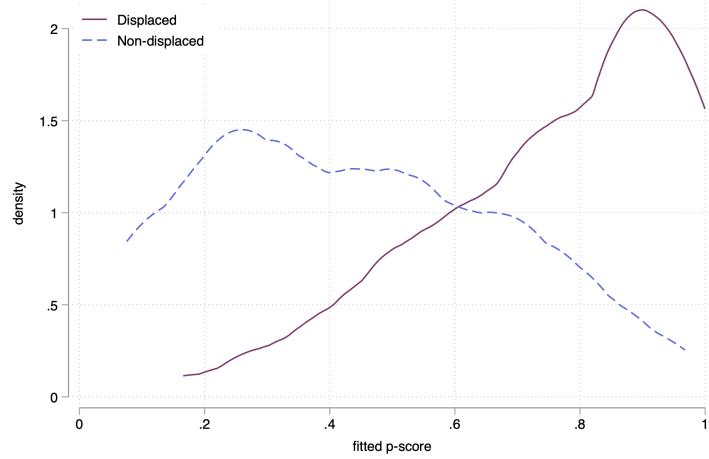


APPENDIX AND SUPPLEMENTARY MATERIAL

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A ADDITIONAL FIGURES AND TABLES

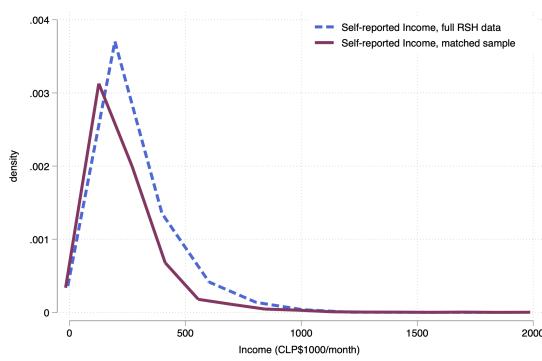
Figure A.1: Distribution of the probability of displacement by treatment



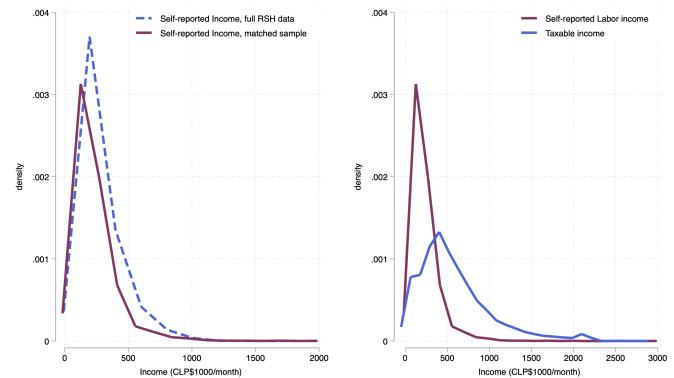
Notes: Figure plots the fitted values of a logit regression that includes controls from regression (4) in Table A.1 by treatment.

Figure A.2: Labor income distribution across different samples

(a) Income distribution in the RSH and matched sample

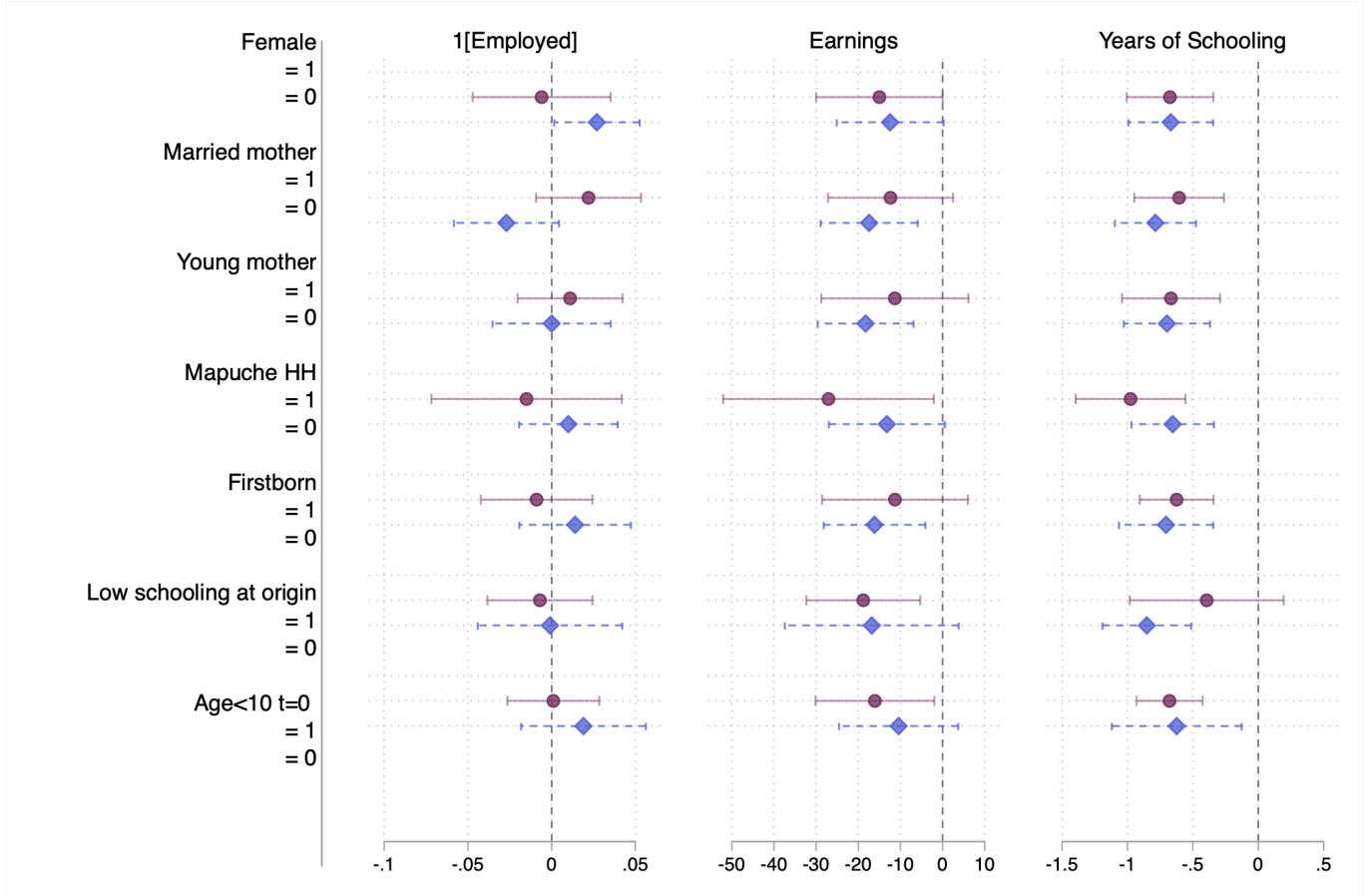


(b) Income distribution in matched sample



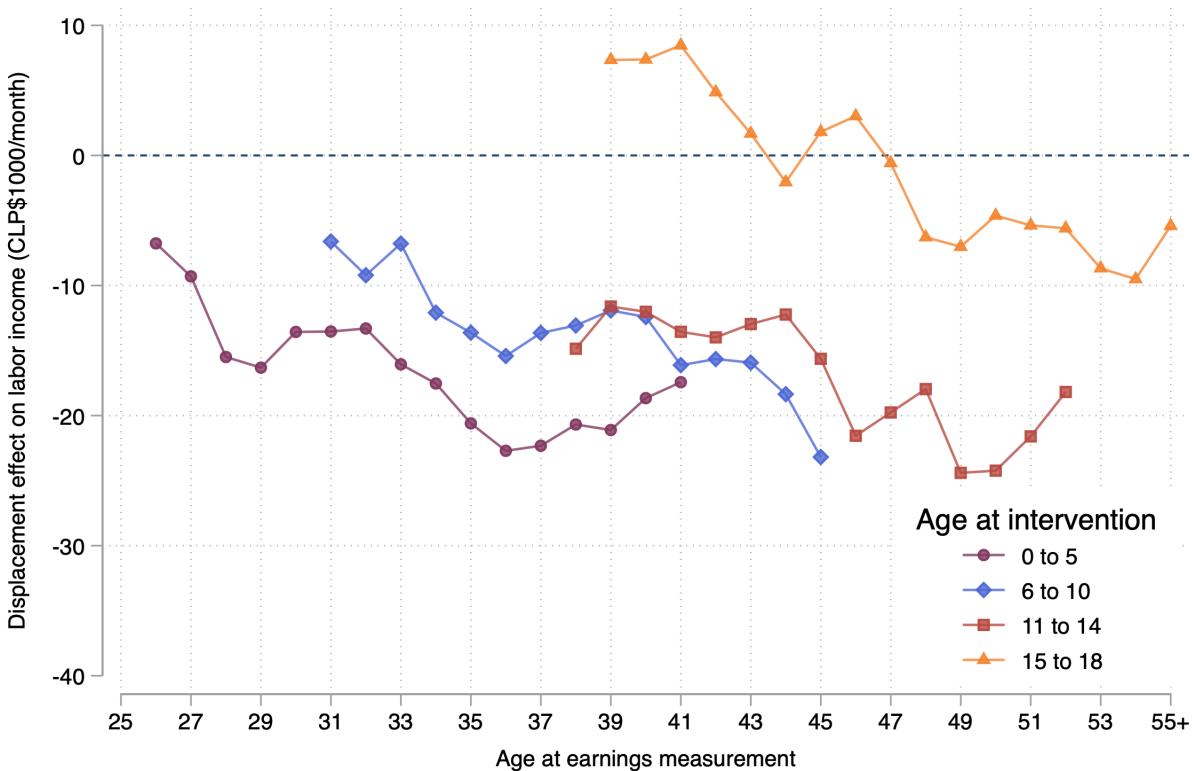
Notes: Income data for the year 2018. Matched sample stands for children aged 0 to 18 at baseline who are matched with the RSH data, and who are 18 or older in 2018. "Full RSH" corresponds to all individuals aged 21 to 60 in the RSH in year 2018 in Greater Santiago.

Figure A.3: Displacement effect by demographic groups on main outcomes



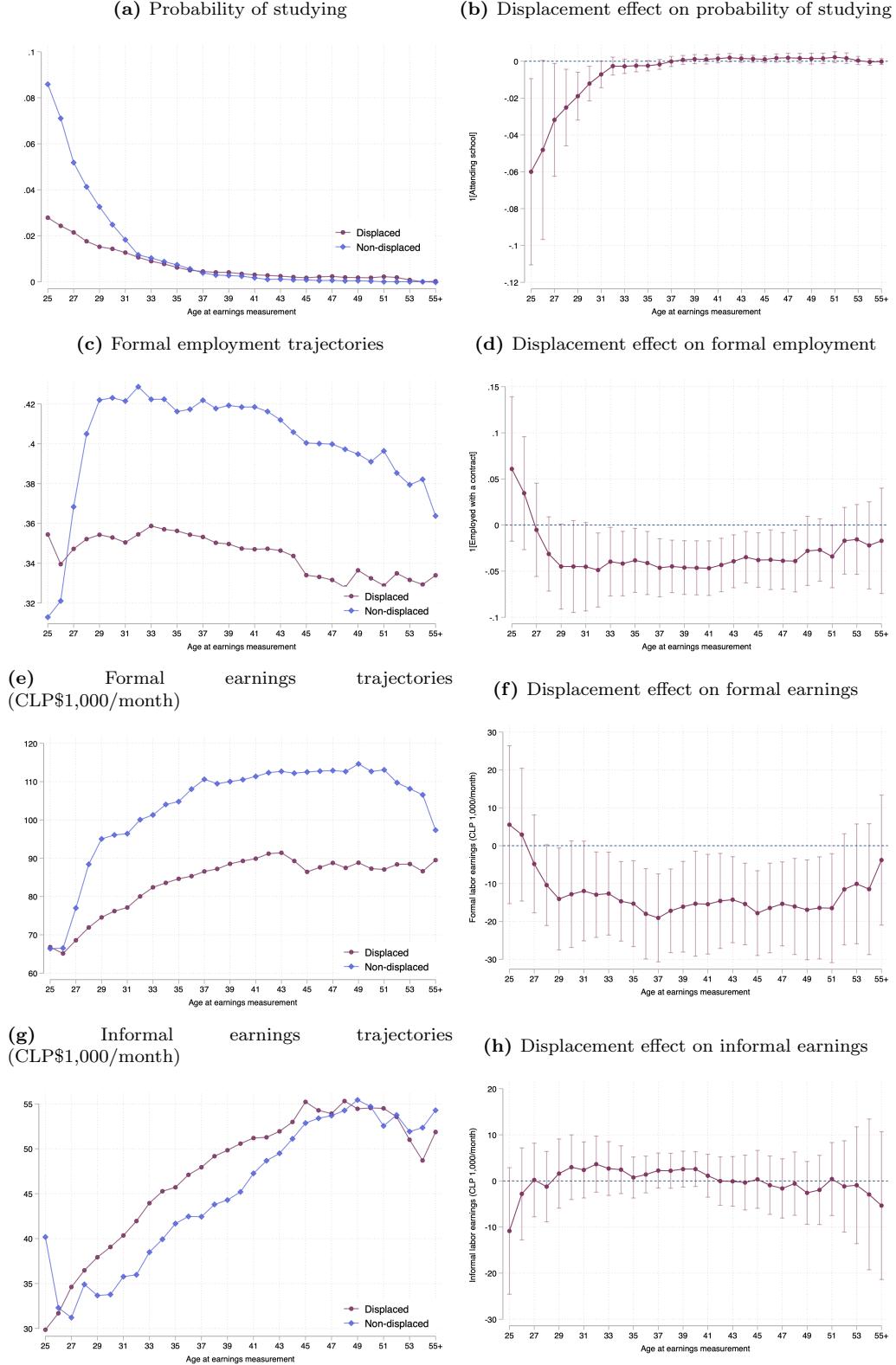
Notes: Regressions for children aged 0 to 18 that are matched to the RSH, and report non-missing schooling. Standard errors are clustered by municipality of origin. Controls include the following: female, mother head of household, single head of household, number of siblings, Mapuche lastname, cohort fixed effects, and time fixed effects. The figure plots the displacement coefficient and its 95% confidence interval resulting from estimating equation (1) stratified by demographic groups. Single mother is measured at the time of intervention, "young mother" stands for mothers younger than 25 (sample median) at the time their child is born, and "Low schooling" at origin stands for municipalities of origin where the population's average schooling is below the sample median.

Figure A.4: Displacement effect on earnings by age at earnings measurement and cohort



Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. Controls include the following: female, mother head of household, married head of household, number of siblings, firstborn dummy, head of household's marital status unknown, and year of birth fixed effects. Figure plots coefficients β_τ and their 95% confidence intervals from the regression: $y_{it} = \sum_{\tau=25}^{55} \beta_\tau Displaced * 1[Age = \tau] + \sum_{\tau=25}^{55} \delta_\tau 1[Age] + \psi_o + X'_{it} \gamma + u_{it}$, for each of the four groups by the age at intervention.

