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## ARTICLE

# Surrogate End Points for Overall Survival in Loco-Regionally Advanced Nasopharyngeal Carcinoma: An Individual Patient Data Meta-analysis

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#### **Abstract**

Background: Our objective was to evaluate progression-free survival (PFS) and distant metastasis-free survival (DMFS) as surrogate end points for overall survival (OS) in randomized trials of chemotherapy in loco-regionally advanced nasopharyngeal carcinomas (NPCs).

Methods: Individual patient data were obtained from 19 trials of the updated Meta-Analysis of Chemotherapy in Nasopharyngeal Carcinoma (MAC-NPC) plus one additional trial (total = 5144 patients). Surrogacy was evaluated at the individual level using a rank correlation coefficient  $\rho$  and at the trial level using a correlation coefficient  $R^2$  between treatment effects on the surrogate end point and OS. A sensitivity analysis was performed with two-year PFS/DMFS and five-year OS.

Results: PFS was strongly correlated with OS at the individual level ( $\rho = 0.93$ , 95% confidence interval [CI] = 0.93 to 0.94) and at the trial level ( $R^2 = 0.95$ , 95% CI = 0.47 to 1.00). For DMFS, too, the individual-level correlation with OS was strong ( $\rho = 0.98$ , 95% CI = 0.98 to 0.98); at trial level, the correlation was high but the regression adjusted for measurement error could not be computed (unadjusted  $R^2 = 0.96$ , 95% CI = 0.94 to 0.99). In the sensitivity analysis, two-year PFS was highly correlated with five-year OS

at the individual level ( $\rho=0.89$ , 95% CI = 0.88 to 0.90) and at the trial level ( $R^2=0.85$ , 95% CI = 0.46 to 1.00); two-year DMFS was highly correlated with five-year OS at the individual level ( $\rho=0.95$ , 95% CI = 0.94 to 0.95) and at the trial level ( $R^2=0.78$ , 95% CI = 0.33 to 1.00).

Conclusions: PFS and DMFS are valid surrogate end points for OS to assess treatment effect of chemotherapy in locoregionally advanced NPC, while PFS can be measured earlier.

Nasopharyngeal carcinoma (NPC) is a rather uncommon cancer and bears a unique pattern of geographical distribution. The International Agency for Research on Cancer (1) estimated 85 500 new cases in 2012, over 60 000 of those in Eastern and South-Eastern Asia. Given these relatively low and geographically heterogeneous incidence rates compared with other cancers and the rapid pace of technical and pharmaceutical developments, conducting research in this disease is challenging. One of the ways to speed up clinical research is through the use of end points that could be obtained earlier than overall survival (OS). Such end points, to be clinically relevant, should be consistently defined and measured across trials, correlated to the gold-standard end point both at the trial and patient levels, and available earlier than this latter end point.

The individual patient data meta-analysis of chemotherapy in nasopharyngeal carcinoma (MAC-NPC) included most randomized trials conducted up to 2010 evaluating the role of adding chemotherapy to radiotherapy in nonmetastatic NPC patients (2). Given its exhaustiveness and the availability of individual patient data, it is a unique database to validate surrogate end points in NPC, as was previously performed by our team for squamous cell head and neck (H&N) carcinoma (3) and lung cancer (4). The primary objective of the present study was to evaluate progression-free survival (PFS) as surrogate end point for OS in randomized trials evaluating cytotoxic chemotherapy in NPC. Given the progressive shift from local to distant progressions observed with the use of intensity modulated radiotherapy and concomitant chemotherapy (5), our secondary objective was to assess the validity of distant metastasis-free survival (DMFS) as a surrogate end point for OS in this setting.

