

# Formula sheet 3

## Matched Pairs

### McNemar's

Test: Let  $n^* = n_{12} + n_{21}$

$$n_{12} \sim \text{Bin}(n^*, 0.5) \Rightarrow z = \frac{n_{12} - n^*/2}{\sqrt{n^* \left(\frac{1}{2}\right) \left(\frac{1}{2}\right)}} \\ = \frac{n_{12} - n_{21}}{\sqrt{n_{12} + n_{21}}} \stackrel{\text{approx.}}{\sim} N(0, 1)$$

or equivalently  $z^2 \sim \chi_1^2$

CI:

$$\underbrace{p_{1+} - p_{+1}}_{\frac{n_{12} - n_{21}}{n}} \mp z_{1-\alpha/2} \frac{1}{n} \sqrt{n_{12} + n_{21} - \frac{(n_{12} - n_{21})^2}{n}}$$

### Rater Agreement - Cohen's Kappa

$$\kappa = \frac{\sum_i \pi_{ii} - \sum_i \pi_{i+} \pi_{+i}}{1 - \sum_i \pi_{i+} \pi_{+i}}$$

Test:

$$\frac{\hat{\kappa} - 0}{s_{\hat{\kappa}}} \stackrel{\text{approx.}}{\sim} N(0, 1)$$

CI:

$$\hat{\kappa} \mp z_{\alpha/2} s_{\hat{\kappa}}$$

## Correlated, Clustered Responses

- GEE is a marginal model allowing for a correlation structure
- GLMM uses a random intercept term

Wald:

$$\bullet TS = \frac{\hat{\beta} - \beta_0}{s_{\hat{\beta}}} \stackrel{H_0}{\sim} N(0, 1)$$

$$\bullet \hat{\beta} \mp z_{1-\alpha/2} (s_{\hat{\beta}})$$

- $\sum_{i=1}^k c_i \hat{\beta}_i \mp z_{1-\alpha/2} \sqrt{\hat{V} \left( \sum_{i=1}^k c_i \hat{\beta}_i \right)}$ , where

$$V \left( \sum_{i=1}^k c_i \hat{\beta}_i \right) = \sum_{i=1}^k c_i^2 V(\hat{\beta}_i) \\ + 2 \sum_{i < j} c_i c_j \text{Cov}(\hat{\beta}_i, \hat{\beta}_j)$$

LRT-Goodness of fit:

$$D(y; \hat{\mu}) := G^2 = -2[L(\hat{\mu}; y) - L(y; y)] \xrightarrow{H_0} \chi_{df}^2$$

LRT:

$$G^2 = D(y; \hat{\mu}_0) - D(y; \hat{\mu}_1) \xrightarrow{H_0} \chi_{df}^2$$

## Loglinear Models

### 2 Way

Independence model:

$$\log(\mu_{ij}) = \lambda + \lambda_i^X + \lambda_j^Y$$

Saturated model:

$$\log(\mu_{ij}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_{ij}^{XY}$$

$$df = \frac{\text{number of Poisson counts}}{\text{number of cells in table}} - \text{number of parameters}$$

Odds ratio (for saturated model):

$$\log \left( \frac{\mu_{ij} \mu_{i'j'}}{\mu_{ij'} \mu_{i'j}} \right) = \log(\mu_{ij}) + \log(\mu_{i'j'}) - \log(\mu_{ij'}) - \log(\mu_{i'j}) \\ = \dots \\ = \lambda_{ij}^{XY} + \lambda_{i'j'}^{XY} - \lambda_{ij'}^{XY} - \lambda_{i'j}^{XY}$$

### 3 Way

- $X, Y, Z$  are *mutual independent*,  $(X, Y, Z)$  if  $\pi_{ijk} = \pi_{i++} \pi_{+j+} \pi_{++k}$

$$\log(\mu_{ijk}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z$$

- $Y$  is *jointly independent* of  $X$  and  $Z$ ,  $(XZ, Y)$   
if  $\pi_{ijk} = \pi_{+j} + \pi_{i+k}$

$$\log(\mu_{ijk}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z + \lambda_{ik}^{XZ}$$

- $X$  and  $Y$  are *conditionally independent* given  $Z$ ,  
 $(XZ, YZ)$  if  $\pi_{ij|k} = \pi_{i+|k} \pi_{+j|k}$

$$\log(\mu_{ijk}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z + \lambda_{ik}^{XZ} + \lambda_{jk}^{YZ}$$

- *Homogeneous association*,  $(XZ, XY, YZ)$  if two  
variables have the same association for all levels  
of the third, e.g.  $\pi_{ij|k} = \pi_{ij|k'}$  same  $\forall k, k'$

$$\log(\mu_{ijk}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z + \lambda_{ik}^{XZ} + \lambda_{ij}^{XY} + \lambda_{ij}^{XY}$$

- *Non restricted association*, (saturated model)  
 $(XYZ)$

$$\begin{aligned} \log(\mu_{ijk}) = & \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z \\ & + \lambda_{ik}^{XZ} + \lambda_{ij}^{YZ} + \lambda_{ij}^{XY} + \lambda_{ijk}^{XYZ} \end{aligned}$$

## Loglinear-Logit Connection

$$\begin{aligned} \log\left(\frac{P(Y=1)}{1-P(Y=1)}\right) &= \underbrace{(\lambda_1^Y - \cancel{\lambda_2^Y})}_{\alpha}^0 \\ &\quad + \underbrace{(\lambda_{i1}^{XY} - \cancel{\lambda_{i2}^{XY}})}_{\beta_i^X}^0 \\ &= \alpha + \beta_i^X \end{aligned}$$

For  $(XY, XZ, YZ)$ ,

$$\log(\mu_{ijk}) = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z + \lambda_{ik}^{XZ} + \lambda_{ij}^{YZ} + \lambda_{ij}^{XY}$$

Suppose  $Y$  is binary, treated as the response, and let

$$\pi_{ik} = P(Y=1|X=i, Z=k)$$

then

$$\begin{aligned} \text{logit}(\pi_{ik}) &= \log(\mu_{i1k}) - \log(\mu_{i2k}) \\ &= \dots \\ &= \underbrace{(\lambda_1^Y - \cancel{\lambda_2^Y})}_{\alpha}^0 + \underbrace{(\lambda_{i1}^{XY} - \cancel{\lambda_{i2}^{XY}})}_{\beta_i^X}^0 \\ &\quad + \underbrace{(\lambda_{1k}^{YZ} - \cancel{\lambda_{2k}^{YZ}})}_{\beta_k^Z}^0 \\ &= \alpha + \beta_i^X + \beta_k^Z \end{aligned}$$