Computational Many-Body Physics - Sheet 2

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The full implementation of all exercises can be found under https://github.com/Fhoeddinghaus/cmbp22-exercises/tree/main/sheet_2.

1. TASEP

As the TASEP corresponds to rule 184, the basis for this exercise is the implementation of rule N from the last sheet, but with one crucial difference: the periodic boundary conditions. To implement these conditions, the values of the left (right) neighbour of the first (last) cell are set to the value of the last (first) cell:

```
function f(N::Int, z::Vector, i::Int)

...

# periodic boundary condition

a = z[end]

...

# periodic boundary condition

c = z[begin]

end
```

Listing 1: The function f(N, z, i) applies the ruleset for rule N and calculates the next cell value of cell i using the given current configuration z. The full implementation can be found in the repository in rule_N.jl.

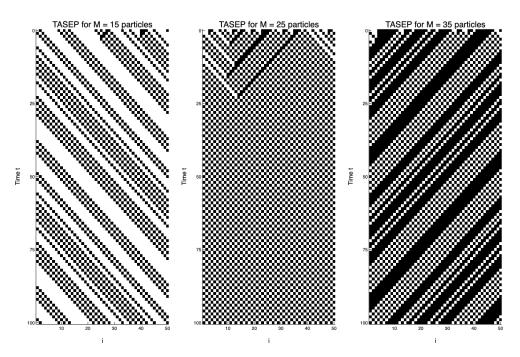


Figure 1.1: Examples of the TASEP (Rule 184) for M=15,25,35.

a). Flow

To generate a random start configuration for a fixed number of particles M, we can use the following code snippet¹:

```
# fill random cells with particle
z_start = zeros(Int, NUMBER_OF_SITES)
while sum(z_start) < M
i = rand(1:NUMBER_OF_SITES)
z_start[i] = 1
end</pre>
```

Listing 2: Generate a random start configuration with M particles.

Using the function calculate_rule_N(z_start, 184) 2 , we can now calculate the time evolution from the start configuration for a globally set NUMBER_OF_TIMESTEPS = 100 timesteps.

To now calculate the flow (number of particles per unit of time transferred from site NUMBER_OF_SITES to site 1), we can use the **flow()** function:

Listing 3: Function to calculate the flow from the configurations stored in zs. To calculate the flow, we only have to watch the value of the first cell and count up, if the value changes from 0 to 1, because then a transfer from site NUMBER_OF_SITES to 1 has occurred.

The flow for M=25 is approximately $0.48 \frac{\text{particles}}{\text{unit of time}}$.

 $^{^1}$ This exercise can be found in the notebook 1_TASEP.ipynb.

 $^{^2}$ Implemented in rule_N.jl

b). Flow vs. density

To calculate the flow vs. density diagram, we iterate through the different values of $M \in [0, \text{NUMBER_OF_SITES}]$ and construct NUMBER_OF_TRIES different random start configurations for each M, calculate for each start configuration the time evolution and the flow and finally average over the calculated flows:

```
NUMBER_OF_TRIES = 1000
M_start, M_end = 0, NUMBER_OF_SITES
      store all calculated flows
    flows = zeros(NUMBER_OF_SITES + 1)
    for M in M_start:M_end
                             flow for any given M NUMBER_OF_TRIES times
         for i in 1:NUMBER_OF_TRIES
                random start config with M particles
10
11
              # fill random cells with particle
              z_start = zeros(Int, NUMBER_OF_SITES)
while sum(z_start) < M
   i = rand(1:NUMBER_OF_SITES)</pre>
12
13
14
15
                   z_start[i] = 1
16
18
              \# calculate NUMBER_OF_TIMESTEPS generations
              zs = calculate_rule_N(z_start, N_rule)
19
20
21
              # calculate the flow for this configuration
22
              flows[M+1] += flow(zs)
23
          # average over the number of tries
24
          flows[M+1] = flows[M+1]/NUMBER_OF_TRIES
```

