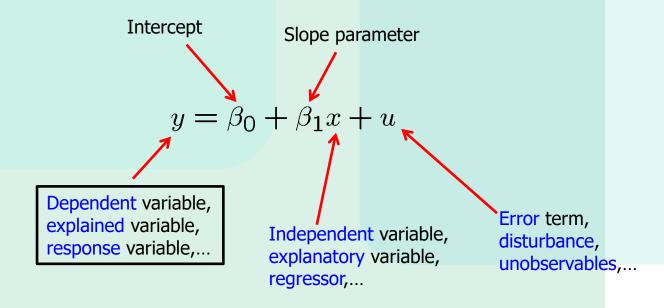


Definition of the simple linear regression model

"Explains variable y in terms of variable x''





Interpretation of the simple linear regression model

"Studies how y varies with changes in x:"

$$\frac{\Delta y}{\Delta x} = \beta_1$$

as long as

By how much does the dependent variable change if the independent variable is increased by one unit? holding other factors fixed (including other x, u)

$$\frac{\Delta u}{\Delta x} = 0$$

Interpretation only correct if all other things remain equal when the independent variable is increased by one unit

The simple linear regression model is rarely applicable in practice but its discussion is useful for pedagogical reasons

(methods suitable for teaching)

Example: Soybean yield and fertilizer

$$yield = \beta_0 + \beta_1 fertilizer + \omega \qquad \qquad \text{Rainfall,} \\ \text{land quality,} \\ \text{presence of parasites,} \dots \\ \text{Measures the effect of fertilizer on} \\ \text{yield, holding all other factors fixed}$$

Example: A simple wage equation

holding all other factors fixed

$$wage = \beta_0 + \beta_1 educ + \omega$$
 Labor force experience, tenure with current employer, work ethic, intelligence, ...

Measures the change in hourly wage given another year of education,



• When is there a causal interpretation?

correlation ≠ causality causality here → ceteris paribus

Conditional mean independence assumption

$$E(u|x) = 0$$
 The explanatory variable must not contain information about the mean of the unobserved factors

Example: wage equation

$$wage = \beta_0 + \beta_1 educ + \omega$$
 e.g. intelligence ...

The conditional mean independence assumption is unlikely to hold because individuals with more education will also be more intelligent on average.

- Population regression function (PFR)
 - The conditional mean independence assumption implies that

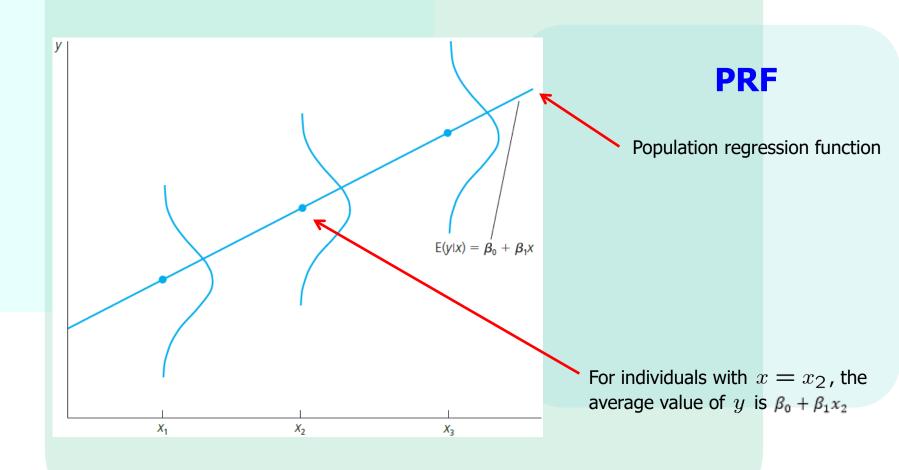
$$E(y|x) = E(\beta_0 + \beta_1 x + u|x)$$
$$= \beta_0 + \beta_1 x + E(u|x)$$
$$= \beta_0 + \beta_1 x$$