Figure A.5: Displacement effects on labor market outcomes by age at earnings measurement



Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. Controls include the following: female, mother head of household, married head of household, number of siblings, firstborn dummy, head of household's marital status unknown, and year of birth fixed effects. We estimate $y_{it} = \sum_{\tau=25}^{55} \beta_\tau Displaced * 1[Age = \tau] + \sum_{\tau=25}^{55} \delta_\tau 1[Age] + \psi_o + X'_{it} \gamma + u_{it}$. Figures (a), (c), (e) and (g) plot the predicted trajectories for the displaced and non-displaced children between ages 25 to 55. Figures (b), (d), (f), and (h) plot coefficients β_τ and their 95% confidence intervals.

Figure A.6: Displacement effect by age at intervention and structural break

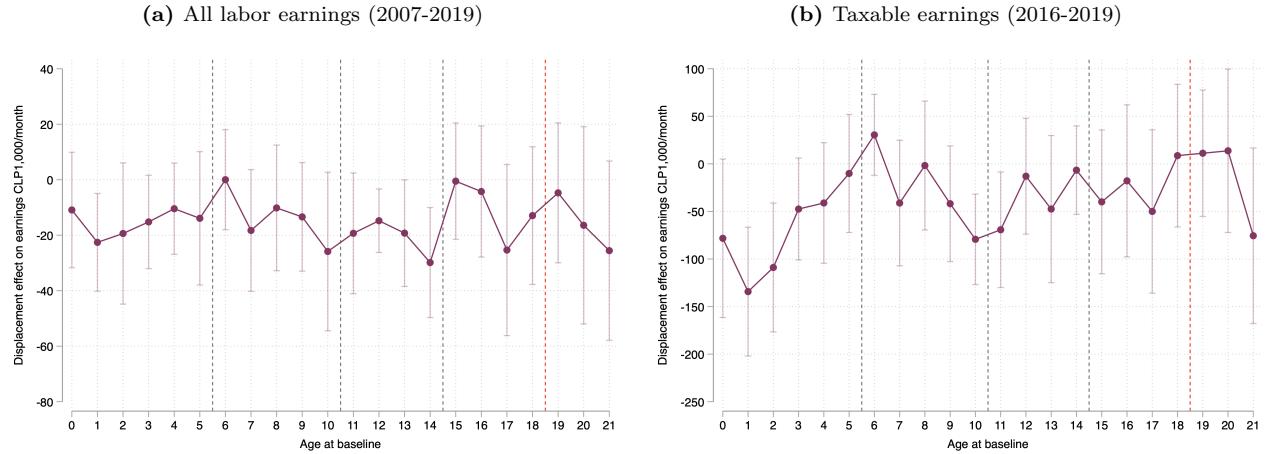


Figure A.7: Relationship between fragmentation and polarization for neighborhoods in the program

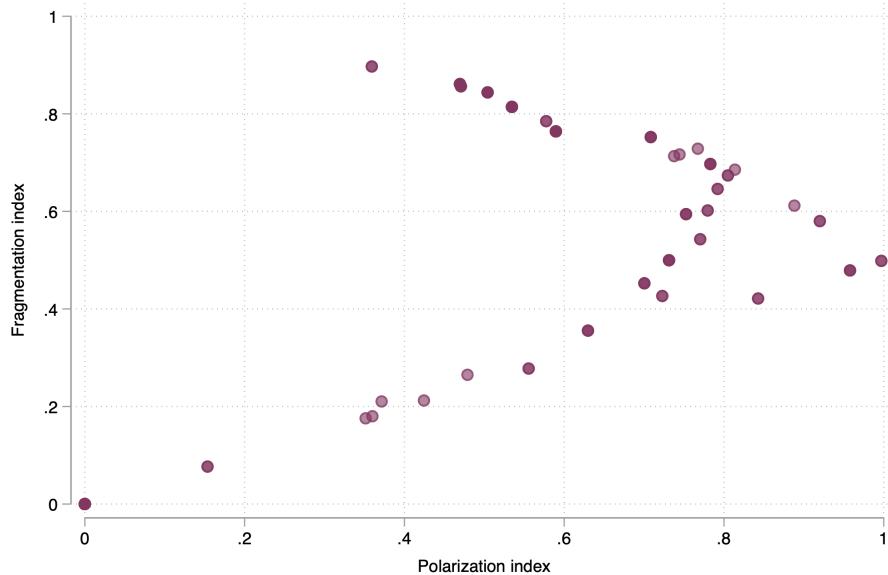
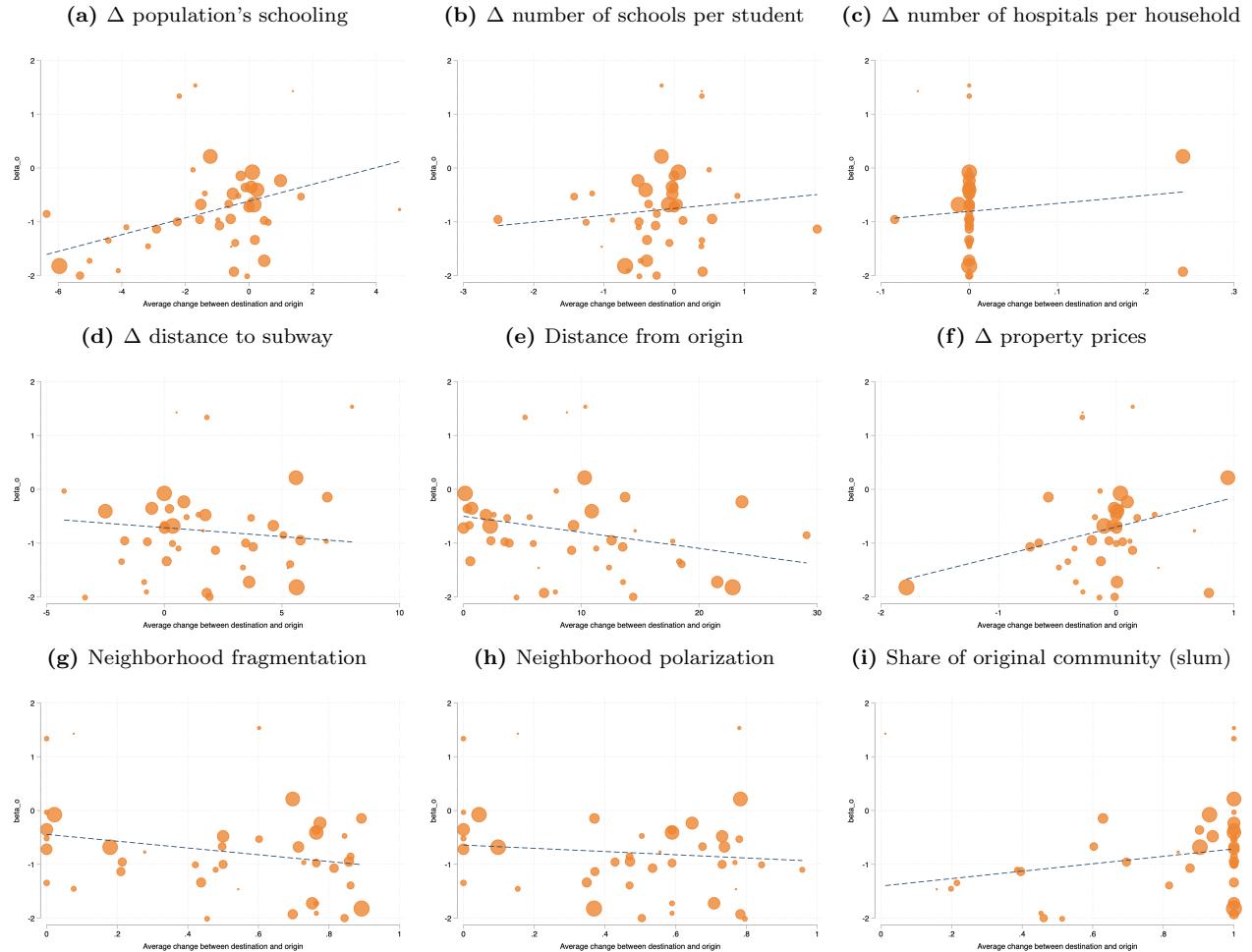
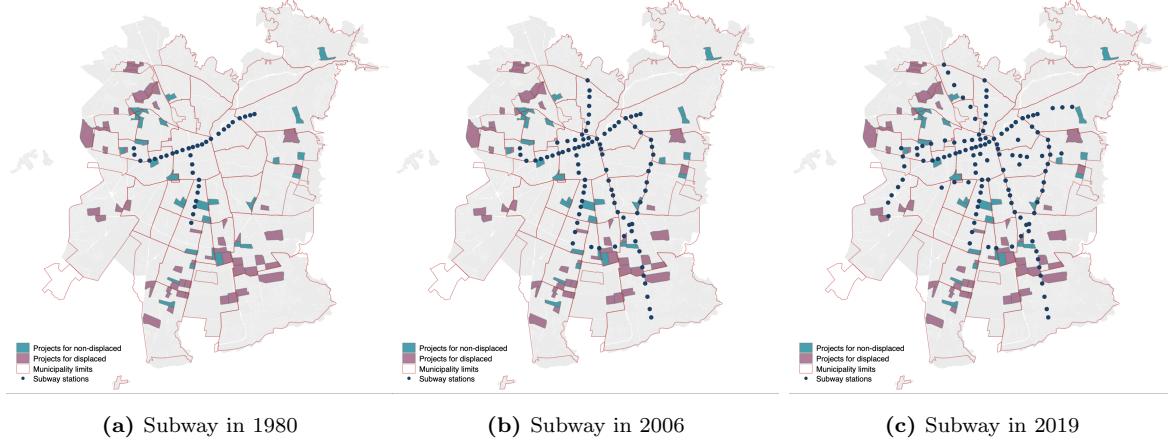


Figure A.8: Displacement effect on years of schooling by municipality of origin and changes in location attributes



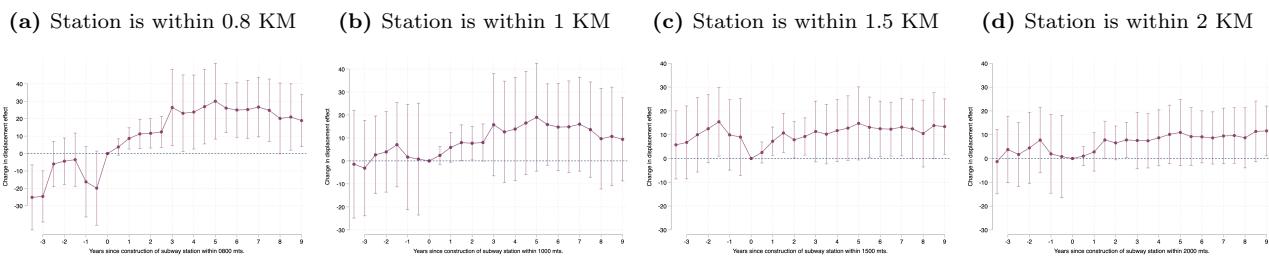
Notes: The figures plot displacement coefficients on years of education stratified by municipality of origin and destination against average changes in location attributes. Coefficients are estimated using the following regression: $y_{it} = \sum_{o=1, d=1} \beta_{od} Displaced * 1[Origin = o, Destination = d] + X'_{it} \gamma + u_{itod}$, where o indexes the municipality of origin, and d municipality of destination for child i . Changes in attributes (x-axis) are computed as $\bar{\Delta}_{od} = \sum_{o=1, d=1}^{30} \Delta_{iod}$. Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Controls include the following: female, mother head of household, married head of household, number of siblings, firstborn dummy, head of household's marital status unknown, and year of birth fixed effects. Coefficients β_{od} are weighted by the number of observations in each cell. Figure V repeats the exercise for earnings.

Figure A.9: Location of public housing projects and subway stations



Notes: This figure shows the rollout of subway stations in Greater Santiago from 1980 to 2019. Red lines represent the urban limits of Greater Santiago and its municipalities in 2019. Colored areas correspond to neighborhoods created by the Program for Urban Marginality between 1979 and 1985. Purple areas correspond to projects that received displaced families, and green areas correspond to projects for the non-displaced families. Blue circles are the locations of subway stations at each moment in time. The data to construct this map come from MINVU (1979), Molina (1986), FLACSO (1982, 1986), and Metro de Santiago.

Figure A.10: Roll out of subway stations between 2007 and 2019 and change in earnings



Notes: Results of equation (2) for different values of λ . Children aged 0 to 18 at baseline that are matched to the RSH, and report non-missing schooling. Standard errors clustered by municipality of origin. 10%*, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, birth order, and year of birth fixed effects.

Table A.1: Determinants of the probability of displacement at the slum level

Sample	Probability of displacement					
	(1)	(2)	(3)	(4)	(5)	(6)
Area (in hectares)	-0.016*** (0.005)	0.012 (0.017)	-0.015** (0.006)	0.017 (0.017)	0.037* (0.021)	0.038 (0.026)
Families (/100)	0.024* (0.012)	0.006 (0.020)	0.024* (0.012)	0.004 (0.019)	-0.019 (0.025)	-0.021 (0.028)
Distance to river	-0.073 (0.042)	-0.118** (0.045)	-0.068 (0.040)	-0.100* (0.048)	-0.125** (0.049)	-0.118* (0.061)
Military name	-0.145 (0.106)	-0.047 (0.113)	-0.162 (0.110)	-0.065 (0.108)	-0.044 (0.121)	-0.007 (0.126)
Log(property prices)	0.122 (0.105)	0.307** (0.127)	0.155 (0.161)	0.290* (0.138)	0.273 (0.223)	0.181 (0.266)
Population's schooling			-0.006 (0.034)	0.016 (0.036)	0.011 (0.043)	0.022 (0.051)
Schools/student			0.032 (0.052)	0.027 (0.054)	0.048 (0.068)	0.065 (0.075)
Distance to subway			0.003 (0.006)	0.010** (0.005)	0.021 (0.016)	0.015 (0.016)
<i>R</i> ²	0.108	0.254	0.116	0.274	0.319	0.301
Sample mean	0.64	0.64	0.64	0.64	0.64	0.61
Observations	133	133	133	133	120	111
Municipality of origin FE	✓		✓	✓	✓	✓

Notes: Regressions for the linear probability of displacement on slums' characteristics. Standard errors are clustered by municipality of origin in parenthesis. Because of the small number of observations, we use the definitions of municipalities before 1980, which correspond to 19 unique urban municipalities of origin. 10%, 5%, 1%. “Matched” stands for the slums in the final sample of children, and “urban” stands for slums in urban municipalities.

