MATU9D2: PRACTICAL STATISTICS: FORMULA SHEET

1. One Sample Summary Statistics

$$\bar{x} = \frac{\sum x_i}{n}$$

$$s^2 = \frac{\sum x_i^2 - \frac{(\sum x_i)^2}{n}}{n-1}$$

2. Two Sample Summary Statistics

$$\bar{x}_{1} = \frac{\sum x_{1i}}{n_{1}} \qquad s_{1}^{2} = \frac{\sum x_{1i}^{2} - \frac{(\sum x_{1i})^{2}}{n_{1}}}{n_{1} - 1}$$

$$\bar{x}_{2} = \frac{\sum x_{2i}}{n_{2}} \qquad s_{2}^{2} = \frac{\sum x_{2i}^{2} - \frac{(\sum x_{2i})^{2}}{n_{2}}}{n_{2} - 1}$$

$$s_{p}^{2} = \frac{(n_{1} - 1)s_{1}^{2} + (n_{2} - 1)s_{2}^{2}}{n_{1} + n_{2} - 2}$$

3. Correlation and Regression

$$S_{XY} = \sum x_i y_i - \frac{\sum x_i \sum y_i}{n}$$

$$S_{XX} = \sum x_i^2 - \frac{\left(\sum x_i\right)^2}{n}$$

$$S_{YY} = \sum y_i^2 - \frac{\left(\sum y_i\right)^2}{n}$$

$$r = \frac{S_{XY}}{\sqrt{S_{XX} S_{YY}}}$$

$$R^2 = \frac{S_{XY}^2}{S_{XX} S_{YY}}$$

The least squares line is estimated using

$$\hat{\beta} = \frac{S_{XY}}{S_{YY}}, \qquad \hat{\alpha} = \bar{y} - \hat{\beta}\bar{x}$$

4. Relationships

The Chi-squared test has test statistic

$$X^{2} = \sum_{a \parallel c dk} \frac{\left(O_{j} - E_{j}\right)^{2}}{E_{j}}$$

Using a significance level of 0.05, this is compared with $\chi^2(df; 0.05)$, where df= (r-1)(c-1).

The test of the Null hypothesis H_0 : $\rho=0$ against the alternative H_1 : $\rho\neq 0$ has test Statistic

$$t = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}}$$

Using significance level of 0.05, this is compared with \pm t(n-2; 0.025).

5. Probability

A random variable X following a Binomial probability model with n trials and probability of success θ has probability distribution

$$\Pr[X = x] = \frac{n!}{x!(n-x)!} \theta^x (1-\theta)^{n-x}, \quad x = 0, \dots, n.$$

and mean and variance given by

$$E(X) = n\theta$$
, $Var(X) = n\theta(1-\theta)$.

If X follows a Normal probability model with mean μ and variance σ^2 , then $Z = \frac{(X - \mu)}{\sigma}$ follows a N(0,1) distribution.

6. One Sample Confidence Intervals

An interval estimate for a population *proportion*, having approximate confidence 0.95, is

$$\hat{\theta} \pm 1.96 \times \sqrt{\frac{\hat{\theta}(1-\hat{\theta})}{n}}$$

An interval estimate for a population mean, having confidence 0.95, is

$$\bar{x} \pm t(n-1;0.025) \times \frac{s}{\sqrt{n}}$$

An interval estimate for a population variance, having confidence 0.95, is

$$\left(\frac{(n-1)s^2}{\chi^2(n-1;0.025)},\frac{(n-1)s^2}{\chi^2(n-1;0.975)}\right).$$

7. Two Sample Tests

The test statistics for the test of the Null Hypothesis H_0 : $\mu_1 = \mu_2$ against the alternative H_1 : $\mu_1 \neq \mu_2$ is

$$t = \frac{\overline{x}_1 - \overline{x}_2}{\sqrt{s_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

For a test having significance level 0.05, we compare this with \pm t($n_1 + n_2 - 2$; 0.025).

The test statistic for the test of the Null Hypothesis $H_0: \sigma_1^2 = \sigma_2^2$ against the alternative $H_1: \sigma_1^2 \neq \sigma_2^2$ is

$$\frac{s_1^2}{s_2^2} \left(\frac{larger}{smaller} \right)$$

For a test having significance level 0.05, we compare this with

$$F(n_1-1,n_2-1;0.025)$$

8. Inference in Simple Linear Regression

The residual variance σ^2 is estimated by

$$\hat{\sigma}^2 = \frac{S_{YY} - \frac{S_{XY}^2}{S_{XX}}}{n - 2}$$

$$\hat{\alpha} = \bar{y} - \hat{\beta}\bar{x}$$

$$\hat{\beta} = \frac{S_{XY}}{S_{YY}}$$

A test of the Null Hypothesis $H_0: \beta = 0$ against the alternative $H_1: \beta \neq 0$ has the test statistic

$$T = \frac{\hat{\beta}}{\sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}}$$

Using a significance level of 0.05, this is compared with \pm t(n -2; 0.025).

<u>Parameter</u>	Estimate	Estimated Standard Error
α	\hat{a}	$\sqrt{\hat{\sigma}^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{S_{XX}} \right)}$
β	\hat{eta}	$\sqrt{rac{\hat{\sigma}^{2}}{S_{_{XX}}}}$
$\alpha + \beta x$	$\hat{\alpha} + \hat{\beta}x$	$\sqrt{\hat{\sigma}^2 \left(\frac{1}{n} + \frac{\left(x - \overline{x} \right)^2}{S_{XX}} \right)}$

In each case interval estimates, having confidence 0.95, are given by

Estimate
$$\pm t(n-2;0.025) \times Estimated Standard Error$$

The formula for a 95% prediction interval for an individual observation is

$$\hat{\alpha} + \hat{\beta}x \pm t(n-2;0.025)\sqrt{\hat{\sigma}^2\left(1 + \frac{1}{n} + \frac{(x-\bar{x})^2}{S_{xx}}\right)}$$

9. Tests for more than Two Samples

The test statistic for the null hypothesis H_o : $\mu_1 = \mu_2 = = \mu_I$ against the alternative H_1 : Not all of the μ_I are equal; when there are J observations in each group is

$$F = \frac{SS_{BG}/(I-1)}{SS_{WG}/(IJ-I)}$$

where
$$SS_{BG} = J \sum_{i=1}^{I} (\bar{y}_{i.} - \bar{y}_{..})^{2}$$

$$SS_{WG} = SS_{TOT} - SS_{BG}$$

$$SS_{TOT} = \sum_{i=1}^{I} \sum_{j=1}^{J} y_{ij}^{2} - \frac{\left(\sum_{i=1}^{I} \sum_{j=1}^{J} y_{ij}\right)^{2}}{IJ}$$

For a test having significance level 0.05, we compare this with F(I-1, IJ-I; 0.05).