

# Quantitative Methods for Finance

## Solutions - Gujarati Appendix D

### D.1

- (a) A single numerical value of a (population) parameter is known as a point estimate. An interval estimate provides a range of values that will include the true parameter with a certain degree of confidence (i.e., probability). A point estimator is a formula or rule that tells how to obtain the point estimate.
- (b) A null hypothesis ( $H_0$ ) is the maintained hypothesis which is tested against another hypothesis, called the alternative hypothesis ( $H_1$  or  $H_a$ ).
- (c) Type I error: The error of rejecting a hypothesis when it is true. A type II error is the error of accepting (i.e., not rejecting) a false hypothesis.
- (d) The probability of committing a type I error is known as the level of significance (or the size of the test). One minus the probability of committing a type I error is called the confidence coefficient.
- (e) The probability of accepting a false hypothesis is called a type II error and (1-prob. of type II error), that is, the probability of not committing a type II error is called the power of the test.

### D.2

- (a) The two branches of classical statistics, estimation of parameters and testing hypothesis about parameters, constitute statistical inference.
- (b) The probability distribution of an estimator.
- (c) It is synonymous with a confidence interval.
- (d) A statistic used to decide whether a null hypothesis is rejected or not.
- (e) That value of the test statistic which demarcates the acceptance region from the rejection region. (Note: demarcate = mark by bounds, set the limits to separate)
- (f) It is the probability of committing a type I error.
- (g) The exact level of significance of a test statistic.

### D.3

- (a) If the average, or expected, value of an estimator coincides with the true value of the parameter, that estimator is known as an unbiased estimator.
- (b) In a group of competing estimators of a parameter the one with the least variance is called a minimum variance estimator.
- (c) In the class of unbiased estimators, the one with the least variance is called an efficient estimator.

- (d) An estimator which is a linear function of the observations.
- (e) An unbiased linear estimator with the least possible variance.

#### D.4

- (a) True. In classical statistics the parameter is assumed to be some fixed number, although unknown.
- (b) False. It is  $E(\hat{\mu}_X) = \mu_X$ , where  $\hat{\mu}_X$  is an estimator. (It is not  $\hat{\mu}_X = \mu_X$ .)
- (c) True.
- (d) False. To be efficient, an estimator must be unbiased and it must have minimum variance.
- (e) False. No probabilistic assumption is required for an estimator to be BLUE.
- (f) True.
- (g) False. A type I error is when we reject a true hypothesis.
- (h) False. A type II error occurs when we do not reject a false hypothesis.

Table 3.3 Classifying hypothesis testing errors and correct conclusions			
		Reality	
		H <sub>0</sub> is true	H <sub>0</sub> is false
Result of test	Significant (reject H <sub>0</sub> )	Type I error = $\alpha$	✓
	Insignificant (do not reject H <sub>0</sub> )	✓	Type II error = $\beta$

Type I and type II errors (Brooks, 3ed, p.110)

- (i) True. This can be proved formally.
- (j) False, generally. Only when the sample size increases indefinitely, the sample mean will be normally distributed. If, however, the sample is drawn from a normal population to begin with, the sample mean is distributed normally regardless of the sample size.
- (k) Uncertain. The  $p$ -value is the exact level of significance. If the chosen level of significance, say,  $\alpha = 5\%$ , coincides with the  $p$ -value, the two will mean the same thing.

**D.5**

See Section D.5.

**D.6**

Disagree. The answer depends on the level of significance ( $\alpha$ ), the probability of committing a type I error, that one is willing to accept. There is nothing sacrosanct about the 5% or 10% level of significance.

(Note: sacrosanct = too important or too special to be changed.)

**D.7**

(a) 1.96 (b) 1.65 (c) 2.58 (d) 2.05.

**D.8**

(a) 3.182 (3 d.f.) (b) 2.353 (3 d.f.) (c) 3.012 (13 d.f.)

(d) 2.650 (13 d.f.) (e) 2.0003 (59 d.f.) (f) 1.972 (199 d.f.)

Note: These critical values have been obtained from electronic statistical tables.

**D.9**

(a)  $P(-2 \leq Z \leq 2) = 0.9544$ .

(b)  $P(Z \geq 2) = 0.0228$ .

(c)  $P(Z \leq -2) = 0.0228$ .

(d) Yes ( $Z = 40$ , an extremely high value).

**D.10**

Note that  $X \sim N(1,000, 10)$ . (Here:  $\sigma^2/n = 10$ .)

(a) Practically zero, because  $P(Z \leq -31.6228)$  is negligible.

(b) The 95% confidence interval is:  $893.8019 \leq \mu_X \leq 906.1981$ . With 95% confidence we can say that the true mean is not equal to 1,000.