 This means that the average value of the dependent variable can be expressed as a linear function of the explanatory variable

linear regression

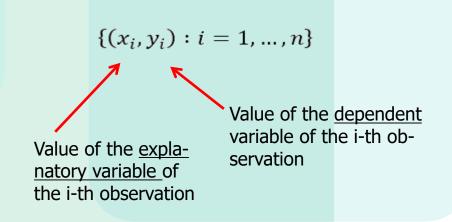
(not to be confused with linear correlation)





- Deriving the ordinary least squares (OLS) estimates
- In order to estimate the regression model one needs data
- A random sample of n observations

$$(x_1,y_1)$$
 First observation (x_2,y_2) Second observation (x_3,y_3) Third observation (x_n,y_n) n-th observation



- What does "as good as possible" mean?
- Regression residuals

$$\widehat{u}_i = y_i - \widehat{y}_i = y_i - \widehat{\beta}_0 - \widehat{\beta}_1 x_i$$

Minimize sum of squared regression residuals

$$\min \sum_{i=1}^{n} \widehat{u}_{i}^{2} \rightarrow \widehat{\beta}_{0}, \widehat{\beta}_{1}$$

Ordinary Least Squares (OLS) estimates

$$\widehat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}, \quad \widehat{\beta}_0 = \bar{y} - \widehat{\beta}_1 \bar{x}$$



Fit as good as possible a regression line through the data points:

For example, the i-th data point (x_i, y_i) $\hat{u}_i = \text{residual}$ \hat{y}_i = fitted value

Fitted regression line

Х

 X_i

CEO Salary and return on equity

$$salary = \beta_0 + \beta_1 roe + u$$

Salary in thousands of dollars

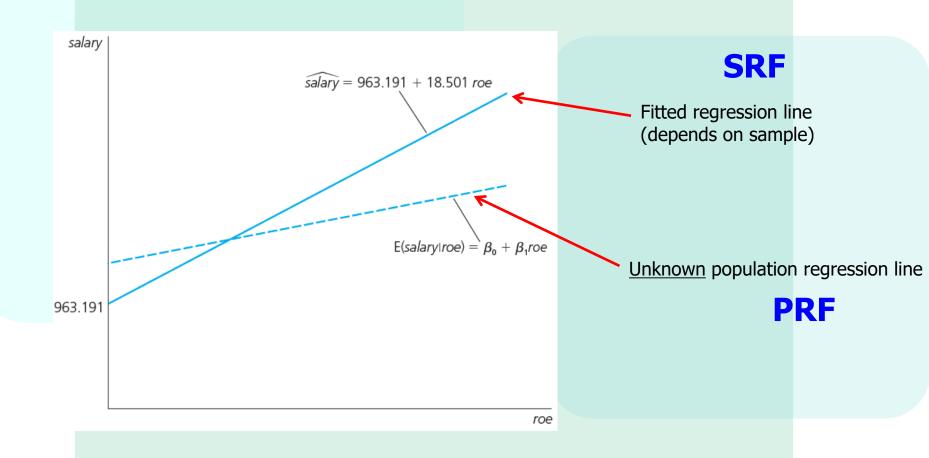
Average return on equity of the CEO's firm

Fitted regression

$$salary = 963.191 + 18.501 \ roe$$
Intercept

If the return on equity increases by 1 percent, then salary is predicted to change by \$18,501

Causal interpretation?



