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# Homework 1

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## 1 MATHEMATICS BASICS

### 1.1 Optimization

Use the Lagrange multiplier method to solve the following problem:

$$\begin{aligned} \min_{x_1, x_2} \quad & x_1^2 + x_2^2 - 1 \\ \text{s.t.} \quad & x_1 + x_2 - 1 = 0 \\ & x_1 - 2x_2 \geq 0 \end{aligned} \tag{1.1}$$

*Proof.* 该问题的对偶问题为:

$$\begin{aligned} \max \quad & \theta(w, v) \\ \text{s.t.} \quad & w \geq 0 \end{aligned} \tag{1.2}$$

其中

$$\theta(w, v) = \inf \{x_1^2 + x_2^2 - 1 - w(x_1 - 2x_2) - v(x_1 + x_2 - 1)\} \tag{1.3}$$

$$= \inf \left\{ \left(x_1 - \frac{2+v}{2}\right)^2 + \left(x_2 + \frac{2w-v}{2}\right)^2 - \frac{(w+v)^2}{4} + \left(w - \frac{v}{2}\right)^2 + v - 1 \right\} \tag{1.4}$$

$$= -\frac{(w+v)^2}{4} + \left(w - \frac{v}{2}\right)^2 + v - 1 \tag{1.5}$$

得到

$$\begin{aligned} \max \quad & -\frac{(w+v)^2}{4} + \left(w - \frac{v}{2}\right)^2 + v - 1 \\ \text{s.t.} \quad & w \geq 0 \end{aligned} \tag{1.6}$$

并且有

$$x_1 = \frac{2+v}{2} \tag{1.7}$$

$$x_2 = -\frac{2w-v}{2} \tag{1.8}$$

计算对目标变量的偏导数

$$\frac{\partial \theta}{\partial w} = \frac{3}{2}w - \frac{3}{2} = 0 \quad (1.9)$$

$$\frac{\partial \theta}{\partial v} = -\frac{3}{2}v - 1 = 0 \quad (1.10)$$

解得

$$w = \frac{2}{3} \quad (1.11)$$

$$v = \frac{2}{3} \quad (1.12)$$

代入方程 1.7 得到

$$x_1 = \frac{2}{3} \quad (1.13)$$

$$x_2 = \frac{1}{3} \quad (1.14)$$

最优值为  $-\frac{4}{9}$

□

## 1.2 Calculus

The gamma function is defined by (assuming  $x > 0$ )

$$\Gamma(x) = \int_0^{\infty} u^{x-1} e^{-u} du \quad (1.15)$$

(1) Prove that  $\Gamma(x+1) = x\Gamma(x)$ .

*Proof.*

$$x\Gamma(x) = \int_0^{\infty} xu^{x-1} e^{-u} du \quad (1.16)$$

$$= \int_0^{\infty} \frac{d}{du}(u^x) e^{-u} du \quad (1.17)$$

$$= u^x e^{-u} \Big|_0^{\infty} - \int_0^{\infty} u^x (-e^{-u}) du \quad (1.18)$$

$$= \int_0^{\infty} u^x e^{-u} du \quad (1.19)$$

$$= \Gamma(x+1) \quad (1.20)$$

□

(2) Also show that

$$\int_0^1 \mu^{a-1} (1-\mu)^{b-1} d\mu = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} \quad (1.21)$$

*Proof.*

$$\Gamma(a)\Gamma(b) = \int_0^{\infty} x^{a-1} e^{-x} dx \int_0^{\infty} y^{b-1} e^{-y} dy \quad (1.22)$$

$$= \int_0^{\infty} \int_0^{\infty} x^{a-1} e^{-x} y^{b-1} e^{-y} dx dy \quad (1.23)$$

取  $x + y = z$  有

$$\Gamma(a)\Gamma(b) = \int_0^\infty x^{a-1} \left( \int_x^\infty (z-x)^{b-1} e^{-z} dz \right) dx \quad (1.24)$$