## **Methods**

## Patients and Study Objectives

Individual patient data (n = 5144) were available from 19 randomized trials and included the updated MAC-NPC metaanalysis (2,6) plus one trial (7). All trials in the MAC-NPC meta-analysis compared radiotherapy (RT) with RT plus chemotherapy (CT), or a treatment strategy (RT plus concomitant or induction or adjuvant CT) with the same strategy plus CT at another timing. Both published and unpublished trials were included. Patients were recruited between 1988 and 2010. The overall median follow-up time was 7.7 years. One more trial that compared two timings of CT (7) was included in the present analysis. One 2x2 trial (8) that randomized the use of CT and RT fractionation was counted as two separate comparisons, one for each RT fractionation regimen. Another 2x2 trial (9) that randomized adjuvant CT and concomitant CT was duplicated and counted as four comparisons. In total, there were 24 comparisons in the analyzed data set. Six patients were excluded because of missing progression data (7). Seventy-four further patients from the HeCOG (10) and INT-0099 (11), having missing information about the type of progression, were excluded from the analyses for DMFS. In total, 5360 patients were analyzed,

with a median of 175.5 patients per trial (range = 65–509). Supplementary Table 1 (available online) describes the comparisons. All patients were analyzed according to the intention-to-treat principle.

The preplanned primary objective was the assessment of PFS as a surrogate end point for OS in patients with locoregionally advanced NPC. The secondary objective was the assessment of DMFS as a surrogate end point for OS in the same population. This retrospective analysis of clinical trial data was part of the meta-analysis protocol, which was approved by the Gustave Roussy Institutional Review Board, and the statistical methodology was preplanned (http://goo.gl/4Vitzs).

## **Sensitivity Analyses**

In order to enhance interpretation from the clinician's point of view, we contrasted in a preplanned sensitivity analysis the two-year surrogate end point vs the five-year OS, which reflect typical trial conditions. Preplanned sensitivity analyses were also performed by calendar period: the nine oldest trials (9,11–17) were reanalyzed separately from the 11 most recent ones (7,8,10,18–25).

#### **End Point Definitions**

Overall survival (OS) was defined as the time from random assignment until death from any cause. Progression-free survival (PFS) was the time from random assignment until first progression (loco-regional or distant) or death from any cause. Distant metastasis-free survival (DMFS) was the time from random assignment until distant progression or death from any cause. Because in some trials only the first event was recorded, patients with a loco-regional progression as first event were censored for DMFS. If both a loco-regional progression and a distant failure were recorded at the same time, patients were considered as having an event for DMFS. Patients alive and free from events at the end of the study were censored at the date of last follow-up. Of note, the previous study in H&N (3) included both trials of patients with resectable tumors and trials of patients with nonresectable tumors. Event-free survival (EFS) was used to denote disease- and progression-free survival, respectively, for resectable and nonresectable tumors. Thus, considering here PFS for trials including only patients with nonresectable tumors coincides with the definition of EFS in the H&N study. To describe the potential benefit of employing each surrogate end point as compared with OS, we computed the number of patients for which the surrogate was observed before death and the ratio of the median surrogate time to the median OS time (Kaplan-Meier estimates). The 95% confidence interval (CI) of the ratio between the median times was based on the delta method.

#### Statistical Methods

A correlation approach was used to assess the validity of each end point as surrogate for OS (26), as already used by Sargent et al. (27) and Buyse et al. (28) in colon cancer, by Burzykowski et al. (29) in breast cancer, by Michiels et al. (3) in locally

advanced H&N cancer, by Oba et al. (30) and Paoletti et al. (31) in gastric cancer, and by Mauguen et al. (4) in lung cancer. This approach investigates the correlation at a trial level and at an individual level.

The association between distributions of OS and the candidate surrogate end point was evaluated by a bivariate survival model (32,33) (copula). Clayton and Plackett copula models were fitted (34); the Plackett copula was chosen because of better convergence. Based on the copula dependence parameter, we computed the Spearman rank correlation coefficient  $\rho$ ; a Spearman coefficient close to 1 indicates a strong correlation.

