Graph Limits & Subgraph Counts

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Introduction

- In 1970 Paul Holland and Samuel Leinhardt (1970, AJS) introduced the triad census.
 - counts of all 4 (16) unique triad isomorphisms in an undirected (directed) graph
 - can construct transitivity index (TI) from triad census...
 - ...as well as the mean and variance of the degree sequence
- Holland and Leinhardt (1976, SM) provide variance expressions for these counts

Introduction (continued)

- In early work normality of these counts was assumed (w/o proof)
- Nowicki (1989, 1991) showed asymptotic normality of counts for homogenous random graphs
- Bickel, Chen & Levina (2011, AS) demonstrated asymptotic normality in the "general" case

Introduction (continued)

- Large literature in sociology which uses triad counts to "test" various hypotheses
 - see Holland and Leinhardt (1976, SM) and Wasserman and Faust (1994)
 - cf., computational biology (e.g., Milo et al., 2002)
- Asymptotic distribution theory puts these tests on firmer ground

Introduction (continued)

- Subgraph frequencies might be used to (partially) identify structural models of network formation (e.g., de Paula et al., 2015)
- indirect inference approach:
 - use structural model to simulate networks...and count subgraphs
 - compare simulated counts with actual counts

Setup

Let $G(\mathcal{V}, \mathcal{E})$ be a finite undirected random graph with

- agents/vertices $V = \{1, ..., N\}$,
- links/edges $\mathcal{E} = \{\{i, j\}, \{k, l\}, \ldots\}$, and
- ullet adjacency matrix $\mathbf{D} = \begin{bmatrix} D_{ij} \end{bmatrix}$ with

$$D_{ij} = \begin{cases} 1 & \text{if } \{i, j\} \in \mathcal{E} \\ 0 & \text{otherwise} \end{cases}$$

Subgraphs

- (Partial Subgraph) Let $V(S) \subseteq V(G)$ be any subset of the vertices of G and $\mathcal{E}(S) \subseteq \mathcal{E}(G) \cap V(S) \times V(S)$, then $S = (V(S), \mathcal{E}(S))$ is an partial subgraph of G.
- (Induced Subgraph) Let $\mathcal{V}(S) \subseteq \mathcal{V}(G)$ be any subset of the vertices of G and $\mathcal{E}(S) = \mathcal{E}(G) \cap \mathcal{V}(S) \times \mathcal{V}(S)$, then $S = (\mathcal{V}(S), \mathcal{E}(S))$ is an *induced subgraph* of G.

Subgraphs (continued)

- The induced subgraph S includes all edges in G connecting any two agents in $\mathcal{V}(S)$
 - a (partial) subgraph may include only a subset of such edges
 - $-S=\bigwedge$ is a partial subgraph of $G=\bigvee$, but not an induced subgraph

Graph Isomorphism

- Consider two graphs, R and S, of the same order.
- Let $\varphi : \mathcal{V}(R) \to \mathcal{V}(S)$ be a bijection from the nodes of R to those of S.
- The bijection $\varphi: \mathcal{V}(R) \to \mathcal{V}(S)$
 - maintains adjacency if for every dyad $i, j \in \mathcal{V}(R)$ if $\{i, j\} \in \mathcal{E}(R)$, then $\{\varphi(i), \varphi(j)\} \in \mathcal{E}(S)$;
 - maintains non-adjacency if for every dyad $i, j \in \mathcal{V}(R)$ if $\{i, j\} \notin \mathcal{E}(R)$, then $\{\varphi(i), \varphi(j)\} \notin \mathcal{E}(S)$.

Graph Isomorphism (continued)

- If the bijection maintains both adjacency and non-adjacency we say it *maintains structure*.
- (<u>Graph Isomorphism</u>) The graphs R and S are isomorphic if there exists a structure-maintaining bijection $\varphi: \mathcal{V}(R) \to \mathcal{V}(S)$.
- Notation: $R \cong S$ means "R is isomorphic to S."