交换积分顺序有

$$\Gamma(a)\Gamma(b) = \int_0^\infty e^{-z} \left( \int_0^z x^{a-1} (z-x)^{b-1} dx \right) dz \quad (1.25)$$

再换元  $x = zu$  有

$$\Gamma(a)\Gamma(b) = \int_0^\infty e^{-z} \left( \int_0^1 (zu)^{a-1} (z-zu)^{b-1} z du \right) dz \quad (1.26)$$

$$= \int_0^\infty e^{-z} z^{a+b-1} dz \int_0^1 u^{a-1} (1-u)^{b-1} du \quad (1.27)$$

$$= \Gamma(a+b) \int_0^1 u^{a-1} (1-u)^{b-1} du \quad (1.28)$$

即有

$$\int_0^1 \mu^{a-1} (1-\mu)^{b-1} d\mu = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} \quad (1.29)$$

□

### 1.3 Probability

Suppose  $\lambda \sim \Gamma(p|\alpha, \beta)$  and  $x|\lambda \sim \text{Poisson}(x|\lambda)$ . Show that  $\lambda|x \sim \Gamma(\lambda|\alpha + x, \beta + 1)$ , which implies that the Gamma distribution can serve as a conjugate prior to the Poisson distribution.

*Proof.* 首先有

$$f(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda} \quad (1.30)$$

$$f(x|\lambda) = \frac{\lambda^x}{x!} \cdot e^{-\lambda} \quad (1.31)$$

则其联合分布密度函数为

$$f(x, \lambda) = f(x|\lambda)f(\lambda) = \frac{\lambda^x}{x!} e^{-\lambda} \cdot \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda} = \frac{\lambda^{x+\alpha-1} \beta^\alpha}{x! \Gamma(\alpha)} e^{-\lambda(\beta+1)} \quad (1.32)$$

则可以求得  $x$  的边缘分布密度函数为

$$f(x) = \int_0^\infty f(x, \lambda) d\lambda = \frac{\Gamma(x+\alpha) \beta^\alpha}{x! (\beta+1)^{x+\alpha} \Gamma(\alpha)} \quad (1.33)$$

根据贝叶斯公式有

$$f(\lambda|x) = \frac{f(x, \lambda)f(x)}{= \frac{\frac{\lambda^{x+\alpha-1} \beta^\alpha}{x! \Gamma(\alpha)} e^{-\lambda(\beta+1)}}{\frac{\Gamma(x+\alpha) \beta^\alpha}{x! (\beta+1)^{x+\alpha} \Gamma(\alpha)}}} = \frac{(\beta+1)^{\alpha+x}}{\Gamma(\alpha+x)} \lambda^{x+\alpha-1} e^{-\lambda(\beta+1)} \quad (1.34)$$

可以看出  $\lambda$  的后验分布满足  $\lambda|x \sim Ga(\alpha + x, \beta + 1)$

即泊松分布式伽马分布的共轭先验分布

□

## 1.4 Stochastic Process

*Proof.* 设定两个事件，即  $A_k$  代表出现序列  $H, \underbrace{T, T, \dots, T}_k$  的平均投掷次数, 而  $B_k$  代表在出现 H 后出现序列  $H, \underbrace{T, T, \dots, T}_k$  的平均投掷次数