Wage and education

$$wage = \beta_0 + \beta_1 educ + u$$
 Hourly wage in dollars Years of education

Fitted regression (fitted to sample data)

$$\widehat{wage} = -0.90 + 0.54 \ educ$$
Intercept

In the sample, one more year of education was associated with an increase in hourly wage by \$0.54

Causal interpretation?

correlation ≠ causality causality here → ceteris paribus



Voting outcomes and campaign expenditures (two parties)

$$voteA = \beta_0 + \beta_1 shareA + u$$

Percentage of vote for candidate A

Percentage of campaign expenditures candidate A

Fitted regression

$$\widehat{voteA} = 26.81 + 0.464 \ shareA$$
Intercept

If candidate A's share of spending increases by one percentage point, he or she receives 0.464 percentage points more of the total vote

Causal interpretation?



- Properties of OLS on any sample of data
- Fitted values and residuals

$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i \qquad \hat{u}_i = y_i - \hat{y}_i$$

Fitted or predicted values

Deviations from regression line (= residuals)

Algebraic properties of OLS regression

$$\sum_{i=1}^{n} \hat{u}_i = 0$$

line sum up to zero

Deviations from regression

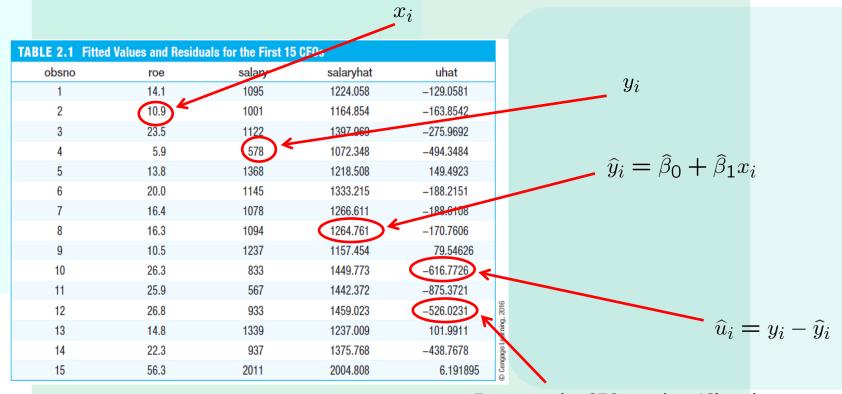
$$\sum_{i=1}^{n} x_i \hat{u}_i = 0$$

Covariance between deviations and regressors is zero

$$\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}$$

Sample averages of y and x lie on regression line





For example, CEO number 12's salary was \$526,023 lower than predicted using the the information on his firm's return on equity



(think: ρ-squared and R-squared) **Goodness-of-Fit**

"How well does the explanatory variable explain the dependent variable?"

Measures of Variation (SS = sum of squares)

$$SST \equiv \sum_{i=1}^{n} (y_i - \bar{y})^2$$
 $SSE \equiv \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2$ $SSR \equiv \sum_{i=1}^{n} \hat{u}_i^2$

Total sum of squares, represents total variation in the dependent variable

(Explained)

Explained sum of squares, represents variation explained by regression

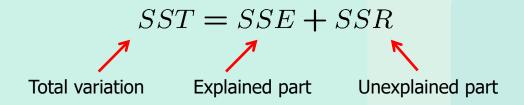
Residual sum of squares, represents variation not explained by regression

(Total)

(Unexplained)



Decomposition of total variation



Goodness-of-fit measure (R-squared)

$$R^2 \equiv \frac{SSE}{SST} = 1 - \frac{SSR}{SST}$$

R-squared measures the fraction of the total variation that is explained by the regression



CEO Salary and return on equity

$$salary = 963.191 + 18.501 \ roe$$

 $n = 209, \quad R^2 = 0.0132$

The regression explains only 1.3% of the total variation in salaries

Voting outcomes and campaign expenditures

$$voiteA = 26.81 + 0.464$$
 shareA
 $n = 173, R^2 = 0.856$

The regression explains 85.6% of the total variation in election outcomes

• Caution: A high R-squared does not necessarily mean that the regression has a causal interpretation! correlation ≠ causality

(相關不等於因果)



