



# **Free Lunch Multivariate Forecast Projection: reducing forecast variance using linear combinations**

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# Free lunch forecast projection

A model-independent post-forecast adjustment method that can reduce forecast variance.

- Averaging indirect forecasts from linear combinations (components)
- Projecting forecasts of augmented series
- Free lunch: no additional data or information needed

# What to expect

- Intuition with data
- Literature
- Method formulation
- Properties
- Empirical applications and simulation

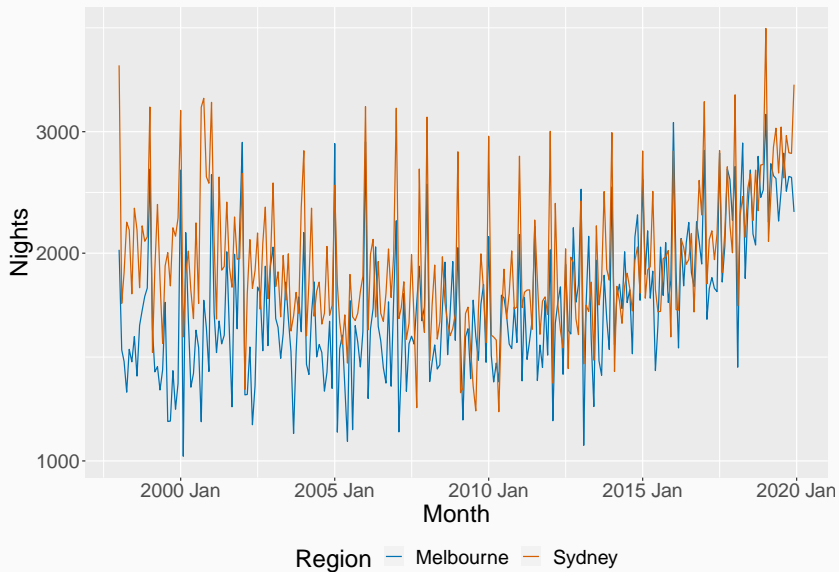
# Australian tourism data

- The data include tourism information on seven states and territories which can be divided into 77 regions
  - ▶ For example, Melbourne, Sydney, East Coast

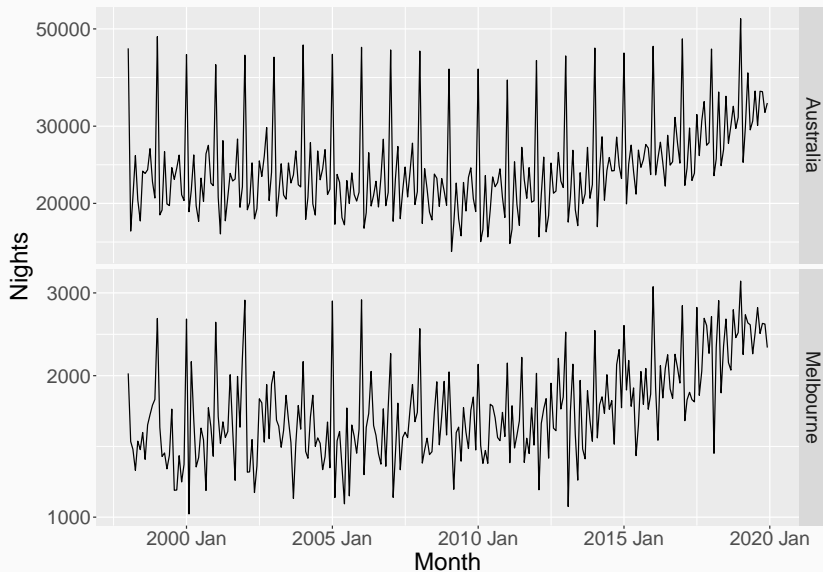
## Visitor nights

The total number of nights spent by Australians away from home recorded monthly

# Melbourne and Sydney



# Total and Region



# Intuition

## Observation

1. Similar patterns are shared by different series.
2. Better signal-noise ratio in the linear combination.

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## One step further

Finding components that

1. are easy to forecast;
2. can capture the common signals;
3. can improve forecast of original series.