显然，考虑第一次投掷有两种情况：

- 结果为H，那么接下来即为  $B_n$
- 结果为T，那么接下来即为  $A_n$

有

$$A_n = 1 + \frac{1}{2}A_n + \frac{1}{2}B_n \quad (1.35)$$

推得

$$A_n = B_n + 2 \quad (1.36)$$

在序列  $B_{n-1}$  后再次进行一次投掷，有两种结果：

- 如果是T，则结束为  $B_{n-1} + 1$
- 如果是H，则重新开始为  $B_{n-1} + 1 + B_n$

接下来得到

$$B_n = \frac{1}{2}(B_{n-1} + 1) + \frac{1}{2}(B_{n-1} + B_n + 1) \quad (1.37)$$

得到递推公式

$$B_n = 2B_{n-1} + 2 \quad (1.38)$$

设  $f(n) = B_n + 2$  则由  $f(0) = B_0 + 2 = 2$  有

$$f(n) = 2^{n+1} \quad (1.39)$$

从而

$$B_n = 2^{n+1} - 2 \quad (1.40)$$

从而得到

$$A_k = B_k + 2 = 2^{k+1} \quad (1.41)$$

其观察到这样的模式的所需要的实验的平均次数为  $2^{k+1}$  □

## 2 SVM

### 2.1 From Primal to Dual

Consider the regression problem with training data  $\{(\mathbf{x}_i, y_i)\}_{i=1}^N$  ( $\mathbf{x}_i \in \mathbb{R}^d, y_i \in \mathbb{R}$ ).  $\epsilon > 0$  debotes a fixed small value. Derive the dual problem of the following primal problem of linear SVM (Please use  $\alpha_i, \hat{\alpha}_i, \beta_i$ , and  $\hat{\beta}_i$  as the Lagrange multipliers):

$$\min_{\mathbf{w}, b, \xi, \hat{\xi}} \frac{1}{2} \|\mathbf{w}\|^2 + C \sum_{i=1}^N (\xi_i + \hat{\xi}_i) \quad (2.1)$$

s.t.

$$y_i \leq \mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i, i = 1, \dots, N \quad (2.2)$$

$$y_i \geq \mathbf{w}^\top \mathbf{x}_i + b - \epsilon - \hat{\xi}_i, i = 1, \dots, N \quad (2.3)$$

$$\xi_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.4)$$

$$\hat{\xi}_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.5)$$

*Proof.* 首先将问题变为标准形式:

$$\min_{\mathbf{w}, b, \xi, \hat{\xi}} \frac{1}{2} \|\mathbf{w}\|^2 + C \sum_{i=1}^N (\xi_i + \hat{\xi}_i) \quad (2.6)$$

s.t.

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i \geq 0, \quad i = 1, \dots, N \quad (2.7)$$

$$y_i - \mathbf{w}^\top \mathbf{x}_i - b + \epsilon + \hat{\xi}_i \geq 0, \quad i = 1, \dots, N \quad (2.8)$$

$$\xi_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.9)$$

$$\hat{\xi}_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.10)$$

问题的对偶问题为

$$\max \quad \theta(\boldsymbol{\alpha}, \hat{\boldsymbol{\alpha}}, \boldsymbol{\beta}, \hat{\boldsymbol{\beta}}) \quad (2.11)$$

s.t.

$$\boldsymbol{\alpha}, \hat{\boldsymbol{\alpha}}, \boldsymbol{\beta}, \hat{\boldsymbol{\beta}} \geq \mathbf{0} \quad (2.12)$$

其中

$$\theta = \inf \left\{ \frac{1}{2} \|\mathbf{w}\|^2 + C \sum_{i=1}^N (\xi_i + \hat{\xi}_i) - \sum_{i=1}^N \alpha_i (\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i) - \right. \quad (2.13)$$

$$\left. \sum_{i=1}^N \hat{\alpha}_i (y_i - \mathbf{w}^\top \mathbf{x}_i - b + \epsilon + \hat{\xi}_i) - \beta_i \xi_i - \hat{\beta}_i \hat{\xi}_i \right\} = \inf \{F\} \quad (2.14)$$

上式中 $F$ 应该对 $\mathbf{w}, b, \xi, \hat{\xi}$ 求最小值, 即偏导数分别为0:

$$\frac{\partial F}{\partial \mathbf{w}} = \mathbf{w} - \sum_i^N \alpha_i \mathbf{x}_i + \sum_i^N \hat{\alpha}_i \mathbf{x}_i = \mathbf{0} \quad (2.15)$$

$$\frac{\partial F}{\partial b} = -\sum_i^N \alpha_i + \sum_i^N \hat{\alpha}_i = 0 \quad (2.16)$$

$$\frac{\partial F}{\partial \xi_i} = C - \alpha_i - \beta_i = 0 \quad (2.17)$$

$$\frac{\partial F}{\partial \hat{\xi}_i} = C - \hat{\alpha}_i - \hat{\beta}_i = 0 \quad (2.18)$$

