

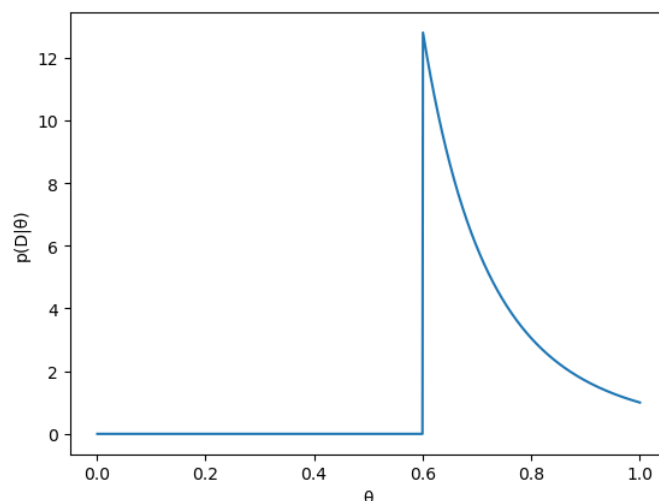
问题1

(1) 可以写出最大似然函数为

$$l(\theta) = \begin{cases} \frac{1}{\theta^n} & 0 \leq x_k \leq \theta (k = 1, 2, \dots, n) \\ 0 & \text{others} \end{cases}$$

易知, 欲使最大似然函数取最大值, θ 应取最最小值, 同时, 由于 $\theta \geq x_k (k = 1, 2, \dots, n)$, 故 θ 的最大似然估计值为 $\max(D)$ 。

(2) 由 (1) 知, $\theta = \max(D) = 0.6$ 时, 最大似然估计取最大值, 在 $n=5$ 时, 最大似然函数如下图所示, 因为在 $\max(D) = 0.6$ 成立的情况下, 其他样本点的值不影响最大似然估计, 因此无需知道。



问题2

(1) μ 的后验概率为

$$\begin{aligned} p(u | D) &= \frac{p(D | u)p(u)}{\int p(D | u)p(u)du} \\ &= \alpha \prod_{k=1}^n p(x_k | u) p(u) \\ &= \alpha' \exp \left(-\frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \Sigma^{-1} (x_k - u) + (u - m_0)^t \Sigma^{-1} (u - m_0) \right) \right) \\ \ln p(u | D) &= \ln \alpha' - \frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \Sigma^{-1} (x_k - u) + (u - m_0)^t \Sigma^{-1} (u - m_0) \right) \end{aligned}$$

去掉无关项后, 最大后验估计(MAP)为

$$\hat{u} = \arg \max_u \left(-\frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \Sigma^{-1} (x_k - u) + (u - m_0)^t \Sigma^{-1} (u - m_0) \right) \right) \quad (1)$$

(2) 在经过线性变换后, 各个相关的参数变化为

$$\begin{aligned}
\boldsymbol{\mu}' &= \mathbf{A}\boldsymbol{\mu} \\
\boldsymbol{\Sigma}' &= \mathcal{E} \left[(\mathbf{x}' - \boldsymbol{\mu}') (\mathbf{x}' - \boldsymbol{\mu}')^t \right] \\
&= \mathcal{E} \left[(\mathbf{A}\mathbf{x} - \mathbf{A}\boldsymbol{\mu}) (\mathbf{A}\mathbf{x} - \mathbf{A}\boldsymbol{\mu})^t \right] \\
&= \mathcal{E} \left[\mathbf{A} (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^t \mathbf{A}^t \right] \\
&= \mathbf{A} \mathcal{E} \left[(\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^t \right] \mathbf{A}^t \\
&= \mathbf{A} \boldsymbol{\Sigma} \mathbf{A}^t
\end{aligned}$$