Induced Subgraph Density

- S is a p^{th} -order graphlet of interest (e.g., $S = \mathbb{N}$ or $S = \mathbb{A}$)
- ullet G_N is the network/graph under study
- $\mathbf{i}_p \subseteq \{1, 2, \dots, N\}$ is a set of p integers with $i_1 < i_2 < \dots < i_p$
 - $-\mathcal{C}_{p,N}$ is set of all $\binom{N}{p}$ such integer sets
 - $G\left[\mathbf{i}_{p}\right]$ is the induced subgraph of G associated with vertex set \mathbf{i}_{p}

Induced Subgraph Density (continued)

• The induced subgraph density of S in G_N , denoted by $t_{\text{ind}}(S, G_N)$ or $P_N(S)$ equals the probability that $G_N[\mathbf{i}_p]$, for \mathbf{i}_p chosen uniformly at random from $C_{p,N}$, is isomorphic to S:

$$t_{\text{ind}}(S, G_N) = {N \choose p}^{-1} \sum_{\mathbf{i}_p \in C_{p,N}} \mathbf{1}(S \cong G_N [\mathbf{i}_p])$$
$$= \Pr(S \cong G_N [\mathbf{i}_p])$$
$$= P_N(S)$$

Induced Subgraph Density (Examples)

•
$$t_{\text{ind}}(\ \ \ \ \ \) = \frac{2}{4}$$
, $t_{\text{ind}}(\ \ \ \ \ \) = \frac{2}{4}$ and $t_{\text{ind}}(\ \ \ \ \ \) = \frac{0}{4}$

•
$$t_{\text{ind}}(, , , ,) = \frac{1}{4}$$
, $t_{\text{ind}}(, , , ,) = \frac{2}{4}$ and $t_{\text{ind}}(, , , ,) = \frac{1}{4}$

Induced Subgraph Density: Graphon Case

- Let $h\left(U_i,U_j\right)$ be a valid graphon.
- ullet iso (S) is the group of isomorphisms of S, and $|\mathrm{iso}\,(S)|$ its cardinality
- \bullet Under the "Aldous-Hoover DGP" the *ex ante* probability that an induced p-subgraph is isomorphic to S is given by

$$t_{\text{ind}}(S, h) = |\text{iso}(S)|$$

$$\times \mathbb{E}\left[\prod_{\{i,j\}\in\mathcal{E}(S)} h\left(U_i, U_j\right) \prod_{\{i,j\}\in\mathcal{E}(\bar{S})} \left[1 - h\left(U_i, U_j\right)\right]\right]$$

$$= P(S).$$

Graph Limits

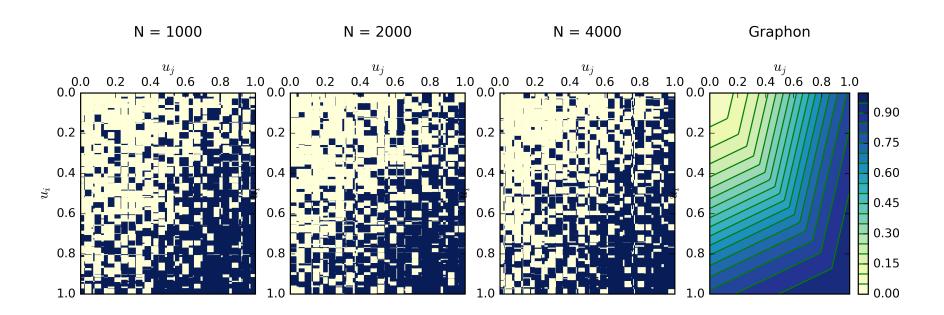
Let $\{G_N\}_{N=1}^{\infty}$ be a sequence of networks. If

$$\lim_{N\to\infty} t_{\text{ind}}(S, G_N) = t_{\text{ind}}(S, h)$$

for some graphon $h(\cdot,\cdot)$ and all fixed subgraphs S, then we say that G_N converges to $h(\cdot,\cdot)$.