Table A.2: Summary statistics for children at the time of intervention by gender

	Women 0 to 18		Men 0 to 18	
	Non-displaced mean	Difference (within municip)	Non-displaced mean	Difference (within municip)
Age	8.60	-0.20 (0.30)	8.71	-0.45 (0.30)
Firstborn	0.35	0.01 (0.01)	0.37	0.01 (0.02)
# Siblings	2.74	0.17 (0.12)	2.72	0.09 (0.15)
HH age	35.76	-0.55 (0.39)	35.83	-0.61 (0.52)
Mother's age at birth	25.02	-0.31** (0.14)	25.03	-0.14 (0.22)
Female HH	0.31	-0.01 (0.03)	0.31	-0.003 (0.03)
Married HH	0.85	-0.07*** (0.02)	0.84	-0.05*** (0.01)
Widowed HH	0.01	0.00 (0.004)	0.01	0.00 (0.003)
HH marital status unknown	0.08	0.01 (0.01)	0.08	0.02 (0.02)
Mapuche HH	0.05	0.02** (0.01)	0.05	0.02* (0.01)
Mother's schooling	6.24	-0.35 (0.26)	6.27	-0.30 (0.22)
Individuals	16,565		16,433	

Notes: Within difference corresponds to the coefficient of *displaced* in equation (1) conditional on municipality of origin and year of intervention fixed effects. All children in matched sample from age 0 to 18 at baseline. Standard errors clustered by municipality of origin in parenthesis. 10%*, 5%**, 1%***.

Table A.3: Variance decomposition of outcomes within municipalities

Outcome (Source)	Household Income/pc (1978 Empl. Survey) (1)	Schooling (Census 1982) (2)	Household Income/pc (CASEN 1990) (3)	Schooling (CASEN 1990) (4)
Mean	13,281.9	6.97	229,720.8	8.37
Std. error	3,104.9	0.30	28,717.0	0.35
% Var. due to municip.	28.92	23.5	21.03	22.3
# of municip.	8	51	42	42

Notes: "% Var. due to municip." stands for the percentage of the variance of each outcome due to variation within municipalities. All outcomes measured for head of households in Greater Santiago. Data sources are 1978 Employment Survey conducted quarterly by University of Chile, Census of Population 1982, and CASEN 1990, which is the Socioeconomic Characterization Survey of 1990. Census data includes all municipalities. Employment Survey groups municipalities geographically in 8 strata. CASEN includes the 42 municipalities of Greater Santiago. Income measured in Chilean pesos in 2018.

Table A.4: Displacement effect on types of occupations/industries

Outcome	Occupation				Industry		
	Employer	Independent worker	Employee	Caregiver	Manufacture	Construction	Services
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Displaced	-0.004*** (0.001)	0.026*** (0.009)	-0.029* (0.014)	0.003 (0.005)	0.008* (0.004)	0.017*** (0.003)	-0.005 (0.005)
Non-displaced mean	0.004	0.193	0.472	0.065	0.040	0.038	0.116
% Var. w.r.t. non-disp.	-100	13.5	-6.1	4.6	21.1	44.7	-4.3
R ²	0.003	0.019	0.094	0.058	0.035	0.093	0.077
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓	✓
Observations	533,444	533,444	533,444	533,444	533,444	533,444	533,444

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%*, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, head of household's age, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects. Row labeled as % Var. w.r.t. non-disp. stands for "percentage variation with respect to non-displaced mean."

Table A.5: Displacement effect on demographic outcomes

Outcome	Ever married	Age at first marriage	Teen parent	#Children	On welfare (2015-2019)	\$Welfare (2015-2019)	Incarcerated (2000-2010)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	-0.007 (0.012)	-0.078 (0.251)	0.058 (0.015)***	0.116 (0.031)***	0.029 (0.013)**	13.338 (5.085)**	0.006 (0.003)**
Non-displaced mean	0.66	24.67	0.34	2.42	0.27	58.29	0.021
% Var. w.r.t. non-disp.	0.9	-3.9	16.8	4.1	10.7	22.9	28.57
R ²	0.064	0.049	0.098	0.041	0.136	0.047	0.031
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓	✓
Observations	26,675	26,675	26,675	26,675	267,074	267,074	26,230
Individuals	26,675	26,675	26,675	26,675	25,433	25,433	26,230

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%*, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, head of household's age, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects. Row labeled as % Var. w.r.t. non-disp. stands for "percentage variation with respect to non-displaced mean."

Table A.6: Displacement effect on household characteristics

Outcome	Homeowner	Renter	Transfer	Squatter	Doubled-up	HH size	Parent in
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Displaced	-0.004 (0.019)	-0.019** (0.007)	0.019 (0.018)	-0.001 (0.002)	-0.001 (0.015)	0.029 (0.062)	-0.010 (0.016)
Non-displaced mean	0.51	0.12	0.35	0.01	0.29	3.87	0.20
% Var. w.r.t. non-disp.	-0.8	-15.8	5.4	-10.0	-0.3	0.7	-5.0
R ²	0.064	0.049	0.098	0.041	0.031	0.057	0.060
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓	✓
Observations	533,444	533,444	533,444	533,444	533,444	533,444	533,444
Individuals	26,675	26,675	26,675	26,675	26,675	26,675	26,675

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%*, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, head of household's age, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects. Row labeled as % Var. w.r.t. non-disp. stands for "percentage variation with respect to non-displaced mean." "Transfer" means current house/apartment is not owned but it has been transferred from a third party. "Parent in" means at least one of the parents lives in the house.

Table A.7: Assignment location attributes and displaced families' characteristics at baseline

Location Atributtes	Population's schooling	Unempl. rate	% Rural	# Primary care cent./1,000HH	# Hospitals/ 1,000HH	# schools/ 1,000 stud.	# Pub. schools/ 1,000 stud.	# Priv. schools/ 1000 stud.	Fragment. index	Polarization index	Prices (in logs)	Distance from origin
HH's age	0.000 (0.001)	-0.000 (0.001)	-0.002 (0.002)	-0.003* (0.002)	-0.003* (0.002)	-0.003 (0.004)	-0.002 (0.003)	-0.004 (0.004)	-0.001 (0.001)	-0.000 (0.001)	-0.001 (0.001)	-0.017** (0.008)
Female HH	-0.011 (0.018)	0.020 (0.018)	-0.002 (0.009)	0.019 (0.027)	0.011 (0.020)	0.040 (0.038)	0.041 (0.033)	0.024 (0.046)	-0.000 (0.007)	0.003 (0.006)	0.000 (0.008)	0.033 (0.089)
Married HH	0.008 (0.017)	0.008 (0.020)	-0.019 (0.020)	0.016 (0.021)	-0.031 (0.026)	-0.035 (0.030)	-0.030 (0.028)	-0.040 (0.028)	-0.012 (0.008)	-0.010 (0.007)	0.001 (0.005)	-0.218* (0.125)
Widowed HH	0.054 (0.056)	-0.060 (0.049)	0.053 (0.054)	0.082 (0.068)	0.056 (0.057)	0.041 (0.026)	0.027 (0.027)	0.076** (0.032)	0.008 (0.015)	0.014 (0.017)	0.023 (0.016)	0.282 (0.345)
Marital status unknown	0.018 (0.021)	-0.023 (0.025)	-0.013 (0.021)	0.018 (0.019)	-0.034 (0.023)	-0.029 (0.021)	-0.028 (0.019)	-0.025 (0.021)	-0.007 (0.007)	-0.008 (0.007)	-0.000 (0.006)	-0.186 (0.159)
# children	0.007 (0.006)	-0.006 (0.008)	-0.004 (0.005)	-0.015* (0.008)	-0.009* (0.005)	-0.005 (0.009)	-0.005 (0.008)	-0.005 (0.010)	-0.001 (0.003)	0.001 (0.002)	-0.002 (0.002)	-0.016 (0.033)
Mapuche HH	0.009 (0.018)	0.006 (0.016)	-0.028* (0.016)	-0.022 (0.020)	-0.056* (0.029)	-0.070 (0.060)	-0.060 (0.053)	-0.080 (0.063)	0.002 (0.011)	0.001 (0.007)	-0.013* (0.006)	-0.110 (0.151)
R ²	0.600	0.588	0.775	0.571	0.666	0.397	0.500	0.325	0.523	0.559	0.721	0.766
Observations	11,327	11,327	11,327	11,327	11,327	11,327	11,327	11,327	11,327	11,327	11,327	11,327
<i>Test of joint significance of baseline controls</i>												
F	0.550	0.905	0.599	1.977	2.013	2.750	2.236	1.466	2.666	1.546	2.317	1.044
p > F	0.788	0.518	0.751	0.099	0.094	0.029	0.065	0.225	0.033	0.198	0.057	0.427
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Year of intervention FE	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓

Notes: Standard errors clustered by municipality of origin. 10%, 5%, 1%. Attributes in columns 1, 2 and 3 are measured at the census district level in 1982; schools, hospitals and subway are measured in 1985. Neighborhood fragmentation and neighborhood polarization are measured at the project level based on data from Molina (1986) and MINVU (1979) (See text for details).

Table A.8: Location attributes at origin by mixed and not mixed neighborhoods

Location Attributes by Census District	Non-displaced mean (1)	Displaced mixed mean at origin (2)	Displaced not-mixed mean at origin (3)	Difference (2)-(1) (within munic.) (4)	Difference (3)-(1) (within munic.) (5)
Schooling HH	7.24	7.54	7.27	0.75 (0.79)	0.23 (0.79)
Unemployed HH	0.18	0.18	0.21	-0.01 (0.02)	0.01 (0.03)
HS dropout students	0.33	0.32	0.32	-0.03 (0.03)	-0.03 (0.03)
Schools per census district	3.89	3.57	3.93	-0.13 (0.90)	0.63 (0.91)
Schools per 1,000 students	1.19	0.84	0.92	-0.54 (0.86)	0.12 (1.74)
Pub. schools per 1,000 students	1.00	0.68	0.86	-0.53 (0.93)	0.17 (1.61)
Priv. schools per 1,000 students	0.18	0.14	0.04	-0.03 (0.12)	-0.05 (0.18)
Family care centers per 1,000 HH	0.01	0.01	0.01	0.00 (0.01)	0.01 (0.02)
Hospitals per 1,000 HH	0.03	0.02	0.02	0.00 (0.02)	-0.03 (0.03)
Distance to (closest) metro station in km	7.95	9.89	8.25	-0.64 (0.38)	1.32 (1.18)
Commuting time to work (min) ^a	42.25	42.14	43.65	-0.11 (0.84)	1.40 (0.83)
Commuting time to study (min) ^a	32.92	33.14	31.87	0.22 (0.61)	-1.05 (0.87)
Observations	53	90	17	143	70
# Slums	47	66	17	113	62
# New projects	47	34	9	77	54

Notes: Each observation is a slum-neighborhood pair. Within difference corresponds to a regression of each location attribute on a displacement dummy conditional on municipality of origin. Standard errors clustered by municipality of origin. 10%, 5%, 1%. All location attributes correspond to population averages by census districts in 1982. ^aMeasured as the weighted average in minutes that takes the average person in each municipality to go to work/study using public transportation; because these two variables are measured at the municipality level, the difference in column (3) does not include municipality fixed effects.

Table A.9: Displacement effect on schooling outcomes by age at intervention

Age group	0-5 (1)	6-10 (2)	11-14 (3)	15-18 (4)
<i>Panel A. Outcome: Years of schooling</i>				
Displaced	-0.741 (0.155)***	-0.644 (0.113)***	-0.488 (0.176)***	-0.822 (0.396)**
Non-displaced mean	11.89	11.57	11.11	10.32
% Var. w.r.t. non-disp.	-6.2	-5.6	-4.4	-8.0
R ²	0.090	0.086	0.105	0.089
<i>Panel B. Outcome: High school graduate</i>				
Displaced	-0.122 (0.018)***	-0.122 (0.021)***	-0.076 (0.027)***	-0.130** (0.053)***
Non-displaced mean	0.75	0.68	0.61	0.50
% Var. w.r.t. non-disp.	-16.3	-17.9	-12.5	-0.26
R ²	0.072	0.069	0.086	0.067
<i>Panel C. Outcome: College attendance</i>				
Displaced	-0.089 (0.022)***	-0.039 (0.014)***	-0.025 (0.016)	-0.027 (0.019)
Non-displaced mean	0.23	0.17	0.14	0.10
% Var. w.r.t. non-disp.	-38.7	-22.9	-17.9	-27.0
R ²	0.041	0.031	0.034	0.035
Municipality of origin FE	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓
Observations (Individuals)	8,665	9,271	5,422	3,317

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects. Row labeled as % Var. w.r.t. non-disp. stands for "percentage variation with respect to non-displaced mean."

Table A.10: Displacement effect on labor market outcomes by gender

Outcome	Labor income (1)	Employed (2)	Has a contract (3)	Temp. worker (4)	Taxable income (5)	Formal income (6)	Informal income (7)
<i>Panel A. Women</i>							
Displaced	-15.479* (7.845)	-0.010 (0.021)	-0.050** (0.020)	0.037 (0.022)	-38.213** (16.997)	-16.813** (7.264)	1.334 (2.040)
Non-displaced mean	109.69	0.55	0.32	0.64	523.09	77.65	32.04
% Variation w.r.t. non-disp.	-14.1	-1.8	-15.6	5.8	-7.3	-21.7	4.2
Observations	312,828	312,828	312,828	312,828	46,930	312,828	312,828
Individuals	14,480	14,480	14,480	14,480	8,626	14,480	14,480
<i>Panel B. Men</i>							
Displaced	-13.240** (6.104)	0.026* (0.013)	-0.025 (0.015)	0.035* (0.018)	-34.423* (19.121)	-13.261** (4.830)	0.021 (3.782)
Non-displaced mean	220.77	0.84	0.53	0.44	631.28	154.23	67.54
% Variation w.r.t. non-disp.	-6.0	3.1	-4.7	8.0	-5.5	-8.6	0.0
Observations	220,616	220,616	220,616	220,616	52,617	220,616	220,616
Individuals	12,195	12,195	12,195	12,195	9,264	12,195	12,195
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓	✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. 10%, 5%, 1%. All regressions control for year of intervention fixed effects and semester fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, head of household's age, age of mother at birth, number of siblings, birth order, and year of birth fixed effects.

Table A.11: Displacement effect on schooling outcomes by gender

Outcome	Years of schooling (1)	1[HS graduate] (2)	1[2y college] (3)	1[5y college] (4)
<i>Panel A. Women</i>				
Displaced	-0.679*** (0.172)	-0.118*** (0.024)	-0.025* (0.013)	-0.022** (0.009)
Non-displaced mean	11.43	0.67	0.12	0.05
% Variation w.r.t. non-disp.	-5.9	-17.6	-20.8	-44.0
R ²	0.121	0.095	0.021	0.029
Individuals	14,480	14,480	14,480	14,480
<i>Panel B. Men</i>				
Displaced	-0.682*** (0.159)	-0.111*** (0.022)	-0.039** (0.012)	-0.024*** (0.007)
Non-displaced mean	11.32	0.65	0.12	0.06
% Variation w.r.t. non-disp.	-6.0	-17.1	-32.5	-40.0
R ²	0.118	0.094	0.029	0.026
Individuals	12,195	12,195	12,195	12,195
Municipality of origin FE	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. 10%, 5%, 1%. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, head of household's age, age of mother at birth, number of siblings, birth order, and year of birth fixed effects.