Lastly, we calculate the densities $\rho = \frac{M}{N}$:

```
1 
hos = collect(M_start:M_end)./NUMBER_OF_SITES # density 
ho = M/N
```

The resulting flow-density-diagram for NUMBER_OF_TRIES = 1000 tries per number of particles M is:

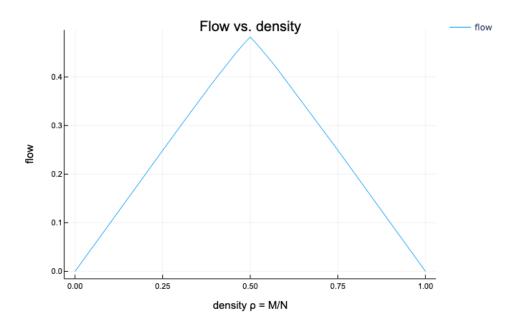


Figure 1.2

2. Metropolis algorithm

a). Implementation

The Metropolis algorithm³ generates a Markov-chain of random numbers $\{x_i\}_i$ with distribution w(x):

- 1. From a random x_i , calculate the next potential random number \bar{x}_{i+1} :
 - (a) choose a random $\delta \in [-1, 1]$ (equally distributed)
 - (b) calculate $\bar{x}_{i+1} = x_i + h\delta$, with step size h
- 2. Calculate $\alpha = \frac{w(\bar{x}_{i+1})}{w(x_i)}$
- 3. Choose a random $\gamma \in [0,1]$ (equally distributed), then: if
 - $\alpha \geq \gamma$: accept new value, $x_{i+1} := \bar{x}_{i+1}$
 - $\alpha < \gamma$: reject new value, $x_{i+1} = x_i$

go to step 1.

At first, we introduce a small helper function $rand_range()$ to calculate equally distributed random numbers in a given interval [a, b]:

```
## random number between a and b
# 0 ≤ rand() ≤ 1 → transform range
function rand_range(a,b)
return abs(b-a) * rand() + a
end
```

Listing 4: Function to calculate a random number in [a, b]

To calculate a Markov-chain with number_xs elements, starting with a value x_start, we use the function markov_chain():

```
# generate a Markov chain from Metropolis algorithm
     function markov_chain(
              x_start, # start value
number_xs, # number of elements in chain
               h, # step size
               w # distribution of the values
          # array to store the random numbers
          xs = zeros(number_xs)
10
          xs[1] = x_start
          for i in 1:(number_xs-1)
              # calculate next potential number \delta = rand_range(-1,1)
14
               y = xs[i] + h * \delta
16
              \alpha = w(y)/w(xs[i]) # probability of acceptance \gamma = rand_range(0, 1)
18
19
               \mathbf{if}\ \alpha\,\geq\,\gamma
                    # accept new value
20
21
                    xs[i+1] = y
22
23
                    # reject new value, keep current
24
                    xs[i+1] = xs[i]
25
               end
26
          end
27
          return xs
```

Listing 5: Function to generate a Markov-chain using the Metropolis algorithm.

³This exercise can be found in the notebook 2_Metropolis.ipynb.

The calculation then looks like this

```
# define the distribution
w(x) = 1/sqrt(\pi) * exp(-x^2)

# generate the markov chain
x_start = 0
number_xs = 1_000_000
step_size = 1
xs = markov_chain(x_start, number_xs, step_size, w);
```

Listing 6: Calculating the Markov-chain for a given distribution w(x) with 1,000,000 elements

To show, that the random values in the Markov-chain are indeed distributed as given by $w(x) = \frac{1}{\sqrt{\pi}}e^{-x^2}$, we only have to plot the histogram of $\{x_i\}_i$ (count of the values in specific intervals (bins) vs. value) and compare that with the given distribution:

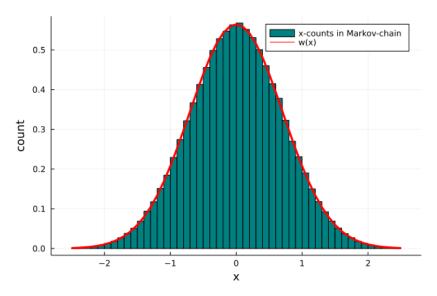


Figure 2.1: Comparison of the Markov-chain and the given distribution.