同理, $m_0' = \mathbf{A}m_0$, $\Sigma_0' = \mathbf{A}\Sigma_0\mathbf{A}^t$

故最大似然估计变为

$$\begin{aligned}
\hat{u}' &= \arg \max_u \left(-\frac{1}{2} \left(\sum_{k=1}^n (Ax_k - Au)^t (\mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^t)^{-1} (Ax_k - Au) + (Au - Am_0)^t \Sigma_0^{-1} (Au - Am_0) \right) \right) \\
&= \arg \max_u \left(-\frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \mathbf{A}^t \left((\mathbf{A}^{-1})^t \boldsymbol{\Sigma}^{-1} \mathbf{A}^{-1} \right) \mathbf{A} (x_k - u) + (u - m_0)^t \mathbf{A}^t \left((\mathbf{A}^{-1})^t \boldsymbol{\Sigma}_0^{-1} \mathbf{A}^{-1} \right) \mathbf{A} (u - m_0) \right) \right) \\
&= \arg \max_u \left(-\frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \left(\mathbf{A}^t (\mathbf{A}^{-1})^t \right) \boldsymbol{\Sigma}^{-1} (\mathbf{A}^{-1} \mathbf{A}) (x_k - u) + (u - m_0)^t \left(\mathbf{A}^t (\mathbf{A}^{-1})^t \right) \boldsymbol{\Sigma}_0^{-1} (\mathbf{A}^{-1} \mathbf{A}) (u - m_0) \right) \right) \quad (2) \\
&= \arg \max_u \left(-\frac{1}{2} \left(\sum_{k=1}^n (x_k - u)^t \boldsymbol{\Sigma}^{-1} (x_k - u) + (u - m_0)^t \boldsymbol{\Sigma}_0^{-1} (u - m_0) \right) \right)
\end{aligned}$$

式(2)与式(1)相等, 故对于 u' 也能给出正确的估计。

问题3

$$\begin{aligned}
Q(\theta; \theta^0) &= \mathcal{E}_{x_{32}} [\ln p(x_g, x_b; \theta) \mid \theta^0, \mathcal{D}_g] \\
&= \int_{-\infty}^{\infty} (\ln p(x_1 \mid \theta) + \ln p(x_2 \mid \theta) + \ln p(x_3 \mid \theta)) p(x_{32} \mid \theta^0, x_{31} = 2) dx_{32} \\
&= \ln p(x_1 \mid \theta) + \ln p(x_2 \mid \theta) + \int_{-\infty}^{\infty} \ln p(x_3 \mid \theta) \cdot p(x_{32} \mid \theta^0, x_{31} = 2) dx_{32} \\
&= \ln p(x_1 \mid \theta) + \ln p(x_2 \mid \theta) + \int_{-\infty}^{\infty} \ln p\left(\begin{pmatrix} 2 \\ x_{32} \end{pmatrix} \mid \theta\right) \cdot \frac{p\left(\begin{pmatrix} 2 \\ x_{32} \end{pmatrix} \mid \theta^0\right)}{\int_{-\infty}^{\infty} p\left(\begin{pmatrix} 2 \\ x'_{32} \end{pmatrix} \mid \theta^0\right) d'_{32}} dx_{32} \\
&= \ln p(x_1 \mid \theta) + \ln p(x_2 \mid \theta) + \frac{1}{4} \int_{-\infty}^{\infty} \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) dx_{32} \\
&= \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) + \ln\left(\frac{1}{\theta_1} e^{-\frac{3}{\theta_1}} \frac{1}{\theta_2}\right) + \frac{1}{4} \int_{-\infty}^{\infty} \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) dx_{32}
\end{aligned}$$

由于已知的样本中 $x_2 = 3$, 故当 $\theta_2 < 3$ 时, 不满足条件。

当 $3 \leq \theta_2 \leq 4$ 时,

$$\begin{aligned}
Q(\theta; \theta^0) &= \ln\left(\frac{1}{\theta_1} e^{-\frac{1}{\theta_1}} \frac{1}{\theta_2}\right) + \ln\left(\frac{1}{\theta_1} e^{-\frac{3}{\theta_1}} \frac{1}{\theta_2}\right) + \frac{1}{4} \int_0^{\theta_2} \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) dx_{32} \\
&= \ln\left(\frac{1}{\theta_1} e^{-\frac{1}{\theta_1}} \frac{1}{\theta_2}\right) + \ln\left(\frac{1}{\theta_1} e^{-\frac{3}{\theta_1}} \frac{1}{\theta_2}\right) + \frac{\theta_2}{4} \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) \\
&= -2 \ln(\theta_1 \theta_2) - \frac{4}{\theta_1} - \frac{\theta_2}{4} \ln(\theta_1 \theta_2) - \frac{\theta_2}{2\theta_1} \\
\frac{\partial Q(\theta; \theta^0)}{\partial \theta_1} &= -\frac{2}{\theta_1} + \frac{4}{\theta_1^2} - \frac{\theta_2}{4\theta_1} + \frac{\theta_2}{2\theta_1^2} \\
&= \frac{-8\theta_1 + 16 - \theta_1 \theta_2 + 2\theta_2}{4\theta_1^2} \\
&= \frac{(2 - \theta_1)(\theta_2 + 8)}{4\theta_1^2}
\end{aligned}$$

故当 $\theta_1 = 2$ 时, 取极值。此时

$$Q(\theta; \theta^0) = -2 - \left(2 \ln 2\theta_2 + \frac{1}{4} \theta_2 (1 + \ln 2\theta_2) \right)$$