则有

$$\mathbf{w} = \sum_i^N \alpha_i \mathbf{x}_i - \sum_i^N \hat{\alpha}_i \mathbf{x}_i \quad (2.19)$$

$$\sum_i^N \alpha_i = \sum_i^N \hat{\alpha}_i \quad (2.20)$$

$$C = \alpha_i + \beta_i \quad (2.21)$$

$$C = \hat{\alpha}_i + \hat{\beta}_i \quad (2.22)$$

代入 $F$ 的表达式得到:

$$\theta = \frac{1}{2} \mathbf{x}^T (\boldsymbol{\alpha} - \hat{\boldsymbol{\alpha}}) (\boldsymbol{\alpha} - \hat{\boldsymbol{\alpha}})^T \mathbf{x} - 2\epsilon \boldsymbol{\alpha}_i^T \mathbf{1} \quad (2.23)$$

即为所求的对偶问题

□

## 2.2 Finding Support Vectors (Optional)

State the corresponding set of KKT conditions. Now please argue from the KKT conditions why the following hold:

$$\alpha_i = 0 \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i \geq 0 \quad (2.24)$$

$$0 < \alpha_i < C \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i = 0 \quad (2.25)$$

$$\alpha_i = C \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i \leq 0 \quad (2.26)$$

$$\hat{\alpha}_i = 0 \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i \leq 0 \quad (2.27)$$

$$0 < \hat{\alpha}_i < C \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i = 0 \quad (2.28)$$

$$\hat{\alpha}_i = C \Rightarrow \mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i \geq 0 \quad (2.29)$$

Please show that either  $\alpha_i$  or  $\hat{\alpha}_i$  (or both) must be zero. What are the support vectors?

*Proof.* 对于 KKT 条件中的导数条件和广义Lagrange函数的导数条件相同即:

$$\mathbf{w} - \sum_i^N \alpha_i \mathbf{x}_i + \sum_i^N \hat{\alpha}_i \mathbf{x}_i = 0 \quad (2.30)$$

$$-\sum_i^N \alpha_i + \sum_i^N \hat{\alpha}_i = 0 \quad (2.31)$$

$$C - \alpha_i - \beta_i = 0 \quad (2.32)$$

$$C - \hat{\alpha}_i - \hat{\beta}_i = 0 \quad (2.33)$$

下面给出 KKT 条件中的互补松弛条件：

$$\alpha_i (\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i) = 0, \quad \forall i = 1, \dots, N \quad (2.34)$$

$$\hat{\alpha}_i (y_i - \mathbf{w}^\top \mathbf{x}_i - b + \epsilon + \hat{\xi}_i) = 0, \quad \forall i = 1, \dots, N \quad (2.35)$$

$$\beta_i \xi_i = 0 \quad \forall i = 1, \dots, N \quad (2.36)$$

$$\hat{\beta}_i \hat{\xi}_i = 0 \quad \forall i = 1, \dots, N \quad (2.37)$$

$$\alpha_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.38)$$

$$\hat{\alpha}_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.39)$$

$$\beta_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.40)$$

$$\hat{\beta}_i \geq 0 \quad \forall i = 1, \dots, N \quad (2.41)$$

如果  $\alpha_i = 0$ ，由互补松弛条件有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i \geq 0 \quad (2.42)$$

而上面的 Lagrange 函数的导数为零可以得到

$$C = \alpha_i + \beta_i \quad (2.43)$$

则有

$$\beta_i = C > 0 \quad (2.44)$$

由于互补松弛条件有

$$\xi_i = 0 \quad (2.45)$$

推得

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i \geq 0 \quad (2.46)$$

如果  $0 < \alpha_i < C$ ，由互补松弛条件有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i = 0 \quad (2.47)$$