Treatment effects were estimated by log hazard ratios in bivariate survival models. The correlation between treatment effect on the surrogate and treatment effect on OS was quantified through a linear regression model accounting via an error-in-variables model (35) for the uncertainty about the effects estimated by copulas. Whenever such a model failed to converge, we approximated it using the simpler model with estimated effects taken as fixed, weighted by the trial size. If the estimated R<sup>2</sup> was close to 1, then the risk reduction for OS was considered strongly correlated with the risk reduction for the candidate surrogate. As done previously (4), the R<sup>2</sup> was considered excellent if higher than 0.9, very good if higher than 0.75, good if higher than 0.5, moderate if higher than 0.25, and poor otherwise.

One objective of a surrogate end point is to predict the treatment effect on OS, based on the treatment effect on the surrogate end point. The surrogate threshold effect (STE) (36) was computed in order to estimate this prediction threshold: The STE is defined as the minimum treatment effect that is necessary on the surrogate to be able to predict a nonzero effect on OS. In other words, the STE corresponds to the smallest estimated treatment effect on the surrogate, which is predictive of a statistically significant effect of the treatment on OS, irrespectively of its clinical significance, which has to be considered case by case. Its calculation was based on the linear regression used for the determination of trial-level surrogacy.

A leave-one-out cross-validation was used to evaluate model predictions (3). For each trial, the bivariate model for OS and the surrogate end point was refitted on the remaining trials to predict the treatment effect on OS in the left-out trial, based on the treatment effect on the surrogate end point. For each trial, the

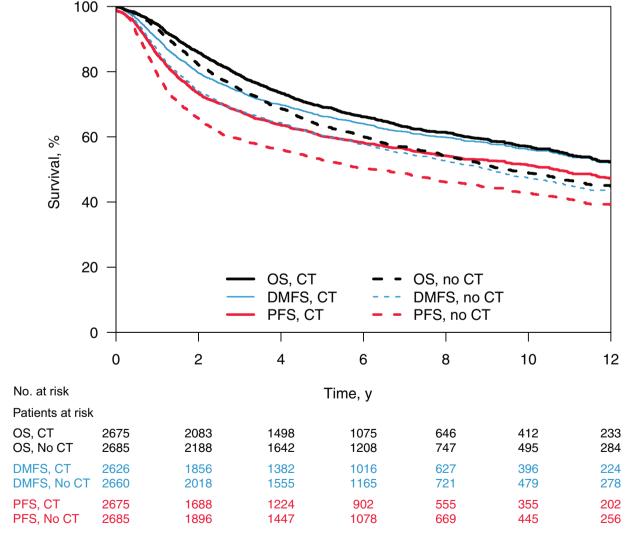


Figure 1. Survival curves for overall survival, progression-free survival, and distant metastasis-free survival in chemotherapy and control arms. Overall survival, progression-free survival, and distant metastasis-free survival in the chemotherapy and control arms. CT = chemotherapy; DMFS = distant metastasis-free survival; no CT = control; PFS = progression-free survival; OS = overall survival.

Table 1. Cumulative number of events for overall, progression-free, and distant metastasis-free survival at 2, 5, and 10 years

End point 2 years No. (%) 5 years No. (%) 10 years No. (%)

End point	2 years No. (%)	5 years No. (%)	10 years No. (%)	Total No. (%)
Overall survival	837 (38.7)	1658 (76.7)	2039 (94.3)	2162 (100.0)
Control arm	468 (40.6)	897 (77.7)	1099 (95.2)	1154 (100.0)
Chemotherapy arm	369 (36.6)	761 (75.5)	940 (93.3)	1008 (100.0)
Progression-free survival	1613 (62.0)	2224 (85.5)	2501 (96.2)	2600 (100.0)
Control arm	907 (65.2)	1206 (86.6)	1350 (97.0)	1392 (100.0)
Chemotherapy arm	706 (58.4)	1018 (84.3)	1151 (95.3)	1208 (100.0)
Distant metastasis-free survival	1270 (54.4)	1891 (81.1)	2218 (95.1)	2333 (100.0)
Control arm	716 (57.0)	1026 (81.7)	1205 (95.9)	1256 (100.0)
Chemotherapy arm	554 (51.4)	865 (80.3)	1013 (94.1)	1077 (100.0)

direct estimate of the hazard ratio was compared with the predicted hazard ratio and 95% prediction interval. The analyses were performed using SAS (version 9.3, SAS Institute, Cary, NC) and R (version 3.2.2, R foundation for Statistical Computing, Vienna, Austria). All statistical tests were two-sided.

