# STA 4241 Lecture, Week 3

September 9th, 2021

### Overview of what we will cover

- Classification approaches
  - Logistic regression

### Classification

- In the previous lecture we focused on predicting a quantitative response variable given a set of predictors
- Classification is the term used for when we want to predict a qualitative or categorical response
- Classification is not that different from regression methods that predict a quantitative response
  - First predict the probability of being in each class
    - Similar to predicting quantitative response
  - Then classify based on this probability

- The setup will be similar to before
- We observe  $(X_1, Y_1), \ldots, (X_n, Y_n)$
- $X_i$  represents a p-dimensional set of predictors for subject i
- We want to predict  $Y_i$  with as little error as possible
  - Especially focused on predicting outcomes for a test set
- ullet Use the training data to build a classifier that predicts  $Y_i$  given  $oldsymbol{X}_i$

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#### Introduction to GLMs

- ullet Generalized linear models are used to relate a response variable Y to a set of predictors  $oldsymbol{X}$ 
  - Understand associations between predictors and outcome
  - Predict the outcome for given predictor levels
- Linear regression is actually a special case of GLMs
- GLMs are a broad class of models that can handle a variety of outcome types

### Introduction to GLMs

- A GLM has three key components
  - Random component
  - Systematic component
  - Link function
- There are a number of options for each component and we must choose one for each
- Thankfully there are standard choices for the random and link components

# Random component

- The random component specifies the probability distribution for Y
- The three most common choices are
  - Normal distribution for continuous data
  - Binomial distribution for binary data
  - Poisson distribution for count data
- Many other probability distributions work with GLMs as well
- These three are by far the most commonly used
  - Adequate for most situations encountered

- This component specifies how the explanatory variables are related to the outcome
- Typically we use a linear function, such as

$$\beta_0 + \beta_1 X_1 + \cdots + \beta_p X_p$$

- The above quantity is sometimes also known as the linear predictor
- This is the same as the manner in which we specify f(X) in linear regression
  - Same issues must be addressed such as linearity, additivity, overfitting, etc.

### Link function

- The link function specifies the functional relationship between the linear predictor and the mean of the outcome
- If we let  $\mu = E(Y|X)$ , then we specify

$$g(\mu) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

- g() is the link function used
- Each type of data has a corresponding link function that is most commonly used

- For the normal distribution, the canonical link is the identity link
  - Ordinary linear regression

$$g(\mu) = \mu = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

For the Poisson distribution, the canonical link is the log link

$$g(\mu) = \log(\mu) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

 For the bernoulli distribution, the canonical link is called the logit or logistic link

$$g(\mu) = log\left[\frac{\mu}{1-\mu}\right] = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

- Let's focus for now on binary outcomes
- The most common classification problem is one in which the outcome only takes two classes
- We will then see extensions to GLMs for multiple classes
- Will also cover classification methods that are not GLMs

- Let's consider the idea of using the identity link (linear regression) for a binary outcome with a single covariate
- The identity link might work well for some X values
- We may obtain probabilities outside of 0 or 1 for some X values
  - Clearly incorrect
- Intuitively the identity link doesn't make a lot of sense for binary data for another reason
- We don't expect a change in X to have the same impact on P(Y=1) when P(Y=1) is close to zero or 1 vs. when P(Y=1) is close to 0.5

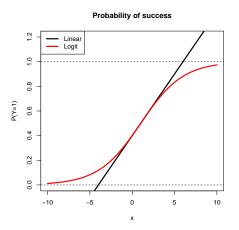
- Due to these two issues, other link functions are more common
  - Logit link
  - Probit link
- · Let's look at the logit link first

$$\log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + \beta_1 X \quad \Rightarrow \quad p(X) = \frac{e^{\beta_0 + \beta_1 X}}{1 + e^{\beta_0 + \beta_1 X}}$$

ullet Note these probabilities are constrained to lie in (0,1)

- GLMs with binary data and the logit link are called logistic regression
- Some key points about logistic regression
  - The parameter  $\beta_1$  determines the rate of increase or decrease of the curve relating X to p(X)
  - When  $\beta_1 > 0$ , p(X) increases as X increases
  - When  $\beta_1 < 0$ , p(X) decreases as X increases
- How does this compare to the linear link?