另外同样由  $C = \alpha_i + \beta_i$  可以得到

$$\beta_i > 0 \quad (2.48)$$

则

$$\xi_i = 0 \quad (2.49)$$

推得

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i = 0 \quad (2.50)$$

如果  $\alpha_i = C$ ，由互补松弛条件有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i = 0 \quad (2.51)$$

即

$$\mathbf{v}^\top \mathbf{x}_i + b + \epsilon - y_i = -\xi_i \quad (2.52)$$

再由  $\alpha_i + \beta_i = C$  可以得到

$$\beta_i = 0 \quad (2.53)$$

有

$$\xi_i \geq 0 \quad (2.54)$$

推得

$$\mathbf{v}^\top \mathbf{x}_i + b + \epsilon - y_i = -\xi_i \leq 0 \quad (2.55)$$

对于  $\hat{\alpha}$  的情况可以完全类似的讨论

如果  $\alpha_i$  和  $\hat{\alpha}_i$  中没有一个等于0，分下面四种情况讨论

- $0 < \alpha_i < C$  和  $0 < \hat{\alpha}_i < C$

此时有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i = 0 \quad (2.56)$$

$$\mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i = 0 \quad (2.57)$$

两式相减得到

$$\epsilon = 0 \quad (2.58)$$

和

$$\epsilon > 0 \quad (2.59)$$

矛盾

- $\alpha_i = C$  和  $\hat{\alpha}_i = C$

有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i \leq 0 \quad (2.60)$$

$$\mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i \geq 0 \quad (2.61)$$

两个不等式相减得到

$$\epsilon \leq 0 \quad (2.62)$$

和

$$\epsilon > 0 \quad (2.63)$$

矛盾

- $\alpha_i = C$  和  $0 < \hat{\alpha}_i < C$

有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i \leq 0 \quad (2.64)$$

$$\mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i = 0 \quad (2.65)$$



推得

$$\epsilon \leq 0 \quad (2.66)$$

和

$$\epsilon > 0 \quad (2.67)$$

矛盾

- $0 < \alpha_i < C$  和  $\hat{\alpha}_i = C$

有

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon - y_i = 0 \quad (2.68)$$

$$\mathbf{w}^\top \mathbf{x}_i + b - \epsilon - y_i \geq 0 \quad (2.69)$$

有

$$\epsilon \leq 0 \quad (2.70)$$

和

$$\epsilon > 0 \quad (2.71)$$

矛盾

综上,  $\alpha_i$  和  $\hat{\alpha}_i$  中至少有一个为0  
满足方程

$$\mathbf{w}^\top \mathbf{x}_i + b + \epsilon + \xi_i - y_i = 0 \quad (2.72)$$

$$y_i - \mathbf{w}^\top \mathbf{x}_i - b + \epsilon + \hat{\xi}_i = 0 \quad (2.73)$$

的点称为支持向量。

□

### 3 IRLS FOR LOGISTIC REGRESSION

#### 3.1 推导

*Proof.* 该迭代方法基于 Newton-Raphson 迭代, 其对权重的更新形式为:

$$\mathbf{w}^{(\text{new})} = \mathbf{w}^{(\text{old})} - \mathbf{H}^{-1} \nabla E(\mathbf{w}) \quad (3.1)$$

对于对数似然函数:

$$\mathcal{L}(\mathbf{w}) = \sum_{i=1}^N (y_i \mathbf{w}^\top \mathbf{x}_i - \log(1 + \exp(\mathbf{w}^\top \mathbf{x}_i))) \quad (3.2)$$

对于 logictic 函数的导数:

$$\frac{\partial \text{sig}(\Phi_n)}{\partial \mathbf{w}} = \text{sig}(\Phi_n)(1 - \Phi_n)\Phi_n \quad (3.3)$$