#### **Results**

Overall 2162 patients (40.3%) died, among which 1690 (78.2%) had a progression recorded before death. This progression (first event) was a loco-regional progression in 684 patients (31.6%) and a distant metastasis in 1006 (46.5%). The estimated median time was 8.2 years for PFS and 11.3 years for OS, with a ratio of 0.726 (95% CI = 0.723 to 0.730), suggesting that a trial based on PFS could be roughly 27% faster than a trial based on OS. The median DMFS time was 10.9 years; the DFMS-to-OS ratio was 0.964 (95% CI = 0.960 to 0.968). Using DMFS instead of OS could speed up a clinical trial only by 4%. Figure 1 shows the OS, PFS, and DMFS curves by treatment group, and Table 1 gives the number of events at two, five, and 10 years. Supplementary Figure 1 (available online) shows the cumulative probability of deaths, metastases, and loco-regional progressions. Most of events occur in the first five years, 10% to 20% after five years, and 5% after 10 years. Of note, the number of PFS events at two years (n = 1613) is similar to the number of OS events at five years (n = 1658).

The results of the surrogacy assessment are summarized in Table 2 and detailed in Supplementary Tables 2 and 3 (available online). Supplementary Table 4 (available online) shows good consistency between results obtained with Clayton and Plackett copulas. The individual-level association between PFS and OS was strong, with rank correlation coefficient  $\rho=$  0.93 (95% CI =0.93 to 0.94) (Table 2). The squared linear correlation coefficient R<sup>2</sup>—adjusted for measurement error—between treatment effects on PFS and OS at the trial level was 0.95 (95% CI = 0.47 to 1. 00) (Figure 2A). The surrogate threshold effect for PFS was 0.89 (Table 2). The individual-level association between DMFS and OS was strong, with a rank correlation coefficient  $\rho$  of 0.98 (95%) CI = 0.98 to 0.98). The regression model adjusted for measurement error failed to converge for DMFS, thus trial-level surrogacy was assessed without adjustment. The squared linear correlation coefficient R2 between treatment effects on DMFS and OS was 0.96 (95% CI = 0.94 to 0.99) (Figure 2B). The STE for DMFS was 0.98 (Table 2).

Figure 3 shows the results from leave-one-trial-out cross-validation. The 95% prediction intervals for the treatment effect on OS contained the observed effect in 22 of 24 comparisons when prediction was based on the observed treatment effect on PFS. When prediction was based on the observed treatment effect on DMFS, the surrogate model failed to converge in two

comparisons. The prediction intervals for the treatment effect on OS contained the observed effect in 19 of the 22 remaining comparisons.

In a sensitivity analysis, we censored the surrogate end points at two years and OS at five years to reflect clinical trial conditions. Under this constraint, PFS was slightly less correlated to OS both at the individual level ( $\rho = 0.89, 95\%$  CI = 0.88 to 0.90) and at the trial level ( $R^2 = 0.85$ , 95% CI = 0.46 to 1.00), but correlations remained very good. DMFS, too, was slightly less correlated with OS with restrained follow-up, with excellent individual-level correlation ( $\rho = 0.95$ , 95% CI = 0.94 to 0.95) and very good trial-level correlation ( $R^2 = 0.78$ , 95% CI = 0.33 to 1.00) (Table 2). As the incidence of events increased up to three years, we also performed an unplanned analysis with three-year PFS vs five-year OS, which showed similar individual correlation ( $\rho$ = 0.90, 95% CI = 0.90 to 0.91) and lower trial correlation (unadjusted  $R^2 = 0.74$ , 95% CI = 0.56 to 0.92). The same analysis for DMFS gave no results because of convergence issues. PFS showed slightly smaller trial-level correlation in the group of old trials ( $R^2 = 0.81$ , 95% CI = 0.62 to 1.00) than within the recent trials ( $R^2 = 0.92$ , 95% CI = 0.84 to 1.00), but confidence intervals largely overlapped (Table 2). Of note, for both subgroups we were able to fit only nonadjusted regression. Accordingly, the STE was lower for old trials (0.84) than for recent ones (0.91), and both STEs were lower with restrained follow-up: 0.74 and 0.87, respectively. The results for DMFS were similar in the two subgroups.