Table A.12: Displacement effect and change in location attributes on main outcomes

Outcome	Employment (1)	Contract (2)	Temp. worker (3)	Years of schooling (4)
Displaced	0.022 (0.012)	-0.009 (0.015)	-0.020 (0.013)	-0.306* (0.165)
* ΔHH years of schooling	0.006** (0.002)	0.007* (0.003)	-0.011*** (0.004)	0.036 (0.052)
* Fragmentation	-0.005 (0.014)	-0.016 (0.016)	0.049*** (0.018)	-0.500* (0.261)
* Distance from origin	0.000 (0.001)	-0.002 (0.001)	0.002** (0.001)	-0.004 (0.007)
* Δ property prices	0.026*** (0.007)	0.005 (0.011)	0.013 (0.010)	0.067 (0.206)
* Δ# schools/child	0.000 (0.003)	-0.003 (0.005)	0.008 (0.005)	0.016 (0.062)
<i>R</i> ²	0.101	0.066	0.076	0.116
Non-displaced mean	0.67	0.41	0.56	11.37
Municipality of origin FE	✓	✓	✓	✓
Controls	✓	✓	✓	✓
Observations	533,444	533,444	533,444	26,675

Notes: This table shows results for $Y_{it} = \alpha + \beta Displaced_{s\{i\}} + \gamma \Delta Attribute_{do} + \psi_o + \psi_\tau + X_i'\theta + \varepsilon_{it}$. All changes in attributes are measured at the census district level which corresponds to a smaller level of aggregation than municipalities. Regressions for children aged 0 to 18 that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%, 5%, 1%***.

Table A.13: Displacement and social capital in the long run

Outcome	Chose neighborhood (1)	Conflictive neighbors (2)	Insecure neighborhood (3)	No trust in neighbors (4)	Trust own child with neighbor (5)	Divided neighborhood (6)
Displaced	-0.037 (0.047)	0.101 (0.277)	-0.099 (0.637)	0.046 (0.032)	-0.100** (0.050)	0.209*** (0.048)
<i>R</i> ²	0.160	0.062	0.071	0.111	0.161	0.224
Fragmentation Index	-0.027 (0.060)	0.877** (0.420)	-0.442 (0.975)	0.083 (0.051)	-0.070 (0.067)	0.284*** (0.069)
<i>R</i> ²	0.142	0.068	0.065	0.108	0.154	0.230
Non-displaced mean	0.780	1.586	1.652	0.116	0.341	0.296
Observations	1,184	1,184	1,184	1,184	1,184	1,184
# neighborhoods	43	43	43	43	43	43

Notes: Results of equation (1) on individuals' perceptions about their neighborhoods in 2012. Data come from Núñez et al. (2012). Each individual in this dataset is matched with a neighborhood in our sample, using current address. Standard errors clustered by municipality of residence in parenthesis. 10%, 5%, 1%***

Table A.14: Displacement effect on children's and parents' locations after 2015

Sample	Parents in RSH (2015-2019)				Children in RSH (2015-2019)			
	Same municipality (1)	Same neighborhood (2)	Distance from assigned neigbh. (3)	Municipality of origin (4)	Same municipality (5)	Same neighborhood (6)	Distance from assigned neigbh. (7)	Municipality of origin (8)
Probability of living in								
Displaced	-0.111 (0.101)	-0.184 (0.130)	1.329 (1.396)	-0.286*** (0.089)	-0.080 (0.088)	-0.121 (0.095)	1.560 (1.403)	-0.215 (0.071)
Non-displaced mean	0.669	0.530	3.156	0.669	0.454	0.309	6.103	0.454
% Var. w.r.t. non-disp.	-16.6	-34.7	42.1	-42.8	17.6	-39.2	25.6	-47.3
<i>R</i> ²	0.196	0.228	0.145	0.452	0.001	0.008	0.002	0.304
Municipality of origin FE	✓	✓	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓	✓	✓
Observations	37,516	37,516	33,256	37,516	90,093	90,093	75,979	90,093

Notes: Regressions for children aged 0 to 18 at baseline, and their parents that are matched to the RSH, and report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis. 10%, 5%, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects. Row labeled as % Var. w.r.t. non-disp. stands for "percentage variation with respect to non-displaced mean."

Table A.15: Displacement effect and subway rollout between 2007 and 2019

Outcome	Labor Earnings					
	0.7 km (1)	0.8 km (2)	1 km (3)	1.2 km (4)	1.5 km (5)	2 km (6)
Distance to new station						
Displaced	-16.311* (8.020)	-16.853** (8.020)	-18.019** (7.620)	-21.338** (8.986)	-21.442** (9.148)	-18.316** (7.854)
Subway station	-5.974 (6.680)	-4.886 (6.877)	-7.321 (5.237)	-10.409* (6.031)	-10.662 (6.297)	-2.378 (4.236)
Displaced*Subway	2.659 (7.143)	11.158* (6.447)	13.657** (5.653)	16.003** (6.486)	14.532** (6.484)	6.071 (4.746)
Non-displaced mean	155.89	155.89	155.89	155.89	155.89	155.89
% Displaced affected by subway	2.2	11.9	13.7	26.6	36.98	53.13
% Non-displaced affected by subway	28.36	28.36	31.58	44.01	48.86	53.59
%Δ Displacement effect	16.3	66.2	75.8	75.0	67.8	33.1
R ²	0.126	0.126	0.126	0.126	0.126	0.126
Municipality of origin FE	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. 10%, 5%, 1%. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, birth order, and year of birth fixed effects.

Table A.16: Comparison of earnings estimates across studies

Study	Setting	% Δ Earnings		% Δ Neighborhood Quality	Elasticity
		(1)	(2)		
Chetty et al. (2016) ^a	MTO (children 7-13 in exp. group)		+14%	-34% (Poverty)	0.41
Chyn (2018) ^b	Public demolition in Chicago (children 7–18)		+16%	-22.2% (Poverty)	0.72
Barnhardt et al. (2016) ^c	Housing lottery Ahmedabad (adults in India)		-14.5%	-37.5% (Urbanicity)— -8.1% (Housing Value)	0.38–1.8
This paper ^d	Program for Urban Marginality (children 0–18 in Chile)		-9.4%	-9.5% (Schooling)	0.99

Notes: Results come from tables in each corresponding paper: ^aTables 2 and 3; ^bTables 2 and 3; ^cTables 5 and 6; ^dTables 3 and 4.

Table A.17: Comparison of schooling estimates across studies

Study	Setting	% Δ Years of Education	% Δ Neighborhood Quality	Elasticity
	(1)	(2)	(3)	(4)
Chetty et al. (2016) ^a	MTO (children 7-12 in Exp. group)	+15% (College Att.)	-34% (Poverty)	0.44
Chyn (2018) ^b	Public demolition in Chicago (children 7-18)	-8.1% (HS dropout) 28% (College Att.)	-22.2% (Poverty)	0.36 1.26
Barnhardt et al. (2016) ^c	Housing lottery Ahmedabad (children in India)	-2.25% (schooling)	-37.5% (Urbanicity)— -8.1% (Housing Value)	0.06–0.27
This paper ^d	Program for Urban Marginality (children 0-18 in Chile)	-6.0% (schooling) -17.6% (HS grad) -32.8% (College att.)	-9.5% (Schooling) -9.5% (Schooling) -9.5% (Schooling)	0.63 1.85 3.45

Notes: Results come from tables in each corresponding paper: ^aTables 2 and 3; ^bTables 2 and 3; ^cTables 5 and 6; ^dTables 3 and 6.

B EVICTION POLICIES

Table B.1: Characteristics of each version of the program

Intervention	Location	Property right	Type of dwelling	Public services	Cost for family
Non-displaced (1/3) (urban renewal)	Same	Yes	Starting kit (*) or apartment	Yes	25% paid in 15 years
Displaced (2/3) (evicted)	New (periphery)	Yes	Apartment or house	Yes	25% paid in 15 years

(*) A starting kit includes a living room, a bathroom, and a kitchen.

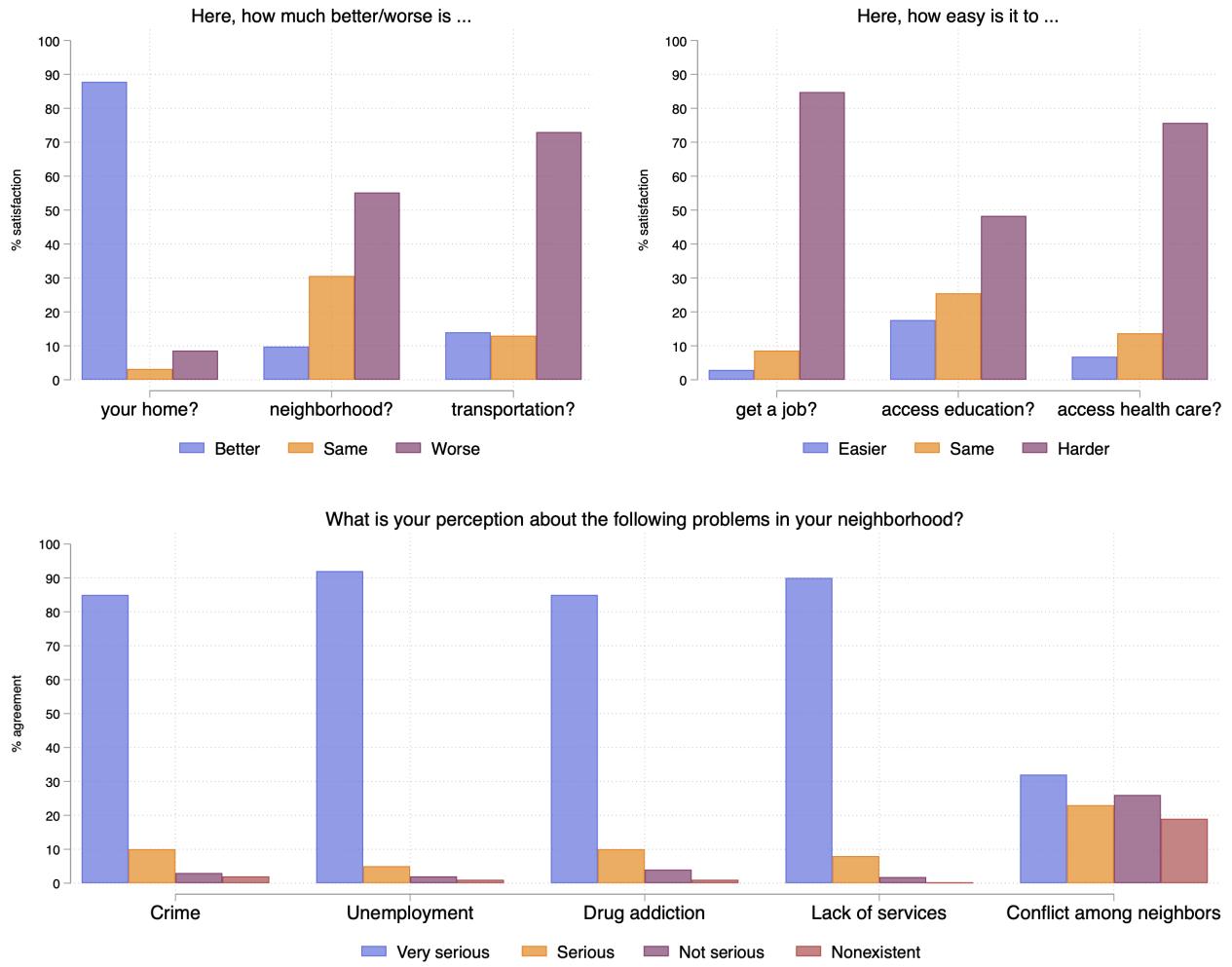
Figure B.1: Example of a slum and new neighborhoods



Notes: Examples of neighborhoods from [Hidalgo \(2019\)](#).

B.1 Evaluation of evictions program in 1987

Figure B.2: Summary of evaluation of the Program for Urban Marginality (Aldunate et al., 1987)



Notes: Summary of results found by Aldunate et al. (1987). The authors interviewed 592 displaced slum dwellers that were relocated into four new neighborhoods.

B.2 Testimonies of slum dwellers after displacement

The following are extracts from testimonies in Álvarez (1988) and Álvarez and Cavieres (2016). The first were recorded in 1988, approximately ten years after displacement, from individuals between 15 and 28 years of age who were displaced from the same slum of origin to a new neighborhood located in the municipality of San Bernardo, as part of a *Operación Confraternidad*. The second set of testimonies was recorded in 2013 from families displaced to neighborhood *El Castillo*.

B.2.1 Álvarez (1988)

Ana

My name is Ana, I was fourteen when I arrived to *Zanjón* with my parents and seven siblings.⁶⁴ We were there for almost four years.

I could say that our life in the slum was good because there were jobs. You could just go out to search and you would always find something. I used to clean houses around the neighborhood. Generally I had two or more *patronas* who needed me to come once or twice a week to clean, wash and iron their clothes. That is how I would put together a salary for myself. And nobody ever, wherever I went, distrusted me as happens to me now when I answer a job ad, just for living in this neighborhood.

At the time, my father worked at the railway company and my mother was a cook. That was enough for us. We did not live in excess, but we had a decent life ... I got married when I was fifteen. At *Zanjón* we had two kids. He had a job at a furniture factory as a loader.

The ten years since we moved to the new neighborhood have been truly bad for us. One of my kids lost his way here because he started hanging out with a gang that smokes *neoprén* ...⁶⁵ My husband lost his job because he was always late, and I had to stop working as a maid because, how was I supposed to pay for all the buses I would need? Just like that, working was off the table. We only had the *mínimo* available.⁶⁶

We were so far away! At the end, you would pay in transportation all the money you made working. That is why I stopped traveling to Santiago and I also joined the *mínimo* program like my husband.⁶⁷

⁶⁴ *Zanjón de la Aguada* was the general area where the slum of origin was located.

⁶⁵ *Neoprén* was the name of a toluene-based industrial adhesive, commonly used as a drug in Chile in the 1980s.

⁶⁶ She is referring to the *Programa de Empleo Mínimo*. A basic-work program implemented at the time to help support low-income families.

⁶⁷ She is referring to the *comuna de Santiago* (Santiago district), one of the 34 *comunas* that form Greater Santiago.

Twenty year-old male

What I can say is that only the houses are better here than in the slum at *Zanjón*, everything else was better there. It was really calm over there, unlike here where everything happens after dark. We have had many crimes, and it is the same young people from here who commit them. Sometimes a person is asked for a cigarette and, if he does not share one, he gets stabbed and left right there on the floor.

In this neighborhood people turned evil; there at the *Zanjón* we were fine. We were peaceful, but here everything changed. I turned evil here. We are doomed. If we had anywhere to go, we would leave.

Another thing that made us worse off is that two slums from the *Zanjón* were combined here: *Isabel Riquelme* and *Centenario*. That started the fights and the gangs because they distrusted each other and both groups wanted to be better than the other. Lately, things have calmed down a bit because many have been taken to *Puente Alto*.⁶⁸ but that never lasts: When the worst are taken away, it doesn't take long for new ones to replace them.

El Chito

People call me *Chito*. I am twenty-eight years old. I am still young but I don't feel like that.

When I was fourteen I had to start working to help support my mother and siblings since I was the oldest of the five (siblings). At that time, I started working as a milkman for *Soprole*.⁶⁹ My dad also worked there loading trucks. Between the two of us we made enough for the family so there was no need for my mom to work. She stayed home and my siblings attended school.