- Lovász (2012) for complete development
- Diaconis and Janson (2008) for connections with Aldous-Hoover Theorem

Graph Limits: Example



(Injective) Homomorphism Density

- The homomorphism density gives the probability that S is (isomorphic to) a subgraph of a randomly selected induced subgraph of G_N of order $p = |\mathcal{V}(S)|$
- Alternatively the homomorphism density equals fraction of injective mappings $\varphi: \mathcal{V}(S) \to \mathcal{V}(G_N)$ that preserve edge adjacency

$$t_{\mathsf{hom}}(S, G_N) = \frac{1}{\binom{N}{p} |\mathsf{iso}(S)|} \sum_{R \subseteq K_N, R \cong S} \mathbf{1}(R \subseteq G_N)$$

$$= \frac{1}{\binom{N}{p} |\mathsf{iso}(S)|} \sum_{R \subseteq K_N, |V(R)| = p} \mathbf{1}(R \cong S) \prod_{\{i,j\} \in \mathcal{E}(R)} D_{ij}$$

$$= Q_N(S)$$

Homomorphism Density (continued)

- Summation in $t_{\text{hom}}(S, G_N) = Q_N(S)$ is over the $\binom{N}{3} | \text{iso}(\bigwedge)| = \frac{3}{6}N(N-1)(N-2)$ (partial) subgraphs of K_N (the complete graph) which are isomorphic to $S = \bigwedge$).
- ullet We count the number of these subgraphs which are also partial subgraphs of G_N

Homomorphism Density (continued)

• The expected value of $Q_N(S)$ is:

$$\mathbb{E}\left[Q_{N}(S)\right] = \frac{1}{\binom{N}{p}|\operatorname{iso}(S)|} \sum_{R \subseteq K_{N}, |V(R)| = p} \{1 (R \cong S) \\ \times \mathbb{E}\left[\mathbb{E}\left[\prod_{\{i,j\} \in \mathcal{E}(R)} D_{ij} \middle| U_{1}, \dots, U_{N}\right]\right]\right\}$$
$$= \mathbb{E}\left[\prod_{\{i,j\} \in \mathcal{E}(S)} h\left(U_{i}, U_{j}\right)\right]$$
$$= Q(S) = t_{\text{hom}}(S, h)$$

• Can use $t_{\text{hom}}(S, G_N)$ to define graph convergence

Recap

- Induced subgraph density, $P_N(S)$: probability that $G_N[\mathbf{i}_p]$, for \mathbf{i}_p chosen uniformly at random from $C_{p,N}$, is isomorphic to S
- Homomorphism density, $Q_N(S)$: probability that a subgraph of $G_N[\mathbf{i}_p]$, for \mathbf{i}_p chosen uniformly at random from $C_{p,N}$, is isomorphic to S
- If $\lim_{N\to\infty} P_N(S) = t_{\text{ind}}(S,h)$ for some graphon $h(\cdot,\cdot)$ and all fixed subgraphs S, then we say that G_N converges to $h(\cdot,\cdot)$.

One more tool! Graphlet Stitchings

- Graph union: $T \cup U = G(\mathcal{V}(T) \cup \mathcal{V}(U), \mathcal{E}(T) \cup \mathcal{E}(U))$
- ullet Let $W_{q,R,S}$ be a union of two isomorphisms, respectively T and U, of the graphlets R and S with
 - 1. |V(R)| = |V(S)| = p
 - 2. $|\mathcal{V}(R) \cap \mathcal{V}(S)| = q$ vertices in common
 - 3. identical structures across all vertices in common
- ullet The multiset of all such graphlet stitchings (including isomorphisms) is denoted by $\mathcal{W}_{q,R,S}$ (with $\mathcal{W}_{q,S,S}=W_{q,S}$)

Graphlet Stitching: Example #1

- Let the graphlets $R= \longrightarrow$ and $S= \longrightarrow$ share one vertex in common.
 - There is just one possible way to join them: $R \cup S \cong \bigwedge$
- \bullet We therefore have that $\mathcal{W}_{1,\, \longleftarrow} = \big\{ \, {\begin{subarray}{c} \end{subarray}} \, \big\}$

Graphlet Stitching: Example #1 (continued)

 \bullet Define the probability of observing an element of $\mathcal{W}_{1,\, \bullet\!\!-\!\!\!-\!\!\!\!-}$ as a subgraph of a randomly sampled triad as

