

# **Exchangeable Random Graphs**

**Econometric Methods for Networks,**

**CORE, Dec 12th - 14th, 2016**

*Bryan S. Graham*

University of California - Berkeley

## Introduction

- First of two lectures on network nonparametrics
- Rest of today:
  - Aldous-Hoover representation
    - \* Orbanz and Roy (2015)
  - Nearest neighborhood smoothing for edge probability estimation
    - \* Zhang, Levina and Zhu (2015)

## Introduction (continued)

- Tomorrow:
  - graph limits (e.g., Lovász, 2012)
  - estimation of network moments
    - \* Holland and Leinhardt (1976)
    - \* Bickel, Chen and Levina (2011)
    - \* Bhattacharya and Bickel (2015)

## Setup

Let  $G(\mathcal{V}, \mathcal{E})$  be a finite undirected random graph with

- agents/vertices  $\mathcal{V} = \{1, \dots, N\}$ ,
- links/edges  $\mathcal{E} = \{\{i, j\}, \{k, l\}, \dots\}$ , and
- adjacency matrix  $\mathbf{D} = [D_{ij}]$  with

$$D_{ij} = \begin{cases} 1 & \text{if } \{i, j\} \in \mathcal{E} \\ 0 & \text{otherwise} \end{cases}$$

## Setup (continued)

- The expected adjacency matrix equals

$$\mathbf{P} = [P_{ij}] = [\mathbb{E} [D_{ij} | U_1, \dots, U_2, \alpha]]$$

for  $i < j$ .

- Here  $\{U_i\}_{i=1}^N$  and  $\alpha$  are *latent* random variables introduced (and explained below).
- Form of  $\mathbf{P}$  might indicate community structure...
- ...or guide other aspects of model formulation

## Exchangeable Networks

- Let  $\pi$  be a permutation of the index set  $\{1, \dots, N\}$ .

- In many situations it is natural to assume that

$$[D_{ij}] \stackrel{d}{=} [D_{\pi(i)\pi(j)}] \quad (1)$$

for every permutation  $\pi$  and  $i < j$ ,  $j = 1, \dots, N$ .

–  $\stackrel{d}{=}$  indicates equality of distribution.

- Condition (1)  $\Rightarrow$  our beliefs about the probability of a link between two agents does not depend on their labels.
- Networks with this property are *jointly exchangeable*.

## Exchangeable Networks (continued)

- Does exchangeability have any modeling implications?
- Does  $\mathbf{D}$  converge to a *graph limit* as  $N \rightarrow \infty$ ?
- Dense graph implication:
  - if  $[D_{ij}] \stackrel{d}{=} [D_{\pi(i)\pi(j)}]$  then  $\rho = \Pr(D_{ij} = 1)$  is either bounded away from zero or zero.
  - exchangeable graphs are either dense or empty!

## Exchangeable Sequences

- The sequence  $Y_1, Y_2, \dots$  is said to be **infinitely exchangeable** if, for every  $N \geq 2$  and permutation  $\pi$ ,

$$(Y_1, Y_2, \dots, Y_N) \stackrel{d}{=} (Y_{\pi(1)}, Y_{\pi(2)}, \dots, Y_{\pi(N)}) .$$

- i.i.d. sequences are exchangeable...
- ... but non i.i.d. sequences can be too:

$$Z + Y_1, Z + Y_2, \dots$$

for  $Z$  some non-trivial random variable, drawn independently of the i.i.d. sequence  $Y_1, Y_2, \dots$



## de Finetti Theorem

- de Finetti (1931): the sequence of binary random variable  $Y_1, Y_2, \dots$  is infinitely exchangeable if, and only if,

$$\Pr(Y_1 = y_1, Y_2 = y_2, \dots, Y_N = y_N) = \int_0^1 \alpha^{t_N} (1 - \alpha)^{N - t_N} d\Pi(\alpha)$$

for  $t_N = \sum_{i=1}^N y_i$ , all  $N \geq 2$ , and  $\Pi$  some measure on  $\alpha \in [0, 1]$ .

- For any infinitely exchangeable sequence we have that – conditional on the random variable  $\alpha$ –

$$\Pr(Y_1 = y_1, Y_2 = y_2, \dots, Y_N = y_N | \alpha) = F_\alpha(y_1) F_\alpha(y_2) \times \dots \times F_\alpha(y_N)$$

for  $F_\alpha(y) = \alpha^y (1 - \alpha)^{1-y}$  if  $y \in \{0, 1\}$  and zero otherwise.