## Forecast reconciliation

- Wickramasuriya, Athanasopoulos, and Hyndman (2019): Projecting forecasts to be consistent with the hierarchical structure

## Forecast combination

- Combining forecasts of the target series
- Hollyman, Petropoulos, and Tipping (2021): Combining direct and indirect forecasts
- Petropoulos and Spiliotis (2021): Combining forecasts of selections and transformations of the target series (“wisdom of data”)

# Literature

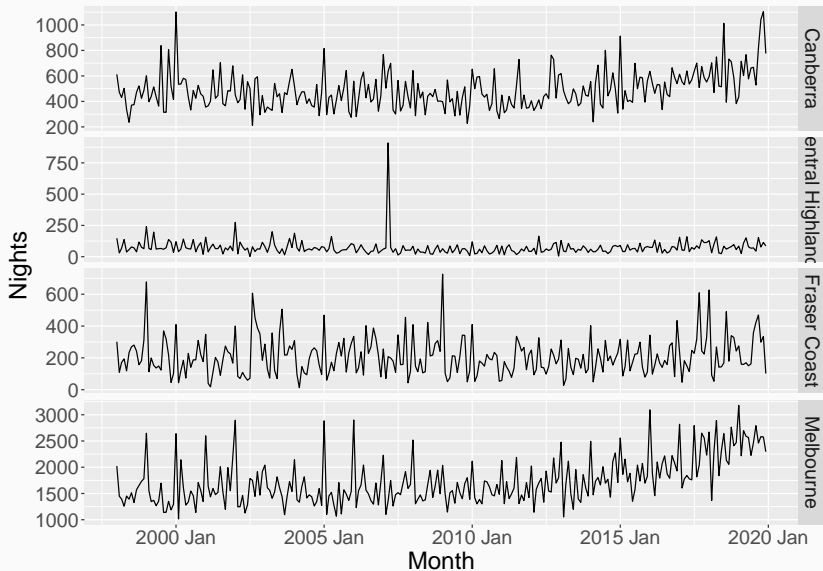
## Bagging

- Bergmeir, Hyndman, and Benítez (2016): Bagging ETS models to forecast
- Petropoulos, Hyndman, and Bergmeir (2018): The benefits of bagging originate from the model uncertainty

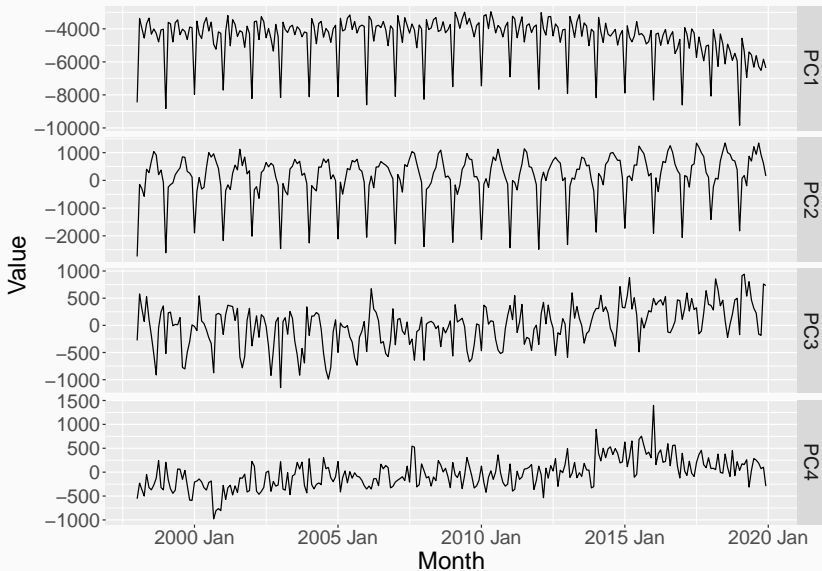
## Dynamic factor model (DFM)

- Stock and Watson (2002a), Stock and Watson (2002b)
- De Stefani et al. (2019): Machine learning extension

# Series $\mathbf{z}_t \in \mathbb{R}^m$



# Components $\mathbf{c}_t = \Phi \mathbf{z}_t \in \mathbb{R}^p$



# Free lunch forecast projection

$$\mathbf{y}_t = \begin{bmatrix} \mathbf{z}_t \\ \mathbf{c}_t \end{bmatrix} \quad \tilde{\mathbf{y}}_{t+h} = \mathbf{M}\hat{\mathbf{y}}_{t+h}$$

$$\tilde{\mathbf{z}}_{t+h} = \mathbf{J}\tilde{\mathbf{y}}_{t+h} = \mathbf{J}\mathbf{M}\hat{\mathbf{y}}_{t+h}$$