At the beginning I liked the houses and I agreed with my mom that moving there had been for the better. However, this enthusiasm didn't last when I realized that we would be surrounded only by countryside. It was a wasteland! None of the young people could get used to the idea that they had left us so far from all the places where we liked to go. Here we learn the true meaning of boredom: standing at the corners without enough money to even ride a bus to go anywhere, when in the *Zanjón* we would walk a few steps and we would be anywhere we wanted.

However, the job situation was even worse. We were so far away that it was a huge success just to be on time. I had to wake up at four in the morning to be at the place where the trucks would unload the milk (for the milkmen to distribute). The thing is that the loaders would unload the milk on the street, even if a milkman had not yet arrived and, of course, in that case the milk would be swiftly stolen. Finally, this happened to me on January 1, when buses started running later than usual. When I arrived, everything had been stolen and that was the end of it.

⁶⁸ *Puente Alto* was the location and informal name of the local jail.

⁶⁹ *Soprole* is a prominent milk company in Chile.

My dad could never adapt to the longer commuting time and started missing work a lot, until one day he simply announced that he was not going to keep working anymore. After that, I stopped trying to simultaneously work and study as I was before; now the whole weight of supporting the household fell on me. And they fired me from *Soprole!* There they really screwed me over. I searched like a madman for some opportunity around here but there was nothing. This sector was made only of farmland and we knew nothing about agriculture. The only option was the *mínimo*.⁷⁰ We ended up doing that together with my mom and dad. Everything we made was for the family and even then it wasn't enough.

Manolo

My name is Manuel, I am twenty-seven years old, I am married and I have a three year-old son. We are living with my parents. If we had waited to have our own place first, we would have never gotten married. This is what all couples do here; they stay living with their in-laws.

I met Tere, my wife, here in the *Confra*⁷¹ and we get along well as a married couple even though there are almost no distractions here. You can't just tell your wife to go the park or the movies because you would have to leave at ten in the morning just to come back before it is dark— it is almost four hours just in the bus. Hence, going out with the family is not easy.

Living here in the *Confra*, my mother started having mental health issues. Now she is always anxious and she has a facial tic. She wakes up at five in the morning or otherwise wouldn't make it to her job on time. But the worst is that she has to walk for four blocks to take the bus at a time when it is dangerous to be on the street. Sometimes, she has a shift and then she comes back around midnight when the street is filled with drug addicts and alcoholics.

At *Zanjón* there were criminals too, but there were only ten out of each hundred, unlike here where almost everybody is a criminal. It's just that wherever you look the only thing you see is misery. We are all poor here. Besides us, they have brought other displaced from many slums. *Confra* is adjacent to the *Santa Marta* slum and the *21 de Mayo* and *La Portada* housing projects. This is also different from living at *Zanjón*, because there you could see other realities. Here it is the opposite: we are all sunk in a hole and there is no exit. Wherever you look it is only misery. How could crime not be rampant?

Osvaldo

My name is Osvaldo. I am twenty-four years old now. Now I work in the warehouse section of a hospital. My family is from Southern Chile. We came to Santiago when my mom became a widow. Then, we moved to *Zanjón*. We are five siblings.

⁷⁰ *Programa de Empleo Mínimo*. A basic-work program for low-income families.

⁷¹ *Confra*, short for *Población Confraternidad*

Here in *Población Confraternidad* what I do not like is what surrounds us, the environment we live in. It is not the poverty—you can get used to that—the terrible thing is the criminality, the drug addiction. It is also that it is not only us, as many more poor people from all over Santiago have also arrived here, to accumulate all of us here and that is the bad thing. All the poor piled together in a single place! And who benefits from that? Not us, in any case.

The despair leads young people to use drugs and their mothers don't know how to guide them. Then, they (young people) lose control and drift towards drugs and crime. This happens because when they reach working age, they need to work and can't find anything; they have no opportunities of any kind. Then they end up standing idle in groups on a corner, and that is how vice and crime begin.

I imagine that those who decided to move us thought that we would have a more organized existence here in San Bernardo. However, they did not think that it is not enough to live in the greater material comfort they provided us; people also need access to culture, education, recreation, and, above all else, work! Without work, without having access to economic means, everything goes downwards. The youth and the families are destroyed. Unfortunately, they still don't understand that. They think since we were worse before, there is nothing to complain about.

What happened with the displacement was that many lost their jobs and the people ended up having to sell whatever they could from their homes. Instead of improving them, people started disassembling them.

B.2.2 Álvarez and Cavieres (2016)

Sixty year-old woman

Yes, the dream of homeownership is fulfilled but what bothered us a lot is that it was too far. From Renca to here is a long way, because from Renca to city center it was only twenty minutes. So it was very difficult for us to travel downtown. So I think that was the main problem we had at the beginning here.

And the people did not adapt here either because there were a lot of quarrels with the other residents who came from other municipalities. There were many fights. There would be a party here and, for example, someone from another municipality would come and a colossal fight would break out.

Informant from Municipality of La Pintana

The movement of people to this neighborhood was not really a migration. They threw these people away. I think that feeling remains forever. So they remain anchored to that event and they can't develop roots, even as several generations have since passed.

C DATA COLLECTION PROCESS AND ATTRITION FROM ARCHIVAL DATA

In this section we provide a more detailed description of the archival collection and database construction.

C.1 Archival Data: Homeowners

The main goal of our data collection is to find the families that participated in the program. The Program for Urban Marginality was implemented under the umbrella of Executive Order 2552. Hence, we search, collect, and digitize archival administrative records generated between 1979 and 1985 associated with Executive Order 2552 from the Metropolitan Regional Housing and Urban Planning Service of Santiago,⁷² located in the National Archives of the Administration (ARNAD), and from historical records kept by the Municipality of Santiago.⁷³

The administrative records consist of lists detailing the names of people and their spouses who received a property deed in a destination neighborhood as part of the Program for Urban Marginality. We are able to collect data for 22,689 unique recipients of social housing. They represent around 56% of the total number of recipients according to the numbers in Molina (1986).

We are not able to find all records for two reasons. First, the original lists of beneficiaries of the program were compiled by the individual municipalities. Some of these records were kept by them and not sent to the central administration. At the time, municipalities were required to keep administrative records for only five years, and after that period ended they were allowed to dispose of them. This issue has been confirmed by several municipalities. A second reason is due to attrition at the ARNAD. There are two instances in the early 1990s in which records were lost: First, when the ARNAD was separated from the National Historical Archives of Chile, and then during a flood that destroyed a fraction of ARNAD's holdings.

A sample of the records is shown in Figure C.1. The archival records contain the information of the recipient of the property deed and their spouse: full names, national identification numbers (NID), place of registration (municipality where the person obtained a national ID), and the address of their new housing unit. These records are grouped by year of eviction or urban renewal and project of destination.

Table C.1 summarizes the total number of recipients and records we are able to find. In the program 65% of the families are displaced and 35% are non-displaced, while in our records, we identify 70% as displaced and 30% as non-displaced. The differences between panels A and

⁷²Each region of Chile (equivalent to a state) has an Urban Development and Housing Service, dependent on the Ministry of Housing and Urban Development. These services administer and implement housing policies at the local level.

⁷³The Municipality of Santiago is one of the 32 municipalities in which Greater Santiago is divided. Hence, it only comprises a small fraction of the population and area of the city.

Figure C.1: Archival Records: Lists of property deeds

Project name (neighborhood)											
ANEXO AL RESUELVO N° 1			Nomina de Asignación de Viviendas Sociales								
Población : José Miguel Infante		Comuna de : Renca		Municipality		New address					
Nº	NOMBRE ASIGNATARIO Y CONYUGE	C.I.DENTIDAD	GABIN.	ROL	DIRECCION MUNICIPAL	CUOTAS DE AHORRO APORT. APLIC. G.NOT.	VALOR	MUTUO			
1	Gaete Dour Cardenio S. Rubilar Figueroa María Y.	1.111.111-1 2.222.222-2	Valdivia Valdivia	6	Toconce N° 1145	220	170	50	283,7663	77,20	
2	Puelma Ibarra Raúl F. Aristegui Palma Silvia	3.333.333-3 4.444.444-4	Thno. Stgo.	9	Toconce N° 1112	100	50	50	283,7663	81,84	
3	Navia Fischer Juan Ovando González Lilia	5.555.555-5 6.666.666-6	Ovalle Ovalle	21	Toconce N° 1968	150	0	150	286,8282	86,83	
4	Díaz José Luis del C. Carrasco Gutierrez Ana	7.777.777-7 8.888.888-8	Stgo. Stgo.	23	Toconce N° 1176	150	100	50	283,7663	79,91	
5	Csses Zúñiga Graciela Aedo Ortiz Modesto	9.999.999-9 10.000.111-1	Nuñoa Stgo.	25	Toconce N° 1184	50	0	50	286,8283	86,83	
6	Araneda Escobar Fernando García Moyano Cristina	11.111.111-2 12.123.456-6	Stgo. Stgo.	27	Toconce N° 1192	505	455	50	283,7663	66,20	

Family ID	Name	Relation to hh	ID	District	House	Address	Aport.	Aplic.	G.Not	Value
1	Gaete Dour Cardenio S.	1	1.111.111	Valdivia	6	Toconce 1145	220	170	50	283.77
1	Rubilar Figueroa María Y.	2	2.222.222	Valdivia	6	Toconce 1145	220	170	50	283.77

B show that we are able to find relatively larger slums and neighborhoods of destination, as we have a smaller share of displaced slums relative to total number of slums in the sample, compared to the same ratio measured by number of families.

Panel C presents our final matched sample. We only keep families in which at least one of the partners has a valid national ID. This variable is key for us because we use it to find homeowners' children. Hence, from the 22,689 observations matched to the slum census only 19,852 families remain in the sample. The 2,837 additional records either do not contain a valid national ID, their national IDs were recorded mistakenly, or they have a legacy ID number from before the creation of the Chilean national ID system of identification. In the last case, we are not always able to validate the ID number using current open data from the electoral records or marriage certificates. Consequently, in this group single adults and the non-displaced are overrepresented. This bias is reflected in the summary statistics of our final sample of children and adults. In consequence, as a robustness check we compute the probability of being found in the archival data at the slum level by estimating a logit regression of the probability of being found as a function of a slum's characteristics, year of intervention, and municipalities of origin fixed effects. Later, we compute the fitted values and use the estimates as a control function in our baseline regressions for earnings. See Table D.7 for results.

Finally, the sample we use corresponds to our matched sample in urban municipalities

Table C.1: Archival Data 1976-1985

Treatment	Displaced	Non-displaced	Total
<i>A. The Program 1979-1985 (Molina, 1986)</i>			
Number of families	26,291	14,200	40,491
Share %	65%	35%	100 %
Number of slums	211	67	278
Number of projects	63	67	130
<i>B. Archival Data 1979-1985</i>			
Number of families	15,866	6,823	22,689
Share %	70%	30%	100%
Number of slums	84	47	130
Number of projects	56	47	96
<i>C. Estimation Sample 1979-1985</i>			
Number of families	13,519	5,468	18,987
Share %	71.2%	28.8%	100%
Number of slums	78	47	124
Number of projects	43	47	84
<i>D. Estimation Sample, urban municipalities</i>			
Number of families	11,327	5,221	16,548
Share %	68.4%	31.6%	100%
Number of slums	62	41	102
Number of projects	32	41	68

Source: [Molina \(1986\)](#) and archival data found by authors.

(panel D) because our measures of neighborhood characteristics are mostly available in urban municipalities of Greater Santiago.

C.2 Locating slums and destination neighborhoods (housing projects)

Archival records are sorted by destination neighborhoods and not by slums. Thus, a key part in the cleaning process of the archival data is to assign each family to a slum of origin. To do this, we use information from three main sources. The housing programs of the Chilean dictatorship in the 1980s were contemporaneously studied by the Latin American Faculty of Social Sciences (FLACSO) in Santiago. We draw intensively from two of their studies: [Benavides et al. \(1982\)](#) compiles a comprehensive list of existing slums in the year 1982, including characteristics such as land size, number of families and location; [Morales and Rojas \(1986\)](#) describe the treatment of each slum, identify neighborhoods of destination, and provide a list of non-displaced slums.

We complement the information from FLACSO with [Molina \(1986\)](#) that studies the experience of displaced families under the program, compiling a list of slums and their locations, the number of families evicted in each episode of displacement, and their neighborhood of destination from administrative records.

Finally, we digitize two slum censuses conducted by Chile's Ministry of Housing and Urban Development (MINVU) in 1979 and 1984. These censuses include the neighborhoods of destination for each slum, which allows us to classify them as displaced or non-displaced.

Combining these three sources we create the treatment variable for each slum. It is important to note that identifying slums is challenging given their dynamic and informal nature. Slums names often changed for a myriad of reasons. For instance, after the military coup of 1973, several slums with left-wing-related names changed their names.

There are two main challenges in the process of allocating families to slums of origin. First, archival records are organized by the date on which the families moved to destination neighborhoods and not by origin and, in most cases, they include groups of families with more than one slum of origin. Second, when a non-displaced slum was treated, very often the new neighborhood had a different name from the original slum name. Moreover, adjacent non-displaced slums were sometimes treated in a single new neighborhood of destination.

To solve the first challenge, we use the number of displaced families treated at every slum of origin, which we obtain from the sources above, jointly with the place of registration variable included on the archival records (see figure C.1). Place of registration is a good proxy for municipality of origin as Greater Santiago was divided in 17 municipalities at the start of the program.⁷⁴ To identify non-displaced slums of origin, we matched the address of the destination neighborhoods to the location of known non-displaced slums and number of families treated. Since families were treated by slum, we know all families in the records classified as non-displaced were part of the same slum of origin. In addition, some of the records for the non-displaced included the type of housing unit received. Thus, when families received a starting-kit (*casetas sanitarias*) we know for certain those families went through an urban renewal process, since that type of housing unit was never given to displaced families.