## de Finetti Theorem (continued)

- Representation result: any exchangeable binary sequence can be modeled ‘as if’ the DGP were:
  1. Draw  $\alpha \sim \Pi$
  2. Draw  $Y_i \sim F_\alpha$  for  $i = 1, \dots, N$
- *Conditional* on  $\alpha$ ,  $Y_1, Y_2, \dots$  is an i.i.d. sequence, where each of its members have the same *random* distribution function  $F_\alpha(y)$ .
- See Orbanz and Roy (2015) for non-technical survey of de Finetti type results

## Alternative Formulation

- The right-continuous inverse of  $F_\alpha(u)$  (i.e., quantile function) is

$$g_\alpha(u) \stackrel{\text{def}}{=} F_\alpha^{-1}(u) = \begin{cases} 0 & \text{if } 0 < u < 1 - \alpha \\ 1 & \text{if } 1 - \alpha \leq u < 1 \end{cases}.$$

- This gives:

$$(Y_1, Y_2, \dots) \stackrel{d}{=} (g_\alpha(U_1), g_\alpha(U_2), \dots)$$

for  $\{U_i\}_{i=1}^\infty$  a sequence of independent  $\mathcal{U}[0, 1]$  random variables.

- We further have that

$$\mathbb{E}[Y_i | U_i = u] = g_\alpha(u).$$

## Alternative Formulation

- The “*sequon*” (sequence function)  $g_\alpha(u)$  is not identifiable...

– consider  $g_\alpha(u)$  above with:

$$g_\alpha^*(u) = \begin{cases} 0 & \text{if } 0 < u < \frac{1-\alpha}{2} \\ 1 & \text{if } \frac{1-\alpha}{2} \leq u < \frac{1}{2} \\ 0 & \text{if } \frac{1}{2} \leq u < \frac{2-\alpha}{2} \\ 1 & \text{if } \frac{2-\alpha}{2} \leq u < 1 \end{cases}.$$

- ...but “moments” are identifiable:

$$- \frac{1}{N} \sum Y_i \xrightarrow{p} \mathbb{E}[g_\alpha(U) | \alpha] = \alpha$$

## Aldous-Hoover

- Aldous (1981) and Hoover (1979) (essentially) showed that a random graph is jointly exchangeable if, and only if, it admits the representation

$$[D_{ij}] \stackrel{d}{=} [g_{\alpha}(U_i, U_j, V_{ij})]$$

for  $\{U_i\}_{i=1}^{\infty}$  and  $\{V_{ij}\}_{i < j}$  sequences of independent  $\mathcal{U}[0, 1]$  random variables.

- Here  $\alpha$  is a mixing parameter as in de Finetti (1931).
  - $g_{\alpha}(\cdot, \cdot, \cdot)$  is a random function

## Aldous-Hoover (continued)

- Averaging over  $V_{ij}$  yields

$$\begin{aligned} h_\alpha(u_i, u_j) &= \mathbb{E} \left[ D_{ij} \mid U_i = u_i, U_j = u_j, \alpha \right] \\ &= \mathbb{E} \left[ g_\alpha(u_i, u_j, V_{ij}) \mid \alpha \right] \\ &= \int_0^1 g_\alpha(u_i, u_j, v) \, dv \end{aligned}$$

from which we get the more convenient representation, for  $i < j$ ,

$$[D_{ij}] \stackrel{d}{=} [\mathbf{1}(V_{ij} \leq h_\alpha(U_i, U_j))]$$

- $h_\alpha(U_i, U_j)$  is a *graphon*: short for **graph function**.