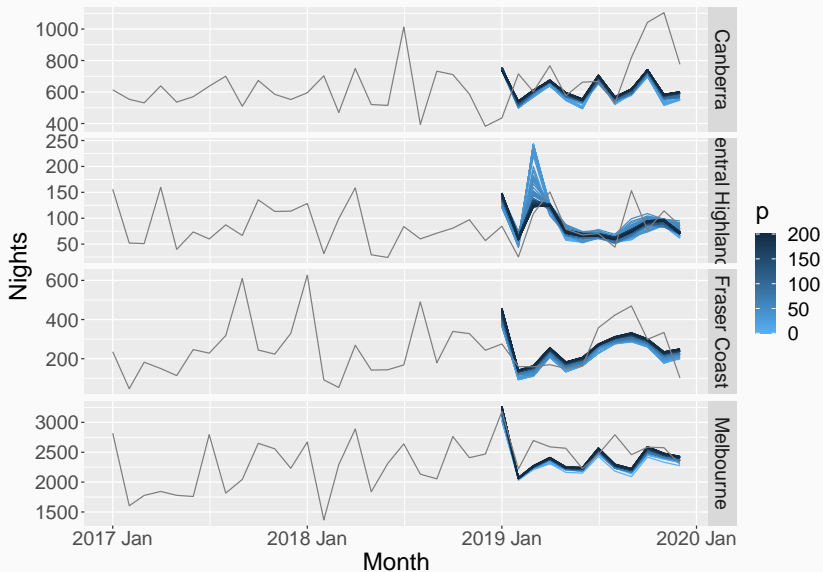
$$\mathbf{M} = \mathbf{I}_{m+p} - \mathbf{W}_h \mathbf{C}' (\mathbf{C} \mathbf{W}_h \mathbf{C}')^{-1} \mathbf{C}$$

$$\mathbf{J} = \mathbf{J}_{m,p} = \begin{bmatrix} \mathbf{I}_m & \mathbf{O}_{m \times p} \end{bmatrix}$$

$$\mathbf{C} = \begin{bmatrix} -\Phi & \mathbf{I}_p \end{bmatrix}$$

$$\mathbf{W}_h = \text{Var}(\hat{\mathbf{y}}_{t+h})$$

# Forecasts and projection of series



# Unbiasedness

## Unbiasedness

If the base forecasts are unbiased, then the projected forecasts are also unbiased.

## Projection matrix

- 1 The matrix  $\mathbf{M}$  is a projection onto the space where the constraint  $\mathbf{C}$  is satisfied.
- 2 The projected forecast  $\tilde{\mathbf{y}}_{t+h}$  satisfies the constraint  $\mathbf{C}$ .
- 3 For  $\mathbf{y}_{t+h}$  that already satisfies the constraint, the projection does not change its value:

$$\mathbf{M}\mathbf{y}_{t+h} = \mathbf{y}_{t+h}$$

# Nonnegative variance reduction

Under unbiasedness, the variance reduction is **positive semi-definite**:

$$\begin{aligned}\text{Var}(\hat{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) - \text{Var}(\tilde{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) \\ = \mathbf{J}\mathbf{W}_h\mathbf{C}'(\mathbf{C}\mathbf{W}_h\mathbf{C}')^{-1}\mathbf{C}\mathbf{W}_h\mathbf{J}'\end{aligned}$$



## Example $W_h = I_{m+p}$

$$\begin{aligned}\text{Var}(\hat{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) - \text{Var}(\tilde{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) \\&= \mathbf{J}\mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1}\mathbf{C}\mathbf{J}' \\&= \Phi'(\Phi\Phi' + \mathbf{I})^{-1}\Phi\end{aligned}$$

Let  $\Phi$  consist of orthogonal unit vectors:

$$\Phi\Phi' = \mathbf{I}_p \text{ when } p \leq m$$

$$\Phi'\Phi = \mathbf{I}_m \text{ when } p = m.$$

$$\begin{aligned}\text{tr}(\text{Var}(\hat{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) - \text{Var}(\tilde{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h})) \\&= \frac{1}{2}\text{tr}(\Phi'\Phi) = \frac{1}{2}p\end{aligned}$$

# Positive condition

For the first component to have a guaranteed reduction of forecast variance, the following condition must be satisfied:

$$\phi_1 \mathbf{W}_{z,h} \neq \mathbf{w}_{c_1 z,h},$$

- $\phi_1$  is the weight vector of the first component
- $\mathbf{W}_{z,h} = \text{Var}(\hat{\mathbf{z}}_{t+h})$
- $\mathbf{w}_{c_1 z,h}$  is the forecast covariance between the first component and the original series.