(c) Reject the null hypothesis. Use the normal distribution because the sample size is reasonably large.

**D.11**

(a) Note that  $(6.5 - \mu)/\sigma = -1.65$  and  $(6.8 - \mu)/\sigma = 1.28$ . Solving these two equations simultaneously, we obtain the two results  $\mu = 6.6689$  and  $\sigma = 0.1024$ .

(b)  $Z = (7 - 6.6689)/0.1024 = 3.2333$ . Therefore,  $P(Z \geq 3.2333)$  is very small, about 0.0006.

**D.12**

Note that  $\bar{X} = 9$ .

(a) If  $X \sim N(5, 2)$ , then  $\bar{X} \sim N(5, 2/10)$ . Therefore,

$$Z = \frac{9 - 5}{0.4472} = 8.9445.$$

Therefore,  $P(Z \geq 8.9445)$  is practically zero. Hence, we reject the null hypothesis that  $\mu = 5$ . Also, note that the 95% confidence interval for  $\bar{X}$  is:

$$\left( 5 \pm 1.96 \frac{1.4142}{\sqrt{10}} \right) = (4.1235, 5.8765).$$

This interval does not include  $\bar{X} = 9$ .

- (b) Also reject the null hypothesis.
- (c) The  $p$ -value is extremely small.

#### D.13

Use the  $t$  distribution, since the true  $\sigma^2$  is unknown. For 9 d.f., the 5% critical  $t$  value is 2.262. Therefore, the 95% CI is:

$$8 \pm 2.262(1.2649) = (5.1388, 10.8612).$$

#### D.14

- (a)  $\bar{X} \sim N(8, 36/25)$ .
- (b)  $Z = (7.5 - 8)/1.2 = -0.4167$ . Therefore,  $P(Z \leq -0.4167) = 0.3372$ .
- (c) The 95% CI for  $\bar{X} = 8 \pm 1.96(1.2) = (5.6480, 10.3520)$ . Since this interval includes the value of 7.5, such a sample could have come from this population.

#### D.15

Using electronic statistical tables, it can be found that

- (a)  $p = 0.0492$ ; (b)  $p = 0.0019$ ; (c)  $p = 0.0814$ ; (d)  $p = 0.9400$ .

#### D.16

The answer depends on the level of  $\alpha$ , the probability of a type I error. The  $p$ -value in this case is about 0.25 (actually, 0.2509). If  $\alpha$  is fixed at 0.25, one could reject the relevant null hypothesis at this level of significance.

#### D.17

- (a)  $E(\hat{\mu}_1) = [E(X_1) + E(X_2) + E(X_3)]/3 = 3\mu/3 = \mu$ . Hence it is an unbiased estimator. Similarly, it can be shown that  $\hat{\mu}_2$  is also unbiased.
- (b)  $\text{Var}(\hat{\mu}_1) = 1/9[\text{Var}(X_1) + \text{Var}(X_2) + \text{Var}(X_3)] = 3\sigma^2/9 = 0.333\sigma^2$ .  
 $\text{Var}(\hat{\mu}_2) = 7/18\sigma^2 = 0.389\sigma^2$ .  
So, choose  $\hat{\mu}_1$  over  $\hat{\mu}_2$ .

**D.18**

- (a) 95% confidence interval for the true mean profit:

900,000  $\pm$  2.262(100,000/10), that is, (828,469, 971,531).

- (b) The  $t$  distribution, since the true  $\sigma^2$  is not known.

**D.19**

- (a) 95% confidence interval for  $\sigma^2$ :

(19)(16)/32.8523  $\leq \sigma^2 \leq$  (19)(16)/8.9065, that is, (9.2535, 34.1324).

Note: For 19 d.f.,  $\chi_{0.025}^2 = 32.8523$  and  $\chi_{0.975}^2 = 8.9065$ .

- (b) Since the preceding interval does not include  $\sigma^2 = 8.2$ , reject the null hypothesis.

**D.20**

- (a) Note that

$$F = \frac{S_1^2}{S_2^2} = \frac{0.3762}{0.3359} = 1.12.$$

The  $p$ -value of obtaining an  $F$  value of 1.12 or greater is 0.4146. Therefore, one may not reject the null hypothesis that the two population variances are the same.

- (b) The  $F$  test is used. The basic assumption is that the two populations are normal.

**D.21**

$$E(X^*) = \frac{1}{n+1} [E(X_1) + \cdots + E(X_n)] = \frac{1}{n+1} (n\mu_X) = \frac{n}{n+1} \mu_X \neq \mu_X.$$