#### Discussion

The present analysis shows that, for trials investigating the role of chemotherapy in nasopharyngeal carcinoma, PFS and DMFS were strongly correlated with OS, both at the trial and the individual level. Leave-one-out cross-validations confirmed the internal validity of the results. The surrogate threshold effect was 0.89 for PFS and 0.98 for DMFS. This suggests that if in a new trial the confidence interval of the hazard ratio on PFS has an upper limit below 0.89, then the benefit on OS is likely to be statistically significant, whereas for DMFS the upper limit of the confidence interval should be below 0.98 to predict a benefit on OS.

Rigorous validation of plausible surrogate end points is important because of unintended, unanticipated, and unrecognized mechanisms of action (37). The theoretical bases of the statistical validation of surrogate end points have been widely discussed since the publication of the Prentice criteria (38). As previously done for H&N squamous cell carcinoma (3) and in a number of other disease sites (4,27–31), we adopted a correlation approach, which deems a surrogate end point acceptable if both the individual-level and the trial-level correlations  $\rho$  and  $\mathbb{R}^2$  are close to 1. Although only superiority trials were included,

Table 2. Surrogate end points in nonmetastatic nasopharyngeal carcinoma: summary of the results\*

Surrogacy measures		By subgroup of trials		
	All trials, all follow-up	Old trials†	Recent trials‡	Surrogate at 2 y vs OS at 5
Events, no.				
OS	2162	1208	954	1658
PFS	2600	1466	1134	1613
DMFS	2333	1285	1048	1270
PFS vs OS				
Individual level				
ρ (95% CI)	0.93 (0.93 to 0.94)	0.93 (0.92 to 0.94)	0.94 (0.93 to 0.95)	0.89 (0.88 to 0.90)
Trial level				
R <sup>2</sup> (95% CI)	0.95 (0.47 to 1.00)	0.81§ (0.62 to 1.00)	0.92§ (0.84 to 1.00)	0.85 (0.46 to 1.00)
STE	0.89	0.84§	0.91§	0.83
DMFS vs OS				
Individual level				
$\rho$ (95% CI)	0.98 (0.98 to 0.98)	0.99 (0.99 to 0.99)	0.97 (0.97 to 0.98)	0.95 (0.94 to 0.95)
Trial level				
R <sup>2</sup> (95% CI)	0.96§ (0.94 to 0.99)	0.98§ (0.95 to 1.00)	0.97§ (0.94 to 1.00)	0.78 (0.33 to 1.00)
STE	0.98§	0.97§	0.99§	0.90

\*Individual- and trial-level surrogacy for progression-free survival and distant metastasis-free survival vs overall survival in the main analysis (20 trials, 5360 patients) and in sensitivity analyses. CI = confidence interval; DMFS = distant metastasis-free survival; OS = oversall survival; PFS = progression-free survival; STE = surrogate threshold effect.

†Nine trials (12 comparisons, 2484 patients) = PWH-88, AOCOA, VUMCA-89, INT-0099, Japan-91, TCOG-94, PWHQEH-94, QMH-95, VUMCA-95.

‡Eleven trials (12 comparisons, 2876 patients) = SQNP01, NPC-9901, NPC-9902, Guangzhou 2001, NPC008, Guangzhou 2002-01, Guangzhou 2002-02, Guangzhou 2003, HeCOG, Shanghai 2004, Guangzhou 2006.