## Aldous-Hoover (continued)

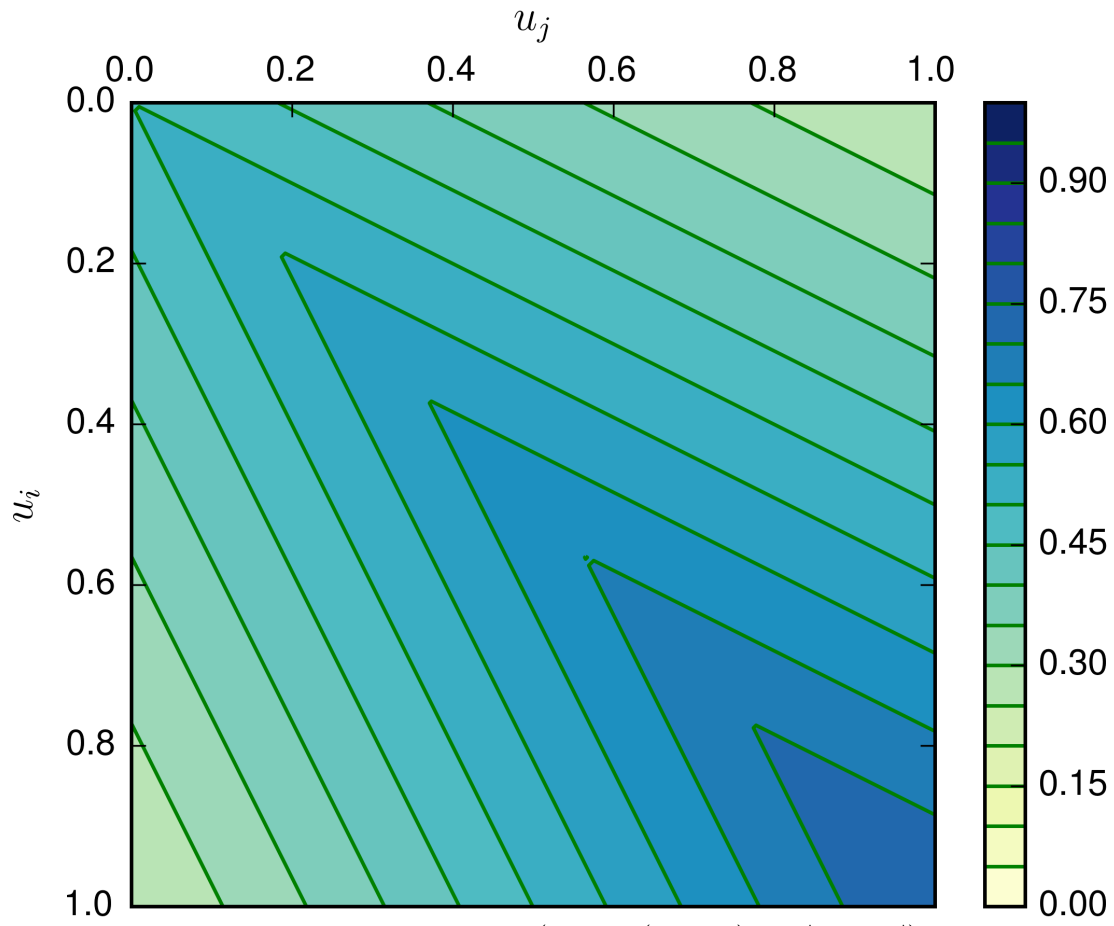
- The Aldous-Hoover representation theorem implies that we can proceed ‘as if’ links formed independently conditional on the agent-specific latent variables  $\{U_i\}_{i=1}^{\infty}$  and  $\alpha$ .
- A network generating process is:
  1. “Draw”  $\alpha$  or choose a graphon;
  2. Draw  $U_i \sim \mathcal{U}[0, 1]$  for agents  $i = 1, \dots, N$ ;
  3. Construct  $\mathbf{D}$ , by sampling  $D_{ij} \mid h_{\alpha}(\bullet, \bullet), U_i, U_j \sim \text{Bernoulli}\left(h_{\alpha}(U_i, U_j)\right)$  for every dyad  $\{i, j\}$  with  $i < j$ .

## Aldous-Hoover (continued)

- *any* exchangeable random graph may be modeled as a mixture conditionally independent edge formation processes
- Conditional independence structure useful for large sample theory
- Representation result: actual network generating process may not coincide with representation (cf., reduced form)



Graphon contour plot



Note:  $h(u_i, u_j) = \frac{\exp(\alpha_0 + \alpha_1(u_i + u_j) + \alpha_2|u_i - u_j|)}{1 + \exp(\alpha_0 + \alpha_1(u_i + u_j) + \alpha_2|u_i - u_j|)}$