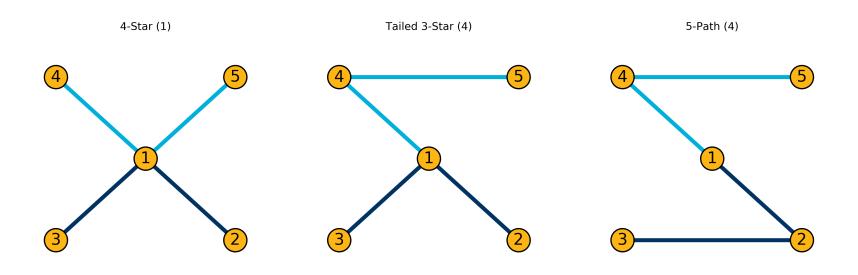
$$Q\left(\mathcal{W}_{1,\bullet,\bullet}\right) = \sum_{W \in \mathcal{W}_{1,\bullet,\bullet}} Q\left(W\right)$$
$$= Q\left(\bigwedge_{1,\bullet,\bullet}\right)$$
$$= \mathbb{E}\left[D_{12}D_{13}\right]$$

with Q(W) the homomorphism density introduced above

 \bullet For q=2 (two nodes in common) we have, of course, $\mathcal{W}_2 = \left\{ \bigodot \right\}$

Graphlet Stitching: Example #2

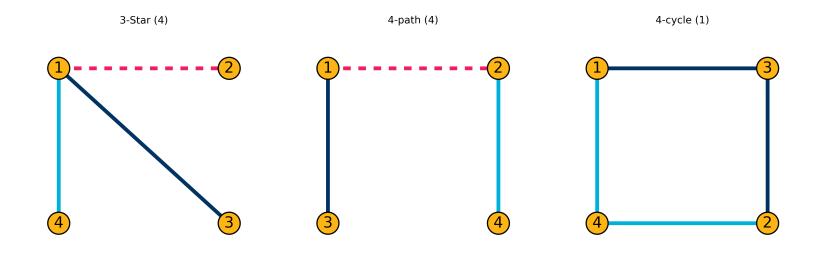
There are nine ways (three up to isomorphisms) to join the graphlets $R = \bigwedge$ and $S = \bigwedge$, sharing one vertex in common.



Notes: Number of isomorphisms of each graphlet in $\mathcal{W}_{1,q}$ given in parentheses.

Graphlet Stitching: Example #2 (continued)

There are nine ways (three up to isomorphisms) to join the graphlets $R = \bigwedge$ and $S = \bigwedge$, sharing two vertices in common.



Notes: Number of isomorphisms of each graphlet in $W_{2,q}$ given in parentheses.

Estimation of Subgraph Frequencies

- We will develop explicit results for two subgraph frequencies
 - the frequency of connected dyads: $S = \longrightarrow$
- General case involves no new ideas...
 - ...but can be very tedious in practice

Density

ullet We estimate $ho_N=\Pr\left(D_{ij}=1
ight)$ by $\widehat{
ho}_N=rac{2}{N\left(N-1
ight)}\sum_{i< j}D_{ij}$

ullet Projecting onto $U_1,....,U_N$ yields the decomposition:

$$\widehat{\rho}_{N} = \underbrace{\frac{2}{N\left(N-1\right)} \sum_{i < j} h_{N}\left(U_{i}, U_{j}\right)}_{\text{U-Statistic}} + \underbrace{\frac{2}{N\left(N-1\right)} \sum_{i < j} \left(D_{ij} - h_{N}\left(U_{i}, U_{j}\right)\right)}_{\text{"Poisson Binomial R.V"}}$$

$$= U_{N} + T_{N}$$

ullet Observe that T_N is mean independent of U_N

Density: Variance Calculation

We have

$$\mathbb{V}(\widehat{\rho}_N) = \mathbb{V}(U_N) + \mathbb{V}(T_N) + 2\mathbb{C}(U_N, T_N)$$
$$= \mathbb{V}(U_N) + \mathbb{V}(T_N).$$