b). Markov-chain without unchanged values

There are two possible variants of this case:

- (a) rejecting the unchanged value and trying again with a new random value, until number_xs random values are calculated ("rejection sampling": number of values is preserved, number of iterations/tries increases)
- (b) removing the unchanged value from the chain, "skipping" the iteration (the number of values reduces for each skip)

In this exercise, we only consider variant $(b)^4$, the adjusted algorithm is implemented in the function $markov_chain_without_unchanged()$:

```
generate a Markov chain from Metropolis algorithm,
      but delete the rejected (unchanged) values (number of elements reduces)
    function markov_chain_without_unchanged(
             x_start, # start value
             number\_tries, # number of tries \neq number of elements in chain
             w # distribution of the values
           array to store the random numbers
9
         xs = Vector{Float64}() # empty array
10
         push!(xs, x_start) # push first element to array
12
         for i in 1:(number_tries-1)
13
             \# calculate next potential number \delta = rand_range(-1,1)
14
16
             y = xs[end] + h * \delta \# instead of x at index i, look at the last element
             \alpha = w(y)/w(xs[end]) # probability of acceptance
19
                = rand_range(0, 1)
             if \alpha \geq \gamma # accept new value
20
21
                  push! (xs, y)
22
                  # reject new value, but DON'T keep current (total possible number of
                        elements reduces by 1)
                  continue
25
             end
26
27
         end
28
         return xs
```

Listing 7: Adjusted Metropolis algorithm with removed unchanged values.

The resulting Markov-chain contains approximately 73% of number_tries elements.

The resulting distribution of random values in comparison to the normal Metropolis algorithm:

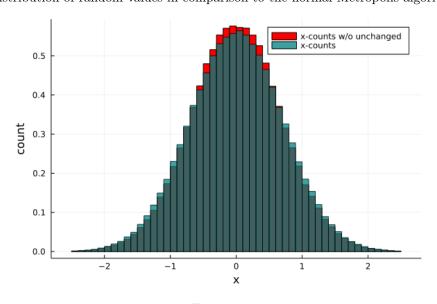


Figure 2.2

 $^{^4\}mathrm{Variant}$ (a) is implemented in 2_Metropolis.ipynb in the function markov_chain_rejection_sampling()

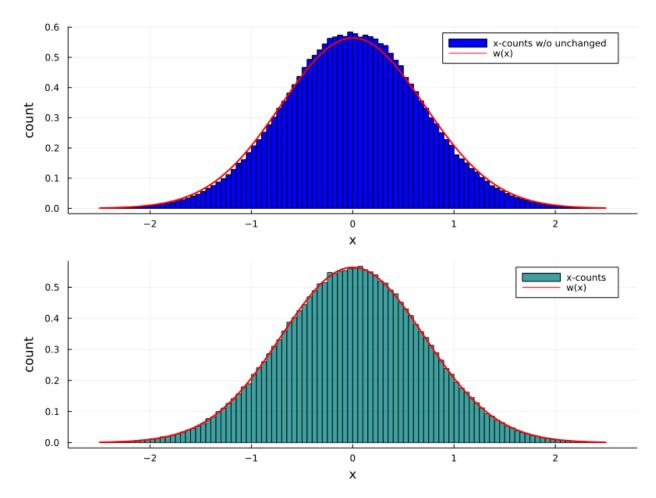


Figure 2.3: Comparison of the Markov-chain generated with the adjusted Metropolis algorithm (top) and the normal Metropolis algorithm (bottom).

Both Markov-chains are very close to the given distribution w(x), but the distribution of the values of the adjusted algorithm (without the unchanged values) is a bit narrower and taller than the given distribution.⁵

⁵An animated side-by-side comparison of both algorithms with the given distribution for different numbers of elements/tries can be found in the repository at 2b.gif or 2b.mp4.