C.3 Matching process: Sample of children

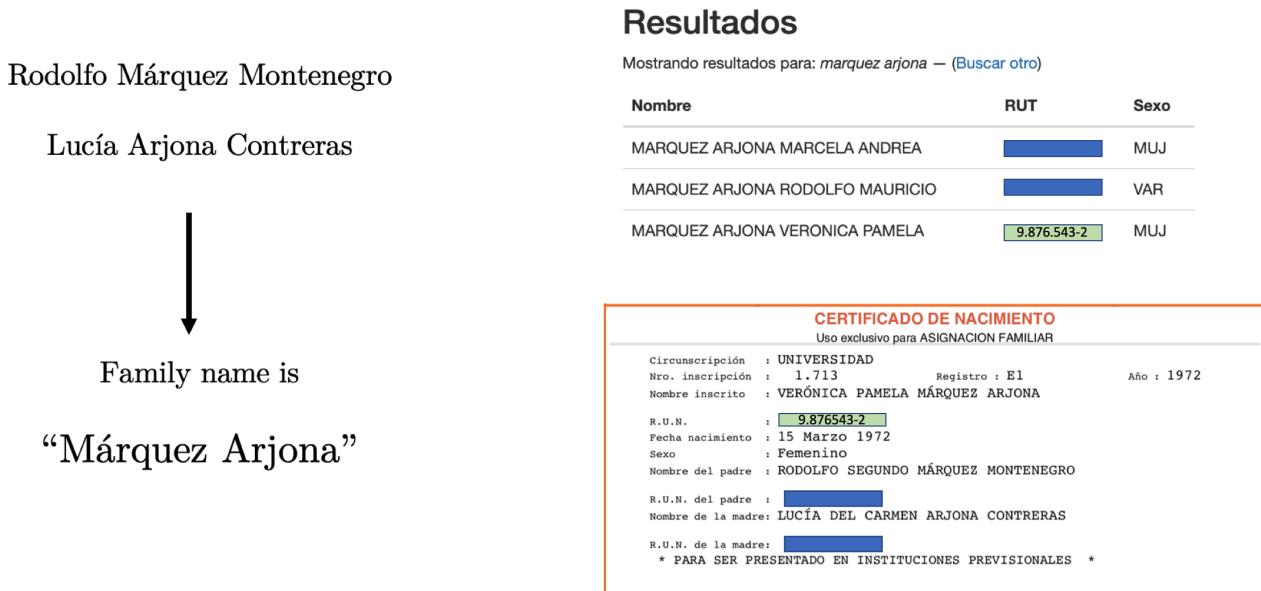
The next step in the construction of the full database consists of finding the children of each family. Our objective is to match spouses to all of their children. Unfortunately, we do not have access to administrative data on family composition at the time of the intervention, so we are forced to reconstruct family structures ourselves. We work with Genealog Chile to web-scrape birth and marriage certificates from Chile's Civil Registry and Identification Service and collect birth certificates for all Chileans who were 18 or older in 2016. The birth certificates contain full name at birth, date of birth, and national ID number, as well as parents' full names, and, in most cases, their national ID numbers (probability of finding parents' national ID numbers decreases by age).

⁷⁴Place of registration is called *Gabinete* in the records and corresponds to the Civil Registry and Identification Service (CRIS) branch where a person was first registered. Most municipalities in Santiago had a CRIS office at the time. After 1980 there was a new geopolitical division of municipalities, and they became 32.

We matched our homeowners' archival data to their children using their national ID numbers. An illustration of the process of finding the children is in Figure C.2. Our initial data consists only of couples. We start by building a data set consisting of all the “family names” for the families in our sample. In Chile, family names are composed of two last names: the first last name (in order from left to right) is the first last name of the father, and the second last name is the first last name of the mother. Hence, both paternal last names are transmitted from parents to children. As an example, consider that María Pérez Rojas (mother) has a child with Juan Rodríguez González (father). In this case, their children will have as family name “Rodríguez Pérez.”

We match family names in the sample to all birth certificates that had the same family name in Chile in 2016. Since birth certificates include the name and national ID numbers of parents, we can corroborate if the parents in a matched birth certificate belong to our sample or not. To find the children of single parents, Genealog was able to obtain all birth certificates for the Chilean population in a second stage of the process. We matched by full name in cases in which the birth certificate did not contain at least one of the parent's national ID numbers. The match in this case was almost exact because Chilean names are composed by a first name, a middle name, and two last names. Thus, the likelihood of an incorrect match is very small.

Figure C.2: Matching process: from parents to children



C.4 Attrition due to missing data from National Archives

There are two potential selection concerns for the data set we assemble. First, we find 56% of the total recipients of the Program of Urban Marginality, which drops to 49% once the data are

cleaned. Second, 81% of the initial sample is later matched successfully with the administrative data we use to evaluate the long-term effects of the program. The 26,676 (80% of 32,998) children in the matched sample would correspond to approximately 40% of the total number of children that would have been part of the program.

In the main body of the text, we discuss whether children in the administrative records are observationally different from children in the archival data. We show that in the administrative data we find more women and younger children. This is the case for both displaced and non-displaced children, meaning that these differences between samples should not pose a problem for the estimation.

Even though children in archives and in administrative data looked similar on their observable characteristics, we have not yet discussed whether the final sample we use in the estimation is a good representation of the individuals in the full program. This is especially relevant if the attrition produced by missing data from the archives generates non-random selection in the origin, and/or in the destination neighborhoods. This is not possible to solve at the demographic level but it is feasible to solve at the slum level because we know which slums are not in our final sample. By looking at the differences between slums in the sample and slums in the full program we can understand whether those differences correlate with the treatment variable.

We use the data compiled by [Morales and Rojas \(1986\)](#) and geolocate all slums in their sample; one caveat is this data set only includes urban areas. Then we match their data with the slums in our sample. The results of this exercise are summarized in Table C.2. In columns (1) and (2) we report the characteristics of both the slums we find in the archives and of those that were missing. In column (3) we report the difference between the two groups conditional on municipality of origin. Notice that we find a greater quantity of displaced slums and larger slums (measured by number of families). This is consistent with the numbers previously shown in Table C.1. On average, slums in our sample were located in census districts with a population that was more educated, had fewer schools per student at their municipalities of origin, and were closer to subway stations (which is positively correlated to being closer to downtown). None of these differences are statistically different from 0 within municipalities (panel A).

We do not find systematic differences in the characteristics of slums at the destination (panel B). On average, the slums we find had families that ended up in municipalities with fewer schools per student, but they were closer to transportation and downtown. This is consistent with finding destination neighborhoods that are less likely to be on the periphery of the city. The destination neighborhoods we find were also less likely to be fragmented compared to neighborhoods we did not find in the archival data. Finally, they also tend to be larger in terms of the number of families (housing units).

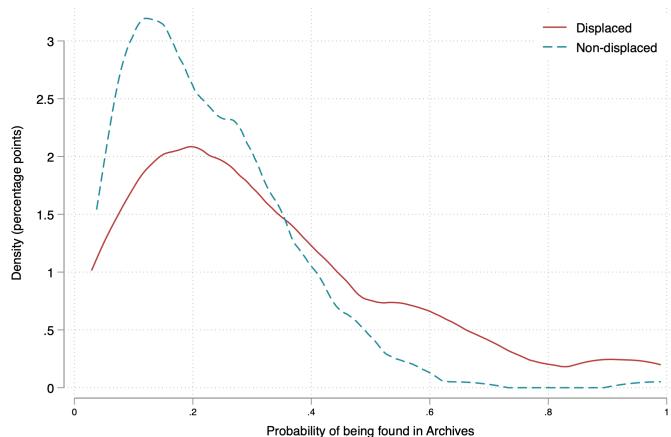
Next we explore if the differences we find vary depending on the treatment received by a slum. In the following columns (4 to 8), we repeat the previous exercise separating the sample now by displaced and non-displaced slums. The numbers in panels B and C show that

the differences between “found” and “not found” are bigger in the displaced group than in the non-displaced.

The relevant question is whether the differences in the probability of finding a slum in the archival data are the reason we find a negative displacement effect instead of a causal effect of the displacement on children’s earnings and education. The signs of the differences make us believe that the selection of slums in our sample would go against finding a negative displacement effect. In column (10) we report the double difference between finding a slum and being displaced: In our sample, displaced slums are more likely to end up in neighborhoods with a less educated population and fewer schools, which would explain a negative displacement effect, but at the same time they ended up closer to transportation and in less fragmented projects relative to the slums that are not found. Thus, if distance to subway and fragmentation correlate negatively with children’s earnings and education, the estimate of a displacement effect we found in our matched sample is an upper bound of the true displacement effect.

The question is which of the previous differences dominate, if any. In Table C.3 we report the same analysis but in regression form in order to control for all the slums’ characteristics at the same time. The results support the validity of our findings, as the probability of finding a slum in the archives is only negatively influenced by features that should make the displacement effect smaller.

Figure C.3: Probability of finding a slum in archival data



Notes: Fitted values correspond to column (5) in Table C.3

Table C.2: Characteristics of slums

Variable	Full sample			Displaced			Non-displaced			Diff-Diff (Displ.*Found) (10)
	in arch. (1)	not in arch. (2)	Diff. (3)	in arch. (4)	not in arch. (5)	Diff. (6)	in arch. (7)	not in arch. (8)	Diff. (9)	
Displaced	0.57	0.42	0.14** (0.06)							
# Families	257.11	168.74	77.91*** (25.91)	243.12	133.60	99.17*** (29.33)	269.10	194.54	85.02 (63.75)	47.43 (87.56)
Land use (hectares)	4.2	3.82	0.59 (0.46)	3.43	3.34	0.42 (0.64)	5.50	4.11	1.44 (1.06)	-1.03 (1.16)
<i>A. Location characteristics at origin</i>										
Population's schooling	7.57	7.37	0.08 (0.31)	7.85	7.90	-0.12 (0.43)	7.25	7.00	0.00 (0.37)	-0.36 (0.36)
Rural (%)	0.03	0.02	0.01 (0.02)	0.04	0.02	0.02 (0.02)	0.01	0.02	-0.01 (0.01)	0.02 (0.02)
# schools per municipality	0.64	0.71	-0.02 (0.05)	0.68	0.75	-0.04 (0.05)	0.57	0.68	-0.04 (0.06)	0.01 (0.05)
# health care centers	0.01	0.01	0.00 (0.00)	0.01	0.01	0.00 (0.00)	0.01	0.01	0.00 (0.00)	0.00 (0.00)
Distance to subway	4.82	6.10	-0.42 (0.49)	4.88	6.08	-0.56 (0.66)	4.78	6.11	-0.62 (0.63)	-0.29 (0.83)
Distance to downtown	9.75	10.77	-0.30 (0.58)	9.48	10.71	-0.59 (0.69)	10.25	10.78	-0.13 (0.67)	-0.72 (0.74)
Property prices (logs)	14.65	14.69	-0.05 (0.05)	14.67	14.77	-0.05 (0.07)	14.64	14.64	-0.08 (0.12)	0.01 (0.12)
<i>B. Location characteristics at destination</i>										
Schooling HH	6.82	6.83	-0.02 (0.15)	6.54	6.59	-0.08 (0.20)	7.25	7.00	0.00 (0.37)	-0.20 (0.41)
Rural	0.03	0.03	0.01 (0.02)	0.04	0.03	0.03 (0.03)	0.01	0.02	-0.01 (0.01)	0.03 (0.03)
# schools per municipality	0.54	0.66	-0.08* (0.04)	0.51	0.63	-0.12* (0.06)	0.57	0.68	-0.04 (0.06)	-0.07 (0.06)
# health care centers	0.01	0.01	0.00 (0.00)	0.01	0.01	0.00 (0.00)	0.01	0.01	0.00 (0.00)	0.00 (0.00)
Distance to subway	5.37	6.75	-1.11*** (0.37)	5.91	7.62	-1.58** (0.51)	4.78	6.11	-0.62 (0.63)	-0.36 (0.85)
Distance to downtown	11.29	12.08	-0.65 (0.62)	12.25	13.84	-1.22 (0.79)	10.25	10.78	-0.13 (0.67)	-0.43 (1.02)
Property prices (logs)	14.64	14.68	-0.08 (0.07)	14.68	14.78	-0.10 (0.10)	14.64	14.64	-0.08 (0.12)	-0.05 (0.13)
<i>C. Project characteristics</i>										
# slums	3.86	3.66	-0.01 (0.37)	6.13	7.38	-1.65** (0.65)	1	1		-1.98 (0.72)
Fragmentation	6636.41	7079.23	-249.12 (423.98)	3992.28	2988.21	1254.44** (420.27)	10000	10000		1469 (467.15)
# families per project	549.69	442.45	-77.78 (51.30)	778	789.35	-56.37 (0.64)	269.10	194.54	85.02 (63.75)	-105.18 (124.33)
Observations	99	222	321	65	108	173	34	114	148	321
Slums						124			148	251
Projects (neighborhoods)						48			148	195

Notes: Each observation is a slum-project pair. Some families from the same slum were sent to more than one project. Summary statistics for all slums reported in Morales and Rojas (1986) that we were able to geolocate. Differences in (3), (7), and (9) correspond to a regression of each attribute on a dummy that indicates if a slum was found in the archival records, conditional on municipality of origin. Column (10) corresponds to the double difference between "found" and "displaced", conditional on municipality of origin. Standard errors clustered at the level of municipality of origin. 10%*, 5%**, 1%***. Morales and Rojas (1986) only include slums in the urban municipalities, and in their data the share of displaced families is smaller than in Molina (1986) because they report more non-displaced slums that we did not find in other sources.

C.5 Variables Definitions

Variable Name	Description
Outcomes and Treatment	
Labor income	Source: RSH. Self-reported labor earnings measured in CLP\$ per month. Original variable corresponds to the sum of all earnings in the last year at the time of the interview. It includes earnings from formal and informal employment, and excludes pensions and transfers. Data available biannually from 2007 to 2019.

Employed	Source: RSH. Person reports to be employed at the time of the interview. It includes any type of employment, formal or informal. Data available biannually from 2007 to 2019.
Taxable income	Source: GRIS Mutuales. Monthly administrative records on taxable earnings for all workers that contribute to Social Security. Data available monthly from 2016 to 2019.
Contract	Source: RSH. Conditional on employment, person reports to work with a formal contract.
Temporary worker	Source: RSH. Conditional on employment, person reports to work on a fixed term.
Years of education	Source: RSH. Completed years of schooling. Constructed based on grade completion and levels. A person can appear multiple times in the RSH with differences in this variable across years. We use the minimum value after the age of 25.
High school graduate	Source: RSH. Person reports to have successfully completed high-school.
College attendance	Source: RSH. Person reports to attend at least one year of tertiary education. This includes 2-3 year colleges or 5-year colleges.
Displaced	Source: Archives and authors calculations. Based on Archival data, MINVU (1979, 1984), Molina (1986), and Morales and Rojas (1986), we construct the displacement dummy at the slum level.

Covariates

Year of intervention	Source: Archives and authors' calculations. Based on Archival data, MINVU (1979, 1984), Molina (1986), and Morales and Rojas (1986).
Municipality of origin	Source: Archives and authors' calculations. Based on Archival data, MINVU (1979, 1984), Molina (1986), and Morales and Rojas (1986).
Slum of origin	Source: Archives and authors' calculations. Based on Archival data, MINVU (1979, 1984), Molina (1986), and Morales and Rojas (1986).
Municipality of destination	Source: Archives and authors calculations'. Based on Archival data, MINVU (1979, 1984), Molina (1986), and Morales and Rojas (1986).
Project of destination	Source: Archival records and electoral records in 2016. We updated the name of the projects using current names reported in families' addresses in 2016 that we observe in the electoral records.
Date of birth	From birth certificates
Age at intervention	Year of intervention minus year of birth

Female	From birth certificates
Mother head of household	We proxy head of household's gender using the gender of the person who received the property deed as it appears in the Archival Record.
Head of household's marital status	From marriage certificates we identify if a person is married or widowed at the time of the intervention. We cannot conclude a person is single if we did not find a marriage certificate because the older the couple, the less likely their marriage certificate is to be available on the Social Registry website.
Age of mother at birth	From birth certificates, year of intervention minus mother's year of birth.
Number of siblings	Number of children from the same couple minus one. We are not always able to observe half-siblings if parents remarried because we only observe the last marriage certificate.
Mother's education	Source: RSH. Constructed the same way as years of education. We correct this variable by weighting the observations by the inverse of the probability of being found in RSH \hat{p} . We compute this probability as the fitted values of a logit regression of the probability of being found in RSH on displaced, dead before 2007 and a full set of demographic controls at the time of the intervention. Then, we weight each observation by $1/\hat{p}$ if mother was displaced, and $1/(1 - \hat{p})$ if mother was non-displaced.
Mapuche last name	Source: Archival Records and Mapuche Data Project. We identify each last name as Mapuche if we find it in the list collected by the <i>Mapuche Data Project</i> . Data available here .

