## Smoothing

Lessons for the master course in Applied Statistics

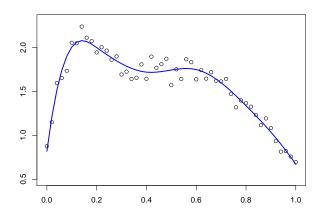
## May 2023

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▶ Data:  $\{(Y_1, x_1), \dots, (Y_n, x_n)\}$ 

#### Model

$$Y = f(x) + \epsilon$$

 $f \in \mathcal{F}$  where  $\mathcal{F}$  is the appropriate functional space

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  $\epsilon_i \sim \text{indep}$   $\mathbb{E}[\epsilon_i] = 0, Var[\epsilon_i] = \sigma^2 < 0$ 

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#### Smoothing problem:

$$\hat{f} = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \operatorname{RSS}(f) = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \left\{ \sum_{i=1}^{n} (Y_i - f(X_i))^2 \right\}$$



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ill posed!



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1st approach: restrict search to  $\mathcal{F}_K$ , with  $dim(\mathcal{F}_K) = K << n$ 



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2nd approach: do not restrict  $\mathcal{F}$  but add a roughness penalty



# SMOOTHING Estimation functional

1st approach:  $\mathcal{F}$  approximated by  $\mathcal{F}_K$ , with  $dim(\mathcal{F}_K) = K << n$ 

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 $\blacktriangleright \psi_1, \ldots, \psi_K$ : K basis functions s.t.  $\mathcal{F}_K = \operatorname{span}(\psi_1, \ldots, \psi_K)$ 

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$$Y_i = f(x_i) + \epsilon_i \qquad \leadsto \qquad Y_i = \sum_{i=1}^K c_i \psi_i(x_i) + \epsilon_i$$



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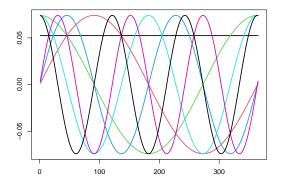
$$= \underset{\boldsymbol{c} \in \mathbb{R}^K}{\operatorname{argmin}} \operatorname{RSS}(\boldsymbol{c}) = \underset{\boldsymbol{c} \in \mathbb{R}^K}{\operatorname{argmin}} \sum_{i=1}^{n} (Y_i - \sum_{j=1}^{K} c_j \psi_j(x_i))^2$$

 $\blacktriangleright \psi_1, \dots, \psi_K$ : K basis functions s.t.  $\mathcal{F}_K = \operatorname{span}(\psi_1, \dots, \psi_K)$ Model

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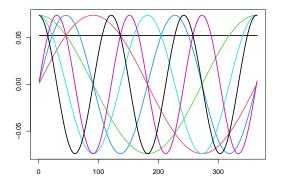
## **FOURIER**



Fourier basis

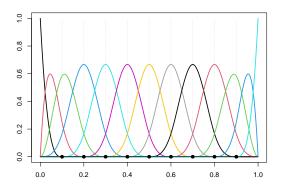


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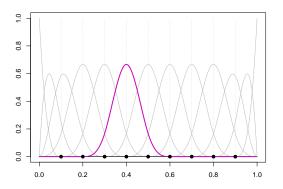
Periodic, localized in frequency, Convenient for differentiation and integration, Computationally efficient (orthogonal)





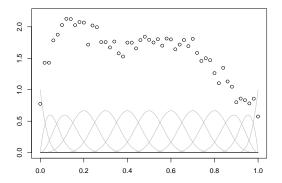
Cubic B-spline basis





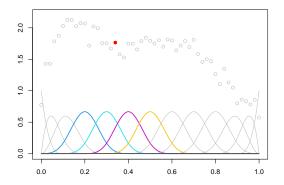
Flexible, localized in space, convenient for differentiation and integration





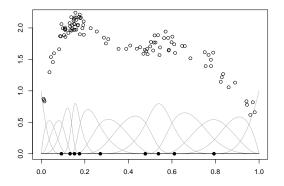
Computationally efficient (local support, band limited structure of key matrices involved, etc)





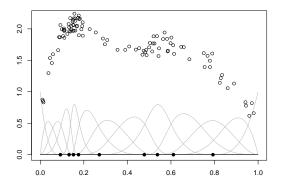
Computationally efficient (local support, band limited structure of key matrices involved, etc)