- Incorporating nonlinearities: Semi-logarithmic form
- Regression of log wages on years of education

$$\log(wage) = \beta_0 + \beta_1 educ + u$$
 Natural logarithm of wage

This changes the interpretation of the regression coefficient:

$$\beta_1 = \frac{\Delta log \ (wage)}{\Delta e duc} = \frac{1}{wage} \cdot \frac{\Delta wage}{\Delta e duc} = \frac{\Delta wage}{\Delta e duc}$$
 — Percentage change of wage ... if years of education are increased by one year

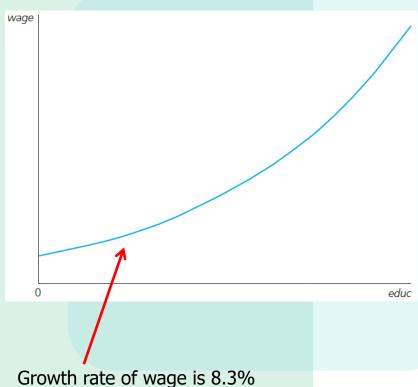
Fitted regression

$$\widehat{\log}(wage) = 0.584 + 0.083 \ educ$$

The wage increases by 8.3% for every additional year of education (= return to another year of education)

For example:

$$\frac{\Delta wage}{wage} = \frac{+0.83\$}{10\$} = 0.083 = +8.3\%$$



Growth rate of wage is 8.3% per year of education



- Incorporating nonlinearities: Log-logarithmic form
- CEO salary and firm sales

This changes the interpretation of the regression coefficient:

$$\beta_1 = \frac{\Delta log \, (salary)}{\Delta log \, (sales)} = \frac{\Delta salary}{\Delta sales} \qquad \qquad \frac{Percentage \, change \, of \, salary}{Logarithmic \, changes \, are} \\ \frac{Logarithmic \, changes \, are}{always \, percentage \, changes}$$



CEO salary and firm sales: fitted regression

$$\widehat{\log}(salary) = 4.822 + 0.257 \log(sales) + 1\% \text{ sales; } + 0.257\% \text{ salary}$$

For example:

$$\frac{\frac{\Delta salary}{salary}}{\frac{\Delta sales}{sales}} = \frac{\frac{+2.570\$}{1,000,000\$}}{\frac{+10.000,000\$}{1,000,000,000\$}} = \frac{+0.257\% \text{ salary}}{+1\% \text{ sales}} = 0.257$$

The log-log form postulates a constant elasticity model,
 whereas the semi-log form assumes a semi-elasticity model

- Expected values and variances of the OLS estimators (biasedness and efficiency)
- The estimated regression coefficients are random variables (why?)
 because they are calculated from a random sample

$$\widehat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}, \quad \widehat{\beta}_0 = \overline{y} - \widehat{\beta}_1 \overline{x}$$

Data is random and depends on particular sample that has been drawn

 The question is what the estimators will estimate on average and how large their variability in repeated samples is

$$E(\widehat{\beta}_0) = ?$$
, $E(\widehat{\beta}_1) = ?$ $Var(\widehat{\beta}_0) = ?$, $Var(\widehat{\beta}_1) = ?$



- Standard assumptions for the linear regression model
- Assumption SLR.1 (Linear in parameters)

$$y = \beta_0 + \beta_1 x + u$$
 In the population, the relationship between y and x is linear

Assumption SLR.2 (Random sampling)

$$\{(x_i,y_i): i=1,\ldots,n\}$$
 The data is a random sample drawn from the population

$$y_i = \beta_0 + \beta_1 x_i + u_i$$
 Each data point therefore follows the population equation



