



# Two-Factor Hull-White Model for Interest Rate Derivative Products in Bloomberg

May 1, 2015

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## Abstract

This document describes the detailed implementation of the Two-Factor Hull-White model for path-dependent interest rate derivatives

**Keywords.** Hull-White, Two-Factor, Monte Carlo, Dual Curve.

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# 1 Two-factor Hull-White model

## 1.1 Model dynamics

### 1.1.1 Risk neutral measure

Under the risk neutral measure  $Q$ , numeraire is given by the money market account, i.e.

$$\mathcal{N}_{rn}(t, \mathcal{F}_t) := \exp \left[ \int_0^t r(s) ds \right]. \quad (1)$$

The two-factor Hull-White model for short rate  $r(t)$  is

$$r(t) = \theta(t) + \mathbf{1}^\top \mathbf{X}(t), \quad (2)$$

where  $\theta(t)$  is a deterministic function to match the initial yield curve.  $\mathbf{X}(t)$  is a two dimensional Ornstein-Uhlenbeck process,

$$d\mathbf{X}(t) = -\boldsymbol{\kappa}\mathbf{X}(t)dt + \boldsymbol{\sigma}(t)d\mathbf{W}^Q(t), \quad \mathbf{X}(0) = \mathbf{0}, \quad (3)$$

$$\left( d\mathbf{W}^Q(t) \right) \left( d\mathbf{W}^Q(t) \right)^\top = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} dt := \boldsymbol{\rho} dt. \quad (4)$$

with parameters  $\boldsymbol{\kappa}$ ,  $\boldsymbol{\sigma}(t)$  and  $\boldsymbol{\rho}$ .

$$\boldsymbol{\kappa} = \begin{bmatrix} \kappa_1 & 0 \\ 0 & \kappa_2 \end{bmatrix}, \quad (5)$$

$$\boldsymbol{\sigma}(t) = \begin{bmatrix} \sigma_1(t) & 0 \\ 0 & \sigma_2(t) \end{bmatrix}. \quad (6)$$

Constants  $\kappa_1$  and  $\kappa_2$  are speeds of mean-reversion for the first and second factor, respectively. We assume that the first factor captures long term dynamics, while the second factor captures short term dynamics. Therefore,  $0 \leq \kappa_1 < \kappa_2$ .  $\sigma_u(t) \geq 0$  ( $u = 1, 2$ ) are volatility related parameters obtained by calibrating to the market volatilities.  $\rho$  is the correlation between the two factors.

The solution to the SDE specified in [Eqs. 3](#) and [Eqs. 4](#) is

$$\mathbf{X}(t) = \mathbf{h}(t - t_0)\mathbf{x} + \int_{t_0}^t \mathbf{h}(t - s)\boldsymbol{\sigma}(s)d\mathbf{W}^Q(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (7)$$

$$= \int_0^t \mathbf{h}(t - s)\boldsymbol{\sigma}(s)d\mathbf{W}^Q(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0}, \quad (8)$$

where

$$\mathbf{h}(t) := e^{-\boldsymbol{\kappa}t} = \begin{bmatrix} h_1(t) & 0 \\ 0 & h_2(t) \end{bmatrix}, \quad (9)$$

$$h_u(t) := h(t; \kappa_u) := e^{-\kappa_u t}.$$

From Eqs. 7,

$$\left[ \mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N}(\boldsymbol{\mu}^Q(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t)),$$

where

$$\boldsymbol{\mu}^Q(t_0, t; \mathbf{x}) = \mathbf{h}(t - t_0)\mathbf{x} \quad (10)$$

$$\boldsymbol{\Sigma}(t_0, t) = \boldsymbol{\nu}(t) - \mathbf{h}(t - t_0)\boldsymbol{\nu}(t_0)\mathbf{h}(t - t_0). \quad (11)$$

where,  $\boldsymbol{\nu}(t)$  is the co-variance of  $\mathbf{X}(t)$ , i.e.

$$\boldsymbol{\nu}(t) := \text{Var}[\mathbf{X}(t)\mathbf{X}^\top(t)] = \int_0^t \mathbf{h}(t-s)\boldsymbol{\sigma}(s)\boldsymbol{\rho}\boldsymbol{\sigma}(s)\mathbf{h}(t-s)\text{d}s = \begin{bmatrix} \nu_{11}(t) & \nu_{12}(t) \\ \nu_{21}(t) & \nu_{22}(t) \end{bmatrix}, \quad (12)$$

$$\nu_{uv}(t) = \rho_{uv} \int_0^t \sigma_u(s)\sigma_v(s)e^{-(\kappa_u + \kappa_v)(t-s)}\text{d}s. \quad (13)$$

### 1.1.2 Terminal forward measure

Under terminal forward measure  $Q_T$ , numeraire is zero coupon bond with maturity  $T$ , i.e.