A Hoeffding (1948) variance decomposition gives

$$\mathbb{V}(U_N) = {N \choose 2}^{-2} \sum_{q=1}^{2} {N \choose 2} {2 \choose q} {N-2 \choose 2-q} \Omega_q$$

for

$$\Omega_q = \mathbb{C}\left(h_N\left(U_{i_1}, U_{i_2}\right), h_N\left(U_{j_1}, U_{j_2}\right)\right)$$

with $\{i_1,i_2\}$ and $\{j_1,j_2\}$ sharing q=1,2 indices in common

Evaluating Ω_1 yields

$$\Omega_{1} = \mathbb{E} \left[h_{N} \left(U_{1}, U_{2} \right) h_{N} \left(U_{1}, U_{3} \right) \right] - \mathbb{E} \left[h_{N} \left(U_{1}, U_{2} \right) \right] \mathbb{E} \left[h_{N} \left(U_{1}, U_{3} \right) \right]$$

$$= Q \left(\mathcal{N}_{1}, \longrightarrow \right) - P \left(\longrightarrow \right) P \left(\longrightarrow \right)$$

$$= Q \left(\mathcal{N}_{1}, \longrightarrow \right) - P \left(\longrightarrow \right) P \left(\longrightarrow \right)$$

Evaluating Ω_2 yields

$$\Omega_{2} = \mathbb{E}\left[h_{N}\left(U_{1}, U_{2}\right)^{2}\right] - \mathbb{E}\left[h_{N}\left(U_{1}, U_{2}\right)\right] \mathbb{E}\left[h_{N}\left(U_{1}, U_{2}\right)\right]$$
$$= \mathbb{V}\left(\mathbb{E}\left[D_{12}|\mathbf{U}\right]\right)$$

Evaluating the variance of $\mathbb{V}(T_N)$ we get

$$\mathbb{V}(T_{N}) = \mathbb{V}(\mathbb{E}[T_{N}|\mathbf{U}]) + \mathbb{E}[\mathbb{V}(T_{N}|\mathbf{U})]$$

$$= 0 + \left(\frac{2}{N(N-1)}\right)^{2} \mathbb{V}\left(\sum_{i < j} \left(D_{ij} - h_{N}\left(U_{i}, U_{j}\right)\right) \middle| \mathbf{U}\right)$$

$$= \left(\frac{2}{N(N-1)}\right)^{2} \sum_{i < j} \mathbb{V}\left(D_{ij} - h_{N}\left(U_{i}, U_{j}\right)\middle| \mathbf{U}\right)$$

$$= \frac{2}{N(N-1)} \mathbb{E}[\mathbb{V}(D_{12}|\mathbf{U})]$$

Collecting terms we have:

$$\mathbb{V}(\widehat{\rho}_{N}) = \frac{4(N-2)}{N(N-1)} \left[Q\left(\bigwedge \right) - P\left(\longrightarrow \right) P\left(\longrightarrow \right) \right]$$

$$+ \frac{2}{N(N-1)} \mathbb{V}\left(\mathbb{E}\left[D_{12} | \mathbf{U} \right] \right) + \frac{2}{N(N-1)} \mathbb{E}\left[\mathbb{V}\left(D_{12} | \mathbf{U} \right) \right]$$

$$= \frac{4(N-2)}{N(N-1)} \left[Q\left(\bigwedge \right) - P\left(\longrightarrow \right) P\left(\longrightarrow \right) \right]$$

$$+ \frac{2}{N(N-1)} P\left(\longrightarrow \right) \left(1 - P\left(\longrightarrow \right) \right)$$

ullet To allow for graph sequences where $ho_N o 0$ as $N o \infty$ we normalize

- Recall that $\lambda_N = (N-1) \rho_N$

• After normalization:

$$\mathbb{V}\left(\frac{\hat{\rho}_{N}}{\rho_{N}}\right) = \frac{4(N-2)}{N(N-1)} \left[\tilde{Q}\left(\frac{\wedge}{\wedge}\right) - \tilde{P}\left(\frac{\wedge}{\wedge}\right)\tilde{P}\left(\frac{\wedge}{\wedge}\right)\right] + \frac{2}{N\lambda_{N}}\tilde{P}\left(\frac{\wedge}{\wedge}\right) - \frac{2}{N(N-1)}\tilde{P}\left(\frac{\wedge}{\wedge}\right)^{2}$$
$$= O\left(\frac{1}{N}\right) + O\left(\frac{1}{N\lambda_{N}}\right) + O\left(\frac{1}{N^{2}}\right)$$