## Graphon

- The graphon  $h_\alpha(u, v)$  is not identifiable...
  - consider the m.p.t.  $\varphi(U) = 1 - U$  or  $\varphi(U) = 2U \bmod 1$
  - $g_\alpha(U_i, U_j, V_{ij})$  and  $g_\alpha(\varphi(U_i), \varphi(U_j), V_{ij})$  generate graphs with the same properties
- ...but link/edge probabilities *are* identifiable (under assumptions).
  - $p_{ij} = \mathbb{E} [D_{ij} | \mathbf{U}] = h_\alpha(U_i, U_j)$

## Graphon (Bickel & Chen, 2009)

- For statistical analysis it is convenient to formulate the graphon somewhat differently.
- Consider the network DGP

$$\Pr(D_{ij} = 1 | U_i, U_j, \alpha) = h_\alpha(U_i, U_j)$$

and define

$$\rho_\alpha = \int_0^1 \int_0^1 h_\alpha(u, v) \, du \, dv$$
$$w_\alpha(u, v) = f_{U_i, U_j | D_{ij}, \alpha}(u, v | D_{ij} = 1, \alpha).$$

- Since  $f_{U_i, U_j | \alpha}(u, v | \alpha) = 1$  on  $[0, 1]^2$  we get the formulation

$$h_\alpha(u, v) = \rho_\alpha w_\alpha(u, v).$$

## Graphon (Bickel & Chen, 2009)

- The Bickel and Chen (2009) formulation is useful for sequences of network GPs where  $\rho_\alpha$ , the network density, is indexed by  $N$ .
  - i.e.,  $\rho_{\alpha,N} \rightarrow 0$  as  $N \rightarrow \infty$
  - in practice we ignore any dependence of  $w_\alpha(u, v)$  on  $N$
- The rate at which  $\rho_{\alpha,N} \rightarrow 0$  controls the sparsity links
- If  $\lambda_N = (N - 1) \rho_{\alpha,N} \rightarrow \lambda > 0$  as  $N \rightarrow \infty$  the graph is *sparse*
  - other cases:  $\lambda_N = O(N)$  (*dense*) or  $\lambda_N = O(\ln N)$  (*semi-dense*)

## Edge Probability Estimation

- Define the inner product

$$\langle f, g \rangle = \int f(u) g(u) du$$

with the associated norm

$$\|f\| = \langle f, f \rangle^{1/2} = \left[ \int f(u)^2 du \right]^{1/2}.$$

- Linking behavior of an agent of type  $u$  is summarized by the *graphon slice*  $\rho w(u, \bullet)$ .
- Measure “distance” between agent  $i$ , with  $U_i = u$ , and agent  $j$ , with  $U_j = v$ , by:

$$\begin{aligned} d(u, v) &= \|\rho w(u, \cdot) - \rho w(v, \cdot)\|_2 \quad (2) \\ &= \rho \left[ \int [w(u, t) - w(v, t)]^2 dt \right]^{1/2} \end{aligned}$$

## Network Neighbors

- $\mathbf{P} = \mathbb{E} [\mathbf{D} | \mathbf{U}]$  denotes the expected adjacency matrix.
- $\mathbf{P}_{i\bullet}$  denotes the  $i^{th}$  row of this matrix
- Distance between  $i$  and  $j$  is

$$\begin{aligned} d_N(i, j) &= \|\mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}\|_2 \\ &= \left[ \frac{1}{N-2} \sum_{k \neq i, j} (P_{ik} - P_{jk})^2 \right]^{1/2} \end{aligned} \quad (3)$$

- $i^{th}$  and  $j^{th}$  elements of both  $\mathbf{P}_{i\bullet}$  and  $\mathbf{P}_{j\bullet}$  are removed prior to calculating  $d(i, j)$ .

## Nearest Network Neighbors

- *j* is an exact neighbor of *i* if  $d_N(i, j) = 0$ 
  - *i* and *j* have identical (expected) adjacency (matrix) slices.
  - i.e., identical *ex ante* linking behavior
  - *realized* links may differ

## Nearest Neighbor Averaging

- In a finite network it may be that agent  $i$  has no exact neighbors, but we can still find a set of *nearest neighbors*:

$$\mathcal{N}_i = \{j : \|\mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}\|_2 \leq q_i(h_N)\} \quad (4)$$

where  $q_i(h_N)$  is the  $h_N^{th}$  sample quantile of  $\{\|\mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}\|_2\}_{j=1, j \neq i}^N$ .