梯度为

$$\nabla E(w) = \sum_{n=1}^N (y_n - t_n) \phi_n = \phi^T (Y - T) \quad (3.4)$$

其中

$$Y = \text{sig}(\Phi W) \quad (3.5)$$

从而

$$H = \nabla \nabla E(w) = \nabla \phi^T (Y - T) = \phi^T \nabla Y = \phi^T \sum_{n=1}^N y_n (1 - y_n) \phi_n = \phi^T \sum_{n=1}^N y_n (1 - y_n) \phi \quad (3.6)$$

其中  $\Phi$  是  $M \times N$  维的数据矩阵，也就是样本的个数为  $M$ ，维度为  $N$ ， $T$  则是  $M \times 1$  的特征的标签矩阵。

用  $R_{nn} = y_n(1 - y_n)$  表示  $\sum_{n=1}^N y_n(1 - y_n)$ ，有

$$w^{new} = w^{old} - H^{-1} \nabla E(w) \quad (3.7)$$

$$= w^{old} - (\phi^T R_{nn} \phi)^{-1} \phi^T (Y - T) \quad (3.8)$$

$$= (\phi^T R_{nn} \phi)^{-1} \phi^T R_{nn} [\phi w^{old} - R_{nn}^{-1} (Y - T)] \quad (3.9)$$

如果加上正则项：

$$L_2 = -\frac{\lambda}{2} \|w\|_2^2 \quad (3.10)$$

由

$$\nabla L_2 = -\lambda w \quad (3.11)$$

$$H_{L_2} = \nabla \nabla L_2 = -\lambda \quad (3.12)$$

得到其更新公式为

$$w^{new} = w^{old} - H^{-1} \nabla E(w) \quad (3.13)$$

$$= w^{old} - (\phi^T R_{nn} \phi - \lambda I_n)^{-1} [\phi^T (Y - T) - \lambda w_{old} I] \quad (3.14)$$

$$= (\phi^T R_{nn} \phi - \lambda I_n)^{-1} [(\phi^T R_{nn} \phi - \lambda I) w_{old} - \phi^T (Y - T) + \lambda w_{old} I] \quad (3.15)$$

$$= (\phi^T R_{nn} \phi + \lambda I_n)^{-1} \phi^T R_{nn} [\phi w^{old} - R_{nn}^{-1} (Y - T)] \quad (3.16)$$

□

### 3.2 程序编写和结果

(1)