$$\mathcal{N}_{Q_T}(t, \mathcal{F}_t) := \mathbb{P}(t, T). \quad (14)$$

Therefore, Eqs. 3 becomes

$$\text{d}\mathbf{X}(t) = \left( -\boldsymbol{\kappa}\mathbf{X}(t) - \boldsymbol{\sigma}(t)\boldsymbol{\rho}\boldsymbol{\sigma}(t)\mathbf{H}(T-t)\mathbf{1} \right)\text{d}t + \boldsymbol{\sigma}(t)\text{d}\mathbf{W}^{Q_T}(t), \quad (15)$$

where,

$$\mathbf{H}(t) := \int_0^t \mathbf{h}(s)\text{d}s = \begin{bmatrix} H_1(t) & 0 \\ 0 & H_2(t) \end{bmatrix}. \quad (16)$$

The solution is,

$$\mathbf{X}(t) = -\boldsymbol{\gamma}^{Q_T}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x} + \int_{t_0}^t \mathbf{h}(t-s)\boldsymbol{\sigma}(s)\text{d}\mathbf{W}^{Q_T}(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (17)$$

$$= -\boldsymbol{\gamma}^{Q_T}(0, t) + \int_0^t \mathbf{h}(t-s)\boldsymbol{\sigma}(s)\text{d}\mathbf{W}^{Q_T}(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0} \quad (18)$$

where,

$$\boldsymbol{\gamma}^{Q_T}(t_0, t) := \left[ \int_{t_0}^t \mathbf{h}(t-s)\boldsymbol{\sigma}(s)\boldsymbol{\rho}\boldsymbol{\sigma}(s)\mathbf{H}(T-s)\text{d}s \right] \mathbf{1}. \quad (19)$$

The drift  $\boldsymbol{\gamma}^{Q_T}(t_0, t)$  can be computed as

$$\boldsymbol{\gamma}^{Q_T}(0, t) = \left[ \boldsymbol{\nu}^h(t) + \boldsymbol{\nu}(t)\mathbf{H}(T-t) \right] \mathbf{1} \quad \text{for } 0 \leq t \leq T \quad (20)$$

$$\boldsymbol{\gamma}^{Q_T}(t_0, t) = \boldsymbol{\gamma}^{Q_T}(0, t) - \mathbf{h}(t - t_0)\boldsymbol{\gamma}^{Q_T}(0, t_0) \quad \text{for } 0 \leq t_0 \leq t \leq T. \quad (21)$$

From [Eqs. 17](#),

$$\left[ \mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N} \left( \boldsymbol{\mu}^{Q^T}(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t) \right),$$

where

$$\boldsymbol{\mu}^{Q^T}(t_0, t; \mathbf{x}) = -\boldsymbol{\gamma}^{Q^T}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x} \quad (22)$$

and  $\boldsymbol{\Sigma}(t_0, t)$  is the same as that in risk neutral measure in [Eqs. 11](#).

### 1.1.3 Spot-Libor measure

Under spot-Libor measure  $Q_\beta$ , numeraire is given by the discretely rebalanced bank account (see [\[Brigo & Mercurio\]](#), Section 6.3)

$$\mathcal{N}_\beta(t, \mathcal{F}_t) := \frac{\mathbb{P}(t, T_{\beta(t)})}{\prod_{j=1}^{\beta(t)} \mathbb{P}(T_{j-1}, T_j)} \quad (23)$$

where  $0 = T_0 < T_1 < \dots$  are fixed as our discrete tenor structure,  $\beta(t)$  is the smallest index such that  $t \leq T_{\beta(t)}$ . From [Eqs. 14](#) and [Eqs. 23](#), note that the stochastic part of  $\mathcal{N}_\beta$  is the same as that in  $\mathcal{N}_{Q_{\beta(t)}}$ . Therefore, under the spot-Libor measure, the dynamics changes to

$$d\mathbf{X}(t) = \left( -\kappa \mathbf{X}(t) - \boldsymbol{\sigma}(t) \boldsymbol{\rho} \boldsymbol{\sigma}(t) \mathbf{H}(T_{\beta(t)} - t) \mathbf{1} \right) dt + \boldsymbol{\sigma}(t) d\mathbf{W}^{Q_\beta}(t). \quad (24)$$

Integrating [Eqs. 24](#), we obtain

$$\mathbf{X}(t) = -\boldsymbol{\gamma}^{Q_\beta}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x} + \int_{t_0}^t \mathbf{h}(t - s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_\beta}(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (25)$$

$$= -\boldsymbol{\gamma}^{Q_\beta}(0, t) + \int_0^t \mathbf{h}(t - s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_\beta}(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0} \quad (26)$$

where  $\boldsymbol{\gamma}^{Q_\beta}(t_0, t)$  can be calculated stepwise. When  $\beta(s) = \beta(t)$  for  $t_0 < s < t$ , we have  $\boldsymbol{\gamma}^{Q_\beta}(t_0, t) = \boldsymbol{\gamma}^{Q_{T_{\beta(t)}}}(t_0, t)$ , which is defined in [Eqs. 19](#).

Therefore, the conditional distribution is

$$\left[ \mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N} \left( \boldsymbol{\mu}^{Q_\beta}(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t) \right),$$

where

$$\boldsymbol{\mu}^{Q_\beta}(t_0, t; \mathbf{x}) = -\boldsymbol{\gamma}^{Q_\beta}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x}. \quad (27)$$

## References

- [Brigo & Mercurio] D. Brigo and F. Mercurio, Interest Rate Models: Theory and Practice, Springer, 2001.
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