- If  $N = 1,000$  and  $h_N = 0.05$ , then we would take the 50 nearest neighbors.
- Estimate  $P_{ij}$  by the local average

$$\hat{P}_{ij}^{\text{oracle}} = \frac{1}{2} \left( \underbrace{\frac{\sum_{k \in \mathcal{N}_i} D_{kj}}{|\mathcal{N}_i|}} + \frac{\sum_{l \in \mathcal{N}_j} D_{il}}{|\mathcal{N}_j|} \right). \quad (5)$$

- Unfortunately  $\mathbf{P}$  is not observed!



## Finding Network Neighbors

- Can we construct a measure of distance between two agents based on the (observed) adjacency matrix alone?

- Zhang et al. (2015) observe that

$$\begin{aligned} d_N^2(i, j) &= \|\mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}\|_2^2 \\ &= \langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet} \rangle \\ &= \langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} \rangle - \langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{j\bullet} \rangle \\ &\leq |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} \rangle| + |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{j\bullet} \rangle| \end{aligned}$$

- Need estimates of  $\langle \mathbf{P}_{i\bullet}, \mathbf{P}_{i\bullet} \rangle$ ,  $\langle \mathbf{P}_{j\bullet}, \mathbf{P}_{j\bullet} \rangle$  and  $\langle \mathbf{P}_{i\bullet}, \mathbf{P}_{j\bullet} \rangle$  to form estimate of  $d_N(i, j)$ .

## Finding Neighbors (continued)

- $\langle \mathbf{P}_{i\bullet}, \mathbf{P}_{i\bullet} \rangle = \frac{1}{N-1} \sum_{i \neq j} P_{ij}^2$  is hard to estimate...
- ...apparently requires estimate of  $P_{ij}$  (which is our target!)
- However the (limit of the) term

$$\langle \mathbf{P}_{i\bullet}, \mathbf{P}_{j\bullet} \rangle = \frac{1}{N-2} \sum_{k \neq i, j} P_{ik} P_{jk}$$

is not hard to estimate since

$$\mathbb{E} \left[ \frac{1}{N-2} \sum_k D_{ik} D_{jk} \right] = \mathbb{E} \left[ \frac{1}{N-2} \sum_{k \neq i, j} P_{ik} P_{jk} \right].$$

- Recall edges form independently conditional on  $\mathbf{U}$ .

## Finding Neighbors (continued)

- Assume that  $w(u, v)$  is Lipschitz continuous:

$$\rho \|w(u, \cdot) - w(v, \cdot)\|_2 \leq C \|u - v\|_2.$$

- With  $N$  large we can find an agent  $l \neq i, j$  such that  $|U_j - U_l| \leq \epsilon_N$  for  $\epsilon_N = o(1)$ .

- We get

$$\begin{aligned} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} \rangle| &= |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{k\bullet} \rangle| \\ &\quad + |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} - \mathbf{P}_{k\bullet} \rangle| \\ \text{(TI)} &\leq |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{k\bullet} \rangle| \\ &\quad + |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{i\bullet} - \mathbf{P}_{k\bullet} \rangle| \\ \text{(CS)} &\leq |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{k\bullet} \rangle| \\ &\quad + \|\mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}\|_2 \|\mathbf{P}_{i\bullet} - \mathbf{P}_{k\bullet}\|_2 \\ &\leq |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{k\bullet} \rangle| + C_{i,j} \epsilon_N \end{aligned}$$

## Finding Neighbors (continued)

- Combining results we have that

$$d^2(i, j) \leq 2 \max_{l \neq i, j} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle| + 2C_{i,j} \epsilon_N$$

- ...if  $2 \max_{l \neq i, j} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle| \approx 0$ , then  $d^2(i, j) \approx 0$  if  $N$  is large.