- We see the logit link respects the (0,1) boundary lines
- Changes in X have smaller effects when P(Y = 1) is closer to 0 or 1



- An alternative link is the probit link
- ullet If we let  $\Phi()$  be the CDF of a standard normal distribution, then the link function is

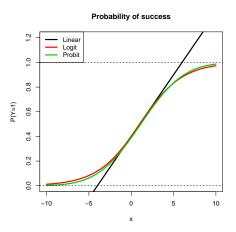
$$g(\cdot) \equiv \Phi^{-1}(\cdot)$$

which implies

$$p(X) = \Phi(\beta_0 + \beta_1 X)$$

ullet Note this again ensures that the probability is inside (0,1)

- We see that the probit looks similar to the logit
- Typically give similar answers



- Let's look at an example using data from the Challenger disaster in 1986
- Is the temperature associated with the probability of a failure of at least one primary O-ring?

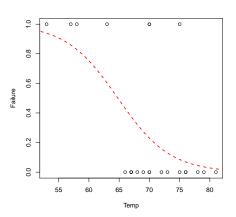
Flight	Temperature	Failure
1	66	0
2	70	1
3	69	0
4	68	0
5	67	0
6	72	0
<u>:</u>	:	:

If we fit a logistic regression model we estimate

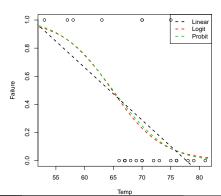
$$\log\left[\frac{\rho(X)}{1-\rho(X)}\right] = \widehat{\beta}_0 + \widehat{\beta}_1 \times \text{temperature}$$
$$= 15.04 - 0.23 \times \text{temperature}$$

- Lower temperatures seem to increase the probability of failure
- ullet The p-value for the test of  $H_0$  :  $eta_1=0$  is 0.03

- Let's look at the predicted probabilities of failure
- Our model predicts the probability of failure was over 0.99 for the temperature found on the day of the disaster (36 degrees)



- Let's try to fit the probit and linear models as well
- Very similar results between probit and logit
- ullet Linear does ok, but predicts that the probability the challenger would fail to be >1



- Be careful about extrapolating results to new input values that were not observed in the data
- The model may fit your observed data well, but not necessarily at predictor values outside the range of your data
- The challenger data point was 15 degrees lower than any point in our observed data

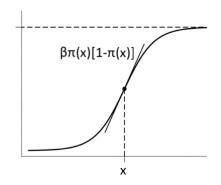
- Logistic regression is far more commonly used in applied research than probit regression
- This is mostly due to interpretation, not due to the ability to predict outcomes better
- We will focus on logistic regression here, but all ideas apply directly to probit regression
  - With the exception of interpretation!

$$\log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + \beta_1 X, \qquad p(X) = \frac{e^{\beta_0 + \beta_1 X}}{1+e^{\beta_0 + \beta_1 X}}$$

- $\beta_0$  represents the log odds of success for subjects with X=0
- $e^{\beta_0}$  is the odds of success for subjects with X=0
- ullet  $eta_1$  represents the strength of association between X and Y
  - $\beta_1 > 0$ , then  $p(X) \uparrow$  as  $X \uparrow$
  - $\beta_1 < 0$ , then  $p(X) \downarrow$  as  $X \uparrow$
  - $\beta_1 = 0$ , then  $p(X) = e^{\beta_0}/(1 + e^{\beta_0})$  which is a constant, with p(X) > 0.5 when  $\beta_0 > 0$

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- By taking derivatives, we can see that the rate of change in p(X) is  $\beta_1 p(X)(1-p(X))$
- Note that this rate of change is maximized when p(X) = 0.5



Rate of change

- $\beta_1$  has a nice interpretation in terms of odds ratios
  - Main reason why the logit link is preferred over the probit link
- $e^{\beta_1}$  is the odds ratio for a one unit change in X
- The odds of success at X are

$$\left(\frac{\rho(X)}{1-\rho(X)}\right)=e^{\beta_0+\beta_1X}$$

• At X + 1 the odds are

$$\Big(rac{p(X+1)}{1-p(X+1)}\Big) = e^{eta_0 + eta_1(X+1)} = e^{eta_0 + eta_1 X} e^{eta_1}$$