因为 $Q(\theta; \theta^0)$ 在 $\theta_2 > 0$ 时单调递减，故当 $\theta_2 = 3$ 时， $Q(\theta; \theta^0)$ 取得最大值 -6.75

当 $\theta_2 > 4$ 时，

$$\begin{aligned} Q(\theta; \theta^0) &= \ln\left(\frac{1}{\theta_1} e^{-\frac{1}{\theta_1}} \frac{1}{\theta_2}\right) + \ln\left(\frac{1}{\theta_1} e^{-\frac{3}{\theta_1}} \frac{1}{\theta_2}\right) + \frac{1}{4} \int_0^4 \ln\left(\frac{1}{\theta_1} e^{-\frac{2}{\theta_1}} \frac{1}{\theta_2}\right) dx_{32} \\ &= -\frac{6}{\theta_1} - 3 \ln(\theta_1 \theta_2) \\ \frac{\partial Q(\theta; \theta^0)}{\partial \theta_1} &= \frac{6 - 3\theta_1}{\theta_1^2} \end{aligned}$$

故当 $\theta_1 = 2$ 时取得极值，此时 $Q(\theta; \theta^0) = -3 - 3 \ln(2\theta_2)$ 为减函数，故当 $\theta_2 = 4$ 时取得最大值 -9.24

1. 在 E-step 中

$$Q(\theta; \theta^0) = \begin{cases} \text{无意义} & \theta_2 < 3 \\ -2 \ln(\theta_1 \theta_2) - \frac{4}{\theta_1} - \frac{\theta_2}{4} \ln(\theta_1 \theta_2) - \frac{\theta_2}{2\theta_1} & 3 \leq \theta_2 \leq 4 \\ -\frac{6}{\theta_1} - 3 \ln(\theta_1 \theta_2) & \theta_2 > 4 \end{cases}$$

2. M-step 运算如上，综上 $\theta = \begin{pmatrix} 2 \\ 3 \end{pmatrix}$

问题4

$$\begin{aligned} \hat{a}_{ij} &= \frac{\sum_{t=1}^T \gamma_{ij}(t)}{\sum_{t=1}^T \sum_k \gamma_{ik}(t)} \\ \hat{b}_{ij} &= \frac{\sum_{t=1}^T \gamma_{ij}(t)}{\sum_{t=1}^T \gamma_{ij}(t)} \\ \gamma_{ij}(t) &= \frac{\alpha_i(t-1) a_{ij} b_{ij} \beta_i(t)}{P(V^T | M)} \end{aligned}$$

$\alpha_i(t)$'s 和 $P(V^T | M)$ 通过前向算法计算得到，分别需要 $O(c^2 T)$ 的复杂度，

```

For t = T to 1 (by - 1)
  For i = 1 to c
     $\beta_i(t) = \sum_j a_{ij} b_{jk} v(t+1) \beta_j(t+1)$ 
  End
End

```

$\beta_i(t)$ 的计算由上式得到，也需要 $O(c^2 T)$ 的复杂度

此时，计算 γ_{ij} 也需要 $O(c^2 T)$ 的复杂度

$$\underbrace{O(c^2 T)}_{\alpha_i(t)'s} + \underbrace{O(c^2 T)}_{\beta_i(t)'s} + \underbrace{O(c^2 T)}_{\gamma_{ij}(t)'s} = O(c^2 T)$$

因此共需要 $O(c^2 T)$ 的复杂度。

问题5

(1) 根据题目可以写出，

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

故估计的期望值

$$\begin{aligned}\bar{p}_n(x) &= \mathcal{E}[p_n(x)] = \frac{1}{nh_n} \sum_{i=1}^n \mathcal{E}\left[\varphi\left(\frac{x-x_i}{h_n}\right)\right] \\ &= \frac{1}{h_n} \int_{-\infty}^{\infty} \varphi\left(\frac{x-v}{h_n}\right) p(v) dv \\ &= \frac{1}{h_n} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-v}{h_n}\right)^2\right] \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{v-\mu}{\sigma}\right)^2\right] dv \\ &= \frac{1}{2\pi h_n \sigma} \exp\left[-\frac{1}{2}\left(\frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2}\right)\right] \int_{-\infty}^{\infty} \exp\left[-\frac{1}{2}v^2\left(\frac{1}{h_n^2} + \frac{1}{\sigma^2}\right) - 2v\left(\frac{x}{h_n^2} + \frac{\mu}{\sigma^2}\right)\right] dv \\ &= \frac{1}{2\pi h_n \sigma} \exp\left[-\frac{1}{2}\left(\frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2}\right) + \frac{1}{2}\frac{\alpha^2}{\theta^2}\right] \int_{-\infty}^{\infty} \exp\left[-\frac{1}{2}\left(\frac{v-\alpha}{\theta}\right)^2\right] dv\end{aligned}$$