- Zhang et al. (2015) estimate

$$2 \max_{l \neq i, j} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle|$$

by  $\hat{d}^2(i, j)$  equal to

$$2 \max_{l \neq i, j} \left| \frac{1}{N-2} \sum_{k \neq i, j} D_{ik} D_{lk} - \sum_{k \neq i, j} D_{jk} D_{lk} \right|$$

- Estimated neighborhood* of agent  $i$  is then

$$\hat{\mathcal{N}}_i = \{j : \hat{d}^2(i, j) \leq q_i(h_N)\}.$$

## Zhang et al. (2015) Estimate

- Estimate  $P_{ij}$  by

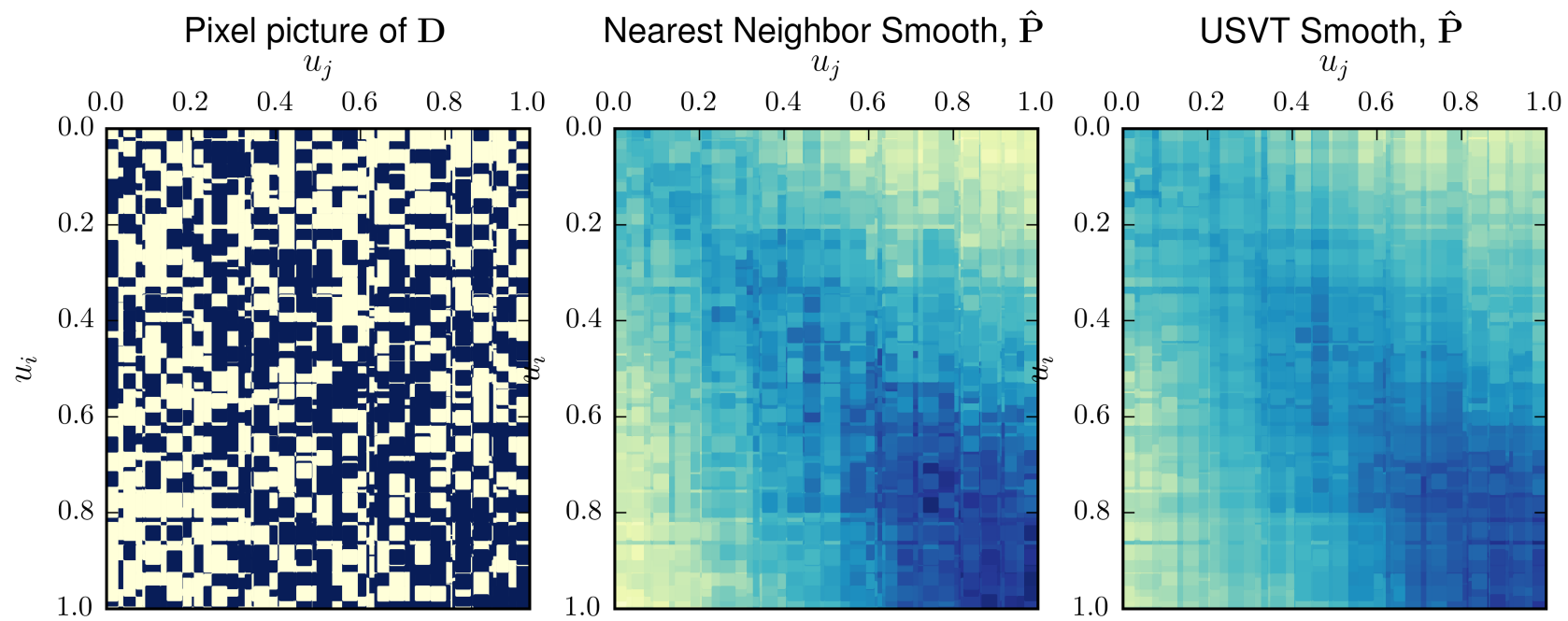
$$\hat{P}_{ij} = \frac{1}{2} \left( \frac{\sum_{k \in \hat{\mathcal{N}}_i} D_{kj}}{|\hat{\mathcal{N}}_i|} + \frac{\sum_{l \in \hat{\mathcal{N}}_j} D_{il}}{|\hat{\mathcal{N}}_j|} \right)$$

- Consistency requires that  $h_N = C\sqrt{\frac{\ln N}{N}}$  for some  $C$ .
- Zhang et al. (2015) suggest that  $C = 0.1$  works well in practice.
  - $K_N = \lfloor Nh_N \rfloor = \lfloor 0.1 (N \ln N)^{1/2} \rfloor$  or  $K_{1000} \approx 8$  and  $K_{2000} \approx 12$ .