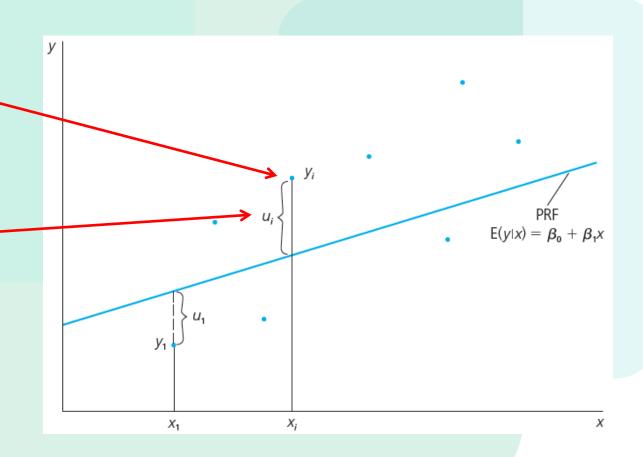
- Discussion of random sampling: Wage and education
 - The population consists, for example, of all workers of country A
 - In the population, a linear relationship between wages (or log wages) and years of education holds
 - Draw completely randomly a worker from the population
 - The wage and the years of education of the worker drawn are random because one does not know beforehand which worker is drawn
 - Throw back worker into population and repeat random draw n times
 - The wages and years of education of the sampled workers are used to estimate the linear relationship between wages and education



The values drawn for the i-th worker (x_i, y_i)

The implied deviation from the population relationship for the i-th worker:

$$u_i = y_i - \beta_0 - \beta_1 x_i$$





- Assumptions for the linear regression model (cont.)
- Assumption SLR.3 (Sample variation in the explanatory variable)

$$\sum_{i=1}^{n} (x_i - \bar{x})^2 > 0$$

The values of the explanatory variables are not all the same (otherwise it would be impossible to study how different values of the explanatory variable lead to different values of the dependent variable)

Assumption SLR.4 (Zero conditional mean)

$$E(u_i|x_i) = 0$$

The value of the explanatory variable must contain no information about the mean of the unobserved factors



Theorem 2.1 (Unbiasedness of OLS)

$$SLR.1-SLR.4 \Rightarrow E(\hat{\beta}_0) = \beta_0, E(\hat{\beta}_1) = \beta_1$$

- Interpretation of unbiasedness
 - The estimated coefficients may be smaller or larger, depending on the sample that is the result of a random draw
 - However, on average, they will be equal to the (true) values that charac-terize the true relationship between y and x in the population
 - "On average" means if sampling was repeated, i.e. if drawing the random sample and doing the estimation was repeated many times
 - In a given sample, estimates may differ considerably from true values



Variances of the OLS estimators

- Depending on the sample, the estimates will be nearer or farther away from the true population values
- How far can we expect our estimates to be away from the true population values on average (= sampling variability)?
- Sampling variability is measured by the estimator's variances

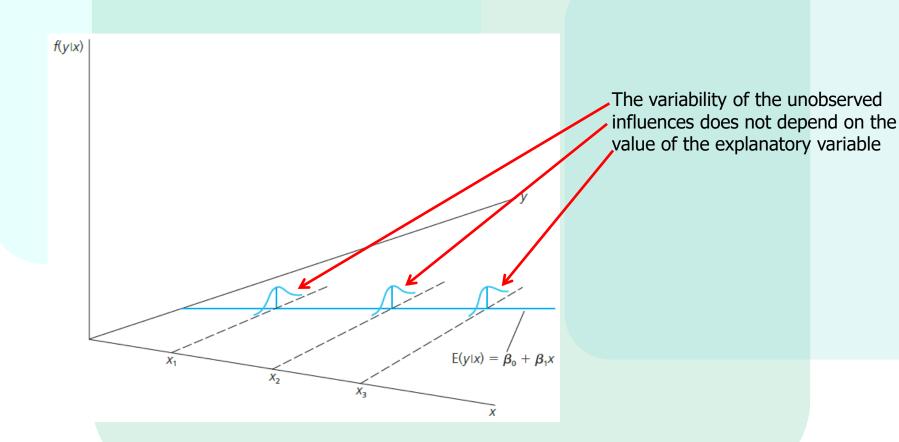
$$Var(\widehat{\beta}_0), \ Var(\widehat{\beta}_1)$$