Slum/Neighborhood Characteristics

Area	Source: MINVU (1979, 1984). Land used by each slum measured in hectares.
# families	Source: MINVU (1979, 1984), Molina (1986). Number of families per slum.
Military name	Constructed by the authors. A slum is considered to have a military name if its name has a reference to any military name or date associated to a military event in the history of Chile.
Distance to river	Measures the distance in kilometers from a slum location to the closest riverbank in Greater Santiago. To georeference slums we use Morales and Rojas (1986), and rivers locations available here .

Census district	Smaller geographic unit than municipality. Source: Census of Population of 1992. Shape files of the census of 1982 were not available in the National Institute of Statistics; thus, we use the corresponding census districts in 1992 because the differences between 1982 and 1992 in the Greater Santiago were minor.
HH's schooling	Source: Census of Population of 1982. Average years of schooling of all heads of households between 18 and 65 years old by municipality, and by census district.
HH's unemployment	Source: Census of Population of 1982. Average unemployment rate of heads of households between 18 and 65 years old by municipality, and by census district.
HS dropout students	Source: Census of Population of 1982. Share of the population that is not in high school but should be as measured by their age. Measures at the level of municipality and census district.
# schools	Source: Ministry of Education. List of all schools in Chile, their location, type (private and public), and their year of inauguration. We keep all schools until year 1985. We measure number of schools per municipality and per census district, as well as the number of schools per 1,000 students by using the schooling population from the 1982 Census as the denominator.
# health care centers	Source: Ministry of Health. List of all public family care centers in Chile, their location, and their year of inauguration. We checked the years of inauguration one by one by calling each of the centers that had incorrect dates. We keep all health care centers until year 1985. We measure number of centers per municipality and per census district, as well as the number of centers per 1,000 households by using the total number of households per municipality/district in the 1982 Census as the denominator.
# hospitals	Source: Ministry of Health. List of all Public Hospitals in Chile, their location, and their year of inauguration. We keep all hospitals built until 1985. We measure number of hospitals per municipality and per census district, as well as the number of hospitals per 1,000 households by using as a denominator the total number of households per municipality/district in the 1982 Census.
Distance to subway	Source: Metro de Santiago. List of metro stations in Greater Santiago, their location, and year of construction. Distance is measured in kilometers as the distance between each slum/project of destination to the closest metro station built in or before 1985.

Waiting time	Source: Origin-Destination Survey, Santiago, 1977. Average waiting time in public transportation at the municipality level, measured in minutes. Unfortunately not available at a more granular geographic level.
Commuting time	Source: Origin-Destination Survey, Santiago, 1991. Average commute time on public transportation at the municipality level. Measured in minutes.
Property prices	Source: El Mercurio newspapers. We digitized and cleaned listings of property sales for years 1978, 1979, 1984 and 1985. Then we geocoded them and took the average of the logarithm of the price residuals at the census district level. To maximize the number of observations, we also use a buffer of 2 kilometers around slums and neighborhoods at baseline. This variable is only available in urban municipalities as we did not find many listings in rural areas.

Table C.3: Probability of finding a slum in archival data

Outcome	1[Found in archives]				
	(1)	(2)	(3)	(4)	(5)
Displaced	0.156** (0.056)	0.151** (0.058)	0.149** (0.057)	0.368*** (0.118)	0.404*** (0.119)
Families (/100)	0.034** (0.012)	0.031** (0.012)	0.026** (0.010)	0.017 (0.013)	0.016 (0.013)
Distance to subway (origin)	-0.020* (0.010)	-0.016 (0.018)	-0.004 (0.013)	-0.005 (0.023)	
<i>Characteristics at destination</i>					
Schooling			0.029 (0.025)	0.024 (0.023)	
Schools per student			-0.180 (0.139)	-0.146 (0.162)	
Distance to subway			-0.031 (0.018)	-0.039* (0.020)	
Families per project (/100)			-0.456*** (0.148)	-0.397*** (0.126)	
Fragmentation index			0.021 (0.014)	0.012 (0.012)	
Log property prices			-0.108 (0.086)	-0.110 (0.110)	
<i>R</i> ²	0.068	0.088	0.171	0.153	0.402
Sample mean	0.31	0.31	0.31	0.31	0.31
Observations	318	318	318	285	285
Municipality of origin FE			✓		✓

Notes: Each observation corresponds to a slum-destination pair. Data found in archives was harmonized with data in Morales and Rojas (1986). See table C.2 for description.

D ROBUSTNESS CHECKS

D.1 Spillovers and mother's years of education

Table D.1: Displacement effect and spillovers

Outcome	Labor income		
	Baseline		
	(1)	(2)	(3)
Displaced	-14.700** (6.701)	-14.442** (7.011)	-13.924 (8.280)
Non-displaced < 0.5km		4.544 (9.556)	5.287 (10.480)
Non-displaced < 1km			3.768 (12.215)
<i>R</i> ²	0.125	0.125	0.125
Observations	533,444	533,444	533,444

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH data and that report non-missing schooling. Standard errors clustered by municipality of origin in parenthesis, 10%*, 5%**, 1%***. This table splits the non-displaced group at baseline into two: Non-displaced without a displaced slum nearby (omitted category), and non-displaced with a displaced slum around 1 km or less. All regressions control for year of intervention fixed effects. Baseline controls include the following: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects.

Table D.2: Displacement effect in the sample of children with mothers in the RSH

Outcome	Income (1)	Income (2)	Employment (3)	Employment (4)	Schooling (5)	Schooling (6)
Displaced	-14.603** (6.747)	-13.852** (6.525)	0.005 (0.014)	0.004 (0.014)	-0.621*** (0.158)	-0.501*** (0.128)
<i>R</i> ²	0.122	0.123	0.098	0.098	0.113	0.138
Non-displaced mean	168.486	168.486	0.663	0.663	11.78	11.78
Observations	465,387	465,387	465,387	465,387	23,220	23,220
Municipality of origin FE	✓	✓	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓	✓	✓
Mother's schooling		✓		✓		✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH, that report non-missing schooling, and whose mothers are in the RSH. Standard errors clustered by municipality of origin in parenthesis, 10%*, 5%**, 1%***. All regressions control for year of intervention fixed effects. Baseline controls include: female, mother head of household, married head of household, head of household's marital status unknown, age of mother at birth, number of siblings, firstborn dummy, and year of birth fixed effects.

D.2 Displacement effect coefficient and sensitivity to omitted variable bias

In this appendix section we discuss a sensitivity analysis in our baseline regressions on earnings and years of schooling. Our goal is to estimate the degree of selection in unobservable characteristics under different scenarios following the framework proposed by [Oster \(2019\)](#).

Consider the following “short” and “long” regressions, of the form

$$Y_{it} = \alpha + \beta Displaced_{s\{i\}} + \psi_o + \varepsilon_{it} \quad (3)$$

$$Y_{it} = \tilde{\alpha} + \tilde{\beta} Displaced_{s\{i\}} + \tilde{\psi}_o + X'_{it}\theta + \tilde{\varepsilon}_{it}, \quad (4)$$

where Y_{it} is current outcome for individual i at time t , such as labor income, or years of schooling, $s(i)$ indexes the slum of origin for individual i 's family. The variable $Displaced_{s\{i\}}$ takes the value of 1 if an individual's family lived in a displaced slum and 0 otherwise. The

variable ψ_o are municipality of origin fixed effects. The matrix X_{it} includes baseline controls for individuals' and families' characteristics, such as gender, child's year of birth, female head of household, married head of household, head of household's age, birth-order dummies, mother's schooling, and year of intervention fixed effects (1979-1985). Under the assumption that X_{it} is uncorrelated with displacement, we would expect that $\beta = \tilde{\beta}$.

Following Oster (2019) we can use β , $\tilde{\beta}$ and the sample R^2 s from each regression to bound the true displacement effect defined by β^* when all confounders have been taken into account,

$$\beta^* \sim \tilde{\beta} + \delta(\tilde{\beta} - \beta) \frac{R_{max} - \tilde{R}}{\tilde{R} - R}, \quad (5)$$

where R and \tilde{R} are the R^2 s from equations (3) and (4) respectively, and R_{max} is the R^2 from the regression that controls for all confounding variables. The coefficient δ is the degree of proportional selection between the unobservable components relative to the observable variables. For example, $|\delta| = 1$ implies that the degree of selection on unobservables is equally important as the observables.

Then, we use equation (5) to bound the true value for β^* . First, we estimate β , β^* , R and \tilde{R} from equations (3) and (4). Second, we vary the values of δ and R_{max} , we choose $R_{max} = 1.3\tilde{R}$, as recommended by Oster (2019), and we also choose $R_{max} = 5\tilde{R}$ as a more conservative case. Then we vary the value of δ to be 1, 2 or 3. For example, Altonji et al. (2005) assume $\delta = 1$. Our results are in D.3.

Table D.3: Displacement effect under different assumptions on selection on unobservables

Outcome	R^2	max	$\hat{\delta}$	δ	$\hat{\beta}^*$	$\hat{\beta}^*$
Labor Earnings	1.3	36.72	1	-14.38	-13.34	
			2	-14.02	-12.73	
	1.3		3	-13.65	-12.09	
		2.80	1	-9.80	-3.55	
			2	-4.52	32.94	
	5		3	1.16	-189.12	
Years of Schooling	1.3	235.62	1	-0.681	-0.450	
			2	-0.680	-0.392	
	1.3		3	-0.680	-0.330	
		19.90	1	-0.676	0.647	
			2	-0.670	-7.664	
	5		3	-0.64	-4.997	
<i>Included controls:</i>						
Baseline controls				✓	✓	
Mother's schooling					✓	

The column labeled as $\hat{\delta}$ reports the estimate for δ for different values of R_{max} assuming the

true value of β^* is equal to 0. The results show that the degree of selection on unobservables would need to be greater than 2 to find a null displacement effect. In other words, under different values of δ that vary between 1 and 3, we find smaller magnitudes for the displacement effect, but they never become non-negative. The only case where we find a positive displacement effect, or very negative effects that are not (economically) plausible, is on years of schooling under the assumption that $R_{max} = 5\tilde{R}$ and when including mother's schooling as a control, which is a very extreme case not even suggested by [Oster \(2019\)](#).

D.3 Lee Bounds (Lee, 2009)

Table D.4: Lee Bounds for displacement effect on earnings

Outcome	Labor Income				
	(1)	(2)	(3)	(4)	(5)
<i>Panel A. Displacement Effect</i>					
Displaced	-13.908*	-12.717	-14.060*	-12.811	-12.077
	(7.438)	(7.369)	(7.483)	(7.423)	(7.081)
<i>Panel B. Lee Bounds</i>					
Lower	-19.628***	-18.718***	-18.076***	-17.076**	-13.629
	(2.426)	(6.680)	(3.447)	(8.049)	(10.392)
Upper	-3.667	-4.292	-4.171***	-4.705	-1.801
	(1.075)	(5.480)	(1.248)	(6.395)	(7.910)
Municipality of origin	✓	✓	✓	✓	✓
Gender		✓		✓	✓
Age			✓	✓	✓
Mother's schooling					✓
Observations (selected)	322,160	322,160	322,160	322,160	283,129

Notes: All regressions include year of intervention fixed effects. The total number of observations before trimming is 418,652.

D.4 Alternative standard error estimates

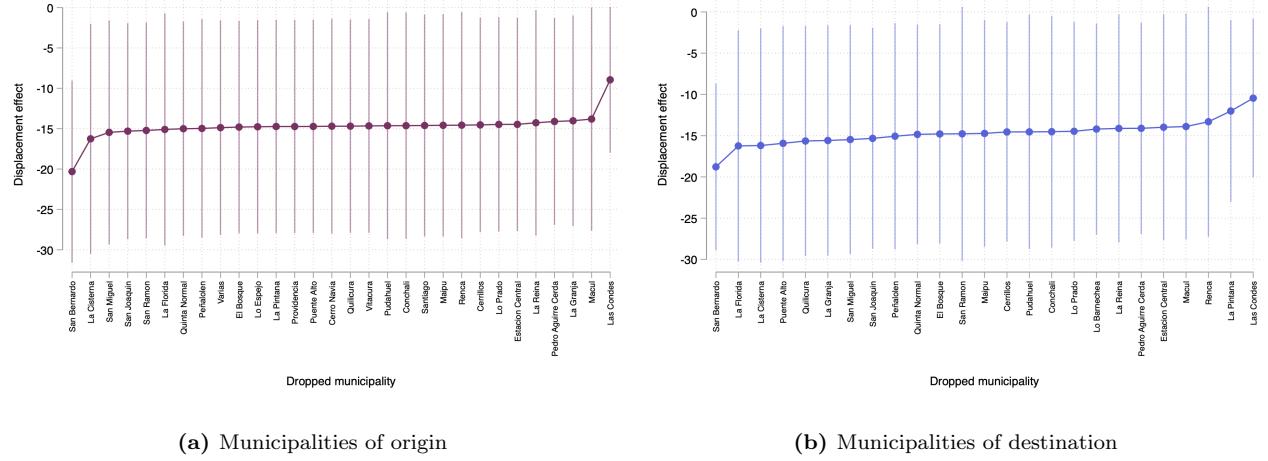
Table D.5: Conley standard errors

Outcome	Labor income	Years of schooling
Displacement coefficient	-15.470	-0.683
Clustered se by municipality of origin	6.701	0.154
Clustered se by slum	4.122	0.090
Conley se (cutoffs in km)		
2	1.134	0.122
3	1.157	0.124
4	1.182	0.122
5	1.210	0.122
6	1.244	0.121
7	1.282	0.122
8	1.325	0.122
9	1.357	0.122
10	1.384	0.122
11	1.410	0.122
12	1.425	0.122
13	1.432	0.123
14	1.432	0.123
15	1.431	0.123

Notes: This table reports estimates of Conley Standard errors on income and schooling for different distance cutoffs (Conley, 1999). The estimation procedure comes from Thiemo Fetzer. For more details see [here](#).

D.5 Other robustness checks

Figure D.1: Results robust to dropping each municipality once from sample. Results for Labor Income.



Notes: The figure plots the displacement coefficient from equation (1) for labor income and its 95% confidence interval dropping each municipality of origin one by one (panel (a)), or each municipality of destination one by one (panel (b)). Standard errors clustered by municipality of origin. All regressions include year of intervention fixed effects. Baseline controls include the following: female, mother head of household, single head of household, number of siblings, firstborn dummy, and cohort fixed effects.