$$\text{其中 } \theta^2 = \frac{1}{1/h_n^2 + 1/\sigma^2} = \frac{h_n^2 \sigma^2}{h_n^2 + \sigma^2}, \quad \alpha = \theta^2 \left(\frac{x}{h_n^2} + \frac{\mu}{\sigma^2}\right)$$

$$\begin{aligned}\bar{p}_n(x) &= \frac{\sqrt{2\pi}\theta}{2\pi h_n \sigma} \exp\left[-\frac{1}{2}\left(\frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2}\right) + \frac{1}{2}\frac{\alpha^2}{\theta^2}\right] \\ &= \frac{1}{\sqrt{2\pi} h_n \sigma} \frac{h_n \sigma}{\sqrt{h_n^2 + \sigma^2}} \exp\left[-\frac{1}{2}\left(\frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2} - \frac{\alpha^2}{\theta^2}\right)\right]\end{aligned}$$

由于

$$\begin{aligned}\frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2} - \frac{\alpha^2}{\theta^2} &= \frac{x^2}{h_n^2} + \frac{\mu^2}{\sigma^2} - \frac{\theta^4}{\theta^2} \left(\frac{x}{h_n^2} + \frac{\mu}{\sigma^2}\right)^2 \\ &= \frac{x^2 h_n^2}{(h_n^2 + \sigma^2) h_n^2} + \frac{\mu^2 \sigma^2}{(h_n^2 + \sigma^2) \sigma^2} - \frac{2x\mu}{h_n^2 + \sigma^2} \\ &= \frac{(x-\mu)^2}{h_n^2 + \sigma^2}\end{aligned}$$

代入原式得到

$$\bar{p}_n(x) = \frac{1}{\sqrt{2\pi} \sqrt{h_n^2 + \sigma^2}} \exp\left[-\frac{1}{2} \frac{(x-\mu)^2}{h_n^2 + \sigma^2}\right]$$

因此 $\bar{p}_n(x) \sim N(\mu, h_n^2 + \sigma^2)$ 得证

(2)

$$\begin{aligned}\text{Var}[p_n(x)] &= \text{Var}\left[\frac{1}{nh_n} \sum_{i=1}^n \varphi\left(\frac{x-x_i}{h_n}\right)\right] \\ &= \frac{1}{n^2 h_n^2} \sum_{i=1}^n \text{Var}\left[\varphi\left(\frac{x-x_i}{h_n}\right)\right] \\ &= \frac{1}{nh_n^2} \text{Var}\left[\varphi\left(\frac{x-v}{h_n}\right)\right] \\ &= \frac{1}{nh_n^2} \left\{ \mathcal{E}\left[\varphi^2\left(\frac{x-v}{h_n}\right)\right] - \left(\mathcal{E}\left[\varphi\left(\frac{x-v}{h_n}\right)\right]\right)^2 \right\}\end{aligned}$$

$$\begin{aligned}
\mathcal{E}\left[\varphi^2\left(\frac{x-v}{h_n}\right)\right] &= \int \varphi^2\left(\frac{x-v}{h_n}\right) p(v) dv \\
&= \int_{-\infty}^{\infty} \frac{1}{2\pi} \exp\left[-\left(\frac{x-v}{h_n}\right)^2\right] \exp\left[-\frac{1}{2}\left(\frac{v-\mu}{\sigma}\right)^2\right] \frac{1}{\sqrt{2\pi}\sigma} dv \\
&= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-v}{h_n/\sqrt{2}}\right)^2\right] \exp\left[-\frac{1}{2}\left(\frac{v-\mu}{\sigma}\right)^2\right] \frac{1}{\sqrt{2\pi}\sigma} dv \\
&= \frac{h_n/\sqrt{2}}{2\pi\sqrt{h_n^2/2 + \sigma^2}} \exp\left[-\frac{1}{2}\frac{(x-\mu)^2}{h_n^2/2 + \sigma^2}\right]
\end{aligned}$$