Assumption SLR.5 (Homoskedasticity)

$$Var(u_i|x_i) = \sigma^2$$
 The value of the explanatory variable must contain no information about the variability of the unobserved factors

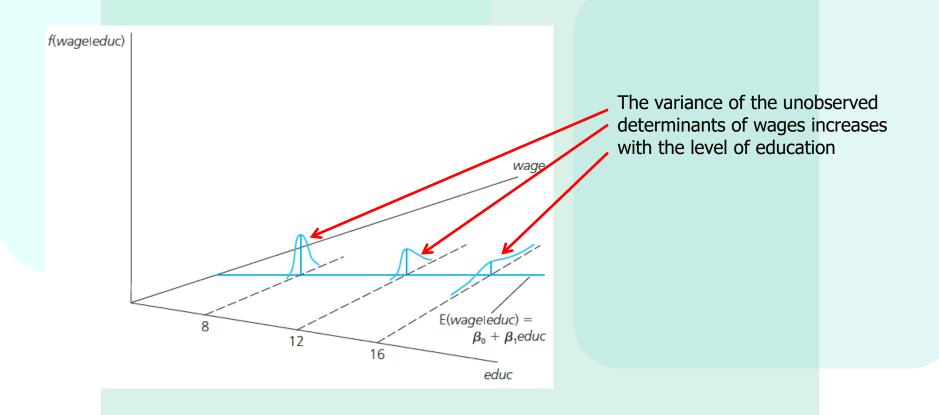


Graphical illustration of homoskedasticity





An example for heteroskedasticity: Wage and education





Theorem 2.2 (Variances of the OLS estimators)

Under assumptions SLR.1 – SLR.5:

$$Var(\widehat{\beta}_{1}) = \frac{\sigma^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} = \frac{\sigma^{2}}{SST_{x}}$$

$$Var(\widehat{\beta}_{0}) = \frac{\sigma^{2}n^{-1}\sum_{i=1}^{n} x_{i}^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} = \frac{\sigma^{2}n^{-1}\sum_{i=1}^{n} x_{i}^{2}}{SST_{x}}$$

Conclusion:

 The sampling variability of the estimated regression coefficients will be the higher, the larger the variability of the unobserved factors, and the lower, the higher the variation in the explanatory variable



Estimating the error variance

$$Var(u_i|x_i) = \sigma^2 = Var(u_i)$$
 The variance of u does not depend on x, i.e. equal to the unconditional variance

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (\hat{u}_i - \bar{\bar{u}}_i)^2 = \frac{1}{n} \sum_{i=1}^n \hat{u}_i^2$$

One could estimate the variance of the errors by calculating the variance of the residuals in the sample; unfortunately this estimate would be biased

$$\widehat{\sigma}^2 = \frac{1}{n-2} \sum_{i=1}^n \widehat{u}_i^2$$
 An unbiased estimate of the error variance can be obtained by substracting the number of estimated regression coefficients from the number of observations



Theorem 2.3 (Unbiasedness of the error variance)

$$SLR.1 - SLR.5 \Rightarrow E(\hat{\sigma}^2) = \sigma^2$$

Calculation of standard errors (s.e.) for regression coefficients

$$se(\hat{\beta}_1) = \sqrt{\widehat{Var}(\hat{\beta}_1)} = \sqrt{\widehat{\sigma}^2/SST_x}$$
 Plug in $\hat{\sigma}^2$ for the unknown σ^2
$$se(\hat{\beta}_0) = \sqrt{\widehat{Var}(\hat{\beta}_0)} = \sqrt{\widehat{\sigma}^2/SST_x}$$

The estimated standard deviations of the regression coefficients are called "standard errors." They measure how precisely the regression coefficients are estimated.