$$Y = \frac{\frac{X_{n+1} - \bar{X}}{\sigma} \cdot \sqrt{\frac{n}{n+1}}}{\sqrt{\frac{(n-1)S^2}{\sigma^2} / (n-1)}} = \frac{X_{n+1} - \bar{X}}{S} \sqrt{\frac{n}{n+1}} \sim t(n-1)$$


# $t$ 分布 (Student 分布)

$X \sim N(0,1)$ ,  $Y \sim \chi^2(n)$ ,  $X, Y$  独立, 则称随机变量

$$T = \frac{X}{\sqrt{Y/n}}$$

为服从自由度是  $n$  的  $t$ -分布, 记作  $T \sim t(n)$ .

其密度函数为

$$f(t) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi} \Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{t^2}{n}\right)^{-\frac{n+1}{2}} \quad -\infty < t < \infty$$



11. 设  $X_1, X_2, \dots, X_9$  为总体  $N(\mu, \sigma^2)$  的样本,

$$Y_1 = \frac{1}{6}(X_1 + \dots + X_6), \quad Y_2 = \frac{1}{3}(X_7 + X_8 + X_9),$$

$$S^2 = \frac{1}{2} \sum_{i=7}^9 (X_i - Y_2)^2, \quad \text{证 } Z = \frac{\sqrt{2}(Y_1 - Y_2)}{S} \sim t(2).$$

证

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \Rightarrow \begin{matrix} Y_1 \sim N\left(\mu, \frac{\sigma^2}{6}\right) \\ Y_2 \sim N\left(\mu, \frac{\sigma^2}{3}\right) \end{matrix} \Rightarrow Y_1 - Y_2 \sim N\left(0, \frac{\sigma^2}{2}\right)$$

$$\therefore U = \frac{Y_1 - Y_2}{\sigma / \sqrt{2}} \sim N(0, 1),$$

$$\therefore S^2 = \frac{1}{2} \sum_{i=7}^9 (X_i - Y_2)^2,$$

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$$

$$\therefore V = \frac{2S^2}{\sigma^2} \sim \chi^2(2),$$

$$Z = \frac{U}{\sqrt{V/2}} = \frac{\frac{Y_1 - Y_2}{\sigma / \sqrt{2}}}{\sqrt{\frac{2S^2}{2\sigma^2}}} = \frac{\sqrt{2}(Y_1 - Y_2)}{S} \sim t(2)$$

12. 设  $X_1, X_2, \dots, X_{2n} (n \geq 2)$  为从正态总体  $X \sim N(\mu, \sigma^2)$  中抽取的简单随机样本, 其样本均值为  $\bar{X} = \frac{1}{2n} \sum_{i=1}^{2n} X_i$ , 求统计量  $Y = \sum_{i=1}^n (X_i + X_{n+i} - 2\bar{X})^2$  的期望.

令  $Z_i = X_i + X_{n+i}$  则  $Z_i \sim N(2\mu, 2\sigma^2)$

$$\therefore \bar{Z} = \frac{1}{n} \sum_{i=1}^n Z_i = \frac{1}{n} \sum_{i=1}^n (X_i + X_{n+i}) = 2\bar{X}.$$

$$\text{则 } Y = \sum_{i=1}^n (Z_i - \bar{Z})^2$$

$$\frac{1}{n-1} E(Y) = E\left(\frac{1}{n-1} Y\right) = E\left(\frac{1}{n-1} \sum_{i=1}^n (Z_i - \bar{Z})^2\right)$$

$$= D(Z_i) = 2\sigma^2$$

$$E(S^2) = D(Z)$$

$$\text{所以 } E(Y) = 2(n-1)\sigma^2$$



17\*. 设  $X_1, X_2, \dots, X_n$  是来自总体  $X \sim N(\mu, \sigma^2)$  的样本, 令

$$Y = \frac{1}{n} \sum_{i=1}^n |X_i - \mu|, \text{ 试证 } E(Y) = \sqrt{\frac{2}{\pi}} \sigma, D(Y) = (1 - \frac{2}{\pi}) \frac{\sigma^2}{n}.$$

解  $Z = \frac{X_i - \mu}{\sigma} \sim N(0, 1), \quad f(z) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{z^2}{2}}$

$$\therefore EY = E\left(\frac{1}{n} \sum_{i=1}^n \sigma |Z|\right) = \frac{\sigma}{n} \sum [E|Z|] = \sqrt{\frac{2}{\pi}} \sigma$$

$$DY = \frac{\sigma^2}{n^2} \sum [D|Z|] = \frac{\sigma^2}{n^2} \sum [E(Z^2) - (E|Z|)^2]$$

$$= \frac{\sigma^2}{n^2} \sum [(E(Z^2) - E^2 Z) + E^2 Z - (E|Z|)^2]$$

$$= \frac{\sigma^2}{n^2} \sum [DZ + (EZ)^2 - (E|Z|)^2] = (1 - \frac{2}{\pi}) \cdot \frac{\sigma^2}{n}$$