- If $\lambda_N \to \infty$ first term dominates
- If $\lambda_N \to \lambda_0 > 0$, first two terms dominate

Asymptotic Inference

 \bullet Asymptotic theory for U-Statistics gives, for $\lambda_N \to \infty$ as $N \to \infty$

$$\sqrt{N} \left(\frac{\widehat{\rho}_{N}}{\rho_{N}} - 1 \right) \stackrel{D}{\to} \mathcal{N} \left(0, 4 \left[\tilde{Q} \left(\right) \right) - \tilde{P} \left(\right) \right) \left[\tilde{P} \left(\right) \right] \right)$$

- Result due to Bickel, Chen and Levina (2011, Annals of Statistics)
- Comment: Under Erdos-Renyi $\tilde{Q}\left(\ref{Q} \right) = \tilde{P}\left(\ref{Q} \right) \tilde{P}\left(\ref{Q} \right)$

Variance Estimation

We can estimate the asymptotic variance using the analog estimators:

$$\widehat{Q}\left(\bigwedge\right) = \binom{N}{3}^{-1} \sum_{i < j < k} \frac{1}{3} \left\{ D_{ij} D_{ik} + D_{ij} D_{jk} + D_{ik} D_{jk} \right\}$$
$$= \binom{N}{3}^{-1} \frac{1}{3} \left[T_{\mathsf{TS}} + 3T_{\mathsf{T}} \right]$$

and

$$\widehat{P}\left(\longrightarrow \right) = \binom{N}{2}^{-1} \sum_{i < j} D_{ij}$$

Variance Estimation for $\widehat{P}(\longrightarrow)$: Nyakatoke

For Nyakatoke we have

$$\widehat{Q}\left(\bigwedge \right) \cong 0.006105$$

and

$$\hat{P}\left(\bullet\bullet\right)\simeq0.0698$$

which gives

$$\hat{\rho}_N$$
 = $\frac{0.0698}{(0.0064)}$

Note: Estimate of excluded higher order variance term is 4.6 \times 10^{-6}

Limit Distribution of $\hat{P}(\nearrow)$

Define the multiset of graphlet stitchings:

$$- \mathcal{W}_{1, \wedge} = (\{ \nwarrow, \sqcap, \square \}, m)$$

– Here $m=\left\{(\nwarrow,4),(\lnot,4),(\lnot,1)\right\}$ gives the multiplicity of each unique graphlet in $\mathcal{W}_{1,}$

Normalize graphlet according to number of edges in it.

$$-\tilde{P}(\boldsymbol{\wedge}) = \frac{P(\boldsymbol{\wedge})}{\rho_N^2} \text{ and } \tilde{Q}\left(\mathcal{W}_{1,\boldsymbol{\wedge}}\right) = \frac{Q\left(\mathcal{W}_{1,\boldsymbol{\wedge}}\right)}{\rho_N^4}$$

Limit Distribution of $\widehat{P}(\bigwedge)$

If $\lambda_N \to \infty$ as $N \to \infty$, then

$$\sqrt{N}\left(\frac{\widehat{P}(\nwarrow)}{\rho_N^2} - \widetilde{P}(\nearrow)\right) \xrightarrow{D} \mathcal{N}\left(0, 9\left[\widetilde{Q}\left(\mathcal{W}_{1, \nearrow}\right) - \widetilde{P}\left(\nearrow\right)\widetilde{P}\left(\nearrow\right)\right]\right).$$

- Analysis involves a variance calculation along the lines outlined above
- And the characterization of the limiting variance of a 3rd order U-Statistics

Wrapping Up

- In large graphs subgraph counting is computationally challenging
 - implications for feasibility of both estimation and inference
 - see Bhattacharya and Bickel (2015) for a subsampling approach
- Very little (i.e., almost none) empirical work using these results
- Tremendous scope for using these methods in empirical analysis; but not easy!