Table D.6: Results robust to dropping municipalities that only expelled/received families

	(1) Baseline	(2) W/o muni. expelled	(3) W/o muni. received	(4) W/o both
<i>Panel A.</i>				
Displaced	-14.700** (6.701)	-13.510* (7.274)	-16.903* (8.449)	-17.024* (8.649)
R ²	0.125	0.120	0.131	0.128
N	533,444	380,000	371,330	319,220
<i>Panel B.</i>				
Displaced	0.005 (0.015)	0.010 (0.016)	0.000 (0.017)	0.002 (0.017)
R ²	0.101	0.105	0.105	0.107
N	533,444	380,000	371,330	319,220
<i>Panel C.</i>				
Displaced	-0.683*** (0.154)	-0.677*** (0.163)	-0.841*** (0.149)	-0.857*** (0.151)
R ²	0.115	0.116	0.127	0.127
N	26,675	19,041	18,659	16,051
Municipality of origin FE	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH and report non-missing schooling. Standard errors clustered at the municipality level. 10%, 5%, 1%. All regressions include year of intervention fixed effects. Baseline controls include the following: female, mother head of household, single head of household, number of siblings, firstborn dummy, mother's age at first birth, head of household's age, and cohort fixed effects.

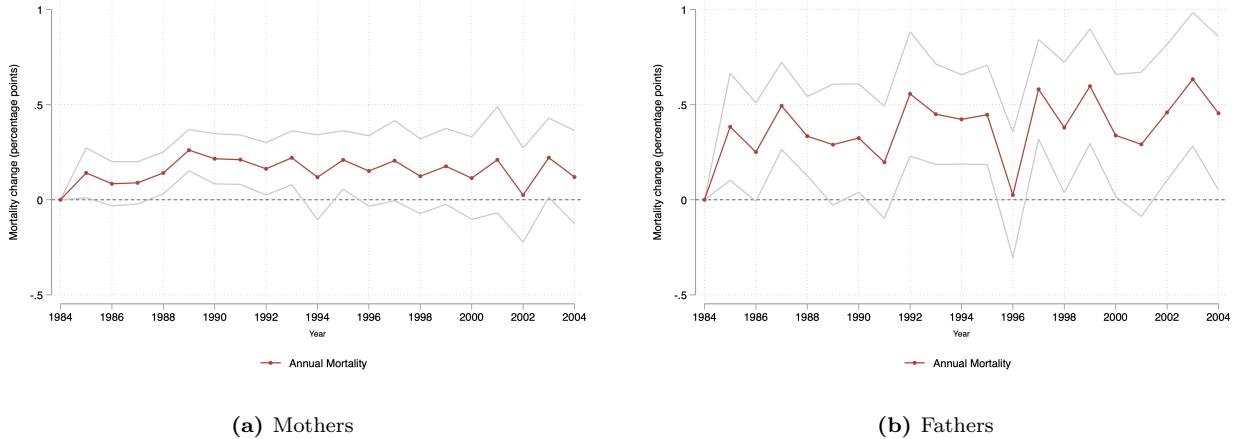
Table D.7: Other robustness checks: Control function and sub-periods

	(1) Baseline	(2) Polynomial match with RSH	(3) Polynomial match with ARNAD	(4) Period 1979-1984	(5) Before 1982
<i>Panel A.</i>					
		Outcome: Labor Income			
Displaced	-14.700** (6.701)	-15.752** (7.189)	-14.315** (6.941)	-14.512** (6.712)	-16.233* (9.201)
R ²	0.125	0.126	0.126	0.126	0.126
Observations	533,444	533,444	533,444	529,901	334,521
<i>Panel B.</i>					
		Outcome: 1[Employed]			
Displaced	0.003 (0.009)	0.000 (0.017)	0.007 (0.019)	0.004 (0.014)	0.006 (0.017)
R ²	0.101	0.099	0.101	0.099	0.101
Observations	533,444	533,444	533,444	529,901	334,521
<i>Panel C.</i>					
		Outcome: Years of Schooling			
Displaced	-0.683*** (0.154)	-0.676*** (0.168)	-0.688*** (0.158)	-0.682*** (0.154)	-0.873*** (0.144)
R ²	0.115	0.118	0.115	0.114	0.120
Observations	26,675	26,675	26,675	26,510	16,822
Municipality of origin FE	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓

Notes: Regressions for children aged 0 to 18 at baseline that are matched to the RSH and report non-missing schooling. Standard errors clustered by municipality of origin. 10%, 5%**, 1%***. All regressions include year of intervention fixed effects. Baseline controls include the following: female, mother head of household, single head of household, number of siblings, firstborn dummy, mother's age at first birth, head of household's age, and cohort fixed effects. Column (1) is the baseline regression, columns (2) and (3) control for the probability of finding a child in the RSH or finding a slum in the Archives, respectively. Column (4) restricts the sample to years 1979 to 1984, and column (5) restricts the sample to years 1979 to 1982.

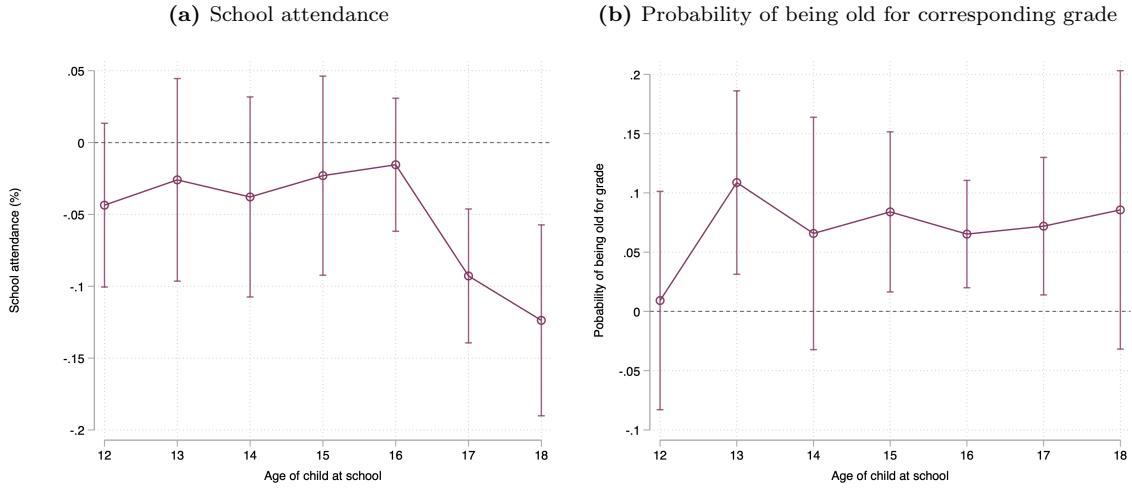
E DISPLACEMENT EFFECT ON OTHER FAMILY MEMBERS

Figure E.1: Annual mortality of parents from 1985 to 2004



Notes: The figure plots the coefficients β_τ and their 95% confidence intervals from regression $Died_{it} = \sum_{\tau=1985}^{2019} \beta_\tau 1(t = \tau) \cdot Displaced_{s\{i\}} + X_i'\theta + \psi_o + \gamma_t + \varepsilon_{it}$. Plotted coefficients until 2004 for better exposition. We follow Deryugina and Molitor (2020) to set up the data. Panel (a) estimates the displacement effect on annual mother's mortality and panel (b) does the same for fathers. These regressions are for households with children at the time of the intervention.

Figure E.2: Displacement effect on school attendance of children born 1 to 5 years after intervention



Notes: Regressions for children born 1 to 5 years after the intervention and matched to school enrollment administrative data from Chile's Ministry of Education (MINEDUC). Clustered standard errors at the municipality level. Controls include: female, mother head of household, married head of household, number of siblings, first born dummy, head of household's marital status unknown, and year of birth fixed effects. Figures plot coefficients β_τ and their 95% confidence intervals from regression: $y_{it} = \sum_{\tau=12}^{18} \beta_\tau Displaced * 1[Age at school = \tau] + \sum_{\tau=12}^{18} \delta_\tau 1[Age at school = \tau] + \psi_o + X_{it}'\gamma + u_{it}$.

Table E.1: Summary statistics for the full sample of families

Variables	Full sample			Families with children		
	Displaced mean	Non-displaced mean	Difference (within municip.)	Displaced mean	Non-displaced mean	Difference (within municip.)
<i>Demographics at baseline</i>						
Head of household age	35.47	37.28	-1.00 (0.58)	34.52	36.33	-1.1 (0.52)
Wife age	34.04	35.78	-0.79 (0.60)	32.98	34.81	-1.02 (0.52)
Husband age	35.29	37.09	-1.11 (0.54)	34.54	36.24	-1.1 (0.50)
Female HH	0.36	0.36	0.01 (0.02)	0.34	0.35	-0.01 (0.02)
Married HH	0.74	0.78	-0.03 (0.01)	0.75	0.80	-0.03 (0.01)
Widowed HH	0.02	0.02	0.00 (0.00)	0.01	0.01	0.00 (0.00)
HH's marital status unknown	0.11	0.08	0.02 (0.01)	0.13	0.10	0.02 (0.01)
Mapuche HH	0.06	0.05	0.02 (0.01)	0.06	0.05	0.02 (0.01)
# children	2.25	2.29	-0.05 (0.06)	2.51	2.59	-0.05 (0.07)
No. children	0.10	0.12	0.01 (0.01)			
Age of youngest child	6.26	7.11	-0.57 (0.28)	6.26	7.11	-0.57 (0.28)
Age of oldest child	12.00	12.86	-0.73 (0.35)	12.00	12.86	-0.73 (0.35)
<i>Demographics measured between 2007 and 2019</i>						
Female's schooling (raw)	6.09	6.39	-0.54 (0.21)	6.15	6.46	-0.50 (0.21)
Female's schooling (corrected)	6.10	6.20	-0.34 (0.25)	6.18	6.26	-0.30 (0.25)
Male's schooling (raw)	6.61	6.99	-0.46 (0.20)	6.65	7.07	-0.43 (0.20)
Male's schooling (corrected)	6.71	6.51	-0.07 (0.29)	6.75	6.65	-0.07 (0.24)
Share HH in the RSH	0.71	0.74	-0.04 (0.02)	0.74	0.77	-0.04 (0.02)
Observations	11,327	5,221	16,548	10,146	4619	14,765

Notes: Within difference correspond to the coefficient of *displaced* in equation (1) conditional on municipality of origin and year of intervention. Standard errors clustered by municipality of origin.

Marital status of married and widowed individuals is computed conditional on finding a marriage certificate or spouse's death certificate. Families with children are all families with at least one child at the time of the intervention regardless of the age of their children. "Corrected" means the difference between displaced and non-displaced controls for the probability of finding a mother or a father in the RSH data.

Table E.2: Annual mortality of adults

Outcome	Mother died (1)	Mother died (2)	Father died (3)	Father died (4)
Panel A. Full sample				
Displaced	0.0014 (0.0009)	0.0021* (0.0011)	0.0042** (0.0019)	0.0047** (0.0019)
<i>R</i> ²	0.0148	0.0149	0.0177	0.0178
Non-displaced mean	0.007	0.007	0.011	0.011
%Var. w.r.t. non-disp.	20.0	30.0	38.2	42.2
Cumulative effect from 1985 to 2019	0.049	0.074	0.153	0.173
Observations	587,062	587,062	478,359	478,359
Individuals	18,080	18,080	15,709	15,709
Panel B. Households with children				
Displaced	0.0007 (0.0010)	0.0010 (0.0012)	0.0039** (0.0018)	0.0042** (0.0019)
<i>R</i> ²	0.0112	0.0113	0.0153	0.0155
Non-displaced mean	0.006	0.006	0.010	0.010
%Var. w.r.t. non-disp.	11.7	16.7	39.0	42.0
Cumulative effect from 1985 to 2019	0.024	0.035	0.141	0.153
Observations	531,650	531,650	435,527	435,527
Individuals	16,149	16,149	14,122	14,122
Municipality of origin FE	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓
Municipality of destination FE		✓		✓

Notes: Separate regressions for men and women in our archival sample. Standard errors clustered by municipality of origin. 10%*, 5%**, 1%***. Outcome is annual mortality, and to account for survival the data is set up as in [Deryugina and Molitor \(2020\)](#).

Regression is $Died_{it} = \alpha + \beta Displaced_{s\{i\}} + X_i'\theta + \psi_o + \gamma_t + \varepsilon_{it}$. All regressions include year of intervention fixed effects and calendar year fixed effects from 1985 to 2019. Baseline controls include the following: marital status, Mapuche last name, head of household dummy, number of children at baseline, and cohort fixed effects.

Table E.3: Adults' labor market outcomes, heads of households in the RSH

Outcome	(1) 1[Employed]	(2) Total income	(3) Labor income	(4) Retirement income
Panel A. All head of households in the RSH				
Displaced	0.036 (0.033)	-16.330*** (5.323)	-6.773 (7.355)	-20.404* (10.234)
Non-displaced mean	0.385	100.05	77.35	72.43
R^2	0.225	0.304	0.207	0.161
Observations	240,568	240,568	240,568	240,568
<i>Individuals</i>	12,515	12,515	12,515	12,515
Panel B. Parents younger than 65 years old				
Displaced	0.035** (0.013)	-14.102*** (3.203)	-9.401* (4.998)	-16.794*** (3.589)
Non-displaced mean	0.602	105.64	128.23	33.19
R^2	0.124	0.295	0.159	0.083
Observations	105,612	105,612	105,612	105,612
<i>Individuals</i>	8,331	8,331	8,331	8,331
Panel C. Parents older than 65 years old				
Displaced	0.065*** (0.016)	-12.006*** (2.574)	3.120 (3.422)	-25.501*** (4.272)
Non-displaced mean	0.286	97.52	53.78	90.61
R^2	0.152	0.318	0.148	0.068
Observations	134,956	134,956	134,956	134,956
<i>Individuals</i>	10,204	10,204	10,204	10,204
Municipality of origin FE	✓	✓	✓	✓
Baseline Controls	✓	✓	✓	✓

Notes: Regressions for head of households matched to the RSH data. Standard errors clustered by municipality level. 10%, 5%, 1%. Controls include the following: female head of household, married head of household, marital status unknown, age at intervention, and cohort fixed effects. All regressions include year of intervention fixed effects.

Table E.4: Displacement effects for children born to treated families

Outcome	Employed (1)	Labor income (2)	Taxable income (3)	Years of schooling (4)	HS graduate (5)	College attendance (6)
Displaced	0.026 (0.021)	-0.074 (6.771)	-19.408 (16.787)	-0.475*** (0.130)	-0.051*** (0.015)	-0.002 (0.020)
Non-displaced mean	0.59	123.20	705.31	12.32	0.80	0.16
% Variation w.r.t. non-disp.	4.4	-0.06	-2.8	-1.1	-6.4	-1.25
<i>R</i> ²	0.093	0.110	0.137	0.076	0.064	0.049
Observations	94,129	94,129	19,346	4,218	4,218	4,218
Individuals	4,665	4,665	4,367	4,218	4,218	4,218
Municipality of origin FE	✓	✓	✓	✓	✓	✓
Controls	✓	✓	✓	✓	✓	✓

Notes: Regressions for children born 1 to 5 years after the intervention that are matched to the RSH data, and report non-missing schooling. Standard errors clustered by municipality of origin. 10%, 5%, 1%. Baseline controls include the following: female, mother head of household, married mother at birth, age of mother at birth, number of siblings, Mapuche last name, cohort fixed effects, year of treatment fixed effects.

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