§Results obtained without adjustment for estimation error because of a lack of convergence of the adjusted model.

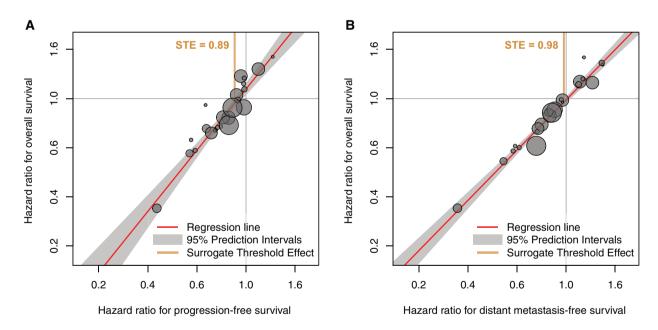
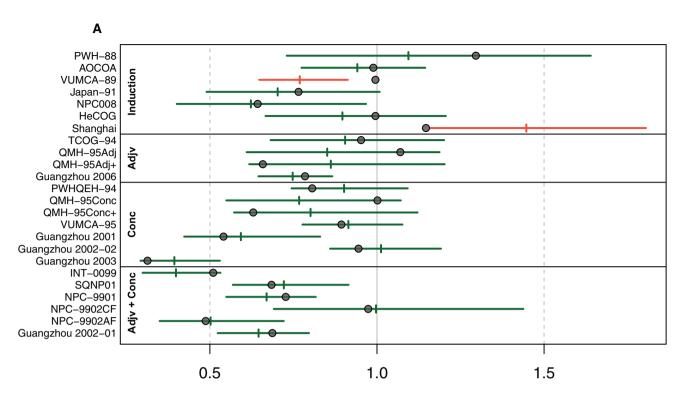


Figure 2. Correlation between treatment effects on the surrogate and overall survival in loco-regionally advanced nasopharyngeal carcinomas. A) Progression-free survival. B) Distant metastasis-free survival. Each circle is a trial, and its size is proportional to the number of patients. STE = surrogate threshold effect.

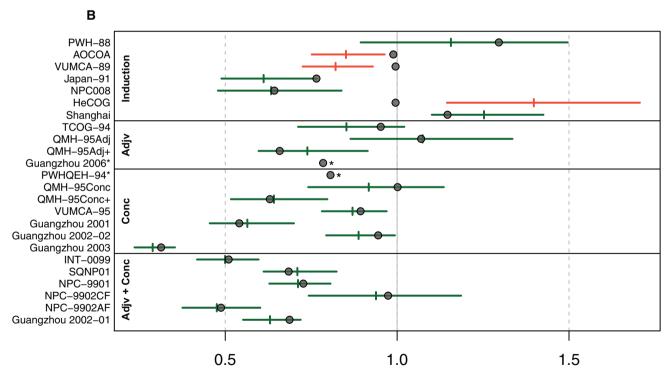
all the results but the surrogate threshold effects are valid also for noninferiority trials.

The clinical usefulness of a surrogate end point depends mostly on its reliability and generalizability to future clinical trials. To be reliable, a surrogate must accurately reflect the final end point. It needs to be validated on high-quality data using a standardized definition and robust statistical methods. We used individual patient data for all trials included in this analysis, we checked the quality of each trial, we recomputed end points consistently using updated follow-up when available,

and we validated surrogacy both at the trial and individual levels. In comparison, a recent analysis used published data only (39) and PFS could only be evaluated in nine trials where the end point was defined consistently. Besides, as always in published data meta-analyses (40), some input data—not clearly mentioned in the text-were estimated based on survival curves. Although this method (41) is frequently used, its limitations are well known (42). This translates into a lower R<sup>2</sup> for PFS than the one presented here, although the STEs are very close. Furthermore, in that paper DMFS could not be evaluated