因此

$$\begin{aligned}
\frac{1}{nh_n^2} \mathcal{E}\left[\varphi^2\left(\frac{x-v}{h_n}\right)\right] &= \frac{1}{nh_n} \frac{1}{2\sqrt{2\pi}\sqrt{h_n^2/2 + \sigma^2}} \exp\left[-\frac{1}{2}\frac{(x-\mu)^2}{h_n^2/2 + \sigma^2}\right] \\
&= \frac{1}{nh_n} \frac{1}{2\sqrt{\pi}} \frac{1}{\sqrt{2\pi}\sqrt{h_n^2/2 + \sigma^2}} \exp\left[-\frac{1}{2}\frac{(x-\mu)^2}{h_n^2/2 + \sigma^2}\right]
\end{aligned}$$

当 $h_n \rightarrow 0$ 时, $\sqrt{h_n^2/2 + \sigma^2} \simeq \sigma$ 上式可以近似为

$$\begin{aligned}
\frac{1}{nh_n^2} \mathcal{E}\left[\varphi^2\left(\frac{x-v}{h_n}\right)\right] &\simeq \frac{1}{nh_n} \frac{1}{2\sqrt{\pi}} \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right] \\
&= \frac{1}{2nh_n\sqrt{\pi}} p(x)
\end{aligned}$$

$$\begin{aligned}
\frac{1}{nh_n^2} \left\{ \mathcal{E}\left[\varphi\left(\frac{x-v}{h_n}\right)\right] \right\}^2 &= \frac{1}{nh_n^2} h_n^2 \left(\frac{1}{\sqrt{2\pi}\sqrt{h_n^2 + \sigma^2}} \exp\left[-\frac{1}{2}\frac{(x-\mu)^2}{h_n^2 + \sigma^2}\right] \right)^2 \\
&= \frac{h_n}{nh_n} \frac{1}{2\pi(h_n^2 + \sigma^2)} \exp\left[-\frac{(x-\mu)^2}{h_n^2 + \sigma^2}\right] \\
&\simeq \frac{h_n}{nh_n} \frac{1}{2\pi\sigma^2} \exp\left[-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right] \simeq 0
\end{aligned}$$

因此, 题目得证:

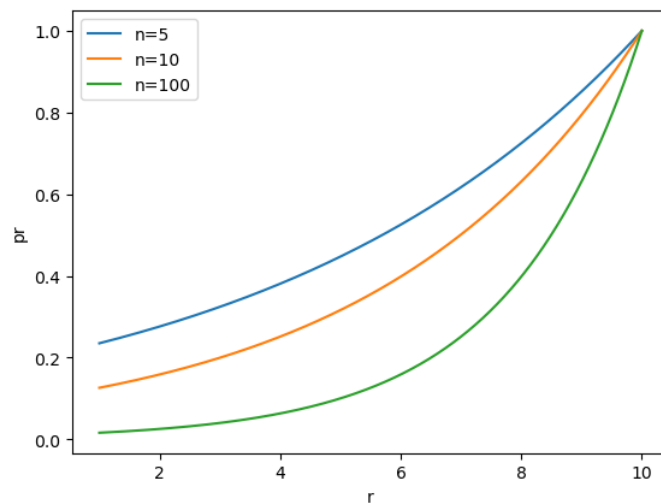
$$\text{Var}[P_n(x)] \simeq \frac{p(x)}{2nh_n\sqrt{\pi}}$$

问题6

(1) 假设 n 个样本点相互独立且均匀分布, 则每个维度有 $n^{\frac{1}{d}}$ 个样本点, 其中 r 个维度有 $n^{\frac{r}{d}}$ 个样本点。选中正确的数据作为最近邻的概率为

$$p = \frac{n^{\frac{r}{d}}}{n} = n^{\frac{r}{d}-1}$$

故当 $d=10$ 时 $p = n^{\frac{r}{10}-1}$, 取不同的 n 画图得到概率 p 关于 r 的曲线如下:



(2) 贝叶斯决策面为

$$\begin{aligned}
 g_1(x) &= g_2(x) \\
 p(\omega_1 | x) &= p(\omega_2 | x) \\
 \frac{p(x | \omega_1) p(\omega_1)}{p(x)} &= \frac{p(x | \omega_2) p(\omega_2)}{p(x)} \\
 p(x | \omega_1) p(\omega_1) &= p(x | \omega_2) p(\omega_2)
 \end{aligned}$$

由于两类样本数均为 $n/2$,故可以假设两类先验概率相等,即 $p(w_1) = p(w_2)$

由题目知概率密度为斜坡函数的乘积, 故可以得到下式

$$\begin{aligned}
 p(x | \omega_1) &= \prod_{i=1}^d k_i x_i \\
 p(x | \omega_2) &= \prod_{i=1}^d k_i (1 - x_i)
 \end{aligned}$$

因此贝叶斯决策面简化为

$$\prod_{i=1}^d x_i = \prod_{i=1}^d (1 - x_i)$$