• If we take the ratio of these two, we can see that the odds ratio for X+1 versus X is simply

$$OR = \frac{p(X+1)/[1-p(X+1)]}{p(X)/[1-p(X)]} = e^{\beta_1}$$

- $\beta_1$  represents the log odds ratio comparing X+1 to X
- Very useful quantity for epidemiologists and other applied researchers
- Probit regression coefficients don't have such a nice interpretation

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- Just as in linear regression, the extension to multiple covariates is straightforward
- Model is now

$$\log\left(\frac{p(\boldsymbol{X})}{1-p(\boldsymbol{X})}\right) = \beta_0 + \sum_{j=1}^{p} \beta_j X_j$$

 The inclusion of categorical covariates, nonlinear terms, or interaction terms is the same as before

- The main difference lies in the interpretation of coefficients
- Now, the individual coefficients must be interpreted conditionally on the remaining covariates
- $\beta_1$  is the conditional log odds ratio for a one unit change in  $X_1$  while conditioning on  $X_2, \ldots, X_p$
- The odds ratio between  $X_1$  and Y is the same at all possible levels of  $X_2, \ldots X_p$
- Is this assumption reasonable?
  - Maybe not, but this model can provide a reasonable approximation to the truth in many cases

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Bias-variance trade-off

- Previously, we used the least squares criterion to estimate the parameters in linear regression
- This criterion doesn't directly apply in the binary outcome setting
- ullet Maximum likelihood is the preferred approach for estimating eta
- Interestingly, for ordinary least squares, the least squares and maximum likelihood approaches coincide
  - Our least squares estimate was also the maximum likelihood estimate (MLE)

- ullet The first step to finding the MLE of eta is writing down the likelihood of the data
- Our outcome is binary and therefore follows a bernoulli distribution
- If Y is bernoulli, then its PMF is given by

$$P(Y = y) = p^{y}(1-p)^{(1-y)}, y = 0, 1$$

where p = P(Y = 1)

• In our setting, we have parameterized this probability as a function of  $X_i$ 

$$P(Y_i = 1 | oldsymbol{X}_i) = p(oldsymbol{X}_i) = rac{e^{eta_0 + \sum_{j=1}^p eta_j X_j}}{1 + e^{eta_0 + \sum_{j=1}^p eta_j X_j}}$$

Since our data are independent, we can write our likelihood as

$$L(\boldsymbol{Y}|\boldsymbol{X}) = \prod_{i=1}^{n} L(Y_i|\boldsymbol{X}_i)$$

$$= \prod_{i=1}^{n} p(\boldsymbol{X}_i)^{Y_i} (1 - p(\boldsymbol{X}_i))^{(1-Y_i)}$$

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- ullet Note this is a function of eta
- ullet To find the MLE, we simply need to maximize  $L(m{Y}|m{X})$  with respect to  $m{eta}$
- ullet Typically this involves taking the derivative with respect to eta, setting equal to zero, and solving for eta
- Unfortunately, this does not have a nice closed-form solution

- One way to proceed is through optimization techniques to find the minimum numerically
  - Newton-Raphson method, many others
  - $\bullet$  Might have trouble finding the maximum as the dimension of  $\beta$  grows
- Another approach is through iteratively re-weighted least squares
  - Successively fit weighted least squares to the model until it converges
  - Used in the glm function in R

# Classification in logistic models

- ullet Once we have an estimate eta, we can proceed with classification
- As discussed in previous lectures, classification should proceed by choosing the most likely class for the outcome
- With only two classes, this amounts to predictions of the form

$$\widehat{Y}_i = \begin{cases} 1, & \widehat{p}(\mathbf{X}_i) \ge 0.5 \\ 0, & \widehat{p}(\mathbf{X}_i) < 0.5 \end{cases}$$

# Classification in logistic models

This can also be framed in terms of the linear predictor of the model

$$\widehat{Y}_i = \begin{cases} 1, & \widehat{\beta}_0 + \sum_{j=1}^p \widehat{\beta}_j X_{ij} \ge 0 \\ 0, & \widehat{\beta}_0 + \sum_{j=1}^p \widehat{\beta}_j X_{ij} < 0 \end{cases}$$

- If our model is correctly specified (we have the correct distribution of Y|X), this should approximate the Bayes classifier
  - Best we could hope to do
- We will see later that other classification techniques lead to similar classifiers

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