## Hazard ratio for overall survival



## Hazard ratio for overall survival

Figure 3. Leave-one-out cross-validation analysis for the model predicting the treatment effect on overall survival based on progression-free survival (A) and distant metastasis-free survival (B) effects. Circles are the observed hazard ratios for the effect on overall survival. Horizontal segments correspond to 95% prediction intervals (PI). The vertical segments are the predicted effects on overall survival using the observed hazard ratio on progression or distant metastasis-free survival of each trial and using the surrogate model fitted on the other trials. \*The surrogate model could not be fitted due to convergence issues. Individual trials are grouped according to the timing of chemotherapy and are named the same way as in our initial publication(2), while references can be found in the Supplementary Table 1. Adjv = adjuvant; Conc = Concomitant.

because of the lack of information on distant progressions in the articles, and no patient-level analysis could be performed for any surrogate end point. To conclude, while a published data meta-analysis on surrogacy is certainly a first step, the surrogacy should be confirmed by an individual patient data analysis.

In the present analysis, two surrogate end points were evaluated, and the question arises as to whether one is preferable. While the correlation coefficients for DMFS are higher than for PFS when considering unrestrained follow-up, this statistical superiority of DMFS is less obvious when considering a "real life" clinical trial situation, which is the prediction of fiveyear OS knowing the efficacy on the two-year surrogate. The trial-level correlation, the one that matters in this situation, is superior for PFS. In general, PFS is measured earlier and provides a higher power, as more events are taken into account, than DMFS. Furthermore, the surrogate-to-OS ratio of median times is lower for PFS, implying that the trial analysis would be more accelerated by using PFS than DMFS. Nevertheless, one can expect that such difference will be smaller in new trials, in which mandatory 3D-conformal/IMRT techniques will likely make DMFS and PFS more closely associated. Lastly, PFS was slightly more robust in cross-validation, with a wrong prediction of treatment effect on OS based on the effect on PFS for two trials, as compared with three for DMFS. Overall, we believe that the use of PFS as a surrogate is more reliable and should be encouraged for future trials evaluating chemotherapy in NPC.

The major limitations of the current study relate to its relevance with regards to the future of clinical research in NPC. Modern imaging techniques could allow the detection of recurrences earlier and the triggering of potentially curative salvage treatments. Then the treatment effect on the surrogate end point in a new trial could be diluted and potentially not detectable on OS. However, at present, salvage loco-regional treatments are seldom curative and can be applied to a highly selective patient population (43). Systemic treatment remains palliative in this setting, as shown by the OS durations in the recurrent and metastatic setting (43). The validity of PFS or DMFS remains strong in our data despite this limitation. Another important issue is the validity in the context of new systemic treatments, although most recent trials in NPC patients still investigate cytotoxic chemotherapies: A trial search performed in the summer of 2015 found that 10 out of 12 identified ongoing or recently completed trials investigated cytotoxic agents, while only two investigated nimotuzumab, a monoclonal antibody directed against the epidermal growth factor. Furthermore, and as mentioned earlier, even if a targeted agent is used for primary treatment in combination with radiotherapy, the relationship between progression and death might unfortunately remain valid because of the absence of broadly applicable effective salvage therapies. When cure is the goal, the switch from "progression-free" to "recurrent/metastatic" will likely remain a major determinant of survival in the context of targeted therapies or immunotherapies should be investigated, especially in future trials that incorporate maintenance systemic therapy. Last, even if an effect is demonstrated on the early surrogate end point, patients should still be followed up, as early end points will not capture long-term treatment-related toxicities and deaths, which are of major relevance in this disease given its high

In conclusion, this individual patient data validation of surrogate end points in localized nasopharyngeal carcinoma treated with radiotherapy and chemotherapy demonstrates that PFS and DMFS are strong surrogates for OS. Two-year PFS could be used as the primary end point in the design of future randomized trials to speed up the research and lower the associated costs.

Whatever the surrogate end point chosen, OS should still be measured and reported at a meaningful time point, such as five years, to detect possible long-term detrimental effects.

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