# Monotonicity

The sum of forecast variance reductions

$$\begin{aligned} & \text{tr}(\text{Var}(\hat{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h}) - \text{Var}(\tilde{\mathbf{z}}_{t+h} - \mathbf{z}_{t+h})) \\ &= \text{tr}(\mathbf{J}\mathbf{W}_h\mathbf{C}'(\mathbf{C}\mathbf{W}_h\mathbf{C}')^{-1}\mathbf{C}\mathbf{W}_h\mathbf{J}') \end{aligned}$$

is non-decreasing as  $p$  increases.

# Minimum variance of individual series

The projection is equivalent to the mapping

$$\tilde{\mathbf{z}}_{t+h} = \mathbf{G}\hat{\mathbf{y}}_{t+h},$$

where  $\mathbf{G} = [\mathbf{g}_1 \ \mathbf{g}_2 \ \dots \ \mathbf{g}_m]' \in \mathbb{R}^{m \times (m+p)}$  is the solution to

$$\arg \min_{\mathbf{G}} \mathbf{G}\mathbf{W}_h\mathbf{G}' \quad \text{s.t. } \mathbf{G}\mathbf{S} = \mathbf{I}$$

or

$$\arg \min_{\mathbf{g}_i} \mathbf{g}_i'\mathbf{W}_h\mathbf{g}_i \quad \text{s.t. } \mathbf{g}_i'\mathbf{s}_j = \mathbf{1}(i=j),$$

$$\text{where } \mathbf{S} = \begin{bmatrix} \mathbf{I}_m \\ \Phi \end{bmatrix} = [\mathbf{s}_1 \ \dots \ \mathbf{s}_m].$$

# Key results

- 1 The forecast variance is **reduced** with forecast projection
- 2 The forecast variance **monotonically** decreases with increasing number of components
- 3 The forecast projection is **optimal** to achieve minimum forecast variance of each series

# Estimation

$$\tilde{\mathbf{z}}_{t+h} = \mathbf{J}\tilde{\mathbf{y}}_{t+h} = \mathbf{J}\mathbf{M}\hat{\mathbf{y}}_{t+h}$$

$$\mathbf{M} = \mathbf{I}_{m+p} - \mathbf{W}_h \mathbf{C}' (\mathbf{C} \mathbf{W}_h \mathbf{C}')^{-1} \mathbf{C}$$

$$\mathbf{W}_h = \text{Var}(\hat{\mathbf{y}}_{t+h})$$

$$\mathbf{C} = \begin{bmatrix} -\Phi & \mathbf{I}_p \end{bmatrix}$$

- Estimation of  $\mathbf{W}_h$
- Construction of  $\Phi$

# Estimation of $\mathbf{W}_h$

**Shrinking variance** towards their median (Opge-Rhein and Strimmer 2007) and **shrinking covariance** towards zero (Schäfer and Strimmer 2005).

The shrinkage estimator is

- Positive definite, and
- Numerically stable.

In empirical applications, we assume

$$\widehat{\mathbf{W}}_h^{shr} = \eta_h \widehat{\mathbf{W}}_1^{shr}.$$

# Construction of $\Phi$

## Principal component analysis (PCA)

Finding the weights matrix so that the resulting components **maximise variance**