## Alternative Distance Measure

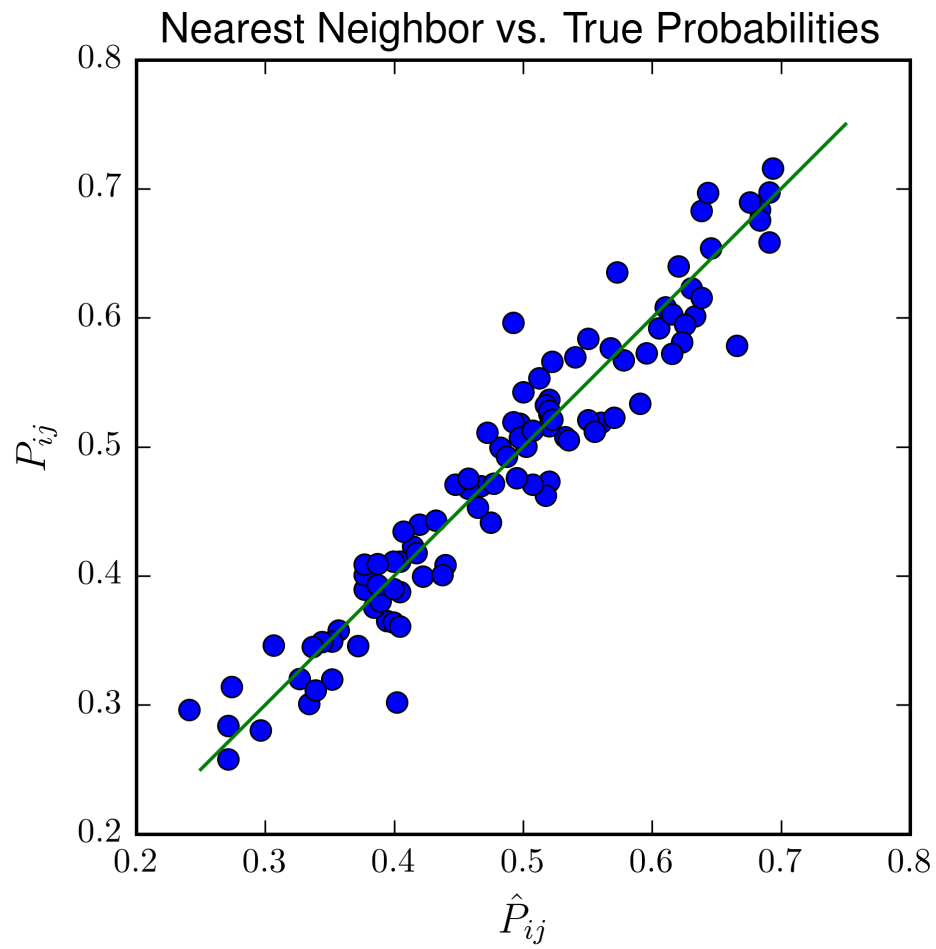
$$\begin{aligned}
 & \max_{l \neq i, j} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle| \\
 & \leq \sum_{l \neq i, j} |\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle| \cdot 1 \\
 \text{(HI)} & \leq \left[ \sum_{l \neq i, j} (\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle)^2 \right]^{1/2} \cdot \left[ \sum_{k \neq i, j} 1^2 \right]^{1/2} \\
 & = \left[ (N-2) \sum_{k \neq i, j} (\langle \mathbf{P}_{i\bullet} - \mathbf{P}_{j\bullet}, \mathbf{P}_{l\bullet} \rangle)^2 \right]^{1/2} \\
 & = \left[ \frac{1}{N-2} \sum_{l \neq i, j} \left( \sum_{k \neq i, l} P_{ik} P_{kl} - \sum_{k \neq j, l} P_{jk} P_{kl} \right)^2 \right]^{1/2} \\
 & = d_N^*(i, j)
 \end{aligned}$$

We can use the ‘smoother’  $\hat{d}_N^*(i, j)$  to find nearest neighbors instead.



## Goodness-of-Fit

( $N = 2,000$ ,  $h_N = 0.1$ )





## Practicalities

- In example it is natural to order  $i$  by their realized values of  $U_i$
- This information is not available in real world examples
- In practice, we can order agents by degree or its smoothed estimate  $\sum_j \hat{P}_{ij}$ 
  - should be sufficient to ‘see’ a block structure (for example) in many cases