该问题的程序见 **LR-IRLS.py**. 其结果为在训练集上的精度为0.84816,在测试集上的预测精度为0.84817，其迭代一次即收敛。

其

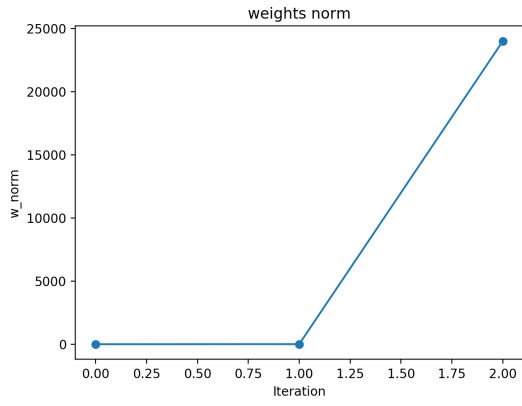


Figure 3.1: Weights norm in iteration

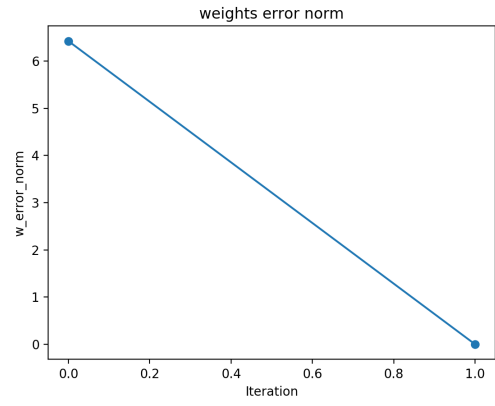


Figure 3.2: Weights error norm in iteration

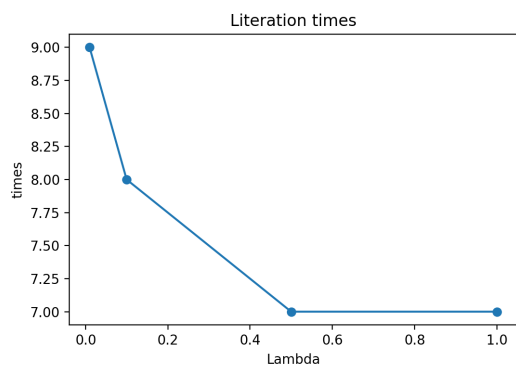


Figure 3.3: Iteration times

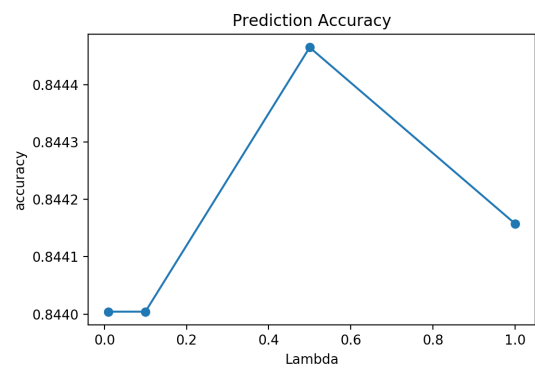


Figure 3.4: Prediction error in different  $\lambda$

(2)

该问题的程序见 **LR-IRLS-L2.py**. 通过比较得出的  $\lambda$  的较优值为 0.5,各个不同的  $\lambda$  的取值分别为[0.01,0.1,0.5,1], 其在不同参数下的迭代次数和预测精度见图3.2可以看出较好的  $\lambda$  的取值为 0.5,其在训练集上的精度为0.8445, 在测试集上的精度为0.8501

对于  $\lambda = 0.5$ 的情况, 其迭代过程中参数  $w$  的范数和更新的  $w$  和之前的  $w$  的差的范数见图3.2

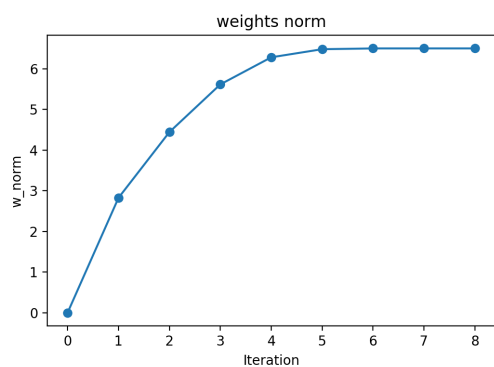


Figure 3.5: weights norm

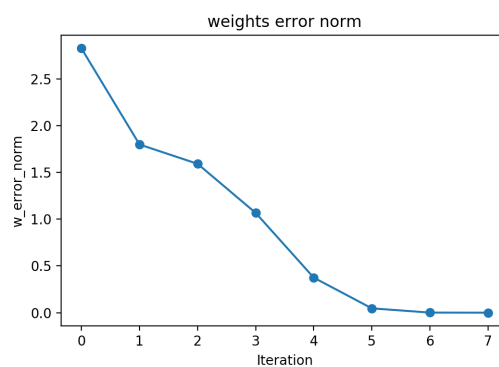


Figure 3.6: weights error norm

### 3.3 (3)

最后利用 **SKlearn** 的结果和前面的精度进行一个参照, 源代码见**LR-SKlearn.py**. 在训练集上的精度为 0.8492, 在测试集上的精度为 0.8498.

### 3.4 代码和文档

代码和文档见 [IRLS-Logistic-Regression](#)