Knots can be placed along the percentiles of X





Knots replication permits discontinuity in derivatives or function itself



$$\Psi = \begin{bmatrix} \psi_{1}(X_{1}) & \psi_{2}(X_{1}) & \cdots & \psi_{K}(X_{1}) \\ \psi_{1}(X_{2}) & \psi_{2}(X_{2}) & \cdots & \psi_{K}(X_{2}) \\ \vdots & \vdots & & \vdots \\ \psi_{1}(X_{n}) & \psi_{2}(X_{n}) & \cdots & \psi_{K}(X_{n}) \end{bmatrix}$$



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Model

$$Y = f + \epsilon = \Psi c + \epsilon$$



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$$\mathbf{Y} = \mathbf{f} + \mathbf{\epsilon} = \Psi \mathbf{c} + \mathbf{\epsilon}$$
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$$\hat{f} = \boldsymbol{\psi}^{\top} \hat{\boldsymbol{c}}$$



#### SMOOTHING Expression of the estimator

$$\hat{\mathbf{c}} = \underset{\mathbf{c} \in \mathbb{R}^K}{\operatorname{argmin}} \ \mathsf{RSS}(\mathbf{c}) = \underset{\mathbf{c} \in \mathbb{R}^K}{\operatorname{argmin}} \ \left\{ (\mathbf{Y} - \Psi \mathbf{c})^\top (\mathbf{Y} - \Psi \mathbf{c}) \right\}$$



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: projection matrix (properties:  $S^{\top} S = S$ )

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$$\hat{\sigma}^2 = \frac{1}{n - K} (\mathbf{Y} - \hat{\mathbf{Y}})^{\top} (\mathbf{Y} - \hat{\mathbf{Y}}) = \frac{1}{n - tr(S)} (\mathbf{Y} - \hat{\mathbf{Y}})^{\top} (\mathbf{Y} - \hat{\mathbf{Y}})$$



#### Properties of the estimator

$$\hat{f}(x) = \psi(x)^{\top} \hat{\mathbf{c}} = \psi(x)^{\top} (\Psi^{\top} \Psi)^{-1} \Psi^{\top} \mathbf{Y}$$

$$\hat{f}'(x) = \psi'(x)^{\top} \hat{\mathbf{c}} \qquad \text{where } \psi' = (\psi'_1, \dots, \psi'_k)^{\top}$$

$$\hat{f}''(x) = \psi''(x)^{\top} \hat{\mathbf{c}} \qquad \text{where } \psi'' = (\psi''_1, \dots, \psi''_k)^{\top}$$

Smoothing requires special care when the curve estimate is asked, not only to provide a good smoothing of the data, but also to reflect the differential features

Curve derivatives (or their functions) are very often:

- objects of analysis
- helpful for further processing and data analysis (curve alignment/clustering)



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Source of bias: discretization

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$$\operatorname{Bias}[\hat{f}(X)] = f(X) - \mathbb{E}[\hat{f}(X)] = f(X) - \psi(X)^{\top} \mathbf{c}$$

Source of bias: discretization

$$\operatorname{Var}[\hat{f}(x)] = \mathbb{E}[\{\hat{f}(x) - \mathbb{E}[\hat{f}(x)]\}^2] = \sigma^2 \psi(x)^\top (\Psi^\top \Psi)^{-1} \psi(x)$$



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## SMOOTHING Bigs-Variance trade-off

Number of bases K controls Bias-Variance trade-off.



Bias-Variance trade-off

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K can be selected by AiC,  $C_p$  Mallows, cross-validation, Generalized Cross Validation:

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K can be selected by AiC,  $C_p$  Mallows, cross-validation, Generalized Cross Validation:

$$GCV(K) = \frac{n}{(n-K)} \frac{1}{(n-K)} (\mathbf{Y} - \hat{\mathbf{Y}})^{\top} (\mathbf{Y} - \hat{\mathbf{Y}})$$
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## SMOOTHING Bigs-Variance trade-off

Number of bases K controls Bias-Variance trade-off.