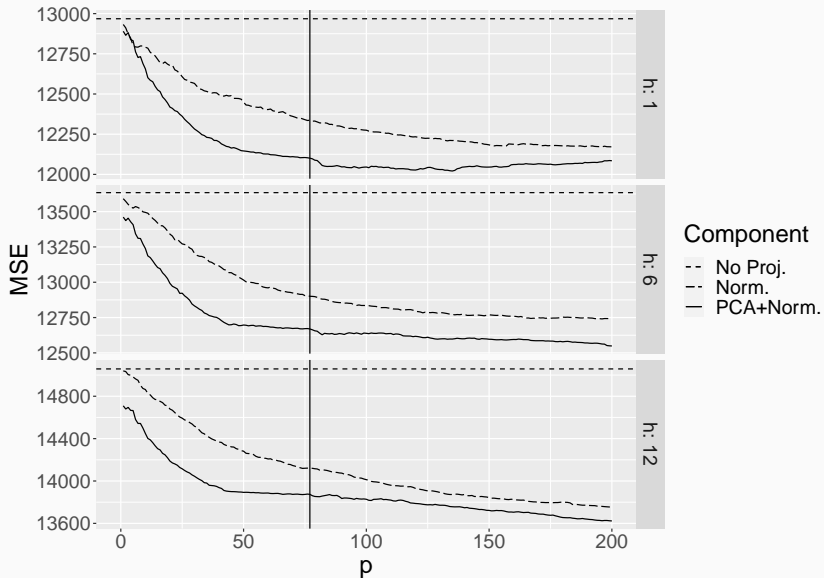
## Simulation

Generating values from a random distribution and normalising them to unit vectors

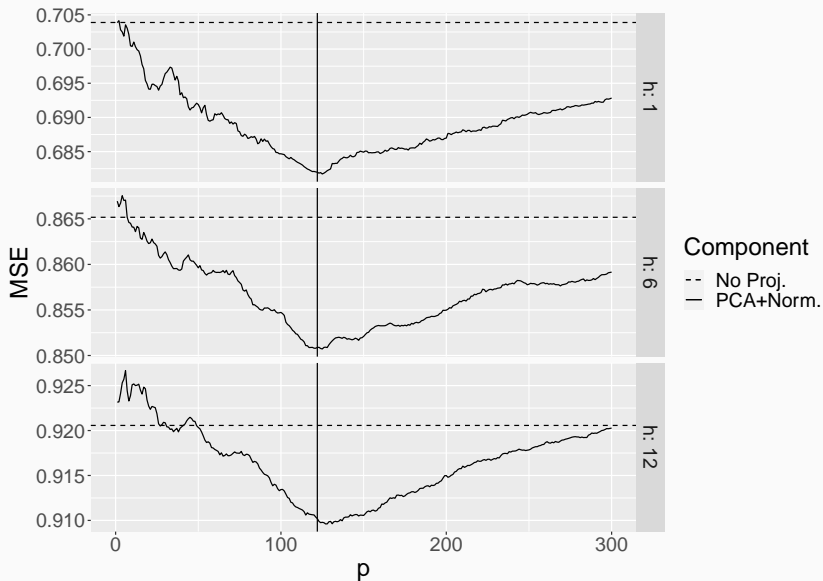
- Normal distribution
- Uniform distribution
- Orthonormal matrix (Borchers 2023)



# Tourism (ETS)



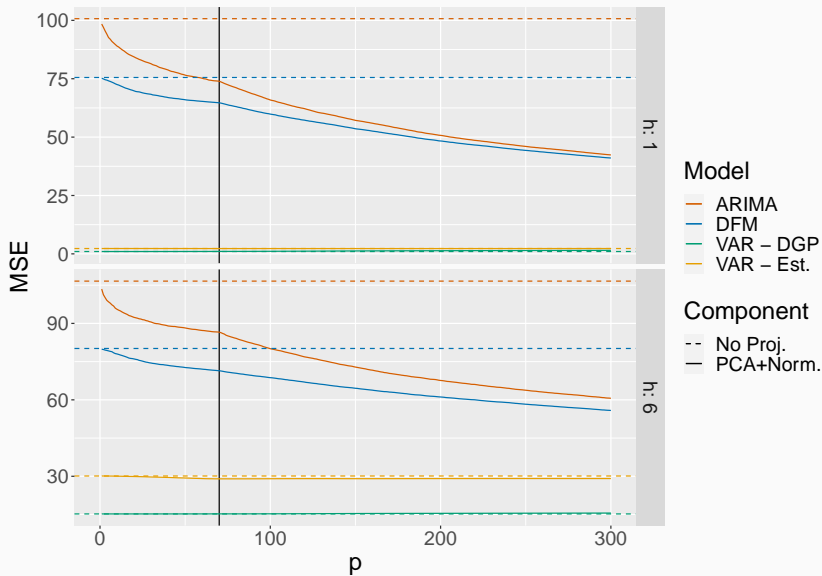
# FRED-MD (DFM)



# Simulation

- Data generating process (DGP): VAR(3) with  $m = 70$  variables
- Sample size:  $T = 400$
- Number of repeated samples: 220
- Base model: ARIMA and DFM

# Simulation



# Future research directions

- Investigate why PCA performs better than random weights
- Find other components that are better than PCA
- Find optimal components by minimising forecast variance with respect to  $\Phi$
- Use forecast projection and forecast reconciliation together

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