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$$= \frac{n}{(n-tr(S))^{2}} (\mathbf{Y} - \hat{\mathbf{Y}})^{\top} (\mathbf{Y} - \hat{\mathbf{Y}})$$

Hopefully the chosen value of K is close to that minimizing

$$\mathrm{MSE}[\hat{f}(x)] = \mathbb{E}[\{\hat{f}(x) - f(x)\}^2] = \mathrm{Bias}^2[\hat{f}(x)] + \mathrm{Var}[\hat{f}(x)]$$



Under regularity conditions, as  $n \to \infty$  and  $K(n) \to \infty$  with appropriate rates

$$\operatorname{Bias}[\hat{f}(x)] \to 0 \qquad \text{and} \qquad \operatorname{Var}[\hat{f}(x)] \to 0$$

$$\mathrm{MSE}[\hat{f}(x)] \to 0$$

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Limiting Gaussian distrib justifies Wald type inference on f(x):



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▶ C.I. of approx level  $(1-\alpha)$  on f(x):  $\hat{f}(x) \pm z_{1-\alpha/2} \sqrt{\operatorname{Var}[\hat{f}(x)]}$ 

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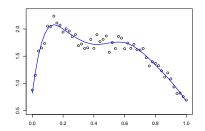
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- ▶ C.I. of approx level  $(1-\alpha)$  on f(x):  $\hat{f}(x) \pm z_{1-\alpha/2} \sqrt{\operatorname{Var}[\hat{f}(x)]}$
- ► Test on  $H_0: f(x) = f_0(x)$  vs  $H_1: f(x) \neq f_0(x)$  of approx level  $\alpha$ Reject  $H_0$  if  $|\hat{f}(x) - f_0(x)| > z_{1-\alpha/2} \sqrt{\widehat{\operatorname{Var}[\hat{f}(x)]}}$



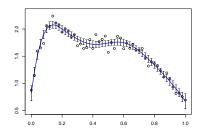
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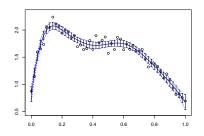
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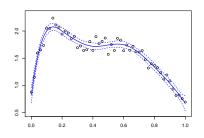


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Even if you join the various C.l.s by bands, do not forget these are C.l.s for f(x) in a specific x.

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You cannot interpret the smooth dashed bands as delimiters of a region that includes the true overall f with a given confidence.

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Various alternative approaches, including bootstrap, as well as undersmoothing or oversmoothing approaches (see, e.g., review in Hall and Horowitz 2013)



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2nd approach: Do not restrict  ${\mathcal F}$  but minimize RSS + roughness penalty

$$\hat{f} = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \ \operatorname{RSS}_{\lambda}(f) = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \ \left\{ \sum_{i=1}^{n} (Y_i - f(x_i))^2 + \lambda \mathcal{P}(f) \right\} \qquad \lambda > 0$$

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THM: If (a, b) s.t.  $a < x_{[1]} < x_{[2]} < \ldots < x_{[n]} < b$ ,  $\mathcal{F} = H^2(a, b)$ , and  $\mathcal{P}(f) = \int_a^b \big(f''(x)\big)^2 dx$  then  $\hat{f}$  is a natural cubic splines over (a, b) with knots at  $x_{[1]}, x_{[2]}, \ldots, x_{[n]}$ .



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$$f = \sum_{i=1}^{K} c_{i} \psi_{i}$$

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$$P \in \mathbb{R}^K \times \mathbb{R}^K$$

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2nd approach:  $\hat{f} \in \mathcal{F}_K$ , with  $dim(\mathcal{F}_K) = K \approx n$  minimize RSS + roughness penalty

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## PENALIZED SMOOTHING

For large datasets, it may be convenient to employ a mixed approach, where one considers a roughness penalty, but also reduces the dimensionality of the data estimation problem by considering a basis of dimension K, where K is still large but somehow smaller than n.



Expression of the estimator

$$\hat{\mathbf{c}} = (\boldsymbol{\Psi}^{\top}\boldsymbol{\Psi} + \lambda P)^{-1}\boldsymbol{\Psi}^{\top}\mathbf{Y}$$



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$$or eq df = tr(S) < K$$
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$$\hat{\sigma}^{2} = \frac{1}{D - tr(S)} (\mathbf{Y} - \hat{\mathbf{Y}})^{\top} (\mathbf{Y} - \hat{\mathbf{Y}})$$



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Physics-informed penantly:  $\mathcal{P}(f) = \int (Lf - u)^2$ Lf = u encodes available problem-specific information

